

AID-521 Mathematics for Data Science

Module: Statistics | Lecture: 5

HYPOTHESIS TESTING & STATISTICAL DECISIONS

hypotheses, power, p-value, testing for a parameter, testing two samples, tests for count data, **interval estimation**

Hypothesis Testing -- Introduction

- Make decisions about the population on the basis of sample information (Statistical Decisions)
- Begin with initial conjectures about the population (Statistical Hypotheses)
- Compare conjecture with sample observations in a probabilistic manner (Tests of significance / Rules of decision)

Elements of a Statistical Test

- Null hypothesis H_0
 - Usually, the nullification of a claim.
- Alternative hypothesis H_1 or H_a
 - The claim itself.
- Test statistic TS
 - Function of the sample measurements used for the statistical decision – to reject H_0 or not.
 - Known distribution under H_0 .
- Rejection region (or critical region) RR
 - Values of the observed TS for which H_0 will be rejected.
 - Such values are usually extreme values of TS , or in other words, highly unlikely values of TS

Usual Alternative Hypotheses

One may have hypotheses such as

$$H_0 : \mu = \mu_0$$

against one of the following alternatives:

→ a two-tailed test/alternative

$$H_1 : \mu \neq \mu_0$$

→ a one-tailed test

$$\begin{cases} H_1 : \mu < \mu_0, & \text{a lower (or left) tailed alternative} \\ H_1 : \mu > \mu_0, & \text{an upper (or right) tailed alternative} \end{cases}$$

The Test Statistic

- A function of random sample (data), hence is a r.v.
 - Usually, an estimator for the unknown parameter
- Its prob. distribution is known under null hypothesis H_0
 - Assume population $\sim \mathcal{N}(\mu, \sigma^2 = \sigma_0^2)$
 - Consider a **simple hypothesis**[†] $H_0 : \mu = \mu_0$ vs. $H_1 : \mu \neq \mu_0$
 - Then $Z(\mu_0 | X_1, \dots, X_n) = \sqrt{n} \frac{\bar{X} - \mu_0}{\sigma_0}$ is a TS with known distribution $\mathcal{N}(0, 1)$
- Observed test statistic is its value when substituted with a given sample's values
 - $z(\mu_0 | x_1, \dots, x_n) = \sqrt{n} \frac{\bar{x} - \mu_0}{\sigma_0}$

[†] A hypothesis that uniquely specifies the distribution from which the sample is taken is called a simple hypothesis.

Interpretation of Statistical Decision

If evidence (sampled data) strongly contradicts H_0 (beyond a reasonable doubt), then we reject H_0 in favor of H_1 .

If H_0 is not rejected, then H_1 is automatically rejected.

Failure to reject H_0 does not necessarily mean that H_0 is true.[†]

[†]For e.g., “not guilty” does not mean a person “is innocent”. This basically means that there is not enough evidence to reject H_0 .

Errors in Statistical Decision

Statistical decision	True state of null hypothesis	
	H_0 true	H_0 false
Do not reject H_0	Correct decision	Type II error (β)
Reject H_0	Type I error (α)	Correct decision

Level of significance = $\alpha = P(\text{reject } H_0 \mid H_0 \text{ is true})$

$\beta = P(\text{don't reject } H_0 \mid H_0 \text{ is false})$

→ For fixed α , as n increases β decreases and vice versa.

Errors in Statistical Decision

- Consequences of different types of errors are, in general, very different.
 - H_0 : Person is innocent, vs. H_1 : Person is guilty
 - H_0 : Person is healthy, vs. H_1 : Person is sick
- In many situations it is possible to determine which of the two errors is more serious.
 - Choose null hypothesis such that its rejection should be considered to be more serious.

Rejection Regions in Statistical Decision

- Given the probability distribution of a TS under H_0 , the **rejection region** consists of those values of TS that are “extremely unlikely”.
- The statistical analyst decides what values of TS are “extreme”.
- The rejection region RR is pre-determined using the analyst's tolerance for error in decision.
 - Usually, the **level of significance** α is used to specify the level of error tolerance, and hence the RR
 - Each value of α corresponds to corresponding critical value(s) of TS

Sample Size

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Power

The **power of a test** is the probability that the test rejects H_0 when the alternative H_1 is true.

→ If $H_0 : \theta = \theta_0$, and $H_1 : \theta = \theta_0$, then the power of the test at some $\theta = \theta_1 \neq \theta_0$ is

$$\pi(\theta_1) = \text{Power}(\theta_1) = P(\text{reject } H_0 \mid \theta = \theta_1)$$

→ A good test will have high power.

Likelihood Ratio Tests

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p-Value

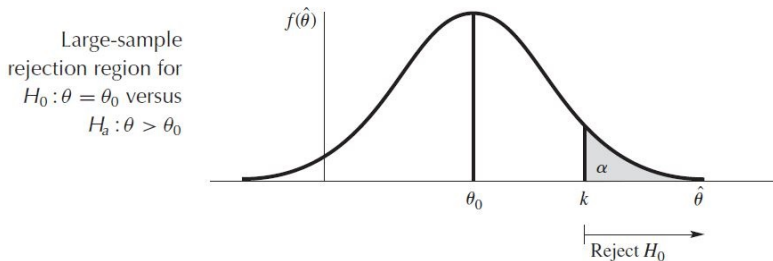
Corresponding to an observed value of a test statistic, the **p-value (or attained significance level)** is the lowest level of significance at which the null hypothesis would have been rejected.

- The maximum value of α , willing to tolerate, is chosen.
- If the p -value of the test is less than the maximum value of α , reject H_0 .

The lower the p-value, the stronger the evidence.

p-Value and Rejection Regions

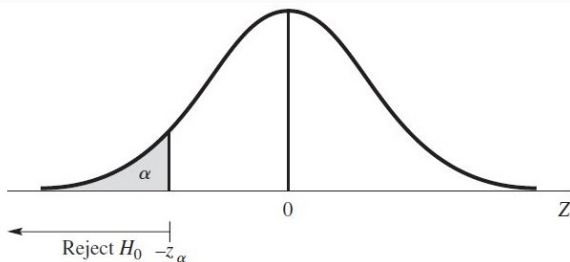
- Large sample
- One-tailed test (right/upper)



p-Value and Rejection Regions

- Large sample
- One-tailed test (left/lower)

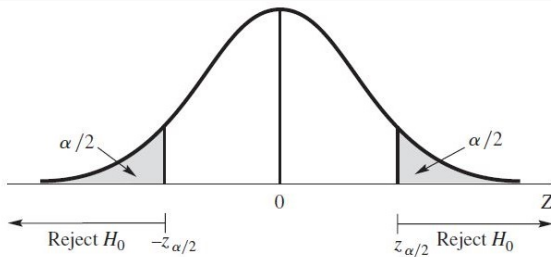
Rejection region for
testing $H_0: \theta = \theta_0$
versus $H_a: \theta < \theta_0$
based on $Z = \frac{\hat{\theta} - \theta_0}{\sigma_{\hat{\theta}}}$



p-Value and Rejection Regions

- Large sample
- Two-tailed test

Rejection region for
testing $H_0: \theta = \theta_0$
versus $H_a: \theta \neq \theta_0$,
based on $Z = \frac{\hat{\theta} - \theta_0}{\sigma_{\hat{\theta}}}$



Hypothesis Test for Parameter μ

SUMMARY OF HYPOTHESIS TESTS FOR μ

Large Sample ($n \geq 30$)

To test

$$H_0 : \mu = \mu_0$$

versus

$$\mu > \mu_0, \text{ upper tail test}$$

$$H_a : \mu < \mu_0, \text{ lower tail test}$$

$$\mu \neq \mu_0, \text{ two-tailed test}$$

$$\text{Test statistic: } Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$$

Replace σ by S , if σ is unknown.

$$\text{Rejection region: } \begin{cases} z > z_{\alpha}, & \text{upper tail RR} \\ z < -z_{\alpha}, & \text{lower tail RR} \\ |z| > z_{\alpha/2}, & \text{two tail RR} \end{cases}$$

Assumption: $n \geq 30$

Assumption: Random sample
comes from a normal
population

Decision: Reject H_0 , if the observed test statistic falls in the RR and conclude that H_a is true with $(1 - \alpha)100\%$ confidence. Otherwise, keep H_0 so that there is not enough evidence to conclude that H_a is true for the given α and more experiments may be needed.

Small Sample ($n < 30$)

To test

$$H_0 : \mu = \mu_0$$

versus

$$\mu > \mu_0, \text{ upper tail test}$$

$$H_a : \mu < \mu_0, \text{ lower tail test}$$

$$\mu \neq \mu_0, \text{ two-tailed test}$$

$$\text{Test statistic: } T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

$$\text{RR: } \begin{cases} t > t_{\alpha, n-1}, & \text{upper tail RR} \\ t < -t_{\alpha, n-1}, & \text{lower tail RR} \\ |t| > t_{\alpha/2, n-1}, & \text{two tail RR} \end{cases}$$

Hypothesis Test for Parameter σ^2

If X_1, \dots, X_n is a random sample from a normal population with the mean μ and variance σ^2 , then

$$\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma^2} = \frac{(n-1) S^2}{\sigma^2} \quad \text{has a chi-square distribution with } (n-1) \text{ degrees of freedom.}$$

We know from Theorem 4.2.7 that $(1/\sigma^2) \sum_{i=1}^n (X_i - \mu)^2$ has a chi-square distribution with n degrees of freedom. Thus,

$$\begin{aligned} \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 &= \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \bar{X} + \bar{X} - \mu)^2 \\ &= \frac{1}{\sigma^2} \left[\sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=1}^n (\bar{X} - \mu)^2 \right] \\ &\quad \left(\text{Since } 2 \sum_{i=1}^n (X_i - \bar{X})(\bar{X} - \mu) = 0 \right) \\ &= \frac{(n-1) S^2}{\sigma^2} + \left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \right)^2. \end{aligned}$$

The left-hand side of this equation has a chi-square distribution with n degrees of freedom. Also, since $(\bar{X} - \mu) / (\sigma/\sqrt{n}) \sim N(0, 1)$ by Theorem 4.2.6 we have $[(\bar{X} - \mu) / (\sigma/\sqrt{n})]^2 \sim \chi^2(1)$. Now from Theorem 4.2.4, $(n-1) S^2 / \sigma^2 \sim \chi^2(n-1)$.

Hypothesis Test for Parameter σ^2

SUMMARY OF HYPOTHESIS TEST FOR THE VARIANCE σ^2

To test

$$H_0 : \sigma^2 = \sigma_0^2$$

versus

$$\begin{aligned} \sigma^2 &> \sigma_0^2, && \text{upper tail test} \\ H_a : \sigma^2 &< \sigma_0^2, && \text{lower tail test} \\ \sigma^2 &\neq \sigma_0^2, && \text{two-tailed test.} \end{aligned}$$

Test statistic:

$$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2}$$

where S^2 is the sample variance.

Observed value of test statistic:

$$\frac{(n-1)s^2}{\sigma_0^2}$$

$$\text{Rejection region : } \begin{cases} \chi^2 > \chi_{\alpha, n-1}^2, & \text{upper tail RR} \\ \chi^2 < \chi_{1-\alpha, n-1}^2, & \text{lower tail RR} \\ \chi^2 > \chi_{\alpha/2, n-1}^2 \text{ or } \chi^2 < \chi_{1-\alpha/2, n-1}^2, & \text{two tail RR} \end{cases}$$

where $\chi_{\alpha, n-1}^2$ is such that the area under the chi-square distribution with $(n-1)$ degrees of freedom to its right is equal to α .

Assumption: Sample comes from a normal population.

Decision: Reject H_0 , if the observed test statistic falls in the RR and conclude that H_a is true with $(1 - \alpha)100\%$ confidence. Otherwise, do not reject H_0 because there is not enough evidence to conclude that H_a is true for given α and more data are needed.

Hypothesis Test for Two Samples

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