



# *Introduction to Statistical Machine Learning*

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## Outlines

- Overview
- Introduction
- Linear Algebra
- Probability
- Linear Regression 1
- Linear Regression 2
- Linear Classification 1
- Linear Classification 2
- Kernel Methods
- Sparse Kernel Methods
- Mixture Models and EM 1
- Mixture Models and EM 2
- Neural Networks 1
- Neural Networks 2
- Principal Component Analysis
- Autoencoders
- Graphical Models 1
- Graphical Models 2
- Graphical Models 3
- Sampling
- Sequential Data 1
- Sequential Data 2

(Many figures from C. M. Bishop, "Pattern Recognition and Machine Learning")



# Part V

## *Linear Regression 1*

*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*



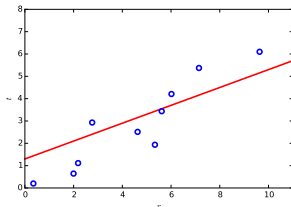
$$N = 10$$

$$\mathbf{x} \equiv (x_1, \dots, x_N)^T$$

$$\mathbf{t} \equiv (t_1, \dots, t_N)^T$$

$$x_i \in \mathbb{R} \quad i = 1, \dots, N$$

$$t_i \in \mathbb{R} \quad i = 1, \dots, N$$



- Predictor  $y(x, \mathbf{w})$ ?
- Performance measure?
- Optimal solution  $\mathbf{w}^*$ ?
- Recall: projection, inverse, eigenvalue decomposition

## Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



- Gaussian Distribution
- Bayes Rule
- Expected Loss
- Cross Validation

## Review

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*

# Linear Curve Fitting - Least Squares



## Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

$$N = 10$$

$$\mathbf{x} \equiv (x_1, \dots, x_N)^T$$

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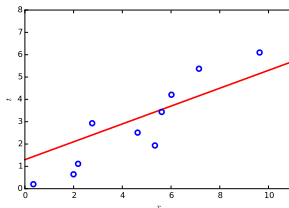
$$x_i \in \mathbb{R} \quad i = 1, \dots, N$$

$$t_i \in \mathbb{R} \quad i = 1, \dots, N$$

$$y(x, \mathbf{w}) = w_1 x + w_0$$

$$X \equiv [\mathbf{x} \quad \mathbf{1}]$$

$$\mathbf{w}^* = (X^T X)^{-1} X^T \mathbf{t}$$



We assume

$$t = \underbrace{y(\mathbf{x}, \mathbf{w})}_{\text{deterministic}} + \underbrace{\epsilon}_{\text{Gaussian noise}}$$



- uncertainty about the parameter  $\mathbf{w}$  captured in the prior probability  $p(\mathbf{w})$
- observed data  $\mathcal{D} = \{t_1, \dots, t_N\}$
- calculate the uncertainty in  $\mathbf{w}$  **after** the data  $\mathcal{D}$  have been observed

$$p(\mathbf{w} | \mathcal{D}) = \frac{p(\mathcal{D} | \mathbf{w})p(\mathbf{w})}{p(\mathcal{D})}$$

- $p(\mathcal{D} | \mathbf{w})$  as a function of  $\mathbf{w}$  : **likelihood function**
- likelihood expresses how probable the data are for different values of  $\mathbf{w}$
- **not** a probability function over  $\mathbf{w}$

## Review

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*



- Consider the linear regression problem, where we have random variables  $\mathbf{x}_n$  and  $t_n$ .
- We assume a conditional model  $t_n|\mathbf{x}_n$
- We propose a distribution, parameterized by  $\theta$

$$t_n|\mathbf{x}_n \sim \text{density}(\theta)$$

For a given  $\theta$  the density defines the probability of observing  $t_n|\mathbf{x}_n$ .

- We are interested in finding  $\theta$  that **maximises** the probability (called the **likelihood**) of the data.

# Likelihood Function - Frequentist versus Bayesian



Likelihood function  $p(\mathcal{D} | \mathbf{w})$

## Frequentist Approach

- $\mathbf{w}$  considered fixed parameter
- value defined by some 'estimator'
- error bars on the estimated  $\mathbf{w}$  obtained from the distribution of possible data sets  $\mathcal{D}$

## Bayesian Approach

- only one single data set  $\mathcal{D}$
- uncertainty in the parameters comes from a probability distribution over  $\mathbf{w}$

## Review

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*



# Frequentist Estimator - Maximum Likelihood



- choose  $\mathbf{w}$  for which the likelihood  $p(\mathcal{D} | \mathbf{w})$  is maximal
- choose  $\mathbf{w}$  for which the probability of the observed data is maximal
- Machine Learning: error function is negative log of likelihood function
- log is a monoton function
- maximising the likelihood  $\iff$  minimising the error
- Example: Fair-looking coin is tossed three times, always landing on heads.
- Maximum likelihood estimate of the probability of landing heads will give 1.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



- including prior knowledge easy (via prior  $\mathbf{w}$ )
- BUT: if prior is badly chosen, can lead to bad results
- subjective choice of prior
- sometimes choice of prior motivated by convenient mathematical form
- need to sum/integrate over the whole parameter space
- advances in sampling (Markov Chain Monte Carlo methods)
- advances in approximation schemes (Variational Bayes, Expectation Propagation)

*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

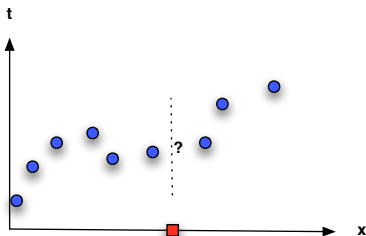
*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*



- Given a training data set of  $N$  observations  $\{\mathbf{x}_n\}$  and target values  $t_n$ .
- Goal : Learn to predict the value of one ore more target values  $t$  given a new value of the input  $\mathbf{x}$ .
- Example: Polynomial curve fitting (see Introduction).



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

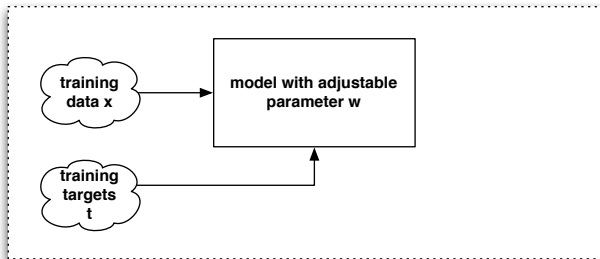
Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

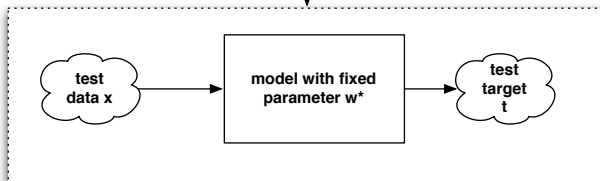
The Bias-Variance  
Decomposition

## Training Phase



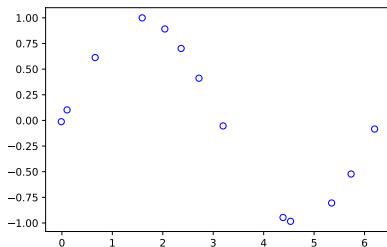
fix the most appropriate  $w^*$

## Test Phase



# Why Linear Regression?

- Analytic solution when using least squares loss
- Well understood statistical behaviour
- Efficient algorithms exist for convex losses and regularizers
- But what if the relationship is non-linear?



*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*



- **Linear** combination of **fixed** nonlinear basis functions

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$$

- parameter  $\mathbf{w} = (w_0, \dots, w_{M-1})^T$
- basis functions  $\boldsymbol{\phi}(\mathbf{x}) = (\phi_0(\mathbf{x}), \dots, \phi_{M-1}(\mathbf{x}))^T$
- convention  $\phi_0(\mathbf{x}) = 1$
- $w_0$  is the **bias parameter**

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

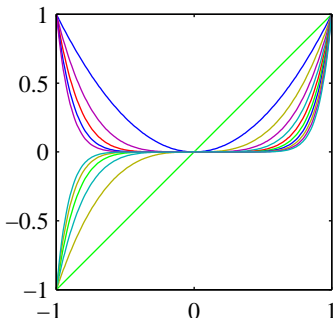
The Bias-Variance  
Decomposition

# Polynomial Basis Functions

- Scalar input variable  $x$

$$\phi_j(x) = x^j$$

- Limitation : Polynomials are global functions of the input variable  $x$ .
- Extension: Split the input space into regions and fit a different polynomial to each region (spline functions).

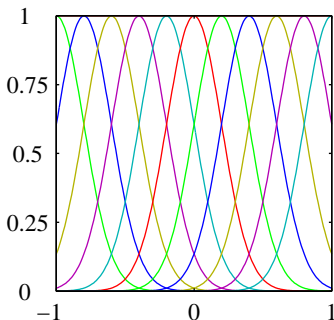


# 'Gaussian' Basis Functions

- Scalar input variable  $x$

$$\phi_j(x) = \exp \left\{ -\frac{(x - \mu_j)^2}{2s^2} \right\}$$

- Not a probability distribution.
- No normalisation required, taken care of by the model parameters  $w$ .





# Sigmoidal Basis Functions

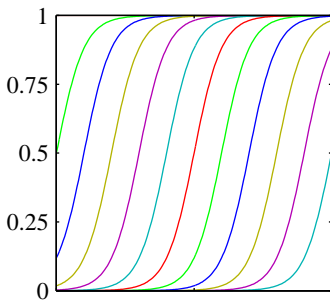
- Scalar input variable  $x$

$$\phi_j(x) = \sigma\left(\frac{x - \mu_j}{s}\right)$$

where  $\sigma(a)$  is the logistic sigmoid function defined by

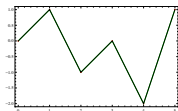
$$\sigma(a) = \frac{1}{1 + \exp(-a)}$$

- $\sigma(a)$  is related to the **hyperbolic tangent**  $\tanh(a)$  by  $\tanh(a) = 2\sigma(a) - 1$ .

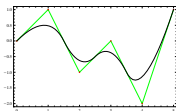




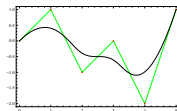
- Fourier Basis : each basis function represents a specific frequency and has infinite spatial extent.
- Wavelets : localised in both space and frequency (also mutually orthogonal to simplify application).
- Splines (piecewise polynomials restricted to regions of the input space; additional constraints where pieces meet, e.g. smoothness constraints  $\rightarrow$  conditions on the derivatives).



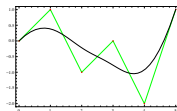
Linear  
Splines



Quadratic  
Splines



Cubic  
Splines



Quartic  
Splines

Approximate the points  
 $\{(0, 0), (1, 1), (2, -1), (3, 0), (4, -2), (5, 1)\}$  by different  
splines.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Maximum Likelihood and Least Squares



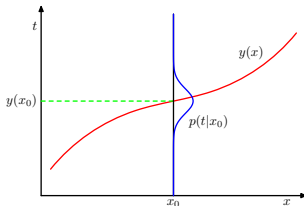
- No special assumption about the basis functions  $\phi_j(\mathbf{x})$ . In the simplest case, one can think of  $\phi_j(\mathbf{x}) = x_j$ , or  $\phi(\mathbf{x}) = \mathbf{x}$ .
- Assume target  $t$  is given by

$$t = \underbrace{y(\mathbf{x}, \mathbf{w})}_{\text{deterministic}} + \underbrace{\epsilon}_{\text{noise}}$$

where  $\epsilon$  is a **zero-mean Gaussian** random variable with **precision** (inverse variance)  $\beta$ .

- Thus

$$p(t | \mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Maximum Likelihood and Least Squares



- Likelihood of one target  $t$  given the data  $\mathbf{x}$  was

$$p(t | \mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

- Now, a **set of inputs**  $\mathbf{X}$  with corresponding target values  $\mathbf{t}$ .
- Assume data are **independent and identically distributed** (i.i.d.) (means : data are drawn independent and from the same distribution). The likelihood of the target  $\mathbf{t}$  is then

$$\begin{aligned} p(\mathbf{t} | \mathbf{X}, \mathbf{w}, \beta) &= \prod_{n=1}^N \mathcal{N}(t_n | y(\mathbf{x}_n, \mathbf{w}), \beta^{-1}) \\ &= \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \phi(\mathbf{x}_n), \beta^{-1}) \end{aligned}$$

- From now on drop the conditioning variable  $\mathbf{X}$  from the notation, as with supervised learning we do not seek to model the distribution of the input data.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Maximum Likelihood and Least Squares



- Consider the **logarithm of the likelihood**  $p(\mathbf{t} | \mathbf{w}, \beta)$  (the logarithm is a monotone function! )

$$\begin{aligned}\ln p(\mathbf{t} | \mathbf{w}, \beta) &= \sum_{n=1}^N \ln \mathcal{N}(t_n | \mathbf{w}^T \phi(\mathbf{x}_n), \beta^{-1}) \\ &= \sum_{n=1}^N \ln \left( \sqrt{\frac{\beta}{2\pi}} \exp \left\{ -\frac{\beta}{2} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2 \right\} \right) \\ &= \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta E_D(\mathbf{w})\end{aligned}$$

where the **sum-of-squares error function** is

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(x_n)\}^2.$$

- $\arg \max_{\mathbf{w}} \ln p(\mathbf{t} | \mathbf{w}, \beta) \rightarrow \arg \min_{\mathbf{w}} E_D(\mathbf{w})$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Maximum Likelihood and Least Squares



- Goal: Find a more compact representation.
- Rewrite the **error function**

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(x_n)\}^2 = \frac{1}{2} (\mathbf{t} - \Phi \mathbf{w})^T (\mathbf{t} - \Phi \mathbf{w})$$

where  $\mathbf{t} = (t_1, \dots, t_N)^T$ , and

$$\Phi = \begin{bmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \dots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \dots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \dots & \phi_{M-1}(\mathbf{x}_N) \end{bmatrix}$$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Maximum Likelihood and Least Squares



- The log likelihood is now

$$\begin{aligned}\ln p(\mathbf{t} | \mathbf{w}, \beta) &= \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta E_D(\mathbf{w}) \\ &= \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta \frac{1}{2} (\mathbf{t} - \Phi \mathbf{w})^T (\mathbf{t} - \Phi \mathbf{w})\end{aligned}$$

- Find **critical points** of  $\ln p(\mathbf{t} | \mathbf{w}, \beta)$ .
- The gradient with respect to  $\mathbf{w}$  is

$$\nabla_{\mathbf{w}} \ln p(\mathbf{t} | \mathbf{w}, \beta) = \beta \Phi^T (\mathbf{t} - \Phi \mathbf{w}).$$

Setting the gradient to zero gives

$$0 = \Phi^T \mathbf{t} - \Phi^T \Phi \mathbf{w},$$

- which results in

$$\mathbf{w}_{ML} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t} = \Phi^\dagger \mathbf{t}$$

where  $\Phi^\dagger$  is the **Moore-Penrose pseudo-inverse** of the matrix  $\Phi$ .

# Maximum Likelihood and Least Squares



- The log likelihood with the optimal  $\mathbf{w}_{ML}$  is now

$$\begin{aligned}\ln p(\mathbf{t} | \mathbf{w}_{ML}, \beta) \\ = \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta \frac{1}{2} (\mathbf{t} - \Phi \mathbf{w}_{ML})^T (\mathbf{t} - \Phi \mathbf{w}_{ML})\end{aligned}$$

- Find critical points of  $\ln p(\mathbf{t} | \mathbf{w}, \beta)$  wrt  $\beta$ ,

$$\frac{\partial \ln p(\mathbf{t} | \mathbf{w}_{ML}, \beta)}{\partial \beta} = 0$$

results in

$$\frac{1}{\beta_{ML}} = \frac{1}{N} (\mathbf{t} - \Phi \mathbf{w}_{ML})^T (\mathbf{t} - \Phi \mathbf{w}_{ML})$$

- Note: We can first find the maximum likelihood for  $\mathbf{w}$  as this does **not depend** on  $\beta$ . Then we can use  $\mathbf{w}_{ML}$  to find the maximum likelihood solution for  $\beta$ .
- Could we have chosen optimisation wrt  $\beta$  first, and then wrt to  $\mathbf{w}$  ?



# Sequential Learning - Stochastic Gradient Descent



- For **large data sets**, calculating the maximum likelihood parameters  $\mathbf{w}_{ML}$  and  $\beta_{ML}$  may be costly.
- For **online** applications, never all data in memory.
- Use a **sequential** algorithms (**online** algorithm).
- If the error function is a sum over data points  $E = \sum_n E_n$ , then
  - 1 initialise  $\mathbf{w}^{(0)}$  to some starting value
  - 2 update the parameter vector at iteration  $\tau + 1$  by

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} - \eta \nabla E_n,$$

where  $E_n$  is the error function after presenting the  $n$ th data set, and  $\eta$  is the **learning rate**.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Sequential Learning - Stochastic Gradient Descent



- For the sum-of-squares error function, stochastic gradient descent results in

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} + \eta \left( t_n - \mathbf{w}^{(\tau)T} \phi(\mathbf{x}_n) \right) \phi(\mathbf{x}_n)$$

- The value for the learning rate must be chosen carefully. A **too large** learning rate may prevent the algorithm from converging. A **too small** learning rate does follow the data too slowly.



- Add regularisation in order to prevent overfitting

$$E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$$

with regularisation coefficient  $\lambda$ .

- Simple quadratic regulariser

$$E_W(\mathbf{w}) = \frac{1}{2} \mathbf{w}^T \mathbf{w}$$

- Maximum likelihood solution

$$\mathbf{w} = (\lambda \mathbf{I} + \Phi^T \Phi)^{-1} \Phi^T \mathbf{t}$$

*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*

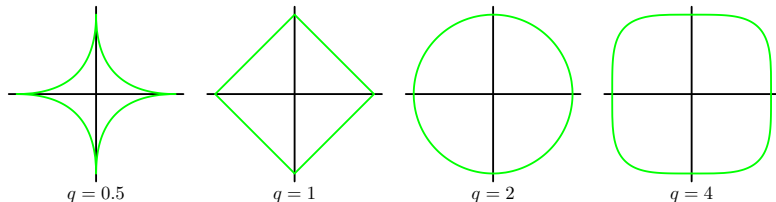
# Regularized Least Squares



- More general regulariser

$$E_W(\mathbf{w}) = \frac{1}{2} \sum_{j=1}^M |w_j|^q$$

- $q = 1$  (lasso) leads to a sparse model if  $\lambda$  large enough.



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

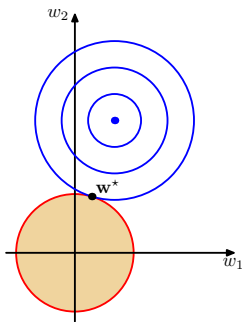
# Comparison of Quadratic and Lasso Regulariser



Assume a sufficiently large regularisation coefficient  $\lambda$ .

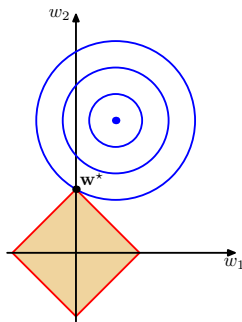
Quadratic regulariser

$$\frac{1}{2} \sum_{j=1}^M w_j^2$$



Lasso regulariser

$$\frac{1}{2} \sum_{j=1}^M |w_j|$$





- More than 1 target variable per data point.
- $\mathbf{y}$  becomes a vector instead of a scalar. Each dimension can be treated with a different set of basis functions (and that may be necessary if the data in the different target dimensions represent very different types of information.)
- Here we restrict ourselves to the SAME basis functions

$$\mathbf{y}(\mathbf{x}, \mathbf{w}) = \mathbf{W}^T \phi(\mathbf{x})$$

where  $\mathbf{y}$  is a  $K$ -dimensional column vector,  $\mathbf{W}$  is an  $M \times K$  matrix of model parameters, and

$\phi(\mathbf{x}) = (\phi_0(\mathbf{x}), \dots, \phi_{M-1}(\mathbf{x}), \phi_0(\mathbf{x}) = 1$ , as before.

- Define target matrix  $\mathbf{T}$  containing the target vector  $\mathbf{t}_n^T$  in the  $n^{th}$  row.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



- Suppose the conditional distribution of the target vector is an isotropic Gaussian of the form

$$p(\mathbf{t} | \mathbf{x}, \mathbf{W}, \beta) = \mathcal{N}(\mathbf{t} | \mathbf{W}^T \phi(\mathbf{x}), \beta^{-1} \mathbf{I}).$$

- The log likelihood is then

$$\begin{aligned} \ln p(\mathbf{T} | \mathbf{X}, \mathbf{W}, \beta) &= \sum_{n=1}^N \ln \mathcal{N}(\mathbf{t}_n | \mathbf{W}^T \phi(\mathbf{x}_n), \beta^{-1} \mathbf{I}) \\ &= \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) - \frac{\beta}{2} \sum_{n=1}^N \|\mathbf{t}_n - \mathbf{W}^T \phi(\mathbf{x}_n)\|^2 \end{aligned}$$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



- Maximisation with respect to  $\mathbf{W}$  results in

$$\mathbf{W}_{ML} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{T}.$$

- For each target variable  $\mathbf{t}_k$ , we get

$$\mathbf{w}_k = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t}_k = \Phi^\dagger \mathbf{t}_k.$$

- The solution between the different target variables decouples.
- Holds also for a general Gaussian noise distribution with arbitrary covariance matrix.
- Why?  $\mathbf{W}$  defines the mean of the Gaussian noise distribution. And the maximum likelihood solution for the mean of a multivariate Gaussian is independent of the covariance.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



# Loss Function for Regression



- Over-fitting results from a large number of basis functions and a relatively small training set.
- Regularisation can prevent overfitting, but how to find the correct value for the regularisation constant  $\lambda$  ?
- Frequentists viewpoint of the model complexity is the **bias-variance** trade-off.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Loss Function for Regression



- Choose an estimator  $y(\mathbf{x})$  to estimate the target value  $t$  for each input  $\mathbf{x}$ .
- Choose a loss function  $L(t, y(\mathbf{x}))$  which measures the difference between the target  $t$  and the estimate  $y(\mathbf{x})$ .
- The **expected loss** is then

$$\mathbb{E}[L] = \int \int L(t, y(\mathbf{x})) p(\mathbf{x}, t) \, d\mathbf{x} \, dt$$

- Common choice: **Squared Loss**

$$L(t, y(\mathbf{x})) = \{y(\mathbf{x}) - t\}^2.$$

- Expected loss for squared loss function

$$\mathbb{E}[L] = \int \int \{y(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt.$$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition



- Expected loss for squared loss function

$$\mathbb{E} [L] = \int \int \{y(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt.$$

- Minimise  $\mathbb{E} [L]$  by choosing the regression function

$$y(\mathbf{x}) = \frac{\int t p(\mathbf{x}, t) \, dt}{p(\mathbf{x})} = \int t p(t | \mathbf{x}) \, dt = \mathbb{E}_t [t | \mathbf{x}]$$

(use calculus of variations to derive this result ;  
alternatively work point-wise by fixing an  $\mathbf{x}$  and using  
stationarity to solve for  $y(\mathbf{x})$ ).

*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

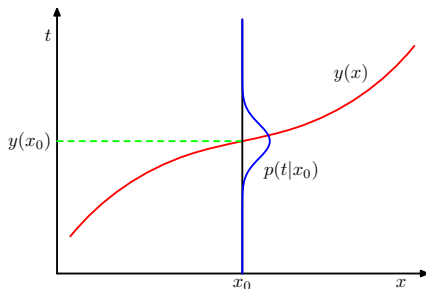
*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*

# Loss Function for Regression



- The regression function which minimises the expected squared loss, is given by the mean of the conditional distribution  $p(t | \mathbf{x})$ .



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Loss Function for Regression



- Analyse the expected loss

$$\mathbb{E} [L] = \int \int \{y(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt.$$

- Rewrite the squared loss

$$\begin{aligned}\{y(\mathbf{x}) - t\}^2 &= \{y(\mathbf{x}) - \mathbb{E} [t | \mathbf{x}] + \mathbb{E} [t | \mathbf{x}] - t\}^2 \\ &= \{y(\mathbf{x}) - \mathbb{E} [t | \mathbf{x}]\}^2 + \{\mathbb{E} [t | \mathbf{x}] - t\}^2 \\ &\quad + 2 \{y(\mathbf{x}) - \mathbb{E} [t | \mathbf{x}]\} \{\mathbb{E} [t | \mathbf{x}] - t\}\end{aligned}$$

- Claim

$$\int \int \{y(\mathbf{x}) - \mathbb{E} [t | \mathbf{x}]\} \{\mathbb{E} [t | \mathbf{x}] - t\} p(\mathbf{x}, t) \, d\mathbf{x} \, dt = 0.$$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Loss Function for Regression



- Claim

$$\int \int \{y(\mathbf{x}) - \mathbb{E}[t | \mathbf{x}]\} \{\mathbb{E}[t | \mathbf{x}] - t\} p(\mathbf{x}, t) \, d\mathbf{x} \, dt = 0.$$

- Separate functions depending on  $t$  from function depending on  $\mathbf{x}$

$$\int \{y(\mathbf{x}) - \mathbb{E}[t | \mathbf{x}]\} \left( \int \{\mathbb{E}[t | \mathbf{x}] - t\} p(\mathbf{x}, t) \, dt \right) \, d\mathbf{x}$$

- Calculate the integral over  $t$

$$\begin{aligned} \int \{\mathbb{E}[t | \mathbf{x}] - t\} p(\mathbf{x}, t) \, dt &= \mathbb{E}[t | \mathbf{x}] p(\mathbf{x}) - p(\mathbf{x}) \int \frac{t p(\mathbf{x}, t)}{p(\mathbf{x})} \, dt \\ &= \mathbb{E}[t | \mathbf{x}] p(\mathbf{x}) - p(\mathbf{x}) \mathbb{E}[t | \mathbf{x}] \\ &= 0 \end{aligned}$$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# Loss Function for Regression



- The expected loss is now

$$\mathbb{E} [L] = \int \{y(\mathbf{x}) - \mathbb{E} [t | \mathbf{x}]\}^2 p(\mathbf{x}) \, d\mathbf{x} + \int \text{var}[t | \mathbf{x}] p(\mathbf{x}) \, d\mathbf{x}$$

- Minimise first term by choosing appropriate  $y(\mathbf{x})$ .
- Second term represents the intrinsic variability of the target data (can be regarded as noise). Independent of the choice  $y(\mathbf{x})$ , can not be reduced by learning a better  $y(\mathbf{x})$ .

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



- Consider now the dependency on the data set  $\mathcal{D}$ .
- Prediction function now  $y(\mathbf{x}; \mathcal{D})$ .
- Consider again squared loss for which the optimal prediction is given by the conditional expectation  $h(\mathbf{x})$

$$h(\mathbf{x}) = \mathbb{E}[t | \mathbf{x}] = \int t p(t | \mathbf{x}) dt.$$

- BUT: we can not know  $h(x)$  exactly, as we would need an infinite number of training data to learn it accurately.
- Evaluate performance of algorithm by taking the expectation  $\mathbb{E}_{\mathcal{D}}[L]$  over all data sets  $\mathcal{D}$

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition





- Taking the expectation over all data sets  $\mathcal{D}$

$$\begin{aligned}\mathbb{E}_{\mathcal{D}} [\mathbb{E} [L]] &= \int \mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2] p(\mathbf{x}) d\mathbf{x} \\ &\quad + \int \int \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) d\mathbf{x} dt\end{aligned}$$

- Again, add and subtract the expectation  $\mathbb{E}_{\mathcal{D}} [y(\mathbf{x}; \mathcal{D})]$

$$\begin{aligned}\{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2 &= \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}} [y(\mathbf{x}; \mathcal{D})] \\ &\quad + \mathbb{E}_{\mathcal{D}} [y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2\end{aligned}$$

and show that the mixed term does vanish under the expectation  $\mathbb{E}_{\mathcal{D}} [\dots]$ .

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



- Expected loss  $\mathbb{E}_{\mathcal{D}} [L]$  over all data sets  $\mathcal{D}$

$$\text{expected loss} = (\text{bias})^2 + \text{variance} + \text{noise}.$$

where

$$(\text{bias})^2 = \int \{\mathbb{E}_{\mathcal{D}} [y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{variance} = \int \mathbb{E}_{\mathcal{D}} \left[ \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}} [y(\mathbf{x}; \mathcal{D})]\}^2 \right] p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{noise} = \int \int \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt.$$

- variance** : How sensitive is the model to small changes in the training set? (How much do solutions for individual data sets vary around their average ?)
- squared bias** : How accurate is a model across different training sets? (How much does the average prediction over all data sets differ from the desired regression function ?)

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

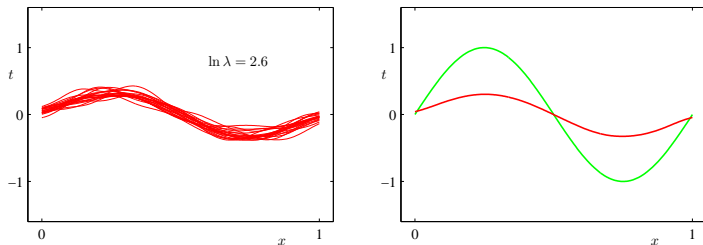
Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



Simple models have low variance and high bias.



Left: Result of fitting the model to 100 data sets, only 25 shown.  
Right: Average of the 100 fits in red, the sinusoidal function from where the data were created in green.

Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

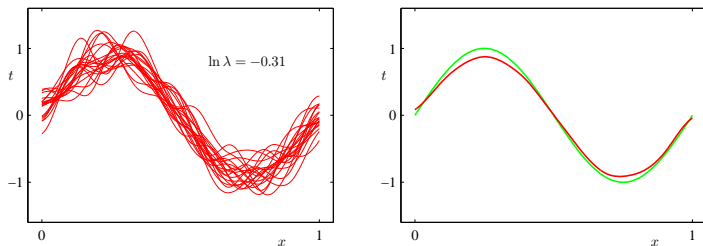
Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



## Dependence of bias and variance on the model complexity



Left: Result of fitting the model to 100 data sets, only 25 shown.  
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Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

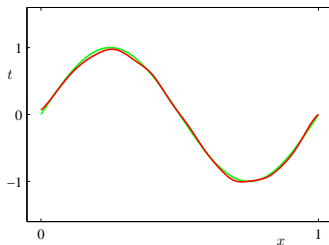
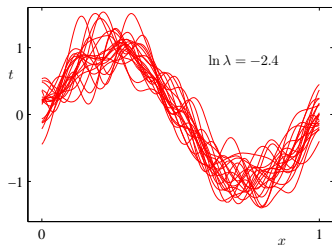
Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



Complex models have high variance and low bias.



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Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

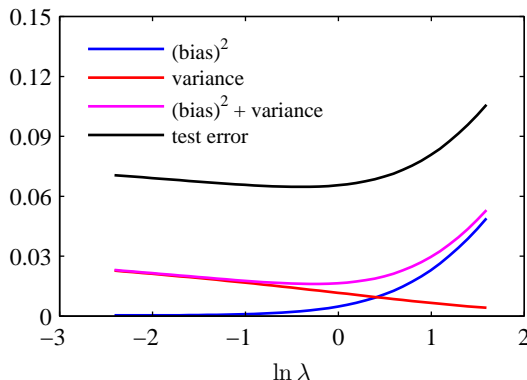
Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



- Squared bias, variance, their sum, and test data
- The minimum for  $(\text{bias})^2 + \text{variance}$  occurs close to the value that gives the minimum error



Review

Linear Basis Function  
Models

Maximum Likelihood and  
Least Squares

Sequential Learning

Regularized Least  
Squares

Multiple Outputs

Loss Function for  
Regression

The Bias-Variance  
Decomposition

# The Bias-Variance Decomposition



- Tradeoff between bias and variance
  - simple models have low variance and high bias
  - complex models have high variance and low bias
- The sum of bias and variance has a minimum at a certain model complexity.
- Expected loss  $\mathbb{E}_{\mathcal{D}} [L]$  over all data sets  $\mathcal{D}$

$$\text{expected loss} = (\text{bias})^2 + \text{variance} + \text{noise}.$$

- The noise comes from the data, and can not be removed from the expected loss.
- To analyse the bias-variance decomposition : many data sets needed, which are not always available.

*Review*

*Linear Basis Function  
Models*

*Maximum Likelihood and  
Least Squares*

*Sequential Learning*

*Regularized Least  
Squares*

*Multiple Outputs*

*Loss Function for  
Regression*

*The Bias-Variance  
Decomposition*