

# 1 General Sampling Methods

With the introduction of random number generators behind us, we assume the availability of an ideal sequence of random numbers. More precisely, we assume the availability of a sequence  $U_1, U_2, \dots$  of independent random variables, each satisfying,

$$P(U_i \leq u) = \begin{cases} 0, & u < 0 \\ u, & 0 \leq u \leq 1 \\ 1, & u > 1, \end{cases}$$

*i.e.*, uniformly distributed between 0 and 1. A typical simulation uses methods for transforming samples from the uniform distribution to samples from other distributions. The two most widely used general techniques are:

1. Inverse Transform Method.
2. Acceptance Rejection Method.

## 1.1 Inverse Transform Method

The inverse transform method is based on the following theorem.

**Theorem 1.** *Let  $F$  be a CDF. Define the quasi-inverse of  $F$  by*

$$F^{-1}(u) = \inf \{x \in \mathbb{R} : F(x) \geq u\} \quad \text{for } 0 < u < 1.$$

*Let  $U \sim U(0, 1)$  and  $X = F^{-1}(U)$ . Then, the CDF of  $X$  is  $F$ .*

Before going in to the proof of the theorem, let us first discuss the inverse transform method. Suppose that we want a sample from a cumulative distribution function  $F(x)$ , *i.e.*, we want to generate a random variable  $X$  with the property that  $P(X \leq x) = F(x)$  for all  $x \in \mathbb{R}$ . Using the Theorem 1, we have the following algorithm.

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### Algorithm 1 Inverse Transform Method

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|--|---------------------------------------|
| 1: Generate $U$ from $U(0, 1)$ distribution. | ▷ Using some random number generator. |
| 2: Set $X = F^{-1}(U)$ .                     |                                       |
| 3: Return $X$ .                              |                                       |
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In principle, we can use this algorithm for generation of all non-uniform random variables. However, there are computational aspects. We generally use this algorithm if  $F^{-1}$  is in closed form and easy to compute.

**Example 1** (Exponential Distribution). The exponential distribution with mean  $\theta$  has distribution

$$F(x) = 1 - e^{-x/\theta}, \quad x \geq 0.$$

Inverting yields

$$X = -\theta \log(1 - U).$$

This can be implemented also as  $X = -\theta \log(U)$ , since  $U$  and  $(1 - U)$  have the same distribution. ||

**Example 2** (Arc Sin Law). Consider the CDF

$$F(x) = \frac{2}{\pi} \arcsin \sqrt{x}, \quad 0 \leq x \leq 1.$$

The inverse transform method for sampling from this distribution is:

$$X = \sin^2 \left( \frac{U\pi}{2} \right) = \frac{1}{2} - \frac{1}{2} \cos(U\pi), \quad U \sim U(0, 1),$$

using the identity,  $2 \sin^2(t) = 1 - \cos(2t)$  for  $0 \leq t \leq \pi/2$ . ||

**Example 3** (Rayleigh Distribution). We consider the Rayleigh Distribution:

$$F(x) = 1 - e^{-2x(x-b)}, \quad x \geq b.$$

Solving the equation  $F(x) = u$ ,  $u \in (0, 1)$ , results in a quadratic with roots:

$$x = \frac{b}{2} \pm \frac{\sqrt{b^2 - 2 \log(1 - u)}}{2}.$$

The inverse is given by the larger of the two roots. Replacing  $U$  with  $(1 - U)$  we get,

$$X = \frac{b}{2} + \frac{\sqrt{b^2 - 2 \log(U)}}{2}. \quad ||$$

**Remark 1.** Note that even if the inverse of  $F$  is not known explicitly, the inverse transform method is still applicable through numerical evaluation of  $F^{-1}$ . Computing  $F^{-1}(u)$  is equivalent to finding a root  $x$  of the equation  $F(x) - u = 0$ . For a CDF  $F$  with PDF  $f$ , Newton's method for finding roots produces a sequence of iterates:

$$x_{n+1} = x_n - \frac{F(x_n) - u}{f(x_n)},$$

given a starting point  $x_0$ . †

Now, we will discuss the prove of the Theorem 1. To prove the above theorem, we need the following lemmas.

**Lemma 1.**  $F$  and  $F^{-1}$  both are non-decreasing.

Proof:  $F$  is a non-decreasing function (a properties of CDF). Now, we will prove that  $F^{-1}$  is a non-decreasing function. Let  $0 < u_1 < u_2 < 1$ . Then

$$\begin{aligned} \{x \in \mathbb{R} : F(x) \geq u_2\} &\subseteq \{x \in \mathbb{R} : F(x) \geq u_1\} \\ \implies \inf \{x \in \mathbb{R} : F(x) \geq u_2\} &\geq \inf \{x \in \mathbb{R} : F(x) \geq u_1\} \\ \implies F^{-1}(u_1) &\leq F^{-1}(u_2). \end{aligned}$$

□

**Lemma 2.**  $F F^{-1}(u) \geq u$  for all  $u \in (0, 1)$ .

Proof: As  $F$  is a right continuous function (a property of CDF), the infimum of the set

$$\{x \in \mathbb{R} : F(x) \geq u\}$$

belongs to the set. That means

$$F^{-1}(u) \in \{x \in \mathbb{R} : F(x) \geq u\}.$$

Therefore, we have the lemma.

□

**Lemma 3.**  $F^{-1} F(x) \leq x$  for all  $x \in \mathbb{R}$ .

Proof: Notice that

$$F^{-1} F(x) = \inf \{y \in \mathbb{R} : F(y) \geq F(x)\}.$$

Moreover,  $x \in \{y \in \mathbb{R} : F(y) \geq F(x)\}$ , which proves the lemma.

□

**Lemma 4.** For  $x \in \mathbb{R}$  and  $0 < u < 1$ ,  $F(x) \geq u$  if and only if  $F^{-1}(u) \leq x$ .

Proof: Suppose that  $F^{-1}(u) \leq x$ . Applying  $F$  on both sides, we get  $F(x) \geq F F^{-1}(u) \geq u$ . The first inequality is due to Lemma 1 and the last inequality is due to Lemma 2.

Now assume that  $F(x) \geq u$ . Then applying  $F^{-1}$  on the both sides,  $F^{-1}(u) \leq F^{-1} F(x) \leq x$ . The first inequality is due to Lemma 1 and the last inequality is due to Lemma 3.

□

Proof (of Theorem 1): Using the Lemma 4, proof of the Theorem 1 is very simple. Let us try to find the CDF of  $X$ . For  $x \in \mathbb{R}$ , the CDF of  $X$  is

$$P(X \leq x) = P(F^{-1}(U) \leq x) = P(U \leq F(x)) = F(x).$$

The first equality is due to the definition of  $X$ , the second is due to Lemma 4, and the last is due to the fact that  $U \sim U(0, 1)$ .

□

## 1.2 Discrete Distribution

Let us start with an example.

**Example 4** (Bernoulli Distribution). Suppose that we want to generate random number from Bernoulli distribution with probability of success  $p$ . Thus,  $P(X = 0) = 1 - p$  and  $P(X = 1) = p$ . The corresponding CDF is

$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1 - p & \text{if } 0 \leq x < 1 \\ 1 & \text{if } x \geq 1. \end{cases}$$

Therefore,

$$F^{-1}(u) = \begin{cases} 0 & \text{if } 0 < u < 1 - p \\ 1 & \text{if } 1 - p \leq u < 1. \end{cases}$$

Thus, generate  $U$  from  $U(0, 1)$ , and then return 0 if  $U < 1 - p$ , return 1 otherwise. ||

In the case of a discrete distribution with finite support, the evaluation of  $F^{-1}$  reduces to a table look up. Note that CDF of a discrete random variable is a step function. Consider, for example, a discrete random variable whose possible values are  $c_1 < c_2 < c_3 < \dots < c_N$ . Let  $p_i$  be the probability attached to  $c_i$ ,  $i = 1, 2, 3, \dots, N$  and set  $q_0 = 0$ . Also,

$$q_i = \sum_{j=1}^i p_j, \quad i = 1, 2, 3, \dots, N.$$

These are cumulative probabilities associated with the  $c_i$ , that is,  $q_i = F(c_i)$ ,  $i = 1, 2, \dots, N$ . To sample from this distribution, we can use the following algorithm.

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**Algorithm 2** Inversion Method for Discrete Random Variable with Finite Support

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- 1: Generate a uniform  $U \sim U(0, 1)$ .
  - 2: Find  $K \in \{1, 2, \dots, N\}$  such that  $q_{K-1} < U \leq q_K$ .
  - 3: Return  $c_K$ .
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If the discrete random variable takes countably infinite values, then table look up does not make sense. When there are infinitely many values, their description can only be a mathematical one. However, sometimes other transformation may help.

**Example 5** (Geometric Distribution). Suppose that we want to generate random number from Geometric distribution with success probability  $p$ . The PMF is given by

$$P(X = i) = p(1 - p)^i \quad \text{for } i = 0, 1, 2, \dots$$

Note that in this case the support is countably infinite. Let  $Y$  be an exponential random variable with mean  $\frac{1}{\lambda}$  and  $W = \lfloor Y \rfloor$ . Then it is easy to see that

$$P(W = i) = e^{-i\lambda} (1 - e^{-\lambda}) \quad \text{for } i = 0, 1, 2, \dots$$

Thus,  $W$  has a Geometric distribution with success probability  $1 - e^{-\lambda}$ . We can use this result to generate random number from a Geometric distribution using the following steps. Generate  $U$  from  $U(0, 1)$  distribution. Then, set  $X = \left\lfloor \frac{\ln U}{\ln(1-p)} \right\rfloor$ . ||

### 1.3 Conditional Distribution

Suppose  $X$  has distribution  $F$  and consider the problem of sampling  $X$  conditional on  $a < X \leq b$  with  $F(a) < F(b)$ . Note that the CDF of  $X$  conditional on  $a < X \leq b$  is given by

$$P(X \leq x | a < X \leq b) = \begin{cases} 0 & \text{if } x < a \\ \frac{P(a < X \leq x)}{P(a < X \leq b)} = \frac{F(x) - F(a)}{F(b) - F(a)} & \text{if } a < X \leq b \\ 1 & \text{if } x > b. \end{cases}$$

Now, using the inverse transform method, this is no more difficult than generating  $X$  unconditionally. We can follow the following algorithm for this purpose.

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**Algorithm 3** Generation from Conditional/Truncated Distribution

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- 1: Generate  $U$  from  $U(0, 1)$  distribution.
  - 2: Set  $X = F^{-1}(F(a) + (F(b) - F(a))U)$ .
  - 3: Return  $X$ .
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