

# An integrated integral projection model (IPM<sup>2</sup>) to disentangle size-structured harvest and natural mortality

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## 1 Abstract

1. Body size is one of the most important traits governing individual-level demographic rates and modulating population-level processes. Multiple size-dependent demographic rates can simultaneously change population structure, so distinguishing their individual contributions to overall population dynamics remains a challenge.

2. Disentangling size-dependent harvest rates from other demographic rates is critical for assessing the impact of removal on populations of invasive species. Inference about invasive populations can be difficult, however, as observations are often collected opportunistically as part of removal programs, rather than experimentally designed. Yet accurate inference is essential for understanding the feasibility of population suppression and optimizing management decisions.

3. We develop an integrated integral projection model (IPM<sup>2</sup>) that leverages the strengths of the integrated population model and integral projection model to enable inference about complex, size-structured demographic rates from imperfect observations. We apply the IPM<sup>2</sup> in the context of invasive European green crab (*Carcinus maenas*), a species for which individual body size strongly regulates both the observation-generating process and latent, population dynamics.

4. The IPM<sup>2</sup> facilitates the distinct estimation of green crab size-structured harvest and natural mortality rates, parameters for which no explicit data is collected and are unidentifiable in component datasets of the integrated population model. The model represents how the green crab population changes over time,

providing the first estimates of size-structured abundance of this high priority species.

5. By forecasting the stable size distribution and equilibrium population size under varying removal efforts, we demonstrate that extremely high levels of removal effort can reduce the equilibrium green crab population size. Yet these high mortality rates also shift the stable size distribution and increase the equilibrium abundance of smaller crabs, since size-selective removal alters intra-specific interactions. These results highlight the value of the IPM<sup>2</sup> framework for inferring complex population dynamics with information needs that outpace information in individual observational datasets, providing a path forward for accurate assessment of conservation programs.

**Key words:** integrated population model, integral projection model, state-space framework, invasive species, European green crab, size structure

## 2 Introduction

Assessing the efficacy of invasive species management interventions requires understanding if and how removal actions change population abundance and dynamics. However, effective population suppression can be challenging to differentiate from natural biological variation, reflecting the perennial challenge of distinguishing harvest and natural mortality (Aanes et al., 2007). Often variability in harvest mortality is confounded with variability in natural mortality (Lewy & Nielsen, 2003), yet separating these fluctuations is crucial for obtaining reliable estimates of population abundance and demographic rates (Walters & Martell, 2004). Invasive species removal program success is highly variable (Prior et al., 2018), so reliably quantifying harvest rates is essential for evaluating intervention success and efficiently allocating limited management resources (Green & Grosholz, 2021).

This challenge of disentangling sources of mortality is amplified for species with complex, size-structured demography. Body size can determine the strength of ecological interactions an individual experiences and can influence its key life-history processes like growth and mortality (De Roos et al., 2003). The ontogenetic scaling of ecological performance with body size ultimately drives patterns at the population level (Werner, 1994). Population dynamics can be shaped by competitive and predatory (cannibalistic) intra-specific interactions between different size cohorts (Claessen et al., 2004), predator-prey size-dependent functional responses (Aljetlawi et al., 2004), variable reproductive capacity across size (Hixon et al., 2014), and size-selective rates of harvest (Tu et al., 2018). Since these size-structured rates and interactions simultaneously modulate population dynamics, unraveling their individual contributions can be difficult.

To monitor invasive populations, ecologists are often limited to removal sampling data (Udell et al., 2022), which are typically collected opportunistically as part of removal programs, rather than experimentally designed (Crall et al., 2010; Rogosch & Olden, 2021; Tiberti et al., 2021). Consequently, these data fail to meet the strict assumptions of models commonly used to estimate abundance with animals removed successively from multiple sites for closed populations (Dorazio et al., 2005). These removal observations are often not collected systematically across time and space and can be associated with large measurement error that can mask biological signals (Auger-Méthé et al., 2016; Katsanevakis et al., 2012; Sibert et al., 2003).

Despite the unique statistical inference challenges associated with invasive species removal data, the costs of under- or over-estimating removal harvest rates can be high. Ineffective invasive population suppression plans waste human and economic resources and can degrade the confidence of people and stakeholders involved in removal (Tiberti et al., 2021). As invasive species management becomes more ambitious in scope and scale, population control can be controversial, stimulating conflicts among people and debates about achievability and efficiency (Crowley et al., 2017). These undesirable social outcomes underscore the need

for rigorous assessments of the effectiveness of control programs.

Combining existing classes of models can be useful for distinguishing process from observation dynamics, facilitating parameter identifiability, and enabling inference about complex, size-structured demographic rates from imperfect observations. First, an integral projection model can be used to make population projections based on vital rates that vary continuously as a function of body size (Merow et al., 2014; Rees et al., 2014). In contrast to matrix population models that can produce artifacts from coarse, arbitrary size class divisions, the integral projection model enables modeling of biological realism by using smooth relationships between individual size and demographic performance (Ellner & Rees, 2006).

To estimate demographic rates, the integral projection model can then be formulated in a state-space framework leveraging size-structured time series data (White et al., 2016). State-space approaches partition process and measurement error to estimate latent process model dynamics, providing a hierarchical structure that separately models 1) an unobserved, process time series that reflects the true, hidden population state and 2) an observation time series that contains imperfect measurements related to the process time series (Auger-Méthé et al., 2021; Newman et al., 2006). By incorporating both process and observation error, the state-space framework allows estimation of harvest mortality — the observation process — separately from natural biological variation.

State-space approaches, however, can suffer from challenges of parameter identifiability, especially in scenarios when information requirements of complex models exceed information available in a single dataset (Auger-Méthé et al., 2016; Knape, 2008). Integrated population models can resolve these identifiability issues by jointly analyzing multiple datasets in an integrated framework (Besbeas et al., 2002). By integrating survey and demographic data, these models increase the precision of parameter estimates and facilitate estimation of additional parameters not identifiable with individual datasets (Abadi et al., 2010; Riecke et al., 2019). By combining an integrated population model with an integral projection model into an integrated integral projection model (IPM<sup>2</sup>), the relationship between individual size and demographic performance can be used to forecast population-level processes (Plard et al., 2019).

An IPM<sup>2</sup> will be useful for understanding population dynamics and removal effectiveness of invasive European green crab (*Carcinus maenas*). The green crab’s population dynamics are highly non-linear and are regulated by complex, size-structured demography. Adult crabs exert direct control of recruitment, largely through strong negative adult-juvenile interactions like cannibalism (Grosholz et al., 2021; Romano & Zeng, 2017). Competition with native crabs is also size-structured, with larger crabs maintaining a competitive advantage for space and resources (Jensen et al., 2007; McDonald et al., 2001). Importantly, these individual, size-structured interactions modulate population-level processes; a field experiment in California, USA found a 30-fold, single-year increase in total green crab abundance in response to removal (Grosholz et al., 2021).

This high level of juvenile survival associated with low adult abundance, often referred to as overcompensation or the “hydra effect”, underscores the importance of quantifying size-structured population response to removal (Abrams, 2009).

The green crab observation process is also size-selective, further complicating distinction between simultaneous size-structured processes. Baited traps used for removal do not catch all individuals with equal probability across size classes, with the larval stage and smaller green crab size classes completely unobserved (Jørgensen et al., 2009). Additionally, the observability of the population changes within a trapping season as individuals grow into size classes capable of being caught in traps. Quantifying the relationship between removal effort and size-structured removal rates will be necessary for assessing the feasibility and impact of removal, understanding long-term dynamics, and optimizing management actions and decisions for this high priority species.

Here we develop an integrated integral projection model (IPM<sup>2</sup>) to quantify size-structured harvest rates of invasive European green crab. As part of a state-space framework, process dynamics are described using an integral projection model, where the population structure changes over time through seasonal growth, natural mortality, and removal, and observations are generated through the use of multiple size-selective removal methods (Figure 1). Combining multiple datasets in an integrated framework allows for distinct inference about size-structured harvest and natural mortality rates, “additional parameters” for which no explicit data is collected. Three datasets contribute to different aspects of the model, with size-at-age data informing seasonal growth rates, mark-recapture data primarily informing natural mortality and trap capture rates, and a time series dataset merging all components and informing latent and observation processes across multiple years (Figure 2). The IPM<sup>2</sup> facilitates prediction of the stable size distribution and equilibrium abundance under different removal strategies, providing a framework for accurate assessment of the effectiveness of invasive species control programs.

## 3 Methods and materials

### 3.1 Study system

Listed as one of the world’s 100 worst invaders (Lowe et al., 2000), the European green crab (*Carcinus maenas*) has successfully colonized all continents except Antarctica (Yamada, 2001), negatively impacting benthic communities (Grosholz, 2005), eelgrass habitat (Garbary et al., 2014; Howard et al., 2019), and commercial shellfish industries (Grosholz et al., 2011). The first introduced green crab population in the Northeast Pacific was established in San Francisco Bay in the early 1990’s. The crab’s range has since

expanded northward through larval transport in the Davidson Current (Yamada et al., 2021); the green crab expanded into Oregon and Washington coastal estuaries in 1998 and into the Salish Sea in 2016 and 2017.

The European green crab is a resilient marine invader, able to withstand a wide range of environmental conditions and rapidly develop cold tolerance adaptations (Tepolt & Palumbi, 2020). Additionally, the green crab can recolonize quickly after removal, largely due to a mismatch in spatial scale between local removal programs and the crab’s wide-ranging dispersal and population processes (Keller et al., 2025). While adult crabs occupy sheltered bays and estuaries, the crab disperses during a larval phase in open marine waters that facilitates recolonization of suppressed populations in neighboring habitats (Yamada et al., 2021). Understanding the relationship between removal effort and population suppression will therefore be critical for identifying the feasibility of suppression and for efficiently allocating conservation resources.

### 3.2 Demographic data

Population-level inference is focused on size-structured green crab abundance at Drayton Harbor in Washington, USA in the Salish Sea (Appendix 1.1, Figure A1.1). The primary dataset (D1) comprises an time series of counts of removed crabs collected in April - October from 2020-2023 and is structured by body size (Figure 1B). Multiple trap types (Fukui, Minnow, and Shrimp) with different size-selective removal rates were used (Figure 1B), and traps were baited and left to “soak” in intertidal and subtidal habitat for 24-48 hours. As Drayton Harbor is a protected bay, nearly enclosed by a land spit, we assume no movement of non-larval crabs in and out of the study site within each year.

Inference was supplemented by two additional datasets, size-at-age data (D2) and batch mark-recapture data (D3). Size-at-age data were collected from records of crab removal observations in northeastern Pacific estuaries from Yamada et al. (2005), when the somatic growth of a large recruitment class was tracked over time (Yamada et al., 2005). Size-structured batch mark-recapture data (D3) was included to inform trap capture rate parameters of Fukui traps. These data were collected by Grosholz et al. 2021 in a mark-recapture experiment in Seadrift lagoon, California, USA (Grosholz et al., 2021), where green crabs were captured with Fukui traps, batch marked (not individually identifiable), released, and then recaptured within a few weeks.

More information on all three datasets can be found in Appendix 1.

### 3.3 Model description

We start by detailing the overall state-space IPM<sup>2</sup> model, including the process and observation sub-models (Figure 1). Description of all model parameters can be found in Table 1.

### 161 3.3.1 Process model

162 The process model describes how the population at Drayton Harbor changes through time due to growth,  
 163 natural mortality, removal, and annual recruitment (Figure 1A). The following process equations describe the  
 164 integral projection model that uses a kernel to project the population forward in time based on seasonal size-  
 165 dependent growth and size-dependent natural survival (Rees et al., 2014). Over the winter between years,  
 166 size-dependent natural survival in the kernel becomes density-dependent. We then describe the annual  
 167 recruitment process and how the time series is initialized with the first population density. These equations  
 168 detail how the population changes within a year (intra-annual change) and between years (inter-annual  
 169 change) (Figure 1A). The model tracks the state of the population in terms of its distribution of carapace  
 170 sizes,  $N(x)_{t,y}$ , which is the density function of individuals of size  $x$  during year  $y$  at time  $t$ .

### 171 Integral projection model

172 The population density,  $N(x)_{t,y}$ , is projected forward in time using an integral projection model. The  
 173 integral projection uses a continuous distribution over  $x$ , and for estimation, the abundance of individuals  
 174 is discretized using a small size interval ( $\Delta x = 5$ ). The discretization is centered on size  $x$ , such that  
 175 the IPM kernel is constructed using the midpoints of the discrete IPM mesh. The total population size is  
 176  $\int_{x \in \Omega} N(x)_{t,y} dx$ , where  $\Omega$  represents all biologically feasible sizes (0 - 110 mm).

177 A kernel,  $K(x', x, D)$ , describes the probability density of moving from size  $x$  to size  $x'$ . This kernel  
 178 is time-dependent, where  $D = \{d_t, d_{t+1}\}$ , a vector of Julian calendar dates associated with  $t$  and  $t + 1$ .  
 179  $N(x)_{t+1,y}$  is therefore a function of  $N(x)_{t,y}$ ,  $K(x', x, T)$ , the total count of removed crabs in each time  
 180 period,  $C^T(x)_{t,y}$ , and recruit abundance,  $R(x)$ . Recruits annually enter the model at time point,  $t_R$ :

$$N(x)_{t+1,y} = \int_{x' \in \Omega} K(x', x, D)(N(x)_{t,y} - C^T(x)_{t,y})dx + R(x)I_{t=t_R} \quad (1)$$

181 The kernel is defined as the product of a growth kernel,  $G(x', x, D)$  and size-dependent natural survival,  
 182  $S(x)$ :

$$K(x', x, D) = G(x', x, D) \times S(x) \quad (2)$$

### 183 Seasonal growth

184 Like many ectotherms, green crab growth is strongly seasonal, with the growth rate peaking in the summer  
 185 due to seasonal variation in temperature, light, and food availability (Contreras et al., 2003; García-Berthou

et al., 2012). We therefore use a seasonal growth model that modifies the traditional von Bertalanffy growth model to allow the growth rate to change as a function of time (Beverton & Holt, 2012; Somers et al., 1988).

$$\mu_{x,D}^G = x + (x_\infty - x)(1 - \exp(-k\Delta t - s(d_{t+1}) + s(d_t))) \quad (3)$$

$$s(d_{t+1}) = \frac{Ak}{2\pi} \sin(2\pi(d_{t+1} - t_s)) \quad (4)$$

$$s(d_t) = \frac{Ak}{2\pi} \sin(2\pi(d_t - t_s)) \quad (5)$$

where  $\mu_{x,D}^G$  is the mean size at  $t + 1$ ,  $x_\infty$  is the asymptotic average size,  $k$  is a measure of the exponential rate of approach to the asymptotic size,  $A$  modulates the amplitude of the growth oscillations, and  $t_s$  is the time between the start of the calendar year and the start of the convex portion of the first sinusoidal growth oscillation (García-Berthou et al., 2012).

To account for variation in growth rate among individuals,  $G(x', x, D)$  is described as:

$$G(x', x, T) = \phi(x'; \mu_{x,D}^G, \sigma^G) \quad (6)$$

where  $\sigma^G$  is the standard deviation in the growth rate.

## Natural mortality

The rate of natural mortality decreases with size, as smaller crabs have lower intra- and inter-specific competitive abilities and are more susceptible to predation and cannibalism (Grosholz et al., 2021; Maszczyk & Brzezinski, 2018). Natural survival,  $S$ , is described as:

$$S(x)_y = \exp\left(-\Delta t(\beta_y + \frac{\alpha}{x^2})\right) \quad (7)$$

where  $\beta$  is the intensity of size-independent natural mortality and  $\alpha$  is the intensity of size-dependent natural mortality (Carlson et al., 2010). Here, process error enters the model as year-specific, size-independent natural mortality:

$$\beta_y \sim \text{Gamma}(\beta_\alpha, \beta_\theta) \quad (8)$$

## Density-dependent overwinter mortality

To transition from year  $y$  to year  $y + 1$ , the population density experiences seasonal growth and density- and size-dependent overwinter survival. The abundance of crabs in each size interval surviving the winter,



206  $N(x)_{t=1,y+1}$ , is drawn from a binomial distribution, where the number of trials is the abundance of crabs after  
 207 seasonal growth and removal in the last time period,  $t_{max}$ , and the probability of success is the probability  
 208 of overwinter survival,  $S^o(x)_y$ . Here, the date interval,  $D$ , used in the seasonal growth equation,  $G(x', x, D)$ ,  
 209 corresponds to the last time point in the year,  $t_{max}$ , and the first time point in the following year (i.e.,  
 210  $D = \{d_{t=t_{max}}, d_{t=1}\}$ ).

$$N(x')_{t,y} = \int_{x' \in \Omega} G(x', x, D)(N(x)_{t_{max},y} - C^T(x)_{t_{max},y})dx \quad (9)$$

$$N(x)_{t=1,y+1} \sim \text{Binomial}(N(x')_{t,y}, S^o(x)_y) \quad (10)$$

212 Due to thermal stress and starvation, the intensity of overwinter mortality is likely stronger than other  
 213 times of the year and plays an important role in population regulation through density-dependent control  
 214 on population size (Henderson et al., 1988). Overwinter mortality is also size-selective; smaller animals tend  
 215 to have lower energy reserves than larger animals and use reserves more rapidly due to the allometry of  
 216 metabolic rate (Hurst, 2007). Probability of overwinter survival,  $S^o$ , is therefore modeled as a density-size  
 217 interaction, such that the intensity of size-dependent overwinter mortality increases at higher population  
 218 densities.  $N_{t_{max},y}^T$  is the total abundance in year  $i$  at the onset of winter.

$$S^o(x, N_{t_{max},y}^T) = \exp\left(-\frac{\alpha_y^o \times N_{t_{max},y}^T}{x^2}\right) \quad (11)$$

219 Process error enters as a year-specific strength of density- and size-dependent overwinter mortality.

$$\alpha_y^o \sim \text{Gamma}(\alpha_\alpha^o, \alpha_\theta^o) \quad (12)$$

220 Since density-dependence during overwinter mortality plays an important role in long-term population dy-  
 221 namics, we performed a model selection procedure to test a variety of formulations for Equations 11-12  
 222 (further described in the Model Fitting section below).

### 223 Initial population density and annual recruitment

224 To initiate the process model time series, we estimate the size distribution and density of adults in the first  
 225 time period. We also estimate the size-structured abundance of recruits that enter the process model in  
 226 mid-May of each year.

227 The first size distribution,  $N(x)_{t=1,y=1}$ , is defined in terms of the abundance of adults in the first time  
 228 period during the first year,  $\lambda^A$ , the mean initial adult size in millimeters,  $\mu^A$ , and the standard deviation  
 229 of initial adult size in millimeters,  $\sigma^A$ . Since the first year of time series data (D1) in 2020,  $y = 1$ , was at

the start of green crab establishment in Drayton Harbor, we expected the size distribution to be dominated by age-one crabs with a few age-two crabs representing the first cohort of colonizing crabs. We therefore assumed a log-normal initial density,  $N(x)$ , to allow for a unimodal size distribution with a longer right tail. Here,  $\phi_L$  represents the log-normal probability density function.

$$N(x)_{t=1,y=1} = \phi_L(x; \mu^A, \sigma^A) \times \lambda^A \quad (13)$$

Ovigerous females spawn in August-December (Klassen & Locke, 2007), and these planktonic larvae exit estuarine habitat to develop in high salinity coastal waters alongside larvae produced by neighboring habitats. Advection then brings larvae back into the estuary during recruitment (Young & Elliott, 2019). Recruitment is therefore modeled as an annual event with open demography, where the annual abundance of recruits,  $\lambda_y^R$ , is independent of adult abundance and follows a normal distribution, truncated such that  $\lambda_y^R > 0$

$$\lambda_y^R \sim Normal(\mu^\lambda, \sigma^\lambda) \quad (14)$$

The annual size distribution of recruits,  $R(x)_y$ , is defined in terms of the annual abundance of recruits,  $\lambda_y^R$ , the mean initial recruit size in millimeters,  $\mu^R$ , and the standard deviation of initial recruit size in millimeters,  $\sigma^R$ .  $\phi$  represents the normal probability density function.

$$R(x)_y = \phi(x; \mu^R, \sigma^R) \times \lambda^R \quad (15)$$

Most crabs will settle from their planktonic larval stage in January to April (Yamada et al., 2005). Instead of estimating the time of planktonic settlement, we represent the recruits annually entering the population at a time point,  $t_R$ , (Equation 1) when 1) it can be assumed that the larvae have already settled as juveniles (i.e., mean initial recruit size,  $\mu^R$ , is greater than zero) and 2) well before the recruits grow into an observable size (Figure 1B). The recruits therefore enter the process model in mid-May, corresponding to  $t_R = 6$ .

### 3.3.2 Observation model

The observation equations describe the data-generating process for the three datasets. A conditional multinomial observation model was used to describe how the time series data (D1) relates to the latent abundance,  $N(x)_{t,y}$ , at Drayton Harbor. Harvest mortality through trapping is described as a size-selective hazard rate, which is informed by both the time series data (D1) and batch mark-recapture data (D3). The seasonal growth process was informed by the size-at-age data (D2). Inference was conducted sequentially, such that

the seasonal growth parameters were fit with the size-at-age records, and the summarized posteriors were used to develop prior distributions in the integrated population model with D1 and D3.

The process model equations in section 3.3.1 describe the latent population dynamics at Drayton Harbor (D1), including the latent states, overwinter mortality, initial population density, and recruitment process (Figure 2). Other process parameters, including the seasonal growth and natural mortality parameters, were informed by multiple datasets (Table 1, Figure 2). Below we also describe how D2 and D3 inform the shared parameters in the process equations (Figure 2).

### Conditional multinomial observation model (D1)

A conditional multinomial observation model was used to describe the data-generating process for the time series data at Drayton Harbor (D1, Figure 1). Here, the removal count data,  $C(x)_{t,j,y}$ , represent the number of crabs of size  $x$ , caught at time  $t$ , during year  $y$ , in trap  $j$ , conditioned on the size-structured latent abundance,  $N(x)_{t,y}$  (Kéry & Royle, 2015).

Multiple traps were placed simultaneously at each time period, so we factor the probability of the removal count data in one time period as 1) a binomial distribution for the total number of removed crabs,  $C^T(x)_{t,y}$  across all traps and 2) a multinomial-Dirichlet distribution to describe which traps the crabs are caught in, given the total number trapped. The total number of removed crabs,  $C(x)_{t,y}$ , follows a binomial distribution with number of trials equal to the latent total abundance of crabs,  $N(x)_{t,y}$ , and total capture probability across all traps set during the same time period,  $p(x)_{t,y}$ .

$$C^T(x)_{t,y} \sim \text{Binomial}(N(x)_{t,y}, p(x)_{t,y}) \quad (16)$$

The size-structured count of crabs in each trap,  $C(x)_{t,j,y}$ , follows a multinomial-Dirichlet mixture distribution where the probability of capture in trap  $j$ ,  $p^C(x)_{t,j,y}$ , is conditioned on being captured at all. Since the trap compositional data are overdispersed due to green crab aggregation and spatial behaviors, the Dirichlet allows for greater variance in the count data than predicted by a multinomial distribution (Thorson et al., 2017). The parameter  $\alpha^D$  describes the amount of over-dispersion. More information on the multinomial-Dirichlet mixture can be found in Appendix 2.2.

$$C(x)_{t,j,y} | C^T(x)_{t,y} \sim \text{MultinomialDirichlet}(C^T(x)_{t,y}, p^C(x)_{t,j,y} | \alpha^D) \quad (17)$$

### Size-selective hazard rates (D1 and D3)

Removal through trapping occurs in continuous time, described as a size-selective hazard rate,  $H(x)$ , representing the instantaneous intensity of capture (Ergon et al., 2018). The shape and magnitude of this size-selective hazard rate varies among the three trap types used for removal: Fukui, Minnow, and Shrimp traps. Both the time series count data (D1) and mark-recapture data (D3) inform the size-selective probability of capture of Fukui traps, and only D1 informs the size-selective probability of capture of Minnow and Shrimp traps (Figure 2).

Fukui and Shrimp traps capture larger crabs at higher rates than smaller crabs. The respective hazard rates of Fukui and Shrimp traps,  $H_F(x)$  and  $H_S(x)$ , are modeled as a logistic function of crab size:

$$H_F(x) = \frac{h_F^{max}}{1 + \exp\{-h_F^k(x - h_F^0)\}} \quad (18)$$

$$H_S(x) = \frac{h_S^{max}}{1 + \exp\{-h_S^k(x - h_S^0)\}} \quad (19)$$

The Minnow trap mesh size is smaller than the maximum crab size, so this trap's size-selective hazard rate,  $H_M(x)$ , follows a bell-shaped curve (Jørgensen et al., 2009):

$$H_M(x) = h_M^{max} \times \exp\left(\frac{x - h_M^A}{2h_M^\sigma}\right) \quad (20)$$

Each baited trap,  $j$ , is placed in the habitat for a short ( $\sim 24$ -48 hr) time interval,  $\Delta b_{t,j,y}$ . The total capture probability,  $p(x)_{t,y}$ , is described as the probability of not surviving the trapping time interval,  $p(x)_{t,y}$ . The probability of surviving the trap interval is the integrated hazard rate, summed across all traps set during the same time period,  $O_{t,y}$ .

$$p(x)_{t,y} = 1 - \exp\left(-\sum_{j=1}^{O_{t,y}} H(x)_{t,j,y} \Delta b_{t,j,y}\right) \quad (21)$$

The conditional probability of capture,  $p^C(x)_{t,j,y}$ , is then  $H(x)_{t,j,y} / \sum_{j=1}^{O_{t,y}} H(x)_{t,j,y}$ .

### Size-at-age observation process (D2)

The size-at-age data,  $W(a)_y$ , indicate the carapace width of an individual crab at age  $a$  in year  $y$ . These data are fit to a similar seasonal growth model as in Equations 3-5, except instead of describing incremental growth from size  $x$  to  $x'$ , the below growth equation describes growth from the age of the crab when it is of size 0,  $t_0$ , to the expected size of a crab at age  $a$ ,  $\widetilde{W}(a)$ . A normally distributed error term,  $\epsilon_i$ , is used to account for non-independence among data collected in the same year, since growth rate is likely affected by

301 water temperature, which varies from year-to-year.

$$\widetilde{W(a)}_i = x_\infty(1 - \exp(-k(a - t_0) - s(a) + s(t_0))) + \epsilon_i \quad (22)$$

302 To account for variation in growth rate, the observed size-at-age data,  $W(a)_i$ , follow a normal distribution,  
 303 with the expected carapace width  $\widetilde{W(a)}_i$ , and standard deviation,  $\tau_w$ .

$$W(a)_i \sim \text{Normal}(\widetilde{W(a)}_i, \tau_w) \quad (23)$$

304 More information on model fitting with D2 can be found in Appendix 1.2.

### 305 **Mark-recapture observation process (D3)**

306 The mark-recapture data (D3) primarily informed the observation parameters that describe the size-selective  
 307 hazard rates of the Fukui trap type,  $H_F(x)$  (Equation 18, Table 1, Figure 2).

308 The batch mark-recapture data consists of 1) the count of marked and released crabs of size  $x$  in the first  
 309 time period,  $n^{mc}(y)$ , and 2) the count of recaptured and marked crabs of size  $y$  in the second time period,  
 310  $m^{mc}(x)$ . These data follow a binomial distribution:

$$m^{mc}(x) \sim \text{Binomial}(n^{mc}(x)_{t_2}, p^{mc}(x)) \quad (24)$$

311 where  $p^{mc}(x)$  is the total probability of capture based on the total number of Fukui traps set,  $O^{mc}$ , over the  
 312 time period  $\Delta b^{mc}$ :

$$p^{mc}(x) = 1 - \exp\left(-\sum_{j=1}^{O^{mc}} \int H_F(x) \Delta b^{mc}\right) \quad (25)$$

313 These data also informed components of the growth and natural mortality kernel (Table 1; Figure 2).  
 314 The number of marked and released crabs,  $n^{mc}(x)$ , at  $t_1^{mc}$  underwent seasonal growth and natural mortality  
 315 to the next time period,  $t_2^{mc}$ .

$$n^{mc}(x)_{t_2} = \int_{x \in \Omega} K(x', x, D^{mc}) n^{mc}(x)_{t_1} dx \quad (26)$$

## 316 **3.4 Model fitting**

317 We fit the IPM<sup>2</sup> in a Bayesian framework using NIMBLE v.1.2.1 (de Valpine et al., 2017). We used vague  
 318 priors for all parameters, which are provided in Appendix 2.1. Parameters were estimated by running four

Markov chain Monte Carlo (MCMC) chains of 100 000 iterations, thinned by a factor of 10. Of these 10 000 samples, 2 000 were discarded as burn-in. We used visual inspection of the MCMC chains and the Brooks and Gelman diagnostic  $\hat{R}$  to assess model convergence, and we found that all parameters had an  $\hat{R} \leq 1.05$  (Brooks & Gelman, 1998). All analyses were conducted in R version 4.4.1 (R Core Team, 2024). Code for the entire model is provided in Appendix 3, and generic, modular code that closely follows the model description is provided in Appendix 4. Posterior summaries, as well as convergence diagnostics and trace plots of model parameters can be found in Appendix 5.

To assess model performance and robustness, we conducted both a model selection and a model checking procedure (Conn et al., 2018). For model selection, we evaluated multiple functional forms of overwinter mortality using Watanabe–Akaike information criterion (WAIC). The inter-annual population transitions (i.e., transition from year  $y$  to year  $y + 1$ ) are largely described by density-dependent overwinter mortality. Since density dependence only enters the model during this process and is therefore likely influential for forecasting the stable size distribution, we compared multiple functional forms for size- and density-dependent overwinter mortality and used the formulation with the lowest WAIC in the final analysis (Eq. 11-12, Appendix 6.1).

To check the model, we calculated posterior predictive p-values using deviance as an omnibus discrepancy function and proportion of zeros as a targeted discrepancy function to check for zero inflation of the count data (Appendix 6.2). We found that the model was an adequate representation of the data-generating process by these measures, with p-values of 0.92 for both the deviance and proportion of zeros discrepancy functions (Appendix 6.2; Figures A.6.1 and A.6.2). These p-values may be conservative, however, as Bayesian p-values tend to be biased toward 0.5 (Conn et al., 2018).

### 3.5 Population forecasts

To evaluate how a green crab population responds to varying removal efforts, we calculated the stable size distribution and equilibrium abundance through stochastic forward simulations with posterior samples. We randomly selected 1000 samples from the full posterior, and for each posterior sample, we generated an initial adult population size and projected the population forward 25 years, applying varying removal efforts for each set of 1000 simulations. These varying removal efforts included a total of 0, 28, 112, 560, 840, 1400, and 2800 annual Shrimp, Fukui, or Minnow traps, applied evenly over the trapping season of 14 biweeks (21 total sets of 1000 simulations; seven removal efforts x three trap types). Year-varying quantities, like recruit abundance, size-independent natural mortality, and size- and density-dependent overwinter mortality were drawn stochastically each year in the forward simulations (Table 1). To reduce simulation noise between

removal effort and trap type combinations, the same set of year-varying stochastic draws for each posterior sample was consistent across the 21 simulation sets.

For each simulation, the size-structured abundance at the end each of year after overwinter mortality,  $N(x)_{t=1,y+1}$ , was recorded (Figure 1A). Simulation outputs were summarized as the mean size-structured abundance at the end of years 6-25, with the first five years treated as transient to ensure the population reached an equilibrium.

## 4 Results

### 4.1 Estimating population-level quantities

The integrated integral projection model tracked the size-structured European green crab abundance at Drayton Harbor (D1) throughout 2020-2023. Figure 3 shows the population density of adults and recruits at the beginning of each year. As the invasion progressed from 2020 to later years in 2021-2023, the size structure of adults shifted toward larger individuals (Figure 3A). This increase in median adult crab size coincided with a decrease in overall adult crab abundance. Total adult crab abundance decreased from 330 (262-508 95% CrI) individuals in 2020 to 248 (235-264 95% CrI), 155 (146-166 95% CrI), and 168 (154-194 95% CrI) in 2021-2023. The size distribution of adults in 2021-2023 is also bimodal, as the recruit age class that survived the winter (year one class) had not yet grown in size to match the sizes of crabs older than one year.

The abundance of recruits varied by multiple orders of magnitude across years, ranging from 1276 (527-2200 95% CrI) and 2267 (1221-3468 95% CrI) in the strong recruitment years of 2020 and 2022, to 51 (29-75 95% CrI) and 121 (39-222 95% CrI) in the weak recruitment years of 2021 and 2023 (Figure 3B). The mean size of recruits when they enter the model at  $t = 6$  each year (mid-May) is 9.4 (3.9-14.4 95% CrI) millimeters (Figure 3B; Appendix 5). The credible intervals around recruit abundance and size are large (Figure 3B), since these recruits are unobserved until they grow into the observable size range in August - September (Figure 1B).

### 4.2 Distinguishing size-structured natural and harvest mortality

By combining information in multiple datasets, the integrated population model allowed for estimation of additional parameters – size-structured natural and harvest mortality – that were not identifiable with any one component dataset (Riecke et al., 2019).

Removal rates were estimated for three trap types – Fukui, Shrimp, and Minnow – with different rates

of removal and size selectivities (Figure 4). Overall, Shrimp traps removed crabs at the highest rate, and Minnow traps were only effective at removing crabs in the 30-60 mm size range. No trap effectively removed crabs smaller than 30 mm, consistent with the completely unobserved small crab portion of the population (Figure 1B).

Overwinter natural survival rates were lower than natural survival rates at other periods of the year (Figure 5), consistent with the expectation that density-dependence in overwinter mortality plays an important role in population regulation. Model comparison with WAIC demonstrated support for overwinter mortality as a function of the interaction between individual crab size and total population density (Appendix 6.1, Eq. 11-12). Overwinter survival rates varied significantly from year to year, and corresponded to total abundance and recruitment strength. Years with higher overall population density coincided with particularly low overwinter survival of small crabs (Figure 5B). In response to a large recruitment event in 2022 (Figure 3), between 2022 and 2023, overwinter survival rates dropped to 0.75 for crabs of the largest size and less than 0.25 for crabs of smaller size (Figure 5B). Conversely, survival rates were high over the winter between 2021 and 2022 in response to a small recruitment event in 2021 (Figure 3B) and an adult size structure biased toward larger crabs (Figure 3A).

### 4.3 Isolating growth in body size as a time- and size-dependent process

Isolating the contribution of growth in body size in individuals to changes in population size structure helped facilitate inference of other size-structured demographic rates. This growth rate was strongly seasonal and varied throughout the year, with growth rate peaking in the summer months and approaching zero in the winter months (Appendix 1; Figure A1.2). The mean asymptotic crab size,  $x_\infty$  was 82.6 mm (80.2-84.7 95% CrI) (Appendix 5). Growth was also variable; the standard deviation in growth rate,  $\sigma^G$ , was 2.8 mm (2.4-3.3 95% CrI).

### 4.4 Population forecasts

Forward simulations with posterior samples were used to forecast the stable size distribution and equilibrium abundance under different levels of removal effort, and subsequently different levels of removal mortality (Figure 6; Figure S1). These simulations indicated that a low removal effort with Fukui and Minnow traps (Figure 6B-C) resulted in only marginal changes in the stable size distribution and equilibrium abundance relative to no removal effort (Figure 6A, Table S1). For example, the mean total equilibrium abundance,  $N_{total}^E$ , across simulation replicates with no removal was 302 crabs (Table S1). While the mean  $N_{total}^E$  across simulation replicates with 2800 annual Fukui and Minnow traps was 275 and 221, respectively (Table S1).



With a high removal effort of Shrimp traps, the total equilibrium abundance,  $N_{total}^E$ , decreases to 114 crabs, and large adult crabs can be completely removed from the population (Figure 6D, Table S1). However, this large crab removal merely shifts the stable size distribution toward smaller crabs; the equilibrium abundance of smaller crabs actually increases relative to no removal effort (Figure 6; Figure S2).

## 5 Discussion

Our integrated integral projection model (IPM<sup>2</sup>) – a framework first described by Plard et al. 2019 – provides important insight for how body size can modulate many individual-level demographic rates that interact to describe population-level dynamics (Plard et al., 2019). The IPM<sup>2</sup> produces an outcome greater than the sum of its parts: combining information in multiple datasets facilitates population inference for a species with complex demography with imperfect measurements. We are able to disentangle multiple size-structured demographic rates that simultaneously change the size distribution over time (Carlson et al., 2010; Sogard, 1997), allowing detailed understanding of the individual contributions of each size-structured demographic rate to overall population dynamics. These results also fill a significant knowledge gap for invasive European green crab (*Carcinus maenas*). While the species is highly studied due to its geographic ubiquity and management concern, this study is the first to quantify removal rates and size-structured population abundance (Young & Elliott, 2019), providing a path forward for assessing feasibility of control and optimizing management actions.

### 5.1 Disentangling size-structured demographic rates

The IPM<sup>2</sup> provides novel insights into previously uncharacterized size-structured European green crab demographic rates, including size-structured natural mortality, size- and density-dependent overwinter mortality, recruitment processes, size-structured growth rate, and size-selective harvest rates.

The model estimates unobserved quantities, most notably the abundance and size (carapace width) of recruits before they grow into observable size classes (Figure 1B, Figure 3). Due to the size-selectivity of most observation and removal methods, very little is known about early stages of the green crab life cycle (Yamada et al., 2005), yet here, we used growth rates estimated by larger sizes to back-calculate size and abundance of recruits before they are observed. However, the model may underestimate non-overwinter natural mortality at smaller sizes, and therefore recruit abundance, because after settlement from the planktonic phase, recruits are unobserved for months and are likely dying naturally at very high rates (Figure 5A). Previous studies have reported extremely high densities of small juveniles (200-2000/m<sup>2</sup>) when counting individuals within quadrats in the intertidal zone at low tide where most juveniles are found (Breteler, 1976; Thiel & Darnedde,

1994). Future work should integrate trap data with measurements of small crabs made by other means to improve estimates of size-dependent, non-overwinter natural mortality.

Overwinter mortality is a major size-selective process for green crab, as both energy reserves and metabolic rate scale with body size, such that larger individuals have higher energy reserves but lower metabolism and are therefore more resilient to starvation and physical extremes (Carlson et al., 2008; Sogard, 1997). Our results provide support for the hypothesis that crab size and overall population density interact to modulate overwinter mortality (Appendix 6.1). Overwinter mortality varied significantly from year to year, and years with higher overall population density coincided with particularly low overwinter survival of small crabs (Figure 5B). Overwinter mortality therefore plays an important role in regulating size-structured green crab population dynamics (Henderson et al., 1988).

Recruitment varies by orders of magnitude from year to year, ranging from about 50 recruits in weak years to 500-1000 in strong years (Figure 3B). The green crab has a long planktonic larval stage, living in open marine waters for months before advection back into the estuarine environment where they settle in the sub- and inter-tidal zone (Yamada, 2001). This inter-annual variability in recruitment is consistent with varying oceanographic conditions; the survival and successful transport of larvae often coincide with oceanographic conditions like El Niño/Southern Oscillation (ENSO) and Pacific Decadal Oscillation (PDO) events (Yamada et al., 2021). Though drawing from a small sample size of only four years, recruit abundance appears to be decoupled from adult abundance; for example, adult abundance was highest in 2020, yet recruit abundance was lowest in 2021 (Figure 3). This decoupling of recruit and adult abundances suggests that recruitment is likely driven by oceanographic conditions and regional population dynamics, rather than local adult abundance. However, a longer time series coupled with genomic information will be necessary for understanding the relative contribution of local and regional processes in population dynamics.

The model also estimates demographic rates that vary across time and size. Marine invertebrates are often subject to seasonal variation in environmental factors like photoperiod, food availability, and temperature, resulting in elevated rates of growth, feeding, and oxygen consumption in the summer season (Brockington & Clarke, 2001). Consistent with this expectation, we find that green crab growth is strongly seasonal, with growth rate peaking in summer months and approaching zero in the winter (Appendix 1, Figure A1.2). Green crab growth rate is also size-dependent, as the molting rate is much higher at smaller crab sizes, and slows down as crabs approach the mean asymptotic crab size (Yamada et al., 2005).

The IPM<sup>2</sup> estimates absolute size-selective capture rates for different trap types, allowing for the first estimates of size-structured abundance of European green crab (Young 2020). These capture rate and abundance estimates mark an important advance for moving beyond catch per unit effort (CPUE) as the primary, yet imperfect, method for measuring population trends (Harley et al., 2001). These results also

highlight the strong size-selectivity of removal methods, since capture rates approaching zero for crabs smaller than 30 millimeters (Figure 4). This size selectivity will have uncertain evolutionary consequences, yet resolving this uncertainty will be important for projecting long-term population dynamics. Fishing-induced life-history evolution is frequently observed in harvested species (Enberg et al., 2012), and previous invasive species management programs have demonstrated rapid ecological and evolutionary changes in response to selective harvesting, including a shift toward earlier size at maturity and an overall slower growing phenotype (Evangelista et al., 2015). These phenotypic responses to removal programs can have strong effects on ecosystem recovery and therefore should be considered in invasive species management (Závorka et al., 2020).

While the IPM<sup>2</sup> is able to represent many simultaneous size-structured demographic rates, the model specification may contain un-modeled heterogeneity in capture probability. To account for the spatial aggregative behavior of green crab, we used a Dirichlet-multinomial mixture to account for overdispersion (i.e., inflation of zeros in count data) across traps set simultaneously (Thorson et al., 2017; Young & Elliott, 2019). However, the presence of large crabs or large numbers of crabs can also dissuade additional crabs from entering a trap, suggesting that the degree of spatial aggregation scales non-linearly with crab abundance and trap count. Additionally, since the traps are baited, these capture rates reflect crab foraging activity, which can vary across time, age, and sex (Young & Elliott, 2019). Newly molted crabs and ovigerous females tend to burrow and not feed, which would contribute to heterogeneity in capture rates across individuals (Ropes, 1968). However, the model structure accounts for both measurement and process error, flexibly accommodating this likely unmodeled heterogeneity across individuals.

## 5.2 Predicting the impact of removal on European green crab dynamics

These estimates of size-structured demographic rates and size-selective harvest rates will be essential for understanding the impact of removal on green crab dynamics and the feasibility of population suppression. Our model results show that harvest mortality associated with low levels of removal effort, especially with Fukui and Minnow traps, only marginally change the equilibrium abundance and size structure, relative to doing nothing (Figure 6A-C). These results highlight that low levels of removal effort can be useful for monitoring population trends but are insufficient for control and population suppression.

Even with extremely high levels of removal effort with the most effective trap type - Shrimp traps - control will mostly shift the stable size distribution of the population toward smaller crab sizes (Figure 6D; Figure S1-2). These results are consistent with observations of decreased median carapace width in response to removal (de Rivera et al. 2007), and they support the prediction that though removal programs may achieve short-term or local benefits, control is likely unable to sustainably suppress populations over larger temporal

and spatial scales (Kanary et al., 2014; Keller et al., 2025; Tummon Flynn et al., 2024). In fact, strong removal pressure increases the equilibrium abundance of small crabs (Figure 6D, Figure S2). Removing adult crabs reduces the intra-specific regulation of recruits, resulting in higher abundance of smaller crabs relative to doing nothing. This increase in small crab abundance is similar to dynamics observed in an intensive control experiment by Grosholz et al. 2021, the first controlled experimental field demonstration of the “hydra effect” (Grosholz et al., 2021).

These results are dependent on the assumption of a completely open population, where recruitment rate is independent of the local population size. This assumption, in general, is likely appropriate, as genomic analyses find high gene flow and low genetic structure across local green crab populations in the Northeast Pacific (Tepolt et al., 2009, 2022). However, these results may not be applicable in some circumstances where a local population is partially closed, often in unique oceanographic conditions that support strong larval retention within an estuary (Grosholz et al., 2021).

The population forecasts assume stationarity when predicting the equilibrium abundance and stable size distributions under varying removal efforts (Figure 6). While these forecasts allow recruitment to vary from year to year, the mean and variance of recruit abundance is constant over time. These forecasts would likely be different if recruitment is non-stationary, either decreasing over time through intensive removal efforts across the region or at known source populations, or increasing over time through increased colonization pressure. Non-stationarity in colonization pressure is likely, as climate and oceanographic models project an increase in the oceanographic events that support survival and transport of larvae (Cai et al., 2021; Du et al., 2024). Other demographic rates, like growth and natural mortality may also be non-stationary. The integral projection model lends itself to understanding how changing environmental factors, like increased temperature, will affect individual size and overall population dynamics (Dahlgren & Ehrlén, 2011; Plard et al., 2019).

### 5.3 Reliability of parameter estimates in integrated population models

Integrated population models are useful for inferring parameters of interest in scenarios where information needs of complex models exceed information in individual datasets. Despite these benefits, often the reliability of parameter estimates remains underexamined. Simulation-based approaches have shown that integrated population model parameters are sensitive to violation of model assumptions, like marker-induced bias in capture–mark–recapture data or heterogeneity in the mortality process (Riecke et al., 2019). Importantly, estimation of “additional”, previously unidentifiable parameters are especially sensitive to violations of model assumptions. In the green crab IPM<sup>2</sup>, management-relevant model parameters, including size-structured

natural and harvest mortality, were not identifiable from one data source but were indirectly identifiable from multiple data sources and the assumed model structure. The IPM<sup>2</sup> uses mark-recapture data from a lagoon with for about 100,000 crabs (Grosholz, D3), a location with about two orders of magnitude greater than the number of crabs in the Drayton Harbor time series dataset (D1). The parameter estimates may be unreliable if the dynamics of these sites may be different. The capture rate may scale non-linearly with abundance, when individuals are not captured independently of one another and the rate of clustering increases with abundance (McCarthy et al., 2013). Future work should investigate the reliability of parameter estimates in scenarios where the ecological dynamics underlying different datasets are not identical.

## 5.4 Embedding the IPM<sup>2</sup> in a decision-making framework

These results demonstrate that removal cannot eradicate an open population of European green crab (Figure 6). While removal can decrease the total equilibrium population size at a location like Drayton Harbor with relatively low colonization pressure (Table S1), removal primarily shifts the stable size distribution toward smaller crabs (Figure 6D). In circumstances where invasive species eradication is infeasible, the management focus often moves toward functional eradication, or suppression of the population below levels that cause unacceptable ecological effects (Green & Grosholz, 2021). The IPM<sup>2</sup> green crab model will be valuable in a decision-making framework to optimize removal actions and find the removal effort – and subsequently the stable size distribution and equilibrium abundance – that both minimizes removal cost and impact on habitat, native species, and natural resources.

Embedding this size-structured model within a decision-making framework, however, will require improved knowledge of size-structured impacts and computational methods to optimize high-dimensional decision problems. Green crab size often mediates its interactions between prey and competitors; other decapod species are preyed upon by green crab as juveniles but outcompete green crab as adults, and green crab often only predate upon bivalves of smaller size (Curtis et al., 2012; Grosholz, 2005; McDonald et al., 2001; Williams et al., 2009). Quantifying size- or biomass-dependent impacts will be critical for optimizing allocation of removal resources in this size-structured system. Additionally, this work highlights that green crab population dynamics cannot be represented in a one-dimensional system (i.e., total abundance), since the size structure of the population plays an important role in long-term dynamics. Techniques like stochastic dynamic programming can be used to optimize sequential decision problems (Marescot et al., 2013), yet due to the curse of dimensionality, these methods will be insufficient for optimizing problems with large state and action spaces. The size-structured green crab abundance changes within a single decision cycle, and multiple trap types have different size-dependent removal rates and different monetary and logistical

565 costs of use (Figure 3). Since the complexity of the decision problem scales non-linearly with the size of  
566 the system, advanced computational methods like neural-network-based reinforcement learning or factored  
567 Markov decision processes will be needed to optimize management actions in this high-dimensional decision  
568 problem (Lapeyrolerie et al., 2022; Nicol et al., 2015).

## 6 Tables

Table 1: Notation and biological meaning of data, latent states, and parameters. Category refers to the parameter categories designated in Figure 2: 1) Init is the size structure of initial population density and annual recruits, 2) Growth is seasonal growth, 2) N. mort is size-dependent and size-independent natural mortality in non-winter months, 3) O. mort is size- and density-dependent overwinter mortality, 4) F obs, M obs, and S obs correspond to the size-selective observation process for Fukui, Minnow, and Shrimp traps, respectively, and 5) Latent corresponds to the latent states in the state-space model (Figure 1).

Symbol	Description	Category
<b>Demographic parameters</b>		
$\mu^A$	Mean adult size in millimeters at $t = 1$ and $i = 1$ .	Init
$\sigma^A$	Standard deviation of adult size in millimeters at $t = 1$ and $i = 1$ .	Init
$\mu^R$	Mean recruit size in millimeters upon entry into the process model at $t = 6$ .	Init
$\sigma^R$	Standard deviation of recruit size in millimeters upon entry into the process model at $t = 6$ .	Init
$x_\infty$	Asymptotic average crab size (carapace width, in mm).	Growth
$k$	Exponential rate of approach to the asymptotic size.	Growth
$A$	Parameter modulating the amplitude of seasonal growth oscillations.	Growth
$t_s$	Time between the start of the calendar year and the start of the convex portion of the first sinusoidal growth oscillation.	Growth
$\sigma^G$	Standard deviation of somatic growth.	Growth
$\beta_y$	Intensity of size-independent natural mortality in year $y$ (not in winter months).	N. mort
$\alpha$	Intensity of size-dependent natural mortality (not in winter months).	N. mort
$\beta_\alpha$	Shape parameter of gamma distribution of intensity of size-independent natural mortality (not in winter months).	N. mort

$\beta_\theta$	Rate parameter of gamma distribution of of intensity of size-independent natural mortality (not in winter months).	N. mort
$\alpha_y^o$	Intensity of overwinter density- and size-dependent natural mortality in year $y$ .	O. mort
$\alpha_\alpha^o$	Shape parameter of gamma distribution of intensity of overwinter density- and size-dependent natural mortality.	O. mort
$\alpha_\theta^o$	Rate parameter of gamma distribution of overwinter intensity of density- and size-dependent natural mortality.	O. mort
<b>Observation parameters</b>		
$h^{max}$	Maximum harvest mortality hazard rate. Maximum rate is trap type-specific, such that $h_F^{max}$ , $h_M^{max}$ , and $h_S^{max}$ correspond to Fukui, Minnow, and Shrimp traps, respectively.	F Obs, M Obs, S Obs
$h^k$	Steepness of change from the minimum to the maximum hazard rate for trap types with a logistic size-selective function, such that $h_F^k$ and $h_S^k$ correspond to Fukui and Shrimp traps, respectively.	F obs, S obs
$h^0$	Midpoint of change from the minimum to the maximum hazard rate for trap types with a logistic size-selective function, such that $h_F^0$ and $h_S^0$ correspond to Fukui and Shrimp traps, respectively.	F obs, S obs
$h_M^A$	Crab size associated with maximum hazard rate with Minnow traps.	M obs
$h_M^\sigma$	Width parameter in the Minnow size-selectivity hazard rate function.	M obs
$\alpha^D$	Parameter that governs the mean and variance of the multinomial-dirichlet mixture distribution.	F obs, M Obs, S Obs
<b>Population-level quantities</b>		



$N(x)_{t,y}$	Population density function of individuals of size $x$ , during year $y$ , at time $t$ .	Latent
$\lambda^A$	Adult abundance at the first time period, $t = 1$ , during the first year, $i = 1$ .	Latent
$\lambda_y^R$	Recruit abundance in year $y$ .	Latent
$\mu^\lambda$	Mean recruit abundance.	Latent
$\sigma^\lambda$	Standard deviation of recruit abundance.	Latent
<b>Observational data</b>		
$C^T(x)_{t,i}$	Total count of removed crabs across all traps during time $t$ , in year $y$ , of size $x$ in the time-series dataset (D1).	-
$C(x)_{t,j,y}$	Count of removed crabs in time $t$ , in trap $j$ , in year $y$ , of size $x$ in the time-series dataset (D1).	-
$O_{t,y}$	Number of observations (traps) in time $t$ , in year $y$ in the time-series dataset (D1).	-
$W(a)_y$	Size of crabs of age $a$ , during year $y$ in the size-at-age dataset (D2).	-
$n^{mc}(x)$	Number of marked and released crabs of size $x$ in the mark-recapture dataset (D3).	-
$m^{mc}(x)$	Number of recaptured marked crabs of size $x$ in the mark-recapture dataset (D3).	-
$O^{mc}$	Number of observations (Fukui traps) in the mark-recapture dataset (D3).	-

## 7 Figures

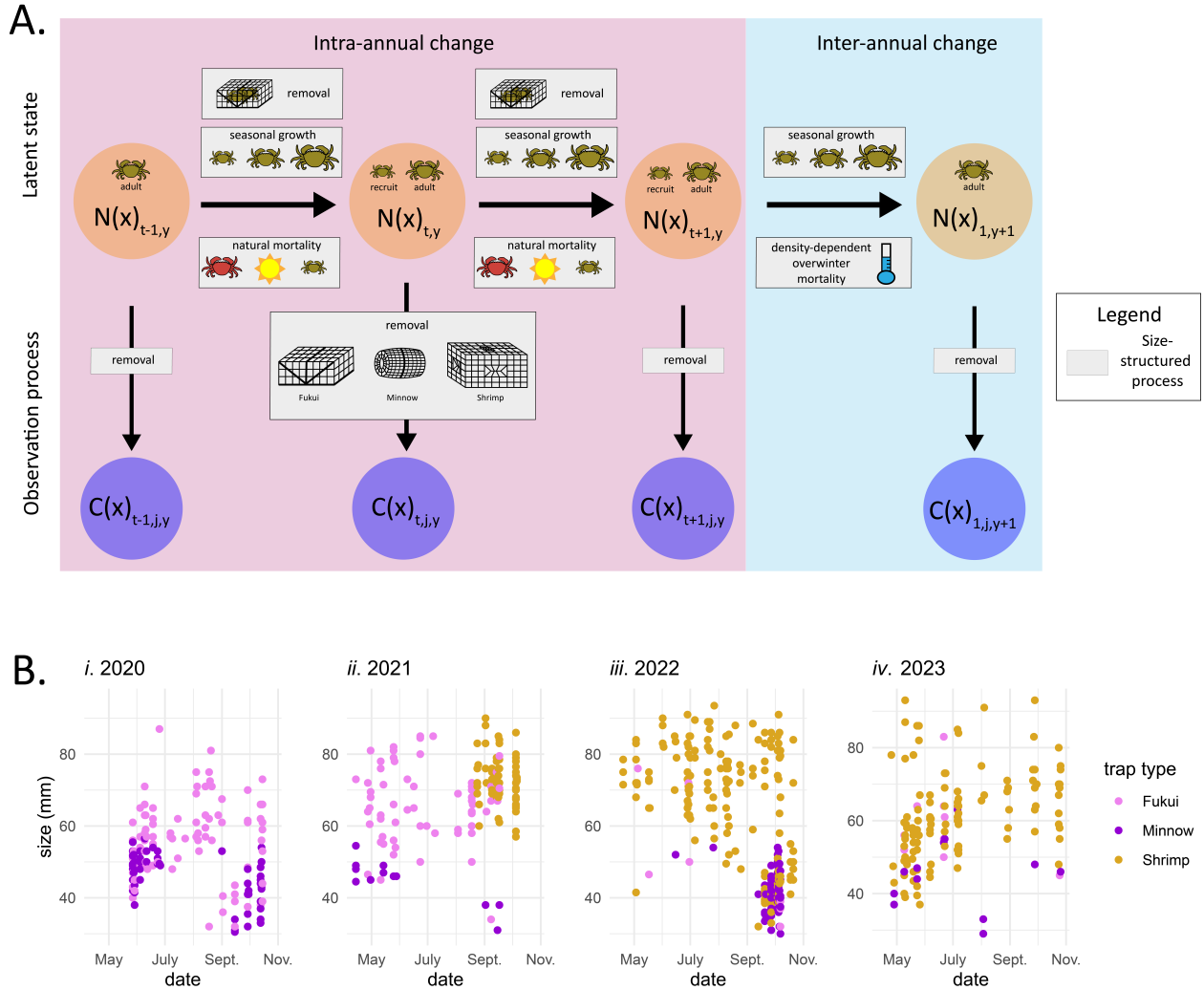


Figure 1: *A.* Conceptual diagram of state-space population model, including the dependence structure in the latent process dynamics and observation process. Orange circles designate the population density of individuals of size  $x$  during year  $y$ , at time  $t$  and are distinguished by dynamics within a year (intra-annual change) and dynamics between years (inter-annual change). Blue circles designate the count of removed crabs during time  $t$ , in trap  $j$ , in year  $y$ , of size  $x$  in the time-series dataset (D1). Grey boxes represent size-structured demographic and observation processes. *B.* Time series data (D1) collected in Drayton Harbor from 2020–2023, highlighting the relationship between time and crab size (carapace width, mm). Each point corresponds to one captured crab, and color corresponds to the type of trap used in capture.

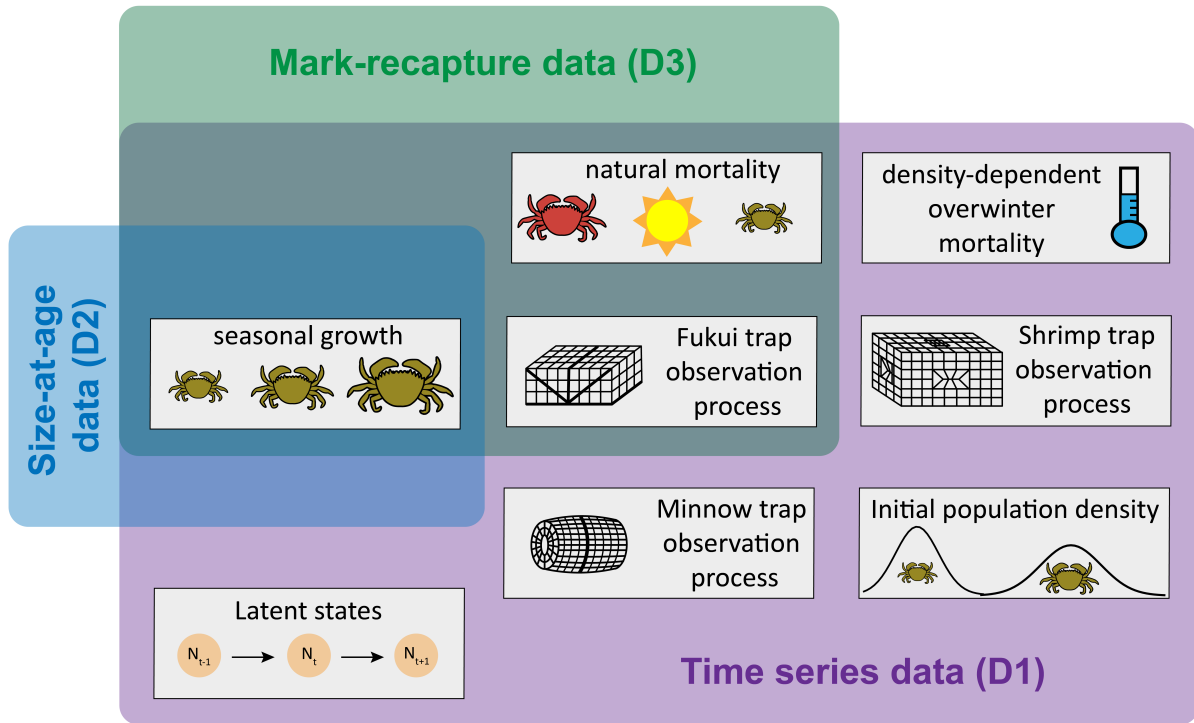


Figure 2: Overview of parameters informed by the three datasets in the integrated population model: time series data (D1) (Figure 1B), size-at-age data (D2) (Figure A1.1), and mark-recapture data (D3). Parameter categories correspond to categories designated in Table 1.

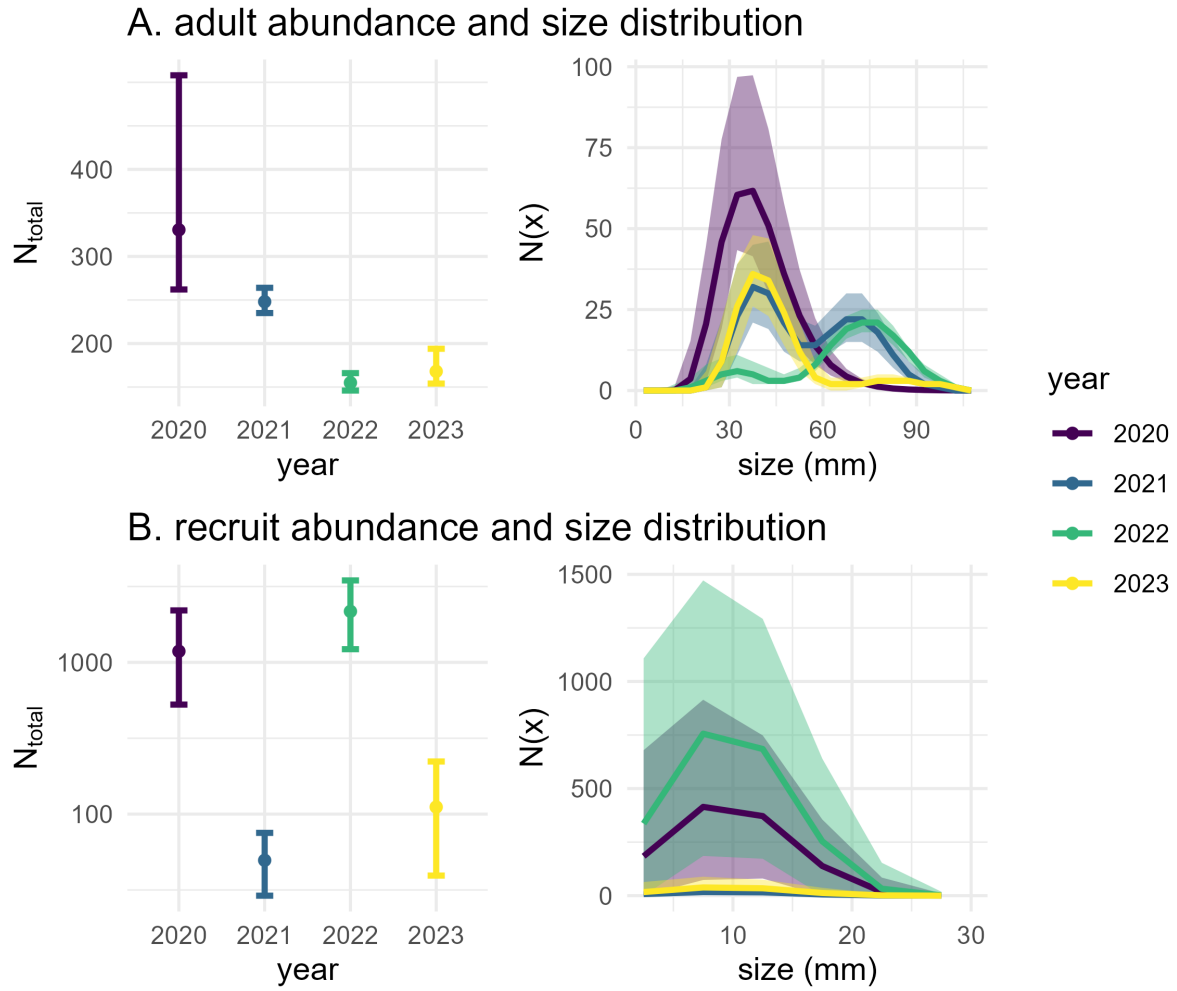


Figure 3: Total abundance and size distribution of *A.* adults and *B.* recruits in the fitted model. The left panels show the total abundance of crabs across all size classes,  $N_{total}$ , at the beginning of each year. The right panels show the size distribution, or the number of individuals in each size class,  $N(x)$ . Size corresponds to crab carapace width. Colors indicate the year, and error bars indicate the 95% credibility interval. Note that the left panel is the integral of the right panel.

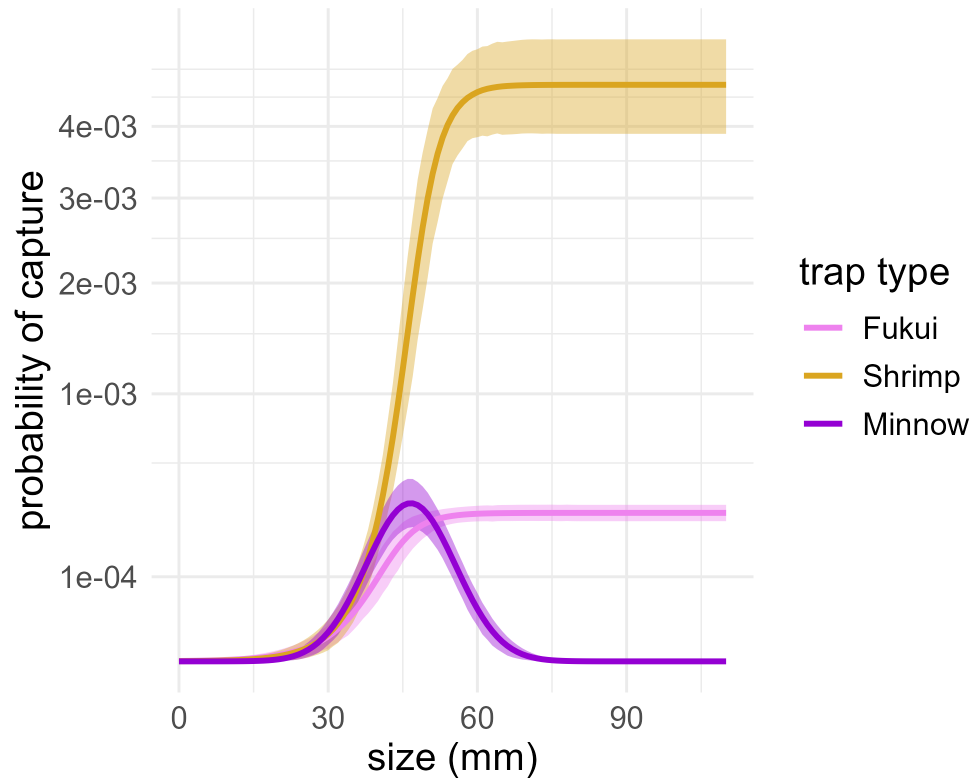


Figure 4: Size-structured probability of capture,  $p_x$ , in one trap over a 24 hours trapping period. Colors indicate the trap type. Note that the y-axis is presented with a square root transformation.

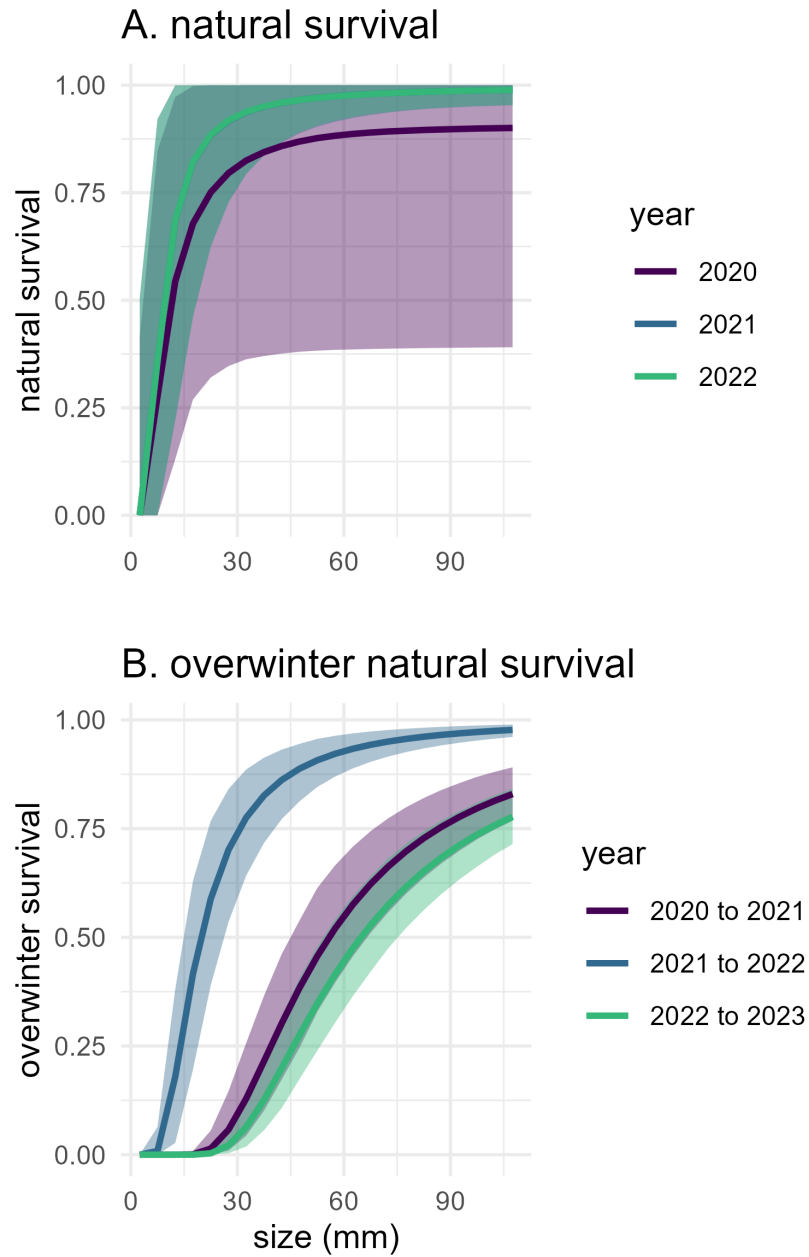


Figure 5: Size-structured natural survival rate in the *A.* non-winter season between Julian days 91 and 305 (Eq. 7) and *B.* overwinter (Eq. 11) between Julian day 306 and Julian day 90 of the following year. Colors indicate year, and error bars indicate the 95% credibility interval.

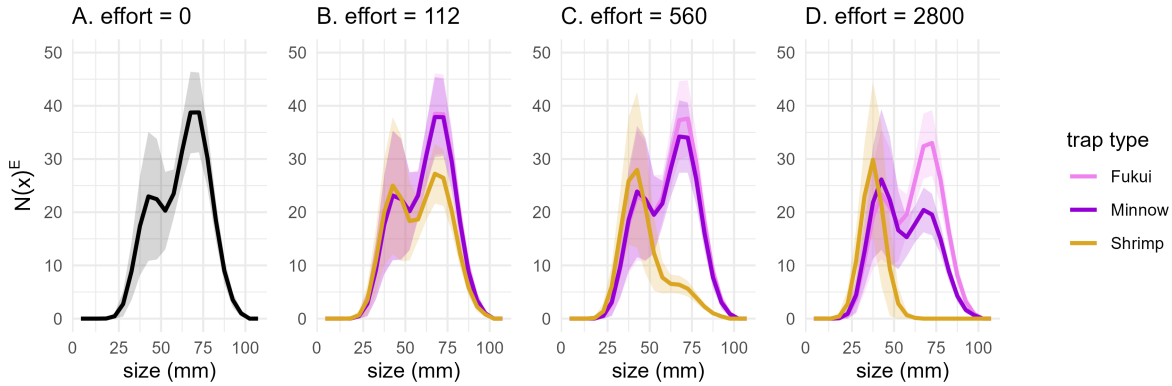


Figure 6: Population forecasts in response to varying removal efforts. Size distributions show the equilibrium crab abundance in each size class,  $N(x)^E$ , at the end of the year after overwinter mortality when *A.* 0 traps, *B.* 112 traps, *C.* 560 traps, and *D.* 2800 traps were applied evenly over a trapping season of 14 biweeks. Solid line indicates the median size-structured abundance across simulation replicates, and the shaded area indicates  $\pm 1$  standard deviation across simulation replicates. Colors indicate trap type used (i.e., in panel B, the purple line shows the resulting size distribution after a trapping effort of 112 Minnow traps).

## 8 Acknowledgements

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