

Lecture 22

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- ① Chi-squared tests for equivalence of two binomial proportions
- ② Chi-squared tests for independence, 2×2 tables
- ③ Chi-squared tests for multiple binomial proportions
- ④ Chi-squared tests for independence, $r \times c$ tables
- ⑤ Chi-squared tests for goodness of fit

Chi-squared testing

- An alternative approach to testing equality of proportions uses the chi-squared statistic

$$\sum \frac{(\text{Observed} - \text{Expected})^2}{\text{Expected}}$$

- “Observed” are the observed counts
- “Expected” are the expected counts under the null hypothesis
- The sum is over all four cells
- This statistic follows a Chi-squared distribution with 1 df
- The Chi-squared statistic is exactly the square of the difference in proportions Score statistic

Example

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Trt	Side Effects	None	Total
X	44	56	100
Y	77	43	120
	121	99	220

- p_1 and p_2 are the cure rates
- $H_0 : p_1 = p_2$

- The χ^2 statistic is $\sum \frac{(O-E)^2}{E}$
- $O_{11} = 44$, $E_{11} = \frac{121}{220} \times 100 = 55$
- $O_{21} = 77$, $E_{21} = \frac{121}{220} \times 120 = 66$
- $O_{12} = 56$, $E_{12} = \frac{99}{220} \times 100 = 45$
- $O_{22} = 43$, $E_{22} = \frac{99}{220} \times 120 = 54$

$$\chi^2 = \frac{(44 - 55)^2}{55} + \frac{(77 - 66)^2}{66} + \frac{(56 - 45)^2}{45} + \frac{(43 - 54)^2}{54}$$

Which turns out to be 8.96. Compare to a χ^2 with one degree of freedom (reject for large values).

```
pchisq(8.96, 1, lower.tail = FALSE)
#result is 0.002
```

```
dat <- matrix(c(44, 77, 56, 43), 2)
chisq.test(dat)
chisq.test(dat, correct = FALSE)
```

Notation reminder

$n_{11} = x$	$n_{12} = n_1 - x$	$n_1 = n_{1+}$
$n_{21} = y$	$n_{22} = n_2 - y$	$n_2 = n_{2+}$
n_{+1}	n_{+2}	

- Reject if the statistic is too large
- Alternative is two sided
- Do not divide α by 2
- A small χ^2 statistic implies little difference between the observed values and those expected under H_0
- The χ^2 statistic and approach generalizes to other kinds of tests and larger contingency tables
- Alternative computational form for the χ^2 statistic

$$\chi^2 = \frac{n(n_{11}n_{22} - n_{12}n_{21})^2}{n_{+1}n_{+2}n_{1+}n_{2+}}$$

- Notice that the statistic:

$$\chi^2 = \frac{n(n_{11}n_{22} - n_{12}n_{21})^2}{n_{+1}n_{+2}n_{1+}n_{2+}}$$

does not change if you transpose the rows and the columns of the table

- Surprisingly, the χ^2 statistic can be used
 - the rows are fixed (binomial)
 - the columns are fixed (binomial)
 - the total sample size is fixed (multinomial)
 - none are fixed (Poisson)
- For a given set of data, any of these assumptions results in the same value for the statistic

Testing independence

- Maternal age versus birthweight¹
- Cross-sectional sample, only the total sample size is fixed

	Birthweight		
Mat. Age	< 2500g	$\geq 2,500g$	Total
< 20y	20	80	100
$\geq 20y$	30	270	300
Total	50	350	400

- H_0 : MA is independent of BW
- H_a : MA is not independent of BW

¹From Agresti Categorical Data Analysis second edition

- Under H_0 (est) $P(MA < 20) = \frac{100}{400} = .25$
- Under H_0 (est) $P(BW < 2500) = \frac{50}{400} = .125$
- Under H_0 (est)

$$P(MA < 20 \text{ and } BW < 2,500) = .25 \times .125$$

- Therefore
 - $E_{11} = \frac{100}{400} \times \frac{50}{400} \times 400 = 12.5$
 - $E_{12} = \frac{100}{400} \times \frac{350}{400} \times 400 = 87.5$
 - $E_{21} = \frac{300}{400} \times \frac{50}{400} \times 400 = 37.5$
 - $E_{22} = \frac{300}{400} \times \frac{350}{400} \times 400 = 262.5$
 - $\chi^2 = \frac{(20-12.5)^2}{12.5} + \frac{(80-87.5)^2}{87.5} + \frac{(30-37.5)^2}{37.5} + \frac{(270-262.5)^2}{262.5} = 6.86$
- Compare to critical value
`qchisq(.95, 1)=3.84`
- Or calculate P-value
`pchisq(6.86, 1, lower.tail = F)=.009`

Chi-squared testing cont'd

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Group	Alcohol use		
	High	Low	Total
Clergy	32	268	300
Educators	51	199	250
Executives	67	233	300
Retailers	83	267	350
Total	233	967	1,200

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- Interest lies in testing whether or not the proportion of high alcohol use is the same in the four occupations
- $H_0 : p_1 = p_2 = p_3 = p_4 = p$
- H_a : at least two of the p_j are unequal
- $O_{11} = 32, E_{11} = 300 \times \frac{233}{1200}$
- $O_{12} = 268, E_{12} = 300 \times \frac{967}{1200}$
- ...
- Chi-squared statistic $\sum \frac{(O-E)^2}{E} = 20.59$
- $df = (Rows - 1)(Columns - 1) = 3$
- Pvalue `pchisq(20.59, 3, lower.tail = FALSE)` ≈ 0

Word distributions

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Word	Book			Total
	1	2	3	
<i>a</i>	147	186	101	434
<i>an</i>	25	26	11	62
<i>this</i>	32	39	15	86
<i>that</i>	94	105	37	236
<i>with</i>	59	74	28	161
<i>without</i>	18	10	10	38
Total	375	440	202	1017

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- H_0 : The probabilities of each word are the same for every book
- H_a : At least two are different
- $O_{11} = 147 \quad E_{11} = 375 \times \frac{434}{1017}$
- $O_{12} = 186 \quad E_{12} = 440 \times \frac{434}{1017}$
- ...
- $\sum \frac{(O-E)^2}{E} = 12.27$
- $df = (6 - 1)(3 - 1) = 10$

Testing independence

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Husband	Wife's Rating				Tot
	N	F	V	A	
N	7	7	2	3	19
F	2	8	3	7	20
V	1	5	4	9	19
A	2	8	9	14	33
	12	28	18	33	91

N=never, F=fairly often, V=very often, A=almost
always

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Independence cont'd

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- H_0 : H and W ratings are independent
- H_a : not independent
- $P(H = N \ \& \ W = A) = P(H = N)P(W = A)$
- $stat = \sum \frac{(O-E)^2}{E}$
- $O_{11} = 7 \ E_{11} = 91 \times \frac{19}{91} \times \frac{12}{91} = 2.51$
- $E_{ij} = n_{i+}n_{+j}/n$
- $df = (Rows - 1)(Cols - 1)$

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```
x<-matrix(c(7,7,2,3,  
            2,8,3,7,  
            1,5,4,9,  
            2,8,9,14),4)  
chisq.test(x)
```

- $\sum \frac{(O-E)^2}{E} = 16.96$
- $df = (4 - 1)(4 - 1) = 9$
- $p - value = .049$
- Cell counts might be too small to use large sample approximation

- Equal distribution and independence test yield the same results
- Same test results if
 - row totals are fixed
 - column totals are fixed
 - total ss is fixed
 - none are fixed
- Note that this is common in statistics; mathematically equivalent results are applied in different settings, but result in different interpretations

- Chi-squared result requires large cell counts
- df is always $(Rows - 1)(Columns - 1)$
- Generalizations of Fishers exact test can be used or continuity corrections can be employed

Chi-squared goodness of fit

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Results from R's RNG

	[0, .25)	[.25, .5)	[.5, .75)	[.75, 1)	Total
Count	254	235	267	244	1000
TP	.25	.25	.25	.25	1

- $H_0 : p_1 = .25, p_2 = .25, p_3 = .25, p_4 = .25$
- $H_a : \text{any } p_i \neq \text{it's hypothesized value}$

- $O_1 = 254$ $E_1 = 1000 \times .25 = 250$
- $O_2 = 235$ $E_2 = 1000 \times .25 = 250$
- $O_3 = 267$ $E_3 = 1000 \times .25 = 250$
- $O_4 = 244$ $E_4 = 1000 \times .25 = 250$
- $\sum \frac{(O-E)^2}{E} = 2.264$
- $df = 3$
- $P - value = .52$

Testing Mendel's hypothesis

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	Phenotype		Total
	Yellow	Green	
Observed	6022	2001	8023
TP	.75	.25	1
Expected	6017.25	2005.75	8023

- $H_0 : p_1 = .75, p_2 = .25$
- $\sum \frac{(O-E)^2}{E} = \frac{(6022-6017.25)^2}{6017.25} + \frac{(2001-2005.75)^2}{2005.75} = .015$

- $df = 1$
- P-value = .90
- Fisher combined several of Mendel's tables
- $\sum \chi^2_{v_i} \sim \chi^2_{\sum v_i}$
- Statistic 42, $df = 84$, P-value = .99996
- Agreement with theoretical counts is perhaps too good?

Notes on GOF

- Test of whether or not observed counts equal theoretical values
- Test statistic is $\sum \frac{(O-E)^2}{E}$
- TS follows χ^2 distribution for large n
- df is the number of cells minus 1
- Undirected alternative is problematic
- Especially useful for testing RNGs
- Kolmogorov/Smirnov test is an alternative test that does not require discretization but often has low power