

→ No lecture on March
No lab 30th

→ Recorded Lecture 3
Thursdays
Lecture

→ PSI is posted April 1

Lecture 2

Review Generalized Linear Models
More on Logistic Regression:
Regression adjustment and continuous covariates

Review of Lecture 1: GLMs

- ▶ Generalized Linear models
 - ▶ Defines a class of regression models for outcomes from the exponential family of distributions
 - ▶ Exponential family includes: Normal, Bernoulli/Binomial, Poisson, Gamma, Beta, among others
- ▶ Requires specification of three components:

Random component: $Y_i \sim \text{distribution}$
mean = μ_i
variance

Systematic component:
 $g(\mu_i) = \beta_0 + \beta_1 X_{i1} + \dots + \beta_p X_{ip}$

Link function: function g that maps the mean μ_i to the linear function of covariates

$$g(\mu_i) = X_i' \beta$$

$$g^{-1}(X_i' \beta) = \mu_i$$

Review of Lecture 1: GLMs

► Linear Model

$$Y_i \sim \mathcal{N}(\mu_i, \sigma^2)$$

$$g(\mu_i) = \mu_i = \beta_0 + \beta_1 X_{i1} + \dots + \beta_p X_{ip}$$

identity link \Rightarrow canonical link
but you could use others

► Logistic Model

$$Y_i = \begin{cases} 0 \\ 1 \end{cases}$$

$$Y_i \sim \text{Bernoulli}(\mu_i)$$

$$\text{mean} = \mu_i = \Pr(Y_i = 1)$$

$$\text{variance} = \mu_i(1 - \mu_i)$$

$$g(\mu_i) = \log \left[\frac{\mu_i}{1 - \mu_i} \right] = \beta_0 + \beta_1 X_{i1} + \dots + \beta_p X_{ip} = X_i' \beta$$

$$\mu_i = g^{-1}(X_i' \beta) = \frac{\exp(X_i' \beta)}{1 + \exp(X_i' \beta)}$$



Review of Lecture 1: Key quantities for simple logistic regression

- Assume your outcome is Y taking values 0 vs. 1 and your primary exposure variable X is also binary taking values 0 vs. 1.

- Mean: $\mu_i = E(Y_i) = \Pr(Y_i = 1 | X_i)$
- Odds: $\text{odds}[\Pr(Y_i = 1 | X_i)] = \frac{\Pr(Y_i = 1 | X_i)}{\Pr(Y_i = 0 | X_i)} = \beta_0 + \beta_1 X_i$

Logit:

- Odds ratio: $\text{Logit}[\Pr(Y_i = 1 | X_i)] = \log[\text{odds}[\Pr(Y_i = 1 | X_i)]]$

$$\frac{\text{odds}[\Pr(Y_i = 1 | X_i = 1)]}{\text{odds}[\Pr(Y_i = 1 | X_i = 0)]} = \frac{\Pr(Y_i = 1 | X_i = 1) / \Pr(Y_i = 0 | X_i = 1)}{\Pr(Y_i = 1 | X_i = 0) / \Pr(Y_i = 0 | X_i = 0)}$$

Review of Lecture 1 + additional models

- ▶ In this lecture we will consider 4 logistic regression models:

▶ Model A: $\text{Logit} [\Pr(Y_i=1)] = \beta_0$

▶ Model B: $\text{Logit} [\Pr(Y_i=1 | X_i)] = \beta_0 + \beta_1 X_i$

▶ Model C: $\text{Logit} [\Pr(Y_i=1 | X_i, Z_i)] = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \beta_3 X_i Z_i$

▶ Model D: $\text{Logit} [\Pr(Y_i=1 | X_i, Z_i)] = \beta_0 + \beta_1 X_i + \beta_2 Z_i$

Y_i = Big expenditure
 X_i = mscd
 Z_i = old: Age 765

- ▶ We will fit models B through D and a few additional models as well.



Revisit Model B

```
## Create the necessary variables:  
data$posexp=ifelse(data$totalexp>0,1,0)  
data$mscd=ifelse(data$l1c5+data$chd5>0,1,0)  
data1=data[!is.na(data$eversmk),]  
data1$older=ifelse(data1$lastage<65,0,1)  
data1$bigexp=ifelse(data1$totalexp>1000,1,0)
```

```
## Model B  
modelB = glm(bigexp~mscd,data=data1,family="binomial")  
lincom(modelB,c("(Intercept)","mscd"))
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## (Intercept)	-0.7395315	-0.7806967	-0.6983663	1239.792	1.372226e-271
## mscd	1.825045	1.694177	1.955913	747.095	1.718138e-164

```
lincom(modelB,c("(Intercept)","mscd"),eform=TRUE)
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## (Intercept)	0.4773375	0.4580868	0.4973973	1239.792	1.372226e-271
## mscd	6.203076	5.442166	7.070374	747.095	1.718138e-164

$$\text{Logit} [Pr(\text{Bigexp}=1|\text{mscd})] \\ = \beta_0 + \beta_1 \text{mscd}$$

$$\hat{\beta}_0, \hat{\beta}_1$$

$$\exp(\hat{\beta}_0) \exp(\hat{\beta}_1)$$

Revisit Model B

```
lincom(modelB,c("(Intercept)","mscd"))
```

Logit $[Pr(0_{\text{exp}}=1 | \text{mscd})]$

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## (Intercept)	-0.7395315	-0.7806967	-0.6983663	1239.792	1.372226e-271
## mscd	1.825045	1.694177	1.955913	747.095	1.718138e-164

$= \beta_0 + \beta_1 \text{mscd}$

```
lincom(modelB,c("(Intercept)","mscd"),eform=TRUE)
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## (Intercept)	0.4773375	0.4580868	0.4973973	1239.792	1.372226e-271
## mscd	6.203076	5.442166	7.070374	747.095	1.718138e-164

$Pr(1_{\text{exp}}=1 | \text{mscd}=0)$
 $= \frac{e^{.48}}{1 + e^{.48}}$

- Interpret beta_0 and exp(beta_0)

$\beta_0 = \log \text{odds of a big expenditure among persons without a major smoking related disease.}$

$\exp(\beta_0) = \text{odds of a Big exp among mscd} = 0 = .48$

- Interpret beta_1 and exp(beta_1)

$\beta_1 = \text{Diff in log odds of a Big exp comparing those with and without a mscd}$

$\hat{\beta}_1 = 1.83$
 $= \log \text{odds ratio}$

$\exp(\beta_1) = \text{relative odds of a Big exp comparing those with and without a mscd}$
 $\exp(\hat{\beta}_1) = 6.2$

Interpretation of $\exp(\hat{\beta}_{mscd})$

Two interpretations:

$$\exp(\hat{\beta}) = 6.2$$

- 1) The odds of a Big expenditure among persons with a MSCD are 6.2 times the odds of a Big expenditure among persons without a MSCD

$$6.2 = \frac{\text{odds}[\text{Pr}(\text{Big exp} = 1 | \text{MSCD} = 1)]}{\text{odds}[\text{Pr}(\text{Big exp} = 1 | \text{MSCD} = 0)]}$$

$$\text{odds}[\text{Pr}(\text{Big exp} = 1 | \text{MSCD} = 1)] = 6.2 \times \text{odds}[\text{Pr}(\text{Big exp} = 1 | \text{MSCD} = 0)]$$

- 2) The odds of a Big expenditure among persons with a MSCD are 520% greater than the odds among persons without a MSCD

Revisit Model C

Logit $[Pr(Bisexp=1 | mscd, older)] = \beta_0 + \beta_1 mscd + \beta_2 older + \beta_3 mscd \times older$

```
## Model C
modelC = glm(bigexp~mscd+older+mscd:older,data=data1,family="binomial")
lincom(modelC,c("mscd","mscd+mscd:older","mscd:older"))
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	1.969895	1.735287	2.204503	270.8301	7.481555e-61
## mscd+mscd:older	1.491115	1.329415	1.652815	326.6619	5.126712e-73
## mscd:older	-0.4787796	-0.7637143	-0.193845	10.84618	0.0009899951

lincom(modelC,c("mscd","mscd+mscd:older","mscd:older"),eform=TRUE)

exp(β_1)

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	7.169921	5.670554	9.065741	270.8301	7.481555e-61
## mscd+mscd:older	4.442046	3.778832	5.221658	326.6619	5.126712e-73
## mscd:older	0.619539	0.4659326	0.8237856	10.84618	0.0009899951

Among persons 65 yrs old or younger, the odds of Bis\$ for those with a mscd are 7.17 times the odds for those without a mscd.

Revisit Model C

Model C

```
modelC = glm(bigexp~mscd+older+mscd:older,data=data1,family="binomial")
lincom(modelC,c("mscd","mscd+mscd:older","mscd:older"))
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
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## mscd:older	-0.4787796	-0.7637143	-0.193845	10.84618	0.000989951

```
lincom(modelC,c("mscd","mscd+mscd:older","mscd:older"),eform=TRUE)
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	7.169921	5.670554	9.065741	270.8301	7.481555e-61
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## mscd:older	0.619539	0.4659326	0.8237856	10.84618	0.000989951

$\exp(\beta_1)$
 $\exp(\beta_1 + \beta_3)$ Among persons older 65 years of age,
 the odds of a Big 5 for those with a mscd are
 4.44 times the odds for those without a mscd.

The relative odds of having a Big 5 expenditure for those w and w/o
 a mscd are 38% smaller among person over 65
 compared to person 65 or younger

older = 1

$(\beta_0 + \beta_2) +$

$(\beta_1 + \beta_3) \text{ mscd}$

older OR 4.44

younger OR 7.17

$\frac{4.44}{7.17} = 0.62$

Model D: Parameter interpretation and estimation

- ▶ Model D: $\text{logit} [\text{Pr}(\text{Big exp} = 1 \mid \text{MSCD}, \text{older})] = \beta_0 + \beta_1 \text{MSCD} + \beta_2 \text{older}$
- ▶ How do you interpret coefficient for MSCD?
log odds ratio for a big exp company, those w and w/o a MSCD
- ▶ How do we estimate this coefficient?
among persons of the same age group (younger vs older)
- ▶ Inverse-variance weighting estimation!
- ▶ Same as linear regression!
- ▶ Need to consider the age (young vs. old) specific 2x2 tables.

model C

	Young		Old	
	MSCD = 1	MSCD = 0	MSCD = 1	MSCD = 0
Bigexp = 1	273	1802	713	1547
Bigexp = 0	101	4780	232	2236

$$\log OR = \log \left[\frac{273 \cdot 4780}{101 \cdot 1802} \right]$$

$$\text{Var}(\log OR) = \frac{1}{273} + \frac{1}{4780} + \frac{1}{101} + \frac{1}{1802}$$

$$\log \left[\frac{713 \cdot 2236}{232 \cdot 1547} \right]$$

$$\frac{1}{713} + \frac{1}{2236} + \frac{1}{232} + \frac{1}{1547}$$

Model D: Parameter interpretation and estimation

Age group	$\log \hat{O}R$	$se(\log \hat{O}R)$	$var(\log \hat{O}R)$	$\frac{1}{var(\log \hat{O}R)}$	$w = \frac{\frac{1}{var(\log \hat{O}R)}}{\sum(\frac{1}{var(\log \hat{O}R)})}$
Younger	<u>1.97</u>	0.12	<u>0.0144</u>	69.4	0.32
Older	<u>1.49</u>	0.083	<u>0.0069</u>	<u>144.9</u>	0.68

$\hat{\beta}_1 = 1.97 \times 0.32 + 1.49 \times 0.68 = 1.64$

Handwritten calculations: $69.4 \times 0.32 = 214.3$, $144.9 \times 0.68 = 98.5$, $214.3 + 98.5 = 312.8$ (Note: The handwritten sum is 312.8, but the text says 214.3, which is likely a typo for 312.8).

$$se(\hat{\beta}_1) = \frac{1}{\sqrt{\sum(\frac{1}{var(\log \hat{O}R)})}} = \frac{1}{\sqrt{214.3}} = \underline{0.068}$$

Model D: Parameter interpretation and estimation

```
modelD = glm(bigexp~mscd+older,data=data1,family="binomial")
summary(modelD)$coeff
```

```
##              Estimate Std. Error   z value    Pr(>|z|)
## (Intercept) -0.9577826 0.02700779 -35.46321 1.81555e-275
## mscd         1.6549130 0.06803662  24.32386 1.096494e-130
## older        0.5638298 0.04104938  13.73540 6.230701e-43
```

```
lincom(modelD,c("mscd","older"),eform=TRUE)
```

```
##      Estimate 2.5 %   97.5 %   Chisq    Pr(>Chisq)
## mscd 5.232625 4.57938  5.979054 591.65   1.096494e-130
## older 1.75739 1.621537 1.904625 188.6613 6.230701e-43
```

You practice: Use the output above, interpret $\exp(\hat{\beta}_2)$.

Among persons with similar disease status, the relative odds of a Big expenditure comparing older to younger persons is 1.76.

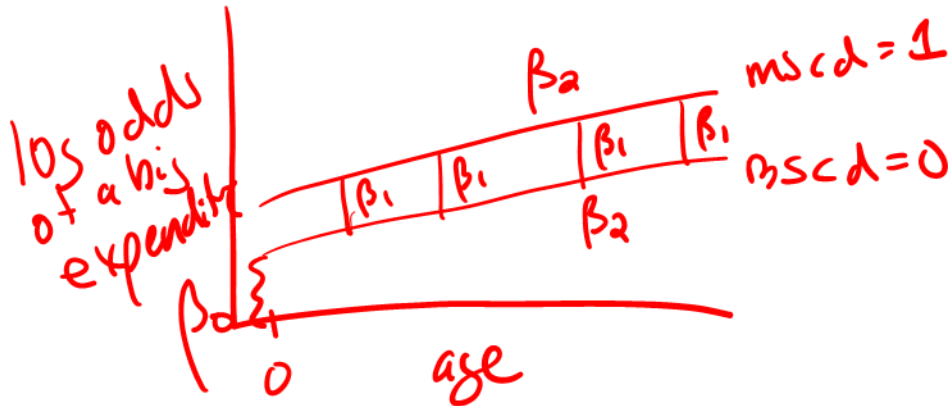
" , the odds of a Big \$ for older persons are 76% greater than the odds for younger person

Among persons of the same age group, the relative odds of a Big expenditure comparing those w and w/o a mscd is 5.23

Model D: Adjustment for continuous covariates

- Now, imagine Model D but where we allow age to be a continuous variable

- Model D with continuous age: $\text{logit} [\Pr(B_i, \text{exp} = 1 \mid \text{mscd}, \text{age})]$
 $= \beta_0 + \beta_1 \text{mscd} + \beta_2 \text{age}$
- Can you draw a picture of this model?



Model D: Adjustment for continuous covariates

```
modelDagecont = glm(bigexp~mscd+lastage,data=data1,family="binomial")  
summary(modelDagecont)$coeff
```

	β_0	Estimate	Std. Error	z value	Pr(> z)
## (Intercept)		-2.27990966	0.099135981	-22.99780	4.903428e-117
## mscd	β_1	1.60502065	0.068269770	23.50998	3.224831e-122
## lastage	β_2	0.02574057	0.001599682	16.09105	2.947835e-58

```
lincom(modelDagecont,c("mscd","lastage"),eform=TRUE)
```

	$\exp(\beta_1)$	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd		<u>4.977962</u>	4.35452	5.690664	552.719	3.224831e-122
## lastage		1.026075	<u>1.022863</u>	<u>1.029297</u>	258.922	2.947835e-58

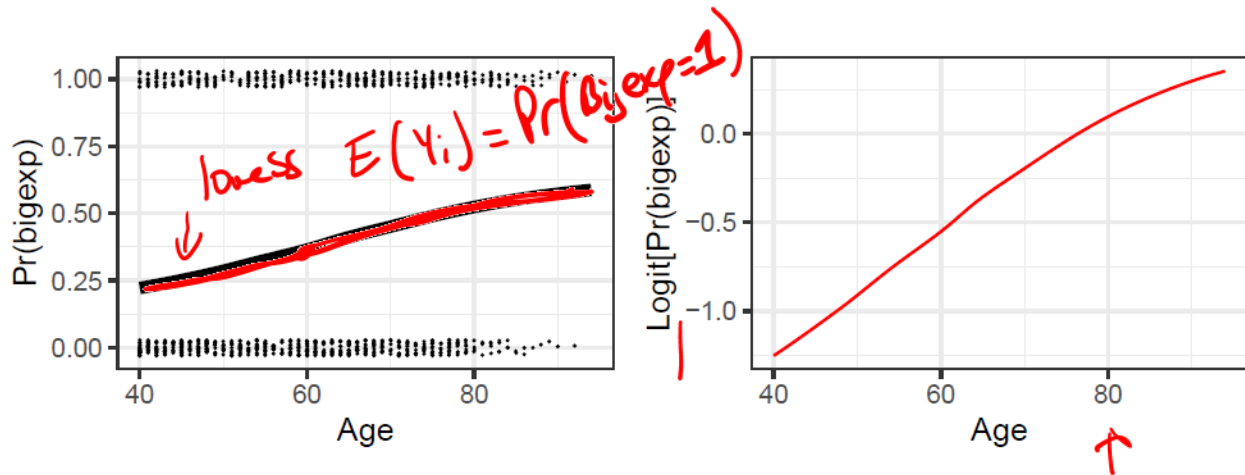
- Interpret both of the coefficients:

Among persons with the same disease status, the odds of a Big expenditure increase by 2.6% per a additional year of age

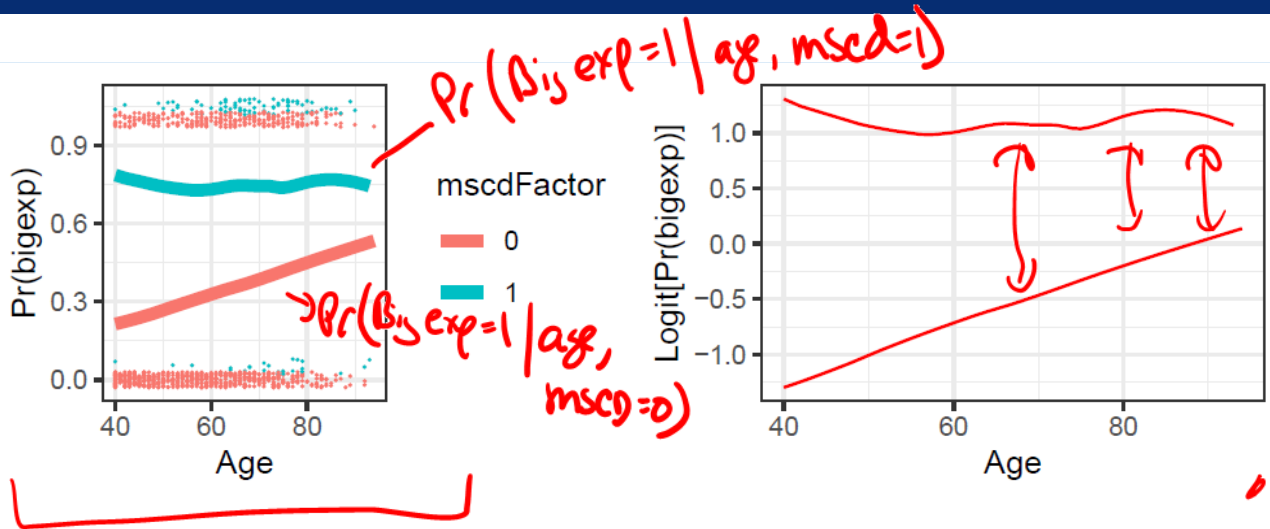
Among persons of the same age, the relative odds of a Big \$ Company those w and w/o a mscd is 4.98

Assessing functional form for continuous covariates

- ▶ How do we know if the relationship between the logit of a big expenditure and age is linear?



Revisit interaction Model C with continuous age



- ▶ What do you think about the MSCD-specific relationship between a big expenditure and age?
 - ▶ Linear? Non-linear?

Revision interaction Model C with continuous age

- Assuming the linear assumption is okay!

```
data1$age_c = data1$lastage - 60
```

```
modelCcont = glm(bigexp~mscd+age_c+mscd:age_c,data=data1,family="binomial")
lincom(modelCcont,c("mscd","mscd+mscd:age_c","mscd+20*mscd:age_c","mscd:age_c"))
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	1.792367	1.625218	1.959516	441.7169	4.579093e-98
## mscd+mscd:age_c	1.768144	1.607967	1.928321	468.0905	8.342527e-104
## mscd+20*mscd:age_c	1.307903	1.113342	1.502464	173.595	1.213445e-39
## mscd:age_c	-0.0242232	-0.03631431	-0.01213209	15.41797	8.616514e-05

lincom(modelCcont,c("mscd","mscd+mscd:age_c","mscd+20*mscd:age_c","mscd:age_c"),eform=TRUE)

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	6.003646	5.079528	7.09589	441.7169	4.579093e-98
## mscd+mscd:age_c	5.859966	4.992649	6.877953	468.0905	8.342527e-104
## mscd+20*mscd:age_c	3.69841	3.044517	4.492744	173.595	1.213445e-39
## mscd:age_c	0.9760678	0.9643371	0.9879412	15.41797	8.616514e-05

Among 60 year olds, relative odds of Big exp comparing those w and w/o a mscd is 6.00
 * among 61 year olds, OR is 5.86

$$\text{Logit}[\text{Pr}(\text{Big exp} = 1 \mid \text{mscd}, (\text{age}-60))] = \beta_0 + \beta_1 \text{mscd} + \beta_2 (\text{age}-60) + \beta_3 (\text{age}-60) \times \text{mscd}$$

$$\frac{5.86}{6.00} = .976$$

Revision interaction Model C with continuous age

- ▶ Assuming the linear assumption is okay!

```
data1$age_c = data1$lastage - 60
```

```
modelCcont = glm(bigexp~mscd+age_c+mscd:age_c,data=data1,family="binomial")  
lincom(modelCcont,c("mscd","mscd+mscd:age_c","mscd+20*mscd:age_c","mscd:age_c"))
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	1.792367	1.625218	1.959516	441.7169	4.579093e-98
## mscd+mscd:age_c	1.768144	1.607967	1.928321	468.0905	8.342527e-104
## mscd+20*mscd:age_c	1.307903	1.113342	1.502464	173.595	1.213445e-39
## mscd:age_c	-0.0242232	-0.03631431	-0.01213209	15.41797	8.616514e-05

```
lincom(modelCcont,c("mscd","mscd+mscd:age_c","mscd+20*mscd:age_c","mscd:age_c"),eform=TRUE)
```

	Estimate	2.5 %	97.5 %	Chisq	Pr(>Chisq)
## mscd	6.003646	5.079528	7.09589	441.7169	4.579093e-98
## mscd+mscd:age_c	5.859966	4.992649	6.877953	468.0905	8.342527e-104
## mscd+20*mscd:age_c	3.69841	3.044517	4.492744	173.595	1.213445e-39
## mscd:age_c	0.9760678	0.9643371	0.9879412	15.41797	8.616514e-05

$\beta_1 + 20\beta_3$

Among 80 yr olds, the OR is 3.70

Revision interaction Model C with continuous age

- How would you rewrite the `lincom` commands to get estimates of the relationship between having a big expenditure and age, separately for those with and without a MSCD?

```
modelCcont = glm(bigexp~mscd+age_c+mscd:age_c,data=data1,family="binomial")  
lincom(modelCcont,c("mscd","mscd+mscd:age_c","mscd+20*mscd:age_c","mscd:age_c"))
```

$$\begin{aligned} \text{logit} [Pr(\text{Big exp}=1 \mid \text{mscd}, \text{age}-60)] \\ = \beta_0 + \beta_1 \text{mscd} + \beta_2 (\text{age}-60) + \underbrace{\beta_3 (\text{age}-60) \text{mscd}} \end{aligned}$$

`lincom(modelCcont, c("age-c", "age-c + mscd:age-c"), eform=TRUE)`

Where to next?

- ▶ Assessing for confounding in logistic regression models
 - ▶ See “Note on confounding and effect modification 2019” by Scott Zeger
 - ▶ Additional references are provided in Lecture 2 Handout
- ▶ Statistical inference in logistic regression models
 - ▶ Maximum likelihood estimation
 - ▶ Iteratively reweighted least squares