

Lecture 4

Lecture 4 Outline

- Focus on separating individual-level and cluster-level covariate effects in multi-level models
 - This continues our discussion from Lecture 3 where we utilized graphical displays to separate the within and between cluster effects for a level-1 covariate
 - Focus will be on estimation and interpretation of fixed effects within a random intercept model (for now)
 - Begg and Parides 2003 Statistics in Medicine paper
- We will also consider the case where the goal is to estimate the effect of a Level-2 covariate while adjusting for level-1 covariates.
- Example: Inner London School Data
 - Students (level 1) nested within schools (level 2)
 - OUTCOME: GCSE score, age 16 exam score
 - Level-1 Covariate: LRT score, age 11 reading test score

Scientific Questions

For now, focus on fixed effects; i.e. interpretation of regression coefficients within mixed model (not interpretation of variance of random effects).

1. Quantify the relationship between GCSE score and LRT score
2. Within a school, quantify the relationship between GCSE score and LRT score
3. Does the “context” of the school matter? i.e. do students from schools with higher school-average LRT scores fair better than otherwise similar students in schools with lower school-average LRT scores
 - Defined as the “contextual” effect (see “Brief conceptual...”

Exploratory Data Analysis (continued)

* Create some new variables

```
sort school student
```

* Generate the number of students within each school
by school: egen totalstudents = count(student)

* Generate a counter for the number of students within each school
by school: gen withinschoolcount = _n

* EDA

* What is the distribution of number of students in each school
summ totalstudents if withinschoolcount==1

Variable	Obs	Mean	Std. Dev.	Min	Max
totalstude~s	65	62.44615	29.74844	2	198

65 schools in the dataset:

Number of students ranges from 2 to 198, average 62

Exploratory Data Analysis

* What is the distribution of the gcse and lrt scores

```
. summ gcse
```

Variable	Obs	Mean	Std. Dev.	Min	Max
gcse	4059	69.99527	9.977929	33.339	100

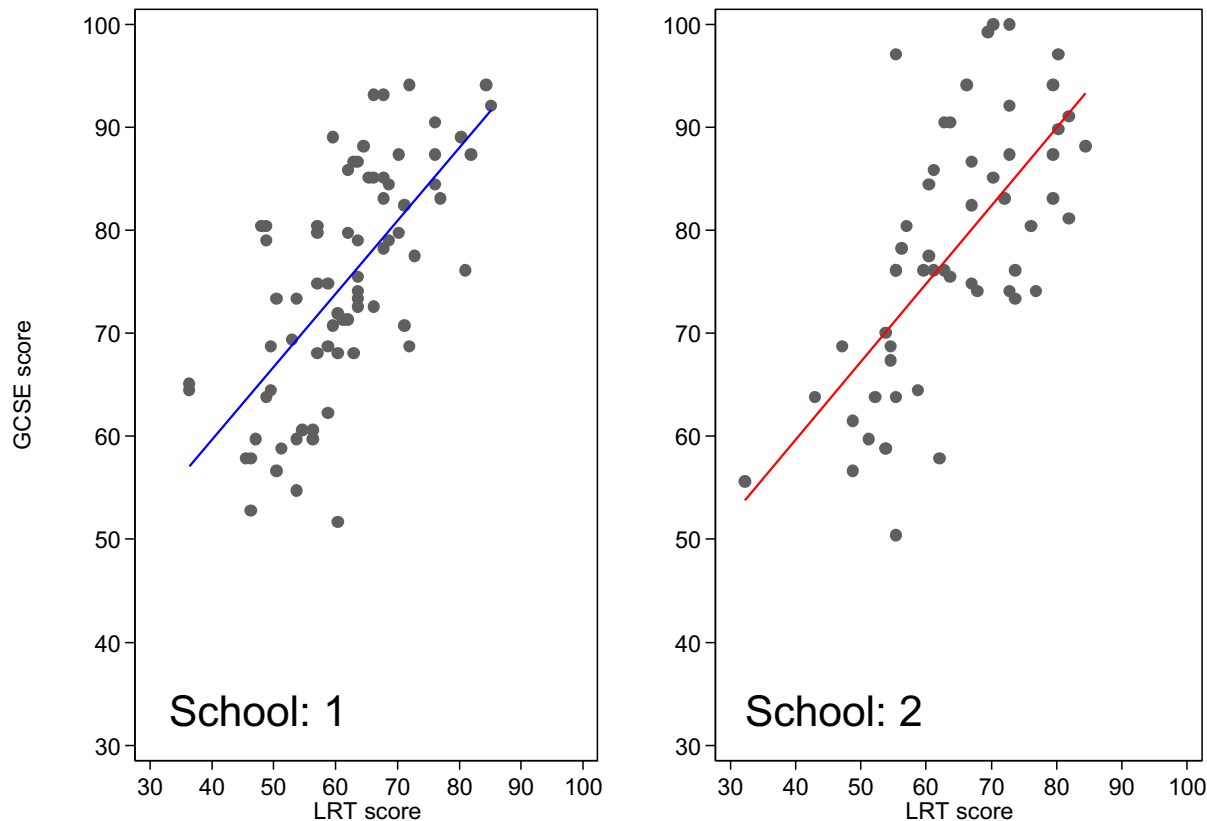
```
. summ lrt
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lrt	4059	60.0181	9.93223	30.65	90.16

NOTE: This data was extracted from your textbook. The data is provided as z-scores with standard deviation of 10 instead of 1. I added 60 to the LRT scores and 70 to the GCSE scores. So we will interpret the data in the lecture as the raw test score.

Exploratory Data Analysis

- Relationship between gcse and lrt among two schools (1 and 2)



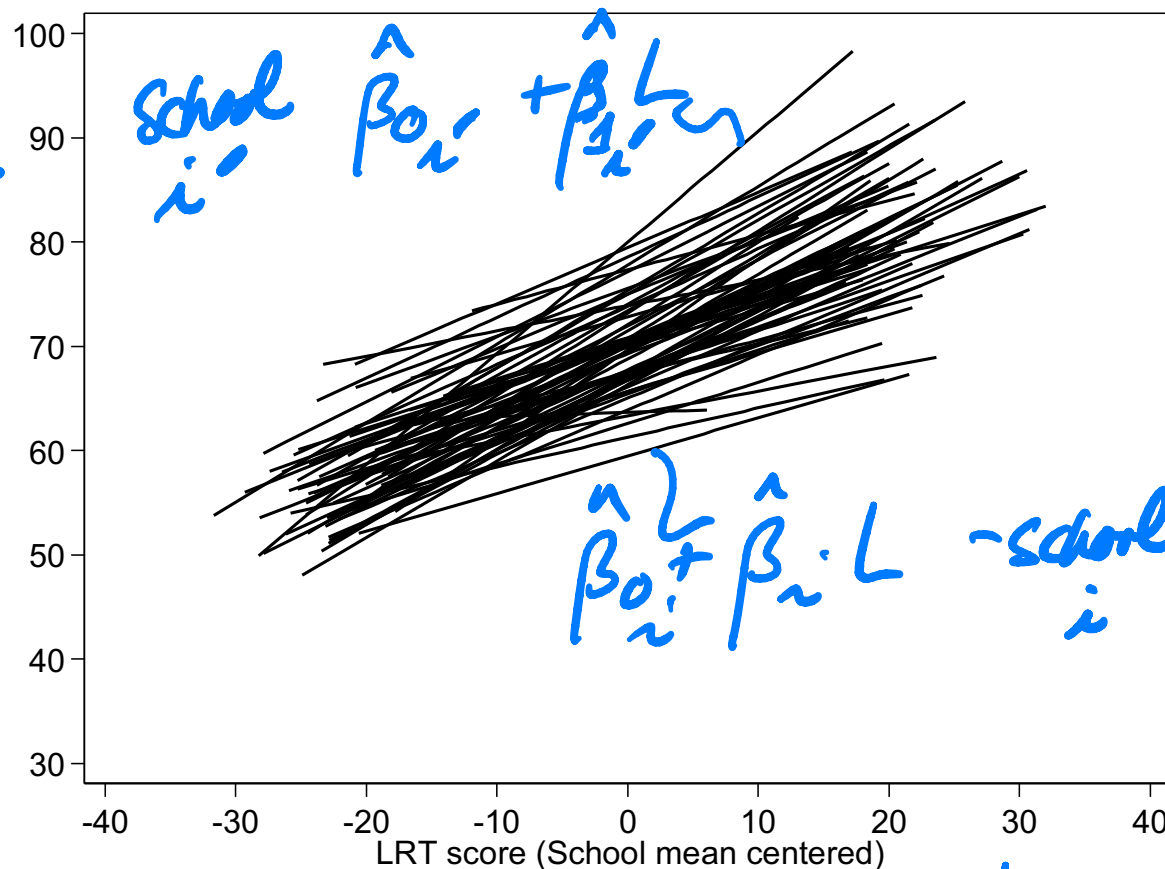
Data by school: $(L_{ij}, G_{ij}, j=1, \dots, n_i)$
Regress G_{ij} on L_{ij} for each i

$$G_{ij} = \beta_{0i} + \beta_{1i} L_{ij} + \epsilon_{ij}, \quad j = 1, \dots, n_i$$

$$\Rightarrow (\hat{\beta}_{0i}, \hat{\beta}_{1i}) \quad (\text{and } \text{Var}(\hat{\beta}_{0i}, \hat{\beta}_{1i}) = V_i)$$

$i = 1, \dots, m$

School-specific relationships among schools with at least 5 students



- One of our objectives is to estimate the within cluster association.

- There appears to be variation in this association across schools ?

- Lecture 6

- For now, we want to estimate the average of these slopes!

Possible Models to Consider

G L

$$\text{Model 1: } E(Y_{ij}|X_{ij}) = \alpha_{1i} + \beta_1 X_{ij}$$

$$\text{Model 2: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{2i} + \beta_2 X_{ij} + \gamma_2 \bar{X}_{i.}$$

$$\text{Model 3: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{3i} + \beta_3 (X_{ij} - \bar{X}_{i.}) + \gamma_3 \bar{X}_{i.}$$

$$\text{Model 4: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{4i} + \beta_4 (X_{ij} - \bar{X}_{i.})$$

$$\text{Model 5: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{5i} + \gamma_5 \bar{X}_{i.}$$

where

$$Y_{ij} = E(Y_{ij}|\dots) + \varepsilon_{ij}, \varepsilon_{ij} \sim N(0, \sigma_k^2)$$

$$\alpha_{ki} \sim N(\alpha_k, \tau_k^2), \text{ for each model } k = 1, \dots, 5$$

Notes on the possible models

Model 1: $E(Y_{ij}|X_{ij}) = \alpha_{1i} + \beta_1 X_{ij}$

- Ignores the clustering of the data when it is estimating β_1 , this is the “total effect”
✗ - does allow for heterogeneity in intercepts.

Model 2: $E(Y_{ij}|X_{ij}, \bar{X}_i) = \alpha_{2i} + \beta_2 X_{ij} + \gamma_2 \bar{X}_i$

Model 3: $E(Y_{ij}|X_{ij}, \bar{X}_i) = \alpha_{3i} + \beta_3 (X_{ij} - \bar{X}_i) + \gamma_3 \bar{X}_i$

- These two models are mathematically equivalent *values for $E(Y_{ij})$*
- In Model 3, we have chosen to “center” the level-1 covariate
 - “cluster-mean” or “cluster-specific” or “within-cluster” centering
- The choice of centering or not will change the interpretation of the intercept and also the interpretation of γ_2 and γ_3
- We will also discuss the option to “grand-mean center” later in the lecture and the impact this can have.

Notes on the possible models

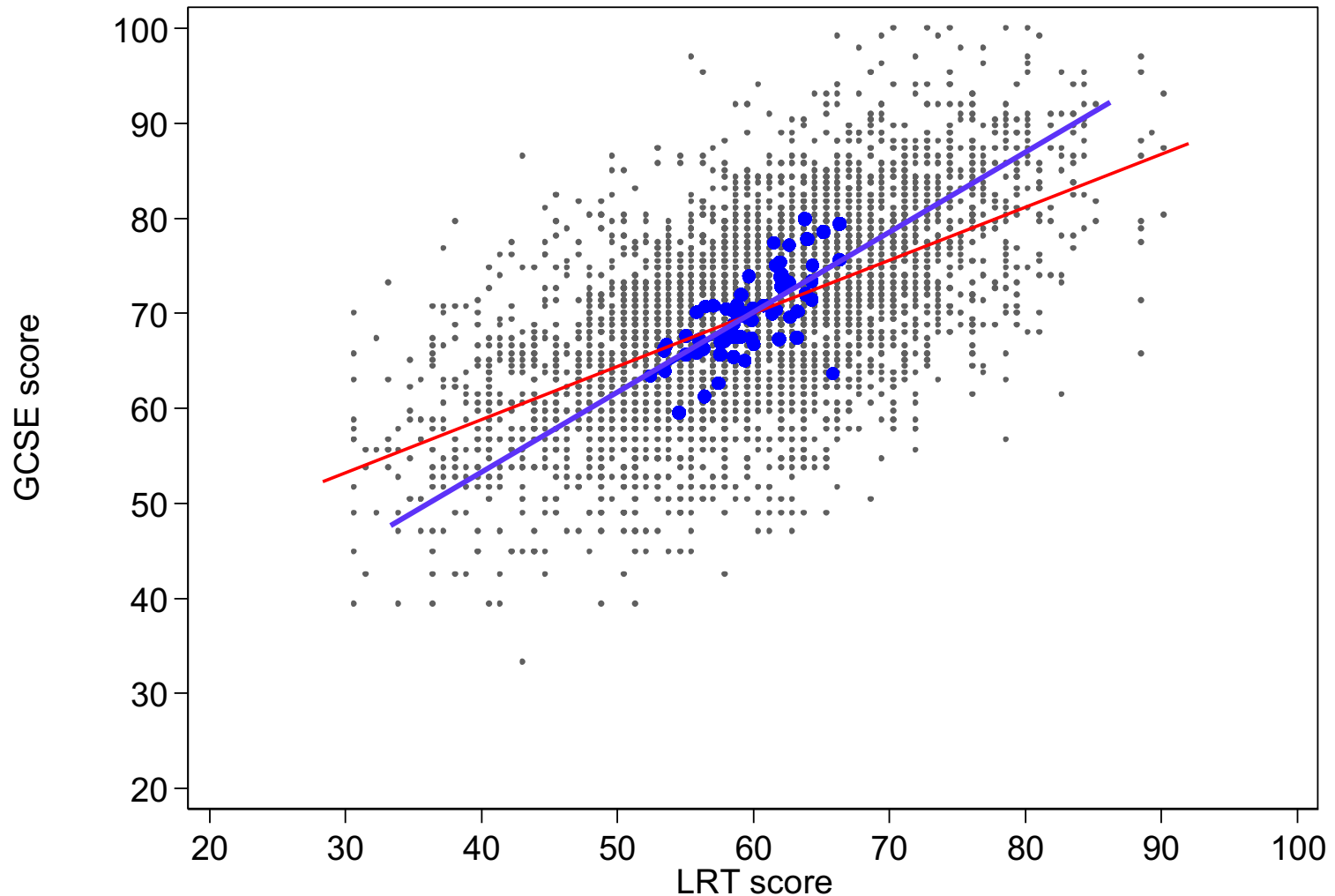
Model 4: $E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{4i} + \beta_4(X_{ij} - \bar{X}_{i.}) + 0 \cdot \bar{X}_{i.}$

- This model ignores the cluster mean covariate
- Includes only the cluster-mean centered level-1 covariate

Model 5: $E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{5i} + \gamma_5 \bar{X}_{i.} + 0 \cdot (X_{ij} - \bar{X}_{i.})$

- This model ignores the cluster-mean centered level-1 covariate
- Includes only the cluster-mean covariate

Recall the EDA: Separation of Between and Within Effects



Fit the models using (xt)mixed in Stata

- * Generate the cluster-mean variable and cluster-mean centered

- * lrt score

```
bys school: egen mean_lrt = mean(lrt)
```

```
gen lrt_within = lrt - mean_lrt
```

***** Model 1

```
mixed gcse lrt || school:
```

***** Model 2

```
mixed gcse lrt mean_lrt || school:
```

***** Model 3

```
mixed gcse lrt_within mean_lrt || school:
```

***** Model 4

```
mixed gcse lrt_within || school:
```

***** Model 5

```
mixed gcse mean_lrt || school:
```

Results

Model	β_k	γ_k
1	0.563 (0.538, 0.587)	
2	0.559 (0.534, 0.583)	0.357 (0.142, 0.573)
3	0.559 (0.534, 0.583)	0.916 (0.702, 1.131)
4	0.559 (0.534, 0.583)	
5		0.925 (0.712, 1.138)

We note first the difference in estimates in the β_k column.

- The estimate from Model 1 (0.563) is the total effect
- This estimate basically ignores the cluster membership
- This estimate is a distorted estimate of the within cluster effect due to confounding by the cluster-average LRT
 - Can show that cluster-average LRT is correlated with both individual LRT and the GCSE score.

Interpretation of Models 2 and 3

$$\text{Model 2: } E(Y_{ij} | X_{ij}, \bar{X}_{i.}) = \alpha_{2i} + \beta_2 X_{ij} + \gamma_2 \bar{X}_{i.}$$

$$\text{Model 3: } E(Y_{ij} | X_{ij}, \bar{X}_{i.}) = \alpha_{3i} + \beta_3 (X_{ij} - \bar{X}_{i.}) + \gamma_3 \bar{X}_{i.}$$

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2	0.559 (0.534, 0.583)	0.357 (0.142, 0.573)
3	0.559 (0.534, 0.583)	0.916 (0.702, 1.131)

β_2 and β_3 have the same estimated values and interpretation.

The effect of LRT score on GCSE within a given cluster:

Within a school, the student's average GCSE scores differ by 0.559 points per additional point on the LRT.

Interpretation of Models 2 and 3

$$\text{Model 2: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{2i} + \beta_2 X_{ij} + \gamma_2 \bar{X}_{i.}$$

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γ_2 and γ_3 have different values, what do they represent?

Model 2: γ_2 represents the contextual effect!

Model 3: γ_3 represents the between effect!
(among) school

Interpretation of Models 2 and 3

$$\text{Model 2: } E(Y_{ij} | X_{ij}, \bar{X}_i) = \alpha_{2i} + \beta_2 X_{ij} + \gamma_2 \bar{X}_i.$$

Model	β_k	γ_k
2	0.559 (0.534, 0.583)	0.357 (0.142, 0.573)

Model 2: Holding X_{ij} fixed, the mean difference in Y_{ij} per unit increase in \bar{X}_i .

Consider two students with the same LRT score but who come from schools that differ in school average LRT score by 1 point.

The student from the school with higher average LRT score is expected to have a GCSE score that is 0.357 points higher than the other student.

Interpretation of Models 2 and 3

$$\text{Model 3: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{3i} + \beta_3(X_{ij} - \bar{X}_{i.}) + \gamma_3\bar{X}_{i.}$$

Model	β_k	γ_k
3	0.559 (0.534, 0.583)	0.916 (0.702, 1.131)

Consider the full model:

$$Y_{ij} = \alpha_{3i} + \beta_3(X_{ij} - \bar{X}_{i.}) + \gamma_3\bar{X}_{i.} + \varepsilon_{3ij}$$

Take the cluster mean:

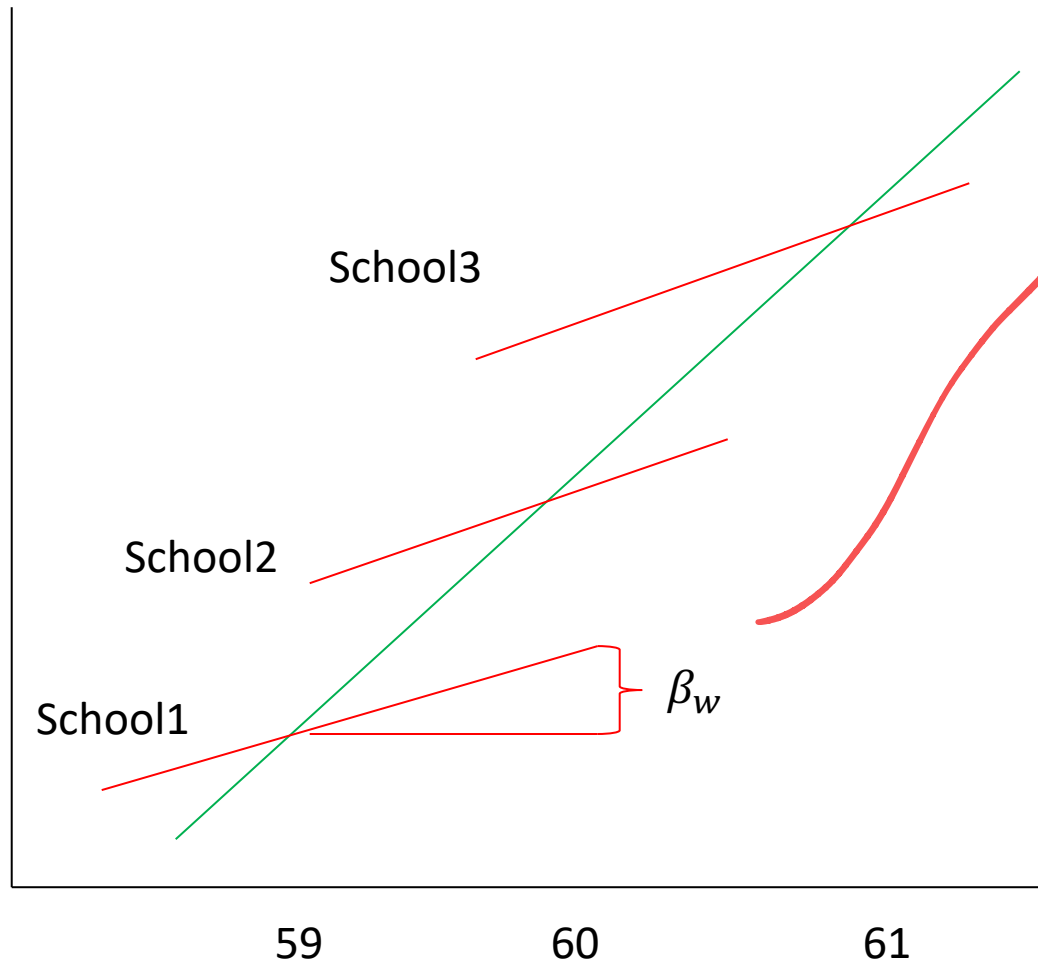
$$\begin{aligned}\bar{Y}_{i.} &= \alpha_{3i} + \beta_3(\bar{X}_{i.} - \bar{X}_{i.}) + \gamma_3\bar{X}_{i.} + \bar{\varepsilon}_{3i.} \\ \bar{Y}_{i.} &= \alpha_{3i} + \gamma_3\bar{X}_{i.} + \bar{\varepsilon}_{3i.}\end{aligned}$$

This is the between effect: The difference in school average GCSE per unit increase in school average LRT score.

Contextual Effects

- Why do these effects occur?
 - Normative effects associated with the cluster/organization/level-2 factor
 - i.e. persons within the cluster tend to be much more like each other than otherwise similar persons from other clusters
 - The mean X within a cluster may act as a proxy for other important cluster level characteristics that are not measured
 - They may signal a statistical artifact if the mean X within a cluster may carry some information if X is measured with error
- Posted a few references for with today's lecture that explore these ideas.

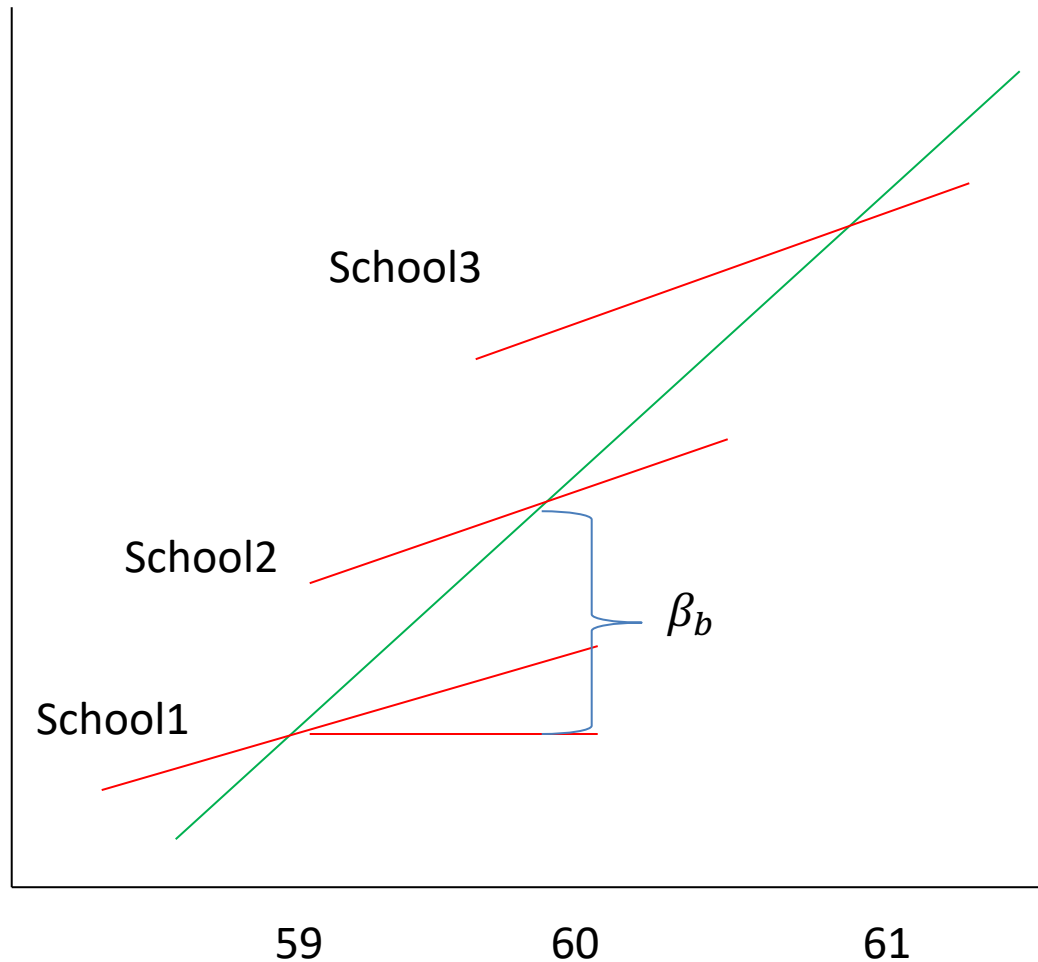
Contextual Effects



Within-effect: Expected difference in Y between two **subjects** from the same **cluster** but who differ in X by one unit

Inner London School EX: expected difference in GCSE score between two students from the same school but who differ in LRT by one point

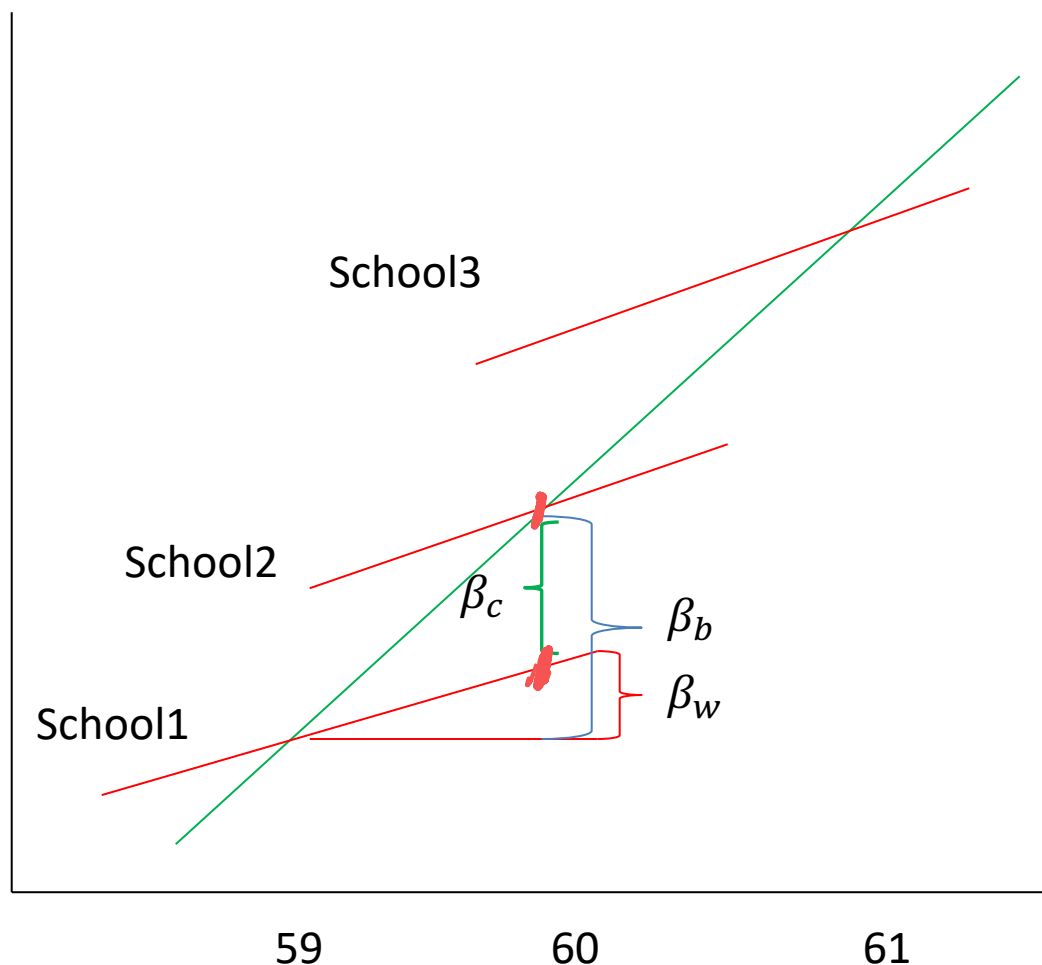
Contextual Effects



Between-effect: Expected difference in mean Y between two **clusters** that differ in average X by one unit

Inner London School EX: expected difference in mean GCSE score between two schools that differ in average LRT by one point

Contextual Effects



Contextual-effect:
Expected difference in Y
between two **subjects** who
have the same value of X
but who come from
clusters that differ by one
unit in mean X

Inner London School EX:
expected difference in
GCSE scores between two
students with the same
LRT score but who come
from two schools that
differ in average LRT by
one point

Interpretation of Models 2 and 3

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γ_2 is the contextual effect; estimated directly within Model 2

β_3 is the within cluster effect and γ_3 is the between cluster effect.

The contextual effect can be estimated within Model 3 by taking:

$$\gamma_2 = \gamma_3 - \beta_3$$

Interpretation of Models 4 and 5

$$\text{Model 3: } E(Y_{ij}|X_{ij}, \bar{X}_{i.}) = \alpha_{3i} + \beta_3(X_{ij} - \bar{X}_{i.}) + \gamma_3\bar{X}_{i.}$$

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5		0.925 (0.712, 1.138)

Models 3 and 4, β_k are both estimating the within effect.

Models 3 and 5, γ_k are both estimating the between effect.

What does this imply about $(X_{ij} - \bar{X}_{i.})$ and $\bar{X}_{i.}$?

Interpretation of Models 4 and 5

What does this imply about $(X_{ij} - \bar{X}_{i.})$ and $\bar{X}_{i.}$?

Level-2 covariates and cluster-mean centered level-1 covariates are independent!

```
corr gcse lrt lrt_within mean_lrt girl schgend  
(obs=4059)
```

		1	2	3	4	5	6
1	gcse	1.0000					
2	lrt	0.5917	1.0000				
3	lrt_within	0.5277	0.9484	1.0000			
4	mean_lrt	0.2879	0.3170	-0.0000	1.0000		
5	girl	0.1144	0.0532	0.0425	0.0407	1.0000	
6	schgend	0.1115	0.0067	-0.0000	0.0210	0.4365	1.0000

Estimating level-2 effects while adjusting for level-1 covariates

* "Compare like to like"

- One of the most common applications of MLM!
- Here it is commonly assumed that no contextual effect exists.
- In this setting, cluster-mean centering is not appropriate
- If you want to center the level-1 covariates, you should grand-mean center

Example: School Type

- Estimate the difference in school-average GCSE score across school-type adjusting for the composition of the schools
- Composition measured by gender and LRT score

Unadjusted model: $Y_{ij} = \beta_{0i} + \varepsilon_{ij}$

$$\beta_{0i} = \beta_0 + \beta_1 I(\text{boys}_i) + \beta_2 I(\text{girls}_i) + b_i$$

Adjusted models add main effects to the level-1 equation

$$\text{Adjusted 1: } Y_{ij} = \beta_{0i} + \alpha_{11} LRT_{ij} + \alpha_{21} \text{girl}_{ij} + \varepsilon_{ij}$$

$$\text{Adjusted 2: } Y_{ij} = \beta_{0i} + \alpha_{12} (LRT_{ij} - \overline{LRT}_{..}) + \alpha_{22} (\text{girl}_{ij} - \overline{\text{girl}}_{..}) + \varepsilon_{ij}$$

$$\text{Adjusted 3: } Y_{ij} = \beta_{0i} + \alpha_{13} (LRT_{ij} - \overline{LRT}_{i.}) + \alpha_{23} (\text{girl}_{ij} - \overline{\text{girl}}_{i.}) + \varepsilon_{ij}$$

Adjusted Model	All Boys (β_1)	All Girls (β_2)
Unadjusted	0.644 (-2.282, 3.569)	2.562 (0.275, 4.849)
1, main effects	1.176 (-0.392, 3.945)	1.575 (-0.133, 3.284)
2, grand mean centered	1.176 (-0.392, 3.945)	1.575 (-0.133, 3.284)
3, cluster mean centered	0.617 (-2.311, 3.546)	2.550 (0.255, 4.845)

- Unadjusted model:
 - The school average GCSE score is 0.644 points higher among all boys schools compared to mixed gender schools.
- Adjusted model:
 - After accounting for the composition of the school, the school average GCSE score is 1.176 points higher among all boys schools compared to mixed gender schools.

NOTE: The adjustment for the level-1 covariates occurs only in the models with no centering or with grand-mean centering.

Lecture 4 Summary

- In linear models, we considered 5 different models that allowed us to estimate
 - The total effect
 - Within-cluster effect
 - Among-cluster effect
 - Contextual effect
- We reviewed the interpretation of these effects within an example; Inner London School data
- We noticed that cluster-mean centered level-1 covariates are independent of level-2 covariates
 - Explored the impact of this observation for studies where our goal is to estimate the association between a level-1 outcome and level-2 covariate but adjusting for the composition of the clusters.

Lecture 5 Introduction

- Our example at the end of the lecture brings up an important topic for MLMs
 - CENTERING!
- In the next lecture, we will consider:
 - the choice of centering and how this choice can affect your solutions in MLMs
 - how the choice of centering can affect your estimates of random intercept and slope variances!