

# <sup>1</sup> Interactive Proofs For Distribution Testing With <sup>2</sup> Conditional Oracles

<sup>3</sup> Ari Biswas  

<sup>4</sup> University Of Warwick, United Kingdom

<sup>5</sup> Mark Bun<sup>1</sup>  

<sup>6</sup> Boston University

<sup>7</sup> Clément Canonne<sup>2</sup>  

<sup>8</sup> University of Sydney

<sup>9</sup> Satchit Sivakumar<sup>3</sup>  

<sup>10</sup> Boston University

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## <sup>11</sup> Abstract

<sup>12</sup> We revisit the framework of interactive proofs for distribution testing, first introduced by Chiesa  
<sup>13</sup> and Gur (ITCS 2018), which has recently experienced a surge in interest, accompanied by notable  
<sup>14</sup> progress (e.g., Herman and Rothblum, STOC 2022, FOCS 2023; Herman, RANDOM 2024). In this  
<sup>15</sup> model, a data-poor verifier determines whether a probability distribution has a property of interest  
<sup>16</sup> by interacting with an all-powerful, data-rich but untrusted prover bent on convincing them that  
<sup>17</sup> it has the property. While prior work gave sample-, time-, and communication-efficient protocols  
<sup>18</sup> for testing and estimating a range of distribution properties, they all suffer from an inherent issue:  
<sup>19</sup> for most interesting properties of distributions over a domain of size  $N$ , the verifier must draw at  
<sup>20</sup> least  $\Omega(\sqrt{N})$  samples of its own. While sublinear in  $N$ , this is still prohibitive for large domains  
<sup>21</sup> encountered in practice.

<sup>22</sup> In this work, we circumvent this limitation by augmenting the verifier with the ability to perform  
<sup>23</sup> an exponentially smaller number of more powerful (but reasonable) *pairwise conditional* queries,  
<sup>24</sup> effectively enabling them to perform “local comparison checks” of the prover’s claims. We system-  
<sup>25</sup> atically investigate the landscape of interactive proofs in this new setting, giving polylogarithmic  
<sup>26</sup> query and sample protocols for (tolerantly) testing all *label-invariant* properties, thus demonstrating  
<sup>27</sup> exponential savings without compromising on communication, for this large and fundamental class  
<sup>28</sup> of testing tasks.

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## <sup>36</sup> 1Introduction

<sup>37</sup> Distribution testing, as introduced by Batu et al. [2000], is a mature subfield of property  
<sup>38</sup> testing [Goldreich et al., 1998, Rubinfeld and Sudan, 1996] aimed at investigating statistical  
<sup>39</sup> properties of an unknown distribution given sample access to it. Given a property (a set of

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<sup>1</sup> Optional footnote, e.g. to mark corresponding author

<sup>2</sup> Optional footnote, e.g. to mark corresponding author

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40 distributions) and a proximity parameter  $\tau \in (0, 0.1]$ , distribution testing algorithms output  
 41 Accept if the distribution is in the property (or close to it), or Reject if the distribution is  
 42  $\tau$ -far from the property, both with high probability. Closeness and farness are quantified with  
 43 respect to a prespecified notion of distance, typically total variation distance. The primary  
 44 motivation behind distribution testing is to design testing algorithms for deciding properties  
 45 with sample complexity sub-linear in the domain size  $N$  (which is demonstrably more  
 46 efficient than learning the distribution, which requires drawing  $\Theta(N)$  samples). Accordingly,  
 47 over the last two decades, researchers have extensively studied the sample complexity of  
 48 numerous distribution properties, such as simple uniformity testing [Goldreich and Ron,  
 49 2011] (testing whether a distribution is uniform over its entire domain), support size decision  
 50 problem [Raskhodnikova et al., 2009, Valiant and Valiant, 2011, Wu and Yang, 2019, Ferreira  
 51 Pinto Jr. and Harms, 2025] (testing whether a distribution's support is within some pre-  
 52 specified range), and many more: see, e.g., [Goldreich, 2017, Chapter 11] and Rubinfeld [2012],  
 53 Canonne [2020, 2022] for a more thorough introduction to distribution testing. Unfortunately,  
 54 although distribution testing is often more efficient than learning the distribution, it is  
 55 still prohibitively expensive for practical use. For example, it is known that generalized  
 56 uniformity testing (testing whether a distribution is uniform over its support) over a domain  
 57 of size  $N$  requires  $\Omega(N^{2/3})$  samples [Batu and Canonne, 2017, Diakonikolas et al., 2018],  
 58 which can be impractical for large domain sizes. Even simple uniformity testing requires  
 59  $\Omega(\sqrt{N})$  samples [Paninski, 2008], and its *tolerant* testing version (which asks to distinguish  
 60 distributions *close* to uniform from those which are far) needs  $\Omega(N/\log N)$  samples [Valiant  
 61 and Valiant, 2017].

62 In the face of these limitations, a nascent line of work [Chiesa and Gur, 2018, Herman  
 63 and Rothblum, 2022, 2023, Herman, 2024] has asked a related question: *with testing being*  
 64 *hard by itself, what is the complexity of verifying the properties of a distribution given*  
 65 *sample access to it?* Here, in addition to drawing samples from the distribution, the tester is  
 66 allowed to interactively communicate with an omniscient but *untrusted* prover that knows  
 67 the distribution in its entirety. The idea here is to leverage the provers extra knowledge  
 68 about the distribution, with the hope that checking the provers' claims is easier than naively  
 69 testing the property. While this model of verifiable computation has only recently been  
 70 explored in the context of distribution testing, it has been an active area of research in other  
 71 areas of theoretical computer science for over 40 years (see for e.g. [Goldwasser et al., 1985,  
 72 Micali, 2000, Rothblum et al., 2013, Goldwasser et al., 2015, Berman et al., 2018, Arun  
 73 et al., 2024]). It models settings where a centralized organization (for example, a company  
 74 turning billions of dollars of profit) has the ability to collect large amounts of data and learn  
 75 distributions to high precision, while end-users may not have the same ability. At the same  
 76 time, the company might have incentives to lie, and so verifying whether the company is  
 77 being truthful is important in this setting. The work of Chiesa and Gur [2018] shows that the  
 78 verification of *any* distribution property over domain  $[N]$  can be reduced to identity testing<sup>4</sup>,  
 79 with communication *superlinear* in the domain size. Follow up work [Herman and Rothblum,  
 80 2022, 2023] recovers this result for the broad class of *label-invariant properties*, while only  
 81 requiring communication *sub-linear* in the domain size. More specifically, the work of Herman  
 82 and Rothblum [2022, 2023] show that for label-invariant properties, verification requires only  
 83  $O(\sqrt{N})$  samples and  $\tilde{O}(\sqrt{N})$  communication, even though, as mentioned earlier, testing some  
 84 properties in this class could require  $\Theta(N/\log N)$  samples. Here, a property is *label-invariant*

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<sup>4</sup> Identity testing refers to the task of testing if a distribution is exactly equal to a pre-specified reference distribution or is  $\tau$ -far from it.

(also known as *symmetric*) if the names of the elements themselves are not significant to the decision outcome (see Definition ??). Testing if a distribution is uniform over its support (also known as generalized uniformity testing) is an example of a label-invariant property.

Unfortunately, while a significant improvement over unaided testing, requiring  $O(\sqrt{N})$  samples from the verifier can still be prohibitive when considering massive domains. Further, there is a matching sample complexity lower bound – verification of even basic label-invariant properties such as checking if a distribution is uniform over its entire domain requires  $\Omega(\sqrt{N})$  samples. To summarise: For most properties, with access to *only* samples from a distribution, it is impossible for any tester to do better than drawing  $\Omega(\sqrt{N})$  samples, with or without the help of a prover. To bypass these limitations and develop more practical algorithms, in this work we study verifiers that can make a very small number of calls to a more powerful *conditional sampling* oracle. These oracles were introduced in the context of distribution testing [Chakraborty et al., 2013, Canonne et al., 2014]; allowing the tester to condition that samples from the oracle come from a subset  $S$  of the domain, of their choosing. The oracle responds with a sample with probability re-normalised over  $S$ . If no element in  $S$  is supported, the oracle responds with FAIL. Since specifying an arbitrary set may considered be unrealistic for practical purposes, a commonly studied restriction is the pairwise conditional sampling model (**PCond**), where the specified sets are restricted to be of size exactly 2 or the entire domain (thus, just a regular sample from the distribution). These oracles can be thought of as allowing for local comparisons between the probabilities of two points. While access to a **PCond** oracle can be significantly helpful for problems like simple uniformity testing, it is unclear from prior work whether it results in more efficient testing for the general class of label-invariant properties. Verification with access to a **PCond** oracle (or any type of conditional sampling oracle) has also, to the best of our knowledge, not been explored. In our quest to find practical algorithms that work for large domains, we thus ask the following question.

*Can label-invariant properties be verified in a (query, sample and communication)-efficient way when the tester has access to a **PCond** oracle?*

## 1.1 Our Results

Our main result is an *exponential* query complexity separation between testing and verification for testing label-invariant properties with access to a **PCond** oracle. A detailed accounting of our results and comparison to existing work can be found in Table 1. A description follows.

One might have initially hoped that the power of a **PCond** oracle allows us to test label-invariant properties efficiently, even without the help of a prover. Indeed, with access to the full power of the **Cond** model (where arbitrary subsets  $S$  can be queried and a sample conditional on  $S$  is obtained), Chakraborty et al. [2013] show that this class of properties over a domain of size  $N$  can be tested with  $O(\text{poly log } N)$  queries to the **Cond** oracle. Our first result dashes this hope – we show a lower bound on the number of **PCond** queries required to test label-invariant properties with constant proximity parameter  $\tau$ , demonstrating that the **PCond** oracle is not much better in the worst case than the sampling oracle for this class of properties. Specifically, we show that a simple variant of the support size distinguishing problem for distributions over a domain of size  $N$  requires  $\Omega(N^{1/3})$  queries to a **PCond** oracle (the exact same as with access to only a sampling oracle). Thus, unaided, there exist (label-invariant) properties for which the **PCond** oracle is not much better than just sample

	Query Complexity Without Prover	Query Complexity With Prover	Communication	Rounds
Samp	$\tilde{\Omega}\left(\frac{N}{\log N}\right)$ Valiant and Valiant [2017]	$\tilde{O}(\sqrt{N})$ [Herman and Rothblum, 2022, 2023]	$\tilde{O}(\sqrt{N})$ [Herman and Rothblum, 2022, 2023]	2 [Herman and Rothblum, 2022, 2023]
PCond	$\Omega(N^{1/3})$ (Theorem ??)	$\text{poly}(\log N, \frac{1}{\tau})$	$\tilde{O}(\sqrt{N})$ (Theorem ??)	$\text{poly}(\log N, \frac{1}{\tau})$

■ **Table 1** Results on testing and verifying label-invariant properties under different types of access to the distribution. We state the best known lower bounds for label invariant properties. For Samp the lower bound is for entropy estimation, whereas for PCond it is the support size decision problem described in Section ???. The upper bounds apply for all label-invariant properties. Our results are highlighted in bold.

127 access.

128 ▶ **Theorem 1 (Informal Version of ??).** *There exists a label-invariant property  $\Pi$  such that*  
 129 *every tester with access to a PCond oracle for  $\Pi$  with proximity parameter  $\tau \leq 1/2$  and failure*  
 130 *probability 0.01 must make  $\Omega(N^{1/3})$  queries.*

131 The above lower bound motivates the investigation of verification with access to a PCond  
 132 oracle. As mentioned earlier, [Chiesa and Gur, 2018, Proposition 3.4] showed that with super-  
 133 linear communication complexity, there exists a reduction from verification to identity testing.  
 134 Instantiating this reduction with an identity tester using PCond oracles [Narayanan, 2021,  
 135 Theorem 1.5], we get that there exists an interactive proof system for every property with  
 136 super-linear communication complexity that makes only  $O(\sqrt{\log N}/\tau^2)$  queries to the PCond  
 137 oracle. However, super-linear communication is also prohibitive for practical algorithms; the  
 138 proof systems by Herman and Rothblum [2022, 2023] require the prover to only communicate  
 139  $\tilde{O}(\sqrt{N})$  domain elements, but still achieve the sample complexity of identity testing (for  
 140 the class of label-invariant properties). Could we also hope to achieve such communication  
 141 while maintaining similar query complexity as that of identity testing? The main result  
 142 of this paper is an affirmative answer to this question. Specifically, we give an interactive  
 143 proof system for tolerantly verifying *any* label-invariant property that has communication  
 144 complexity  $\tilde{O}(\sqrt{N})$  and query complexity  $\text{poly}(\log N)$  (suppressing the dependence on the  
 145 proximity parameter).

146 ▶ **Theorem 2 (Informal Label-Invariant Tolerant Verification Theorem (Theorem ??)).** *Fix a*  
 147 *label-invariant property  $\Pi$  over a domain  $[N]$  and proximity parameters  $\tau_c, \tau_f \in (0, 1/2]$ .*  
 148 *There exists a polylogarithmic (in  $N$ ) round interactive protocol  $\Pi$  between an honest verifier*  
 149  *$\mathsf{T}$ , and an omniscient untrusted prover  $\mathsf{P}^\mathcal{D}$ , where the verifier has PCond access to  $\mathcal{D}$ , such*  
 150 *that at the end of the interaction the verifier satisfies the following conditions:*

151 1. **Completeness:** *If the prover follows the protocol as prescribed, and  $d_{TV}(D, \Pi) \leq \tau_c$ ,*  
 152 *then*

$$153 \Pr \left[ \text{out} \left[ \Pi \left( \mathsf{P}^\mathcal{D}, \mathsf{T}^{(\mathcal{D})}; \tau_c, N \right) \right] = \text{Accept} \right] \geq 2/3$$

154 2. **Soundness:** *If  $d_{TV}(D, \Pi) \geq \tau_f$ , then for any prover  $\widetilde{\mathsf{P}}^\mathcal{D}$*

$$155 \quad \Pr \left[ \text{out} \left[ \Pi \left( \tilde{\mathbf{P}}^{\mathcal{D}}, \mathbf{T}^{(\mathcal{D})}; \tau_c, N \right) \right] = \text{Reject} \right] \geq 2/3$$

156     The complexity of the verifier is as follows:

- 157     1. **Query Complexity + Sample Complexity:**  $O(\text{poly}(\log N, 1/(\tau_f - \tau_c)))$
- 158     2. **Communication Complexity:**  $\tilde{O}(\sqrt{N} \text{poly}(1/(\tau_f - \tau_c)))$

159     In the process of proving this result, we give protocols for more basic primitives that  
160     may be of independent interest. Most significantly, we give an interactive proof system that  
161     is able to calculate the approximate probability mass of any point<sup>5</sup> in the domain using  
162     communication complexity  $\tilde{O}(\sqrt{N})$  and query complexity  $O(\text{poly}(\log N, 1/\tau))$ . As we will  
163     explain in the techniques section to follow, this is a key technical workhorse in our protocol  
164     for verifying label-invariant properties.

165     ► **Theorem 3** (Informal Version of ??). *Fix a domain  $[N]$ . For every  $\tau > 0$  and  $\delta \in (0, 1/2)$ ,  
166     there exists a  $O(\text{poly}(\log N, \log 1/\delta, \frac{1}{\tau}))$ -round interactive proof system such that the verifier  
167      $\mathbf{T}$  with access to a  $PCond$  oracle satisfies the following.*

- 168     1. **Completeness:** For every distribution  $\mathcal{D}$ , if the prover  $\mathbf{P}^{\mathcal{D}}$  is honest, then

$$169 \quad \Pr \left[ \mathbf{T} \text{ outputs } (y^*, \tilde{\mathcal{D}}[y^*]) \text{ s.t. } \frac{\tilde{\mathcal{D}}[y^*]}{\mathcal{D}[y^*]} \in \left[ \frac{1}{(1+\tau)}, 1+\tau \right] \right] \geq 1-\delta$$

- 170     2. **Soundness:** For any cheating prover  $\tilde{\mathbf{P}}^{\mathcal{D}}$ , then

$$171 \quad \Pr \left[ \mathbf{T} \text{ outputs Reject} \vee \mathbf{T} \text{ outputs } (y^*, \tilde{\mathcal{D}}[y^*]) \text{ s.t. } \frac{\tilde{\mathcal{D}}[y^*]}{\mathcal{D}[y^*]} \in [1/(1+\tau)^2, (1+\tau)^2] \right] \geq 1-\delta$$

172     The complexity of the verifier is as follows:

- 173     1. **Query Complexity + Sample Complexity:**  $O(\text{poly}(\log N, 1/(\tau)))$
- 174     2. **Communication Complexity:**  $\tilde{O}(\sqrt{N} \text{poly}(1/(\tau)))$

## 175     1.2 Technical Overview

176     In this section, we give an overview of our protocol for verifying label-invariant properties.

### 177     1.2.0.1 Unlabelled Bucket Histogram:

178     There is now a long line of work on the testing and verification of label-invariant properties  
179     [Batu et al., 2000, Valiant, 2008, Chakraborty et al., 2013, Herman and Rothblum, 2022,  
180     2023], and a key object used in this work is the unlabelled *approximate*  $\tau$ -bucket histogram  
181     of a distribution. Bucketing corresponds to partitioning the interval  $[0, 1]$  into smaller  
182     multiplicative probability intervals (see Definition ??). The  $\tau$ -bucket histogram divides the  
183     interval  $[0, 1]$  into  $\tilde{O}(\log N/\tau)$  buckets where the  $\ell^{\text{th}}$  bucket is a set of domain elements  
184     with individual probability mass in the range  $(\tau(1+\tau)^\ell/N, \tau(1+\tau)^{\ell+1}/N]$ . The *approximate*

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<sup>5</sup> Provided it does not have prohibitively small probability mass in its neighborhood.

185 unlabelled bucket histogram of a distribution then corresponds to a list of  $\tilde{O}(\log N/\tau)$   
 186 fractions, where the  $\ell^{\text{th}}$  element of the list is the fraction of domain points whose probability  
 187 lies in the range specified by the  $\ell^{\text{th}}$  bucket (see Definition ??). It is well known (see [Valiant,  
 188 2008]) that the *approximate* unlabelled  $\tau$ -bucket histogram of a distribution is a sufficient  
 189 statistic to (tolerantly) test any label-invariant property with proximity parameter(s)  $O(\tau)$ .  
 190 Thus, similar to prior work [Herman and Rothblum, 2022, 2023, Herman, 2024], our protocol  
 191 also focuses on efficiently verifying an unlabelled  $\tau$ -bucket histogram given to us by the  
 192 prover.

### 193 1.2.0.2Using Pairwise Comparisons to Learn Bucket Histogram:

194 Note that the unlabelled bucket histogram is a distribution over the buckets of the  $\tau$ -bucket  
 195 histogram and is hence a distribution over a domain of size  $\tilde{O}(\log \frac{N}{\tau})$ . Hence, by standard  
 196 results in distribution learning,  $\tilde{O}(\log N/\tau^2)$  samples from this bucket distribution would be  
 197 sufficient to learn it. However, sampling from this bucket distribution is non-trivial since a  
 198 sample from the original distribution  $\mathcal{D}$  does not come with information about histogram  
 199 bucket index.

200 While sampling from the bucket distribution might be hard with `Samp` access to the  
 201 distribution, one might be more optimistic about the possibility of sampling from the  $\tau$ -  
 202 approximate histogram with `PCond` queries. In particular, one approach that we might take is  
 203 the following: the verifier draws a dataset of size  $\tilde{O}(\log N/\tau^2)$  (enough to learn the histogram),  
 204 and sends these samples to the prover, who responds with the bucket index of each sample.  
 205 From how a  $\tau$ -histogram is defined, if  $x$  and  $y$  belong to buckets  $i$  and  $j$  respectively, then  
 206 this implies that the ratio of the probability masses of  $x$  and  $y$  under  $\mathcal{D}$  is guaranteed to  
 207 be in the interval  $\left[ \frac{1}{(1+\tau)^{|j-i|}}, (1+\tau)^{|j-i|} \right]$ . As `PCond` access allows us to conditionally sample  
 208 from a set restricted to two domain elements, it allows us to approximately learn the ratio of  
 209 their probability masses up to a multiplicative constant (see Lemma ??). Equipped with this  
 210 power, for each pair  $x \neq y$  from the set of drawn samples, the verifier uses the `PCond` oracle to  
 211 check if the learned ratios align with the provers claims. If the prover were to lie significantly,  
 212 then for at least one pair of samples, the claimed ratio would significantly different from  
 213 the learned ratio. Unfortunately, this simplistic strategy comes with two pitfalls. Firstly,  
 214 assuming the above strategy was sound, naively comparing elements with arbitrary bucket  
 215 indices could require  $\Omega(N)$  `PCond` queries if the elements being compared had significantly  
 216 different probability masses (see the ??, wherein  $K$  could be as large as  $N$ ). This would be as  
 217 bad as learning the distribution itself. Secondly, and more importantly, the above strategy is  
 218 *not* sound. It does not catch a prover that “slides” all samples into different buckets *in the*  
 219 *same way*, i.e., it lies about *every* bucket index by the same offset. As an example, consider  
 220 two distributions:  $\mathcal{D}_1$  is the uniform distribution over  $N^{1/4}$  domain elements and  $\mathcal{D}_2$  is the  
 221 uniform distribution over  $\sqrt{N}$  domain elements. The bucket histograms of both involve a  
 222 unit mass on a single (but different) bucket. However, pairwise comparisons between samples  
 223 taken from either distribution will always reveal a ratio that is approximately 1, since the  
 224 probabilities of all elements in the support of both distributions are identical. Hence, given  
 225 distribution  $\mathcal{D}_1$ , the prover can output the bucket histogram of distribution  $\mathcal{D}_2$ , and it is  
 226 impossible for a verifier to catch it purely by using the test described above.

227 A remedy to these obstacles lies in the following observation. If we had a good estimate  
 228 of the probability mass of a *single point*  $y$  in the domain, then we could resolve the soundness  
 229 issue discussed above. We simply use the `PCond` oracle to learn the ratio of probabilities  
 230 between each of our samples and  $y$ . Using this ratio, and knowledge of the approximate mass

231 of  $y$ , we can compute estimates of the probabilities of all samples. This would squash the  
 232 sliding attack described above. To deal with the first issue (that of the probability mass of  $y$   
 233 being very far from that of a sample), we would need to do more than learn just one value of  
 234  $y$ . Instead, we could learn the mass of a point  $y_j$ , for every bucket  $j$  that has large enough  
 235 mass. This way, for any sample  $x$  in the tester's set of samples (which are likely to come  
 236 from buckets with sufficiently large mass), we can find some  $y_j$  that is in a near-by bucket  
 237 with high probability.

#### 238 **1.2.0.3 Verifying the probability of points:**

239 Given that we simply need to identify the probability of a few points in the domain, one  
 240 might expect that this could be done even without access to a prover — this would give a  
 241 query-efficient tester for label-invariant properties. However, this ends up being a surprisingly  
 242 challenging task. Indeed, our lower bound in Theorem 1 shows that this is impossible (if we  
 243 wanted to bypass the  $\text{poly}(N)$  lower bound). This indicates a power of an untrusted prover;  
 244 it is able to certify the probability of a few points in the distribution support. Indeed, the  
 245 proof system with super-linear communication Chiesa and Gur [2018] achieves something  
 246 stronger- it certifies the entire distribution. Since we only need to certify the mass of at most  
 247  $O(\log N)$  points, we ask if this be done in a more communication efficient way?

#### 248 **1.2.0.4 Support Size Verification:**

249 Inspired by the “sliding” cheating prover from earlier, we consider the orthogonal but related  
 250 problem of verifying the support size of a flat distribution.<sup>6</sup> We will subsequently show  
 251 that a protocol for this problem can be combined with the ability of a PCond oracle to learn  
 252 neighbourhoods around points, enabling us to solve the probability approximation problem  
 253 for “relevant” domain elements.

254 Given a support size claim represented by four numbers  $A', A, B, B'$ , corresponding to  
 255 the claim that  $A' < A \leq \text{Supp}(\mathcal{D}) \leq B < B'$ , our hope is to accept if the claimed support  
 256 size range is accurate, and reject if the true support is larger than  $B'$  or smaller than  $A'$ .  
 257 Given different values of  $A$  and  $B$ , we develop a number of tests to verify with *sub-linear*  
 258 communication, the support size assuming the distribution is uniform over its support<sup>7</sup>. We  
 259 summarize the ideas below.

260 If the claimed support size upper bound  $B$  is small (that is,  $O(\sqrt{N})$ ), we could ask the  
 261 prover to send us the claimed support of the distribution. If the prover lies, and the true  
 262 support is actually much larger, then taking a few samples from the distribution would give  
 263 us a domain element outside the claimed support, thus catching the lie of the prover. If  
 264 the true support is much smaller than  $A$ , then taking a number of uniform samples from  
 265 the claimed support sent by the prover would result in a sample outside the support of  
 266 the distribution, which could be easily detected with a few PCond queries (see Lemma ??).  
 267 This gives us a protocol with a constant number of queries, and communication complexity  
 268 roughly  $O(B)$ .

269 On the other hand, if the claimed support size lower bound  $A$  is large (that is,  $\omega(\sqrt{N})$ ),  
 270 then asking the prover to send the support is not communication-efficient. Our approach  
 271 instead involves using uniform samples from the domain. The first test is to catch provers

<sup>6</sup> A distribution is *flat* if it is uniform over its support.

<sup>7</sup> We relax this condition in the main protocol, but assuming uniformity makes the description more intuitive.

that lie that the support is much larger than it really is. It involves drawing  $O(N/A)$  uniform samples  $S_1$  from the domain, and sending them to the prover. We ask the prover to send back a sample in this subset that is in the support of the distribution. If the true support is much smaller than  $A$ , there are (with high probability) no samples in the support of the distribution in  $S_1$ , and we can ensure the prover does not cheat by checking whether any element it sends back is in the support using a constant number of  $\text{PCond}$  queries. The second test catches provers that lie that the support is much smaller than it really is. It involves drawing  $O(N/B')$  samples  $z_1, \dots, z_m$  uniformly from the domain and permuting them with one sample  $x$  taken from the distribution. We ask the prover to identify the index of  $x$  in the permuted set. If the true support is smaller than  $B$ , then w.h.p, we expect that none of  $z_1, \dots, z_s$  are in the support of the distribution and hence, the honest prover can identify  $x$  exactly. On the other hand, if the support is larger than  $B'$ , we expect that at least one of  $z_1, \dots, z_s$  is in the support of the distribution, and the prover is unable to tell what the sample inserted by the prover was (since it could be  $x$  or  $z_i$ ). This gives us a protocol with a constant number of queries, and communication complexity roughly  $O(N/A)$ . Balancing parameters in these tests to optimize the communication complexity, we get an overall protocol for support size verification with communication complexity roughly  $\sqrt{N}$ .

We emphasize that the above outline is a simplification of the truth, and sweeps important details under the rug. Recall that our main goal is to certify the histogram of a distribution. In an attempt to do so, we will use the support size protocol above repeatedly as a sub-routine. This requires bounding the soundness and completeness errors in a meaningful way. As described above, these protocols do not have low enough soundness and completeness error to be used as sub-routines. Additionally, the above description assumes that distributions are exactly flat. In practice this will not be the case, and we need to be able to handle distributions where the probability ratio between any two elements in the support is upper bounded by some constant  $\alpha$  (nearly flat). In Section ??, we show how to analyse the protocols above to facilitate amplification of soundness and completeness errors, and make the protocol work for the more general class of  $\alpha$ -flat distributions.

### 1.2.0.5 Estimating Probability of a Point using Support Size Verification:

Finally, we explain how we can use the support size protocols to estimate the probability of a point. Prior work [Canonne et al., 2015] using the  $\text{PCond}$  oracle has shown that it can be used to estimate the mass within a multiplicative  $(1+\tau)$ -neighbourhood of any point<sup>8</sup> (see ??). We ask the prover to send us a point  $y^*$ <sup>9</sup> with sufficiently large mass in its neighbourhood (by an averaging argument, at least one histogram bucket needs to have  $\tau/\log N$  mass, ensuring that such a point exists, see Claim ??), and tell us the bucket the  $y^*$  belongs to. We then use the  $\text{PCond}$  oracle to estimate the mass within the multiplicative neighbourhood of  $y^*$ . The learned mass of the neighborhood divided by the prover's claimed probability mass<sup>10</sup> of  $y^*$  gives us bounds on the number of elements in  $y^*$ 's neighbourhood. Additionally, by definition the neighbourhood of  $y^*$  will be nearly flat. Hence, we have reduced the problem to a claim about the support size of a nearly-uniform distribution over a subset of the domain. Observe that we can sample from the distribution restricted to this bucket using  $\text{PCond}$  queries—since

<sup>8</sup> As long as the point does not have prohibitively small probability mass under the distribution.

<sup>9</sup> Recall that in the final protocol, we will ask the prover to send us points from every bucket with sufficiently large mass. For simplicity, we consider a single point in this description.

<sup>10</sup>The bucket index gives a lower and upper bound on the true probability mass of  $y^*$ . This is sufficient to catch a cheating prover.

313 there is sufficient mass in the neighbourhood of the point,  $O(\text{poly log } N)$  samples from the  
 314 distribution will contain at least one sample from the bucket, and we can use PCond queries  
 315 between the samples and  $y^*$  to find out which sample it is (see ??). If the prover was telling  
 316 the truth (or rather did not lie egregiously), then the support size claim holds and the verifier  
 317 will accept, thereby giving us a point and its approximate probability mass<sup>11</sup>. If the prover  
 318 significantly lied, then the support size claim will be false and the prover will be caught. We  
 319 note that the final analysis needs to handle some additional subtleties, since the PCond oracle  
 320 is itself only approximate and does not provide perfect comparisons (which can affect the  
 321 sampling), and additionally the bucket boundaries and the neighborhood of the provided  
 322 point do not overlap precisely. We refer to Sections ?? and ?? for the details. Once we  
 323 have estimated the probability of important points, we can use the ratio learning techniques  
 324 discussed earlier to certify the prover’s claim about the bucket histogram of the distribution.

### 325 1.3 Other Related Work:

#### 326 1.3.0.1 Interactive Proofs for Distribution Testing:

327 As mentioned earlier, interactive proofs for verifying distribution properties were first in-  
 328 troduced in the work of Chiesa and Gur [2018]. Follow-up work [Herman and Rothblum,  
 329 2022, 2023] studied interactive proofs for verifying *label-invariant* properties, focusing on  
 330 sublinear communication, and doubly efficient protocols (i.e. computationally-efficient and  
 331 sample-efficient generation of a proof). Herman [2024] studied *public-coin* interactive proofs  
 332 for testing label-invariant properties, where the verifier has no private randomness. Herman  
 333 and Rothblum [2024] gives communication-efficient and sample-efficient interactive proofs  
 334 for more general distribution properties that can be decided by uniform polynomial-size  
 335 circuits with bounded depth. Herman and Rothblum [2025] introduces *computationally*  
 336 *sound* interactive proofs and shows communication-, time-, and sample-efficient protocols for  
 337 any distribution property that can be decided in polynomial time given explicit access to the  
 338 full list of distribution probabilities. All of the above works assume that the verifier only  
 339 has sample access to the distribution, and the verifier sample complexity in all of them is  
 340  $\Omega(\sqrt{N})$  (where  $N$  is the size of the domain).

#### 341 1.3.0.2 Interactive Proofs for Learning:

342 A related but orthogonal line of work [Goldwasser et al., 2021, Mutreja and Shafer, 2023, Gur  
 343 et al., 2024, Caro et al., 2024b,a] focuses on interactive proofs for verifying learning problems –  
 344 for a specific hypothesis class  $H$ , given access to an untrusted prover, the goal is for a verifier  
 345 to output an accurate hypothesis from  $H$  for an underlying unknown distribution  $D$  if the  
 346 resource-rich prover is honest, and to reject if the prover lies egregiously. Different types  
 347 of resource asymmetry between the prover and the verifier are explored in these papers –  
 348 including differing number of samples, different computational complexities, different types  
 349 of access to the underlying function (sample vs query), and differing access to computational  
 350 resources (classical vs quantum computation and communication).

#### 351 1.3.0.3 Distribution testing under conditional oracles:

352 Our work focuses on interactive proofs for *verifying* distribution properties under conditional  
 353 sampling models. There is a long line of work on *testing* with access to conditional samples.

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<sup>11</sup>The claimed mass of  $y^*$  is derived from the bucket sent by the prover.

## 8:10 REFERENCES

354 Chakraborty et al. [2013], Canonne et al. [2014] introduced the conditional sampling (**Cond**)  
355 model and its more restricted variants (**PCond**, **ICond**). They gave algorithms for uniformity  
356 testing, tolerant uniformity testing, identity testing etc. in these conditional sampling models  
357 with query and sample complexity significantly better than the **Samp** model. Follow-up work  
358 shows improved bounds for identity testing (and its tolerant version), tolerant uniformity  
359 testing, and new algorithms for other tasks such as equivalence testing and support size  
360 problem in the **Cond** model [Falahatgar et al., 2015, Narayanan, 2021, Chakraborty et al.,  
361 2023, 2024]. There is also a line of work studying the power of non-adaptive queries in the  
362 conditional sampling model [Acharya et al., 2015, Kamath and Tzamos, 2019]. The **PCond**  
363 model was studied in detail by Narayanan [2021] who gave optimal bounds for identity testing  
364 and tolerant uniformity testing in this model, improving on results from Canonne et al. [2015].  
365 Testing under other types of conditional sampling has also been studied in the literature  
366 including subcube conditioning, where the distribution is supported on the hypercube, and  
367 the tester is allowed to ask for conditional samples from subcubes [Bhattacharyya and  
368 Chakraborty, 2018, Canonne et al., 2021, Chen et al., 2021, Kumar et al., 2023, Chakraborty  
369 et al., 2025], and coordinate conditional sampling [Blanca et al., 2025] (a version of subcube  
370 conditioning where all but one coordinate is fixed to a specific configuration and a sample  
371 is obtained from the remaining coordinate). Testing under other types of access to the  
372 distribution such as Probability Mass Function queries or Cumulative Distribution Function  
373 queries has also been studied in the literature [Batu et al., 2002, Rubinfeld and Servedio,  
374 2005, Guha et al., 2009, Canonne and Rubinfeld, 2014, Onak and Sun, 2018].

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