

# 2023 554 R Notes on Mapping for Point Data

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## Overview

In these notes we will consider mapping and modeling of point data in which the (nominal) exact locations are known.

We will look at modeling a spatially-indexed continuous response via:

- Conventional Kriging via MLE and variants
- A generalized additive model (GAM)
- A Bayesian approach using stochastic partial differential equations (SPDE)

## Continuous Response: Motivating Example

We illustrate methods for continuous data using on Zinc levels in the Netherlands.

This data set gives locations and top soil heavy metal concentrations (in ppm), along with a number of soil and landscape variables, collected in a flood plain of the river Meuse, near the village Stein in the South of the Netherlands.

Heavy metal concentrations are bulk sampled from an area of approximately  $28\text{km} \times 39\text{km}$ .

The Meuse data are in a variety of packages. The version in the `geoR` library are not a spatial object, but can be used with likelihood and Bayes methods.

## Meuse analysis using `geostat` functions

We look at the sampling locations and then examine variograms.

```
library(tidyverse)
library(ggpubr)
library(viridis)
library(geoR)
data("meuse")
library(sp)
library(INLA)
```

```

pal <- function(n = 9){ brewer.pal(n, "Reds") }
data(meuse)
coords <- SpatialPoints(meuse[,c("x", "y")])
meuse1 <- SpatialPointsDataFrame(coords, meuse)
data(meuse.riv)
river_polygon <- Polygons(list(Polygon(meuse.riv)), ID="meuse")
rivers <- SpatialPolygons(list(river_polygon))
coordinates(meuse) = ~x+y

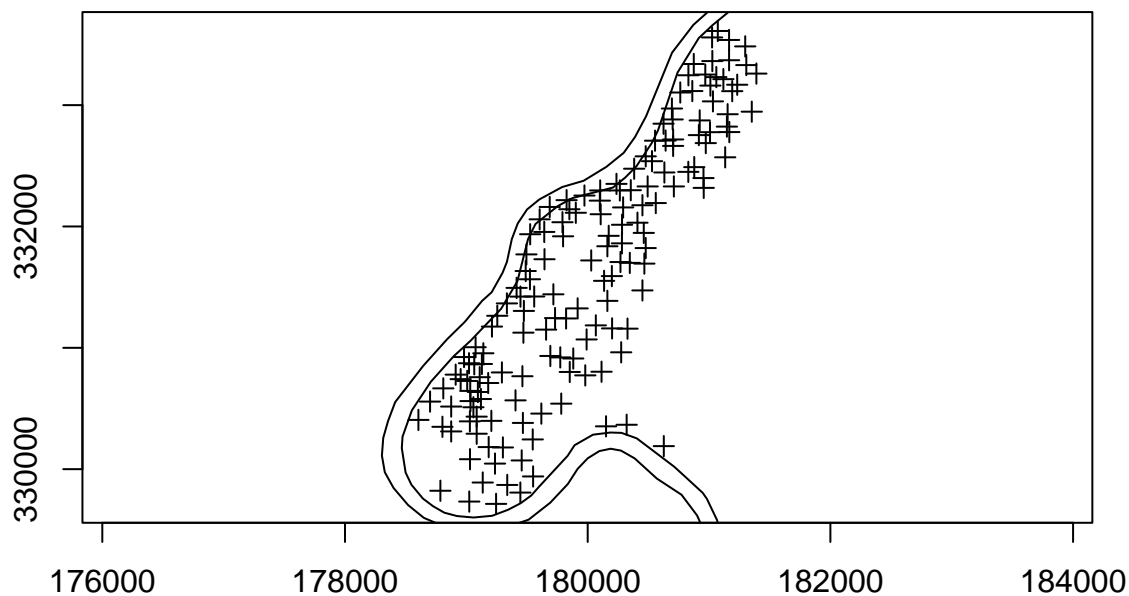
```

## Zinc: Sampling locations

```

plot(meuse1, axes=T)
plot(rivers, add=T)

```



## Exploratory analysis

We work with  $\log(\text{zinc})$  as the distribution is more symmetric than on the original scale, and the variance more constant across levels of covariates.

It's often a good idea to do some exploratory data analysis (EDA), so let's see how  $\log(\text{zinc})$  varies by two possible covariates:

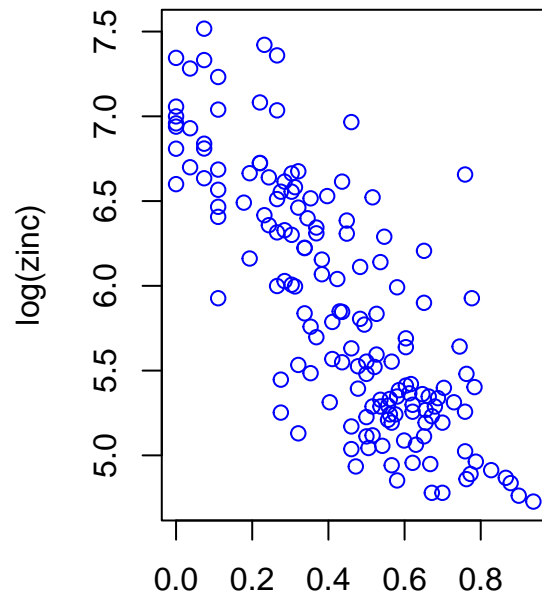
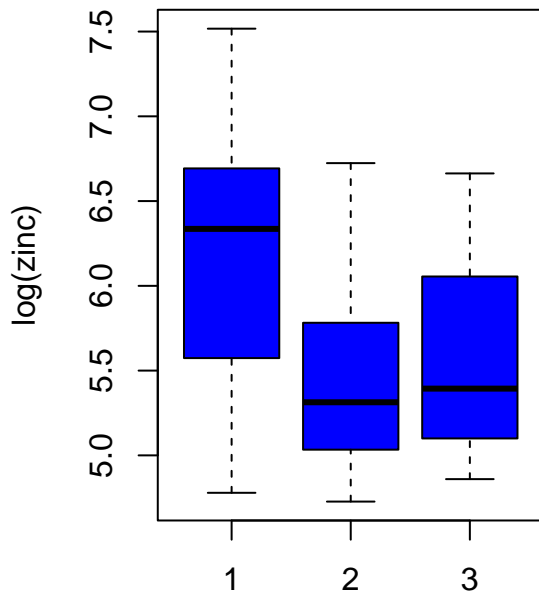
- Flooding frequency (`ffreq`); 1 = once in two years; 2 = once in ten years; 3 = one in 50 years
- Distance to the Meuse river (`dist`); normalized to [0, 1]

We focus on these, since they are available across the study region and so can be used for prediction. Following previous authors, we take `sqrt(dist)` which is closer to linearity.

```

par(mfrow=c(1,2))
plot(log(meuse$zinc)~meuse$ffreq, ylab="log(zinc)", xlab="Flooding Frequency", col="blue")
plot(log(meuse$zinc)~sqrt(meuse$dist), ylab="log(zinc)", xlab="Sqrt distance to river", col="blue")

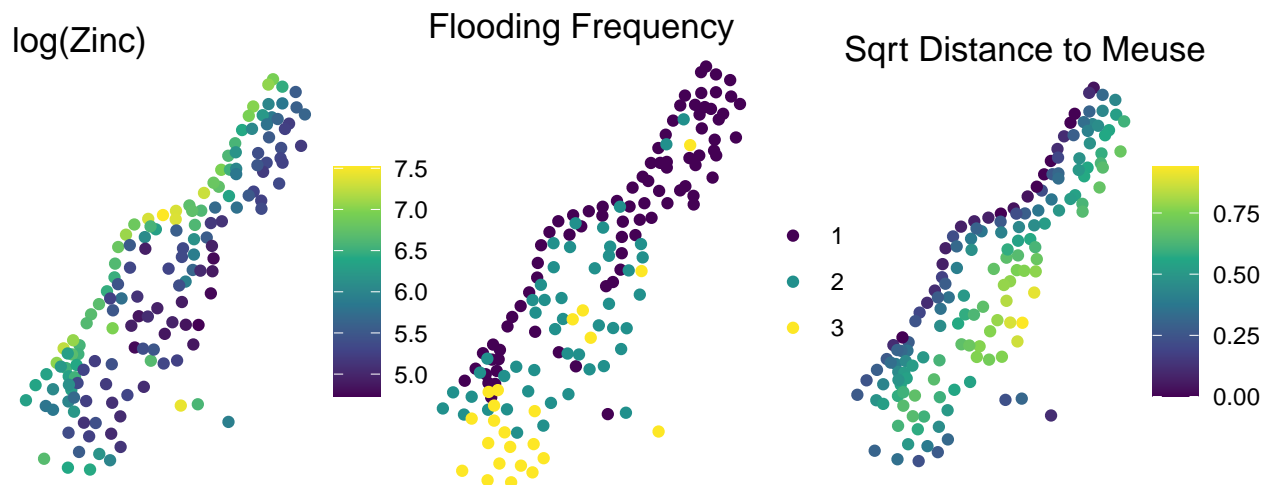
```



Also map log(zinc) and these covariates.

```
m.sf <- sf::st_as_sf(meuse, coords = c("x", "y"))
m.sf$logzinc <- log(m.sf$zinc)
m.sf$sdist <- sqrt(m.sf$dist)

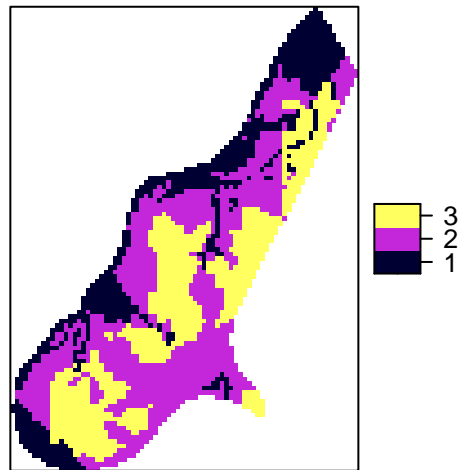
a <- ggplot() + geom_sf(data = m.sf[, "logzinc"], aes(color = logzinc)) +
  theme_void() + scale_color_viridis_c() + labs(title = "log(Zinc)", color=NULL)
b <- ggplot() + geom_sf(data = m.sf[, "ffreq"], aes(color = ffreq)) +
  theme_void() + scale_color_viridis(discrete=T) + labs(title = "Flooding Frequency", color=NULL)
c <- ggplot() + geom_sf(data = m.sf[, "sdist"], aes(color = sdist)) +
  theme_void() + scale_color_viridis_c() + labs(title = "Sqrt Distance to Meuse", color=NULL)
ggpubr::ggarrange(a,b,c, nrow=1, ncol=3)
```



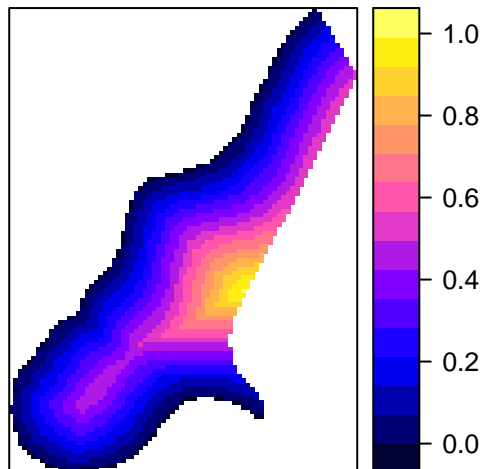
These two covariates are also available for a grid, in the `meuse.grid` data object. We will make use of this when we get to making predictions. Load in `meuse.grid` and take a look at the grid covariates.

```
library(sp)
data(meuse.grid)
coordinates(meuse.grid) = ~x+y
proj4string(meuse.grid) <- CRS("+init=epsg:28992")
gridded(meuse.grid) = TRUE
```

```
spplot(meuse.grid["ffreq"])
```

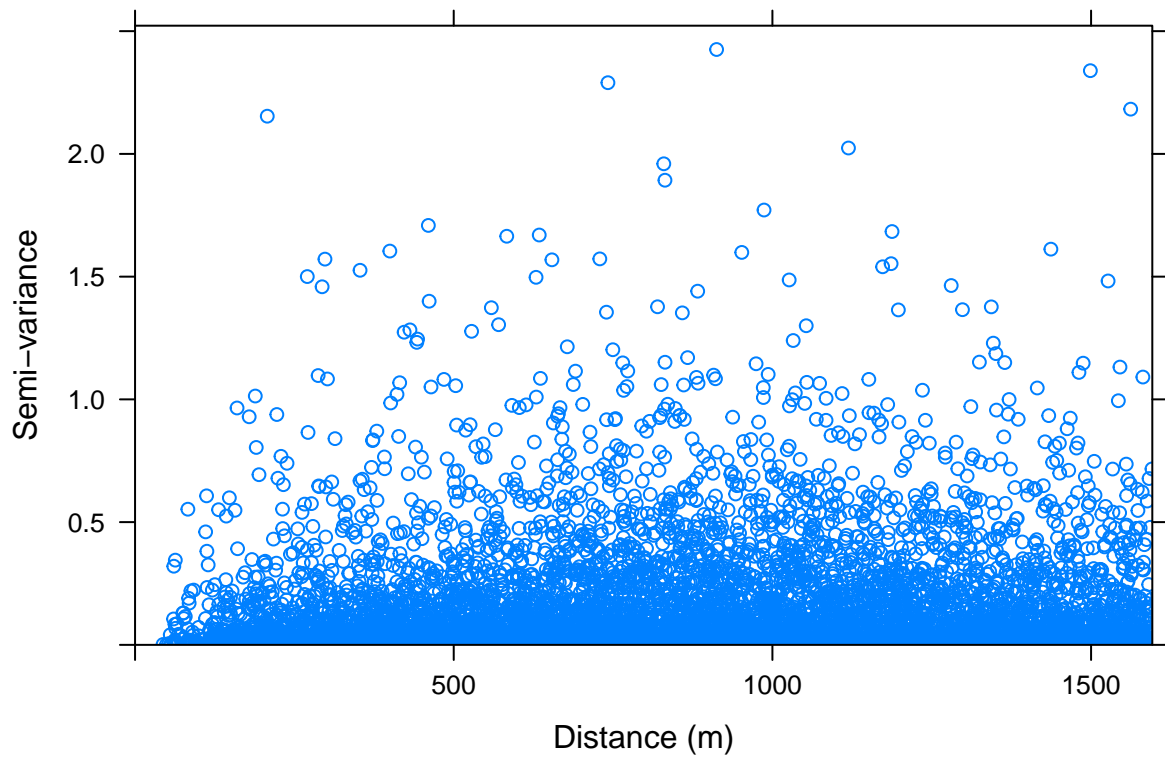


```
spplot(meuse.grid["dist"])
```



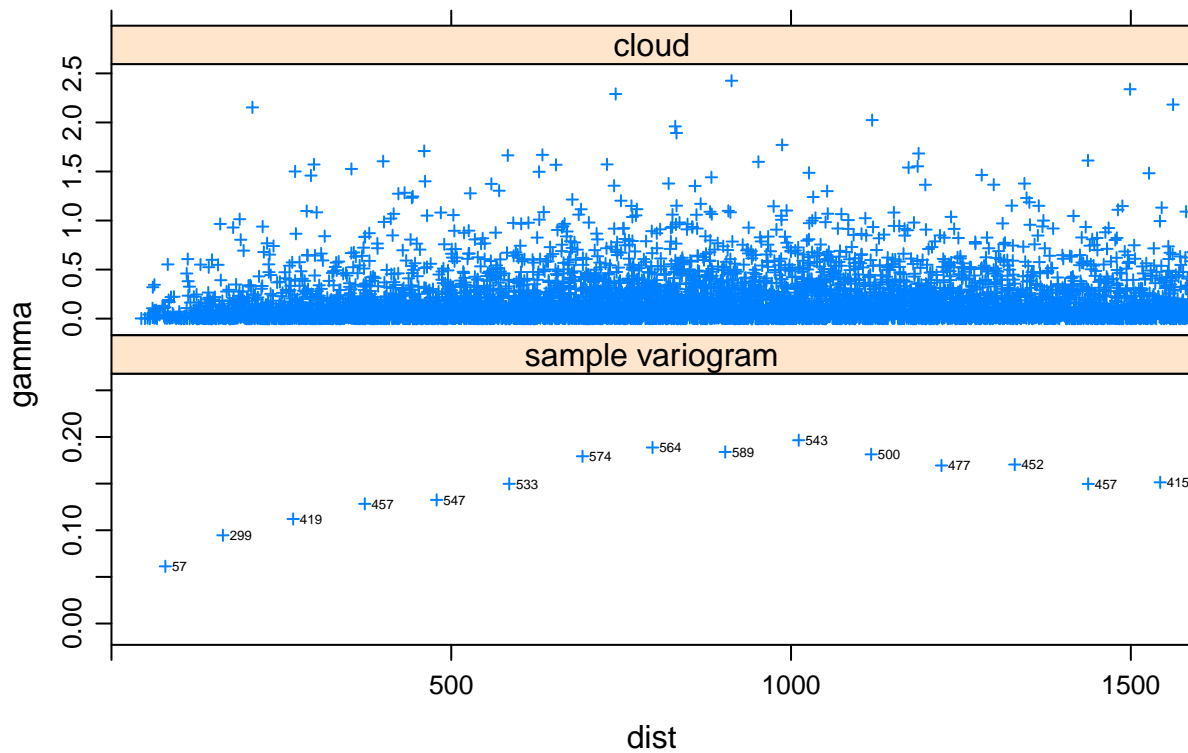
log(zinc): Variogram cloud, trend removed

```
library(gstat)
cld <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse, cloud = TRUE)
plot(cld, ylab="Semi-variance", xlab="Distance (m)")
```



More variograms, with sample sizes

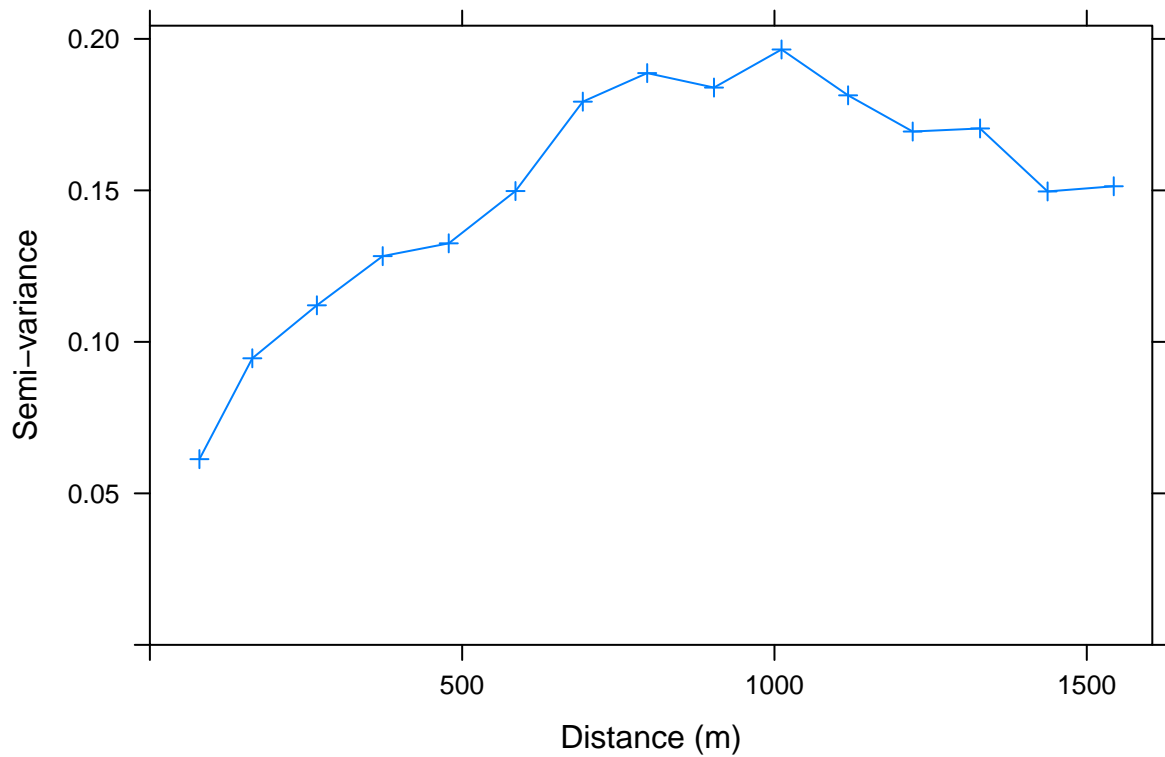
```
library(lattice)
cld <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse, cloud = TRUE)
svgm <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse)
d <- data.frame(gamma = c(cld$gamma, svgm$gamma),
  dist = c(cld$dist, svgm$dist),
  id = c(rep("cloud", nrow(cld)), rep("sample variogram", nrow(svgm)))
)
xyplot(gamma ~ dist | id, d,
  scales = list(y = list(relation = "free",
    #ylim = list(NULL, c(-.005,0.25)))),
    limits = list(NULL, c(-.005,0.25))),
  layout = c(1, 2), as.table = TRUE,
  panel = function(x,y, ...) {
    if (panel.number() == 2)
      ltext(x+10, y, svgm$np, adj = c(0,0.5), cex=.4) #
    panel.xyplot(x,y,...)
  },
  xlim = c(0, 1590),
  cex = .5, pch = 3
)
```



## Monte Carlo simulations of semi-variogram

We simulate 100 datasets with random relabeling of points, and then form variograms for each.

```
v <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse)
plot(v, type = 'b', pch = 3, xlab="Distance (m)", ylab="Semi-variance")
fn = function(n = 100) {
  for (i in 1:n) {
    meuse$random = sample(meuse$zinc)
    v = variogram(log(random) ~ 1, meuse)
    trellis.focus("panel", 1, 1, highlight = FALSE)
    llines(v$dist, v$gamma, col = 'grey')
    trellis.unfocus()
  }
}
fn()
```



Monte Carlo envelopes under no spatial dependence - it is clear there is dependence here.

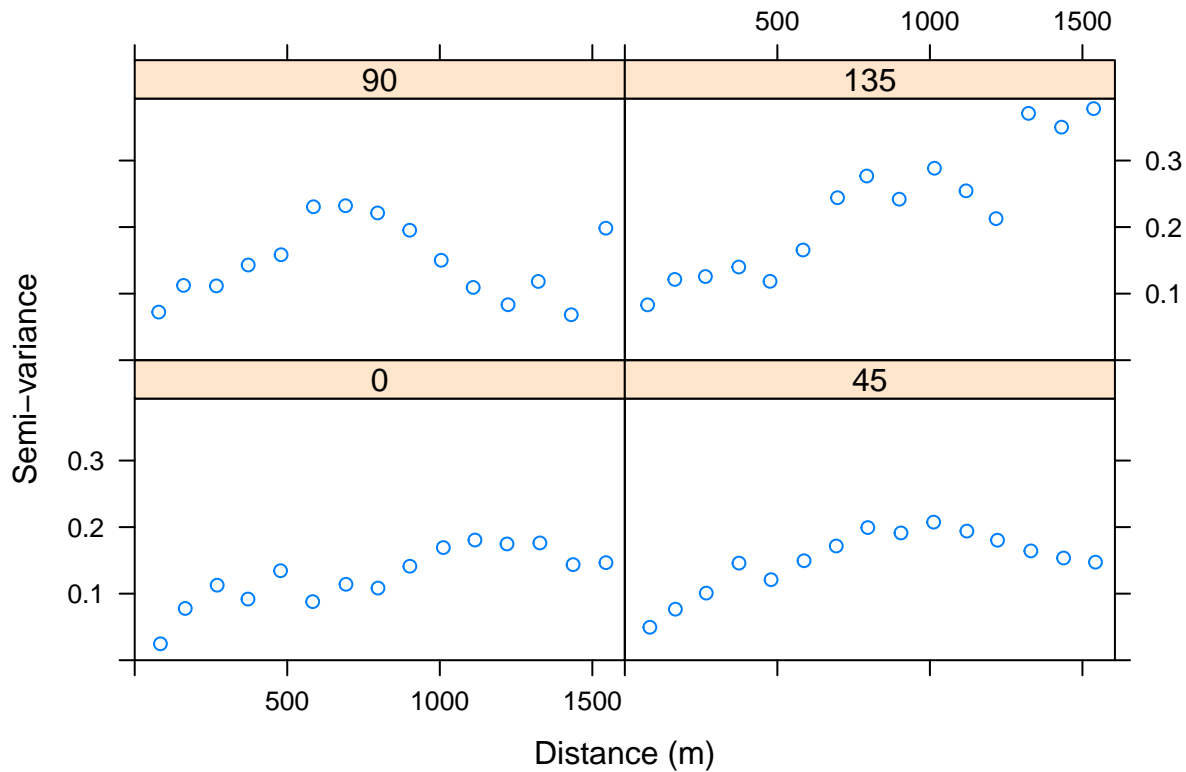
### Directional variogram with linear trend removed

We form 4 variograms with data taken from different directions, with 0 and 90 corresponding to north and east, respectively.

Note that 0 is the same as 180.

```
dirclld <- variogram(log(zinc)~sqrt(meuse$dist)+as.factor(ffreq), meuse, alpha=c(0,45,90,135))
```

```
plot(dirclld,xlab="Distance (m)",ylab="Semi-variance")
```



## Other capabilities in gstat

See

- `fit.variogram` for estimation from the variogram
- `krige` (and associated functions) for Kriging,
- `vgm` generates variogram models

## geoR for geostatistics

We continue the analysis using functions from the `geoR` library, for which a `geodata` data type is required. There are 155 observations (sampling locations)

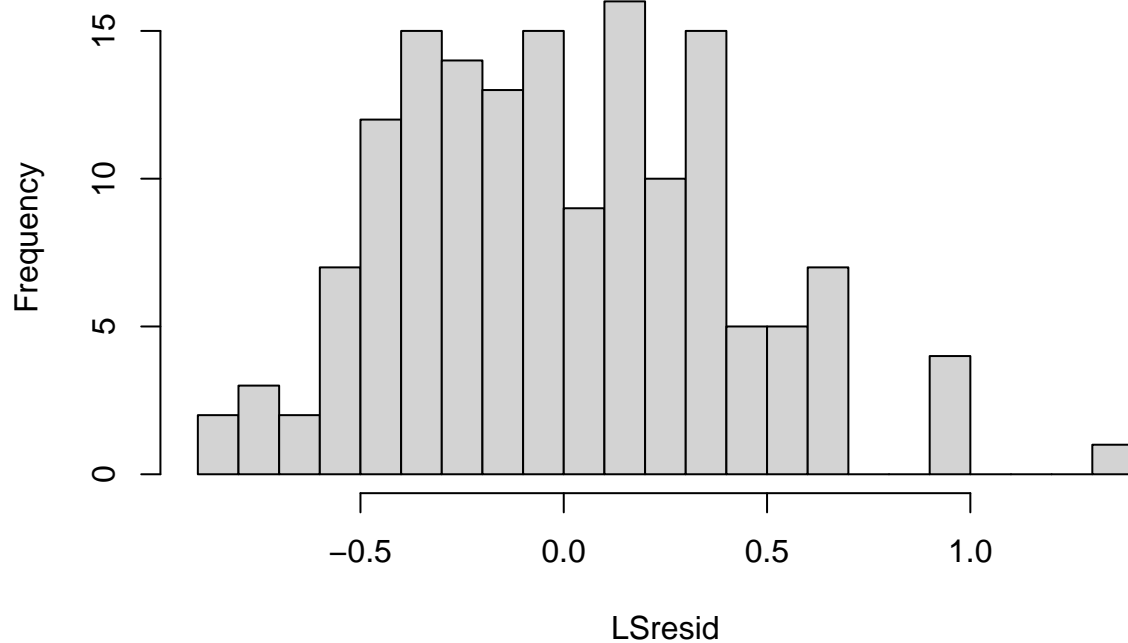
```
zmat <- matrix(cbind(meuse$x,meuse$y,log(meuse$zinc)),
               ncol=3,nrow=155,byrow=F)
geozinc <- as.geodata(zmat,coords.col=c(1,2),data.col=c(3))
```

First we will be assuming a spatial model on the residuals with `ffreq` and `sqrt(dist)` in the model. Fit this initial linear model, and view the residuals by histogram and map.

```
LSmod <- lm(log(meuse$zinc)~sqrt(meuse$dist)+as.factor(meuse$ffreq))
LSresid <- residuals(LSmod,type="pearson")
hist(LSresid,nclass=25)
```

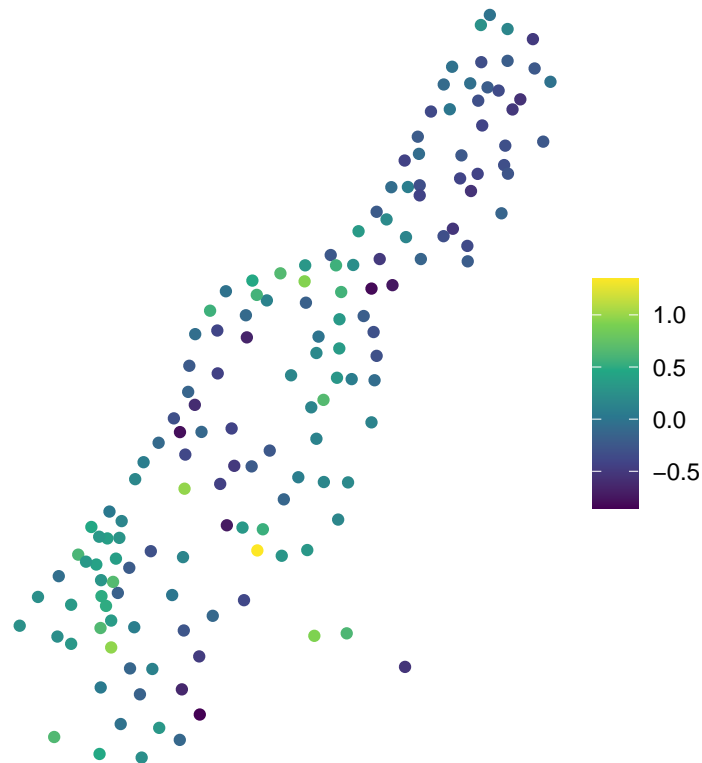


## Histogram of LSresid



```
m.sf$resid <- LSresid  
ggplot() + geom_sf(data = m.sf[, "resid"], aes(color = resid)) +  
  theme_void() + scale_color_viridis_c() + labs(title = "Residuals from Linear Model", color=NULL)
```

## Residuals from Linear Model



## Moran's I

We can calculate Moran's I to see the strength of spatial dependence in the residuals. Our results show that we can reject the null hypothesis that there is zero spatial autocorrelation present in LSresid at the  $\alpha = 0.05$  level.

```
library(ape)
dists <- as.matrix(dist(cbind(geozinc[1]$coords[, 1], geozinc[1]$coords[, 2])))
dists.inv <- 1/dists
diag(dists.inv) <- 0
dists.inv[1:5, 1:5]
##           1           2           3           4           5
## 1 0.000000000 0.014116748 0.008414063 0.003857440 0.002729898
## 2 0.014116748 0.000000000 0.007063831 0.003535423 0.002757553
## 3 0.008414063 0.007063831 0.000000000 0.006984644 0.003983684
## 4 0.003857440 0.003535423 0.006984644 0.000000000 0.006482446
## 5 0.002729898 0.002757553 0.003983684 0.006482446 0.000000000

Moran.I(LSresid, dists.inv)
## $observed
## [1] 0.1210698
##
## $expected
## [1] -0.006493506
##
## $sd
```

```
## [1] 0.01060145
##
## $p.value
## [1] 0
```

## Maximum likelihood for log(zinc)

We fit with the Matern covariance model with

$$\rho(h) = \frac{1}{2^{\kappa-1}\Gamma(\kappa)} \left(\frac{h}{\phi}\right)^{\kappa} K_{\kappa}\left(\frac{h}{\phi}\right),$$

where  $h$  is the distance between 2 points and we take  $\kappa = 0.5$  so that  $\phi$  is the only estimated parameter. The other parameters estimated are  $\tau^2$ , the nugget, and  $\sigma^2$ , the spatial variance. The practical range reported is the distance at which the correlations drop to 0.05, and is a function of  $\phi$ .

We suppress the output from the call.

```
mlfit <- likfit(geozinc, cov.model="matern", ini=c(.2,224), trend=~sqrt(meuse$dist)+as.factor(meuse$ffreq))
```

We examine the results, specifically point estimates and standard errors.

```
mlfit$parameters.summary
##           status    values
## beta0    estimated    7.0712
## beta1    estimated   -2.1208
## beta2    estimated   -0.5154
## beta3    estimated   -0.5266
## tausq    estimated    0.0372
## sigmasq  estimated    0.1316
## phi      estimated  300.5381
## kappa      fixed    0.5000
## psiA      fixed    0.0000
## psiR      fixed    1.0000
## lambda    fixed    1.0000
for (i in 1:3){
cat(cbind(mlfit$beta[i],sqrt(mlfit$beta.var[i,i])), "\n")
}
## 7.071249 0.1326244
## -2.120796 0.2365602
## -0.5153624 0.06780136
mlfit$practicalRange
## [1] 900.3318
```

Note that the standard errors change from the least squares fit.

## Restricted maximum likelihood for log(zinc)

Restricted MLE is preferable in general for estimation when there are variance parameters.

```
remlfit <- likfit(geozinc,cov.model="matern",ini=c(.55,224),lik.method="RML",
  trend=~sqrt(meuse$dist)+as.factor(meuse$ffreq))
```

The results show slight differences from ML.

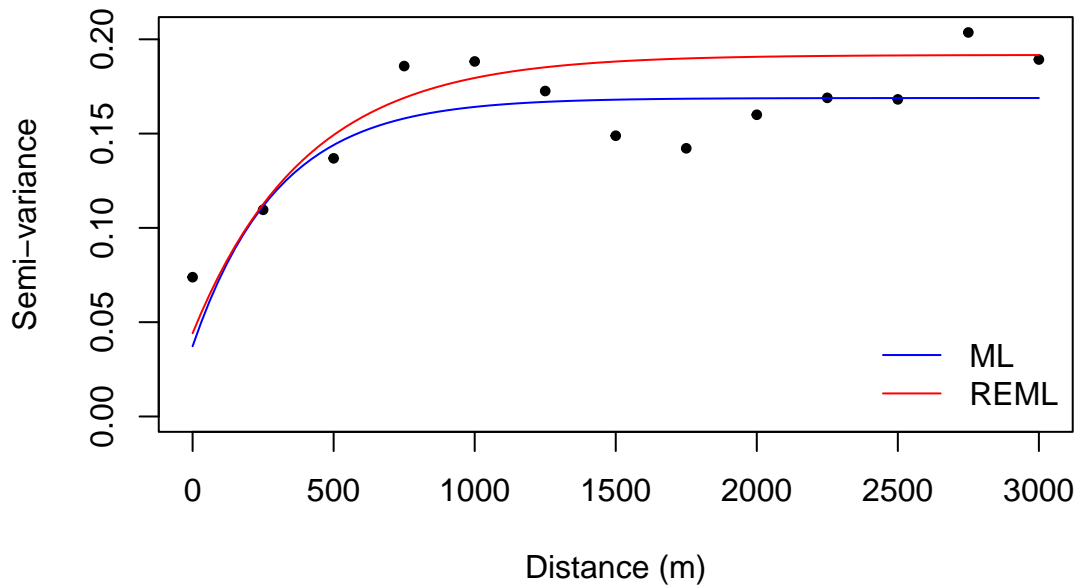
```
remlfit$parameters.summary
##           status    values
## beta0    estimated    7.0823
## beta1    estimated   -2.1109
## beta2    estimated   -0.5280
## beta3    estimated   -0.5455
## tausq    estimated    0.0442
## sigmasq  estimated    0.1476
## phi      estimated  401.4313
## kappa     fixed     0.5000
## psiA      fixed     0.0000
## psiR      fixed     1.0000
## lambda    fixed     1.0000
remlfit$practicalRange
## [1] 1202.581
```

```
for (i in 1:3){
cat(cbind(mlfit$beta[i],sqrt(mlfit$beta.var[i,i]),remlfit$beta[i],sqrt(remlfit$beta.var[i,i])), "\n")
}
## 7.071249 0.1326244 7.082279 0.1527493
## -2.120796 0.2365602 -2.110917 0.2515499
## -0.5153624 0.06780136 -0.5280202 0.06913553
```

## Comparison of estimates

```
binzinc <- variog(geozinc,uvec=seq(0,5000,250),
  trend=~sqrt(meuse$dist)+as.factor(meuse$ffreq))
## variog: computing omnidirectional variogram

plot(binzinc,max.dist=3000,xlab="Distance (m)",ylab="Semi-variance",pch=19,cex=.6)
lines(mlfit,max.dist=3000,lty=1,col="blue")
lines(remlfit,max.dist=3000,lty=1,col="red")
legend("bottomright",legend=c("ML","REML"),
  lty=c(1,1),bty="n",col=c("blue","red"),cex=1.0)
```



## Prediction for log(zinc) by Kriging

We now fit a linear model in distance and elevation to log(zinc). We then form a `geodata` object with the residuals as the response.

```
lmdata <- data.frame(logzinc=geozinc$data, sqrtdist=sqrt(meuse$dist),
                     ffreq=as.factor(meuse$ffreq))
lmfit <- lm(logzinc~sqrtdist+ffreq, data=lmdata)
lmfit
##
## Call:
## lm(formula = logzinc ~ sqrtdist + ffreq, data = lmdata)
##
## Coefficients:
## (Intercept)      sqrtdist      ffreq2      ffreq3
##      7.0299      -2.2660     -0.3605     -0.3167
detrend <- as.geodata(cbind(geozinc$coords,lmfit$residuals))

likfit(geozinc, ini = c(1000, 50))
## -----
## likfit: likelihood maximisation using the function optim.
## likfit: Use control() to pass additional
##         arguments for the maximisation function.
##         For further details see documentation for optim.
## likfit: It is highly advisable to run this function several
##         times with different initial values for the parameters.
## likfit: WARNING: This step can be time demanding!
## -----
## likfit: end of numerical maximisation.
## likfit: estimated model parameters:
##      beta      tausq      sigmasq      phi
## " 6.6360" " 0.0347" " 1.8478" "2142.6149"
## Practical Range with cor=0.05 for asymptotic range: 6418.701
```

```
##
## likfit: maximised log-likelihood = -99.13
```

We can obtain spatial predictions on a grid, using the parameter estimates from the ML fit.

```
# get meuse.grid and expand to all combinations
data(meuse.grid)
pred.grid <- expand.grid(sort(unique(meuse.grid$x)),
                        sort(unique(meuse.grid$y)))

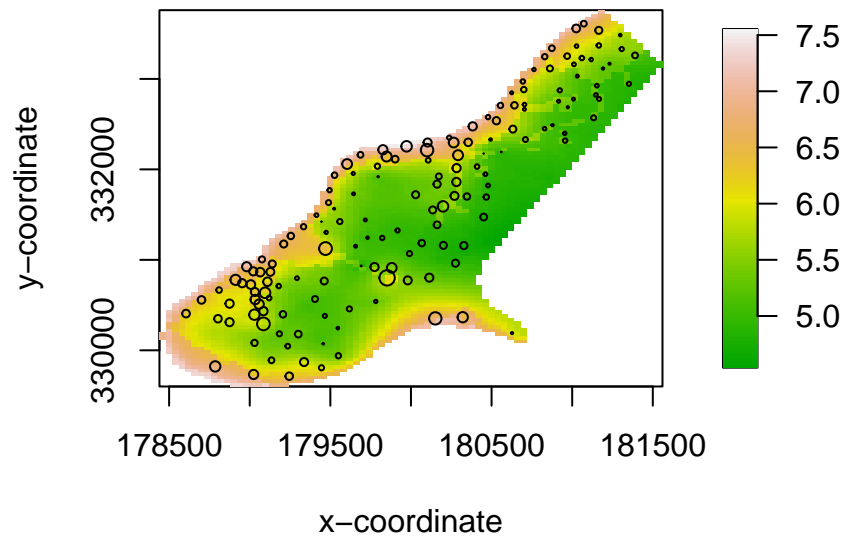
# kriging
kc <- krige.conv(detrend, loc=pred.grid,
                krige=krige.control(obj.m=mlfit))
## krige.conv: model with constant mean
## krige.conv: Kriging performed using global neighbourhood
```

However, since we're predicting the residuals, to get back to the original task of predicting log(zinc), we can add this value back onto our predictions from `lmfit`, which yields the following:

```
# merge linear predictions onto grid
pred_lm <- meuse.grid[, c("x", "y", "dist", "ffreq")]
pred_lm$sqrtdist <- sqrt(pred_lm$dist)
pred_lm$pred_lm <- predict(lmfit, newdata = pred_lm)
pred_resid <- data.frame(pred_resid = kc$predict, x = pred.grid$Var1,
                        y = pred.grid$Var2)
pred_logzinc <- merge(pred_resid, pred_lm, by = c("x", "y"), all=T)
pred_logzinc$pred_logzinc <- pred_logzinc$pred_lm + pred_logzinc$pred_resid
pred_logzinc <- pred_logzinc %>% arrange(y,x)
```

To view the results produce an image plot of the predictions, with the data superimposed.

```
library(fields)
image.plot(x=unique(pred.grid[["Var1"]]), y=unique(pred.grid[["Var2"]]),
           z=matrix(pred_logzinc$pred_logzinc, nrow=78, ncol=104), col=terrain.colors(100),
           xlab="x-coordinate", ylab="y-coordinate")
symbols(detrend$coords[,1], detrend$coords[,2],
        circles=(detrend$data-min(detrend$data))/1, add=T, inches=0.04)
```

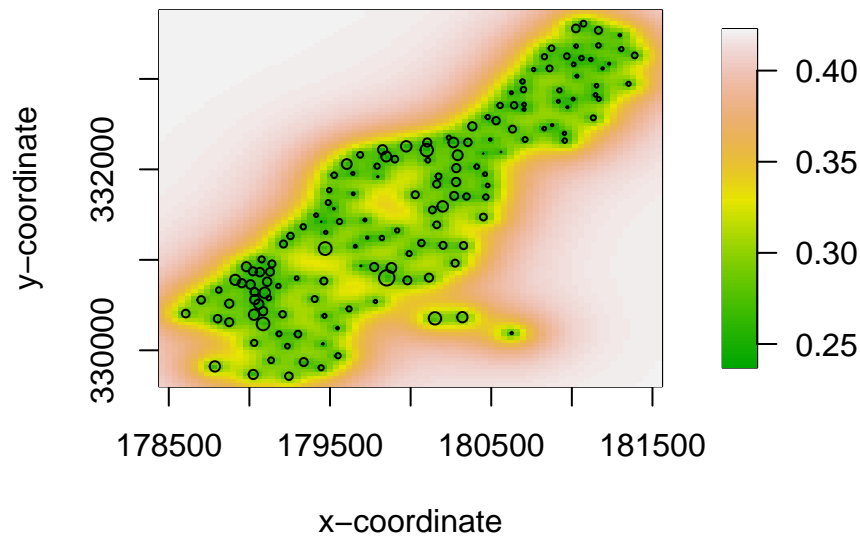


```
# library(sf)
# df <- st_as_sf(zinc.dat, coords=c("x", "y"))
# ggplot() +
#   geom_sf(data = df, col = "black", fill = NA) +
#   theme_bw() +
#   coord_sf() +
#   xlab("") +
#   ylab("") +
#   theme(legend.title = element_blank(),
#         panel.grid = element_blank(),
#         panel.border = element_blank(),
#         axis.ticks = element_blank(),
#         axis.text = element_blank())
```

## Standard deviations of prediction for log(zinc)

We now plot the Kriging standard deviations of the predictions.

```
image.plot(x=unique(pred.grid[["Var1"]]),y=unique(pred.grid[["Var2"]]),
           z=matrix(sqrt(kc$krige.var),nrow=78,ncol=104),col=terrain.colors(100),
           xlab="x-coordinate",ylab="y-coordinate")
symbols(detrend$coords[,1],detrend$coords[,2],
        circles=(detrend$data-min(detrend$data))/1,add=T,inches=0.04)
```



The standard deviation is smallest close to the datapoints, as expected.

## GAM model for $\log(\text{zinc})$

We now model the  $\log(\text{zinc})$  surface as linear in the square root of distance to the Meuse and flooding frequency, and with the spatial surface modeled with a thin plate regression spline, with the smoothing parameter estimated using REML.

```
library(mgcv)
library(lattice)
library(latticeExtra)
library(RColorBrewer)
zinc.dat <- data.frame(x=meuse$x,
                      y=meuse$y, lzinc=log(meuse$zinc),
                      sqrtsdist=sqrt(meuse$dist),
                      ffreq=as.factor(meuse$ffreq))
gam.mod <- gam(lzinc ~ s(x,y, bs="tp") + sqrtsdist + ffreq,
              data=zinc.dat, method="REML")
```

```
summary(gam.mod)
##
## Family: gaussian
## Link function: identity
##
## Formula:
## lzinc ~ s(x, y, bs = "tp") + sqrtsdist + ffreq
##
## Parametric coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   6.9831     0.1522  45.869 < 2e-16 ***
## sqrtsdist    -1.8917     0.3496  -5.411 2.87e-07 ***
## ffreq2       -0.5879     0.0721  -8.154 2.43e-13 ***
## ffreq3       -0.6241     0.1129  -5.530 1.66e-07 ***
```



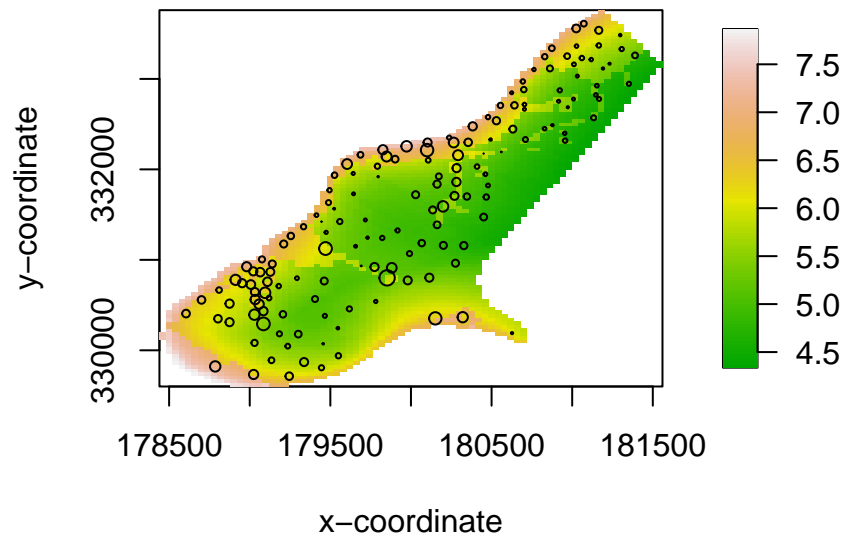
```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##           edf Ref.df      F p-value
## s(x,y) 19.27  24.06 5.803  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.835   Deviance explained = 85.9%
## -REML = 58.733   Scale est. = 0.085927   n = 155
```

## GAM prediction

```
pred.grid.gam <- expand.grid(sort(unique(meuse.grid$x)),
                             sort(unique(meuse.grid$y)))
meuse.grid$sqrtdist <- sqrt(meuse.grid$dist)
meuse.grid$zinc.pred.gam <- predict.gam(gam.mod, meuse.grid)
meuse.grid$zinc.pred.gam.sd <- predict.gam(gam.mod, se.fit=T, meuse.grid)$se.fit

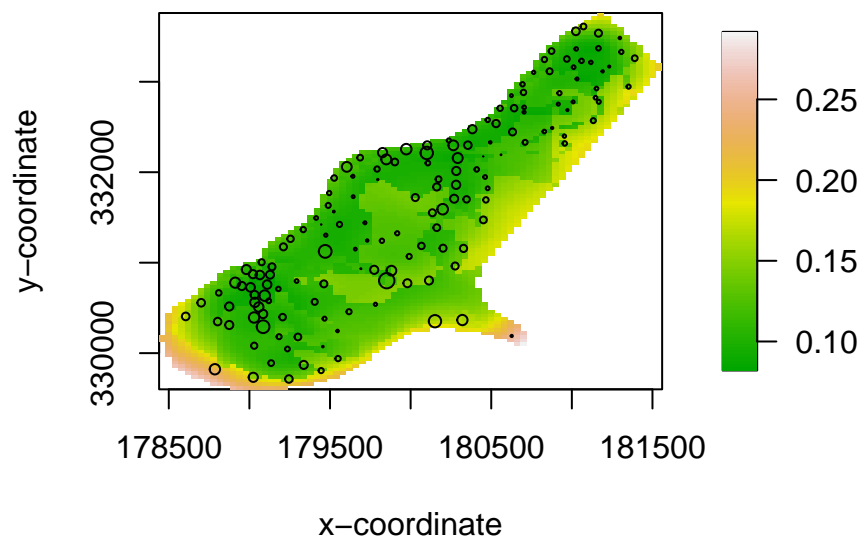
zinc.pred.gam <- xtabs(zinc.pred.gam ~ x + y, data = meuse.grid)
zinc.pred.gam[zinc.pred.gam==0] <- NA
zinc.pred.gam.sd <- xtabs(zinc.pred.gam.sd ~ x + y, data = meuse.grid)
zinc.pred.gam.sd[zinc.pred.gam.sd==0] <- NA
```

```
image.plot(x=pred.grid.gam[["Var1"]][1:78],
           y=unique(pred.grid.gam[["Var2"]]),
           z=matrix(zinc.pred.gam,nrow=78,ncol=104),
           col=terrain.colors(100),
           xlab="x-coordinate",
           ylab="y-coordinate")
symbols(detrend$coords[,1],
        detrend$coords[,2],
        circles=(detrend$data-min(detrend$data))/1,
        add=T,inches=0.04)
```



### Standard deviations of prediction from GAM

```
image.plot(x=pred.grid.gam[["Var1"]][1:78],
           y=unique(pred.grid.gam[["Var2"]]),
           z=matrix(zinc.pred.gam.sd,nrow=78,ncol=104),
           col=terrain.colors(100),
           xlab="x-coordinate",
           ylab="y-coordinate")
symbols(detrend$coords[,1],
        detrend$coords[,2],
        circles=(detrend$data-min(detrend$data))/1,
        add=T,inches=0.04)
```



### SPDE Model for $\log(\text{zinc})$

We not create a Bayesian SPDE model for  $\log(\text{zinc})$ .

```

zincdf <- zinc.dat %>%
  dplyr::select(y = lzinc,
                locx = x,
                locy = y,
                sqrtldist = sqrtldist,
                ffreq = ffreq)

# Create mesh
max.edge <- 200
mesh <- inla.mesh.2d(
  loc = zincdf[, c("locx", "locy")],
  offset = c(100, 500), max.edge = c(max.edge, max.edge *
    3)
)

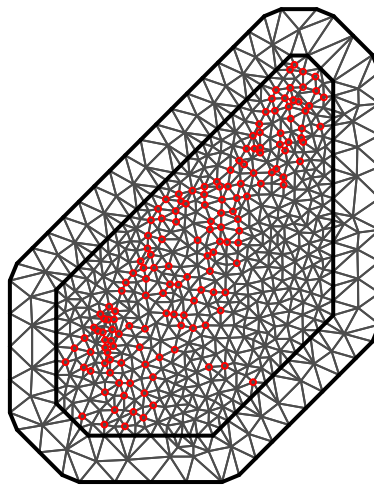
```

We can visualize the mesh we created as follows:

```

plot(mesh, asp = 1, main = "")
points(zincdf[, c("locx", "locy")], col = "red", cex = 0.4)

```



Next, we create an  $A$  matrix and the stack, which we need in order to define our model.

```

A <- inla.spde.make.A(mesh = mesh, loc = data.matrix(zincdf[, c("locx", "locy")])
)

# Define X matrix
Xcov <- data.frame(intercept = 1,
                  sqrtldist = zincdf$sqrtldist)
Xcov <- as.matrix(Xcov)
colnames(Xcov)
## [1] "intercept" "sqrtldist"

# Create the stack
stack <- inla.stack(
  tag = "est",

```

```

# - Name (nametag) of the stack
# - Here: est for estimating
data = list(y = zincdf$y),
effects = list(
  # - The Model Components
  s = 1:mesh$n,
  Xcov = Xcov
),
# - The second is all fixed effects
A = list(A, 1)
# - First projector matrix is for 's'
# - second is for 'fixed effects'
)

```

Next, we can set a prior and define our model.

```

# Define a prior
prior.median.sd <- 0.07
prior.median.range <- 2000
spde <- inla.spde2.pcmatern(mesh, alpha = 2, prior.range = c(
  prior.median.range,
  0.5
), prior.sigma = c(prior.median.sd, 0.5), constr = T)

# Define the model
formula <- y ~ -1 + Xcov + f(s, model = spde)
prior.median.gaus.sd <- 1 # Prior for measurement error
family <- "gaussian"
control.family <- list(hyper = list(prec = list(
  prior = "pc.prec",
  fixed = FALSE, param = c(
    prior.median.gaus.sd,
    0.5
  )
))
))

```

Finally, we can fit our model:

```

res <- inla(formula,
  data = inla.stack.data(stack, spde = spde),
  control.predictor = list(A = inla.stack.A(stack), compute = T),
  # compute=T to get posterior for fitted values
  family = family,
  control.family = control.family,
  # control.compute = list(config=T, dic=T, cpo=T, waic=T),
  # if Model comparisons wanted
  control.inla = list(int.strategy = "eb"),
  # - faster computation
  # control.inla = list(int.strategy='grid'),
  # - More accurate integration over hyper-parameters
  verbose = F
)
summary(res)

```

```

##
## Call:
## c("inla.core(formula = formula, family = family, contrasts = contrasts,
##   ", " data = data, quantiles = quantiles, E = E, offset = offset, ", "
##   scale = scale, weights = weights, Ntrials = Ntrials, strata = strata,
##   ", " lp.scale = lp.scale, link.covariates = link.covariates, verbose =
##   verbose, ", " lincomb = lincomb, selection = selection, control.compute
##   = control.compute, ", " control.predictor = control.predictor,
##   control.family = control.family, ", " control.inla = control.inla,
##   control.fixed = control.fixed, ", " control.mode = control.mode,
##   control.expert = control.expert, ", " control.hazard = control.hazard,
##   control.lincomb = control.lincomb, ", " control.update =
##   control.update, control.lp.scale = control.lp.scale, ", "
##   control.pardiso = control.pardiso, only.hyperparam = only.hyperparam,
##   ", " inla.call = inla.call, inla.arg = inla.arg, num.threads =
##   num.threads, ", " blas.num.threads = blas.num.threads, keep = keep,
##   working.directory = working.directory, ", " silent = silent, inla.mode
##   = inla.mode, safe = FALSE, debug = debug, ", " .parent.frame =
##   .parent.frame)")
## Time used:
##   Pre = 4.21, Running = 0.864, Post = 0.0443, Total = 5.12
## Fixed effects:
##           mean      sd 0.025quant 0.5quant 0.975quant mode kld
## Xcov1  6.997 0.117      6.768   6.997      7.225   NA    0
## Xcov2 -2.579 0.240     -3.050  -2.579     -2.109   NA    0
##
## Random effects:
##   Name      Model
##   s SPDE2 model
##
## Model hyperparameters:
##
##           mean      sd 0.025quant 0.5quant
## Precision for the Gaussian observations 11.728  2.815      7.148  11.406
## Range for s                             582.939 173.761    314.523 558.496
## Stdev for s                             0.321  0.056      0.225  0.316
##
##           0.975quant mode
## Precision for the Gaussian observations 18.185   NA
## Range for s                             994.002   NA
## Stdev for s                             0.444   NA
##
## Marginal log-Likelihood: -94.24
## is computed
## Posterior summaries for the linear predictor and the fitted values are computed
## (Posterior marginals needs also 'control.compute=list(return.marginals.predictor=TRUE)')

```