# 2023 554 R Notes on Mapping for Point Data

## Jon Wakefield Departments of Biostatistics and Statistics University of Washington

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#### Overview

In these notes we will consider mapping and modeling of point data in which the (nominal) exact locations are known.

We will look at modeling a spatially-indexed continuous response via:

- Conventional Kriging via MLE and variants
- A generalized additive model (GAM)
- A Bayesian approach using stochastic partial differential equations (SPDE)

# Continuous Response: Motivating Example

We illustrate methods for continuous data using on Zinc levels in the Netherlands.

This data set gives locations and top soil heavy metal concentrations (in ppm), along with a number of soil and landscape variables, collected in a flood plain of the river Meuse, near the village Stein in the South of the Netherlands.

Heavy metal concentrations are bulk sampled from an area of approximately  $28 \text{km} \times 39 \text{km}$ .

The Meuse data are in a variety of packages. The version in the geoR library are not a spatial object, but can be used with likelihood and Bayes methods.

# Meuse analysis using geostat functions

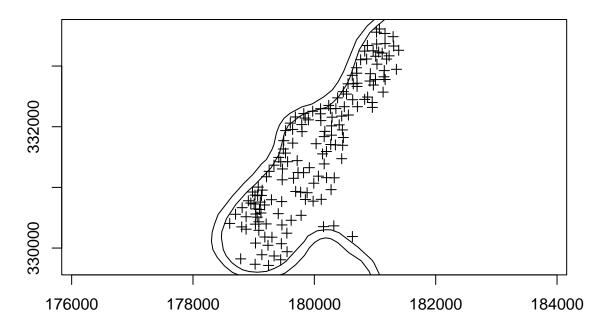
We look at the sampling locations and then examine variograms.

```
library(tidyverse)
library(ggpubr)
library(viridis)
library(geoR)
data("meuse")
library(sp)
pal <- function(n = 9){ brewer.pal(n, "Reds") }
data(meuse)</pre>
```

```
coords <- SpatialPoints(meuse[,c("x","y")])
meuse1 <- SpatialPointsDataFrame(coords,meuse)
data(meuse.riv)
river_polygon <- Polygons(list(Polygon(meuse.riv)),ID="meuse")
rivers <- SpatialPolygons(list(river_polygon))
coordinates(meuse) = ~x+y</pre>
```

### Zinc: Sampling locations

```
plot(meuse1,axes=T)
plot(rivers,add=T)
```



#### Exploratory analysis

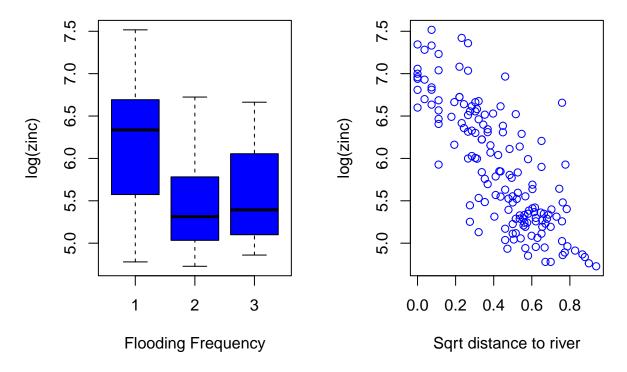
We work with log(zinc) as the distribution is more symmetric than on the original scale, and the variance more constant across levels of covariates.

It's often a good idea to do some exploratory data analysis (EDA), so let's see how  $\log(zinc)$  varies by two possible covariates:

- Flooding frequency (ffreq); 1 = once in two years; 2 = once in ten years; 3 = one in 50 years
- Distance to the Meuse river (dist); normalized to [0,1]

We focus on these, since they are available across the study region and so can be used for prediction. Following previous authors, we take sqrt(dist) which is closer to linearity.

```
par(mfrow=c(1,2))
plot(log(meuse$zinc)~meuse$ffreq,ylab="log(zinc)",xlab="Flooding Frequency",col="blue")
plot(log(meuse$zinc)~sqrt(meuse$dist),ylab="log(zinc)",xlab="Sqrt distance to river",col="blue")
```



# plot(log(meuse\$zinc)~meuse\$elev,ylab="log(zinc)",xlab="Elevation",col="blue")

Also map log(zinc) and these covariates.

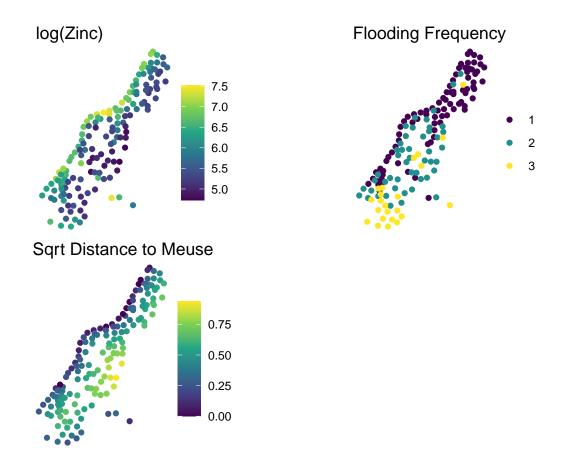
```
m.sf <- sf::st_as_sf(meuse, coords = c("x","y"))
m.sf$logzinc <- log(m.sf$zinc)
m.sf$sdist <- sqrt(m.sf$dist)

a <- ggplot() + geom_sf(data = m.sf[,"logzinc"], aes(color = logzinc)) +
    theme_void() + scale_color_viridis_c() + labs(title = "log(Zinc)", color=NULL)

b <- ggplot() + geom_sf(data = m.sf[,"ffreq"], aes(color = ffreq)) +
    theme_void() + scale_color_viridis(discrete=T) + labs(title = "Flooding Frequency", color=NULL)

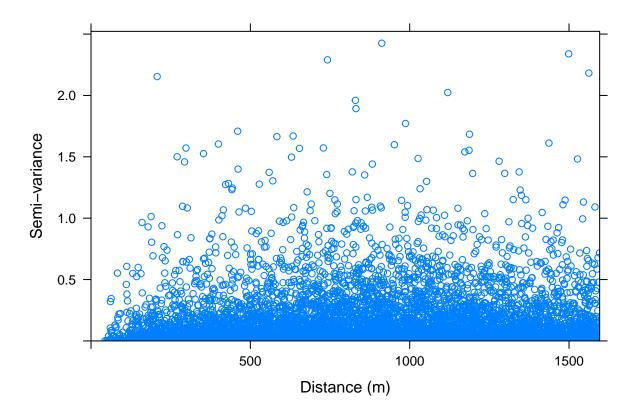
c <- ggplot() + geom_sf(data = m.sf[,"sdist"], aes(color = sdist)) +
    theme_void() + scale_color_viridis_c() + labs(title = "Sqrt Distance to Meuse", color=NULL)

# d <- ggplot() + geom_sf(data = m.sf[,"elev"], aes(color = elev)) + theme_void() + scale_color_viridis_ggpubr::ggarrange(a,b,c, nrow=2, ncol=2)</pre>
```



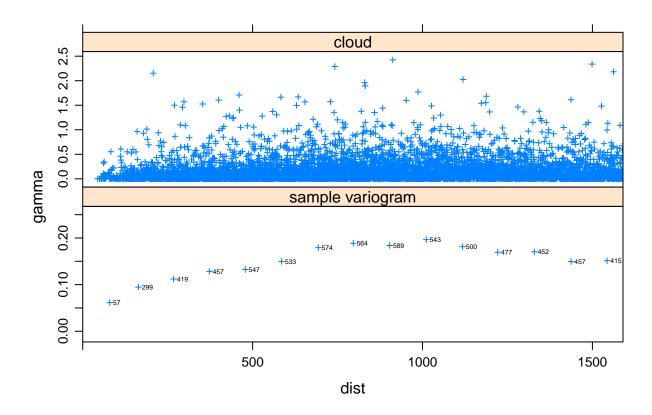
log(zinc): Variogram cloud, tremd removed

```
library(gstat)
cld <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse, cloud = TRUE)
plot(cld,ylab="Semi-variance",xlab="Distance (m)")</pre>
```



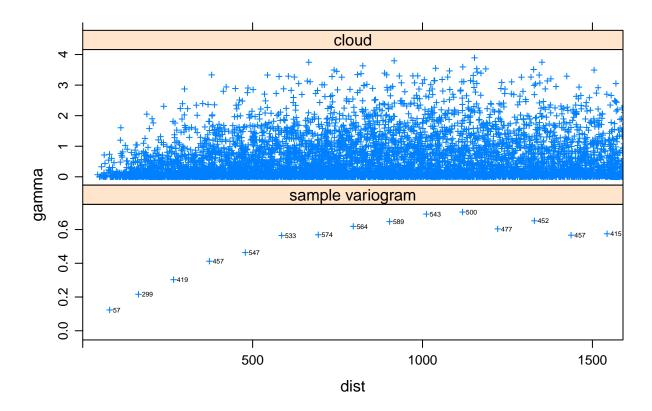
More variograms, with sample sizes

```
library(lattice)
cld <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse, cloud = TRUE)</pre>
svgm <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse)</pre>
d <- data.frame(gamma = c(cld$gamma, svgm$gamma),</pre>
    dist = c(cld$dist, svgm$dist),
    id = c(rep("cloud", nrow(cld)), rep("sample variogram", nrow(svgm)))
xyplot(gamma ~ dist | id, d,
    scales = list(y = list(relation = "free",
      #ylim = list(NULL, c(-.005, 0.25)))),
      limits = list(NULL, c(-.005,0.25)))),
    layout = c(1, 2), as.table = TRUE,
    panel = function(x,y, ...) {
        if (panel.number() == 2)
            ltext(x+10, y, svgmnp, adj = c(0,0.5), cex=.4) #$
        panel.xyplot(x,y,...)
    },
    xlim = c(0, 1590),
    cex = .5, pch = 3
)
```



## More variograms

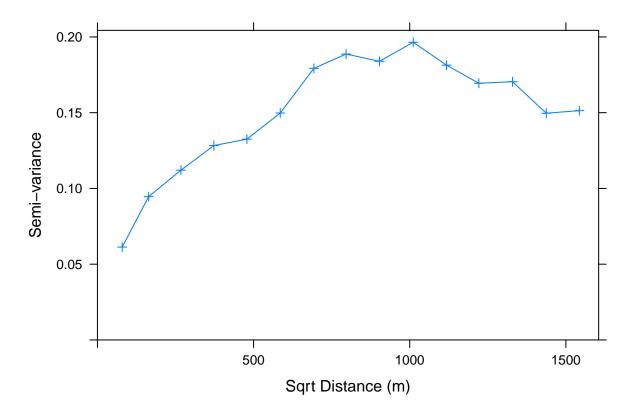
```
cld <- variogram(log(zinc) ~ 1, meuse, cloud = TRUE)</pre>
svgm <- variogram(log(zinc) ~ 1, meuse)</pre>
d <- data.frame(gamma = c(cld$gamma, svgm$gamma),</pre>
    dist = c(cld$dist, svgm$dist),
    id = c(rep("cloud", nrow(cld)), rep("sample variogram", nrow(svgm)))
xyplot(gamma ~ dist | id, d,
    scales = list(y = list(relation = "free",
      #ylim = list(NULL, c(-.005, 0.7)))),
      limits = list(NULL, c(-.005, 0.7)))),
    layout = c(1, 2), as.table = TRUE,
    panel = function(x,y, ...) {
        if (panel.number() == 2)
            ltext(x+10, y, svgmnp, adj = c(0,0.5), cex=.4) #$
        panel.xyplot(x,y,...)
    },
    xlim = c(0, 1590),
    cex = .5, pch = 3
```



## Monte Carlo simulations of semi-variogram

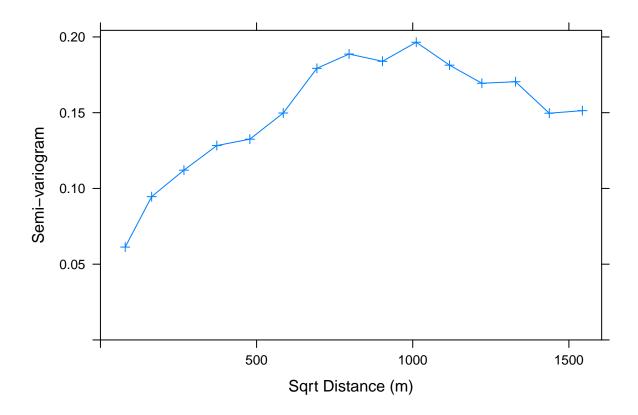
We simulate 100 datasets with random relabeling of points, and then form variograms for each.

```
v <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse)
plot(v, type = 'b', pch = 3,xlab="Sqrt Distance (m)",ylab="Semi-variance")
fn = function(n = 100) {
    for (i in 1:n) {
        meuse$random = sample(meuse$zinc)
        v = variogram(log(random) ~ 1, meuse)
        trellis.focus("panel", 1, 1, highlight = FALSE)
        llines(v$dist, v$gamma, col = 'grey')
        trellis.unfocus()
    }
}
fn()</pre>
```



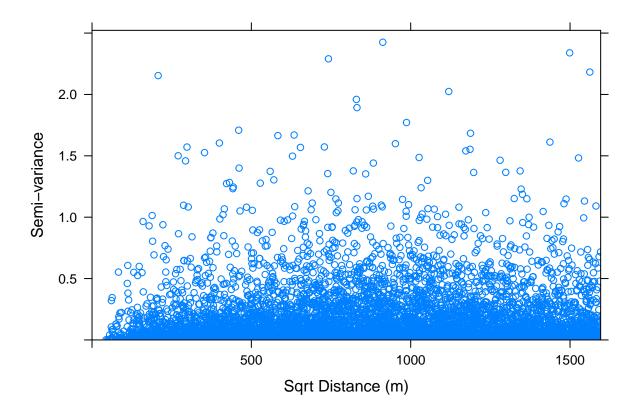
Monte Carlo envelopes under no spatial dependence - it is clear there is dependence here.

```
v <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(meuse$ffreq), meuse)
plot(v, type = 'b', pch = 3,xlab="Sqrt Distance (m)",ylab="Semi-variogram")
fn = function(n = 100) {
    for (i in 1:n) {
        meuse$random = sample(meuse$zinc)
        v = variogram(log(random) ~ 1, meuse)
        trellis.focus("panel", 1, 1, highlight = FALSE)
        llines(v$dist, v$gamma, col = 'grey')
        trellis.unfocus()
    }
}
fn()</pre>
```



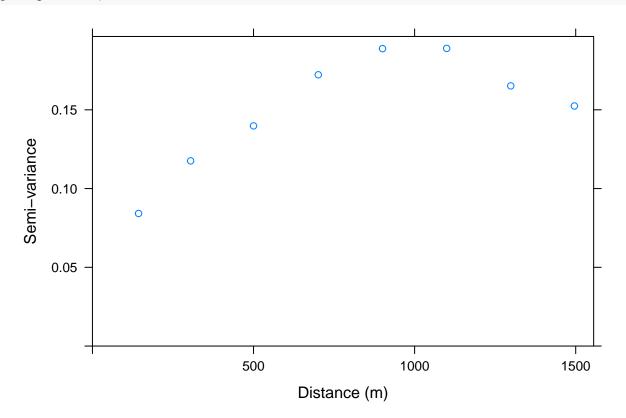
 $\log(\mathrm{zinc})$ : Variogram cloud, detrended

```
cld2 <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(ffreq), meuse, cloud = TRUE)
plot(cld2,ylab="Semi-variance",xlab="Sqrt Distance (m)")</pre>
```



Binned variogram for log(zinc), detrended

```
gstatbin <- variogram(log(zinc) ~ sqrt(meuse$dist)+as.factor(ffreq), meuse, width=200)
plot(gstatbin,ylab="Semi-variance",xlab="Distance (m)")</pre>
```



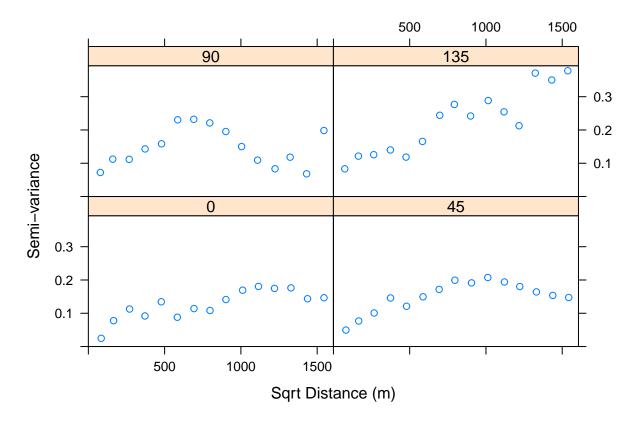
Directional variogram with linear trend removed

We form 4 variograms with data taken from different directions, with 0 and 90 corresponding to north and east, respectively.

Note that 0 is the same as 180.

```
dircld <- variogram(log(zinc)~sqrt(meuse$dist)+as.factor(ffreq), meuse, alpha=c(0,45,90,135))
```

Directional variogram with linear trend removed



### Other capabilities in gstat

See

- fit.variogram for estimation from the variogram
- krige (and associated functions) for Kriging,
- vgm generates variogram models

### geoR for geostatistics

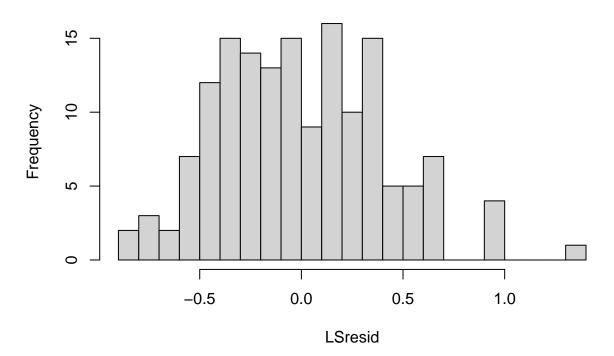
We continue the analysis using functions from the geoR library, for which a geodata data type is required. There are 155 observations (sampling locations)

# log(zinc) Variogram

First we will be assuming a spatial model on the residuals with ffreq and sdist in the model. Fit this initial linear model, and view the residuals by histogram and map.

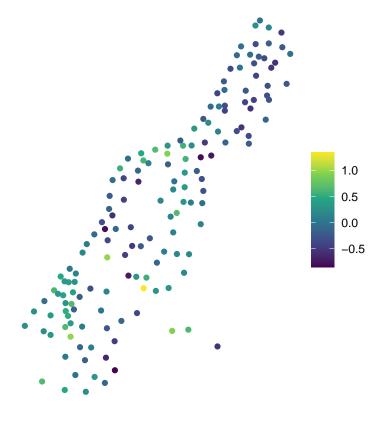
```
LSmod <- lm(log(meuse$zinc)~sqrt(meuse$dist)+as.factor(meuse$ffreq))
LSresid <- residuals(LSmod,type="pearson")
hist(LSresid,nclass=25)</pre>
```

# **Histogram of LSresid**



```
m.sf$resid <- LSresid
ggplot() + geom_sf(data = m.sf[,"resid"], aes(color = resid)) +
   theme_void() + scale_color_viridis_c() + labs(title = "Residual", color=NULL)</pre>
```

### Residual



#### Moran's I

We can calculate Moran's I to see the strength of spatial dependence in the residuals. Our results show that we can reject the null hypothesis that there is zero spatial autocorrelation present in LSresid at the  $\alpha=0.05$  level.

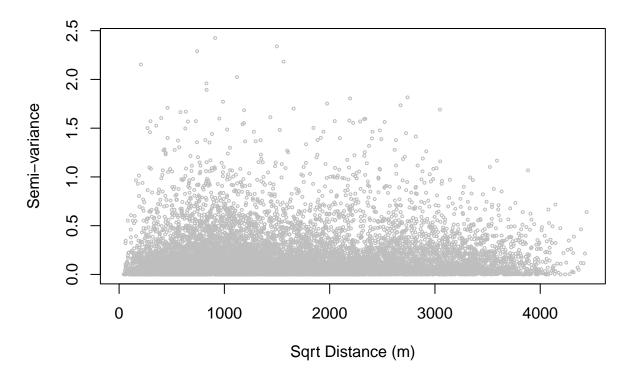
```
library(ape)
dists <- as.matrix(dist(cbind(geozinc[1]$coords[, 1], geozinc[1]$coords[, 2])))</pre>
dists.inv <- 1/dists</pre>
diag(dists.inv) <- 0</pre>
dists.inv[1:5, 1:5]
## 1 0.000000000 0.014116748 0.008414063 0.003857440 0.002729898
## 2 0.014116748 0.000000000 0.007063831 0.003535423 0.002757553
## 3 0.008414063 0.007063831 0.000000000 0.006984644 0.003983684
## 4 0.003857440 0.003535423 0.006984644 0.000000000 0.006482446
## 5 0.002729898 0.002757553 0.003983684 0.006482446 0.000000000
Moran.I(LSresid, dists.inv)
## $observed
## [1] 0.1210698
##
## $expected
## [1] -0.006493506
##
## $sd
```

```
## [1] 0.01060145
##
## $p.value
## [1] 0
```

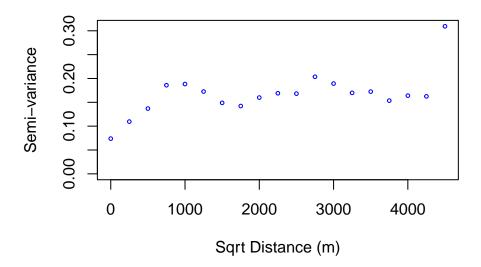
# Variogram Cloud

Variogram cloud for log zinc residuals, after taking out ffreq and dist.

```
cloudzinc <- variog(geozinc,option="cloud", trend=~sqrt(meuse$dist)+as.factor(meuse$ffreq))
## variog: computing omnidirectional variogram
plot(cloudzinc,ylab="Semi-variance",xlab="Sqrt Distance (m)",col="grey",cex=.4)</pre>
```



## Binned Variogram



## Maximum likelihood for log(zinc)

We fit with the Matern covariance model with

$$\rho(h) = \frac{1}{2^{\kappa - 1} \Gamma(\kappa)} \left( \frac{h}{\phi} \right)^{\kappa} K_{\kappa} \left( \frac{h}{\phi} \right),$$

where h is the distance between 2 points and we take  $\kappa = 0.5$  so that  $\phi$  is the only estimated parameter. The other parameters estimated are  $\tau^2$ , the nugget, and  $\sigma^2$ , the spatial variance. The practical range reported is the distance at which the correlations drop to 0.05, and is a function of  $\phi$ .

We suppress the output from the call.

```
mlfit <- likfit(geozinc,cov.model="matern",ini=c(.2,224),trend=~sqrt(meuse$dist)+as.factor(meuse$ffreq)</pre>
```

We examine the results, specifically point estimates and standard errors.

```
mlfit$parameters.summary
              status
                        values
##
## beta0
           estimated
                        7.0712
## beta1
           estimated
                       -2.1208
                       -0.5154
## beta2
           estimated
## beta3
           estimated
                       -0.5266
## tausq
           estimated
                        0.0372
## sigmasq estimated
                        0.1316
## phi
           estimated 300.5381
## kappa
               fixed
                        0.5000
## psiA
               fixed
                        0.0000
               fixed
                        1.0000
## psiR
## lambda
               fixed
                        1.0000
for (i in 1:3){
cat(cbind(mlfit$beta[i],sqrt(mlfit$beta.var[i,i])),"\n")
}
## 7.071249 0.1326244
## -2.120796 0.2365602
## -0.5153624 0.06780136
mlfit$practicalRange
## [1] 900.3318
```

Note that the standard errors change from the least squares fit.

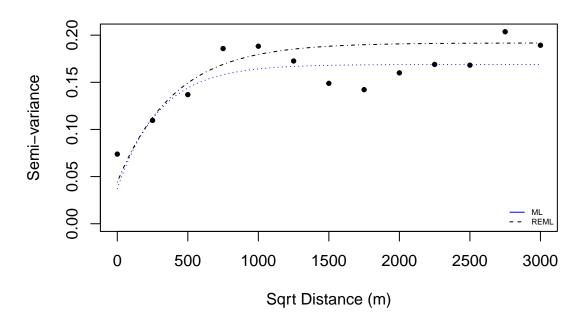
### Restricted maximum likelihood for log(zinc)

Restricted MLE is preferable in general for estimation when there are variance parameters.

The results show slight differences from ML.

```
remlfit$parameters.summary
##
             status values
## beta0 estimated 7.0823
## beta1 estimated -2.1109
## beta2 estimated -0.5280
## beta3 estimated -0.5455
## tausq estimated 0.0442
## sigmasq estimated 0.1476
## phi
        estimated 401.4313
             fixed 0.5000
## kappa
             fixed 0.0000
## psiA
## psiR
             fixed 1.0000
## lambda
             fixed 1.0000
remlfit$practicalRange
## [1] 1202.581
for (i in 1:3){
cat(cbind(mlfit$beta[i],sqrt(mlfit$beta.var[i,i]),remlfit$beta[i],sqrt(remlfit$beta.var[i,i])),"\n")
}
## 7.071249 0.1326244 7.082279 0.1527493
## -2.120796 0.2365602 -2.110917 0.2515499
## -0.5153624 0.06780136 -0.5280202 0.06913553
```

#### Comparison of estimates



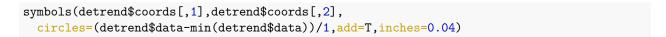
# Prediction for log(zinc) by Kriging

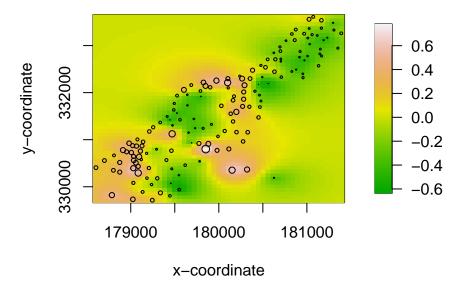
We now fit a linear model in distance and elevation to log(zinc). We then form a geodata object with the residuals as the response.

```
lmfit <- lm(geozinc$data ~ sqrt(meuse$dist) + as.factor(meuse$ffreq))</pre>
lmfit
##
## Call:
## lm(formula = geozinc$data ~ sqrt(meuse$dist) + as.factor(meuse$ffreq))
##
## Coefficients:
##
                                     sqrt(meuse$dist) as.factor(meuse$ffreq)2
                (Intercept)
##
                     7.0299
                                              -2.2660
                                                                         -0.3605
## as.factor(meuse$ffreq)3
                    -0.3167
detrend <- as.geodata(cbind(geozinc$coords,lmfit$residuals))</pre>
```

Finally, we can obtain spatial predictions on a grid, using the parameter estimates from the ML fit.

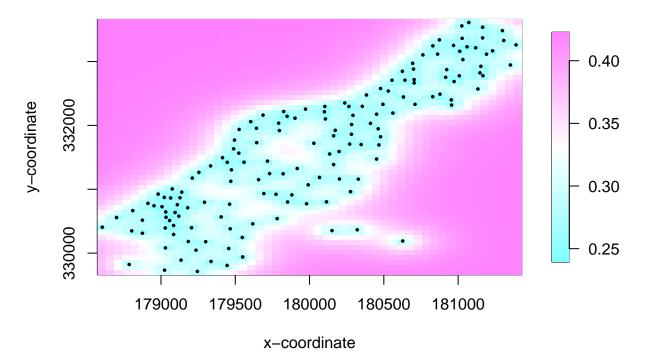
To view the results produce an image plot of the predictions, with the data superimposed.





### Standard deviations of prediction for log(zinc)

We now plot the Kriging standard deviations of the predictions.



The standard deviation is smallest close to the datapoints, as expected.

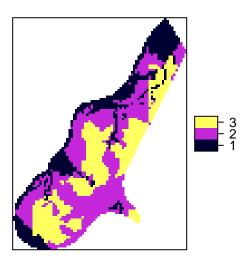
### Another example, using categorical flooding frequency covariate

Now, we'll repeat, but use categorical flooding frequency and distance to the Meuse River as covariates. These two covariates are available for a grid, in the meuse.grid data object. This means we can predict log(Zinc) for a grid, using the sum of the linear model fit and the interpolated spatial field.

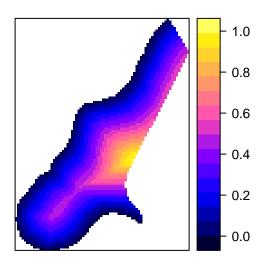
First take a look at the grid covariates.

```
library(sp)
data(meuse.grid)
coordinates(meuse.grid) = ~x+y
proj4string(meuse.grid) <- CRS("+init=epsg:28992")
gridded(meuse.grid) = TRUE</pre>
```

## spplot(meuse.grid["ffreq"])



#### spplot(meuse.grid["dist"])



Now proceed with Kriging for log(Zinc).

```
# (label objects associated with this version with a "3")
lmdata <- data.frame(logzinc=geozinc$data, dist=sqrt(meuse$dist), ffreq=as.factor(meuse$ffreq))</pre>
lmfit3 <- lm(logzinc~dist+ffreq, data=lmdata)</pre>
detrend3 <- as.geodata(cbind(geozinc$coords,lmfit3$residuals))</pre>
mlfit3 <- likfit(detrend3,ini=c(.2,224))</pre>
## -----
## likfit: likelihood maximisation using the function optim.
## likfit: Use control() to pass additional
           arguments for the maximisation function.
          For further details see documentation for optim.
## likfit: It is highly advisable to run this function several
          times with different initial values for the parameters.
## likfit: WARNING: This step can be time demanding!
## likfit: end of numerical maximisation.
# get meuse.grid and expand to all combinations
data(meuse.grid)
pred.grid3 <- expand.grid(sort(unique(meuse.grid$x)),</pre>
                          sort(unique(meuse.grid$y)))
# kriqinq
kc3 <- krige.conv(detrend3,loc=pred.grid3,</pre>
                  krige=krige.control(obj.m=mlfit3))
## krige.conv: model with constant mean
## krige.conv: Kriging performed using global neighbourhood
# merge linear predictions onto grid
pred_lm <- meuse.grid[, c("x", "y", "dist", "ffreq")]</pre>
pred_lm$pred_lm <- predict(lmfit3, newdata = pred_lm)</pre>
pred_resid <- data.frame(pred_resid = kc3$predict, x = pred.grid3$Var1,</pre>
                          y = pred.grid3$Var2)
pred_logzinc <- merge(pred_resid, pred_lm, by = c("x","y"), all=T)</pre>
pred_logzinc$pred_logzinc <- pred_logzinc$pred_lm + pred_logzinc$pred_resid</pre>
pred_logzinc <- pred_logzinc %>% arrange(y,x)
# plot prediction
image.plot(x=unique(pred.grid3[["Var1"]]), y=unique(pred.grid3[["Var2"]]),
           z=matrix(pred logzinc$pred logzinc,nrow=78,ncol=104),col=terrain.colors(100),
           xlab="x-coordinate",ylab="y-coordinate")
symbols(detrend3$coords[,1],detrend3$coords[,2],
        circles=(detrend3$data-min(detrend3$data))/1,add=T,inches=0.04)
```

