MATH 323: Probability

Anna Brandenberger

September 25, 2018

This is a transcript of the lectures given by Prof. David Stephens¹ during the fall semester of the 2018-2019 academic year (09-12 2018) for the Probability class (MATH 240). **Subjects covered** are: sample space, events, conditional probability, independence of events, Bayes' Theorem; basic combinatorial probability, random variables, discrete and continuous univariate and multivariate distributions; independence of random variables; inequalities, weak law of large numbers, central limit theorem.

¹ david.stephens@mcgill.ca, BH 1225

Basics of Probability

.1 Review of Set Theory

Definition 1.1. A **set** *S* is a collection of elements $s \in S$:

- FINITE: finite number of elements
- COUNTABLE: countably infinite number of elements
- UNCOUNTABLE: uncountably infinite number of elements

Definition 1.2. The **power set** $\mathcal{P}(A)$ of A is the set of all subsets of A: $\mathcal{P}(A) = \{B \mid B \subseteq A\}.$

SET OPERATIONS

- 1. INTERSECTION: $s \in A \cap B \iff s \in A \text{ and } s \in B$ extends to A_1, A_2, \ldots, A_K (finite): $s \in \bigcap_{k=1}^K A_k \iff s \in A_k \ \forall k$
- 2. UNION: $s \in A \cup B \iff s \in A \text{ or } s \in B$ extends to A_1, A_2, \ldots, A_K (finite²): $s \in \bigcup_{k=1}^K A_k \iff \exists k \text{ s.t. } s \in A_k$
- 3. Complement: for $A \subseteq S$, $s \in A' \iff s \in S$ but $s \notin A$
- 4. SET DIFFERENCE $A B \equiv A \setminus B \equiv A \cap B'$
- 5. EXCLUSIVE OR (XOR): $A \oplus B \equiv A \cup B A \cap B$

Definition 1.3. A_1, A_2, \ldots, A_K is called a **partition** of S if these sets are pairwise disjoint $A_j \cap A_k = \emptyset \ \forall j \neq k$ and exhaustive $\bigcup_{k=1}^K A_k = S$.

Theorem 1.1 (De Morgan's Laws).

$$A' \cap B' = (A \cup B)'$$
 $A \cup B' = (A \cap B)'$

Probability is a numerical value representing the chance of a particular event occurring given a particular set of circumstances.

Let $A \subseteq S$:

$$A \cap \emptyset = \emptyset$$
 $A \cup \emptyset = A$
 $A \cap S = A$ $A \cup S = S$

Intersection \cap and union \cup are commutative and associative.

² technically also extends to a countably infinite number of sets.

Proof. Let $A_1 = A \cup B$, $A_2 = A' \cap B'$. To show that $A_1' = A_2$, one must show $A_1 \cup A_2 = \emptyset$ and $A_1 \cup A_2 = S$.

 $= S \cap S = S$

$$A_1 \cap A_2 = (A \cup B) \cap A_2 = (A \cap A_2) \cup (B \cap A_2)$$

$$= (A \cap A' \cap B') \cup (B \cap A' \cap B')$$

$$= \emptyset \cup \emptyset = \emptyset$$

$$A_1 \cup A_2 = A_1 \cup (A' \cap B') = (A_1 \cup A') \cap (A_1 \cup B')$$

$$= (A \cup B \cup A') \cap (A \cup B \cup B')$$

1.2 Sample Spaces and Events

Definition 1.4. The set *S* of possible outcomes of an experiment is the **sample space** of an experiment, and the individual elements in *S* are **sample points** or **outcomes**.

Definition 1.5. An **event** A is a collection of sample outcomes $A \subseteq S$. Individual sample outcomes $E_1, E_2, \ldots, E_K, \cdots \in A$ are termed **simple** (**elementary**) events. We say that A **occurs** if the outcome $s \in A$. S is the **certain** event; \emptyset the **impossible** event.

Definition 1.6. The **probability function** P(.) assigns numerical values to events: $P: \mathcal{A} \longrightarrow \mathbb{R}$, $A \longrightarrow p$ (i.e. P(A) = p, $A = \mathcal{P}(\mathcal{S})$).

1.3 Axioms and Consequences

Theorem 1.2. Basic probability axioms:

- 1. $P(A) \ge 0$
- 2. P(S) = 1
- 3. *P* is countably additive: if A_1, A_2, \ldots s.t. $A_i \cap A_k = \emptyset \ \forall j \neq k$:

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i)$$

Corollary 1.3.
$$\forall A \ P'(A) = 1 - P(A)$$

Corollary 1.4. $P(\emptyset) = 0$

Corollary 1.5. $\forall A \ P(A) \leq 1$

Corollary 1.6. $\forall A \subseteq B, P(A) \leq P(B)$

Corollary 1.7 (General Addition Rule).

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

Using inductive arguments, one can construct a formula for $P(\bigcup_{i=1}^{n} A_i)$ (c.f. MATH 240) but we will mostly use:

Theorem 1.8 (Boole's Inequality).
$$P(\bigcup_{i=1}^n A_i) \leq \sum_{i=1}^n P(A_i)$$

As a method, one can lay out probabilities in table form such as in Table 1, ensuring that $0 \le p_{\star} \le 1$ and probabilities along rows and column add up correctly.

Proof.
$$S = A \cup A'$$
. By Theorem 1.2(3), $P(S) = P(A) + P(A') = 1$ $\implies P'(A) = 1 - P(A)$

Proof. Simple from Corollary 1.3. □

Proof. By Theorem 1.2 and 1.3,
$$1 = P(S) = P(A) + P(A') \ge P(A)$$

Proof. Take
$$B = A \cup (A' \cap B)$$
. Then, $P(B) = P(A) + P(A' \cap B) \ge P(A)$

Proof. First,
$$(A \cup B) = A \cup (A' \cap B)$$

 $\implies P(A \cup B) = P(A) + P(A' \cap B)$ [1
Then, take $B = (A \cap B) \cup (A' \cap B)$

$$\Rightarrow P(B) = P(A \cap B) + P(A' \cap B)$$
 [2] So taking [1] - [2], we get:

$$P(A \cup B) - P(B) = P(A) - P(A \cap B)$$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B) \quad \Box$$

	A	A'	Total
В	$p_{A\cap B}$	$p_{A'\cap B}$	p_B
B'	$p_{A\cap B'}$	$p_{A'\cap B'}$	$p_{B'}$
Total	p_A	$p_{A'}$	p_S

Table 1: Probabilities laid out in a table.

Specifying probabilities

Definition 1.7 (Equally likely sample outcomes). Suppose *S* finite, with sample outcomes E_1, \ldots, E_N , then $P(E_i) = \frac{1}{N} \ \forall i = 1, \ldots, N$. For an event $A \in S$, $\exists n \leq N$ s.t. $A = \bigcup_{i=1}^{n} E_{i_A}$, $i_A \in \{1, ..., n\}$ $\implies P(A) = \frac{n}{N} = \frac{\text{\# sample outcomes in } A}{\text{\# sample outcomes in } S}$

Definition 1.8 (Relative frequencies). For a given experiment with a sample space S and interest event A, consider a finite sequence of N repeat experiments and *n* the number of times that *A* occurs. The **fre**quencist definition of probability, generalizes the previous definition to $P(A) = \lim_{N \to \infty} \frac{n}{N}$, the relative frequency of appearance of A.

Definition 1.9 (Subjective assessment). For a given experiment with sample space S, the probability of event A is further generalized to a numerical representation of one's own personal degree of belief that the actual outcome lies in A, assuming one is **internally consistent**, i.e. rational and coherent.

Examples of experiments with equally likely sample outcomes include fair coins, dice and cards.

Combinatorial Rules

- 1. Multiplication principle: a sequence of k operations i, each with n_i possible outcomes, results in $n_1 \times n_2 \times \cdots \times n_k$ possible sequences of outcomes.
- 2. Selection principles: selecting from a finite set can be done:
 - WITH REPLACEMENT: each selection is independent;
 - WITHOUT REPLACEMENT: set is depleted by each selection.
- 3. Ordering: the order of a sequence of selections can be:
 - IMPORTANT (312 distinct from 123);
 - UNIMPORTANT (312 identical to 123).
- 4. DISTINGUISHABLE ITEMS: objects being selected can be:
 - DISTINGUISHABLE: individually labelled (e.g. lottery balls);
 - INDISTINGUISHABLE: labelled according to a type (e.g. colors).

Definition 1.10. A **permutation** is an ordered arrangement of *r* distinct objects. The number of possible permutations for a set *r* is $P_r^n = \frac{n!}{(n-r)!}$.

c.f. MATH 240 notes for examples.

Definition 1.11. The number of ways of partitioning n distinct elements into k disjoint subsets of sizes n_1, \ldots, n_k is the **multinomial coefficient** $N = \binom{n}{n_1, \ldots, n_k} \coloneqq \frac{n!}{n_1! \times n_2! \times \cdots \times n_k!}$. This is a generalization of the **binomial coeff.** $\binom{n}{j} = \frac{n!}{j!(n-j)!} \coloneqq \binom{n}{j,n-j}$.

Definition 1.12. The number of **combinations** C_r^n is the number of subsets of size r that can be selected from n objects $C_r^n = \binom{n}{r} = n! \cdot P_r^n$.

1.6 Conditional Probability and Independence

Definition 1.13. The **conditional probability** of *A* given *B* is $P(A \mid B) = \frac{P(A \cap B)}{P(B)}$ when P(B) > 0.

Corollary 1.9.
$$P(A | S) = P(A)$$
 $P(A | A) = 1$ $P(A | B) \le 1$

Warning 1. Important distinction between $P(A \cap B)$ and $P(A \mid B)$:

- $P(A \cap B)$: chance of A and B occurring RELATIVE TO S
- $P(A \mid B)$: chance of A and B occurring RELATIVE TO B non-symmetric: in general, $P(A \mid B) \neq P(B \mid A)$.

Theorem 1.10. The conditional probability function satisfies the probability axioms (Theorem 1.2).

Proof. 1.
$$P(A \mid B) = \frac{P(A \cap B)}{P(B)} \ge 0$$

2.
$$P(S \mid B) = \frac{P(S \cap B)}{P(B)} = \frac{P(B)}{P(B)} = 1$$

3.
$$P\left(\bigcup_{i=1}^{\infty}(A_i\mid B)\right) = \frac{P\left(\bigcup_{i=1}^{\infty}A_i\cap B\right)}{P(B)} = \frac{\sum_{i=1}^{\infty}P(A_i\cap B)}{P(B)} = \sum_{i=1}^{\infty}P(A_i\mid B)$$
 and since A_i disjoint $\forall i, (A_i, B)$ also disjoint.

Definition 1.14. Two events $A, B \in S$ are **independent** if $P(A \mid B) = P(A) \iff P(A \cap B) = P(A)P(B)$.

Remark 2. Notice then that $P(B \mid A) = \frac{P(A \cap B)}{P(A)} = \frac{P(A)P(B)}{P(A)} = P(B)$. *Warning* 3. Not the same as mutual exclusivity where $P(A \cap B) = 0$!

Warning 4. For multiple events $A_1, \ldots, A_n \in S$, we can talk about pairwise independence between any A_i, A_j with $i \neq j$, but this does not automatically generalize (c.f. Example 1.2).

Definition 1.15. Events $A_1, ..., A_K$ are **mutually independent** if $P(\bigcap_{k \in \mathcal{T}} A_k) = \prod_{k \in \mathcal{T}} P(A_k)$ for all subsets \mathcal{I} of $\{1, 2, ..., K\}$.

Given partial knowledge (B happened) i.e. knowing $s \in B \subseteq S$, one can make a more accurate probability assessment of some other event A, as $s \in A \cap B$.

Example 1.1. Given three two-sided cards, one RR, one RB and one BB, one card is selected randomly and one side displayed: it is red (R). Find the probability that the other side is red.

Answer. $S = \{R_1, R_2, B, R, B_1, B_2\}$ Card 1 (RR) selected $A = \{R_1, R_2\} \implies P(A) = 2/6$ Red side exposed: $B = \{R_1, R_2, R\} \implies P(B) = 3/6$ $P(A \mid B) = \frac{P(A \cap B)}{P(B)} = \frac{2/6}{3/6} = 2/3$

Example 1.2. Rolling two dice with independent outcomes A_1 : first roll outcome is odd; A_2 : second roll outcome is odd and A_3 : total outcome is odd. Then $P(A_1) = P(A_2) = \frac{1}{2}$ and $P(A_1 \mid A_3) = P(A_2 \mid A_3) = \frac{1}{2}$. But $P(A_1 \cap A_2 \cap A_3) = 0$, $P(A_1 \mid A_2 \cap A_3) = 0$, etc.

Definition 1.16. Consider events A_1 , A_2 , $B \in S$ with P(B) > 0. A_1 and A_2 are **conditionally independent** given B if $P(A_1 | A_2 \cap B) = P(A_1 | B) \iff P(A_1 \cap A_2 | B) = P(A_1 | B)P(A_2 | B)$

Theorem 1.11 (General Multiplication Rule).

$$P(A_1, A_2, ..., A_K) = P(A_1)P(A_2 | A_1)P(A_3 | A_1 \cap A_2) \cdots \cdots P(A_K | A_1 \cap A_2 \cap \cdots \cap A_{K-1})$$

We can use probability trees to display joint probabilities of multiples, where the junctions are the events and the branches correspond to possible sequences of choices of events. We then multiply along a branch to get the joint probability.

Important Theorems

Theorem 1.12 (Theorem of Total Probability). Given a partition B_1, \ldots, B_n of the sample space S into n subsets s.t. $P(B_i) > 0$, then for any event $A \in S$, $P(A) = \sum_{i=1}^{n} P(A \mid B_i) P(B_i)$.

Proof. Trivial, recursively (Chain Rule for probabilities).

Proof.
$$A = (A \cap B_1) \cup \cdots \cup (A \cap B_n)$$

$$P(A) = P(A \cap B_1) + \cdots + P(A \cap B_n)$$

$$= P(A \mid B_1)P(B_1) + \cdots$$

$$+ P(A \mid B_n)P(B_n)$$

$$= \sum_{i=1}^{n} P(A \mid B_i)P(B_i) \qquad \Box$$