

2021 中级计量经济学作业 3 参考答案

1. (渐进一致的权重矩阵 $\widehat{\mathbf{W}}_1$ 和 $\widehat{\mathbf{W}}_2$) 假设 $\widehat{\mathbf{W}}_1 - \widehat{\mathbf{W}}_2 \xrightarrow{p} \mathbf{0}$. 证明 GMM 估计量

$$\sqrt{n}\hat{\beta}(\widehat{\mathbf{W}}_1) - \sqrt{n}\hat{\beta}(\widehat{\mathbf{W}}_2) \xrightarrow{p} \mathbf{0}$$

解:

$$\sqrt{n}\hat{\beta}(\widehat{\mathbf{W}}_1) - \sqrt{n}\hat{\beta}(\widehat{\mathbf{W}}_2) = [(\mathbf{S}'_{\mathbf{XZ}}\widehat{\mathbf{W}}_1\mathbf{S}_{\mathbf{XZ}})^{-1}\mathbf{S}'_{\mathbf{XZ}}\widehat{\mathbf{W}}_1 - (\mathbf{S}'_{\mathbf{XZ}}\widehat{\mathbf{W}}_2\mathbf{S}_{\mathbf{XZ}})^{-1}\mathbf{S}'_{\mathbf{XZ}}\widehat{\mathbf{W}}_2] \sqrt{n}\bar{\mathbf{g}} \quad (1)$$

其中, 中括号部分依概率收敛于 0, $\sqrt{n}\bar{\mathbf{g}}$ 依分布收敛于正态分布, 因此上式依概率收敛于 0.

2. 下列说法是否正确? 为什么?

即使在过度识别时, 也可用如下方法找到方程 $\mathbf{g}_n(\hat{\beta}) \equiv \frac{1}{n} \sum_{i=1}^n \mathbf{z}_i(y_i - \mathbf{x}'_i\hat{\beta}) = \mathbf{0}$ 的解. 等式可写为

$$\mathbf{S}_{ZX}\hat{\beta} = \mathbf{S}_{Zy}, \quad (2)$$

其中 $\mathbf{S}_{ZX} = \frac{1}{n}\mathbf{z}_i\mathbf{x}'_i$, $\mathbf{S}_{Zy} = \frac{1}{n}\sum_{i=1}^n \mathbf{z}_iy_i$. 等式两边都乘以 \mathbf{S}'_{ZX} 得到

$$\mathbf{S}'_{ZX}\mathbf{S}_{ZX}\hat{\beta} = \mathbf{S}'_{ZX}\mathbf{S}_{Zy}, \quad (3)$$

因为 \mathbf{S}_{ZX} 满秩, $\mathbf{S}'_{ZX}\mathbf{S}_{ZX}$ 可逆, 因此得到

$$\hat{\beta} = (\mathbf{S}'_{ZX}\mathbf{S}_{ZX})^{-1}\mathbf{S}'_{ZX}\mathbf{S}_{Zy} \quad (4)$$

解: 有误。虽然 $\hat{\beta} = (\mathbf{S}'_{ZX}\mathbf{S}_{ZX})^{-1}\mathbf{S}'_{ZX}\mathbf{S}_{Zy}$ 是(3)式的解, 但 $\hat{\beta}$ 并不是(2)式的解。

3. 课本 271 页习题 15.3

1. 混合回归。

```
reg lwage exp exp2 wks ed, vce(cluster id)
```

```
. reg lwage exp exp2 wks ed, vce(cluster id)
```

```
Linear regression               Number of obs   =      4,165
                                F(4, 594)         =      72.58
                                Prob > F           =      0.0000
                                R-squared          =      0.2836
                                Root MSE       =      .39082
```

(Std. Err. adjusted for 595 clusters in id)

lwage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
exp	.044675	.0054385	8.21	0.000	.0339941	.055356
exp2	-.0007156	.0001285	-5.57	0.000	-.0009679	-.0004633
wks	.005827	.0019284	3.02	0.003	.0020396	.0096144
ed	.0760407	.0052122	14.59	0.000	.0658042	.0862772
_cons	4.907961	.1399887	35.06	0.000	4.633028	5.182894

2. 随机效应模型的 FGLS 估计以及检验

```
xtreg lwage exp exp2 wks ed, re r
xttest0
```

```
. xtreg lwage exp exp2 wks ed, re r
```

```
Random-effects GLS regression   Number of obs   =      4,165
Group variable: id              Number of groups =      595
```

```
R-sq:                               Obs per group:
    within = 0.6340                      min =      7
    between = 0.1716                     avg =     7.0
    overall = 0.1830                     max =      7
```

```
corr(u_i, X) = 0 (assumed)           Wald chi2(4)     =    1598.50
                                      Prob > chi2       =      0.0000
```

(Std. Err. adjusted for 595 clusters in id)

lwage	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
exp	.0888609	.0039992	22.22	0.000	.0810227	.0966992
exp2	-.0007726	.0000896	-8.62	0.000	-.0009481	-.000597
wks	.0009658	.0009259	1.04	0.297	-.000849	.0027806
ed	.1117099	.0083954	13.31	0.000	.0952552	.1281647
_cons	3.829366	.1333931	28.71	0.000	3.567921	4.090812
sigma_u	.31951859					
sigma_e	.15220316					
rho	.81505521	(fraction of variance due to u_i)				

```
. xttest0
```

Breusch and Pagan Lagrangian multiplier test for random effects

```
lwage[id,t] = Xb + u[id] + e[id,t]
```

Estimated results:

	Var	sd = sqrt(Var)
lwage	.2129935	.4615122
e	.0231658	.1522032
u	.1020921	.3195186

Test: Var(u) = 0

chibar2(01) = 5192.13
Prob > chibar2 = 0.0000

拒绝“不存在个体随机效应”的原假设。因此在 pooled OLS 和 RE 之间，应选择 RE。

3. 随机效应模型的 MLE 估计

```
xtreg lwage exp exp2 wks ed, mle
```

lwage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
exp	.1079955	.0024806	43.54	0.000	.1031335	.1128574
exp2	-.0005202	.0000546	-9.53	0.000	-.0006272	-.0004132
wks	.0008365	.0006042	1.38	0.166	-.0003477	.0020208
ed	.1378558	.0125933	10.95	0.000	.1131735	.1625382
_cons	2.989859	.1720638	17.38	0.000	2.65262	3.327097
/sigma_u	.8509013	.0278622			.7980078	.9073006
/sigma_e	.1536109	.0018574			.1500132	.1572949
rho	.9684385	.002199			.9638788	.9725117

LR test of sigma_u=0: chibar2(01) = 4576.13 Prob >= chibar2 = 0.000

exp 参数估计的比上面的 FGLS 结果更大。其他的检验结果差不多。

4. 固定效应模型组内估计量。在 stata 中可看到个体虚拟变量大多数显著异于 0(未截图)，因此可以认为个体效应存在。

```
xtreg lwage exp exp2 wks ed, fe vce(cluster id)
```

```

Fixed-effects (within) regression
Group variable: id

Number of obs   =    4,165
Number of groups =     595

R-sq:
    within = 0.6566
    between = 0.0276
    overall = 0.0476

Obs per group:
    min = 7
    avg = 7.0
    max = 7

F(3,594) = 1059.72
Prob > F = 0.0000
corr(u_i, Xb) = -0.9107

```

(Std. Err. adjusted for 595 clusters in id)

lwage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
exp	.1137879	.0040289	28.24	0.000	.1058753	.1217004
exp2	-.0004244	.0000822	-5.16	0.000	-.0005858	-.0002629
wks	.0008359	.0008697	0.96	0.337	-.0008721	.0025439
ed	0	(omitted)				
_cons	4.596396	.0600887	76.49	0.000	4.478384	4.714408
sigma_u	1.0362039					
sigma_e	.15220316					
rho	.97888036	(fraction of variance due to u_i)				

5. LSDV

```
reg lwage exp exp2 wks ed i.id, vce(cluster id)
```

```

. reg lwage exp exp2 wks ed i.id, vce(cluster id)
note: 595.id omitted because of collinearity

```

```

Linear regression
Number of obs   =    4,165
F(2, 594)       =          .
Prob > F        =          .
R-squared       =    0.9068
Root MSE       =    .1522

```

(Std. Err. adjusted for 595 clusters in id)

lwage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
exp	.1137879	.0043514	26.15	0.000	.1052418	.1223339
exp2	-.0004244	.0000888	-4.78	0.000	-.0005988	-.00025
wks	.0008359	.0009393	0.89	0.374	-.0010089	.0026806
ed	.1022134	.0046744	21.87	0.000	.093033	.1113938

个体的虚拟变量大多数都是显著异于 0 的，因此可以认为存在个体效应。另外可以发现 LSDV 和 FE 的估计参数是一样的。标准差不一样是因为 FE 和 LSDV 估

计的夹心估计量的表达式中的数据矩阵 \mathbf{X} 不一样，从而最终的对角线上的元素 (参数估计的方差) 不一样。一般在短面板中，个体的固定效应选择 FE。

6. 带时间效应的固定效应模型

```
xtreg lwage exp exp2 wks ed i.t, fe r
```

```
Fixed-effects (within) regression              Number of obs   =      4,165
Group variable: id                            Number of groups =       595

R-sq:                                         Obs per group:
    within = 0.6599                          min =           7
    between = 0.0275                         avg =          7.0
    overall = 0.0480                         max =           7

corr(u_i, Xb) = -0.9089                      F(8,594)        =      412.33
                                              Prob > F         =      0.0000
```

(Std. Err. adjusted for 595 clusters in id)

lwage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
exp	.1119927	.0041184	27.19	0.000	.1039043	.1200812
exp2	-.0004051	.0000834	-4.86	0.000	-.0005688	-.0002413
wks	.00068	.0008812	0.77	0.441	-.0010506	.0024105
ed	0	(omitted)				
t						
2	-.0083984	.0049321	-1.70	0.089	-.0180849	.0012881
3	.0259652	.0084359	3.08	0.002	.0093974	.0425329
4	.0289134	.0078093	3.70	0.000	.0135762	.0442506
5	.0239406	.0065275	3.67	0.000	.0111208	.0367604
6	.0069955	.0064617	1.08	0.279	-.0056949	.019686
7	0	(omitted)				
_cons	4.618339	.0599451	77.04	0.000	4.500609	4.736069
sigma_u	1.0268811					
sigma_e	.15159041					
rho	.97867247	(fraction of variance due to u_i)				

7. 一阶差分估计量。也可以用课本中提到的 xtserial

```
reg d.(lwage exp exp2 wks ed), noconst vce(cluster id)
```

```
. reg d.(lwage exp exp2 wks ed), noconst vce(cluster id)
```

note: D.ed omitted because of collinearity

```
Linear regression               Number of obs   =      3,570
                                F(3, 594)         =     1035.19
                                Prob > F           =      0.0000
                                R-squared           =      0.2209
                                Root MSE        =      .18156
```

(Std. Err. adjusted for 595 clusters in id)

D.lwage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
exp D1.	.1170654	.0040974	28.57	0.000	.1090182	.1251126
exp2 D1.	-.0005321	.0000808	-6.58	0.000	-.0006908	-.0003734
wks D1.	-.0002683	.0011783	-0.23	0.820	-.0025824	.0020459
ed D1.	0 (omitted)					

8. 组间估计量

```
xtreg lwage exp exp2 wks ed, be
```


10. 稳健 Hausman 检验

```
xtreg lwage exp exp2 wks ed, re r  
xtoverid
```

```
. xtoverid
```

```
Test of overidentifying restrictions: fixed vs random effects  
Cross-section time-series model: xtreg re robust cluster(id)  
Sargan-Hansen statistic 1792.412 Chi-sq(3) P-value = 0.0000
```

4. 考察 Stata 的 abdata 数据 (在 Stata 中输入 webuse abdata 即可获得数据)。数据中 n 为企业雇员数, w 是企业工资水平, k 为公司总资本, ys 为公司所在行业的总产出。所有变量均取了 log。变量名称后缀 L1, L2 表示滞后 1 阶, 2 阶。以 n 为因变量, $nL1$, $nL2$, w , $wL1$, k , $kL1$, $kL2$, ys , $ysL1$, $ysL2$, 以及时间虚拟变量 $yr1976, \dots, yr1984$ 为自变量, 做如下回归并比较和讨论结果:

1. OLS 回归。注意 $nL1$ 的系数估计大于 1, 从理论上讲, 滞后项系数应当小于 1, 否则 n 会越来越大, 不符合常识。由于未考虑到 fixed effect, n 和 $nL1$ 之间的相关性被放大了。也就是说, $nL1$ 的系数估计偏大。

```
reg n nL1 nL2 w wL1 k kL1 kL2 ys ysL1 ysL2 yr*
```


Source	SS	df	MS	Number of obs	=	751
Model	1343.31797	16	83.9573732	F(16, 734)	=	8136.58
Residual	7.57378164	734	.010318504	Prob > F	=	0.0000
				R-squared	=	0.9944
				Adj R-squared	=	0.9943
Total	1350.89175	750	1.801189	Root MSE	=	.10158

n	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
nL1	1.044643	.0336647	31.03	0.000	.9785523	1.110734
nL2	-.0765426	.0328437	-2.33	0.020	-.1410214	-.0120639
w	-.5236727	.0487799	-10.74	0.000	-.6194374	-.427908
wL1	.4767538	.0486954	9.79	0.000	.381155	.5723527
k	.3433951	.0255185	13.46	0.000	.2932972	.3934931
kL1	-.2018991	.0400683	-5.04	0.000	-.2805613	-.123237
kL2	-.1156467	.0284922	-4.06	0.000	-.1715826	-.0597107
ys	.4328752	.1226806	3.53	0.000	.1920285	.673722
ysL1	-.7679125	.1658165	-4.63	0.000	-1.093444	-.4423813
ysL2	.3124721	.111457	2.80	0.005	.0936596	.5312846
yr1976	0	(omitted)				
yr1977	0	(omitted)				
yr1978	-.0153956	.0230101	-0.67	0.504	-.060569	.0297779
yr1979	.0004932	.0219057	0.02	0.982	-.0425121	.0434986
yr1980	.0065977	.0222523	0.30	0.767	-.0370881	.0502835
yr1981	-.0375487	.0231813	-1.62	0.106	-.0830582	.0079608
yr1982	-.0304299	.0218943	-1.39	0.165	-.0734128	.0125529
yr1983	-.0080024	.0214113	-0.37	0.709	-.0500371	.0340323
yr1984	0	(omitted)				
_cons	.2901212	.3418808	0.85	0.396	-.3810596	.9613019

2. 固定效应回归。固定效应把 u_i 去掉了，但此时 $y_{i,t-1} - \overline{Ly}_i$ 会和 $\varepsilon_{it} - \bar{\varepsilon}_i$ 相关。
nL1 的系数此时小于 1。侧面印证之前 OLS 的估计过大。

```
xtreg n nL1 nL2 w wL1 k kL1 kL2 ys ysL1 ysL2 yr*, fe
```

```

Fixed-effects (within) regression               Number of obs   =       751
Group variable: id                             Number of groups =       140

R-sq:                                           Obs per group:
    within = 0.7973                             min =           5
    between = 0.9809                             avg =          5.4
    overall = 0.9758                             max =           7

corr(u_i, Xb) = 0.5459                         F(16,595)       =      146.27
                                                Prob > F        =      0.0000

```

n	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
nL1	.7329476	.039304	18.65	0.000	.6557563	.810139
nL2	-.1394773	.040026	-3.48	0.001	-.2180867	-.0608678
w	-.5597445	.057033	-9.81	0.000	-.6717551	-.4477339
wL1	.3149987	.0609756	5.17	0.000	.1952451	.4347522
k	.3884188	.0309544	12.55	0.000	.3276256	.4492119
kL1	-.0805185	.0384648	-2.09	0.037	-.1560618	-.0049751
kL2	-.0278013	.0328257	-0.85	0.397	-.0922695	.036667
ys	.468666	.1231278	3.81	0.000	.2268481	.7104839
ysL1	-.6285587	.15796	-3.98	0.000	-.9387856	-.3183318
ysL2	.0579764	.1345353	0.43	0.667	-.2062454	.3221982
yr1976	0	(omitted)				
yr1977	0	(omitted)				
yr1978	.0119152	.0261724	0.46	0.649	-.0394862	.0633167
yr1979	.0165714	.024721	0.67	0.503	-.0319795	.0651224
yr1980	.0231479	.0249969	0.93	0.355	-.0259449	.0722408
yr1981	-.013454	.0260231	-0.52	0.605	-.0645623	.0376542
yr1982	-.0224821	.0229812	-0.98	0.328	-.0676163	.0226521
yr1983	-.0161192	.0209498	-0.77	0.442	-.0572636	.0250253
yr1984	0	(omitted)				
_cons	1.780205	.5014522	3.55	0.000	.7953734	2.765036
sigma_u	.22568151					
sigma_e	.09395847					
rho	.85227336	(fraction of variance due to u_i)				

```

F test that all u_i=0: F(139, 595) = 1.89                Prob > F = 0.0000

```

3. Anderson-Hsiao 两阶段最小二乘法回归。AH 回归的结果理论上来说是 consistent，但滞后项的系数变为 2.3，远大于 1。结果仍然是很奇怪的。由此可以看出，模型设定是否和真实数据情况一样，是一个很难判断的问题。

```

ivregress 2sls d.n (d.nL1=nL2) d.(nL2 w wL1 k kL1 kL2 ys ///
ysL1 ysL2 yr1979 yr1980 yr1981 yr1982 yr1983)

```

Instrumental variables (2SLS) regression

Number of obs = 611
Wald chi2(15) = 89.93
Prob > chi2 = 0.0000
R-squared = .
Root MSE = .247

D.n	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
nL1						
D1.	2.307626	1.973193	1.17	0.242	-1.559762	6.175013
nL2						
D1.	-.2240271	.179043	-1.25	0.211	-.5749448	.1268907
w						
D1.	-.8103626	.261805	-3.10	0.002	-1.323491	-.2972342
wL1						
D1.	1.422246	1.179492	1.21	0.228	-.8895156	3.734007
k						
D1.	.2530975	.1447404	1.75	0.080	-.0305884	.5367835
kL1						
D1.	-.5524613	.6154929	-0.90	0.369	-1.758805	.6538825
kL2						
D1.	-.2126364	.2397909	-0.89	0.375	-.6826179	.2573451
ys						
D1.	.9905803	.4630105	2.14	0.032	.0830965	1.898064
ysL1						
D1.	-1.937912	1.438225	-1.35	0.178	-4.75678	.8809566
ysL2						
D1.	.4870838	.5099415	0.96	0.339	-.5123832	1.486551
yr1979						
D1.	.0467148	.0448599	1.04	0.298	-.0412089	.1346385
yr1980						
D1.	.0761344	.0624919	1.22	0.223	-.0463474	.1986163
yr1981						
D1.	.022623	.0557394	0.41	0.685	-.0866242	.1318701
yr1982						
D1.	.0127801	.0548402	0.23	0.816	-.0947048	.120265
yr1983						
D1.	.0099072	.0456113	0.22	0.828	-.0794894	.0993037
_cons	.0159337	.0273445	0.58	0.560	-.0376605	.0695279

4. 差分 GMM 回归。n 的滞后项为 GMM 形式工具变量，其他的工具（外生）变量为常规工具变量形式。注意 nL1 的系数估计较为合理。

```
xtabond2 n L.n L2.n w L.w L(0/2).(k ys) yr*, gmm(L.n) ///
iv(w L.w L(0/2).(k ys) yr*) nolevel robust
```

Dynamic panel-data estimation, one-step difference GMM

Group variable: id				Number of obs	=	611
Time variable : year				Number of groups	=	140
Number of instruments = 41				Obs per group: min	=	4
Wald chi2(19) = 1727.45				avg	=	4.36
Prob > chi2 = 0.000				max	=	6
n	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
n						
L1.	.6862261	.1445943	4.75	0.000	.4028266	.9696257
L2.	-.0853582	.0560155	-1.52	0.128	-.1951467	.0244302
w						
--.	-.6078208	.1782055	-3.41	0.001	-.9570972	-.2585445
L1.	.3926237	.1679931	2.34	0.019	.0633632	.7218842
k						
--.	.3568456	.0590203	6.05	0.000	.241168	.4725233
L1.	-.0580012	.0731797	-0.79	0.428	-.2014308	.0854284
L2.	-.0199475	.0327126	-0.61	0.542	-.0840631	.0441681
ys						
--.	.6085073	.1725313	3.53	0.000	.2703522	.9466624
L1.	-.7111651	.2317163	-3.07	0.002	-1.165321	-.2570095
L2.	.1057969	.1412021	0.75	0.454	-.1709542	.382548
yr1976	0	(omitted)				
yr1977	0	(omitted)				
yr1978	0	(omitted)				
yr1979	.0095545	.0102896	0.93	0.353	-.0106127	.0297217
yr1980	.0220152	.0177104	1.24	0.214	-.0126966	.056727
yr1981	-.0117743	.0295079	-0.40	0.690	-.0696086	.04606
yr1982	-.0270588	.0292751	-0.92	0.355	-.0844369	.0303193
yr1983	-.0213204	.0304599	-0.70	0.484	-.0810207	.0383798
yr1984	-.0077033	.0314106	-0.25	0.806	-.069267	.0538604

需要注意扰动项是否存在 2 阶自相关。这里的结果认为不存在 2 阶自相关。符合差分 GMM 的假定。

```
Arellano-Bond test for AR(1) in first differences: z = -3.60 Pr > z = 0.000
Arellano-Bond test for AR(2) in first differences: z = -0.52 Pr > z = 0.606
```

过度识别检验。Hansen test 认为可以在 5% 水平下接受所有工具变量外生。

```
Sargan test of overid. restrictions: chi2(22) = 67.59 Prob > chi2 = 0.000
(Not robust, but not weakened by many instruments.)
Hansen test of overid. restrictions: chi2(22) = 31.38 Prob > chi2 = 0.089
(Robust, but weakened by many instruments.)
```

5. 系统 GMM 回归。此时 nL1 系数符号又大于 1 了。

```
xtabond2 n L.n L2.n w L.w L(0/2).(k ys) yr*, gmm(L.n) ///
iv(w L.w L(0/2).(k ys) yr*) robust twostep
```

Dynamic panel-data estimation, two-step system GMM

Group variable: id	Number of obs	=	751
Time variable : year	Number of groups	=	140
Number of instruments = 49	Obs per group: min	=	5
Wald chi2(19) = 75477.28	avg	=	5.36
Prob > chi2 = 0.000	max	=	7

n	Coef.	Corrected Std. Err.	z	P> z	[95% Conf. Interval]	
n						
L1.	1.057271	.0572935	18.45	0.000	.9449778	1.169564
L2.	-.098028	.0368302	-2.66	0.008	-.1702138	-.0258421
w						
--.	-.537141	.1667007	-3.22	0.001	-.8638684	-.2104137
L1.	.5011521	.1765282	2.84	0.005	.1551631	.847141
k						
--.	.2996729	.0601121	4.99	0.000	.1818554	.4174904
L1.	-.1453414	.0788808	-1.84	0.065	-.2999449	.0092621
L2.	-.1194829	.0424068	-2.82	0.005	-.2025988	-.0363671
ys						
--.	.5180324	.1996673	2.59	0.009	.1266916	.9093731
L1.	-.7292613	.246855	-2.95	0.003	-1.213088	-.2454344
L2.	.255183	.1481775	1.72	0.085	-.0352395	.5456055
yr1976	0	(omitted)				
yr1977	0	(omitted)				
yr1978	.01419	.0312493	0.45	0.650	-.0470575	.0754374
yr1979	.0290609	.0273843	1.06	0.289	-.0246113	.0827332
yr1980	.0421484	.0197262	2.14	0.033	.0034857	.080811
yr1981	0	(omitted)				
yr1982	.0075506	.0178061	0.42	0.672	-.0273487	.0424499
yr1983	.0452061	.0240706	1.88	0.060	-.0019714	.0923835
yr1984	.0275589	.0279575	0.99	0.324	-.0272369	.0823547
_cons	-.0701535	.3408083	-0.21	0.837	-.7381255	.5978186

6. 讨论：滞后项的系数估计可以提供估计是否合理的证据。如果模型是稳定的，那么滞后项的系数估计一般应当小于 1。另外，在这组数据中，工资 w 对 n 的影响是负的，尽管在不同的模型中估计的参数不一样。 k ，资本的影响是正的。动态面板估计的工具变量可以调整的空间很大，不同的工具变量设置有时会带来很不一样的估计结果。有兴趣的同学可以自行设置不同的工具变量，不同的滞后阶数来观察结果的异同。