

① Given

$$x_1, \dots, x_m \stackrel{iid}{\sim} N(\mu_1, \sigma_1^2) \quad > \text{indp.}$$

$$y_1, \dots, y_n \stackrel{iid}{\sim} N(\mu_2, \sigma_2^2)$$

$$\text{Then } \bar{x} \sim N(\mu_1, \sigma_1^2/m)$$

$$\bar{y} \sim N(\mu_2, \sigma_2^2/n)$$

$$\bar{x} - \bar{y} \sim N(\mu_1 - \mu_2, \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n})$$

$$\frac{(m-1)s_1^2}{\sigma_1^2} \sim \chi_{m-1}^2, \quad \frac{(n-1)s_2^2}{\sigma_2^2} \sim \chi_{n-1}^2$$

$$\Rightarrow \frac{(m-1)s_1^2}{\sigma_1^2} + \frac{(n-1)s_2^2}{\sigma_2^2} \sim \chi_{m+n-2}^2$$

Hence

$$\frac{\bar{x} - \bar{y} - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} \sim \frac{(m-1)s_1^2 + (n-1)s_2^2}{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}} \sim t_{m+n-2}$$

$$\Rightarrow \frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{\sqrt{\frac{(m-1)s_1^2}{\sigma_1^2} + \frac{(n-1)s_2^2}{\sigma_2^2}}} \sim t_{m+n-2}$$

$$\textcircled{2} \quad X_1, \dots, X_n, X_{n+1} \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$$

$$\bar{X} \sim N(\mu, \sigma^2/n), \text{ where } \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

$$\frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1} \quad \text{where } s^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

$$X_{n+1} - \bar{X} \sim N(0, \sigma^2 \left(\frac{n+1}{n}\right))$$

$$S_0 \quad \frac{X_{n+1} - \bar{X}}{\sqrt{\sigma^2 \left(\frac{n+1}{n}\right)}} \sim \frac{\frac{(n-1)s^2}{\sigma^2}}{n-1} \sim t_{n-1}$$

$$\Rightarrow \left(\frac{X_{n+1} - \bar{X}}{S} \right) \sqrt{\frac{n}{n+1}} \sim t_{n-1}$$

$$\textcircled{3} \quad X_1, \dots, X_m \stackrel{\text{iid}}{\sim} N(\mu_1, \sigma^2) \quad Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} N(\mu_2, \sigma^2) \quad \text{y indep}$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{m}\right), \quad \bar{Y} \sim N\left(\mu_2, \frac{\sigma^2}{n}\right)$$

$$\alpha \bar{X} + \beta \bar{Y} \sim N\left(\alpha \mu_1 + \beta \mu_2 + \sigma^2 \left(\frac{\alpha^2}{m} + \frac{\beta^2}{n} \right)\right)$$

$$\Rightarrow \frac{\alpha(\bar{X} - \mu) + \beta(\bar{Y} - \mu_2)}{\sigma \sqrt{\frac{\alpha^2}{m} + \frac{\beta^2}{n}}} \sim N(0, 1)$$

$$\frac{(m-1) S_1^2}{\sigma^2} + \frac{(n-1) S_2^2}{\sigma^2} \sim \chi^2_{m+n-2}$$

$$\Rightarrow \frac{\alpha(\bar{X} - \mu_1) + \beta(\bar{Y} - \mu_2)}{\sqrt{\frac{(m-1) S_1^2 + (n-1) S_2^2}{m+n-2} \left(\frac{\alpha^2}{m} + \frac{\beta^2}{n} \right)}} \sim t_{m+n-2}$$

④ Let X_1, \dots, X_5 denote the life of tires. Then

$$X_1, \dots, X_5 \stackrel{\text{iid}}{\sim} N(38500, 2500^2)$$

$$\bar{X} \sim N\left(38500, \frac{2500^2}{5}\right)$$

Required prob is

$$P(\bar{X} < 36000)$$

$$= P\left(\frac{\bar{X} - 38500}{\sqrt{\frac{2500^2}{5}}} < \frac{36000 - 38500}{\sqrt{\frac{2500^2}{5}}}\right)$$

$$= P(Z < -\sqrt{5}) = P(Z < -2.234)$$

$$= 0.0129.$$

⑤ Let the population random variable is X . Population mean is μ , population s.d. is σ .

$$\text{Given } P(|\bar{X} - \mu| \leq 25\% \sigma) = 0.95$$

$$\Rightarrow P(|\bar{X} - \mu| \leq 0.25\sigma) = 0.95$$

$$\Rightarrow P\left(\left|\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}\right| \leq 0.25\sqrt{n}\right) = 0.95$$

As the sample size is sufficiently large

$$\Rightarrow P(|Z| \leq 0.25\sqrt{n}) = 0.95$$

$$\Rightarrow P(-0.25\sqrt{n} \leq Z \leq 0.25\sqrt{n}) = 0.95$$

i.e. $P(Z \geq 0.25\sqrt{n}) = 0.025$

or $P(Z \leq -0.25\sqrt{n}) = 0.025$

From the table we have

$$0.25\sqrt{n} = 1.96$$

$$\Rightarrow n \approx 64.$$

⑥ $x_1, \dots, x_n \stackrel{iid}{\sim} \text{Exp}(\mu, 1)$

We know the density of $x_{(1)}$ is

$$f_{X_{(1)}}(y) = \begin{cases} ne^{-n(y-\mu)}, & y > \mu \\ 0, & \text{otherwise} \end{cases}$$

$$P(X_{(1)} > 3\mu) = \int_{3\mu}^{\infty} ne^{-n(y-\mu)} dy$$

$$= e^{-2n\mu}.$$

⑦ Given $X \sim N(68, 6.25)$

Let n be the sample size.

$$\bar{X} \sim N(68, 6.25/n)$$

Required prob is

$$P(\bar{X} - 68 \leq 1) \geq 0.95$$

$$\Rightarrow P\left(\frac{\bar{X} - 68}{\sqrt{\frac{6.25}{n}}} \leq \frac{\sqrt{n}}{6.25}\right) \geq 0.95$$

$$\Rightarrow P\left(Z \leq \frac{\sqrt{n}}{6.25}\right) \geq 0.95$$

From the table value

$$\frac{\sqrt{n}}{6.25} = 1.65$$

$$n \approx 17.$$

⑧ Let $x_1, \dots, x_{25} \sim N(\mu, 6)$

Required pmb is

$$P(S^2 > 9.105)$$

$$= P\left(\frac{24}{6}S^2 > \frac{24 \times 9.105}{6}\right)$$

$$= P(\chi^2_{24} > 36.42) = 0.05$$

From the chisquare table.

⑨ Given $n_1 = 5, n_2 = 4$.

Required pmb is

$$P\left(\frac{s_1^2}{s_2^2} < \frac{1}{5.2} \text{ or } > 6.5\right)$$

$$= 1 - P \left(\frac{1}{S_2} \leq \frac{S_1^2}{S_2^2} \leq 6.5 \right)$$

$$= 1 - P \left(\frac{1}{S_2} \leq F_{4,3} \leq 6.25 \right)$$

$$\left[\frac{4S_1^2}{\sigma^2} \sim \chi^2_4, \quad \frac{3S_2^2}{\sigma^2} \sim \chi^2_3 \right]$$

$$= 1 - P \left(0.1923 < F_{4,3} \leq 6.25 \right)$$

$$= 1 - [P(F_{4,3} \leq 6.25) - P(F_{4,3} \leq 0.1923)]$$

This values can be obtained from the

table for F -distⁿ

$$① 8 \quad X_1, \dots, X_5 \stackrel{iid}{\sim} N(2.5, 36)$$

$$② a \quad P(30 \leq S^2 \leq 40)$$

$$= P\left(\frac{30 \times 4}{36} \leq \chi^2_4 \leq \frac{40 \times 4}{36}\right) \because \frac{4S^2}{36} \sim \chi^2_4$$

$$= P(3.3 \leq \chi^2_4 \leq 4.4)$$

$$= P(\chi^2_4 \geq 3.3) - P(\chi^2_4 \geq 4.4)$$

$$= 0.5 - 0.35 = 0.15 \text{ (approx)} \\ \text{(From table).}$$

We have $\bar{X} \sim N(2.5, \frac{36}{5})$

$$③ b \quad P(1.3 \leq \bar{X} \leq 3.5, 30 \leq S^2 \leq 44)$$

$$= P(1.3 \leq \bar{X} \leq 3.5) P(30 \leq S^2 \leq 44)$$

$[\bar{X} \text{ & } S^2 \text{ are indep}]$

$$= 0.15 P(1.3 \leq \bar{X} \leq 3.5)$$

$$= 0.15 P\left(\frac{1.3 - 2.5}{6/\sqrt{5}} \leq Z \leq \frac{3.5 - 2.5}{6/\sqrt{5}}\right)$$

$$= 0.15 P(-0.4472 \leq Z \leq 0.372)$$

$$= 0.15 [\Phi(0.372) - \Phi(-0.4472)]$$

$$= 0.15 [0.64431 - 0.32997]$$

$$= 0.15 \times 0.31441.$$

(ii) The joint distⁿ of (x_1, x_2) is

$$f_{\underline{x}}(x_1, x_2) = \begin{cases} \frac{1}{\Gamma(n_1)\Gamma(n_2)} \frac{x_1^{n_1-1} x_2^{n_2-1}}{\theta^{n_1+n_2}} e^{-\frac{x_1+x_2}{\theta}}, & x_1 > 0, x_2 > 0 \\ 0, & \text{otherwise} \end{cases}$$

Hence $S_{\underline{x}} = (0, \infty)^2$. Let

$$h_1(x_1, x_2) = y_1 = x_1 + x_2, \quad h_2(x_1, x_2) = y_2 = \frac{x_1}{x_1 + x_2}$$

$$x_1 = y_1, y_2, \quad x_2 = y_1(1-y_2)$$

$$J = \begin{vmatrix} y_2 & y_1 \\ 1-y_2 & -y_1 \end{vmatrix} = -y_1$$

$$y_1 > 0, 0 < y_2 < 1. \quad S_0 \subseteq \mathbb{R}^2 = (0, \infty) \times (0, 1)$$

The joint pdf of (x_1, x_2) is

$$f_{\underline{y}}(y_1, y_2) = \frac{(y_1 y_2)^{n_1-1} (y_1(1-y_2))^{n_2-1}}{\Gamma(n_1) \Gamma(n_2) \theta^{n_1+n_2}} e^{-\frac{y_1 y_2 + y_1(1-y_2)}{\theta}}$$

$$| -y_1 | \times$$

$$\mathbb{I}_{(0, \infty) \times (0, 1)}$$

$$= \left\{ \frac{e^{-y_1/\theta} y_1^{n_1+n_2-1}}{\Gamma(n_1+n_2)} \mathbb{I}_{(0, \infty)}(y_1) \right\} \times$$

$$\left\{ \frac{1}{B(n_1, n_2)} y_1^{n_1-1} (1-y_1)^{n_2-1} I_{(0,1)}(y_1) \right\}$$

$$= f_{Y_1}(y_1) f_{Y_2}(y_2)$$

where $Y_1 \sim \text{Gamma}(n_1 + n_2, \theta)$

$Y_2 \sim \text{Beta}(n_1, n_2)$

clearly y_1 & y_2 are indep.