

# Thank you, next: demographic consequences of partner diversity and turnover in a multi-species ant-plant mutualism

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## Abstract

The diversity of partners in a multi-species mutualism causes varied demographic effects on the population of the focal mutualist which can be explained by several mechanisms: portfolio effect, complementarity, and sampling effect. *Cylindropuntia imbricata* (tree cholla) produce extrafloral nectar and various ant partners provide defense from herbivores and seed predators. We used plant demographic censuses to parameterize a series of Bayesian hierarchical generalized linear vital rate models to determine the impacts of different partners on the focal mutualists. We constructed an Integral Projection Model in which we simulate different combinations of ant partners that don't occur in nature. The hierarchical models revealed that different ant partners had different impacts on the cholla vital rates. Specifically, *C. opuntiae* tended plants had advantages in both growth and survival when small, and large *L. apiculatum* tended plants had floral viability advantages. The IPM results revealed that despite these differences in vital rates, the presence of any ant partner is beneficial, the identity and number of partners do not matter. This suggests that there are no benefits of diversity in this system. This study highlights that partner diversity does not always increase the benefits a focal mutualist receives.

## Introduction

Mutualisms are species interactions where all participants receive net benefits, leading to higher individual fitness and increased population growth rates. They are widespread species interactions but can deteriorate into commensalism or parasitism under conditions that elevate costs or dampen benefits (Rodriguez-Rodriguez et al., 2017; Song et al., 2020; Mandyam and Jumpponen, 2014; Thrall et al., 2007; Bahia et al., 2022; Bronstein, 1994; Chamberlain et al., 2014; Frederickson, 2013; Axelrod and Hamilton, 1981; Leigh, 2010). Mutualisms are considered more context dependent than other species interactions (Chamberlain et al., 2014; Frederickson, 2013), meaning the magnitude and sign of interaction strength are often determined by environmental conditions and species' identities (Noe and Hammerstein, 1994; Leigh, 2010).

Mutualism is defined at the level of a species pair (+/+) but these interactions are embedded within multi-species communities, and growing evidence suggests that pairwise interactions are poor predictors of the net effects of multi-species mutualism (Afkhami, 2014; Palmer et al., 2010; Bascompte, 2009; Dattilo et al., 2014). A focal mutualist may interact with multiple guilds of partner types (e.g., plants that interact with pollinators, seed dispersers, soil microbes, and ant defenders) or with multiple partner species within the same guild (e.g., plants visited by multiple pollinator species). Within a mutualist guild, partner species often differ in the amount or type of goods or services they provide, making partner identity an important source of contingency in mutualism (Stanton, 2013). Whether and how partner diversity modifies the demographic effects of mutualistic interactions remain open questions within relevance in applied settings (Rogers et al., 2014; Thibaut et al., 2012; Frederickson, 2005; Palmer et al., 2010).

There are multiple mechanisms by which partner diversity can influence the net benefits accrued by a focal mutualist, mirroring the mechanisms by which, at a larger scale of organization, biodiversity can influence ecosystem function (Yeung et al. (2006); Barrett et al. (2015); Ushio (2020)). First, when there is a hierarchy of fitness effects (a consistent ranking of best to worst mutualists), a more diverse sample of the partner community may be more likely to include the best partner

Frederickson (2013). This can lead to an apparent benefit of diversity driven by a sampling effect Batstone (2018) – but there is no benefit of diversity *per se*, only better and worse partners. If partner associations are mutually exclusive (a focal mutualist may engage with only one partner at a time), then partner diversity may impose opportunity costs, leading to negative effects of a diverse mutualist assemblage relative to exclusive association with the single best partner (Miller, 2007). Second, even within a single mutualist guild, the benefits conferred by alternative partner species can vary in type and not just degree Stachowicz and Whitlatch (2005); Bronstein et al. (2006); Stanton (2013). This can lead to a positive effect of partner diversity through complementarity of alternative functions Batstone (2018). Interference or synergies between partners can make their combined effect different than the expected from the sum of complementary functions Afkhami (2014). Third, partner species can have species-specific responses to environmental variation, either spatially (Ollerton et al., 2006) or temporally (Alarcón et al., 2008). Multiple partners can therefore act as a 'portfolio' that stabilizes fitness benefits across spatial or temporal heterogeneity, leading to positive effects of partner diversity through the portfolio effect Batstone (2018); Lázaro et al. (2022); Horvitz and Schemske (1990).

Partner diversity can have different effects depending on whether partners are present simultaneously or sequentially (partner turnover) (Djiéto-Lordon et al., 2005; Ness et al., 2006; Bruna et al., 2014; Barrett et al., 2015; Ushio, 2020; Dattilo et al., 2014). Sequential associations are likely when alternative partners engage in interference competition for access to a shared mutualist Kiers et al. (2003); Batstone (2018); Trøjelsgaard et al. (2015); Wulff (2008). Turnover can happen at different timescales, from minutes to years (Oliveira et al., 1999; Horvitz and Schemske, 1986). The frequency of partner turnover can impact the level of benefits received by the focal mutualist, particularly if the benefits continue to accumulate with successive turnover (e.g., when sequential partners provide complementary functions) or if they saturate over time (Sachs et al., 2004; Fiala et al., 1994). Directionality of turnover can also influence effects of partner diversity if partner identity changes consistently across ontogeny of a focal mutualist (Fonseca et al., 2003; Noe and Hammerstein, 1994; Dejean et al., 2008). For example, plant susceptibility to enemies can change

across life stages (Boege and Marquis, 2005; Barton and Koricheva, 2010), so the benefits of a diverse guild of defensive mutualists are greatest when more defensive partner species align with more vulnerable life stages (Djiéto-Lordon et al., 2005; Dejean et al., 2008).

Defensive ant-plant mutualisms – where plants provide food and/or housing to ants that in turn defend them from enemies – are widespread interactions that offer valuable model systems for the ecology and evolution of mutualism (Bronstein, 1998; Bronstein et al., 2006). Extrafloral nectar (EFN) -bearing plants can serve as dietary resources that promote ant abundance and colony size (Byk and Del-Claro, 2011; Ness et al., 2009, 2006; Donald and Miller, 2022). In turn, presence of defensive ant partners is often linked to reductions in herbivory (Trager et al., 2010; Rudgers, 2004) and demographic advantages for the plant partner (Báez et al., 2016). Defensive ant-plant mutualisms are commonly multi-species, where a guild of ant partner species share, and often compete for, a plant mutualist (Bronstein, 1998; Beattie, 1985; Trager et al., 2010; Agrawal and Rutter, 1998). Ant partners can vary in their ability to deter herbivores (Bruna et al., 2014), and visitation by low quality ant partners can prevent visitation by higher quality partners, consequently causing a reduction in fitness through missed opportunity costs (Fraser et al., 2001; Frederickson, 2005). Susceptibility to herbivory can also vary significantly throughout the life stages of the plant (Boege and Marquis, 2005), suggesting that the order and timing of successive partners is important to the fitness impacts of the combined partner guild (Barton and Koricheva, 2010; Boege and Marquis, 2005; Fonseca et al., 2003). Finally, herbivore identity and pressure can vary inter-annually (Wetzel et al., 2023), much like mutualist identity and presence, meaning the threat plants face can vary just as much as the protection they receive due to temporal stochasticity. Previous studies have investigated how ant partner diversity affects plant fitness (Palmer et al., 2010; Afkhami, 2014; Fiala et al., 1994; Gaume et al., 1998; Dattilo et al., 2014; Ludka et al., 2015). However, little is known about the combined effects of partner identity, directional partner turnover, and temporal stochasticity, particularly because the necessary long-term data are rarely available.

This study examined the consequences of partner diversity in a food-for-protection mutualism

96 between the tree cholla cactus (*Cylindropuntia imbricata*), a long-lived EFN-bearing plant, and  
multiple species of ant partners. Previous studies have shown that herbivory by specialized in-  
sect herbivores negatively affects cactus fitness Miller et al. (2009), and that ant defense reduces  
99 herbivore damage Miller (2007). Tree cholla are tended by two common ant species (*Liometopum*  
*apiculatum* and *Crematogaster opuntiae*) and several additional rarer species, all of which collect  
EFN during foraging visits but their colonies are ground-nesting and not housed by the plants.  
102 These ant species locally co-occur but individual plants are typically tended by only one species  
that patrols the plant around-the-clock and maintains control of the plant’s nectar resources, usu-  
ally for an entire growing season (Ohm and Miller, 2014; Donald and Miller, 2022). Switches  
105 between partner species, or between vacancy and ant occupancy, commonly occur from one grow-  
ing season to the next (Miller, 2007). Prior experiments suggested a hierarchy of mutualist quality,  
with *L. apiculatum* providing strong anti-herbivore defense and *C. opuntiae* having net negative  
108 effects because herbivore deterrence is outweighed by deterrence of pollinators (Miller, 2007; Ohm  
and Miller, 2014). However, all of our previous studies in this system have focused on single life  
stages (adult plants) or vital rates (seed production) and did not integrate the demographic effects  
111 of ant defense across the life cycle, which may be essential for understanding net fitness effects of a  
diverse partner guild (e.g., Palmer et al., 2010). To our knowledge no previous study has incorpo-  
rated inter-annual stochasticity into models of ant-plant dynamics, which limits our understanding  
114 of diversity benefits that may arise through the portfolio effect.

Here we used a unique long-term data set that allows us to explore mutualistic associations  
with multiple partner species, longitudinal turnover in partner identity, and how the demographic  
117 effects of alternative partners varies across plant size structure and nearly 20 years of inter-annual  
fluctuations. We used this observational data set of plant demography and ant-plant associations,  
contextualized by previous ant exclusion experiments, to investigate whether and through which  
120 mechanism(s) partner diversity affects the fitness benefits of ant visitation. Specifically, we asked:

1. What are the demographic effects of association with alternative partners and how do these  
effects fluctuate across years?

- 123 2. What are the frequency and direction of partner turnover across the plant life cycle?
3. What is the net effect of partner diversity on plant fitness, and what mechanism(s) explain(s) this effect?

126 To answer these questions we used a hierarchical Bayesian statistical approach to estimate demographic vital rates for hosts in different states of ant occupancy and to quantify state-dependent partner turnover from the long-term data. We then used a stochastic, multi-state integral projection model (IPM) that combines diverse effects on vital rates and pathways of partner turnover to  
129 quantify effects of partner diversity on plant fitness.

## Methods

### *Study System*

132 This study was conducted in the Los Piños mountains, a small mountain chain located on the Sevilleta National Wildlife Refuge, a Long-term Ecological Research site (SEV-LTER) in central New Mexico, USA. This is an area characterized by steep, rocky slopes, and perennial vegetation  
135 including grasses (*Bouteloua eriopoda* and *B. gracilis*), yuccas, cacti, and junipers. Tree cholla cacti are common in high Chihuahuan desert habitats of the southwestern USA (Benson, 1982).  
138 These arborescent plants produce cylindrical segments with large spines. In the growing season (May to August in New Mexico), the plants initiate new vegetative segments and flowerbuds at the ends of existing segments. While most plants produce new segments every season, only those  
141 which are reproductively mature produce flowerbuds. Like other EFN-bearing cacti, tree cholla secrete nectar from specialized glands on young vegetative segments and flowerbuds (Ness et al., 2006; Oliveira et al., 1999). Flowerbuds produce more and higher-quality EFN than vegetative  
144 segments, making reproductive cholla valuable mutualist partners (Miller, 2014). Smaller, non-flowering cholla produce little to no EFN and are commonly vacant (no ant visitation at the time of our census) (Miller, 2014).

147 Tree cholla EFN is harvested by various ground-nesting ant species. At SEV-LTER, cholla  
are visited primarily by two species, *Crematogaster opuntiae* and *Liometopum apiculatum*, as  
well as other rarer species, including *Forelius pruinosus* and unidentified species in the genera  
150 *Aphaenogaster* and *Camponotus*. *L. apiculatum* are the most frequent visitors with 25% – 60% of  
tree cholla tended by these ants in a given year, followed by *C. opuntiae* (5% – 20%) (Donald and  
Miller, 2022). Between 30% – 80% of cacti are vacant in any given year. Workers of different species  
153 rarely co-occur on individual plants, likely due to interspecific competition. For example, staged  
introductions of *C. opuntiae* to *L. apiculatum*-tended plants, and vice versa, provoke aggressive  
responses by residents (A. Campbell, *personal observation*).

156 Several insect herbivores and seed predators specialize on tree cholla (Mann, 1969), and defense  
against these enemies is the main pathway by which ant visitation affects plant demography. The  
Cerambycid beetle *Moneilema appressum* and an unidentified weevil (Coleoptera: Curculionidae)  
159 of the genus *Gerstaekeria* feed on vegetative and reproductive structures as adults and their larvae  
feed internally. Two species of cactus bugs, *Narnia pallidicornis* and *Chelinidea vittiger* (Hemiptera:  
Coreidae), feed on all cholla parts with a preference for flower buds; their damage can induce floral  
162 abortion (Miller et al., 2006). A seed predator, *Cahela ponderosella* (Lepidoptera: Pyralidae),  
oviposits in open flowers and larvae eat seeds in developing fruits. These consumers can have  
significant negative impacts on plant fitness and depress population growth (Miller et al., 2009).  
165 Prior experiments showed that ant-tended tree cholla experience less herbivory and seed predation  
than plants from which ants were excluded (Miller, 2007; Ohm and Miller, 2014).

## *Data Collection*

168 This study is based on long-term demographic data spanning 2004 to 2023 at SEV-LTER. From  
2004 to 2008, we censused 134 plants distributed across three spatial blocks. This initial census  
group was discontinued in 2009, when we established six 30 × 30-meter plots and tagged all tree  
171 cholla within those plots. Two additional 30 × 30-meter plots were added in 2011, and this group  
of eight plots has since been censused annually through 2023 (with the exception of 2020 due to the



pandemic shutdown). For all plants, in May or early June of each year we recorded plant survival since the last survey and, for survivors, we recorded height (cm), maximum crown width (cm), and crown width perpendicular to the maximum (cm). Size measurements were used to calculate plant volume ( $cm^3$ ) based on the volume of an elliptical cone. We measured reproduction by counting flowerbuds, and in most years we distinguished between flowerbuds that were viable and aborted. We recorded the ant species present (or vacancy if no ants present). Occurrences of more than one ant species on one plant were rare (less than 5% of observations), and for the purpose of this analysis we classified the plant as being occupied by the more abundant species. Plots were searched for new recruits each year, and these were added to the census. These data allowed us to link each plant's demographic fate (survival, growth, and reproduction) to its state of ant visitation. In total, the data set includes a total of 9,787 observations of 1141 unique individuals across 15 complete transition years (spanning May/June of year  $t$  to May/June of year  $t + 1$ ) (Campbell and Miller, 2024). In addition to missing the year 2020, there are gaps in the time series where we were switched plots or plants (and thus broke up transition years for growth and survival) or where we did not distinguish between viable and aborted flower buds (Table ??).

We used additional, smaller data sets from previously published studies to estimate seed and seed bank parameters. Ohm et al. 2014 provide data on the number of seeds per fruit for plants tended by *L. apiculatum*, *C. opuntiae*, or no ants (experimental exclusion), accounting for their effects on pollinator visitation. Elder and Miller 2016 provide data on seed entry to the seed bank and seedling germination and survival rates.

### *Multi-state Integral Projection Model*

The demographic data were used to parameterize a multi-state Integral Projection Model (IPM). IPMs describe population dynamics in discrete time with functions that relate vital rates to continuous state variables, typically size (Ellner et al., 2016). While IPMs are a natural choice for populations with continuous size structure, they can also be modified to accommodate a combination of continuous and discrete state variables, as we do here. We constructed a stochastic,

multi-state IPM that stitches together population structure associated with plant size and ant state, allowing us to determine the individual fitness effects of each ant species and the composite effects of multiple partners, with ant transition dynamics and inter-annual variability modeled explicitly.

Given the low frequency of ant occupancy states other than *L. apiculatum* and *C. opuntiae* (<8% of observations) we combined all other ants into an “other” category, such that our multi-state IPM included four ant states: vacant, *L. apiculatum*, *C. opuntiae*, and Other. The “Other” category was made up of *Forelius pruinosus* (3.5% of observations), unidentified species belonging to the genera *Camponotus* (0.9%), *Aphaenogaster* (0.4%), *Myrmecocystus* (0.08%), *Tetramorium* (0.02%), *Brachymyrmex* (0.02%), and additional ants not identified to genus or species (2.8%).

Ant state is included as a predictor variable in IPM sub-models where there are biologically realistic pathways through which ants could impact the outcome of that process. For example, ant partners defend cacti from herbivores, and prior experimental work indicates that ant tending can reduce vegetative tissue loss and floral abortion. Therefore, ant state was included in sub-models for survival, growth, flowerbud viability, and seed number per flowerbud. In contrast, we have no reason to expect that ant tending can directly influence the probability of flowering or flowerbud production independently of its influence on plant size, so these sub-models do not include ant state as a predictor variable.

We modeled the tree cholla life cycle using continuously size-structured plants where number of plants of size  $x$  and ant state  $a$  in year  $t$  ( $n(x, a)_t$ ) predicts the number of plants of size  $x'$  and ant state  $a'$  in year  $t + 1$  ( $n(x', a')_{t+1}$ ) based on a size- and ant-specific vital rates. The models also includes two discrete seed banks ( $B_t^1$  and  $B_t^2$ ) corresponding to 1 and 2-year old seeds. Seed bank dynamics are given by:

$$B_{t+1}^1 = \delta \sum_{a=1}^A \int_L^U \kappa(a) P(x; \boldsymbol{\tau}^P) F(x; \boldsymbol{\tau}^F) V(a; \boldsymbol{\tau}_a^V) n(x, a)_t dx$$

$$B_{t+1}^2 = (1 - \gamma_1) B_t^1$$

Functions  $P(x; \boldsymbol{\tau}^P)$  and  $F(x; \boldsymbol{\tau}^F)$  give the probability of flowering in year  $t$  and the number of  
 225 flowerbuds produced in year  $t$ , respectively, by plants of size  $x$  in year  $t$ . The proportion of  
 flowerbuds that remain viable through fruit set ( $V(a; \boldsymbol{\tau}_a^V)$ ) and the number of seeds per fruit ( $\kappa(a')$ )  
 is dependent on ant state  $a$ . The vectors  $\boldsymbol{\tau}$  give year-specific deviates (mean zero) and appear in  
 228 functions for which we can estimate temporal stochasticity from the long-term data; superscripts  
 indicate the corresponding vital rate and, when present, the  $a$  subscript indicates that deviates are  
 specific to plants in ant state  $a$ . For example, temporal deviates  $\boldsymbol{\tau}_a^V$  describe better- and worse-than-  
 231 average years for flowerbud viability and plants in different ant states can fluctuate independently  
 (good years for *L. apiculatum* -occupied plants may not be good years for *C. opuntiae*-occupied  
 plants, for example). Seed production is integrated over the size distribution, from the lower  $L$   
 234 to upper  $U$  size limits, and summed over all possible initial ant states ( $A = 4$ ) giving total seed  
 production. Seeds are multiplied by the probability of escaping post-dispersal seed predation ( $\delta$ )  
 to give the number of seeds that enter the one-year old seed bank. Plants can recruit out of the  
 237 year-one seed bank with probability  $\gamma_1$  or transition to the two-year seed bank with a probability  
 of  $1 - \gamma_1$ . Seeds in the two-year seed bank are assumed to either germinate with probability  $\gamma_2$  or  
 die.

240 For the above-ground part of the life cycle, the number of plants of size  $x'$  and ant state  $a'$   
 in year  $t + 1$  is given by survival/growth transitions from size  $x$  and ant state  $a$  in year  $t$ , plus  
 germination out of the seed banks:

$$243 \quad n(x', a')_{t+1} = (\gamma_1 B_t^1 + \gamma_2 B_t^2) \eta(x') \omega \rho_0(a') +$$

$$\sum_{a=1}^A \int_L^U S(x, a; \boldsymbol{\tau}_a^S) G(x', x, a; \boldsymbol{\tau}_a^G) \rho(x, a, a'; \boldsymbol{\tau}^e) n(x, a)_t dx$$

246

The first term in Eq. <sup>1</sup> estimates the number of individuals recruiting from a one or two-year  
 seed bank to a plant of size  $x'$  and ant state  $a'$  based on the recruit size distribution  $\eta(x')$  and

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<sup>1</sup> *We should label s. I am not sure why the equation label is not working here and I did not try to figure it out. It is probably something with the linenomath formatting.*

the probability of over-winter seedling survival ( $\omega$ ) from germination (late summer) to the census (May). This term is multiplied by  $\rho_0(a')$ , which gives the probability that a new recruit has ant state  $a'$  ( $\sum \rho_0(a') = 1$ ). The second term represents all possible transitions from size  $x$  and ant  $a$  to size  $x'$  and ant  $a'$ , conditioned on survival. Survival ( $S(x, a; \tau_a^S)$ ) and growth from size  $x$  to  $x'$  ( $G(x', x, a; \tau_a^G)$ ) are both dependent on initial size and ant state. As above, these functions include inter-annual variability through year-specific deviates that can vary by ant state ( $\tau_a$ ). Finally, ant transition function  $\rho(a', a, x; \tau^\rho)$  gives the probability that an individual transitions from ant state  $a$  to  $a'$  in the next census, conditional on initial size  $x$ . This function includes inter-annual variability through year-specific intercepts which are consistent across initial ant states ( $\tau^\rho$ ).

### Statistical modeling and parameter estimation

We parameterized the IPM using a series of generalized linear mixed models in a hierarchical Bayesian framework. Vital rate models included spatial and temporal random effects associated with plot and year variation, respectively (only year variation is used in the IPM), and included plant size (the natural logarithm of volume,  $\log(cm^3)$ ;  $x, x'$ ), ant partner state ( $a, a'$ ), or both as fixed-effect predictor variables. As in the IPM, our statistical modeling assumed that demographic effects of ant occupancy are limited to survival, growth, and flowerbud viability.

*Growth.* We fit the growth sub-model ( $G(x', x, a; \tau_a^G)$ ) to data on size in year  $t + 1$  ( $y^G$ ) using the skewed normal distribution to account for left-skewed size transitions (at some initial sizes, transitions below the expected future size were more common than transitions above it). The skew-normal has three parameters corresponding to location ( $\hat{G}$ ), shape ( $\sigma$ ), and scale ( $\alpha$ ):

$$y_i^G \sim SkewedNormal(\hat{G}_i, \sigma_i, \alpha_i)$$

$$\hat{G}_i = \beta_{a[i]}^0 + \beta_{a[i]}^1 x_i + \beta_{a[i]}^2 x_i^2 + u_{year[i], a[i]} + w_{plot[i]}$$

$$\log(\sigma_i) = \beta^3 + \beta^4 x_i$$

$$\alpha_i = \beta^5 + \beta^6 x_i$$

Here, the location parameter for the  $i$ th observation  $\hat{G}_i$  is defined as a second-order polynomial with ant-size interactions because preliminary analysis found this was an improvement over a linear relationship (note that the location parameter of the skew-normal is not the mean, but the mean can be derived as  $\hat{G} + \frac{\sigma\alpha}{\sqrt{1+\alpha^2}}\sqrt{\frac{2}{\pi}}$ ). The year- and ant-specific random effect  $u$  (which parameterizes the  $\boldsymbol{\tau}_a^G$  vectors) and plot-specific random effect  $w$  are normally distributed with variances  $\sigma_{year}^2$  and  $\sigma_{plot}^2$ , respectively. Parameters  $\sigma_i$  and  $\alpha_i$  control residual variance and skewness, respectively, and were defined as linear functions of initial size  $x_i$  ( $\sigma_i$  is strictly positive and was modeled with a log link function). We assume growth variance and skewness were not dependent on ant occupancy state.

*Survival.* The survival sub-model ( $S(a, x; \boldsymbol{\tau}_a^S)$ ) estimates the probability of survival from year  $t$  to year  $t + 1$ , with fixed effects of size  $x$  and ant partner  $a$  in year  $t$ . We fit this model to the survival data (alive or dead) using a Bernoulli distribution with a similar linear predictor for the probability of survival as in the growth model but with a logit link function and without the second-order influence of size.

*Reproduction.* The flowering sub-model ( $P(x; \boldsymbol{\tau}^P)$ ) estimates the probability of reproducing in year  $t$ , with fixed effects size  $x$  in year  $t$  and random effects of plot and year. We fit this model to the reproductive status data (vegetative or flowering) using a Bernoulli distribution and a logit link function, similar to the survival model above but with no ant effects. The flowerbud function  $F(x; \boldsymbol{\tau}^F)$  estimates the total flowers produced by a reproducing plant in year  $t$ , with fixed effects of size  $x$  in year  $t$ . We fit this model to flowerbud count data (sum of viable and aborted buds) using a zero-truncated negative binomial distribution with a log link and normally distributed year and plot random effects.

The flowerbud viability sub-model ( $V(a; \boldsymbol{\tau}_a^V)$ ) estimates the probability that flowerbuds produced in year  $t$  remain viable (not aborted), with fixed effects of ant partner  $a$  in year  $t$ . We fit this model to floral viability data using a binomial distribution where trials and successes are given by the total number of flowerbuds and the number that are viable, respectively. This model used

a logit link function and included random effects for plot and year following the same structure as the growth and survival models, with ant-specific year random effects.

Estimates for the number of seeds per fruit were obtained from a field experiment which excluded ants (Miller et al., 2006). This data only included *L. apiculatum* and *C. opuntiae* ants, so we had to make an assumption about the seeds per fruit on plants tended by Other ants. We chose to use the vacancy estimates for plants tended by other ants, a decision which does not have a significant impact on the final results. Additional reproductive parameters for the number of seeds per fruit, probability of entry to the seed bank, germination rates, and recruit size were estimated following methods described in Appendix A .

*Ant Transitions.* The ant transition model ( $\rho(x, a, a'; \boldsymbol{\tau}^e)$ ) estimates the probability of a cactus being occupied by ant partner  $a'$  in year  $t + 1$  given that it was occupied by initial ant partner  $a$  in year  $t$ , with fixed effect of initial size  $x$ . We fit this model to ant partner data using a multinomial distribution with a logit link function.

*Parameter estimation.* We fit models using Stan run through version 4.0.2 of R R Core Team (2023); Stan Development Team (2023). We used vague priors for all parameters. For each model, we obtained three chains of 10,000 iterations, discarding the first 1,500 iterations. We visually assessed parameter convergence between and within chains (Appendix : Figures 8 – 14 b) and assessed overall model fit with posterior predictive checks to examine how well the fitted model can generate simulated data similar to the real data (Appendix : Figures 8 – 14 a).

## IPM Analysis

Analyzing the IPM required that we discretize the continuous IPM kernel into an approximating matrix. Size variable  $x$  was discretized into  $b$  bins, resulting in a  $b \times b$  matrix. In our model there is additional complexity in the form of transitions between  $A$  ant states and two additional discrete states (year one and year two seed banks), leading to a matrix size of  $A(b + 2) \times A(b + 2)$ . We

used  $b = 500$  bins, which we found to be sufficient for numerically stable outputs, and extended the  
 327 integration limits beyond the minimum ( $L$ ) and maximum ( $U$ ) observed sizes to avoid unintentional  
 eviction using the “floor-and-ceiling” method Williams et al. (2012).

For stochastic analyses, we estimated the approximating matrix corresponding to each  $t$  to  $t + 1$   
 330 transition year. To estimate population mean fitness in a stochastic environment ( $\lambda_S$ ) we simulated  
 population dynamics for 500 years by randomly sampling among the 15 annual transition matrices,  
 discarding the first 100 years of the simulation to minimize the influence of initial conditions.  
 333 Sampling observed transition matrices (rather than independently sampling regression coefficients)  
 produces demographic time series that realistically capture inter-annual variation by preserving  
 correlations between vital rates Metcalf et al. (2015). We tallied the total population size at each  
 336 time step as  $N_t = B_t^1 + B_t^2 + \sum_{a=1}^A \int n(x, a)_t dx$  and calculated the stochastic growth rate as  
 $\log(\lambda_S) = E[\log(\frac{N_{t+1}}{N_t})]$  (Mark and Ellner, 2009). We propagated uncertainty from the vital rate  
 models using 100 draws from the joint posterior distribution of model parameters, resulting in a  
 339 posterior distribution of  $\lambda_S$  and other derived quantities.

### *Partner diversity simulation experiments*

Using the fully parameterized multi-state IPM, we conducted simulation experiments to quantify  
 342 how diversity and identity of ant partners influenced plant fitness. From the full version of the model  
 (described above) corresponding to the observed assemblage of partners and observed frequencies  
 of partner transition, we created treatments corresponding to all eight “counter-factual” scenarios  
 345 of diversity and composition: no ant partners (complete vacancy); one ant partner (*C. opuntiae*  
 only, *L. apiculatum* only, Other only); two partners (all pairwise combinations of *C. opuntiae*, *L.*  
*apiculatum*, and Other); and three partners (observed scenario of all ant states). These simulation  
 348 experiments were made possible by extrapolating ant-specific demographic performance across the  
 size distribution, even for combinations of size and ant occupancy that were rarely observed. For  
 example, the no-partner scenario modeled a hypothetically ant-free cactus population, even though  
 351 no such population exists to our knowledge, by applying the statistical knowledge gleaned from

vacant plants across the size distribution. We refer to stochastic fitness associated with partner number or identity using superscripts, e.g.  $\lambda_S^0$  for vacant plants (zero partners),  $\lambda_S^{1+}$  for any state of ant tending,  $\lambda_S^C$  for tending by only *C. opuntiae*,  $\lambda_S^{CO}$  for tending by *C. opuntiae* and Other ants, etc.

In all scenarios that included any ant partners, we preserved the observed pattern of size-dependent vacancy/occupancy (estimated through the ant transition sub-model) and manipulated partner identity conditional on occupancy. This means, for example, that the *C. opuntiae*-only scenario included two possible states, vacancy and occupied by *C. opuntiae*. While our statistical models allow us to extrapolate the demographic performance of ant-tended plants to small sizes that are typically vacant, the natural history of this system tells us that this is not biologically sensible. Small, non-reproductive plants are typically vacant because they do not produce extrafloral nectar, and once plants begin producing nectar they are nearly always ant-tended (Miller, 2014). Our simulation experiments preserved this basic biology, avoiding tiny ant-occupied plants that do not and could not occur in nature.

The partner diversity treatment scenarios required additional assumptions about the mechanisms that give rise to observed occupancy patterns. Based on evidence that EFN-bearing cacti are nearly always ant-occupied (Miller, 2014), we assume that ant partners competitively exclude one another from EFN-bearing cacti and that competition is zero-sum. This means that, in scenarios that remove species from the partner community, remaining species gain access to plants that the removed species would have tended. In Appendix , we present results under an alternative assumption, that ant visitation is limited by factors other than availability of cactus EFN (e.g., nesting sites or off-plant dietary resources), such that when a species is removed from the partner community, the plants it would have occupied remain vacant.

### *Temporal stochasticity experiments*

Under the portfolio effect hypothesis, partner diversity may confer a fitness advantage when the benefits of alternative partners are not perfectly synchronized across temporal environmental vari-



ation, yielding an advantage of a diverse “portfolio” of partners when the environment fluctuates. Our statistical estimation of ant-specific year random effects in the vital rates allows for this possibility. We constructed two versions of the stochastic, multi-state IPMs that allowed us to test this hypothesis. The baseline, ‘non-synchronous’ model described above included ant effects that could vary uniquely across time. In this scenario, good years and bad years can differ between ant states, according to the parameter estimates for the random effects ( $\tau_a$ ). We quantified from the fitted random effects how tightly inter-annual variation was correlated between ant states for each vital rate. The ‘synchronous’ (S) version included ant effects that were forced to be the same across species. To synchronize ant states, we averaged the ant-specific year random effects, thus ensuring that plants in all ant states fluctuated synchronously in response to temporal environmental variation. This second, “synchronized” version of the model effectively turns off any portfolio effect, holding all else equal. Both scenarios of temporal stochasticity, non-synchronized and synchronized, were run for all eight ant partner scenarios described above. We indicate stochastic fitness under the synchronized scenario with  $\lambda_S^{C(s)}$ , for example.

### *Statistical inference on fitness consequences of partner identity and diversity*

The range of models we created generated many outputs; we focus our inference on the following specific contrasts. First, to determine whether ant occupancy and partner diversity are beneficial, we calculated a posterior distribution of  $\lambda_S$  for each of four partner richness levels (zero, one, two, three), averaging over composition scenarios within each level. If cactus fitness increases with partner richness, this would be interpreted as evidence for benefits of partner diversity. Second, to determine whether each partner, in isolation, confers a fitness advantage and to rank alternative partners, we contrasted the fitness of each single partner scenario (*C. opuntiae* only, *L. apiculatum* only, Other only) against vacancy (zero partners). Third, to determine whether any benefits of diversity are due to the sampling effect or complementarity, we contrasted the fitness of multi-partner scenarios against the single best partner scenario. If the best multi-partner scenario exceeds the fitness associated with the best single partner, this would be interpreted as evidence of comple-

mentarity, a true benefit of diversity *per se*. Alternatively, the sampling effect hypothesis predicts  
 405 that no multi-partner scenario yields higher plant fitness than the best single partner. It is also  
 possible that multi-partner scenarios yield lower fitness than the single best partner, which would  
 be consistent with an opportunity cost of diversity. Fourth, to quantify any contribution of the  
 408 portfolio effect, we contrasted  $\lambda_S$  of the full (four-state) scenario to vacancy (as a measure of how  
 much benefit is received from all partners being present) for synchronized and non-synchronized  
 responses to temporal stochasticity. If the portfolio effect confers a benefit of diversity, the fit-  
 411 ness advantage of having all vs. no partners should be greater when temporal fluctuations are not  
 synchronized across ant states.

We base our statistical inferences on the posterior probability distributions of the contrasts  
 414 described above. For example, the contrast of *C. opuntiae* ( $\lambda_S^C$ ) with vacancy ( $\lambda_S^V$ ) yields a posterior  
 distribution of the difference  $\Delta\lambda_S^{C-V}$ . We can quantify from this distribution our certainty in the  
 mutualistic effect of *C. opuntiae*, given the data, as  $Pr(\Delta\lambda_S^{C-V} > 0)$ . We apply similar logic to  
 417 other contrasts described above.

## Results

*What are the demographic effects of association with alternative partners  
 420 and how do these effects fluctuate across years?*

Over the 20-year data set, we found that ant partners influenced demographic performance of  
 cactus hosts, and different ant partners had contrasting demographic effects across host vital rates.  
 423 Ant-tended plants had a growth advantage over vacant plants, especially at smaller sizes (Figure  
 1). Over the size range for which plants were likely ant-tended (larger than the minimum observed  
 size of  $0.8 \log(\text{cm}^3)$ ; solid lines in Figure 1) there were modest differences between partner species,  
 426 with the greatest growth advantage associated with *C. opuntiae* followed by *L. apiculatum* and  
 then Other ants. At large sizes the growth trajectories of ant-tended and vacant plants were nearly

indistinguishable. For all ant states, growth was left-skewed, with transitions to sizes below the  
429 mean were more common than sizes above the mean.

Similarly, for plants which were large enough to have ant visitors, visitation enhanced cactus  
survival (Figure 2). Mean survival rates ranged from 7.7% to 99.9%, with the smallest plants the  
432 most vulnerable to mortality. *C. opuntiae*-occupied plants had a survival advantage over other ant-  
tended plants, particularly at smaller sizes, consistent with the positive effects on growth. At larger  
sizes, plants in any state of ant occupancy had a survival advantage over vacant plants. Plants that  
435 were smaller than  $-2 \log(cm^3)$  were predicted to experience a survival advantage from vacancy, but  
apparent benefits of occupancy were based on extrapolated survival estimates of ant-tended plants.  
Plants in this size range were never observed to be ant-tended and, because the IPM preserves the  
438 size-dependence of vacancy, benefits of vacancy at small sizes are never realized in the IPM.

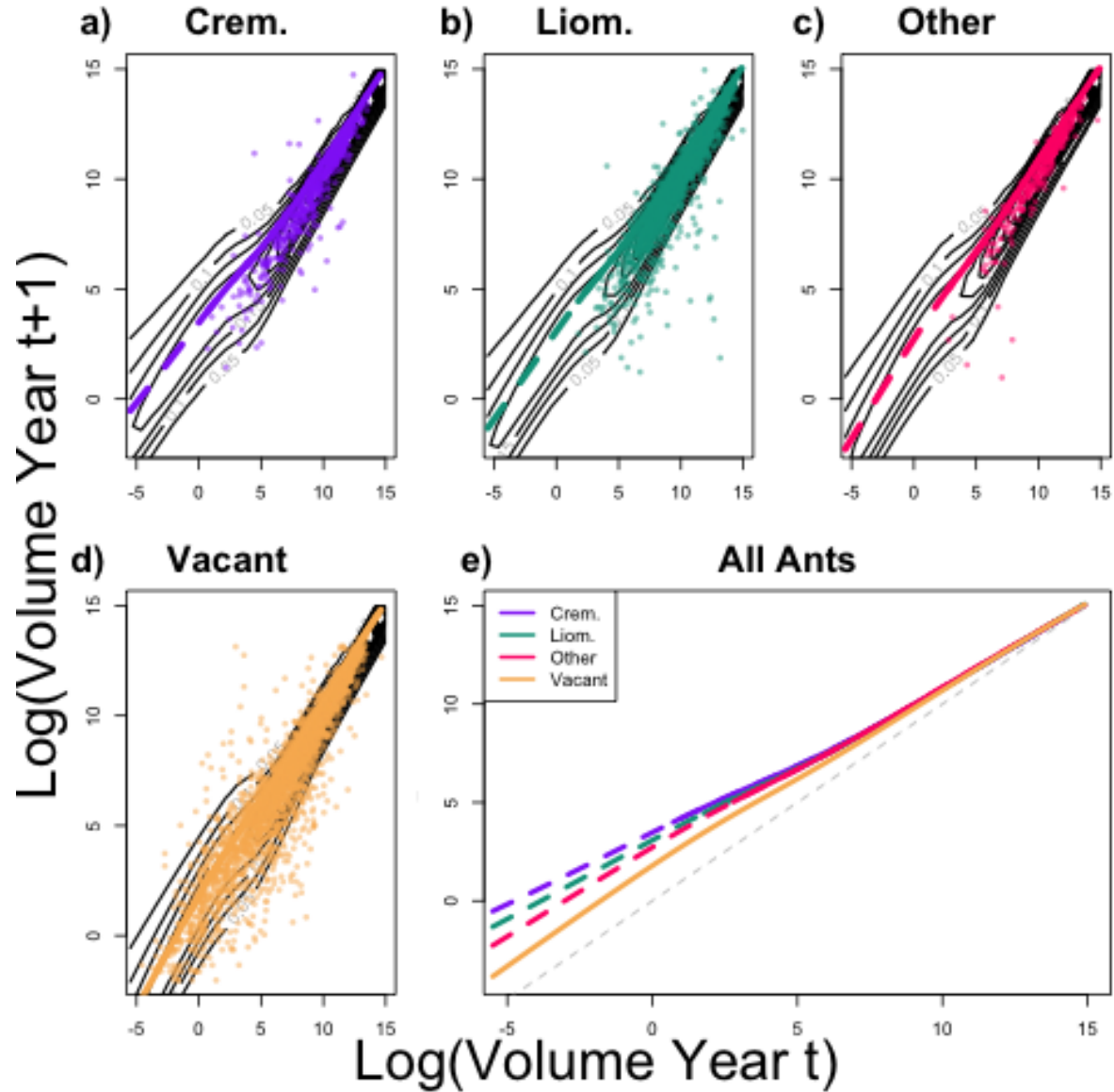


Figure 1: This figure shows the next predicted size of cholla based on previous size with each individual ant partner. The colored lines (seen in all panels) are the next mean predicted size of cholla. Dashed lines indicate extrapolation while solid lines indicate estimations within the range of observed data. The points (seen in panels a-d) are the observed data which informs these estimates. The black contour lines (seen in a-d) appear at 5% increments showing where 5%, 10%, etc. of the data is expected to fall. The gray dashed line (in panel e only) shows the line where the next predicted size is the same as the previous (aka there is no growth on this line and below this line is shrinkage).

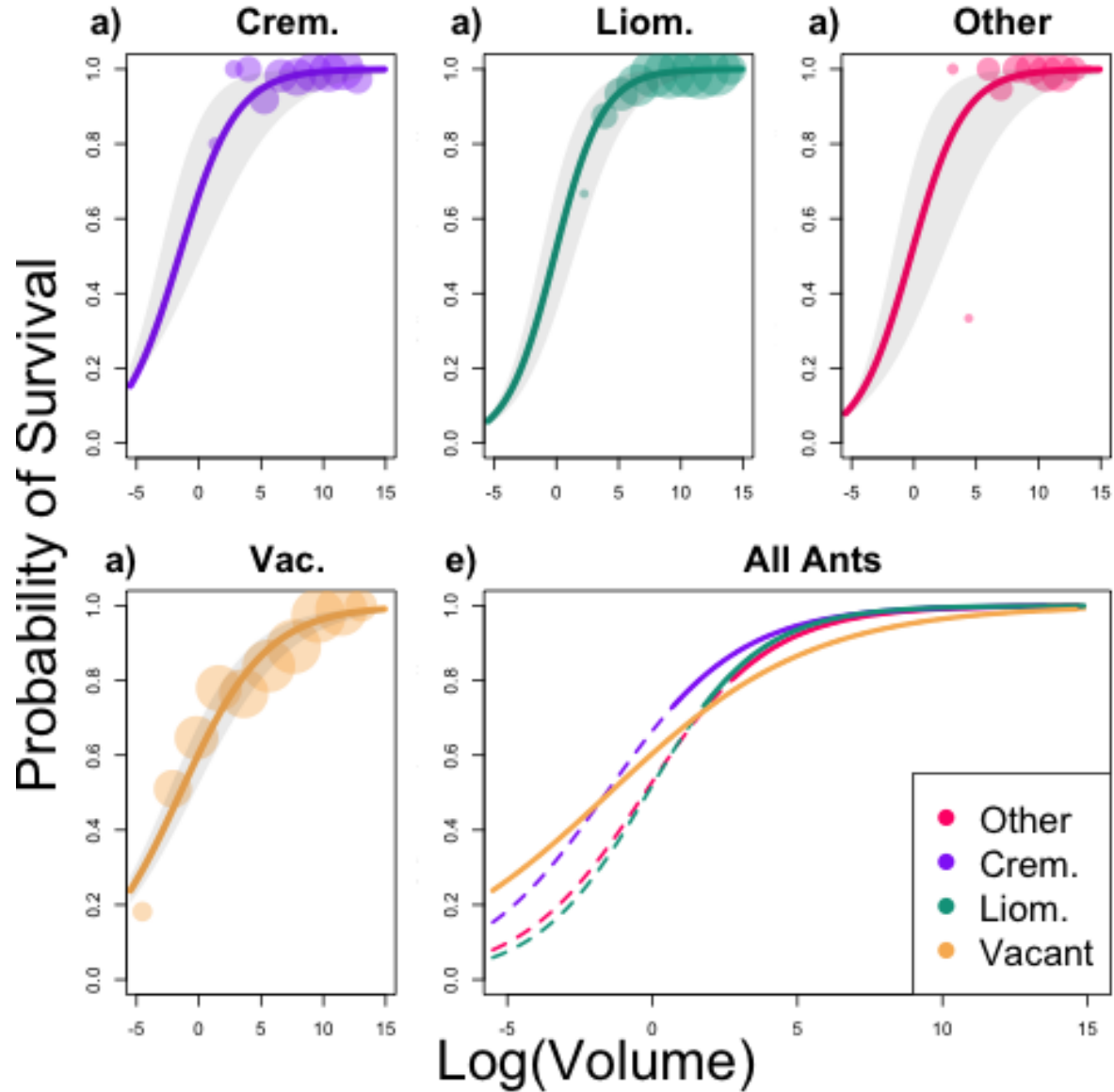


Figure 2: This figure shows the estimated survival rates based on the size of the cactus with each individual ant partner. The solid colored lines (shown on all panels) indicate the mean estimated survival rates. The dashed lines (shown in panel e) indicate extrapolations beyond existing data (where we estimated survival for plants tended by ants where we had never seen a tended cactus of that size). The grey area around the solid lines (shown in panels a-d) show the 90% confidence interval for the estimates. The colored dots are the real data binned by size to show how our estimates align with real survival observations. A larger circle means we had more data on survival of plants of this size with this partner.

We found evidence that ant tending was associated with increased floral viability rates and that ant identity can influence the strength of viability benefits. We observed mean viability rates  
441 of flowers between 40% and 92% (Figure 3). *L. apiculatum*-tended plants had the highest mean  
viability rate (86.1% [95% credible interval: 77.6–92.4%]), and there were similar viability rates for  
vacant (60.0% [44.3–75.0%]), Other-tended plants (at 60.6% [43.7–75.5%]), and *C. opuntiae* tended  
444 plants (57.1% [40.6–72%]). Furthermore, *C. opuntiae*-tended plants had fewer seeds per fruit  
(115.0[79.5–165.5] seeds) than vacant (147.2[114.1–189.9] seeds) or *L. apiculatum*-tended plants  
(142.4[100.7–200.2] seeds) (Figure 1).

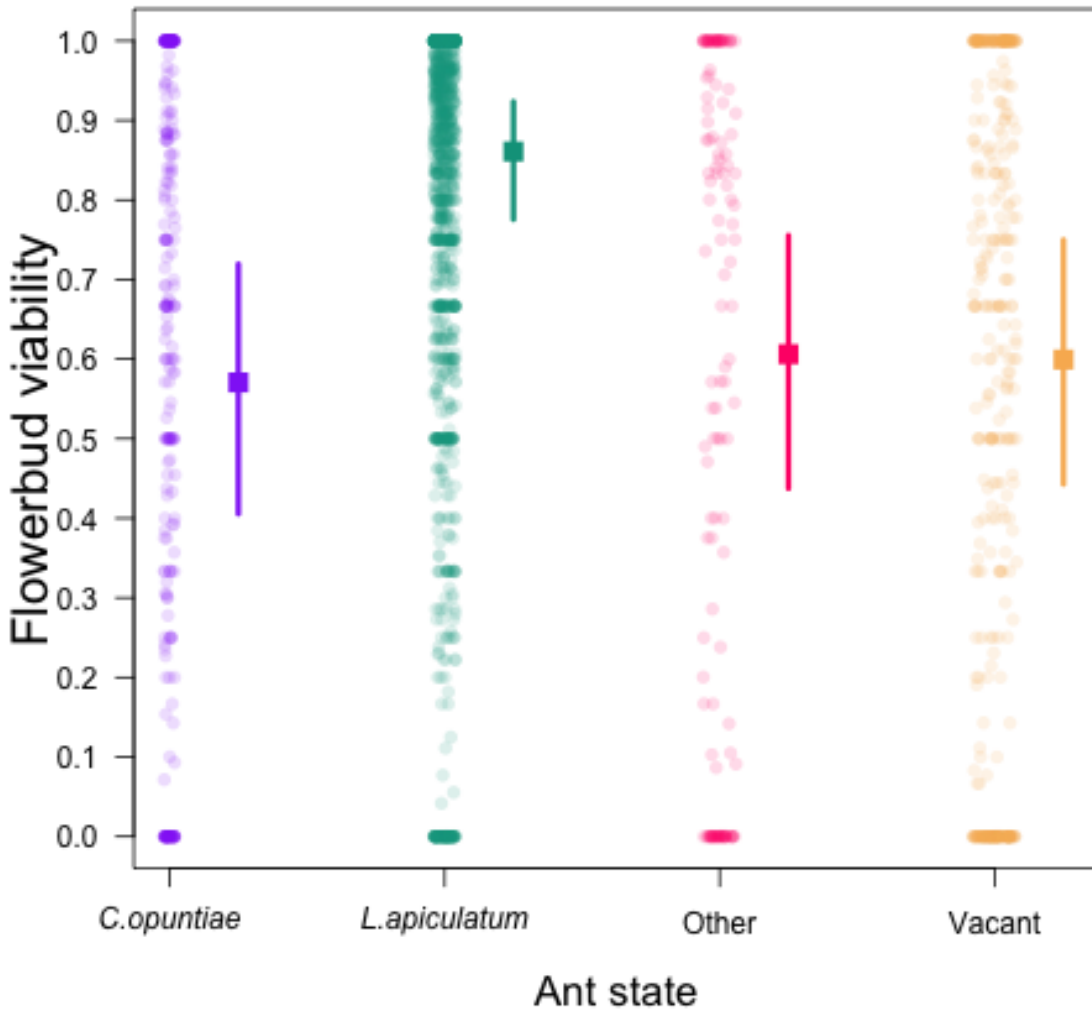


Figure 3: This figure shows the estimated distributions of floral viability rates compared to observed distributions of floral viability rates of cholla based on ant partner identity. The clouds of points show the observed data, with the color representing the ant state. The solid lines indicate 95% of the estimated viability posterior distribution.

447 The results of all other vital rates as well as the posterior predictive checks are all included in the Appendix : Figures 1 – 4.

We analyzed the correlation coefficients of all models that included ant state as a predictor  
450 and found that, across vital rates, the annually varying effects on growth were the most correlated between ant states (mean pairwise correlation: 0.63) and effects on survival were the least correlated (0.36) (Figure 4). For growth, annually varying random effects of Other ants were the  
453 least synchronous with other ants states (correlation coefficients ranging from 0.28 to 0.39) while the effects of vacancy were the most synchronous with other states (correlation coefficients ranging from 0.39 to 0.81). For survival, effects of *C. opuntiae* were the least synchronous with other ant  
456 states (correlation coefficients ranging from 0 to 0.15) while the effects of vacancy were the most synchronous (correlation coefficients ranging from 0.13 to 2.6). For floral viability, effects of *L. apiculatum* were the least synchronous with other ant states (correlation coefficients ranging from  
459 0.06 to 0.44) while effects of vacancy were the most synchronous (correlation coefficients ranging from 0.44 to 0.62).

The variation in synchrony across ant states and vital rates indicates there is potential for  
462 portfolio effect, as many of the ant effects revealed low synchronicity, particularly in survival.



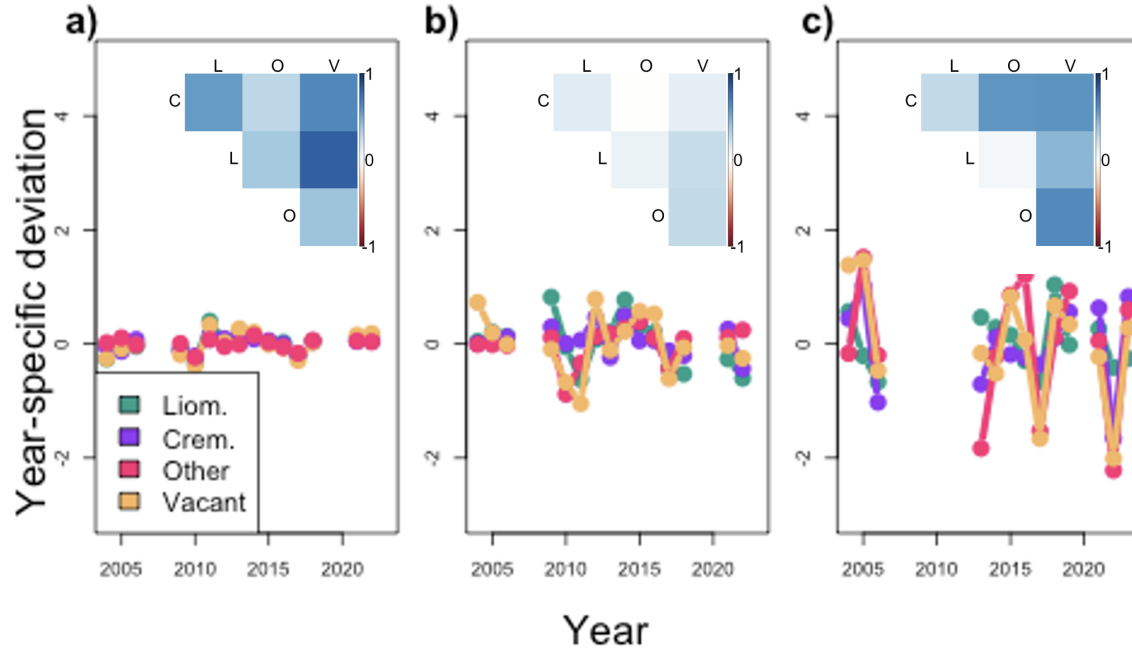


Figure 4: The colored lines in each panel show the mean effect of each ant partner on a vital rate across years (blank periods indicate a lack of transition data). The blue squares in the upper corner of each panel show the correlation matrix of how similar ant-specific annual responses are for each vital rate. Panel a) shows the effects on the estimated next size, b) the effects on estimated survival, and c) the effects on floral viability of cacti across every year of our study. These values are estimated from the fitted random effects of ant and year in our models. Each point represents the mean of the random effect of the identified model, ant, and year (e.g. the lowest dot in panel b) represents the mean effect of vacancy on survival rates in year 2011).

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<sup>2</sup>*Honestly not sure if I should include an image for this one or just report some values? We should discuss.*

*What are the frequency and direction of partner turnover across the plant  
life cycle?*

465

We found that 55% of individual plants surveyed in the long-term data experienced at least one ant state transition on average, with very distinct size-dependence and directional patterns (Figure 5). Vacancy was the most likely ant state of small plants ( $\leq 10 \log(\text{cm}^3)$ ). Even when small plants were ant-tended at the start of the transition year, they were most likely to transition back to vacancy (Figure 5b-d). The probability of becoming ant-tended increases with size, though it was not equally likely to be tended by all partners. For large plants that were initially vacant or tended by *L. apiculatum* or Other ants, *L. apiculatum* was the most likely next partner, suggesting that this species is able to colonize plants that were previously vacant or occupied by Other ants, and effectively retain plants that it previously occupied. *C. opuntiae* were also able to retain plants they previously occupied, but not as well as *L. apiculatum*: for plants that begin the transition year with *C. opuntiae*, the probability that those plants remain occupied by *C. opuntiae* at the end of the transition year is only slightly greater than the probability of take-over by *L. apiculatum*, while take-over in the other direction is extremely rare. It is also notable that transitions away from the initial state of *L. apiculatum* were almost always transitions to vacancy (Figure 5d), while transitions away from the initial states of *C. opuntiae* and Other were often transitions to other ants. This suggests a competitive hierarchy whereby *L. apiculatum* may abandon low-value plants with little nectar production but is almost never displaced from high-value plants.

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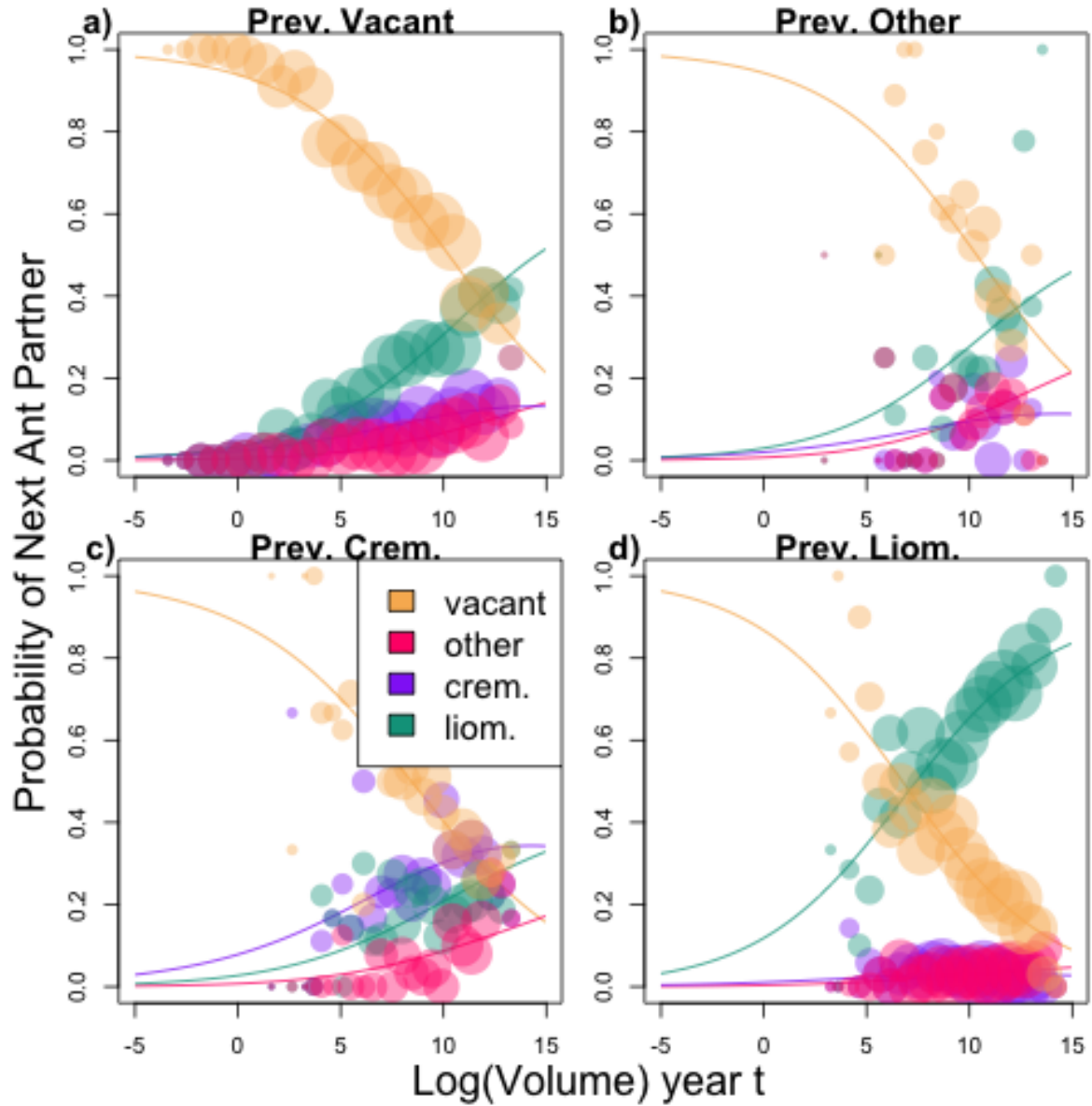


Figure 5: This figure shows the probability of being tended by each ant partner or vacant based on the size of the plant. Each panel shows these probabilities for a different previous ant state. The solid lines represent the mean probability of being tended by a specific partner. The colored points are the real data binned by size to show how our estimates align with real visitation observations. A larger circle means we had more data on visitation of plants of this size with this previous partner.

*What is the net effect of partner diversity on plant fitness, and what  
mechanism(s) explain(s) this effect?*

By integrating vital rate results, temporal fluctuations, and ant transition dynamics into the stochastic, multi-state IPM we can evaluate the fitness implications of different scenarios of partner diversity and identity. First, there was strong evidence that ant visitation had mutualistic fitness effects on plant partners. The lowest mean stochastic fitness was  $\lambda_S^0$ , the fitness of the cholla with no partners (Figure 6a). Across all 1+ partner scenarios, we found that the  $\lambda_S^{1+}$  posterior distributions were greater than  $\lambda_S^0$  with nearly 100% certainty. This indicates that ant visitation elevates fitness no matter the identity or number of partners. Furthermore, we found no benefits of partner diversity, with the fitness of cholla 1, 2, and 3 partners roughly equivalent (Figure 6 a). The 1- and 2-partner scenarios were statistically consistent but there was a modest reduction in fitness in the 3-partner scenario compared to 2 partners ( $Pr(\Delta\lambda_S^{2-3} = 0.95)$ ).

Among the one-partner scenarios, there was no strong evidence for any single best partner species. While *L. apiculatum*-tended plants had the highest mean fitness, it was not significantly higher than *C. opuntiae*- ( $Pr(\Delta\lambda_S^{L-C} = 0.7)$ ) or Other- ( $Pr(\Delta\lambda_S^{L-O} = 0.69)$ ) tended plants. Furthermore, the fitness of *L. apiculatum*-tended plants was consistent with 1- and 2-partner scenarios ( $Pr(\Delta\lambda_S)$  ranged from 0.37 to 0.83). However, as above, there was evidence for an opportunity cost of diversity wherein fitness of the three-partner scenario was lower than any of the 2-partner scenarios ( $Pr(\Delta\lambda_S) > 0.95$ ).

We found that the difference in posterior distributions of any ant scenario ( $\lambda_{NS,A} - \lambda_{NS,B}$ , where A and B are arbitrary ant diversity scenarios that are not vacancy) were between 31% and 83% different. The posterior distributions of  $\lambda$  for each diversity scenario are visualized in Figure 6 b. This high level of overlap between posterior distributions indicates that while each of these partners are beneficial, there does not appear to be a significant benefit of partner diversity within this system.

The lack of diversity benefits are not driven by the high overall frequency of *L. apiculatum*.

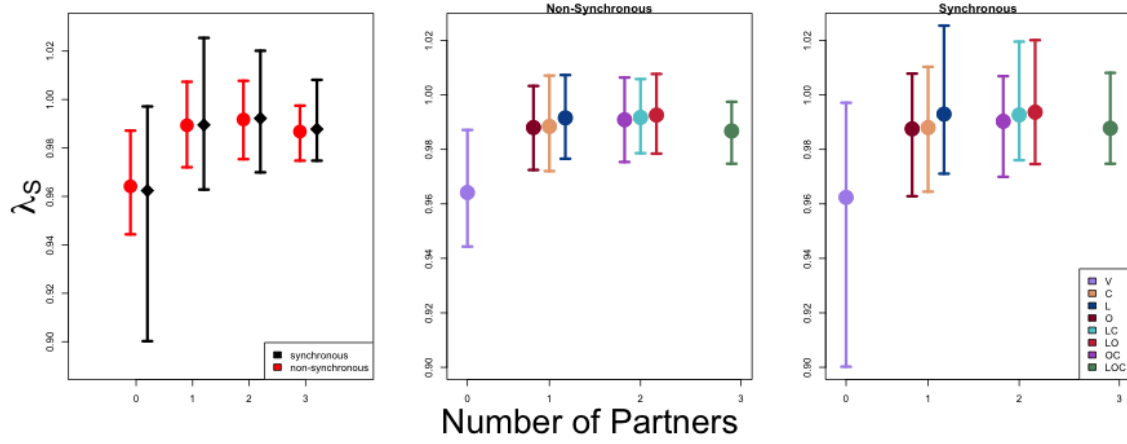


Figure 6: Panel a) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated  $\lambda$  values. Panels b-c) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated  $\lambda$  values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

Using the simulations where all ants had equal frequencies across sizes (further explained and  
 510 analyzed in Appendix C), we found the same fitness patterns as in the competitive exclusion model  
 described above. Equal probability for transitioning into any ant state meant that the numbers  
 of *C. opuntiae* and Other ants were boosted significantly. Despite this, the fitness of scenarios  
 513 including these less frequent ants were not increased meaningfully.

We found no evidence of portfolio effect. The effect of all ant partners can be measured as  
 $\lambda_{All} - \lambda_{Vacant}$  (Figure 7). We are 100% confident that when all ants are present the cholla  
 516 experience higher fitness than when no ants are present according to both the synchronized (S)  
 and non-synchronized (NS) model scenarios. When subtracting these two resulting vectors from  
 each other  $((\lambda_{NS,All} - \lambda_{NS,Vacant}) - (\lambda_{S,All} - \lambda_{S,Vacant}))$ , we found that there is no real difference  
 519 between the two scenarios, meaning we have no evidence of portfolio effect.

## Discussion

Mutualisms commonly involve multiple partners but the ecological consequences of partner diversity  
 522 remain poorly understood. Here we show that while alternative partners may be ecologically  
 different, their fitness effects on a shared mutualist can be effectively equivalent and interchangeable.  
 The results of our hierarchical models revealed that different ant partners had different effects on  
 525 vital rates, with *C. opuntiae* tended plants **experiencing slight advantages in growth**<sup>3</sup> and survival  
 when small, and *L. apiculatum*-tended plants experiencing floral viability advantages. The results  
 of our synchronous and non-synchronous stochastic IPM revealed that all diversity scenarios which  
 528 included any partners resulted in the highest possible fitness for tree cholla, suggesting that while  
 ant visitation is beneficial, partner identity and diversity are inconsequential in this system. These  
 results highlight that partners can differ meaningfully while still resulting in the same fitness benefits  
 531 for a focal mutualist.

Previous studies have reported complementarity (Palmer et al., 2010; Rez, 2007; Afk, 2021)

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<sup>3</sup> *I don't think this is true if you look at my v1 figure.*

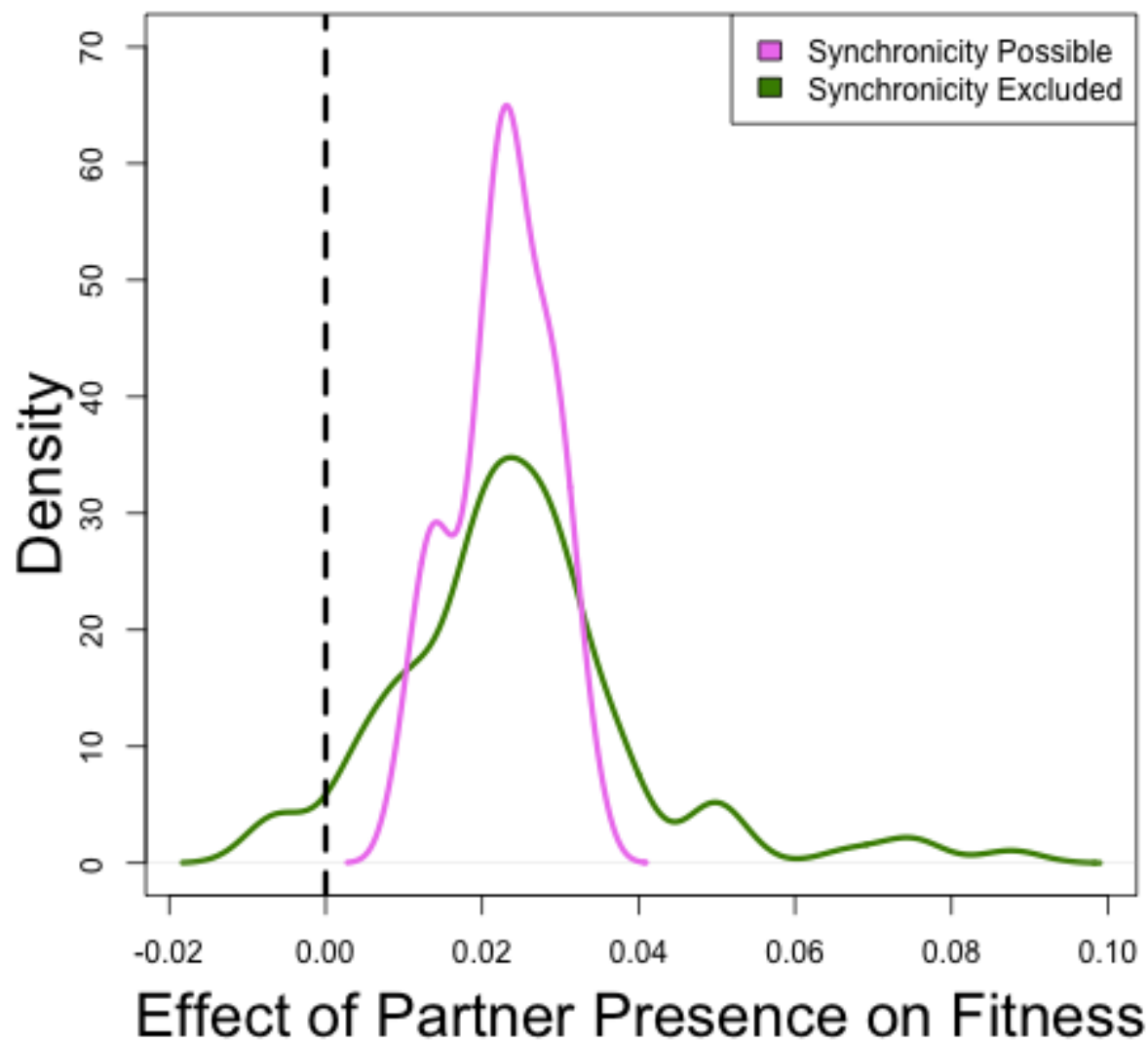


Figure 7: This figure shows the distribution of  $\lambda_{NS,All} - \lambda_{NS,Vacant}$  in pink and  $\lambda_{S,All} - \lambda_{S,Vacant}$  in green. The vertical dashed line shows where the effect of partners on the fitness of the cholla is 0 (to the left the partners have a negative effect, to the right the partners are beneficial).

while we found that there are no benefits of diversity in our system. Vital rate results indicated  
534 that different ant partners affected different vital rates uniquely (Figures 1,2,3). On their own,  
these results suggested a likely benefit of diversity through complementarity, such that each ant  
partner offered a unique form of herbivore protection across ontogeny and or across different life  
537 history processes. However, our multi-state IPM results, which synthesize all life history processes  
and accounts for partner identity and turnover, revealed an overall positive effect of ant visitation  
but no apparent benefit of diversity, raising questions about how to resolve the tension between  
540 these two sets of results.

We speculate that our results are driven by the vital rate sensitivity structure of this population.  
Like other long-lived, iteroparous species (Franco and Silvertown, 2004), tree cholla fitness is most  
543 sensitive to the growth and survival of established individuals (Miller et al., 2009; Elderd and  
Miller, 2016), which are virtually guaranteed many years of reproductive opportunities once they  
reach a size that is protected from mortality. Differences between alternative partners were most  
546 pronounced either in reproductive rates (Figure 3), which contribute relatively weakly to fitness  
(because a long reproductive lifespan overrides individual reproductive bouts), or in survival at small  
sizes, where mortality risk was high (Figure 2) and the probability of ant tending was low (Figure  
549 5). At larger sizes with low mortality risk, ant-tended plants had a modest survival advantage over  
vacant plants regardless of ant identity, and that result dominates integrated measures of fitness  
because of the high sensitivity of established plant survival. The very real demographic differences  
552 between partner species in other vital rates and at other sizes do not register nearly as strongly in  
the currency of fitness.

Our work explicitly incorporated temporal environmental stochasticity, which raises the oppor-  
555 tunity for portfolio effect as a mechanism of diversity benefits. Yet, we find no evidence of portfolio  
effect within our system. When partners exhibit different reactions to varying environments, in-  
teracting with multiple partners can lead to more consistent benefits across time (Batstone, 2018).  
558 We did not directly measure the reactions of ant partners to temporal environmental stochasticity,  
rather we measure how the effects of each partner on cholla demography fluctuate through time,



and whether those fluctuations are correlated. The vital rate results led us to believe that portfolio  
561 effect may be at play due to the differences in ant responses to annual variation. However, there is  
no stabilization of fitness across temporal heterogeneity and therefore no benefits of diversity due  
to the portfolio effect. Other portfolio effect studies (Lázaro et al., 2022; Tornos-Estupiña et al.,  
564 2023) found that asynchrony was linked to weak evidence of portfolio effect. One study Dallas and  
Kramer (2022) found that while portfolio effect was easy to show in theoretical models, it is often  
very weak or nonexistent in empirical data across many systems. This indicates that it may be  
567 very difficult to detect, disguised by different mechanisms, or uncommon in nature.

Partner turnover is likely a significant driver behind the fitness we see within Cholla populations,  
however the processes which drive the actual turnover frequency and directions remain a mystery in  
570 this system. In the literature, it is clear that the frequency of partner turnover can have big effects  
on the fitness of the focal mutualist Fiala et al. (1994); Horvitz and Schemske (1986); Oliveira et al.  
(1999); Sachs et al. (2004). The direction of partner turnover is also important when the identity of  
573 partners impacts the quality of benefits recieved Fonseca et al. (2003); Alonso (1998); Dejean et al.  
(2008); Noe and Hammerstein (1994). In our system we found that there are distinct patterns to the  
directions and frequency of partner turnover, but we still don't understand the mechanisms driving  
576 these turnovers. In the model presented in this paper we assumed that competitive exclusion was  
the driving mechanism behind the patterns we see. This assumption would indicate that ants are  
competing for each individual cacti, likely based on EFN quality and composition (Heil et al., 2004,  
579 2010), and that the more common an ant is the better a competitor it is. In our simulations this  
means that when one ant is removed the ants which are left occupy the newly vacated cacti so that  
the proportion of vacant cacti never changes. We included this assumption because we believe it  
582 to be the most biologically reasonable based on the personal observations of multiple researchers.  
However, it is possible this assumption is incorrect, so we created two alternative simulations to  
determine how robust our results were to various assumptions about partner turnover drivers. The  
585 first alternative we tested is a model that assumes all ants have the same competitive ability rather  
than the differences in competitive ability we assume cause observed variety in ant frequency. In

this hypothetical simulation all ants have equal probability of occupying any given cactus, but the level of vacancy across the population remains consistent with observed data. This simulation was designed to test if the overwhelming frequency of *L. apiculatum* may be overriding any potential advantages of other partners. The results of this simulation (Appendix : Figure 7) were effectively the same as the competitive exclusion model presented in the body of this paper (Figure: 6). The second alternative we tested was a model that assumed rather than competition driving ant occupations, it was the population size of each ant species alone. In this hypothetical simulation when one ant species is removed from the model the cacti are left vacant. We believe this is an unlikely scenario as there appear to be vast populations of the ant species in the area.<sup>4</sup> This simulation was designed to test if the fitness was more sensitive to changes in vacancy across the population than the frequency of any particular ant species. The results of this simulation revealed that when vacancy increased sampling effect is at play (Appendix : Figure 6). This result highlights that the primary driver behind the results we presented are ant presence and that it is only in cases of unrealistically high vacancy that there are benefits to partner diversity. Further information about both of these models can be found in Appendix C. Together, these results indicate that our assertions that the fitness of the cacti are more sensitive to changes in vacancy than partner identity are robust across multiple assumptions.

This paper shows the importance of long-term datasets in investigating species interactions and calls for further use of long-term data. Previously studies have analyzed how partner identity and partner turnover impact focal mutualist fitness Fonseca et al. (2003); Dejean et al. (2008); Noe and Hammerstein (1994); Barrett et al. (2015); Bruna et al. (2014); Trøjelsgaard et al. (2015). Separate studies have analyzed how inter-annual variability impacts focal mutualists Alonso (1998); Alarcón et al. (2008); Ollerton et al. (2006); Horvitz and Schemske (1990); Lázaro et al. (2022). The long term data set we used gave us the unique ability to consider the combined effects of partner identity, partner turnover, and temporal stochasticity. By piecing together complete life cycle information from long-term data, we gain a more nuanced understanding of the fitness consequences

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<sup>4</sup> *This is just personal observation so it might be best to remove this?*

of specific demographic effects. For example, our previous study suggested that *C. opuntiae* has overall parasitic fitness effects because activity of this species within tree cholla flowers can deter pollinators and reduce seed set (Ohm and Miller, 2014). Yet, the more complete analysis presented here, which accounts for reduced seed set alongside other demographic advantages, indicates that this species is clearly a mutualist and nearly as strong a mutualist as *L. apiculatum*.

Herbivory is an important driver of the fitness of cholla in this system, as herbivores directly impact the growth, survival, and reproductive efforts of the cacti (Miller et al., 2009). We have not explicitly incorporated herbivory but rather assume that herbivory can be captured in the effects they have on growth, survival, and reproductive effort observations. Analysis of our observational herbivore data shows that vacant plants experience elevated levels of herbivory compared to tended ones (Appendix : Figure 5). In the future, further studies on the direct impacts of herbivory would bolster the results reported here and expand our knowledge within this system.

As with any study, there are limitations to consider when interpreting the results reported here. These results highlighted in the paper are based on observational data regarding ant effects on plant demography rather than experimental data, meaning we are able to determine correlations but not causation. However, we have previously conducted experimental manipulations which revealed that ant presence has direct impacts on plant demography in our system (Miller, 2007). This combination of observational data backed up by experimental results gives us greater confidence in our causal interpretations than if we had only observational data. Specifically, we are confident in claiming that ant presence causes increased performance in cactus vital rates and the fitness signals we are detecting are directly related to ant partner presence. One further expansion could include nectar analysis in conjunction with ant interactions to look not only at the ant partner impacts on the plant demography, but also at the ability of plants to attract specific partners through nectar composition shifts.

This study highlights that while partners within a mutualistic guild can be ecologically different, they may still be interchangeable to a focal mutualist. The individual estimation of each vital rate and explicit inclusion of ant partners in this IPM allowed us to determine that partner diversity is

not beneficial in this system and may highlight circumstances under which we would expect similar results in other systems. The individual vital rate estimation approach allowed us to determine that while different ant partners did impact each vital rate differently, the largest differences were estimated to occur at sizes when ants would not be present in real life or in vital rates that do not have a significant impact on the overall fitness. Considering partner impacts on vital rates across the entire life cycle of the plant allows more power in determining the importance of these differences. Based on our results we would expect to see that partner diversity does not matter when partner impacts occur during low sensitivity life stages.

## Acknowledgments

We are grateful to the Sevilleta community (LTER and US Fish and Wildlife Service) for providing a stimulating research environment and logistical support. Jeremiah Dye identified ants and Massa Takahashi identified herbivores. We acknowledge the many students and technicians who have helped collect annual census data for this project, including M. Donald, J. Fowler, T. Jordan-Millet, J. Moutouama, C. Oxley, K. Schraeder, B. Scherick, A. Sears, M. Tucker, and J. Xiong. Financial support for this work came from the Sevilleta LTER (NSF DEB-0217774)

## Data and Code Availability

The data that support the findings of this study are openly available as a data package on the Environmental Data Initiative website, package id: knb-lter-sev.323.1

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## Tables

### Figure legends

### Appendix A: Additional Methods and Parameters

849 In addition to the models described in the body of the paper, we fit several other models using data from previous studies. These models are described below.

*Seeds Per Fruit.* With data from Miller et al. (2006), we fit a model for the number of seeds produced by every fruit on a cholla ( $\kappa(a')$ ) in year  $t + 1$  based on the ant partner  $a'$  in year  $t + 1$ . We fit this model to seed data  $y^\kappa$  using a Negative Binomial distribution and the log link function:

$$y^\kappa \sim \text{NegativeBinomial}(\hat{\kappa}, \hat{\phi})$$

$$\hat{\phi} = \beta_0^\phi$$

The data used for this model did not include data on ants in the “other” category, so we used the data from vacant plants to parameterize seeds per flower for plants with “other” ants in the IPM.

We found that vacant plants produced the most mean seeds (147.2 per fruit), followed by *L. apiculatum* tended plants (142.4 per fruit), and finally, *C. opuntiae* tended plants (115.0 per fruit) (Figure 1).

*Recruit Size Distribution.* We fit this model to recruit size data  $y^\eta$  using a Normal distribution with the identity link function:

$$y^\eta \sim N(\hat{\eta}, \hat{\sigma})$$

where  $\hat{\sigma}$  is estimated with a non-informative prior.

We found that the mean size of recruits is  $\log(-2.097)m^3$  with an interquartile range from  $\log(-2.173)m^3$  to  $\log(-1.712)m^3$  (Figure 2).

Parameter	Median (95% CI)	Prior Distribution
growth xi intercept vacant $\beta_{01}^g$	-5.210899(-5.686865, -5.491787)	sDE
growth xi intercept other $\beta_{02}^g$	-5.8288(-5.956217, 1.766021)	asdf
growth xi intercept <i>C. opuntiae</i> $\beta_{03}^g$	-4.529523(-6.0770390, 0.1222112)	asdf
growth xi intercept <i>L. apiculatum</i> $\beta_{04}^g$	-5.106802(-5.4499944, 0.5453901)	asdf
growth xi size dependent vacant $\beta_{11}^g$	asdf	asdf
growth xi size dependent other $\beta_{12}^g$	asdf	asdf
growth xi size dependent <i>C. opuntiae</i> $\beta_{13}^g$	asdf	asdf
growth xi size dependent <i>L. apiculatum</i> $\beta_{14}^g$	sadf	asdf
growth omega intercept $\omega_0^g$		
growth omega size dependent $\omega_1^g$		
growth alpha intercept $\alpha_0^g$		
growth alpha size dependent $\alpha_1^g$		
1-year germination intercept $\alpha^{\gamma_1}$		
2-year germination intercept $\alpha^{\gamma_2}$		
survival intercept vacant $\beta_{01}^s$		
survival intercept other $\beta_{02}^s$		
survival intercept <i>C.opuntiae</i> $\beta_{03}^s$		
survival intercept <i>L. apiculatum</i> $\beta_{04}^s$		
survival size dependent vacant $\beta_{11}^s$		
survival size dependent other $\beta_{12}^s$		
survival size dependent <i>C. opuntiae</i> $\beta_{13}^s$		
survival size dependent <i>L. apiculatum</i> $\beta_{14}^s$		
flowering intercept $\beta_0^f$		
flowering size dependent $\beta_1^f$		
viability intercept vacant $\beta_{01}^v$		
viability intercept other $\beta_{02}^v$		
viability intercept <i>C. opuntiae</i> $\beta_{03}^v$		
viability intercept <i>L. apiculatum</i> $\beta_{04}^v$	46	

Table 1: This table includes the median estimates, the 95% confidence intervals, and the prior distribution for each parameter in each model.

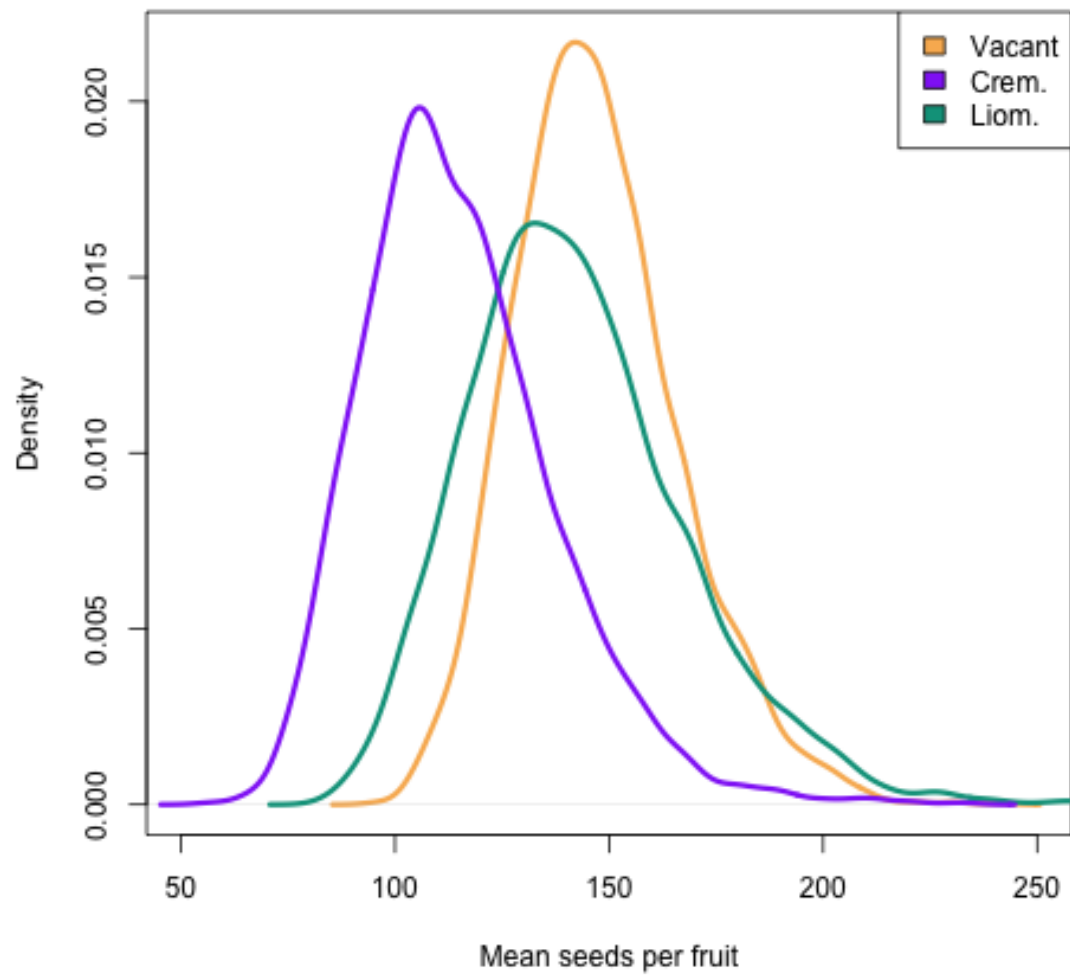


Figure 1: Shows the distribution of seeds per fruit on plants tended by *C. opuntiae* (purple), *L. apiculatum* (Teal), and Vacant plants (Yellow).

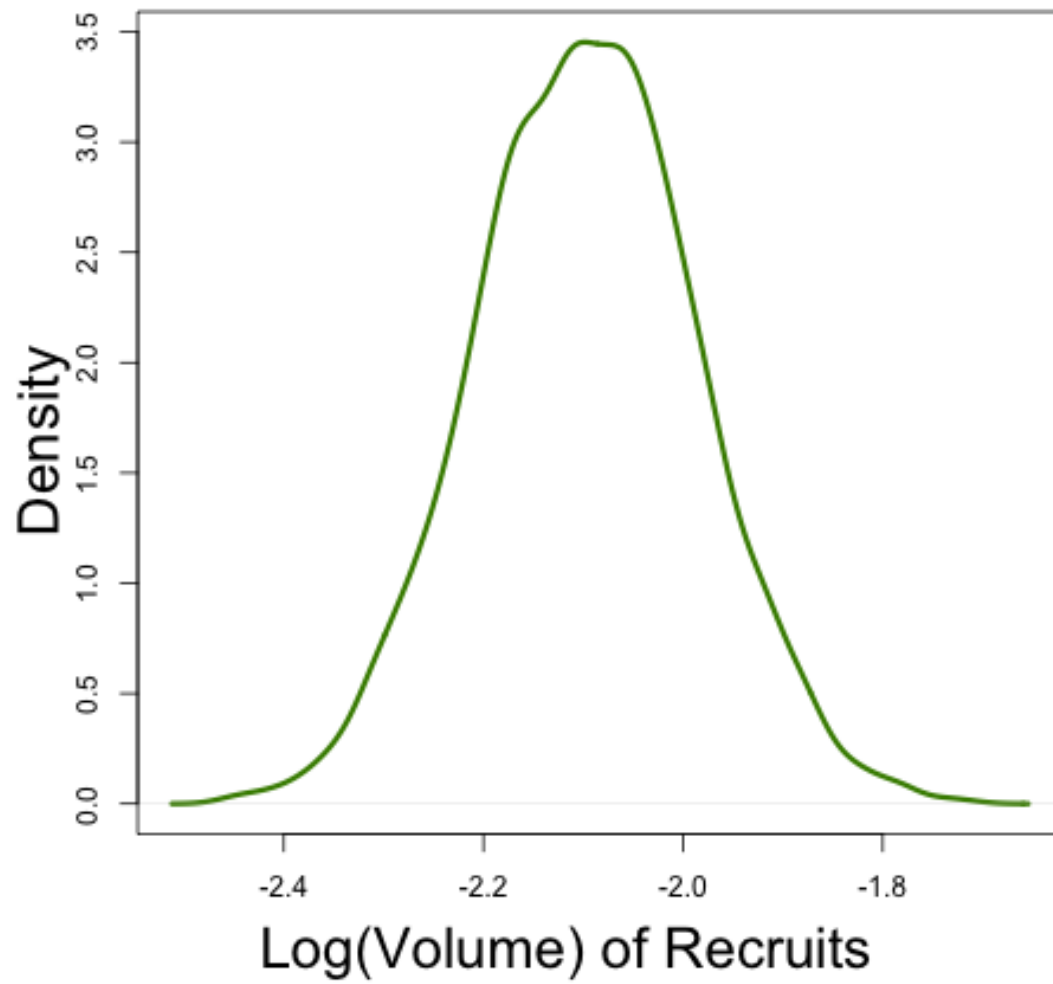


Figure 2: Shows the distribution of recruit sizes. The mean size of a recruit is marked by the dark line in the green box, the green box shows the interquartile range of recruit sizes, and the dashed lines show the minimum and maximum of estimated recruit sizes. The points beyond these dashed lines are considered outliers.



*Germination.* With germination data Miller (2007), we fit two models for the probability of germinating from the first year seedbank ( $\gamma_1$ ) or the second year seedbank ( $\gamma_2$ ) in year  $t + 1$ , with no fixed or random effects. These models were fit to germination data  $y^{\gamma_1}, y^{\gamma_2}$  using the binomial distribution with logit link functions:

$$y^{\gamma_1} \sim \text{Binomial}(\hat{\gamma}_1)$$

$$y^{\gamma_2} \sim \text{Binomial}(\hat{\gamma}_2)$$

We found that the mean germination rates for seeds in the seedbank for one year is 0%, with an interquartile range of 0% and 1%. We found that the mean germination rates for seeds in the  
861 seedbank for a second year is 0%, with an interquartile range of 0% to 0.4% (Figure 3). Seeds are more likely to germinate in their first year in the seedbank, but most seeds will never germinate.

*Pre-Census Survival.* With recruit census data Miller et al. (2006), we fit a model for the probability of a seedling (which germinates in early Fall) surviving to when we census in May ( $\delta$ ) of year  $t + 1$  (accounting for missed mortality events), with fixed effects of the previous size  $x$  and random effects of the transect  $m$ . We fit this model to pre-census survival data  $y^\delta$  using a Bernoulli distribution with a logit link function:

$$y^\delta \sim \text{Bern}(\hat{\delta})$$

where  $m \sim N(0, \sigma_{\text{transect}}^2)$  is the random effect of transect where the recruited individual was  
864 analyzed for survival.

We found that plants have a 16.2% probability of surviving from germination to the next census. Our model estimated this very well, expecting a 16.3% probability (Figure 4).

## 867 **Appendix B: Observed Herbivory Data**

Herbivory is an important driver in this population and shapes the range and demography of cholla. Herbivore presence has been shown to negatively impact growth and fecundity of cholla populations  
870 (Miller et al., 2009). Ant visitors are believed to offer defensive benefits to the plants they tend

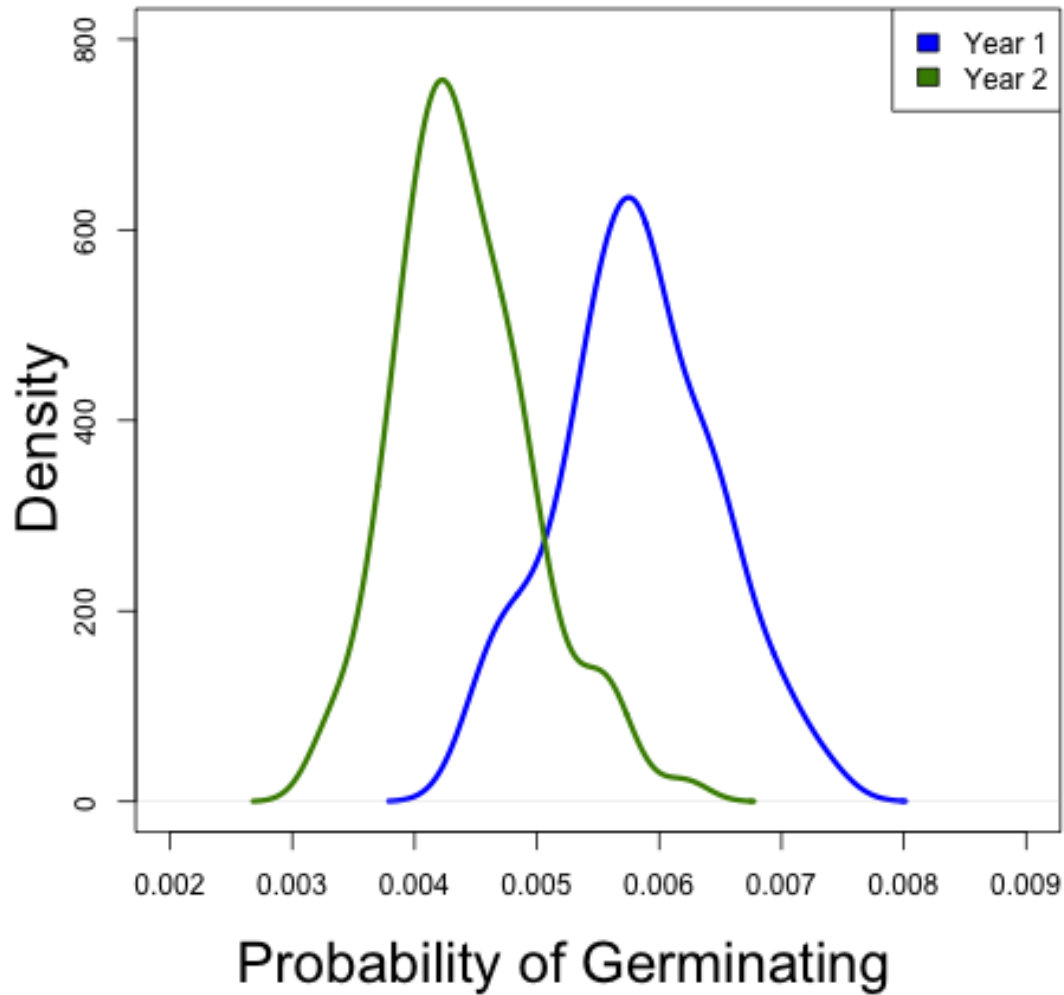


Figure 3: Shows the probability of a plant germinating during its first year in the seedbank (left) or during the second year in the seedbank (right). The mean germination rates are shown by the dark line near the bottom of the green box. The green boxes show the interquartile ranges of the germination rates for year 1 and 2 plants. The dashed line shows the maximum germination rates observed. The points above that show the outliers of our data.

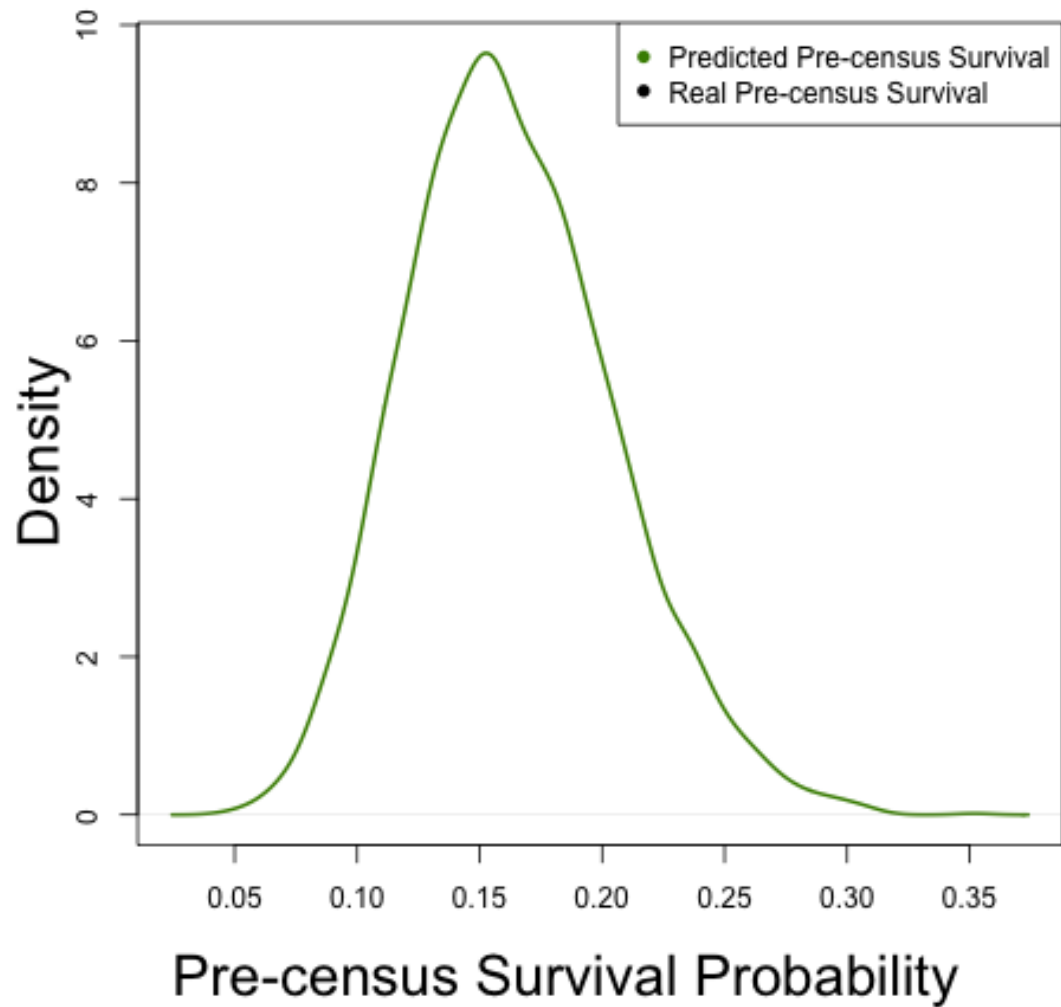


Figure 4: Shows the distribution of the probability of plants which germinated between censuses surviving to the next census to be counted. The green line shows the predicted pre-census survival and the grey dashed line shows the actual observed pre-census survival.

in this system, leading to the hypothesis that ant presence would be correlated with reduced herbivory. Herbivory data was collected during censuses any time herbivores were identified on a plant. This involved noting the type and quantity of herbivores observed. This data has been taken consistently since 2017, so the analysis below considers 6 years of data. We considered only plants which were reproducing, as they were likely to produce the highest quantity of EFN (Miller, 2014). The proportion of reproducing plants that experienced herbivory was calculated for each ant state separately. Analysis showed that ant presence is correlated with lower herbivore visitation. 40% of vacant cacti experienced herbivory. Plants tended by Other ants experienced similar, though lower, levels of herbivory on reproducing plants, with herbivores detected on 37.5% of plants. Herbivores were detected on 25% of plants tended by *C. opuntiae* ants and on 11% of plants tended by *L. apiculatum* ants. These results indicate that ant presence is correlated with lower levels of herbivory and that partner identity has an impact on the level of herbivory. They also indicate that the partner correlated with the lowest levels of herbivory is *L. apiculatum* ants. These findings are consistent with literature findings which show that *L. apiculatum* ants are the most aggressive (therefore the most effective against herbivores), but differ from previous findings that *C. opuntiae* may not offer defensive benefits (Miller, 2007).

## Appendix C: Alternative Ant Transition Simulations

In addition to the competitive exclusion model defined and analyzed in the body of the paper, we simulated results from several other potential models. We chose to include competitive exclusion as our primary results in the paper because we believe it to be the most biologically realistic. However, in building and testing of alternative models we found that the method of ants occupying plants significantly impacts the fitness of the population. We tested two alternative transition models, one called the frequency based model and one called the equal likelihood model.

*Frequency Based Model.* The first alternative hypothesis we tested was what we called the frequency based model. In this model rather than the proportion of vacant cacti being maintained, the

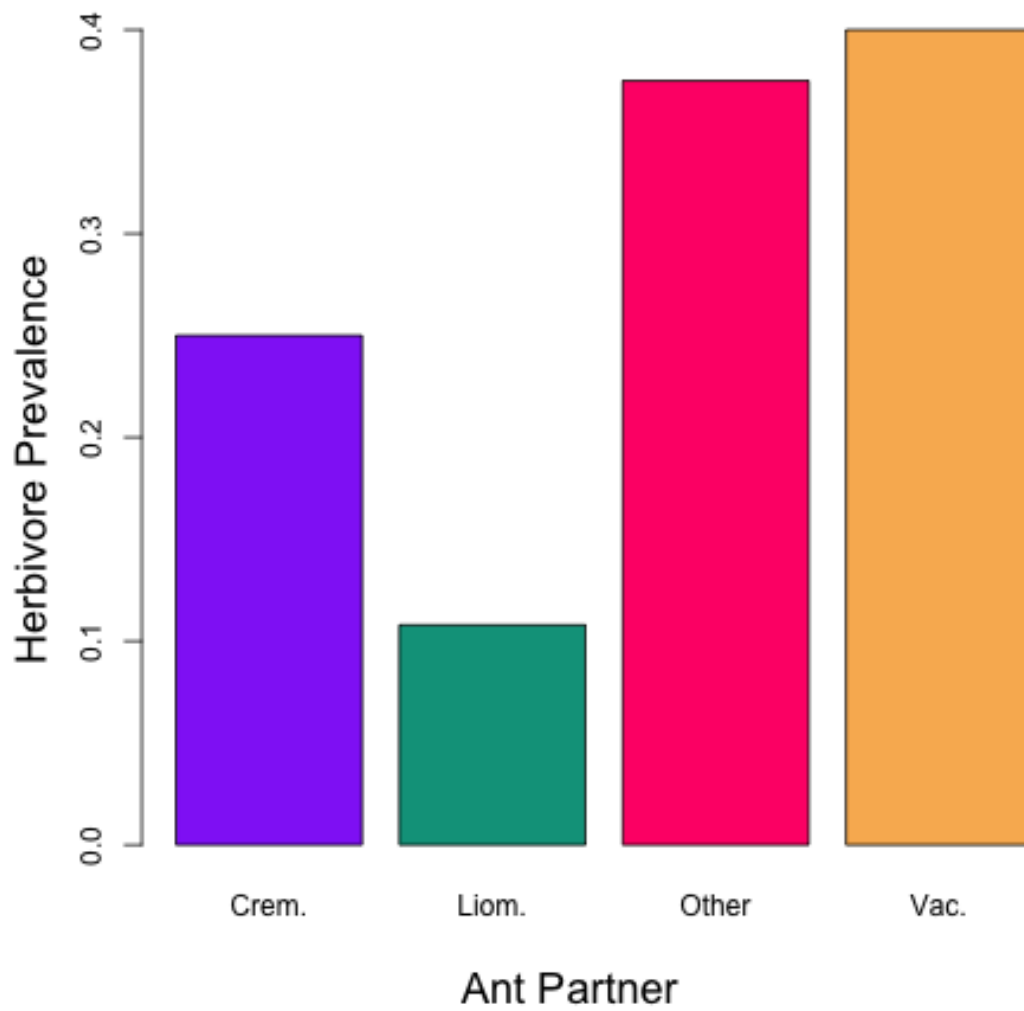


Figure 5: Shows the proportion of reproducing plants which are visited by herbivores. Each bar represents the subset of the cacti population in a different ant state.

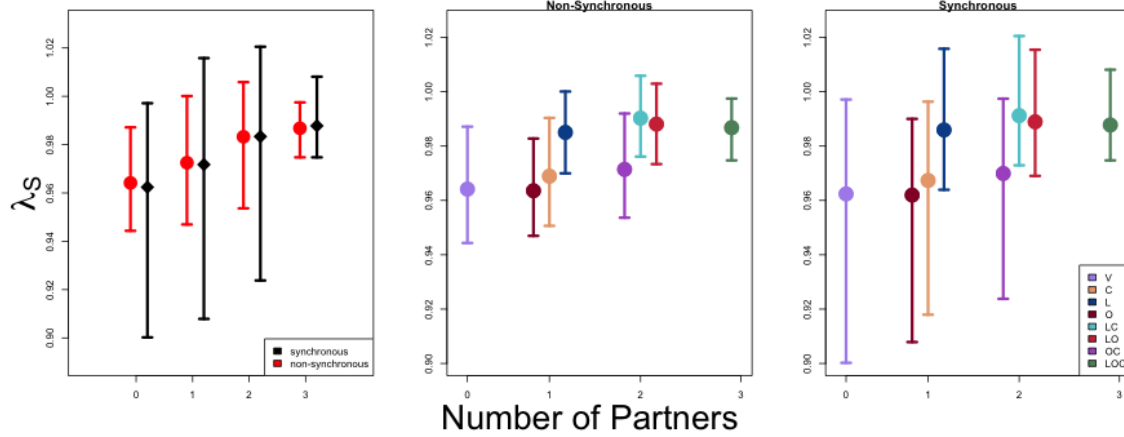


Figure 6: Panel a) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated  $\lambda$  values. Panels b-c) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated  $\lambda$  values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

proportion of cacti occupied by each species is maintained and when one is removed it is replaced with vacancy. This version of the model assumes that the frequency of each ant we see is reflective of the real frequency of populations rather than some other mechanism. With this model we found very clear evidence of Sampling Effect in the system. When only *C. opuntiae*, Other ants, or both ants are present, there is very little difference in the fitness of the cacti from when no partners are present. Only when *L. apiculatum* ants are present do we see an increase in the fitness of the focal mutualist (Figure ??a). In this simulation, the more partners that are present the higher the fitness of the focal mutualist is, confirming that partner diversity would be beneficial through sampling effect if this transition model were correct. (Figure ??b).

*Equal Likelihood Model.* The second alternative hypothesis we tested was what we called the equal likelihood model. In this model we preserved the observed pattern of size-dependent va-

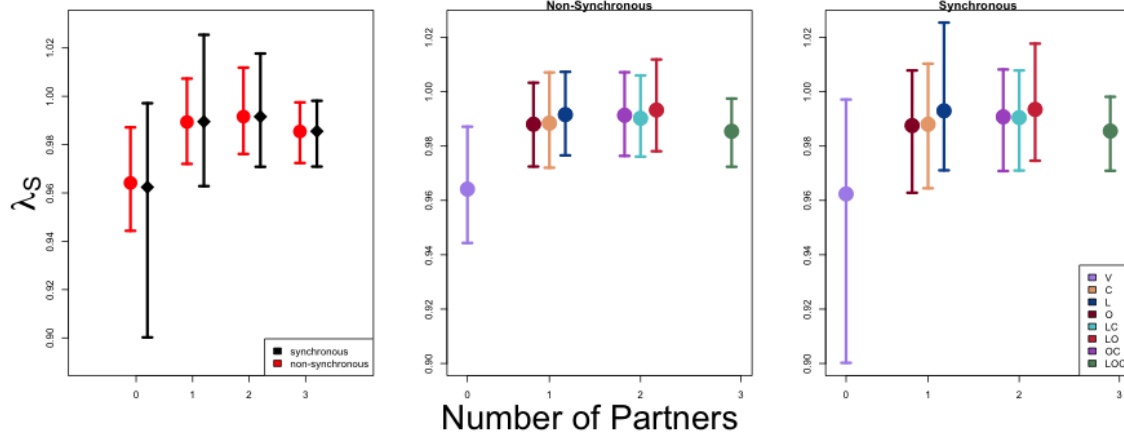


Figure 7: Panel a) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated  $\lambda$  values. Panels b-c) shows the mean estimated  $\lambda_{NS}$  and  $\lambda_S$  respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated  $\lambda$  values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

cancy/occupancy, but occupancy was manipulated to be equally likely for all partner identities. This was designed to remove the effect overwhelming numbers of *L. apiculatum* ants may have. Despite very different proportions, we found very similar outcomes to the competitive exclusion model analyzed in the paper. All ants are beneficial, but having more than one is not necessarily any better than having an individual species as a partner (Figure ??b). Partner presence is beneficial, but neither identity nor number of partners appears to be important (Figure ??a).

## Appendix D: Posterior Checks and Model Validation

For each model fitted, we conducted two tests to determining if the fit was acceptable to use in our IPM. First, we checked the convergence of each parameter. Below we show the convergence of all  $\beta$  terms listed in the Statistical Modeling subsection of Methods. Second, we checked the posterior

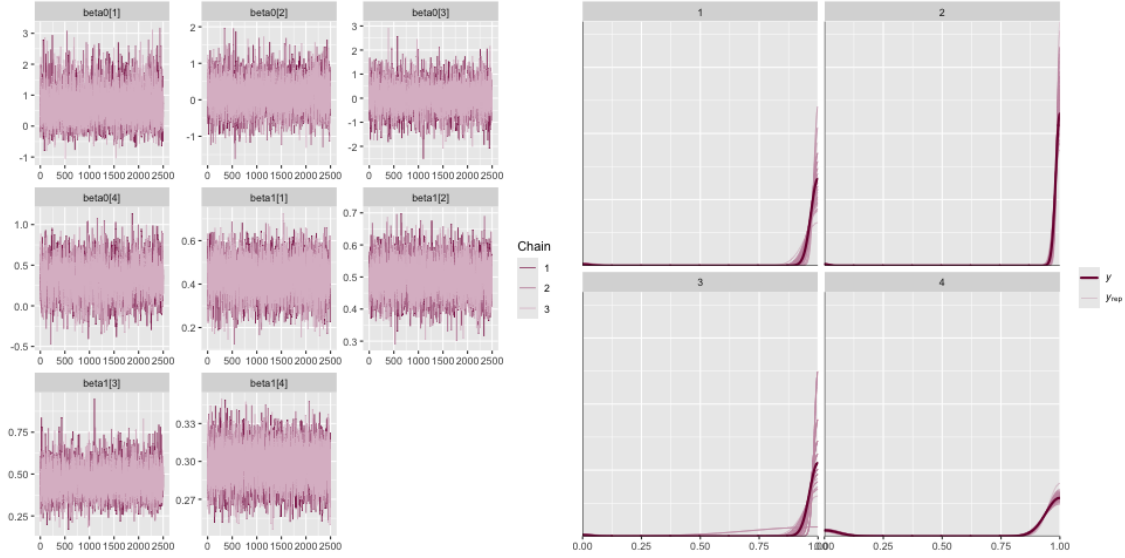


Figure 8: The a) posterior convergence of the parameters estimated by the survival model and the b) posterior distribution of survival estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean survival distribution (black line) of the real data.

fit, comparing the estimated values of each model to the  $y$  values of the actual data. We show these  
 918 posterior checks below, split by ant partner where relevant.

## Statistical Models – Results

Below are the results reorted of all statistical models not described in the main body of the text.

921 *Reproduction Model.* The probability of a plant reproducing in a given year is highly size dependent. The mean probability of reproducing remains at about 0% until the plant reaches a medium size, after which the mean probability of reproducing increases steadily before reaching about 100%  
 924 at large sizes.

*Seeds Per Flower Model.* Each viable flower on a plant produces between 97 and 257 seeds. This number is affected by the ant partner present, as shown in previous work (Ohm and Miller, 2014).



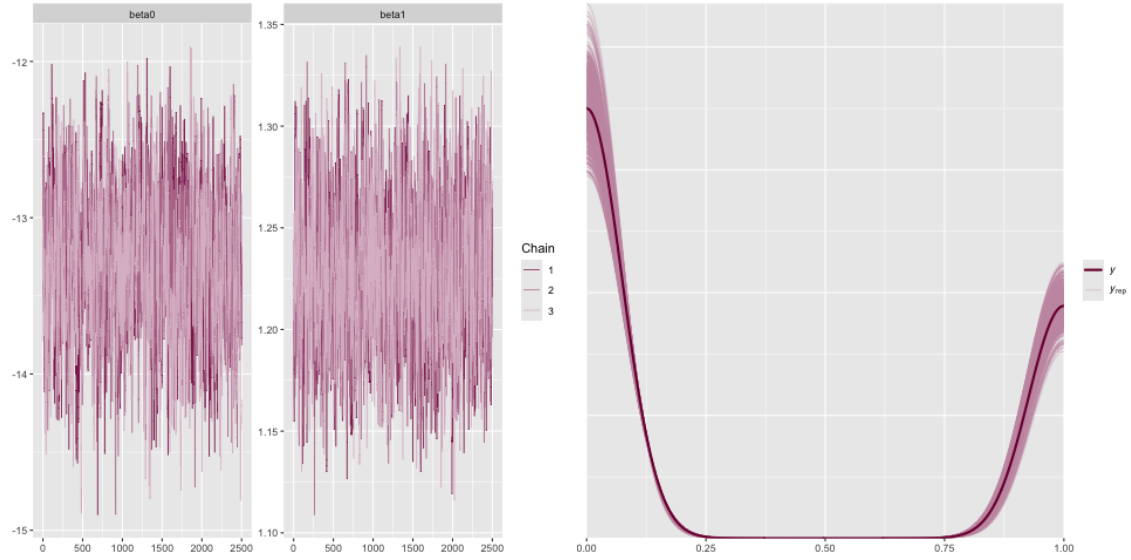


Figure 9: The a) posterior convergence of the parameters estimated by the reproduction model and the b) posterior distribution of reproductive status estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean reproductive status distribution (black line) of the real data.

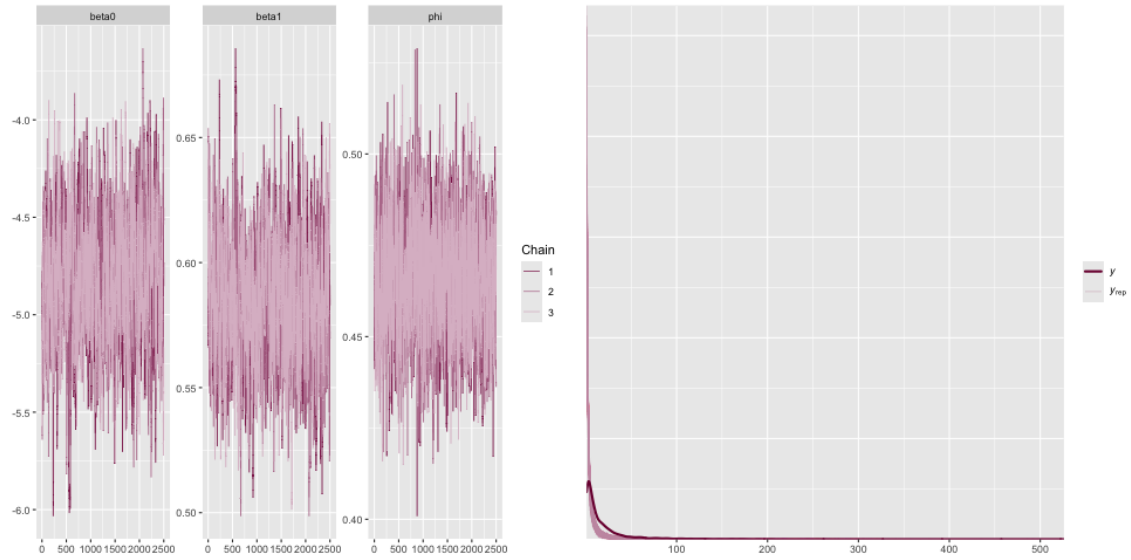


Figure 10: The a) posterior convergence of the parameters estimated by the number of flowers model and the b) posterior distribution of the number of flowers estimated (pink lines) compared to the mean distribution of observed flowers (black line).

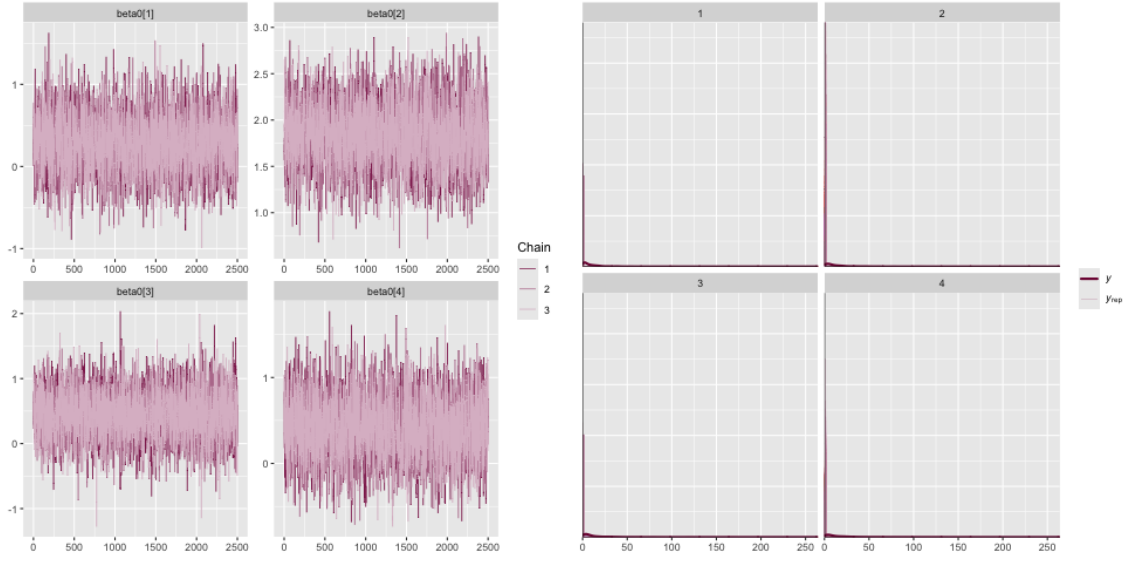


Figure 11: The a) posterior convergence of the parameters estimated by the viability model and the b) posterior distributions of floral viability estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean floral viability distribution (black line) of the real data.

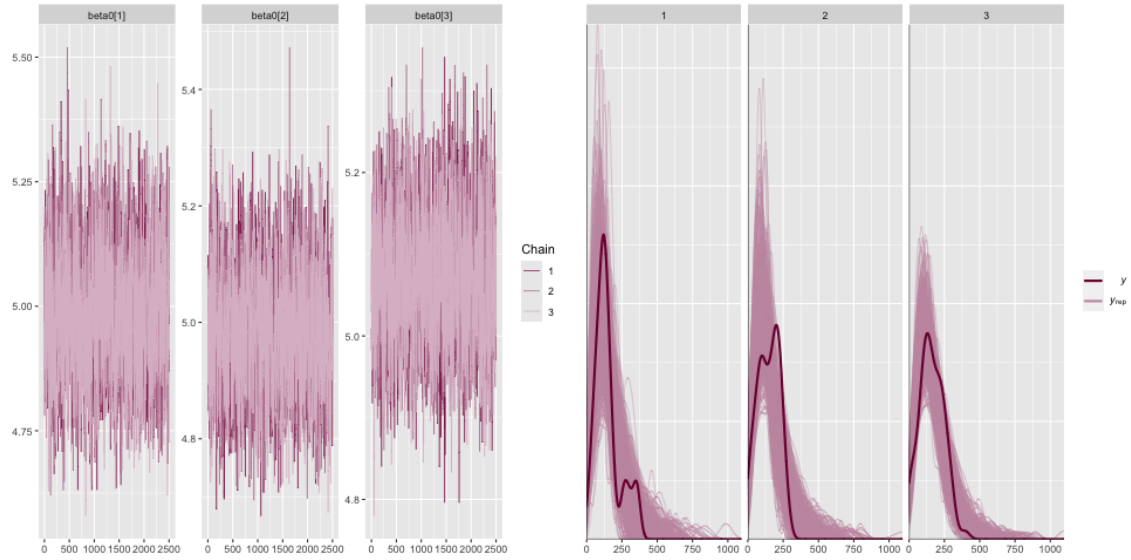


Figure 12: The a) posterior convergence of the parameters estimated by the seeds per fruit model and the b) posterior distributions of seeds per fruit estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = vacant) compared to the mean seeds per fruit distribution (black line) of the real data.

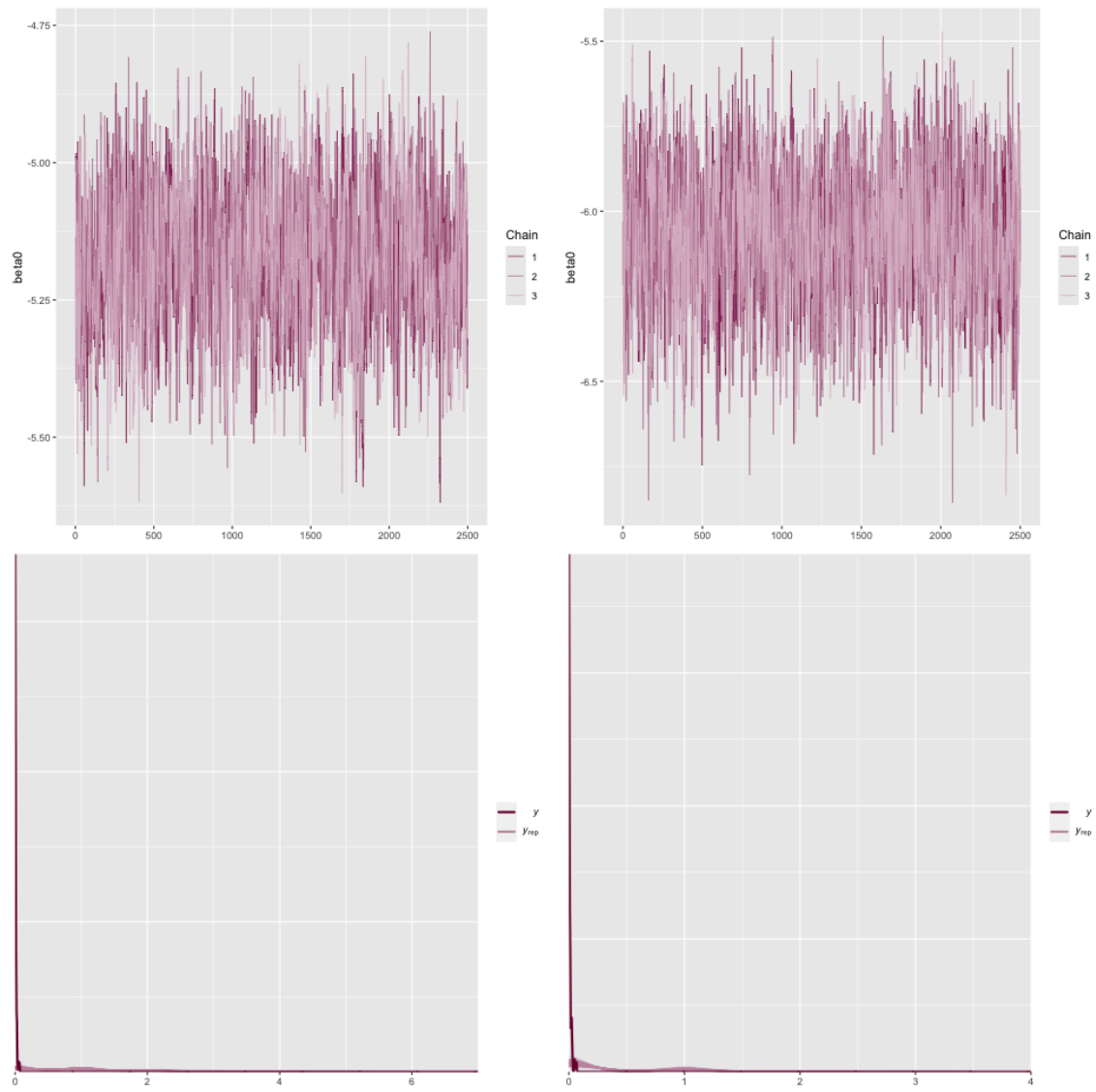


Figure 13: The a-b) posterior convergence of the parameters estimated by the germination from year one seedbank and germination from year two seedbank models respectively. The c-d) posterior distributions of floral viability estimates (pink lines) compared to the mean germination distribution (black line) of the real data for first year germinants and second year germinants respectively.

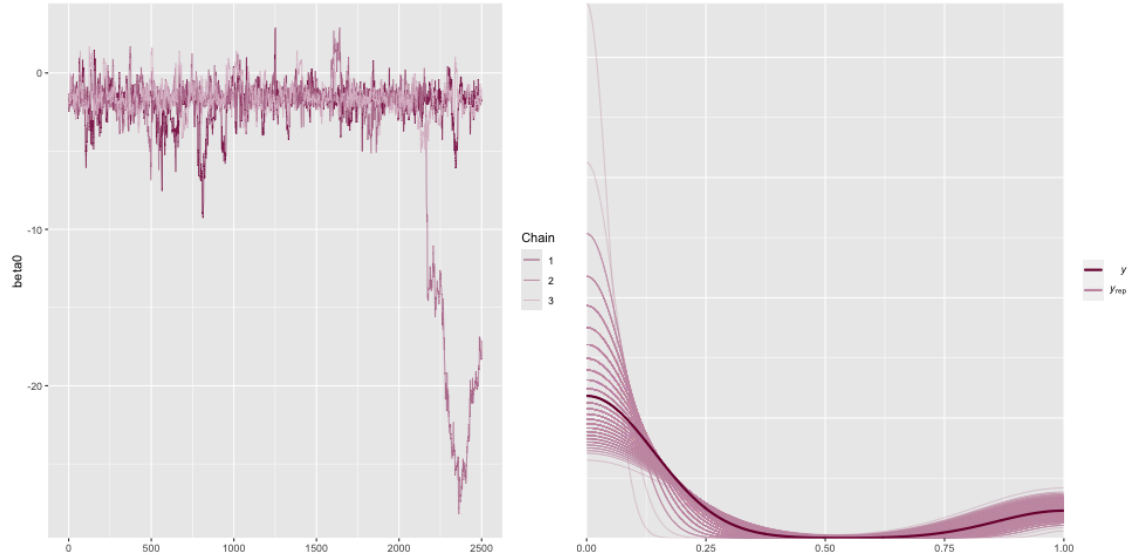


Figure 14: The a) posterior convergence of the parameters estimated by the pre-census survival model and the b) posterior distribution of the pre-census survival estimated (pink lines) compared to the mean distribution of observed pre-census survival (black line).

927 *C. opuntiae* tended plants produce a mean of 115 seeds per flower. *L. apiculatum* tended plants produce a mean of 143 seeds per flower. Vacant plants produce a mean of 148 seeds per flower. Comparison between posterior distributions revealed that *C. opuntiae* tended plants produced fewer  
 930 seeds per flower than *L. apiculatum* tended plants and vacant plants 80% and 87% of the time. Vacant plants produced more seeds per flower than *L. apiculatum* tended plants only 57% of the time. We are confident that *C. opuntiae* tended plants produce the fewest seeds per flowers.

933 *Precensus Survival Model.* Pre-census seed survival rates fall between 0% and 95% with the mean pre-census seed survival at 18%.

*Germination Model.* Seeds have a significantly higher probability of germinating in year one than  
 936 in year two. Seeds in year one experience germination rates between 50% and 100% with a mean of 62% germination. Seeds in year two experience germination rates between 50% and 98% with a mean of 58% germination.

939      New recruits are expected to be between the sizes of  $0.11 \text{ cm}^3$  and  $0.38 \text{ cm}^3$  with a mean size of  $0.20 \text{ cm}^3$ .