

Thank you, next: demographic consequences of partner diversity and turnover in a multi-species ant-plant mutualism

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Abstract

The diversity of partners in a multi-species mutualism causes varied demographic effects on the population of the focal mutualist which can be explained by several mechanisms: portfolio effect, complementarity, and sampling effect. *Cylindropuntia imbricata* (tree cholla) produce extrafloral nectar and various ant partners provide defense from herbivores and seed predators. We used plant demographic censuses to parameterize a series of Bayesian hierarchical generalized linear vital rate models to determine the impacts of different partners on the focal mutualists. We constructed an Integral Projection Model in which we simulate different combinations of ant partners that don't occur in nature. The hierarchical models revealed that different ant partners had different impacts on the cholla vital rates. Specifically, *C. opuntiae* tended plants had advantages in both growth and survival when small, and large *L. apiculatum* tended plants had floral viability advantages. The IPM results revealed that despite these differences in vital rates, the presence of any ant partner is beneficial, the identity and number of partners do not matter. This suggests that there are no benefits of diversity in this system. This study highlights that partner diversity does not always increase the benefits a focal mutualist receives.

Introduction

Mutualisms are species interactions where all participants receive net benefits, leading to higher individual fitness and increased population growth rates. They are widespread species interactions but can deteriorate into commensalism or parasitism under conditions that elevate costs or dampen benefits (Rodriguez-Rodriguez et al., 2017; Song et al., 2020; Mandyam and Jumpponen, 2014; Thrall et al., 2007; Bahia et al., 2022; Bronstein, 1994; Chamberlain et al., 2014; Frederickson, 2013; Axelrod and Hamilton, 1981; Leigh, 2010). Mutualisms are considered more context dependent than other species interactions (Chamberlain et al., 2014; Frederickson, 2013), meaning the magnitude and sign of interaction strength are often determined by environmental conditions and species' identities (Noe and Hammerstein, 1994; Leigh, 2010).

Mutualism is defined at the level of a species pair (+/+) but these interactions are embedded within multi-species communities, and growing evidence suggests that pairwise interactions are poor predictors of the net effects of multi-species mutualism (Afkhami, 2014; Palmer et al., 2010; Bascompte, 2009; Dattilo et al., 2014). A focal mutualist may interact with multiple guilds of partner types (e.g., plants that interact with pollinators, seed dispersers, soil microbes, and ant defenders) or with multiple partner species within the same guild (e.g., plants visited by multiple pollinator species). Within a mutualist guild, partner species often differ in the amount or type of goods or services they provide, making partner identity an important source of contingency in mutualism (Stanton, 2013). Whether and how partner diversity modifies the demographic effects of mutualistic interactions remain open questions within relevance in applied settings (Rogers et al., 2014; Thibaut et al., 2012; Frederickson, 2005; Palmer et al., 2010).

There are multiple mechanisms by which partner diversity can influence the net benefits accrued by a focal mutualist, mirroring the mechanisms by which, at a larger scale of organization, biodiversity can influence ecosystem function (Yeung et al. (2006); Barrett et al. (2015); Ushio (2020)). First, when there is a hierarchy of fitness effects (a consistent ranking of best to worst mutualists), a more diverse sample of the partner community may be more likely to include the best partner

Frederickson (2013). This can lead to an apparent benefit of diversity driven by a sampling effect Batstone (2018) – but there is no benefit of diversity *per se*, only better and worse partners. If partner associations are mutually exclusive (a focal mutualist may engage with only one partner at a time), then partner diversity may impose opportunity costs, leading to negative effects of a diverse mutualist assemblage relative to exclusive association with the single best partner (Miller, 2007). Second, even within a single mutualist guild, the benefits conferred by alternative partner species can vary in type and not just degree Stachowicz and Whitlatch (2005); Bronstein et al. (2006); Stanton (2013). This can lead to a positive effect of partner diversity through complementarity of alternative functions Batstone (2018). Interference or synergies between partners can make their combined effect different than the expected from the sum of complementary functions Afkhami (2014). Third, partner species can have species-specific responses to environmental variation, either spatially (Ollerton et al., 2006) or temporally (Alarcón et al., 2008). Multiple partners can therefore act as a 'portfolio' that stabilizes fitness benefits across spatial or temporal heterogeneity, leading to positive effects of partner diversity through the portfolio effect Batstone (2018); Lázaro et al. (2022); Horvitz and Schemske (1990).

Partner diversity can have different effects depending on whether partners are present simultaneously or sequentially (partner turnover) (Djiéto-Lordon et al., 2005; Ness et al., 2006; Bruna et al., 2014; Barrett et al., 2015; Ushio, 2020; Dattilo et al., 2014). Sequential associations are likely when alternative partners engage in interference competition for access to a shared mutualist Kiers et al. (2003); Batstone (2018); Trøjelsgaard et al. (2015); Wulff (2008). Turnover can happen at different timescales, from minutes to years (Oliveira et al., 1999; Horvitz and Schemske, 1986). The frequency of partner turnover can impact the level of benefits received by the focal mutualist, particularly if the benefits continue to accumulate with successive turnover (e.g., when sequential partners provide complementary functions) or if they saturate over time (Sachs et al., 2004; Fiala et al., 1994). Directionality of turnover can also influence effects of partner diversity if partner identity changes consistently across ontogeny of a focal mutualist (Fonseca et al., 2003; Noe and Hammerstein, 1994; Dejean et al., 2008). For example, plant susceptibility to enemies can change

across life stages (Boege and Marquis, 2005; Barton and Koricheva, 2010), so the benefits of a diverse guild of defensive mutualists are greatest when more defensive partner species align with more vulnerable life stages (Djiéto-Lordon et al., 2005; Dejean et al., 2008).

Defensive ant-plant mutualisms – where plants provide food and/or housing to ants that in turn defend them from enemies – are widespread interactions that offer valuable model systems for the ecology and evolution of mutualism (Bronstein, 1998; Bronstein et al., 2006). Extrafloral nectar (EFN) -bearing plants can serve as dietary resources that promote ant abundance and colony size (Byk and Del-Claro, 2011; Ness et al., 2009, 2006; Donald and Miller, 2022). In turn, presence of defensive ant partners is often linked to reductions in herbivory (Trager et al., 2010; Rudgers, 2004) and demographic advantages for the plant partner (Báez et al., 2016). Defensive ant-plant mutualisms are commonly multi-species, where a guild of ant partner species share, and often compete for, a plant mutualist (Bronstein, 1998; Beattie, 1985; Trager et al., 2010; Agrawal and Rutter, 1998). Ant partners can vary in their ability to deter herbivores (Bruna et al., 2014), and visitation by low quality ant partners can prevent visitation by higher quality partners, consequently causing a reduction in fitness through missed opportunity costs (Fraser et al., 2001; Frederickson, 2005). Susceptibility to herbivory can also vary significantly throughout the life stages of the plant (Boege and Marquis, 2005), suggesting that the order and timing of successive partners is important to the fitness impacts of the combined partner guild (Barton and Koricheva, 2010; Boege and Marquis, 2005; Fonseca et al., 2003). Finally, herbivore identity and pressure can vary inter-annually (Wetzel et al., 2023), much like mutualist identity and presence, meaning the threat plants face can vary just as much as the protection they receive due to temporal stochasticity. Previous studies have investigated how ant partner diversity affects plant fitness (Palmer et al., 2010; Afkhami, 2014; Fiala et al., 1994; Gaume et al., 1998; Dattilo et al., 2014; Ludka et al., 2015). However, little is known about the combined effects of partner identity, directional partner turnover, and temporal stochasticity, particularly because the necessary long-term data are rarely available.

This study examined the consequences of partner diversity in a food-for-protection mutualism

96 between the tree cholla cactus (*Cylindriopuntia imbricata*), a long-lived EFN-bearing plant, and
multiple species of ant partners. Previous studies have shown that herbivory by specialized in-
sect herbivores negatively affects cactus fitness Miller et al. (2009), and that ant defense reduces
99 herbivore damage Miller (2007). Tree cholla are tended by two common ant species (*Liometopum*
apiculatum and *Crematogaster opuntiae*) and several additional rarer species, all of which collect
EFN during foraging visits but their colonies are ground-nesting and not housed by the plants.
102 These ant species locally co-occur but individual plants are typically tended by only one species
that patrols the plant around-the-clock and maintains control of the plant’s nectar resources, usu-
ally for an entire growing season (Ohm and Miller, 2014; Donald and Miller, 2022). Switches
105 between partner species, or between vacancy and ant occupancy, commonly occur from one grow-
ing season to the next (Miller, 2007). Prior experiments suggested a hierarchy of mutualist quality,
with *L. apiculatum* providing strong anti-herbivore defense and *C. opuntiae* having net negative
108 effects because herbivore deterrence is outweighed by deterrence of pollinators (Miller, 2007; Ohm
and Miller, 2014). However, all of our previous studies in this system have focused on single life
stages (adult plants) or vital rates (seed production) and did not integrate the demographic effects
111 of ant defense across the life cycle, which may be essential for understanding net fitness effects of a
diverse partner guild (e.g., Palmer et al., 2010). To our knowledge no previous study has incorpo-
rated inter-annual stochasticity into models of ant-plant dynamics, which limits our understanding
114 of diversity benefits that may arise through the portfolio effect.

Here we used a unique long-term data set that allows us to explore mutualistic associations
with multiple partner species, longitudinal turnover in partner identity, and how the demographic
117 effects of alternative partners varies across plant size structure and nearly 20 years of inter-annual
fluctuations. We used this observational data set of plant demography and ant-plant associations,
contextualized by previous ant exclusion experiments, to investigate whether and through which
120 mechanism(s) partner diversity affects the fitness benefits of ant visitation. Specifically, we asked:

1. What are the demographic effects of association with alternative partners and how do these
effects fluctuate across years?

- 123 2. What are the frequency and direction of partner turnover across the plant life cycle?
3. What is the net effect of partner diversity on plant fitness, and what mechanism(s) explain(s) this effect?

126 To answer these questions we used a hierarchical Bayesian statistical approach to estimate demographic vital rates for hosts in different states of ant occupancy and to quantify state-dependent partner turnover from the long-term data. We then used a stochastic, multi-state integral projection model (IPM) that combines diverse effects on vital rates and pathways of partner turnover to
129 quantify effects of partner diversity on plant fitness.

Methods

Study System

132 This study was conducted in the Los Piños mountains, a small mountain chain located on the Sevilleta National Wildlife Refuge, a Long-term Ecological Research site (SEV-LTER) in central New Mexico, USA. This is an area characterized by steep, rocky slopes, and perennial vegetation
135 including grasses (*Bouteloua eriopoda* and *B. gracilis*), yuccas, cacti, and junipers. Tree cholla cacti are common in high Chihuahuan desert habitats of the southwestern USA (Benson, 1982).
138 These arborescent plants produce cylindrical segments with large spines. In the growing season (May to August in New Mexico), the plants initiate new vegetative segments and flowerbuds at the ends of existing segments. While most plants produce new segments every season, only those
141 which are reproductively mature produce flowerbuds. Like other EFN-bearing cacti, tree cholla secrete nectar from specialized glands on young vegetative segments and flowerbuds (Ness et al., 2006; Oliveira et al., 1999). Flowerbuds produce more and higher-quality EFN than vegetative
144 segments, making reproductive cholla valuable mutualist partners (Miller, 2014). Smaller, non-flowering cholla produce little to no EFN and are commonly vacant (no ant visitation at the time of our census) (Miller, 2014).

147 Tree cholla EFN is harvested by various ground-nesting ant species. At SEV-LTER, cholla
are visited primarily by two species, *Crematogaster opuntiae* and *Liometopum apiculatum*, as
well as other rarer species, including *Forelius pruinosus* and unidentified species in the genera
150 *Aphaenogaster* and *Camponotus*. *L. apiculatum* are the most frequent visitors with 25% – 60% of
tree cholla tended by these ants in a given year, followed by *C. opuntiae* (5% – 20%) (Donald and
Miller, 2022). Between 30% – 80% of cacti are vacant in any given year. Workers of different species
153 rarely co-occur on individual plants, likely due to interspecific competition. For example, staged
introductions of *C. opuntiae* to *L. apiculatum*-tended plants, and vice versa, provoke aggressive
responses by residents (A. Campbell, *personal observation*).

156 Several insect herbivores and seed predators specialize on tree cholla (Mann, 1969), and defense
against these enemies is the main pathway by which ant visitation affects plant demography. The
Cerambycid beetle *Moneilema appressum* and an unidentified weevil (Coleoptera: Curculionidae)
159 of the genus *Gerstaekeria* feed on vegetative and reproductive structures as adults and their larvae
feed internally. Two species of cactus bugs, *Narnia pallidicornis* and *Chelinidea vittiger* (Hemiptera:
Coreidae), feed on all cholla parts with a preference for flower buds; their damage can induce floral
162 abortion (Miller et al., 2006). A seed predator, *Cahela ponderosella* (Lepidoptera: Pyralidae),
oviposits in open flowers and larvae eat seeds in developing fruits. These consumers can have
significant negative impacts on plant fitness and depress population growth (Miller et al., 2009).
165 Prior experiments showed that ant-tended tree cholla experience less herbivory and seed predation
than plants from which ants were excluded (Miller, 2007; Ohm and Miller, 2014).

Data Collection

168 This study is based on long-term demographic data spanning 2004 to 2023 at SEV-LTER. From
2004 to 2008, we censused 134 plants distributed across three spatial blocks. This initial census
group was discontinued in 2009, when we established six 30 × 30-meter plots and tagged all tree
171 cholla within those plots. Two additional 30 × 30-meter plots were added in 2011, and this group
of eight plots has since been censused annually through 2023 (with the exception of 2020 due to

the pandemic shutdown). For all plants, in May or early June of each year we recorded plant survival since the last survey and, for survivors, we recorded height (cm), maximum crown width (cm), and crown width perpendicular to the maximum (cm). Size measurements were used to calculate plant volume (cm^3) based on the volume of an elliptical cone. We measured reproduction by counting flowerbuds, and we distinguished between flowerbuds that were viable and aborted. We recorded the ant species present (or vacancy if no ants present). Occurrences of more than one ant species on one plant were rare (less than 5% of observations), and for the purpose of this analysis we classified the plant as being occupied by the more abundant species. Plots were searched for new recruits each year, and these were added to the census. These data allowed us to link each plant's demographic fate (survival, growth, and reproduction) to its state of ant visitation. In total, the data set includes a total of 9,787 observations of 1141 unique individuals across 15 complete transition years (spanning May/June of year t to May/June of year $t + 1$) (Campbell and Miller, 2024).

We used additional, smaller data sets from previously published studies to estimate seed and seed bank parameters. Ohm et al. 2014 provide data on the number of seeds per fruit for plants tended by *L. apiculatum*, *C. opuntiae*, or no ants (experimental exclusion), accounting for their effects on pollinator visitation. Elderd and Miller 2016 provide data on seed entry to the seed bank and seedling germination and survival rates.

Multi-state Integral Projection Model

The demographic data were used to parameterize a multi-state Integral Projection Model (IPM). IPMs describe population dynamics in discrete time with functions that relate vital rates to continuous state variables, typically size (Ellner et al., 2016). While IPMs are a natural choice for populations with continuous size structure, they can also be modified to accommodate a combination of continuous and discrete state variables, as we do here. We constructed a stochastic, multi-state IPM that stitches together population structure associated with plant size and ant state, allowing us to determine the individual fitness effects of each ant species and the composite effects

of multiple partners, with ant transition dynamics and inter-annual variability modeled explicitly.

Given the low frequency of ant occupancy states other than *L. apiculatum* and *C. opuntiae* ($<8\%$ of observations) we combined all other ants into an “other” category, such that our multi-state IPM included four ant states: vacant, *L. apiculatum*, *C. opuntiae*, and Other. The “Other” category was made up of *Forelius pruinosus* (3.5% of observations), unidentified species belonging to the genera *Camponotus* (0.9%), *Aphaenogaster* (0.4%), *Myrmecocystus* (0.08%), *Tetramorium* (0.02%), *Brachymyrmex* (0.02%), and additional ants not identified to genus or species (2.8%).

Ant state is included as a predictor variable in IPM sub-models where there are biologically realistic pathways through which ants could impact the outcome of that process. For example, ant partners defend cacti from herbivores, and prior experimental work indicates that ant tending can reduce vegetative tissue loss and floral abortion. Therefore, ant state was included in sub-models for survival, growth, flowerbud viability, and seed number per flowerbud. In contrast, we have no reason to expect that ant tending can directly influence the probability of flowering or flowerbud production independently of its influence on plant size, so these sub-models do not include ant state as a predictor variable.

We modeled the tree cholla life cycle using continuously size-structured plants where number of plants of size x and ant state a in year t ($n(x, a)_t$) predicts the number of plants of size x' and ant state a' in year $t + 1$ ($n(x', a')_{t+1}$) based on a size- and ant-specific vital rates. The models also includes two discrete seed banks (B_t^1 and B_t^2) corresponding to 1 and 2-year old seeds. Seed bank dynamics are given by:

$$B_{t+1}^1 = \delta \sum_{a=1}^A \int_L^U \kappa(a) P(x; \boldsymbol{\tau}^P) F(x; \boldsymbol{\tau}^F) V(a; \boldsymbol{\tau}_a^V) n(x, a)_t dx$$

$$B_{t+1}^2 = (1 - \gamma_1) B_t^1$$

Functions $P(x; \boldsymbol{\tau}^P)$ and $F(x; \boldsymbol{\tau}^F)$ give the probability of flowering in year t and the number of flowerbuds produced in year t , respectively, by plants of size x in year t . The proportion of

flowerbuds that remain viable through fruit set ($V(a; \boldsymbol{\tau}_a^V)$) and the number of seeds per fruit ($\kappa(a')$) is dependent on ant state a . The vectors $\boldsymbol{\tau}$ give year-specific deviates (mean zero) and appear in functions for which we can estimate temporal stochasticity from the long-term data; superscripts indicate the corresponding vital rate and, when present, the a subscript indicates that deviates are specific to plants in ant state a . For example, temporal deviates $\boldsymbol{\tau}_a^V$ describe better- and worse-than-average years for flowerbud viability and plants in different ant states can fluctuate independently (good years for *L. apiculatum*-occupied plants may not be good years for *C. opuntiae*-occupied plants, for example). Seed production is integrated over the size distribution, from the lower L to upper U size limits, and summed over all possible initial ant states ($A = 4$) giving total seed production. Seeds are multiplied by the probability of escaping post-dispersal seed predation (δ) to give the number of seeds that enter the one-year old seed bank. Plants can recruit out of the year-one seed bank with probability γ_1 or transition to the two-year seed bank with a probability of $1 - \gamma_1$. Seeds in the two-year seed bank are assumed to either germinate with probability γ_2 or die.

For the above-ground part of the life cycle, the number of plants of size x' and ant state a' in year $t + 1$ is given by survival/growth transitions from size x and ant state a in year t , plus germination out of the seed banks:

$$n(x', a')_{t+1} = (\gamma_1 B_t^1 + \gamma_2 B_t^2) \eta(x') \omega \rho_0(a') + \sum_{a=1}^A \int_L^U S(x, a; \boldsymbol{\tau}_a^S) G(x', x, a; \boldsymbol{\tau}_a^G) \rho(x, a, a'; \boldsymbol{\tau}^\epsilon) n(x, a)_t dx$$

The first term in Eq. ¹ estimates the number of individuals recruiting from a one or two-year seed bank to a plant of size x' and ant state a' based on the recruit size distribution $\eta(x')$ and the probability of over-winter seedling survival (ω) from germination (late summer) to the census (May). This term is multiplied by $\rho_0(a')$, which gives the probability that a new recruit has ant

¹ We should label s . I am not sure why the equation label is not working here and I did not try to figure it out. It is probably something with the `linenomath` formatting.

state a' ($\sum \rho_0(a') = 1$). The second term represents all possible transitions from size x and ant a to size x' and ant a' , conditioned on survival. Survival ($S(x, a; \boldsymbol{\tau}_a^S)$) and growth from size x to x' ($G(x', x, a; \boldsymbol{\tau}_a^G)$) are both dependent on initial size and ant state. As above, these functions include inter-annual variability through year-specific deviates that can vary by ant state ($\boldsymbol{\tau}_a$). Finally, ant transition function $\rho(a', a, x; \boldsymbol{\tau}^\rho)$ gives the probability that an individual transitions from ant state a to a' in the next census, conditional on initial size x . This function includes inter-annual variability through year-specific intercepts which are consistent across initial ant states ($\boldsymbol{\tau}^\rho$).

Statistical modeling and parameter estimation

We parameterized the IPM using a series of generalized linear mixed models in a hierarchical Bayesian framework. Vital rate models included spatial and temporal random effects associated with plot and year variation, respectively (only year variation is used in the IPM), and included plant size (the natural logarithm of volume, $\log(\text{cm}^3)$; x, x'), ant partner state (a, a'), or both as fixed-effect predictor variables. As in the IPM, our statistical modeling assumed that demographic effects of ant occupancy are limited to survival, growth, and flowerbud viability.

Growth. We fit the growth sub-model ($G(x', x, a; \boldsymbol{\tau}_a^G)$) to data on size in year $t + 1$ (y^G) using the skewed normal distribution to account for left-skewed size transitions (at some initial sizes, transitions below the expected future size were more common than transitions above it). The skew-normal has three parameters corresponding to location (\hat{G}), shape (σ), and scale (α):

$$y_i^G \sim \text{SkewedNormal}(\hat{G}_i, \sigma_i, \alpha_i)$$

$$\hat{G}_i = \beta_{a[i]}^0 + \beta_{a[i]}^1 x_i + \beta_{a[i]}^2 x_i^2 + u_{\text{year}[i], a[i]} + w_{\text{plot}[i]}$$

$$\log(\sigma_i) = \beta^3 + \beta^4 x_i$$

$$\alpha_i = \beta^5 + \beta^6 x_i$$

Here, the location parameter for the i th observation \hat{G}_i is defined as a second-order polynomial with ant-size interactions because preliminary analysis found this was an improvement over a linear

relationship (note that the location parameter of the skew-normal is not the mean, but the mean
 can be derived as $\hat{G} + \frac{\sigma\alpha}{\sqrt{1+\alpha^2}}\sqrt{\frac{2}{\pi}}$). The year- and ant-specific random effect u (which parameterizes
 the τ_a^G vectors) and plot-specific random effect w are normally distributed with variances σ_{year}^2 and
 σ_{plot}^2 , respectively. Parameters σ_i and α_i control residual variance and skewness, respectively, and
 were defined as linear functions of initial size x_i (σ_i is strictly positive and was modeled with a log
 link function). We assume growth variance and skewness were not dependent on ant occupancy
 state.

Survival. The survival sub-model ($S(a, x; \tau_a^S)$) estimates the probability of survival from year t to
 year $t + 1$, with fixed effects of size x and ant partner a in year t . We fit this model to the survival
 data (alive or dead) using a Bernoulli distribution with a similar linear predictor for the probability
 of survival as in the growth model but with a logit link function and without the second-order
 influence of size.

Reproduction. The flowering sub-model ($P(x; \tau^P)$) estimates the probability of reproducing in
 year t , with fixed effects size x in year t and random effects of plot and year. We fit this model
 to the reproductive status data (vegetative or flowering) using a Bernoulli distribution and a logit
 link function, similar to the survival model above but with no ant effects. The flowerbud function
 $F(x; \tau^F)$ estimates the total flowers produced by a reproducing plant in year t , with fixed effects of
 size x in year t . We fit this model to flowerbud count data (sum of viable and aborted buds) using
 a zero-truncated negative binomial distribution with a log link and normally distributed year and
 plot random effects.

The flowerbud viability sub-model ($V(a; \tau_a^V)$) estimates the probability that flowerbuds pro-
 duced in year t remain viable (not aborted), with fixed effects of ant partner a in year t . We fit
 this model to floral viability data using a binomial distribution where trials and successes are given
 by the total number of flowerbuds and the number that are viable, respectively. This model used
 a logit link function and included random effects for plot and year following the same structure as
 the growth and survival models, with ant-specific year random effects.

Estimates for the number of seeds per fruit were obtained from a field experiment which excluded
 303 ants (cite Tom’s data). This data only included *L. apiculatum* and *C. opuntiae* ants, so we had to
 make an assumption about the seeds per fruit on plants tended by Other ants. We chose to use
 the vacancy estimates for plants tended by other ants, a decision which does not have a significant
 306 impact on the final results. Additional reproductive parameters for the number of seeds per fruit,
 probability of entry to the seed bank, germination rates, and recruit size were estimated following
 methods described in Appendix XX.

309 *Ant Transitions.* The ant transition model ($\rho(x, a, a'; \boldsymbol{\tau}^\epsilon)$) estimates the probability of a cactus
 being occupied by ant partner a' in year $t + 1$ given that it was occupied by initial ant partner a in
 year t , with fixed effect of initial size x . We fit this model to ant partner data using a multinomial
 312 distribution with a logit link function.

Parameter estimation. We fit models using Stan run through version 4.0.2 of R R Core Team
 (2023); Stan Development Team (2023). We used vague priors for all parameters. For each model,
 315 we obtained three chains of 10,000 iterations, discarding the first 1,500 iterations. We visually
 assessed parameter convergence between and within chains (Appendix : Figures 8 – 14 b) and
 assessed overall model fit with posterior predictive checks to examine how well the fitted model can
 318 generate simulated data similar to the real data (Appendix : Figures 8 – 14 a).

IPM Analysis

Analyzing the IPM required that we discretize the continuous IPM kernel into an approximating
 321 matrix. Size variable x was discretized into b bins, resulting in a $b \times b$ matrix. In our model there is
 additional complexity in the form of transitions between A ant states and two additional discrete
 states (year one and year two seed banks), leading to a matrix size of $A(b + 2) \times A(b + 2)$. We
 324 used $b = 500$ bins, which we found to be sufficient for numerically stable outputs, and extended the
 integration limits beyond the minimum (L) and maximum (U) observed sizes to avoid unintentional

eviction using the “floor-and-ceiling” method Williams et al. (2012).

For stochastic analyses, we estimated the approximating matrix corresponding to each t to $t + 1$ transition year. To estimate population mean fitness in a stochastic environment where ants responded to annual variation uniquely $(\lambda_{NS})^2$ we simulated population dynamics for 500 years by randomly sampling among the 16 annual transition matrices, discarding the first 100 years of the simulation to minimize the influence of initial conditions. Sampling observed transition matrices (rather than independently sampling regression coefficients) produces demographic time series that realistically capture inter-annual variation by preserving correlations between vital rates Metcalf et al. (2015). We tallied the total population size at each time step as $N_t = B_t^1 + B_t^2 + \sum_{a=1}^A \int n(x, a)_t dx$ and calculated the stochastic growth rate as $\log(\lambda_S) = E[\log(\frac{N_{t+1}}{N_t})]$ (Mark and Ellner, 2009). We propagated uncertainty from the vital rate models using 100 draws from the joint posterior distribution of model parameters, resulting in a posterior distribution of λ_{NS} and other derived quantities.

Partner diversity simulation experiments

Using the fully parameterized multi-state IPM, we conducted simulation experiments to quantify how diversity and identity of ant partners influenced plant fitness in a stochastic environment (λ_{NS}). From the full version of the model (described above) corresponding to the observed assemblage of partners and observed frequencies of partner transition, we created treatments corresponding to all eight “counter-factual” scenarios of diversity and composition: no ant partners (complete vacancy); one ant partner (*C. opuntiae* only, *L. apiculatum* only, Other only); two partners (all pairwise combinations of *C. opuntiae*, *L. apiculatum*, and Other); and three partners (observed scenario of all ant states). These simulation experiments were made possible by extrapolating ant-specific demographic performance across the size distribution, even for combinations of size and ant occupancy that were rarely observed. For example, the no-partner scenario modeled a hypothetically ant-free cactus population, even though no such population exists to our knowledge,

²*Let’s discuss this symbology.*

by applying the statistical knowledge gleaned from vacant plants across the size distribution.

In all scenarios that included any ant partners, we preserved the observed pattern of size-dependent vacancy/occupancy (estimated through the ant transition sub-model) and manipulated partner identity conditional on occupancy. This means, for example, that the *C. opuntiae*-only scenario included two possible states, vacancy and occupied by *C. opuntiae*. While our statistical models allow us to extrapolate the demographic performance of ant-tended plants to small sizes that are typically vacant, the natural history of this system tells us that this is not biologically sensible. Small, non-reproductive plants are typically vacant because they do not produce extrafloral nectar, and once plants begin producing nectar they are nearly always ant-tended (Miller, 2014). Our simulation experiments preserved this basic biology, avoiding tiny ant-occupied plants that do not and could not occur in nature.

The partner diversity treatment scenarios required additional assumptions about the mechanisms that give rise to observed occupancy patterns. Based on evidence that EFN-bearing cacti are nearly always ant-occupied (Miller, 2014), we assume that ant partners competitively exclude one another from EFN-bearing cacti and that competition is zero-sum. This means that, in scenarios that remove species from the partner community, remaining species gain access to plants that the removed species would have tended. In Appendix , we present results under an alternative assumption, that ant visitation is limited by factors other than availability of cactus EFN (e.g., nesting sites or off-plant dietary resources), such that when a species is removed from the partner community, the plants it would have occupied remain vacant.

Temporal stochasticity experiments

Under the portfolio effect hypothesis, partner diversity may confer a fitness advantage when the benefits of alternative partners are not perfectly synchronized across temporal environmental variation, yielding an advantage of a diverse “portfolio” of partners when the environment fluctuates. Our statistical estimation of ant-specific year random effects in the vital rates allows for this possibility. We constructed two versions of the stochastic, multi-state IPMs that allowed us to test

this hypothesis by exploring two scenarios of environmental variation. The 'non-synchronous' (NS) version included ants effects that varied uniquely across time and the 'synchronous' (S) version included ant effects that were forced to be the same across species. First, we evaluated the model using empirical estimates for the τ_a vectors that describe ant-specific year deviates. In this scenario, good years and bad years can differ between ant states, according to the empirical parameter estimates. We also quantified from the fitted random effects how tightly inter-annual variation was correlated between ant states. Second, we averaged the ant-dependent vital rates (survival, growth, flowerbud viability) across ant-specific year random effects, thus ensuring that plants in all ant states fluctuated synchronously in response to temporal environmental variation. We evaluated a second, "synchronized" version of the model that effectively turns off any portfolio effect, holding all else equal. Both scenarios of temporal stochasticity, non-synchronized and synchronized, were run for all eight ant partner scenarios described above.

Statistical inference on fitness consequences of partner identity and diversity

The range of models we created generated many outputs; we focus our inference on the following specific contrasts. First, to determine whether ant occupancy and partner diversity are beneficial, we calculated a posterior distribution of λ_{NS} for each of four partner richness levels (zero, one, two, three), averaging over composition scenarios within each level. If cactus fitness increases with partner richness, this would be interpreted as evidence for benefits of partner diversity. Second, to determine whether each partner, in isolation, confers a fitness advantage and to rank alternative partners, we contrasted the fitness of each single partner scenario (*C. opuntiae* only, *L. apiculatum* only, Other only) against vacancy (zero partners). Third, to determine whether any benefits of diversity are due to the sampling effect or complementarity, we contrasted the fitness of multi-partner scenarios against the single best partner scenario. If the best multi-partner scenario exceeds the fitness associated with the best single partner, this would be interpreted as evidence of complementarity, a true benefit of diversity *per se*. Alternatively, the sampling effect hypothesis predicts that no multi-partner scenario yields higher plant fitness than the best single partner. It

is also possible that multi-partner scenarios yield lower fitness than the single best partner, which would be consistent with an opportunity cost of diversity. Fourth, to quantify any contribution of the portfolio effect, we contrasted λ_{NS} of the full (four-state) scenario to vacancy for synchronized and non-synchronized responses to temporal stochasticity (as a measure of how much benefit is recieved from all partners being present). If the portfolio effect confers a benefit of diversity, the fitness advantage of having all vs. no partners should be greater when temporal fluctuations are not synchronized across ant states.

We base our statistical inferences on the posterior probability distributions of the contrasts described above. For example, the contrast of *C. opuntiae* with vacancy yields a posterior distribution of the difference in λ_{NS} ($\Delta\lambda_{C-V}$). We can quantify from this distribution our certainty in the mutualistic effect of *C. opuntiae*, given the data, as $Pr(\Delta\lambda_{C-V} > 0)$. We apply similar logic to other contrasts described above.

Results

*What are the demographic effects of association with alternative parnters
and how do these effects fluctuate across years?*

Over the 20-year data set, we found that ant partners influenced demographic performance of cactus hosts, and different ant partners had contrasting demographic effects across host vital rates. Plants tended by *C. opuntiae* had a growth advantage, particularly for small plants, while plants in states of *L. apiculatum*, Other ants, and vacancy had indistinguishable growth trajectories (Figure 1). For all ant states, growth was left-skewed, with transitions to sizes below the mean were more common than sizes above the mean. Similarly, for plants which were large enough to have ant visitors (larger than $0.8 \log(cm^3)$), visitation enhanced cactus survival (Figure 2). Plants that were smaller than $-2 \log(cm^3)$ (indicated by the dashed lines for ant estimates, Figure 2) were estimated to experience survival boosts from vacancy. Due to the assumptions in our models based on real

ant presence data, tended plants of the size where vacancy would offer a boost are never seen and therefore do not influence the fitness estimates of cacti. ³ The smallest plants showed evidence

429 that vacancy resulted in the highest survival rates, however these plants are never tended, meaning the models of ant-tended cacti survival before the solid lines are extrapolation. Mean survival rates ranged from 7.7% to 99.9%, with the smallest plants the most vulnerable to mortality. *C. opuntiae*-occupied plants had a survival advantage over other ant-tended plants, particularly at
432 smaller sizes, consistent with the positive effects on growth. At larger sizes, plants in any state of ant occupancy had a survival advantage over vacant plants.

³*I'm not 100% certain on how to discuss this. I know that this still shows that crem offer a survival advantage, but I see that it could look like vacancy offers survival advantages.*

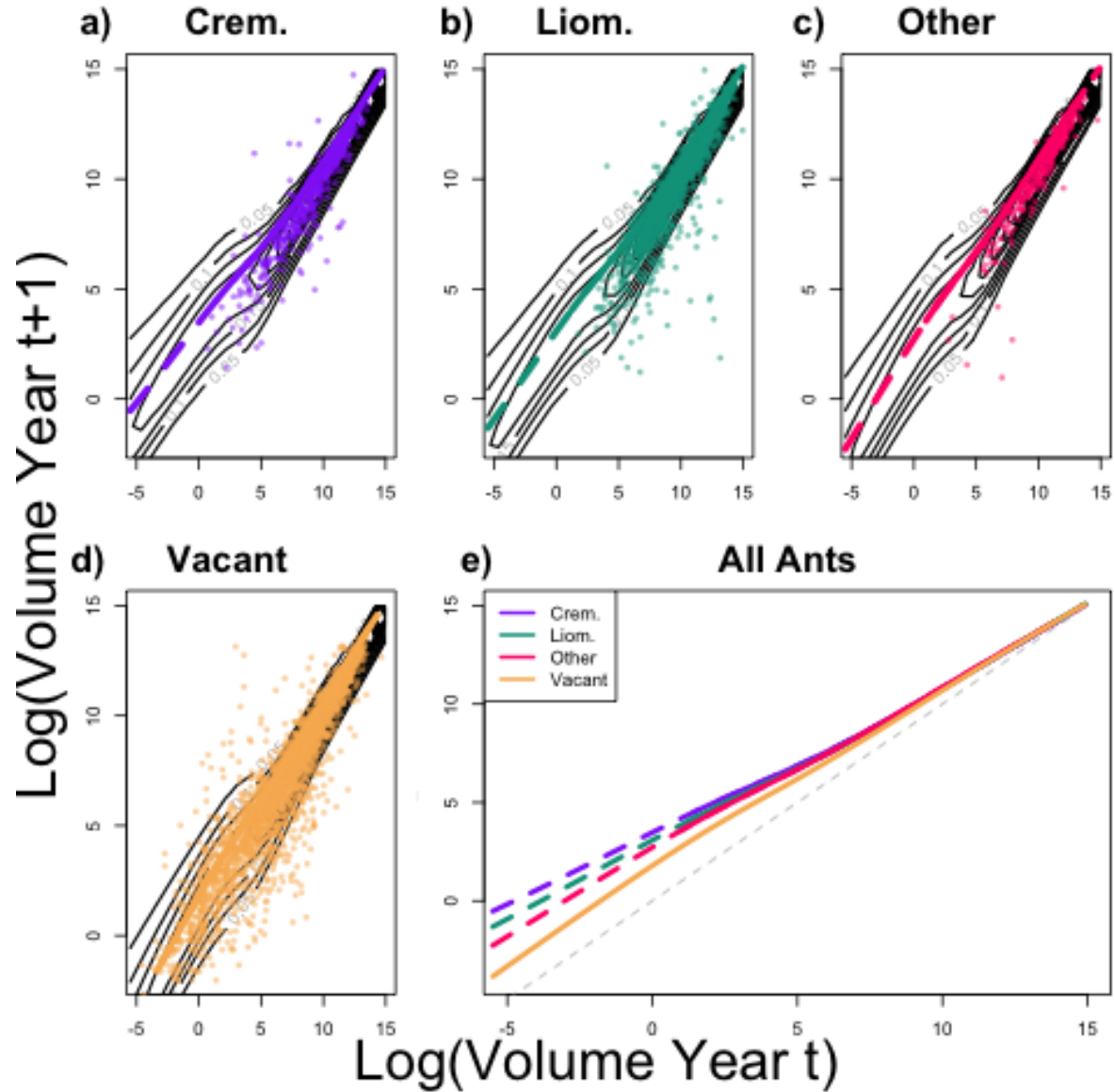


Figure 1: This figure shows the next predicted size of cholla based on previous size with each individual ant partner. The solid colored lines (seen in all panels) are the next mean predicted size of cholla. The points (seen in panels a-d) are the observed data which informs these estimates. The black countour lines (seen in a-d) appear at 5% increments showing where 5%, 10%, etc. of the data is expected to fall. They grey dashed line (in panel e only) shows the line where the next predicted size is the same as the previous (aka there is no growth on this line and below this line is shrinkage).

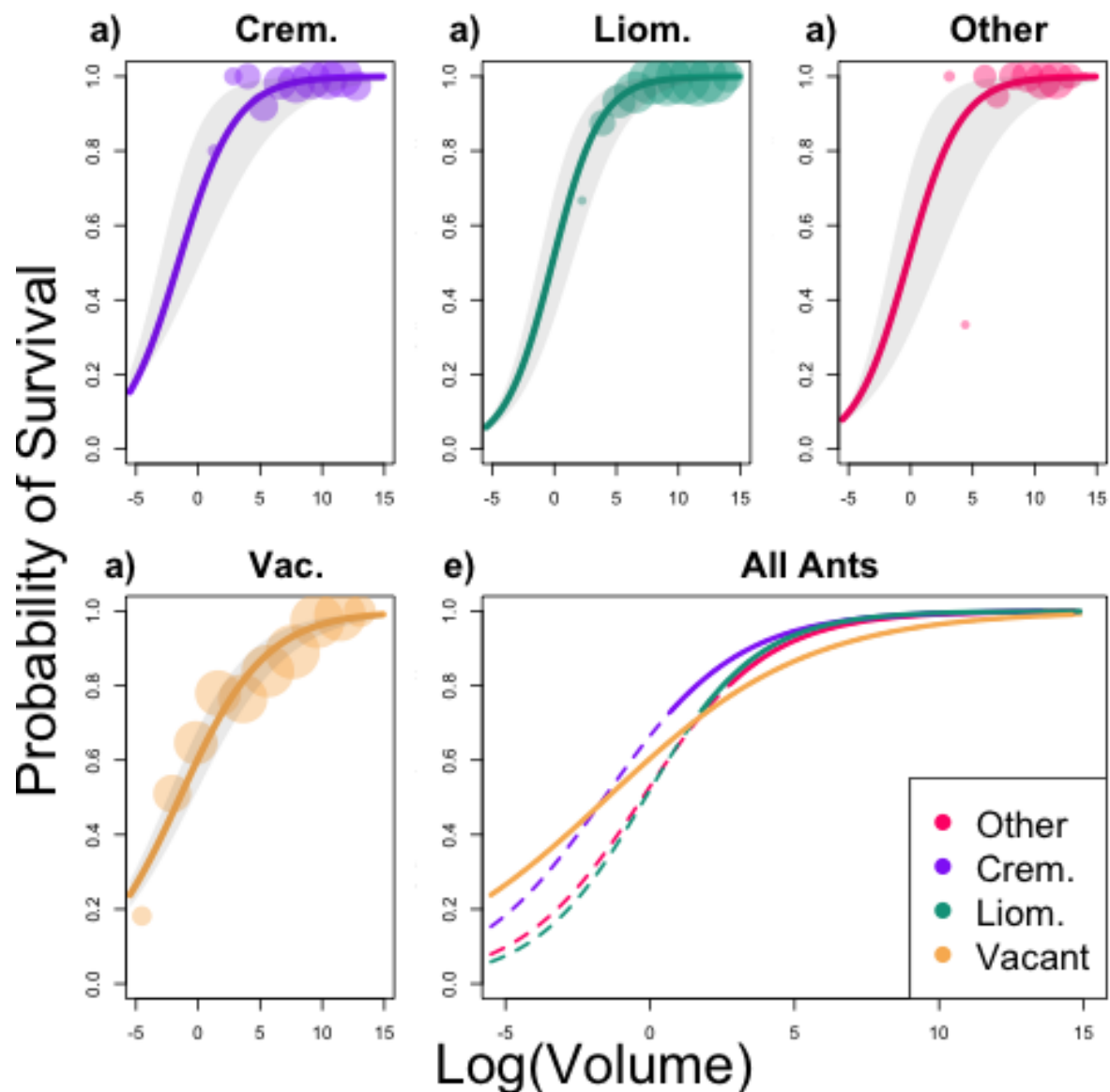


Figure 2: This figure shows the estimated survival rates based on the size of the cactus with each individual ant partner. The solid colored lines (shown on all panels) indicate the mean estimated survival rates. The dashed lines (shown in panel e) indicate extrapolations beyond existing data (where we estimated survival for plants tended by ants where we had never seen a tended cactus of that size). The grey area around the solid lines (shown in panels a-d) show the 90% confidence interval for the estimates. The colored dots are the real data binned by size to show how our estimates align with real survival observations. A larger circle means we had more data on survival of plants of this size with this partner.

435 We found evidence that ant visitation by *L. apiculatum* ants leads to increased floral viability
rates and that ant identity can influence the strength of viability benefits. We observed mean
viability rates of flowers between 40% and 92% (Figure 3). Ant partners influence the mean
438 viability rate of flowers, with *L. apiculatum*-tended plants experiencing the highest mean viability
rate (86%), followed by vacant and Other tended plants (at 60-61%), and *C. opuntiae* tended plants
had the lowest floral viability rate (57%).

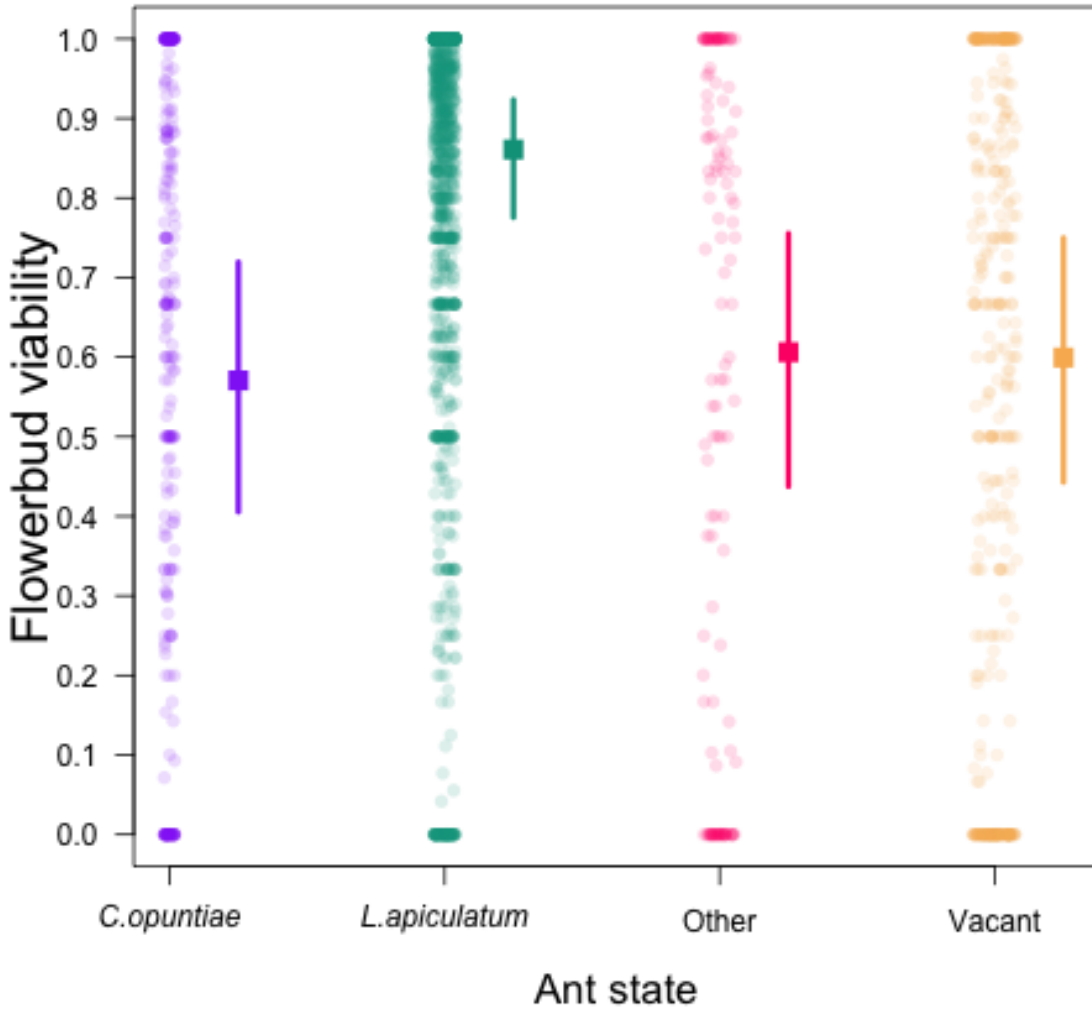


Figure 3: This figure shows the estimated distributions of floral viability rates compared to observed distributions of floral viability rates of cholla based on ant partner identity. The solid lines indicate the estimated viability distribution. The colored histograms represent the observed viability rates of plants with that partner.

441 We analyzed all other vital rates mentioned in the methods as well. The results of each of these as well as the posterior predictive checks are all included in the Appendix : Figures 1 – 4.

We analyzed the correlation coefficients of all models which included ant state as a predictor and found that the annually varying effects of each ant on growth were the most correlated (correlation coefficients averaged 0.63), and that the effects on survival were the least correlated (correlation coefficients averaged 0.36). High correlation indicates synchronicity, which is necessary for portfolio effect to occur because a central thesis of this effect is that partners must react differently to temporal environmental stochasticity. For the growth model, annually varying random effects of Other ants were the least synchronous (correlation coefficients ranging from 0.28 to 0.39) while the effects of vacancy were the most synchronous (correlation coefficients ranging from 0.39 to 0.81). For the viability model, annually varying random effects of *L. apiculatum* ants were the least synchronous (correlation coefficients ranging from 0.06 to 0.44) while the effects of vacancy were the most synchronous (correlation coefficients ranging from 0.44 to 0.62). For the survival model, annually varying random effects of *C. opuntiae* ants were the least synchronous (correlation coefficients ranging from 0 to 0.15) while the effects of vacancy were the most synchronous (correlation coefficients ranging from 0.13 to 2.6). This variety of synchronicity across ant states and vital rates indicates there is potential for portfolio effect as many of the ants effects revealed low synchronicity (low correlation of annual effects, indicating that ant identity impacts annual random effects), particularly in the survival model.

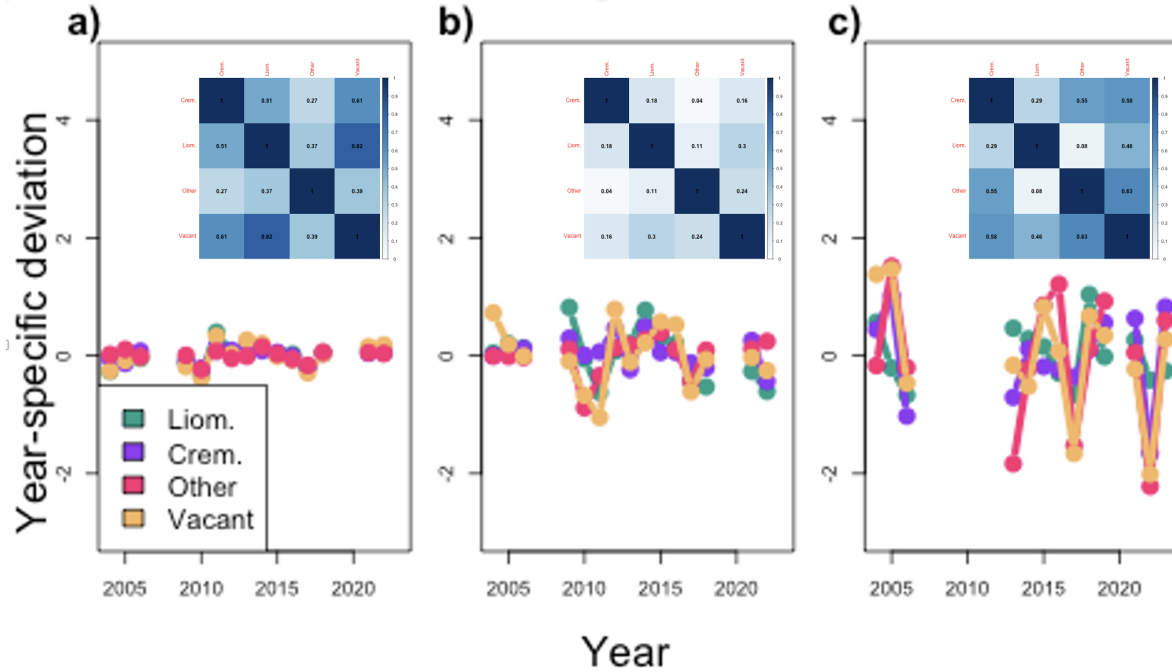


Figure 4: This figure shows the mean affect of each ant partner on a) the estimated next size, b) the estimated survival, and c) the floral viability of cacti across every year of our study. These values are estimated from the fitted random effects of ant and year in our models. Each point represents the mean of the random effect of the identified model, ant, and year (e.g. the lowest dot in panel b) represents the mean effect of vacancy on survival rates in year 2011).

4

What are the frequency and direction of partner turnover across the plant life cycle?

462

We found that 55% of individual plants surveyed in the long-term data experienced at least one ant state transition on average, with very distinct size-dependence and directional patterns (Figure 5).

465 Vacancy is the most likely ant state of small plants ($\leq 10 \log(\text{cm}^3)$). Even when small plants are ant-tended at the start of the transition year, they are most likely to transition back to vacancy

⁴*Honestly not sure if I should include an image for this one or just report some values? We should discuss.*

(Figure 5b-d). The probability of becoming ant-tended increases with size, though it is not equally
468 likely to be tended by all partners. For large plants that are initially vacant or tended by *L.*
apiculatum or Other ants, *L. apiculatum* is the most likely next partner, suggesting that this
partner species is able to colonize plants that were previously vacant or occupied by Other ants,
471 and effectively retain plants that it previously occupied. *C. opuntiae* were also able to retain plants
they previously occupied, but not as well as *L. apiculatum*: for plants that begin the transition
year with *C. opuntiae*, the probability that those plants remain occupied by *C. opuntiae* at the end
474 of the transition year is only slightly greater than the probability of take-over by *L. apiculatum*,
while take-over in the other direction is extremely rare. It is also notable that transitions away
from the initial state of *L. apiculatum* were almost always transitions to vacancy (Figure 5d), while
477 transitions away from the initial states of *C. opuntiae* and Other were often transitions to other
ants. This suggests a competitive hierarchy whereby *L. apiculatum* may abandon low-value plants
with little nectar production but is almost never displaced from high-value plants.

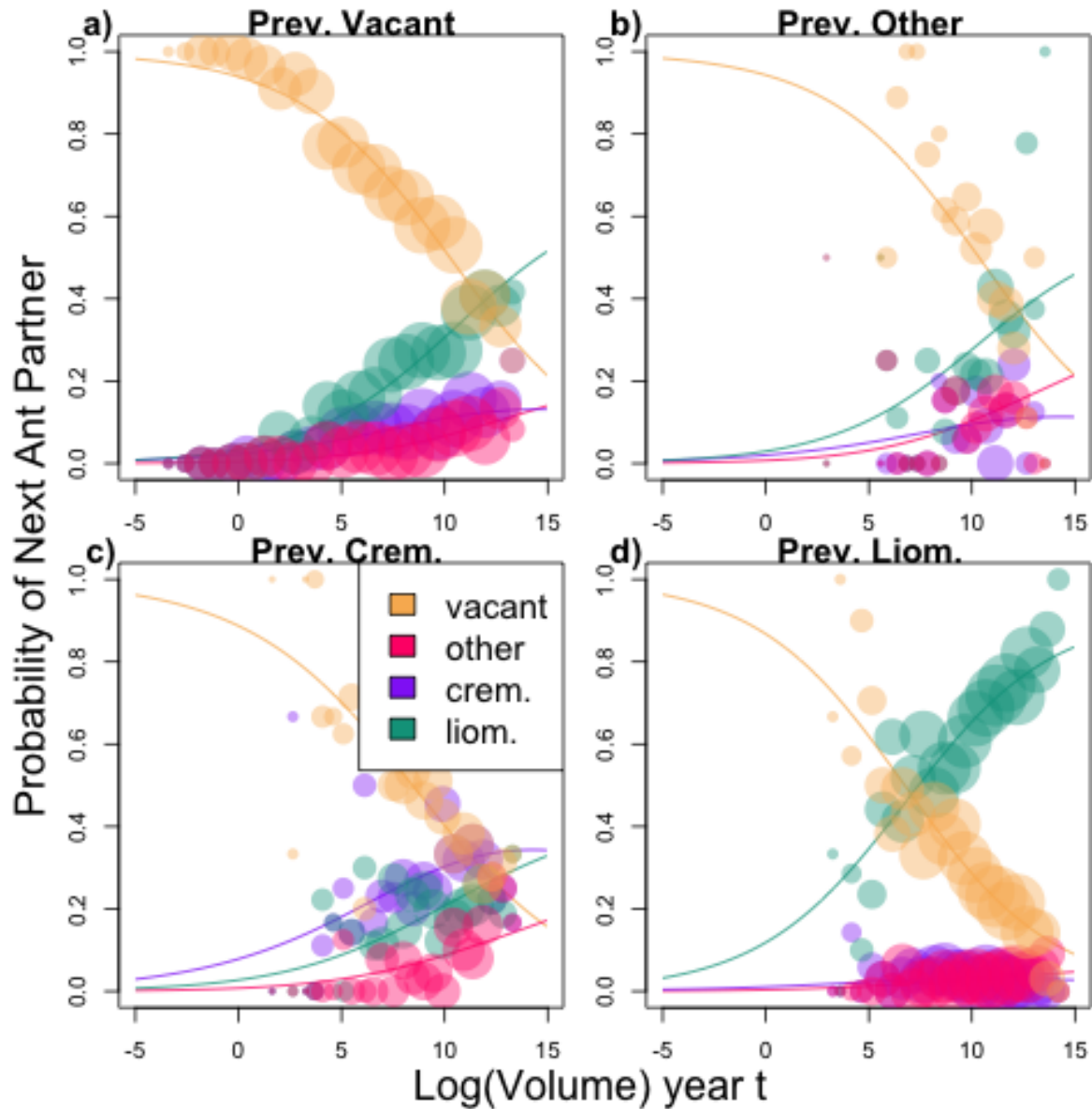


Figure 5: This figure shows the probability of being tended by each ant partner or vacant based on the size of the plant. Each panel shows these probabilities for a different previous ant state. The solid lines represent the mean probability of being tended by a specific partner. The colored points are the real data binned by size to show how our estimates align with real visitation observations. A larger circle means we had more data on visitation of plants of this size with this previous partner.

480 *What is the net effect of partner diversity on plant fitness, and what*
mechanism(s) explain(s) this effect?

By integrating vital rate results and ant transition dynamics into the multi-state IPM we can
483 evaluate the fitness implications of different scenarios of partner diversity and identity. First, there
was strong evidence that ant visitation had mutualistic fitness effects on plant partners. The
lowest mean fitness was $\lambda_{NS,Vacant}$, the fitness of the cholla with no partners (Figure 6b). Across
486 all 1+ partner scenarios, we found that the posterior $\lambda_{NS,Occupied}$ distributions were greater than
 $\lambda_{NS,Vacant}$ distributions 100% of the time. This indicates that ant visitation elevates fitness no
matter the identity or number of partners. Furthermore, we find no benefits of partner diversity
489 here, with the fitness of the cholla with 1, 2, and 3 partners roughly equivalent (Figure 6 a). From
the means alone it appears that there is a reduction in fitness when all ants are present compared
to 1 or 2 partners. The contrasts of λ_{NS} posterior distributions of 0 - 2 partner diversity scenarios
492 to full diversity show that the overlap ranges from 58% - 100%, indicating high similarity between
the full partner diversity scenario and others. This is highlighted in Figure 6 panel b.

We found that the difference in posterior distributions of any ant scenario ($\lambda_{NS,A} - \lambda_{NS,B}$,
495 where A and B are arbitrary ant diversity scenarios that are not vacancy) were between 31% and
83% different. The posterior distributions of λ for each diversity scenario are visualized in Figure
6 b. This high level of overlap between posterior distributions indicates that while each of these
498 partners are beneficial, there does not appear to be a significant benefit of partner diversity within
this system.

The lack of diversity benefits are not driven by the high overall frequency of *L. apiculatum*.
501 Using the simulations where all ants had equal frequencies across sizes (further explained and
analyzed in Appendix C), we found the same fitness patterns as in the competitive exclusion model
described above. Equal probability for transitioning into any ant state meant that the numbers
504 of *C. opuntiae* and Other ants were boosted significantly. Despite this, the fitness of scenarios
including these less frequent ants were not increased meaningfully.

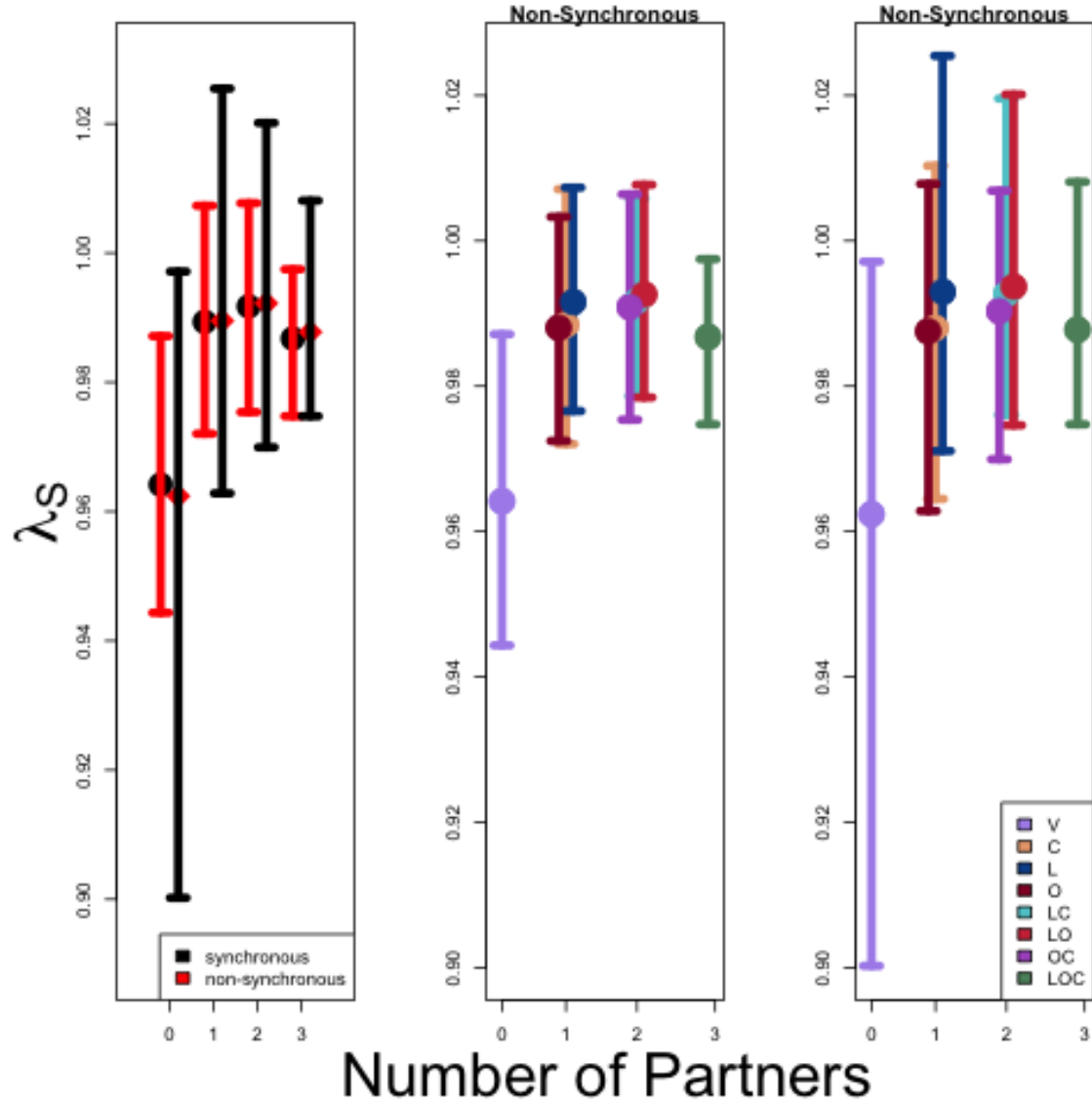


Figure 6: Panel a) shows the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). Panels b-c) shows the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

We found no evidence of portfolio effect. The effect of all ant partners can be measured as $\lambda_{All} - \lambda_{Vacant}$ (Figure 7). We are 98% confident that when all ants are present the cholla experience higher fitness than when no ants are present according to both the synchronized (S) and non-synchronized (NS) model scenarios. When subtracting these two resulting vectors from each other $((\lambda_{NS,All} - \lambda_{NS,Vacant}) - (\lambda_{S,All} - \lambda_{S,Vacant}))$, we found that there is no real difference between the two scenarios, meaning we have no evidence of portfolio effect.

Discussion

Mutualisms commonly involve multiple partners but the ecological consequences of partner diversity remain poorly understood. Here we show that while alternative partners may be ecologically different, their fitness effects on a shared mutualist can be effectively equivalent and interchangeable. The results of our hierarchical models revealed that different ant partners had different effects on vital rates, with *C. opuntiae* tended plants **experiencing slight advantages in growth**⁵ and survival when small, and *L. apiculatum*-tended plants experiencing floral viability advantages. The results of our synchronous and non-synchronous stochastic IPM revealed that all diversity scenarios which included any partners resulted in the highest possible fitness for tree cholla, suggesting that while ant visitation is beneficial, partner identity and diversity are inconsequential in this system. These results highlight that partners can differ meaningfully while still resulting in the same fitness benefits for a focal mutualist.

Previous studies have reported complementarity (Palmer et al., 2010; Rez, 2007; Afk, 2021) while we found that there are no benefits of diversity in our system. Vital rate results indicated that different ant partners affected different vital rates uniquely (Figures 1,2,3). On their own, these results suggested a likely benefit of diversity through complementarity, such that each ant partner offered a unique form of herbivore protection across ontogeny and or across different life history processes. However, our multi-state IPM results, which synthesize all life history processes

⁵ *I don't think this is true if you look at my v1 figure.*

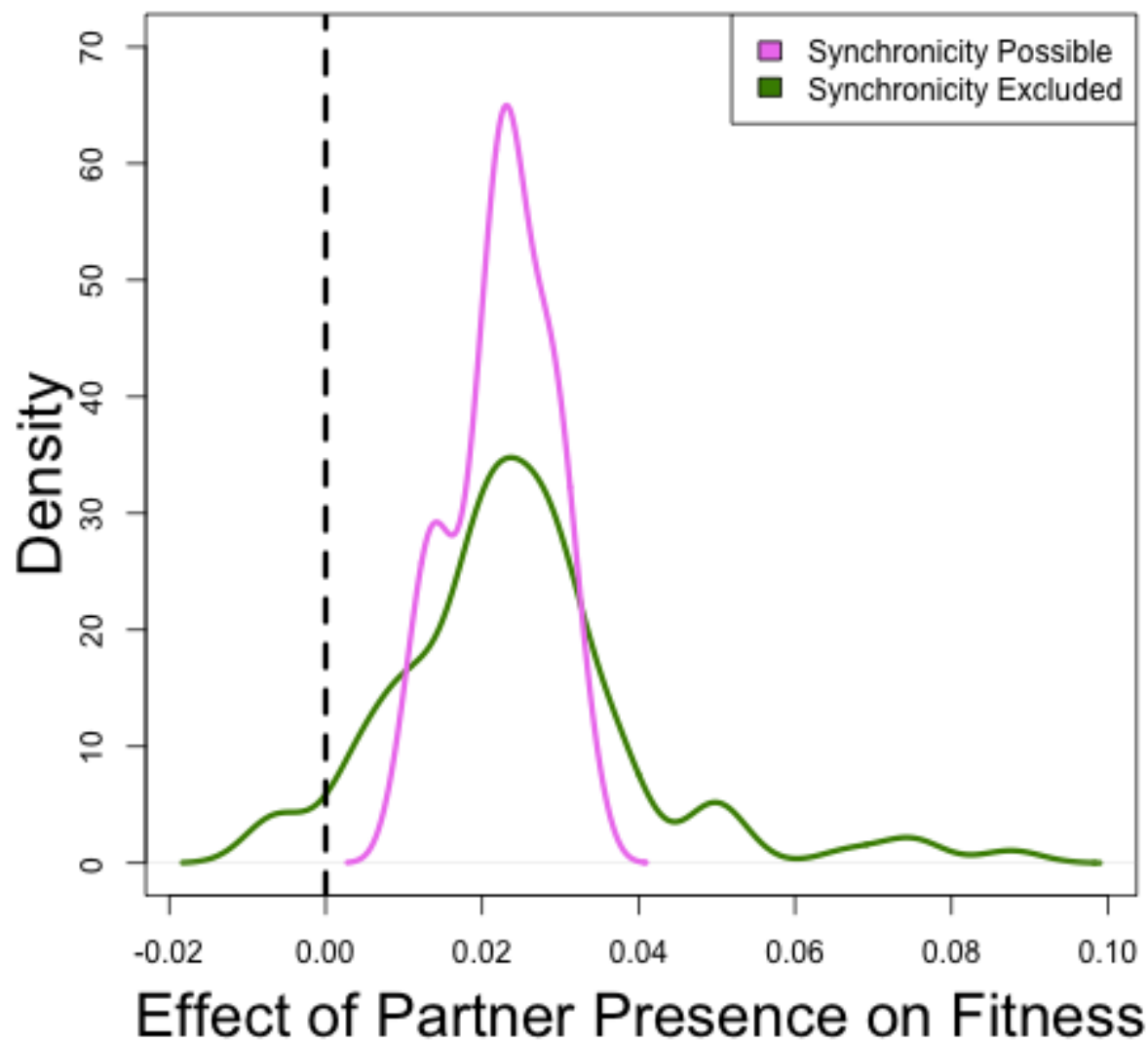


Figure 7: This figure shows the distribution of $\lambda_{NS,All} - \lambda_{NS,Vacant}$ in pink and $\lambda_{S,All} - \lambda_{S,Vacant}$ in green. The vertical dashed line shows where the effect of partners on the fitness of the cholla is 0 (to the left the partners have a negative effect, to the right the partners are beneficial).

and accounts for partner identity and turnover, revealed an overall positive effect of ant visitation
531 but no apparent benefit of diversity, raising questions about how to resolve the tension between
these two sets of results.

We speculate that our results are driven by the vital rate sensitivity structure of this population.
534 Like other long-lived, iteroparous species (Franco and Silvertown, 2004), tree cholla fitness is most
sensitive to the growth and survival of established individuals (Miller et al., 2009; Elderd and
Miller, 2016), which are virtually guaranteed many years of reproductive opportunities once they
537 reach a size that is protected from mortality. Differences between alternative partners were most
pronounced either in reproductive rates (Figure 3), which contribute relatively weakly to fitness
(because a long reproductive lifespan overrides individual reproductive bouts), or in survival at small
540 sizes, where mortality risk was high (Figure 2) and the probability of ant tending was low (Figure
5). At larger sizes with low mortality risk, ant-tended plants had a modest survival advantage over
vacant plants regardless of ant identity, and that result dominates integrated measures of fitness
543 because of the high sensitivity of established plant survival. The very real demographic differences
between partner species in other vital rates and at other sizes do not register nearly as strongly in
the currency of fitness.

546 Our work explicitly incorporated temporal environmental stochasticity, which raises the oppor-
tunity for portfolio effect as a mechanism of diversity benefits. Yet, we find no evidence of portfolio
effect within our system. When partners exhibit different reactions to varying environments, in-
549 teracting with multiple partners can lead to more consistent benefits across time (Batstone, 2018).
We did not directly measure the reactions of ant partners to temporal environmental stochasticity,
rather we measure how the effects of each partner on cholla demography fluctuate through time,
552 and whether those fluctuations are correlated. The vital rate results led us to believe that portfolio
effect may be at play due to the differences in ant responses to annual variation. However, there is
no stabilization of fitness across temporal heterogeneity and therefore no benefits of diversity due
555 to the portfolio effect. Other portfolio effect studies (Lázaro et al., 2022; Tornos-Estupiña et al.,
2023) found that asynchrony was linked to weak evidence of portfolio effect. One study Dallas and

Kramer (2022) found that while portfolio effect was easy to show in theoretical models, it is often
558 very weak or nonexistent in empirical data across many systems. This indicates that it may be
very difficult to detect, disguised by different mechanisms, or uncommon in nature.

Partner turnover is likely a significant driver behind the fitness we see within Cholla populations,
561 however the processes which drive the actual turnover frequency and directions remain a mystery in
this system. In the literature, it is clear that the frequency of partner turnover can have big effects
on the fitness of the focal mutualist Fiala et al. (1994); Horvitz and Schemske (1986); Oliveira et al.
564 (1999); Sachs et al. (2004). The direction of partner turnover is also important when the identity of
partners impacts the quality of benefits recieved Fonseca et al. (2003); Alonso (1998); Dejean et al.
(2008); Noe and Hammerstein (1994). In our system we found that there are distinct patterns to the
567 directions and frequency of partner turnover, but we still don't understand the mechanisms driving
these turnovers. In the model presented in this paper we assumed that competitive exclusion was
the driving mechanism behind the patterns we see. This assumption would indicate that ants are
570 competing for each individual cacti, likely based on EFN quality and composition (Heil et al., 2004,
2010), and that the more common an ant is the better a competitor it is. In our simulations this
means that when one ant is removed the ants which are left occupy the newly vacated cacti so that
573 the proportion of vacant cacti never changes. We included this assumption because we believe it
to be the most biologically reasonable based on the personal observations of multiple researchers.
However, it is possible this assumption is incorrect, so we created two alternative simulations to
576 determine how robust our results were to various assumptions about partner turnover drivers. The
first alternative we tested is a model that assumes all ants have the same competitive ability rather
than the differences in competitive ability we assume cause observed variety in ant frequency. In
579 this hypothetical simulation all ants have equal probability of occupying any given cactus, but the
level of vacancy across the population remains consistent with observed data. This simulation was
designed to test if the overwhelming frequency of *L. apiculatum* may be overriding any potential
582 advantages of other partners. The results of this simulation (Appendix : Figure 7) were effectively
the same as the competitive exclusion model presented in the body of this paper (Figure: 6).

The second alternative we tested was a model that assumed rather than competition driving ant occupations, it was the population size of each ant species alone. In this hypothetical simulation when one ant species is removed from the model the cacti are left vacant. We believe this is an unlikely scenario as there appear to be vast populations of the ant species in the area.⁶ This simulation was designed to test if the fitness was more sensitive to changes in vacancy across the population than the frequency of any particular ant species. The results of this simulation revealed that when vacancy increased sampling effect is at play (Appendix : Figure 6). This result highlights that the primary driver behind the results we presented are ant presence and that it is only in cases oh unrealistically high vacancy that there are benefits to partner diversity. Further information about both of these models can be found in Appendix C. Together, these results indicate that our assertions that the fitness of the cacti are more sensitive to changes in vacancy than partner identity are robust across multiple assumptions.

This paper shows the importance of long-term datasets in investigating species interactions and calls for further use of long-term data. Previously studies have analyzed how partner identity and partner turnover impact focal mutualist fitness Fonseca et al. (2003); Dejean et al. (2008); Noe and Hammerstein (1994); Barrett et al. (2015); Bruna et al. (2014); Trøjelsgaard et al. (2015). Separate studies have analyzed how inter-annual variability impacts focal mutualists Alonso (1998); Alarcón et al. (2008); Ollerton et al. (2006); Horvitz and Schemske (1990); Lázaro et al. (2022). The long term data set we used gave us the unique ability to consider the combined effects of partner identity, partner turnover, and temporal stochasticity. By piecing together complete life cycle information from long-term data, we gain a more nuanced understanding of the fitness consequences of specific demographic effects. For example, our previous study suggested that *C. opuntiae* has overall parasitic fitness effects because activity of this species within tree cholla flowers can deter pollinators and reduce seed set (Ohm and Miller, 2014). Yet, the more complete analysis presented here, which accounts for reduced seed set alongside other demographic advantages, indicates that this species is clearly a mutualist and nearly as strong a mutualist as *L. apiculatum*.

⁶ *This is just personal observation so it might be best to remove this?*

Herbivory is an important driver of the fitness of cholla in this system, as herbivores directly impact the growth, survival, and reproductive efforts of the cacti (Miller et al., 2009). We have not explicitly incorporated herbivory but rather assume that herbivory can be captured in the effects they have on growth, survival, and reproductive effort observations. Analysis of our observational herbivore data shows that vacant plants experience elevated levels of herbivory compared to tended ones (Appendix : Figure 5). In the future, further studies on the direct impacts of herbivory would bolster the results reported here and expand our knowledge within this system.

As with any study, there are limitations to consider when interpreting the results reported here. These results highlighted in the paper are based on observational data regarding ant effects on plant demography rather than experimental data, meaning we are able to determine correlations but not causation. However, we have previously conducted experimental manipulations which revealed that ant presence has direct impacts on plant demography in our system (Miller, 2007). This combination of observational data backed up by experimental results gives us greater confidence in our causal interpretations than if we had only observational data. Specifically, we are confident in claiming that ant presence causes increased performance in cactus vital rates and the fitness signals we are detecting are directly related to ant partner presence. One further expansion could include nectar analysis in conjunction with ant interactions to look not only at the ant partner impacts on the plant demography, but also at the ability of plants to attract specific partners through nectar composition shifts.

This study highlights that while partners within a mutualistic guild can be ecologically different, they may still be interchangeable to a focal mutualist. The individual estimation of each vital rate and explicit inclusion of ant partners in this IPM allowed us to determine that partner diversity is not beneficial in this system and may highlight circumstances under which we would expect similar results in other systems. The individual vital rate estimation approach allowed us to determine that while different ant partners did impact each vital rate differently, the largest differences were estimated to occur at sizes when ants would not be present in real life or in vital rates that do not have a significant impact on the overall fitness. Considering partner impacts on vital rates across the

entire life cycle of the plant allows more power in determining the importance of these differences. Based on our results we would expect to see that partner diversity does not matter when partner impacts occur during low sensitivity life stages.

Acknowledgments

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Data and Code Availability

The data that support the findings of this study are openly available as a data package on the Environmental Data Initiative website, package id: knb-lter-sev.323.1

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Tables

Figure legends

Appendix A: Additional Methods and Parameters

In addition to the models described in the body of the paper, we fit several other models using data from previous studies. These models are described below.

Seeds Per Fruit. With data from Miller et al. (2006), we fit a model for the number of seeds produced by every fruit on a cholla ($\kappa(a')$) in year $t + 1$ based on the ant partner a' in year $t + 1$. We fit this model to seed data y^κ using a Negative Binomial distribution and the log link function:

$$y^\kappa \sim \text{NegativeBinomial}(\hat{\kappa}, \hat{\phi})$$

$$\hat{\phi} = \beta_0^\phi$$

The data used for this model did not include data on ants in the “other” category, so we used the data from vacant plants to parameterize seeds per flower for plants with “other” ants in the IPM.

We found that vacant plants produced the most mean seeds (147.2 per fruit), followed by *L. apiculatum* tended plants (142.4 per fruit), and finally, *C. opuntiae* tended plants (115.0 per fruit) (Figure 1).

Recruit Size Distribution. We fit this model to recruit size data y^η using a Normal distribution with the identity link function:

$$y^\eta \sim N(\hat{\eta}, \hat{\sigma})$$

where $\hat{\sigma}$ is estimated with a non-informative prior.

We found that the mean size of recruits is $\log(-2.097)m^3$ with an interquartile range from $\log(-2.173)m^3$ to $\log(-1.712)m^3$ (Figure 2).

Parameter	Median (95% CI)	Prior Distribution
growth xi intercept vacant β_{01}^g	-5.210899(-5.686865, -5.491787)	sDE
growth xi intercept other β_{02}^g	-5.8288(-5.956217, 1.766021)	asdf
growth xi intercept <i>C. opuntiae</i> β_{03}^g	-4.529523(-6.0770390, 0.1222112)	asdf
growth xi intercept <i>L. apiculatum</i> β_{04}^g	-5.106802(-5.4499944, 0.5453901)	asdf
growth xi size dependent vacant β_{11}^g	asdf	asdf
growth xi size dependent other β_{12}^g	asdf	asdf
growth xi size dependent <i>C. opuntiae</i> β_{13}^g	asdf	asdf
growth xi size dependent <i>L. apiculatum</i> β_{14}^g	sadf	asdf
growth omega intercept ω_0^g		
growth omega size dependent ω_1^g		
growth alpha intercept α_0^g		
growth alpha size dependent α_1^g		
1-year germination intercept α^{γ_1}		
2-year germination intercept α^{γ_2}		
survival intercept vacant β_{01}^s		
survival intercept other β_{02}^s		
survival intercept <i>C.opuntiae</i> β_{03}^s		
survival intercept <i>L. apiculatum</i> β_{04}^s		
survival size dependent vacant β_{11}^s		
survival size dependent other β_{12}^s		
survival size dependent <i>C. opuntiae</i> β_{13}^s		
survival size dependent <i>L. apiculatum</i> β_{14}^s		
flowering intercept β_0^f		
flowering size dependent β_1^f		
viability intercept vacant β_{01}^v		
viability intercept other β_{02}^v		
viability intercept <i>C. opuntiae</i> β_{03}^v		
viability intercept <i>L. apiculatum</i> β_{04}^v	46	

Table 1: This table includes the median estimates, the 95% confidence intervals, and the prior distribution for each parameter in each model.

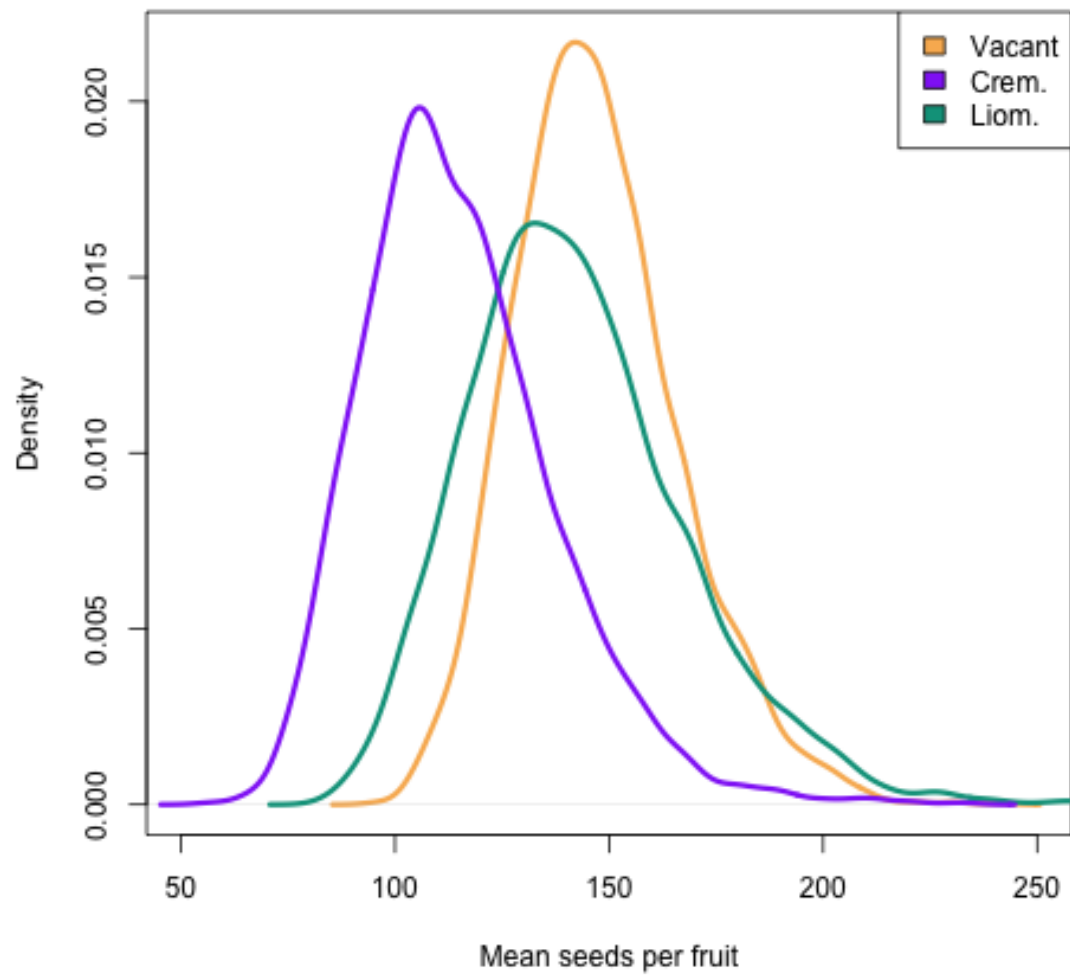


Figure 1: Shows the distribution of seeds per fruit on plants tended by *C. opuntiae* (purple), *L. apiculatum* (Teal), and Vacant plants (Yellow).

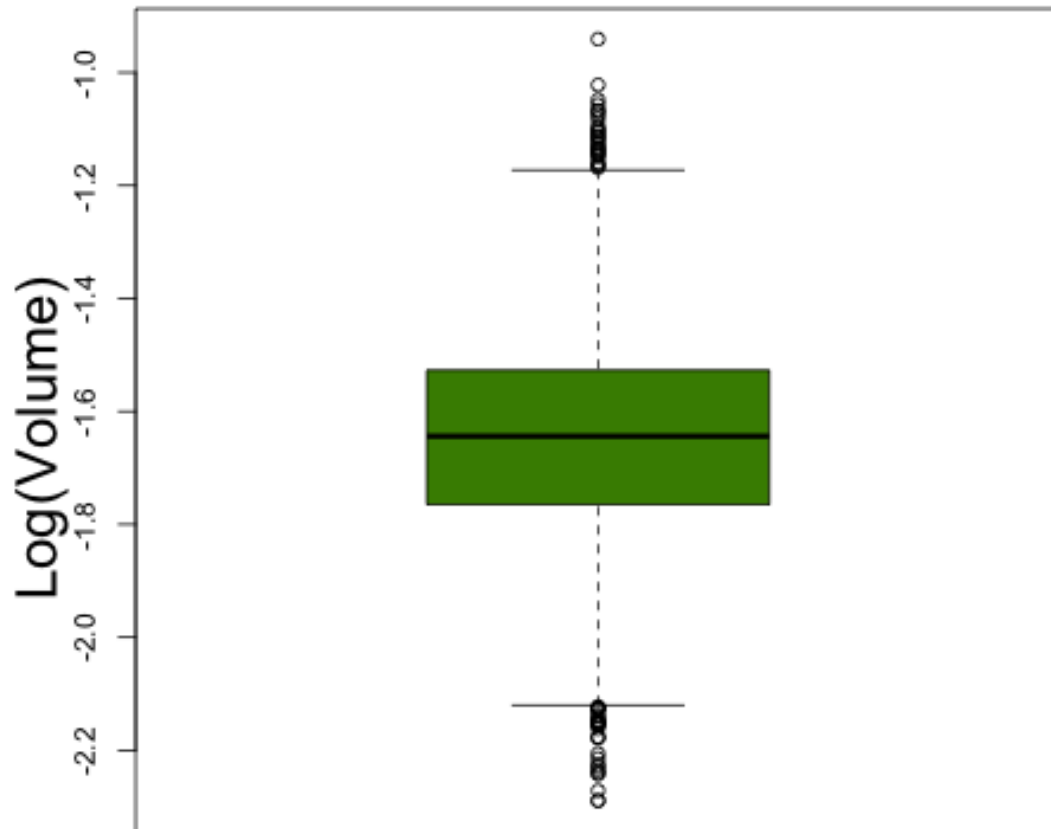


Figure 2: Shows the distribution of recruit sizes. The mean size of a recruit is marked by the dark line in the green box, the green box shows the interquartile range of recruit sizes, and the dashed lines show the minimum and maximum of estimated recruit sizes. The points beyond these dashed lines are considered outliers.

Germination. With germination data Miller (2007), we fit two models for the probability of germinating from the first year seedbank (γ_1) or the second year seedbank (γ_2) in year $t + 1$, with no fixed or random effects. These models were fit to germination data $y^{\gamma_1}, y^{\gamma_2}$ using the binomial distribution with logit link functions:

$$y^{\gamma_1} \sim \text{Binomial}(\hat{\gamma}_1)$$

$$y^{\gamma_2} \sim \text{Binomial}(\hat{\gamma}_2)$$

We found that the mean germination rates for seeds in the seedbank for one year is 0%, with
852 an interquartile range of 0% and 1%. We found that the mean germination rates for seeds in the seedbank for a second year is 0%, with an interquartile range of 0% to 0.4% (Figure 3). Seeds are more likely to germinate in their first year in the seedbank, but most seeds will never germinate.

Pre-Census Survival. With recruit census data Miller et al. (2006), we fit a model for the probability of a seedling (which germinates in early Fall) surviving to when we census in May (δ) of year $t + 1$ (accounting for missed mortality events), with fixed effects of the previous size x and random effects of the transect m . We fit this model to pre-census survival data y^δ using a Bernoulli distribution with a logit link function:

$$y^\delta \sim \text{Bern}(\hat{\delta})$$

855 where $m \sim N(0, \sigma_{\text{transect}}^2)$ is the random effect of transect where the recruited individual was analyzed for survival.

We found that plants have a 16.2% probability of surviving from germination to the next census.
858 Our model estimated this very well, expecting a 16.3% probability (Figure 4).

Appendix B: Observed Herbivory Data

Herbivory is an important driver in this population and shapes the range and demography of cholla.
861 Herbivore presence has been shown to negatively impact growth and fecundity of cholla populations (Miller et al., 2009). Ant visitors are believed to offer defensive benefits to the plants they tend

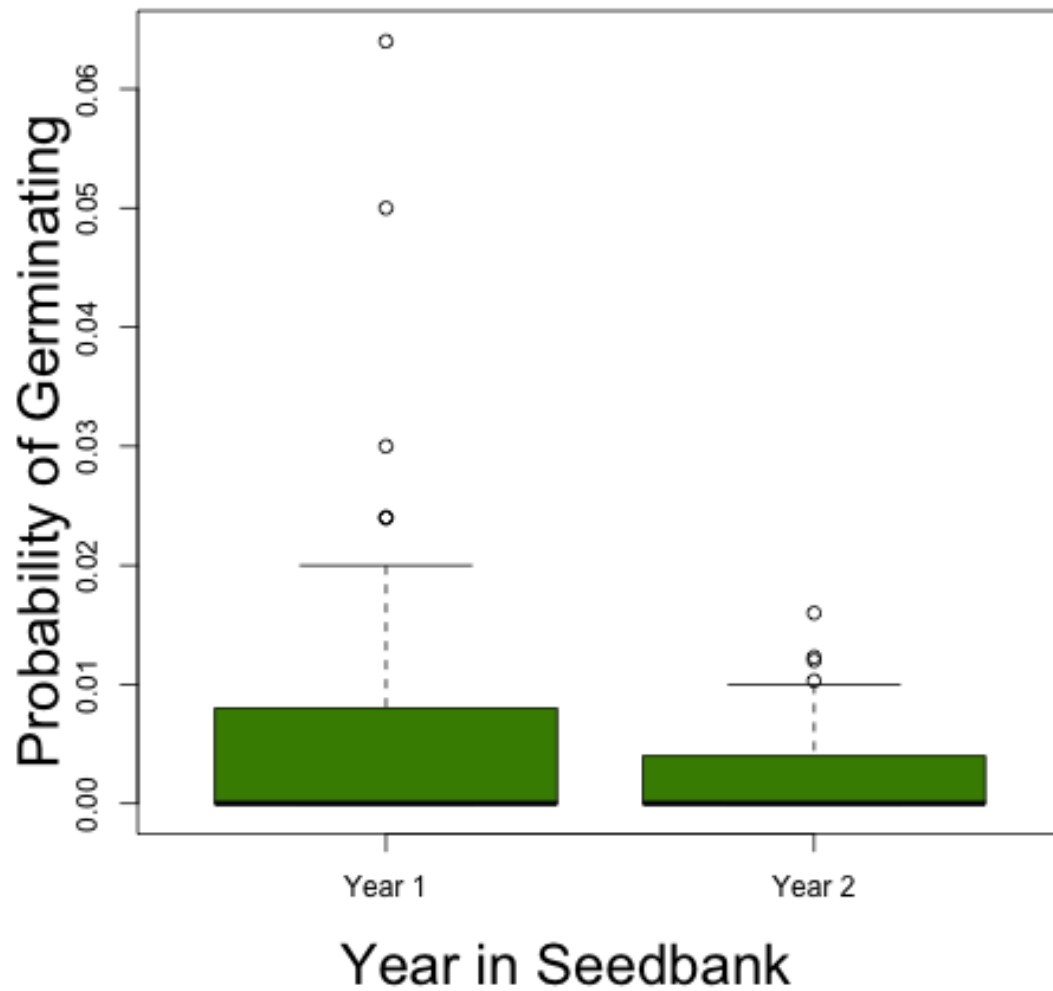


Figure 3: Shows the probability of a plant germinating during its first year in the seedbank (left) or during the second year in the seedbank (right). The mean germination rates are shown by the dark line near the bottom of the green box. The green boxes show the interquartile ranges of the germination rates for year 1 and 2 plants. The dashed line shows the maximum germination rates observed. The points above that show the outliers of our data.

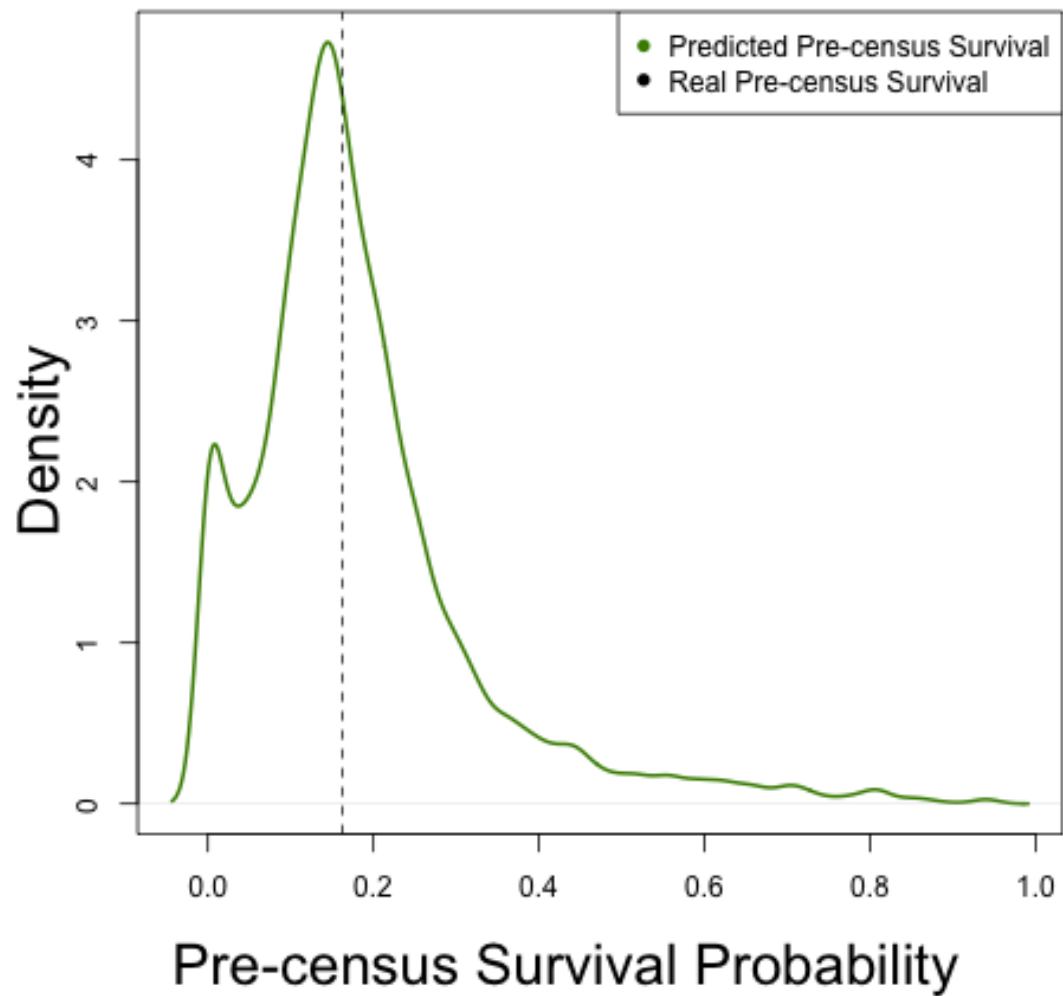


Figure 4: Shows the distribution of the probability of plants which germinated between censuses surviving to the next census to be counted. The green line shows the predicted pre-census survival and the grey dashed line shows the actual observed pre-census survival.

in this system, leading to the hypothesis that ant presence would be correlated with reduced herbivory. Herbivory data was collected during censuses any time herbivores were identified on a plant. This involved noting the type and quantity of herbivores observed. This data has been taken consistently since 2017, so the analysis below considers 6 years of data. We considered only plants which were reproducing, as they were likely to produce the highest quantity of EFN (Miller, 2014). The proportion of reproducing plants that experienced herbivory was calculated for each ant state separately. Analysis showed that ant presence is correlated with lower herbivore visitation. 40% of vacant cacti experienced herbivory. Plants tended by Other ants experienced similar, though lower, levels of herbivory on reproducing plants, with herbivores detected on 37.5% of plants. Herbivores were detected on 25% of plants tended by *C. opuntiae* ants and on 11% of plants tended by *L. apiculatum* ants. These results indicate that ant presence is correlated with lower levels of herbivory and that partner identity has an impact on the level of herbivory. They also indicate that the partner correlated with the lowest levels of herbivory is *L. apiculatum* ants. These findings are consistent with literature findings which show that *L. apiculatum* ants are the most aggressive (therefore the most effective against herbivores), but differ from previous findings that *C. opuntiae* may not offer defensive benefits (Miller, 2007).

Appendix C: Alternative Ant Transition Simulations

In addition to the competitive exclusion model defined and analyzed in the body of the paper, we simulated results from several other potential models. We chose to include competitive exclusion as our primary results in the paper because we believe it to be the most biologically realistic. However, in building and testing of alternative models we found that the method of ants occupying plants significantly impacts the fitness of the population. We tested two alternative transition models, one called the frequency based model and one called the equal likelihood model.

Frequency Based Model. The first alternative hypothesis we tested was what we called the frequency based model. In this model rather than the proportion of vacant cacti being maintained, the

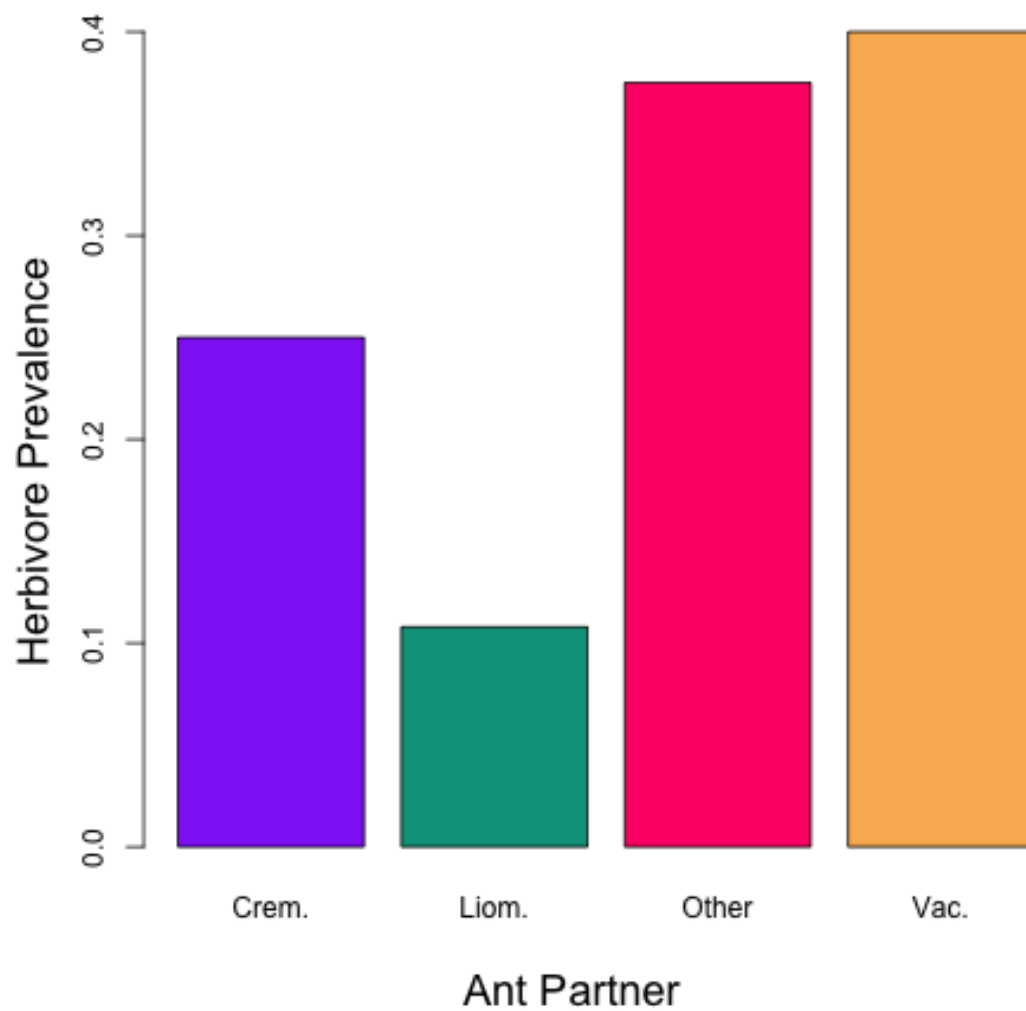


Figure 5: Shows the proportion of reproducing plants which are visited by herbivores. Each bar represents the subset of the cacti population in a different ant state.

proportion of cacti occupied by each species is maintained and when one is removed it is replaced with vacancy. This version of the model assumes that the frequency of each ant we see is reflective of the real frequency of populations rather than some other mechanism. With this model we found very clear evidence of Sampling Effect in the system. When only *C. opuntiae*, Other ants, or both ants are present, there is very little difference in the fitness of the cacti from when no partners are present. Only when *L. apiculatum* ants are present do we see an increase in the fitness of the focal mutualist (Figure ??a). In this simulation, the more partners that are present the higher the fitness of the focal mutualist is, confirming that partner diversity would be beneficial through sampling effect if this transition model were correct. (Figure ??b).

Equal Likelihood Model. The second alternative hypothesis we tested was what we called the equal likelihood model. In this model we preserved the observed pattern of size-dependent vacancy/occupancy, but occupancy was manipulated to be equally likely for all partner identities. This was designed to remove the effect overwhelming numbers of *L. apiculatum* ants may have. Despite very different proportions, we found very similar outcomes to the competitive exclusion model analyzed in the paper. All ants are beneficial, but having more than one is not necessarily any better than having an individual species as a partner (Figure ??b). Partner presence is beneficial, but neither identity nor number of partners appears to be important (Figure ??a).

Appendix D: Posterior Checks and Model Validation

For each model fitted, we conducted two tests to determine if the fit was acceptable to use in our IPM. First, we checked the convergence of each parameter. Below we show the convergence of all β terms listed in the Statistical Modeling subsection of Methods. Second, we checked the posterior fit, comparing the estimated values of each model to the y values of the actual data. We show these posterior checks below, split by ant partner where relevant.

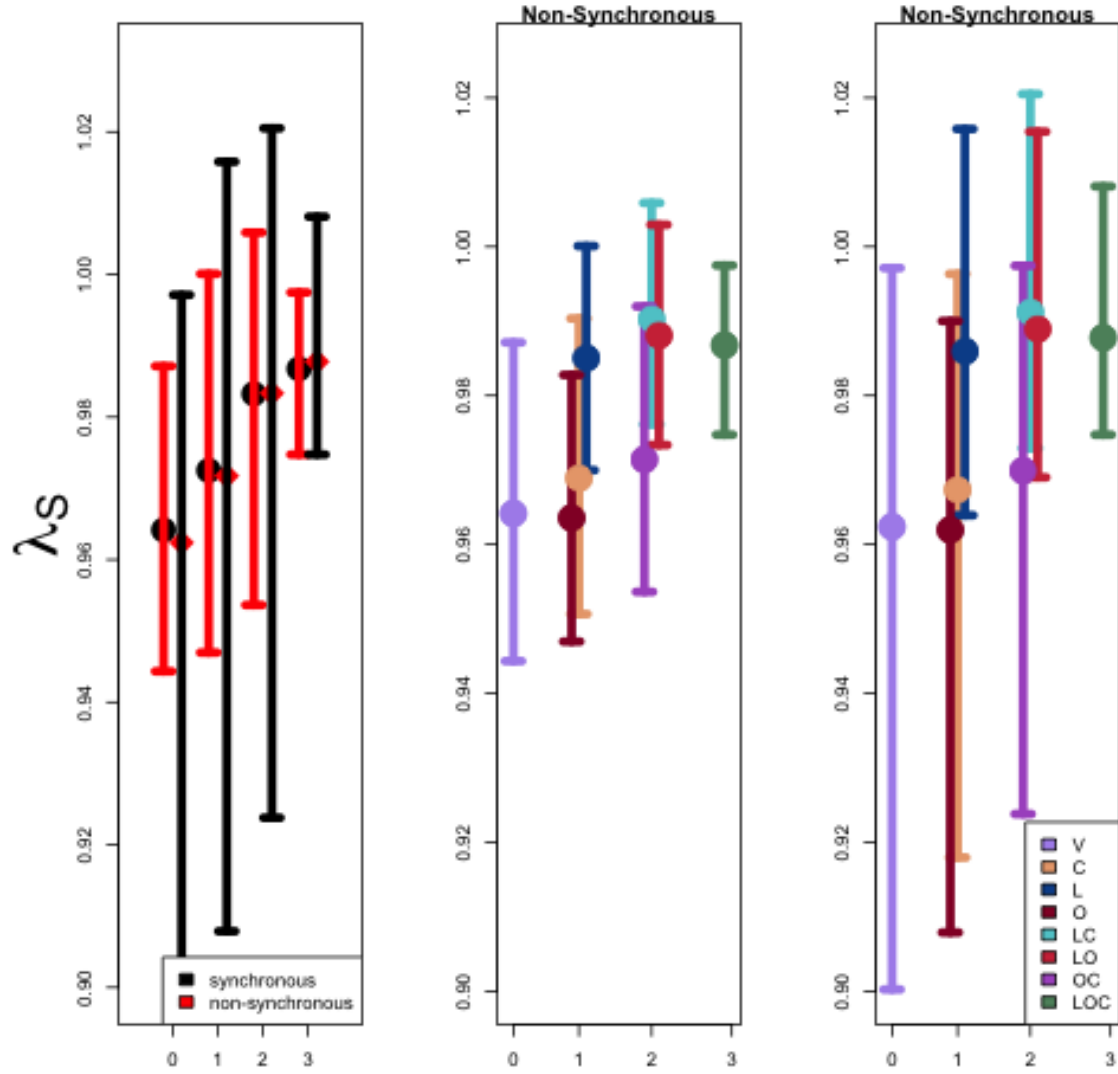


Figure 6: Panel a) shows the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). Panels b-c) shows the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

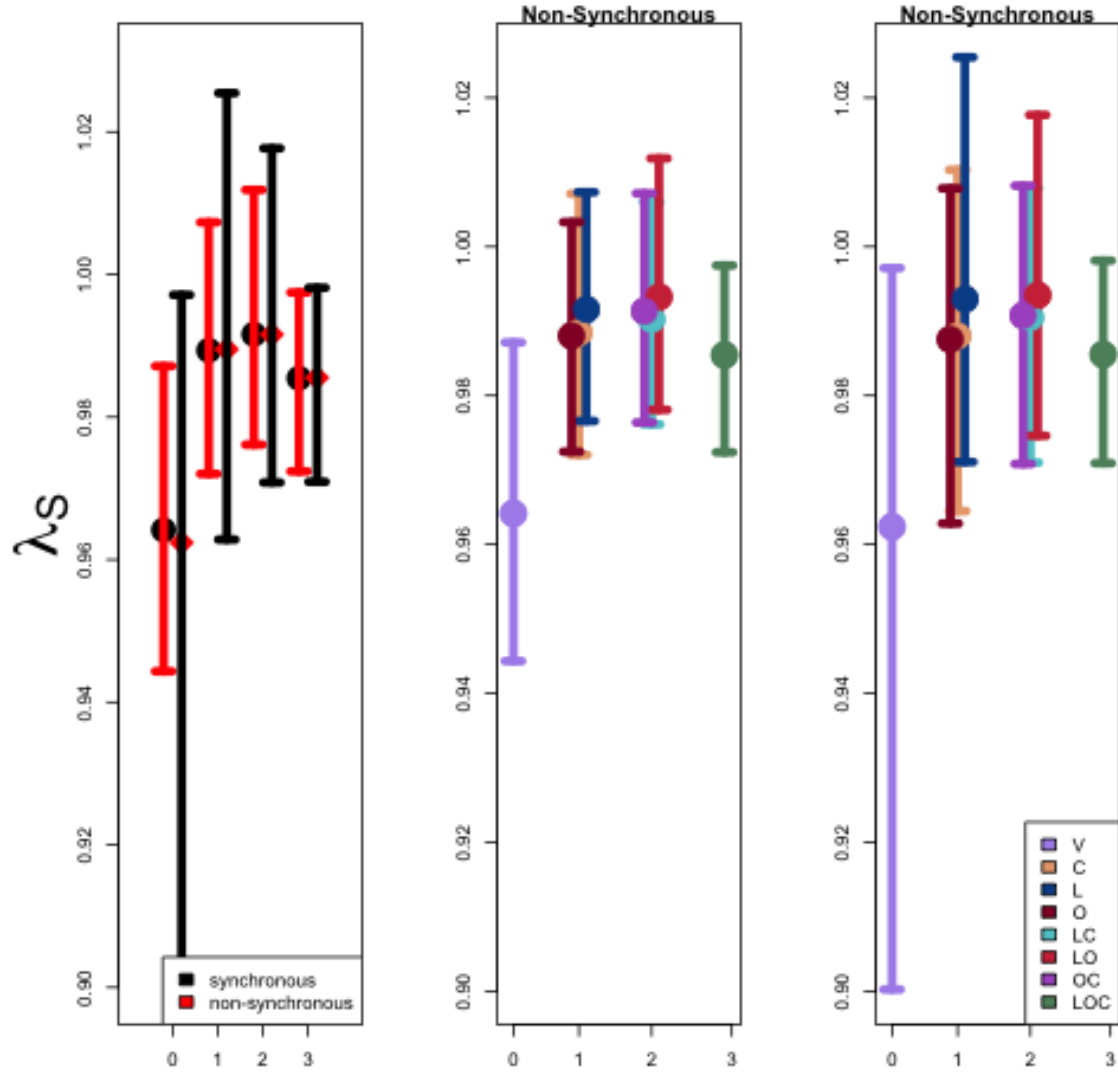


Figure 7: Panel a) shows the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). Panels b-c) shows the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

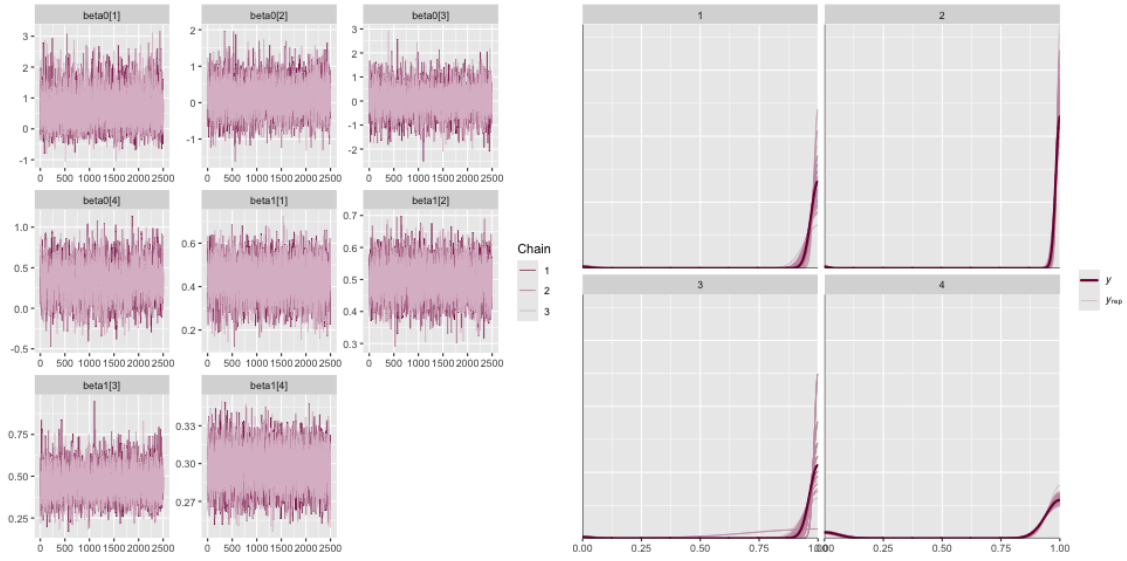


Figure 8: The a) posterior convergence of the parameters estimated by the survival model and the b) posterior distribution of survival estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean survival distribution (black line) of the real data.

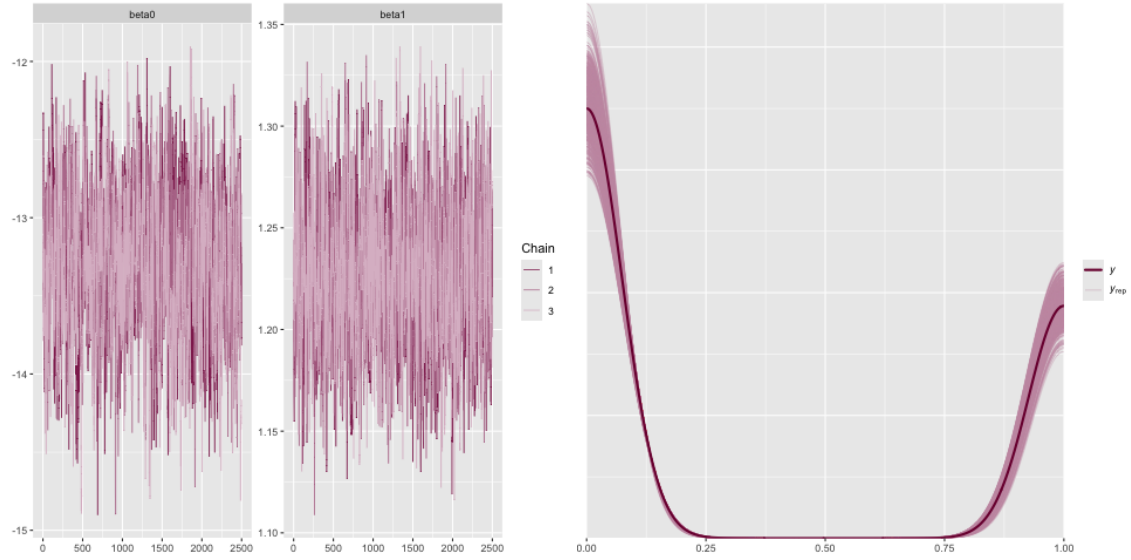


Figure 9: The a) posterior convergence of the parameters estimated by the reproduction model and the b) posterior distribution of reproductive status estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean reproductive status distribution (black line) of the real data.

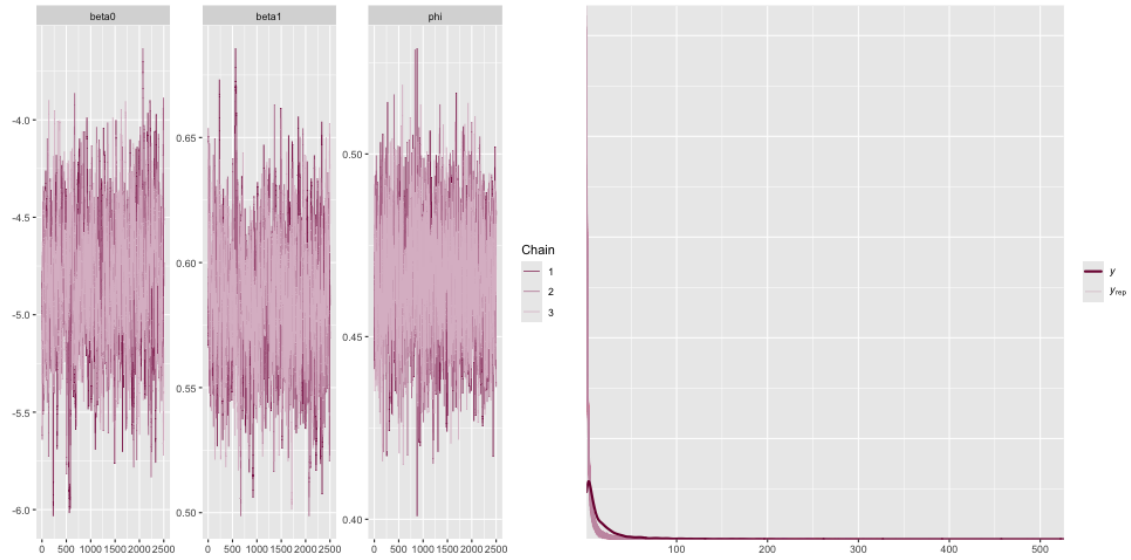


Figure 10: The a) posterior convergence of the parameters estimated by the number of flowers model and the b) posterior distribution of the number of flowers estimated (pink lines) compared to the mean distribution of observed flowers (black line).

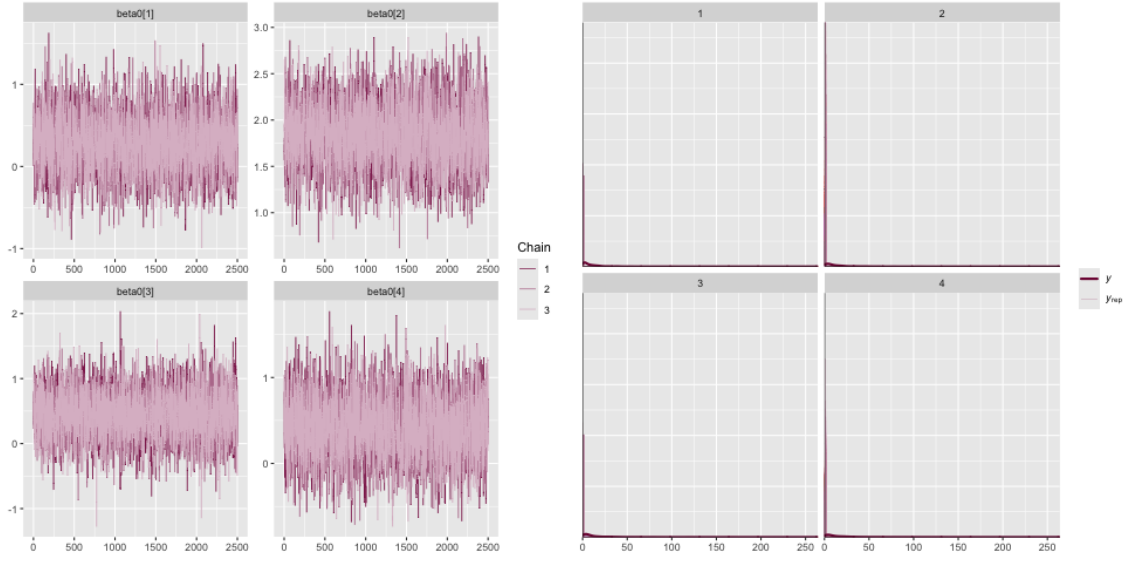


Figure 11: The a) posterior convergence of the parameters estimated by the viability model and the b) posterior distributions of floral viability estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean floral viability distribution (black line) of the real data.

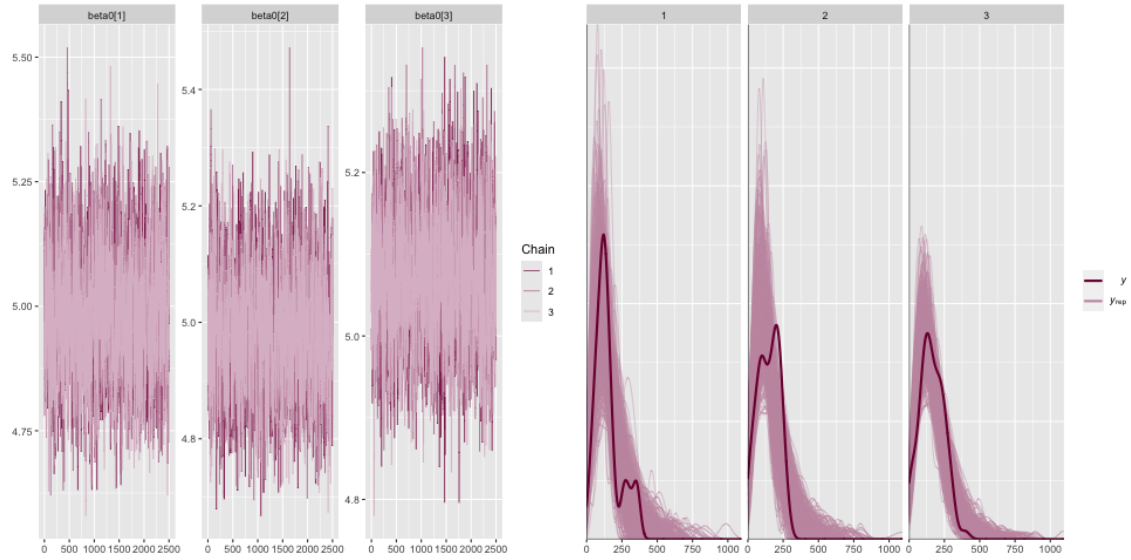


Figure 12: The a) posterior convergence of the parameters estimated by the seeds per fruit model and the b) posterior distributions of seeds per fruit estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = vacant) compared to the mean seeds per fruit distribution (black line) of the real data.

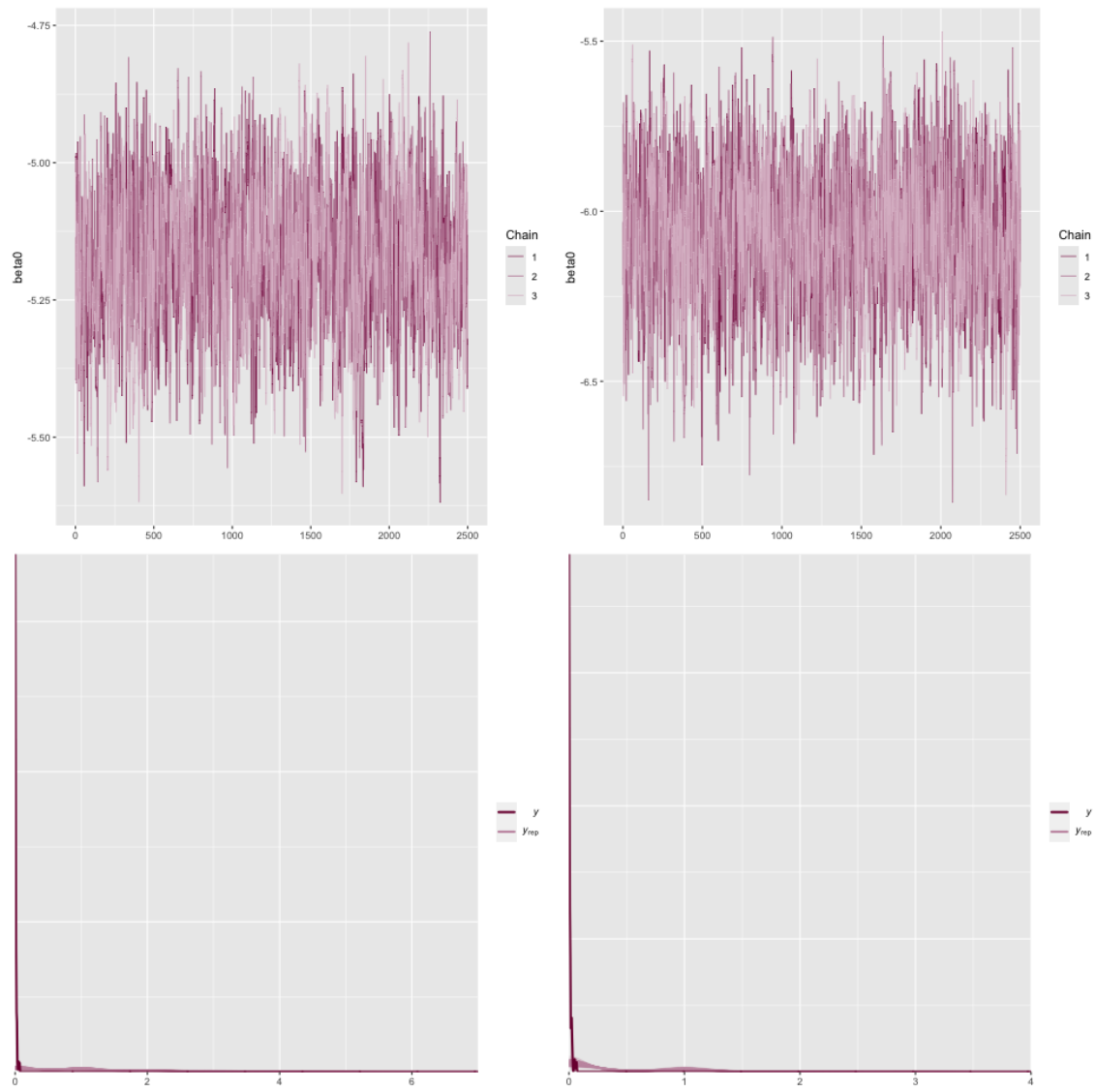


Figure 13: The a-b) posterior convergence of the parameters estimated by the germination from year one seedbank and germination from year two seedbank models respectively. The c-d) posterior distributions of floral viability estimates (pink lines) compared to the mean germination distribution (black line) of the real data for first year germinants and second year germinants respectively.

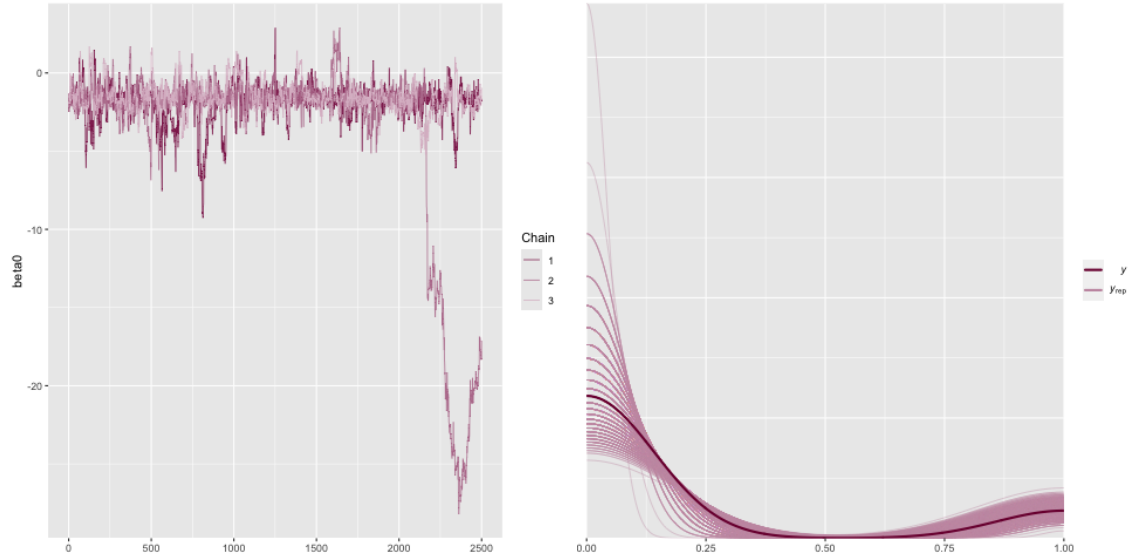


Figure 14: The a) posterior convergence of the parameters estimated by the pre-census survival model and the b) posterior distribution of the pre-census survival estimated (pink lines) compared to the mean distribution of observed pre-census survival (black line).

Statistical Models – Results

Below are the results reorted of all statistical models not described in the main body of the text.

Reproduction Model. The probability of a plant reproducing in a given year is highly size dependent. The mean probability of reproducing remains at about 0% until the plant reaches a medium size, after which the mean probability of reproducing increases steadily before reaching about 100% at large sizes.

Seeds Per Flower Model. Each viable flower on a plant produces between 97 and 257 seeds. This number is affected by the ant partner present, as shown in previous work (Ohm and Miller, 2014).

C. opuntiae tended plants produce a mean of 115 seeds per flower. *L. apiculatum* tended plants produce a mean of 143 seeds per flower. Vacant plants produce a mean of 148 seeds per flower. Comparison between posterior distributions revealed that *C. opuntiae* tended plants produced fewer

seeds per flower than *L. apiculatum* tended plants and vacant plants 80% and 87% of the time.

Vacant plants produced more seeds per flower than *L. apiculatum* tended plants only 57% of the

924 time. We are confident that *C. opuntiae* tended plants produce the fewest seeds per flowers.

Precensus Survival Model. Pre-census seed survival rates fall between 0% and 95% with the mean pre-census seed survival at 18%.

927 *Germination Model.* Seeds have a significantly higher probability of germinating in year one than in year two. Seeds in year one experience germination rates between 50% and 100% with a mean of 62% germination. Seeds in year two experience germination rates between 50% and 98% with a

930 mean of 58% germination.

New recruits are expected to be between the sizes of 0.11 cm^3 and 0.38 cm^3 with a mean size of 0.20 cm^3 .