

Thank you, next: demographic consequences of partner diversity and turnover in a multi-species ant-plant mutualism

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Abstract

We expect having multiple mutualistic partners to be beneficial for a species. These benefits be
3 explained by several mechanisms: portfolio effect, complementarity, and sampling effect. To study
multi-species demographic effects, we considered *Cylindropuntia imbricata* (tree cholla), which pro-
duce extrafloral nectar and various ant partners (*Crematogaster opuntiae*, *Liometopum apiculatum*,
6 and other less frequent ants) provide defense from herbivores and seed predators. We used plant
demographic censuses to parameterize a series of Bayesian hierarchical generalized linear vital rate
models to determine the impacts of different partners on the focal mutualists. We constructed an
9 Integral Projection Model (IPM) in which we simulate different combinations of ant partners (both
those that do and don't occur in nature). Different ant partners had different impacts on the cholla
vital rates. Specifically, *C. opuntiae* tended plants had advantages in both growth and survival
12 when small, and large *L. apiculatum* tended plants had floral viability advantages. The IPM results
revealed that despite these differences in vital rates, the presence of any ant partner is beneficial,
the identity and number of partners do not matter. This suggests that there are no benefits of
15 diversity in this system and therefore than none of the highlighted mechanisms explain the benefits
of multi-species mutualisms. This study highlights that partner diversity does not always increase
the benefits a focal mutualist receives.

Introduction

Cooperations between species, aka mutualisms, occur on a pairwise basis or in a more complex network and often play integral and foundational roles in ecological communities. Mutualisms are species interactions where all participants receive net benefits, leading to higher individual fitness and increased population growth rates. They are widespread species interactions but can deteriorate into commensalism or parasitism under conditions that elevate costs or dampen benefits (Rodriguez-Rodriguez et al., 2017; Song et al., 2020; Mandyam and Jumpponen, 2014; Thrall et al., 2007; Bahia et al., 2022; Bronstein, 1994; Chamberlain et al., 2014; Frederickson, 2013; Axelrod and Hamilton, 1981; Leigh, 2010). Mutualisms are considered more context dependent than other species interactions (Chamberlain et al., 2014; Frederickson, 2013), meaning the magnitude and sign of interaction strength are often determined by species' identities (Noe and Hammerstein, 1994; Leigh, 2010). Whether and how partner diversity modifies the demographic effects of mutualistic interactions remain open questions within relevance in applied settings (Rogers et al., 2014; Thibaut et al., 2012; Frederickson, 2005; Palmer et al., 2010). A focal mutualist may interact with multiple guilds of partner types (e.g., plants that interact with pollinators, seed dispersers, soil microbes, and ant defenders) or with multiple partner species within the same guild (e.g., plants visited by multiple pollinator species). Within a mutualist guild, partner species often differ in the amount or type of goods or services they provide, making partner identity an important source of contingency in mutualism (Stanton, 2013).

There are multiple mechanisms by which partner diversity can influence the net benefits accrued by a focal mutualist, mirroring the mechanisms by which, at a larger scale of organization, biodiversity can influence ecosystem function (Yeung et al., 2006; Barrett et al., 2015; Ushio, 2020). First, when there is a hierarchy of fitness effects (a consistent ranking of best to worst mutualists), a more diverse sample of the partner community may be more likely to include the best partner (Frederickson, 2013). This can lead to an apparent benefit of diversity driven by a sampling effect (Batstone, 2018) – but there is no benefit of diversity *per se*, only better and worse partners. If

partner associations are mutually exclusive (a focal mutualist may engage with only one partner
at a time), then partner diversity may impose opportunity costs, leading to negative effects of a
diverse mutualist assemblage relative to exclusive association with the single best partner (Miller,
2007). Second, even within a single mutualist guild, the benefits conferred by alternative partner
species can vary in type and not just degree (Stachowicz and Whitlatch, 2005; Bronstein et al., 2006;
Stanton, 2013). This can lead to a positive effect of partner diversity through complementarity of
alternative functions (Batstone, 2018). Interference or synergies between partners can make their
combined effect different than the expected from the sum of complementary functions (Afkhani,
2014). Third, partner species can have species-specific responses to environmental variation, ei-
ther spatially (Ollerton et al., 2006) or temporally (Alarcón et al., 2008). Multiple partners can
therefore act as a 'portfolio' that stabilizes fitness benefits across spatial or temporal heterogeneity,
leading to positive effects of partner diversity through the portfolio effect (Batstone, 2018; Lázaro
et al., 2022; Horvitz and Schemske, 1990). When different partners have different quality or types
of benefits, the order in which interactions takes place is extremely important.

Partner diversity can have different effects depending on whether partners are present simul-
taneously or sequentially (partner turnover) (Djiéto-Lordon et al., 2005; Ness et al., 2006; Bruna
et al., 2014; Barrett et al., 2015; Ushio, 2020; Dattilo et al., 2014). Sequential associations are
likely when alternative partners engage in interference competition for access to a shared mutualist
(Kiers et al., 2003; Batstone, 2018; Trøjelsgaard et al., 2015; Wulff, 2008). Turnover can happen
at different timescales, from minutes to years (Oliveira et al., 1999; Horvitz and Schemske, 1986).
The frequency of partner turnover can impact the level of benefits received by the focal mutualist,
particularly if the benefits continue to accumulate with successive turnover (e.g., when sequential
partners provide complementary functions) or if they saturate over time (Sachs et al., 2004; Fiala
et al., 1994). Directionality of turnover can also influence effects of partner diversity if partner
identity changes consistently across ontogeny of a focal mutualist (Fonseca et al., 2003; Noe and
Hammerstein, 1994; Dejean et al., 2008). For example, plant susceptibility to enemies can change
across life stages (Boege and Marquis, 2005; Barton and Koricheva, 2010), so the benefits of a

diverse guild of defensive mutualists are greatest when more defensive partner species align with
72 more vulnerable life stages (Djiéto-Lordon et al., 2005; Dejean et al., 2008). The length and order
of interactions are just as important as the identities of the interacting partners.

This study examined the consequences of partner diversity in a food-for-protection mutualism
75 between the tree cholla cactus (*Cylindropuntia imbricata*), a long-lived EFN-bearing plant, and
multiple species of ant partners. There are two common ant partners and several additional rarer
species, all of which collect EFN in exchange for defense against herbivores. These ant species locally
78 co-occur but individual plants are typically tended by only one species that patrols the plant around-
the-clock and maintains control of the plant’s nectar resources, usually for an entire growing season
(Ohm and Miller, 2014; Donald and Miller, 2022). Switches between partner species, or between
81 vacancy and ant occupancy, commonly occur from one growing season to the next (Miller, 2007).
No other study has integrated analyses of the demographic effects of partners on entire life cycle
with inter-annual stochastic population levels to understand multiple possible mechanistic dynamics
84 of a system. Our novel approach allows for a deeper understanding of the mechanisms at play and
the ability to pinpoint our results to specific life cycle stages.

Defensive ant-plant mutualisms – where plants provide food and/or housing to ants that in turn
87 defend them from enemies – are widespread interactions that offer valuable model systems for the
ecology and evolution of mutualism (Bronstein, 1998; Bronstein et al., 2006). Extrafloral nectar
(EFN) -bearing plants can serve as dietary resources that promote ant abundance and colony size
90 (Byk and Del-Claro, 2011; Ness et al., 2009, 2006; Donald and Miller, 2022). In turn, presence
of defensive ant partners is often linked to reductions in herbivory (Trager et al., 2010; Rudgers,
2004) and demographic advantages for the plant partner (Báez et al., 2016). Defensive ant-plant
93 mutualisms are commonly multi-species, where a guild of ant partner species share, and often
compete for, a plant mutualist (Bronstein, 1998; Beattie, 1985; Trager et al., 2010; Agrawal and
Rutter, 1998). Ant partners can vary in their ability to deter herbivores (Bruna et al., 2014), and
96 visitation by low quality ant partners can prevent visitation by higher quality partners, consequently
causing a reduction in fitness through missed opportunity costs (Fraser et al., 2001; Frederickson,

2005). Susceptibility to herbivory can also vary significantly throughout the life stages of the
 99 plant (Boege and Marquis, 2005), suggesting that the order and timing of successive partners
 is important to the fitness impacts of the combined partner guild (Barton and Koricheva, 2010;
 Boege and Marquis, 2005; Fonseca et al., 2003). Finally, herbivore identity and pressure can vary
 102 inter-annually (Wetzel et al., 2023), much like mutualist identity and presence, meaning the threat
 plants face can vary just as much as the protection they receive due to temporal stochasticity.
 Previous studies have investigated how ant partner diversity affects plant fitness (Palmer et al.,
 105 2010; Afkhami, 2014; Fiala et al., 1994; Gaume et al., 1998; Dattilo et al., 2014; Ludka et al.,
 2015) However, little is known about the combined effects of partner identity, directional partner
 turnover, and temporal stochasticity, particularly because the necessary long-term data are rarely
 108 available.

Here we used a unique long-term data set that allows us to explore mutualistic associations
 with multiple partner species, longitudinal turnover in partner identity, and how the demographic
 111 effects of alternative partners varies across plant size structure and nearly 20 years of inter-annual
 fluctuations. We used this observational data set of plant demography and ant-plant associations,
 contextualized by previous ant exclusion experiments, to investigate whether and through which
 114 mechanism(s) partner diversity affects the fitness benefits of ant visitation. Specifically, we asked:

1. What are the vital rate effects of association with alternative partners and how do these
 effects fluctuate across years?
- 117 2. What are the frequency and direction of partner turnover across the plant life cycle?
3. What is the net effect of partner diversity on plant fitness, and what mechanism(s) explain(s)
 this effect?

120 To answer these questions we used a hierarchical Bayesian statistical approach to estimate demo-
 graphic vital rates for hosts in different states of ant occupancy and to quantify state-dependent
 partner turnover from the long-term data. We then used a stochastic, multi-state integral projec-
 123 tion model (IPM) that combines diverse effects on vital rates and pathways of partner turnover to

quantify effects of partner diversity on plant fitness.

Methods

Study System

This study was conducted in the Los Piños mountains, a small mountain chain located on the Sevilleta National Wildlife Refuge, a Long-term Ecological Research site (SEV-LTER) in central New Mexico, USA. This is an area characterized by steep, rocky slopes, and perennial vegetation including grasses (*Bouteloua eriopoda* and *B. gracilis*), yuccas, cacti, and junipers. Tree cholla cacti are common in high Chihuahuan desert habitats of the southwestern USA (Benson, 1982). These arborescent plants produce cylindrical segments with large spines. In the growing season (May to August in New Mexico), the plants initiate new vegetative segments and flowerbuds at the ends of existing segments. While most plants produce new segments every season, only those which are reproductively mature produce flowerbuds. Like other EFN-bearing cacti, tree cholla secrete nectar from specialized glands on young vegetative segments and flowerbuds (Ness et al., 2006; Oliveira et al., 1999). Flowerbuds produce more and higher-quality EFN than vegetative segments, making reproductive cholla valuable mutualist partners (Miller, 2014). Smaller, non-flowering cholla produce little to no EFN and are commonly vacant (no ant visitation at the time of our census) (Miller, 2014).

Tree cholla EFN is harvested by various ground-nesting ant species. At SEV-LTER, cholla are visited primarily by two species, *Crematogaster opuntiae* and *Liometopum apiculatum*, as well as other rarer species, including *Forelius pruinosus* and unidentified species in the genera *Aphaenogaster* and *Camponotus*. *L. apiculatum* are the most frequent visitors with 25% – 60% of tree cholla tended by these ants in a given year, followed by *C. opuntiae* (5% – 20%) (Donald and Miller, 2022). Between 30% – 80% of cacti are vacant in any given year. Workers of different species rarely co-occur on individual plants, likely due to interspecific competition. For example, staged introductions of *C. opuntiae* to *L. apiculatum*-tended plants, and vice versa, provoke aggressive

responses by residents (A. Campbell, *personal observation*).

Several insect herbivores and seed predators specialize on tree cholla (Mann, 1969), and defense against these enemies is the main pathway by which ant visitation affects plant demography. The Cerambycid beetle *Moneilema appressum* and an unidentified weevil (Coleoptera: Curculionidae) of the genus *Gerstaeckeria* feed on vegetative and reproductive structures as adults and their larvae feed internally. Two species of cactus bugs, *Narnia pallidicornis* and *Chelinidea vittiger* (Hemiptera: Coreidae), feed on all cholla parts with a preference for flower buds; their damage can induce floral abortion (Miller et al., 2006). A seed predator, *Cahela ponderosella* (Lepidoptera: Pyralidae), oviposits in open flowers and larvae eat seeds in developing fruits. These consumers can have significant negative impacts on plant fitness and depress population growth (Miller et al., 2009). Prior experiments showed that ant-tended tree cholla experience less herbivory and seed predation than plants from which ants were excluded (Miller, 2007; Ohm and Miller, 2014).

Data Collection

This study is based on long-term demographic data spanning 2004 to 2023 at SEV-LTER. From 2004 to 2008, we censused 134 plants distributed across three spatial blocks. This initial census group was discontinued in 2009, when we established six 30 × 30-meter plots and tagged all tree cholla within those plots. Two additional 30 × 30-meter plots were added in 2011, and this group of eight plots has since been censused annually through 2023 (with the exception of 2020 due to the pandemic shutdown). For all plants, in May or early June of each year we recorded plant survival since the last survey and, for survivors, we recorded height (cm), maximum crown width (cm), and crown width perpendicular to the maximum (cm). Size measurements were used to calculate plant volume (cm^3) based on the volume of an elliptical cone. We measured reproduction by counting flowerbuds, and in most years we distinguished between flowerbuds that were viable and aborted. We recorded the ant species presence and absence (or vacancy if no ants present). Occurrences of more than one ant species on one plant were rare (less than 5% of observations), and for the purpose of this analysis we classified the plant as being occupied by the more abundant

species. Plots were searched for new recruits each year, and these were added to the census. These data allowed us to link each plant’s demographic fate (survival, growth, and reproduction) to its state of ant visitation. In total, the data set includes a total of 9,787 observations of 1,141 unique individuals across 15 complete transition years (spanning May/June of year t to May/June of year $t + 1$) (Campbell and Miller, 2024). In addition to missing the year 2020, there are gaps in the time series where we switched plots or plants (and thus broke up transition years for growth and survival) or where we did not distinguish between viable and aborted flower buds (Table ??).

We used additional, smaller data sets from previously published studies to estimate seed and seed bank parameters. Ohm et al. 2014 provide data on the number of seeds per fruit for plants tended by *L. apiculatum*, *C. opuntiae*, or no ants (experimental exclusion), accounting for their effects on pollinator visitation. Elder and Miller 2016 provide data on seed entry to the seed bank and seedling germination and survival rates.

Multi-state Integral Projection Model

The demographic data were used to parameterize a multi-state Integral Projection Model (IPM). IPMs describe population dynamics in discrete time with functions that relate vital rates to continuous state variables, typically size (Ellner et al., 2016). While IPMs are a natural choice for populations with continuous size structure, they can also be modified to accommodate a combination of continuous and discrete state variables, as we do here. We constructed a stochastic, multi-state IPM that stitches together population structure associated with plant size and ant state, allowing us to determine the individual fitness effects of each ant species and the composite effects of multiple partners, with ant transition dynamics and inter-annual variability modeled explicitly.

Given the low frequency of ant occupancy states other than *L. apiculatum* and *C. opuntiae* (<8% of observations) we combined all other ants into an “other” category, such that our multi-state IPM included four ant states: vacant, *L. apiculatum*, *C. opuntiae*, and Other. The “Other” category was made up of *Forelius pruinosus* (3.5% of observations), unidentified species belonging to the genera *Camponotus* (0.9%), *Aphaenogaster* (0.4%), *Myrmecocystus* (0.08%), *Tetramorium*

(0.02%), *Brachymyrmex* (0.02%), and additional ants not identified to genus or species (2.8%).

We initially tried to split this category up further, defining other ant states, but found that error estimates were too high given low frequencies.

Ant state is included as a predictor variable in IPM sub-models where there are biologically realistic pathways through which ants could impact the outcome of that process. For example, ant partners defend cacti from herbivores, and prior experimental work indicates that ant tending can reduce vegetative tissue loss and floral abortion. Therefore, ant state was included in sub-models for survival, growth, flowerbud viability, and seed number per flowerbud. In contrast, we have no reason to expect that ant tending can directly influence the probability of flowering or flowerbud production independently of its influence on plant size, so these sub-models do not include ant state as a predictor variable.

We modeled the tree cholla life cycle using continuously size-structured plants where number of plants of size x and ant state a in year t ($n(x, a)_t$) predicts the number of plants of size x' and ant state a' in year $t + 1$ ($n(x', a')_{t+1}$) based on a size- and ant-specific vital rates. The models also includes two discrete seed banks (B_t^1 and B_t^2) corresponding to 1 and 2-year old seeds. Seed bank dynamics are given by:

$$\begin{array}{l} \text{First Year Seedbank} \\ \underbrace{B_{t+1}^1} \end{array} = \delta \sum_{a=1}^A \int_L^U \underbrace{\kappa(a)}_{\text{\# seeds per fruit}} \underbrace{P(x; \boldsymbol{\tau}^P)}_{\text{prob. of flowering}} \underbrace{F(x; \boldsymbol{\tau}^F)}_{\text{\# flowerbuds}} \underbrace{V(a; \boldsymbol{\tau}_a^V)}_{\text{flower viability}} \underbrace{n(x, a)_t}_{\text{\# size x plants}} dx$$

$$\begin{array}{l} \text{First Year Seedbank (yr t+1)} \\ \underbrace{B_{t+1}^1} \end{array} = \overbrace{\kappa(a)P(x; \boldsymbol{\tau}^P)F(x; \boldsymbol{\tau}^F)V(a; \boldsymbol{\tau}_a^V)n(x, a)_t}^{\text{\# of viable seeds produced by plants of size x (yr t)}}$$

$$\begin{array}{l} \text{Second Year Seedbank (yr t+1)} \\ \underbrace{B_{t+1}^2} \end{array} = \underbrace{(1 - \gamma_1)B_t^1}_{\text{Seeds which did not germinate in the first year in the seedbank}}$$

$$B_{t+1}^2 = (1 - \gamma_1)B_t^1$$

Functions $P(x; \boldsymbol{\tau}^P)$ and $F(x; \boldsymbol{\tau}^F)$ give the probability of flowering in year t and the number of flowerbuds produced in year t , respectively, by plants of size x in year t . The proportion of

flowerbuds that remain viable through fruit set ($V(a; \tau_a^V)$) and the number of seeds per fruit ($\kappa(a)$) is dependent on ant state a . The vectors τ give year-specific deviates (mean zero) and appear in functions for which we can estimate temporal stochasticity from the long-term data; superscripts indicate the corresponding vital rate and, when present, the a subscript indicates that deviates are specific to plants in ant state a . For example, temporal deviates τ_a^V describe better- and worse-than-average years for flowerbud viability and plants in different ant states can fluctuate independently (good years for *L. apiculatum* -occupied plants may not be good years for *C. opuntiae*-occupied plants, for example). Seed production is integrated over the size distribution, from the lower L to upper U size limits, and summed over all possible initial ant states ($A = 4$) giving total seed production. Seeds are multiplied by the probability of escaping post-dispersal seed predation (δ) to give the number of seeds that enter the one-year old seed bank. Plants can recruit out of the year-one seed bank with probability γ_1 or transition to the two-year seed bank with a probability of $1 - \gamma_1$. Seeds in the two-year seed bank are assumed to either germinate with probability γ_2 or die.

For the above-ground part of the life cycle, the number of plants of size x' and ant state a' in year $t + 1$ is given by survival/growth transitions from size x and ant state a in year t , plus germination out of the seed banks:

$$\begin{aligned}
 \text{Plants of size } x' \text{ and ant state } a' &= \text{new plants coming from the seedbanks} \\
 n(x', a')_{t+1} &= (\gamma_1 B_t^1 + \gamma_2 B_t^2) \eta(x') \omega \rho_0(a') + \\
 &\quad \overbrace{\sum_{a=1}^A \int_L^U S(x, a; \tau_a^S) G(x', x, a; \tau_a^G) \rho(x, a, a'; \tau^e) n(x, a)_t dx}^{\text{Already existing plants growing}}
 \end{aligned}$$

The first term in Eq. ?? estimates the number of individuals recruiting from a one or two-year seed bank to a plant of size x' and ant state a' based on the recruit size distribution $\eta(x')$ and the probability of over-winter seedling survival (ω) from germination (late summer) to the census (May). This term is multiplied by $\rho_0(a')$, which gives the probability that a new recruit has ant state a' ($\sum \rho_0(a') = 1$). The second term represents all possible transitions from size x and ant a

to size x' and ant a' , conditioned on survival. Survival ($S(x, a; \boldsymbol{\tau}_a^S)$) and growth from size x to x' ($G(x', x, a; \boldsymbol{\tau}_a^G)$) are both dependent on initial size and ant state. As above, these functions include inter-annual variability through year-specific deviates that can vary by ant state ($\boldsymbol{\tau}_a$). Finally, ant transition function $\rho(a', a, x; \boldsymbol{\tau}^\rho)$ gives the probability that an individual transitions from ant state a to a' in the next census, conditional on initial size x . This function includes inter-annual variability through year-specific intercepts which are consistent across initial ant states ($\boldsymbol{\tau}^\rho$).

Statistical modeling and parameter estimation

We parameterized the IPM using a series of generalized linear mixed models in a hierarchical Bayesian framework. Vital rate models included spatial and temporal random effects associated with plot and year variation, respectively (only year variation is used in the IPM), and included plant size (the natural logarithm of volume, $\log(\text{cm}^3)$; x, x'), ant partner state (a, a'), or both as fixed-effect predictor variables. As in the IPM, our statistical modeling assumed that demographic effects of ant occupancy are limited to survival, growth, and flowerbud viability.

Growth. We fit the growth sub-model ($G(x', x, a; \boldsymbol{\tau}_a^G)$) to data on size in year $t + 1$ (y^G) using the skewed normal distribution to account for left-skewed size transitions (at some initial sizes, transitions below the expected future size were more common than transitions above it). The skew-normal has three parameters corresponding to location (\hat{G}), shape (σ), and scale (α):

$$y_i^G \sim \text{SkewedNormal}(\hat{G}_i, \sigma_i, \alpha_i)$$

$$\hat{G}_i = \beta_{a[i]}^0 + \beta_{a[i]}^1 x_i + \beta_{a[i]}^2 x_i^2 + u_{\text{year}[i], a[i]} + w_{\text{plot}[i]}$$

$$\log(\sigma_i) = \beta^3 + \beta^4 x_i$$

$$\alpha_i = \beta^5 + \beta^6 x_i$$

Here, the location parameter for the i th observation \hat{G}_i is defined as a second-order polynomial with ant-size interactions because preliminary analysis found this was an improvement over a linear

relationship. The location parameter of the skew-normal is not the mean, but the mean can be
 276 derived as $\hat{G} + \frac{\sigma\alpha}{\sqrt{1+\alpha^2}}\sqrt{\frac{2}{\pi}}$. The year- and ant-specific random effect u (which parameterizes the $\boldsymbol{\tau}_a^G$
 vectors) and plot-specific random effect w are normally distributed with variances σ_{year}^2 and σ_{plot}^2 ,
 respectively. Parameters σ_i and α_i control residual variance and skewness, respectively, and were
 279 defined as linear functions of initial size x_i (σ_i is strictly positive and was modeled with a log link
 function). After several rounds of model selection, we chose to model growth variance and skewness
 as not dependent on ant occupancy state.

282 *Survival.* The survival sub-model ($S(a, x; \boldsymbol{\tau}_a^S)$) estimates the probability of survival from year t to
 year $t + 1$, with fixed effects of size x and ant partner a in year t . We fit this model to the survival
 data (alive or dead) using a Bernoulli distribution with a similar linear predictor for the probability
 285 of survival as in the growth model but with a logit link function and without the second-order
 influence of size.

Reproduction. The flowering sub-model ($P(x; \boldsymbol{\tau}^P)$) estimates the probability of reproducing in
 288 year t , with fixed effects size x in year t and random effects of plot and year. We fit this model
 to the reproductive status data (vegetative or flowering) using a Bernoulli distribution and a logit
 link function, similar to the survival model above but with no ant effects. The flowerbud function
 291 $F(x; \boldsymbol{\tau}^F)$ estimates the total flowers produced by a reproducing plant in year t , with fixed effects of
 size x in year t . We fit this model to flowerbud count data (sum of viable and aborted buds) using
 a zero-truncated negative binomial distribution with a log link and normally distributed year and
 294 plot random effects.

The flowerbud viability sub-model ($V(a; \boldsymbol{\tau}_a^V)$) estimates the probability that flowerbuds pro-
 duced in year t remain viable (not aborted), with fixed effects of ant partner a in year t . We fit
 297 this model to floral viability data using a binomial distribution where trials and successes are given
 by the total number of flowerbuds and the number that are viable, respectively. This model used
 a logit link function and included random effects for plot and year following the same structure as
 300 the survival model, with ant-specific year random effects.

Estimates for the number of seeds per fruit were obtained from a field experiment which excluded ants (Miller et al., 2006). This data only included *L. apiculatum* and *C. opuntiae* ants, so we had
 303 to make an assumption about the seeds per fruit on plants tended by Other ants. We chose to use the vacancy estimates for plants tended by other ants, a decision which does not have a significant impact on the final results. Additional reproductive parameters for the number of seeds per fruit, probability of entry to the seed bank, germination rates, and recruit size were estimated following
 306 methods described in Appendix A.

Ant Transitions. The ant transition model ($\rho(x, a, a'; \boldsymbol{\tau}^e)$) estimates the probability of a cactus
 309 being occupied by ant partner a' in year $t + 1$ given that it was occupied by initial ant partner a in year t , with fixed effect of initial size x . We fit this model to ant partner data using a multinomial distribution with a logit link function.

Parameter estimation. We fit models using Stan run through version 4.0.2 of R (R Core Team, 2023; Stan Development Team, 2023). We used vague priors for all parameters. For each model, we obtained three chains of 10,000 iterations, discarding the first 1,500 iterations. We visually assessed
 315 parameter convergence between and within chains (Figures S9 – 15b) and assessed overall model fit with posterior predictive checks to examine how well the fitted model can generate simulated data similar to the real data (Figures S9 – 15a).

318 *IPM Analysis*

Analyzing the IPM required that we discretize the continuous IPM kernel into an approximating matrix. Size variable x was discretized into b bins, resulting in a $b \times b$ matrix. In our model there is
 321 additional complexity in the form of transitions between A ant states and two additional discrete states (year one and year two seed banks), leading to a matrix size of $A(b + 2) \times A(b + 2)$. We used $b = 500$ bins, which we found to be sufficient for numerically stable outputs, and extended the
 324 integration limits beyond the minimum (L) and maximum (U) observed sizes to avoid unintentional

eviction using the “floor-and-ceiling” method (Williams et al., 2012).

For stochastic analyses, we estimated the approximating matrix corresponding to each t to $t + 1$ transition year. To estimate population mean fitness in a stochastic environment (λ_S) we simulated population dynamics for 500 years by randomly sampling among the 15 annual transition matrices, discarding the first 100 years of the simulation to minimize the influence of initial conditions. Sampling observed transition matrices (rather than independently sampling regression coefficients) produces demographic time series that realistically capture inter-annual variation by preserving correlations between vital rates (Metcalf et al., 2015). We tallied the total population size at each time step as $N_t = B_t^1 + B_t^2 + \sum_{a=1}^A \int n(x, a)_t dx$ and calculated the stochastic growth rate as

$$\log(\lambda_S) = E[\log(\frac{N_{t+1}}{N_t})]$$

(Mark and Ellner, 2009). We propagated uncertainty from the vital rate models using 100 draws from the joint posterior distribution of model parameters, resulting in a posterior distribution of λ_S and other derived quantities.

Partner diversity simulation experiments

Using the fully parameterized multi-state IPM, we conducted simulation experiments to quantify how diversity and identity of ant partners influenced plant fitness. From the full version of the model (described above) corresponding to the observed assemblage of partners and observed frequencies of partner transition, we created treatments corresponding to all eight “counter-factual” scenarios of diversity and composition: no ant partners (complete vacancy); one ant partner (*C. opuntiae* only, *L. apiculatum* only, Other only); two partners (all pairwise combinations of *C. opuntiae*, *L. apiculatum*, and Other); and three partners (observed scenario of all ant states). These simulation experiments were made possible by extrapolating ant-specific demographic performance across the size distribution, even for combinations of size and ant occupancy that were rarely observed. For example, the no-partner scenario modeled a hypothetically ant-free cactus population, even though no such population exists to our knowledge, by applying the statistical knowledge gleaned from

vacant plants across the size distribution. We refer to stochastic fitness associated with partner
 342 number or identity using superscripts, e.g. λ_S^0 for vacant plants (zero partners), λ_S^{1+} for any state
 of ant tending, λ_S^C for tending by only *C. opuntiae*, λ_S^{CO} for tending by *C. opuntiae* and Other
 ants, etc.

345 In all scenarios that included any ant partners, we preserved the observed pattern of size-
 dependent vacancy/occupancy (estimated through the ant transition sub-model) and manipulated
 partner identity conditional on occupancy. This means, for example, that the *C. opuntiae*-only
 348 scenario included two possible states, vacancy and occupied by *C. opuntiae*. While our statistical
 models allow us to extrapolate the demographic performance of ant-tended plants to small sizes that
 are typically vacant, the natural history of this system tells us that this is not biologically sensible.
 351 Small, non-reproductive plants are typically vacant because they do not produce extrafloral nectar,
 and once plants begin producing nectar they are nearly always ant-tended (Miller, 2014). Our
 simulation experiments preserved this basic biology, avoiding tiny ant-occupied plants that do not
 354 and could not occur in nature.

The partner diversity treatment scenarios required additional assumptions about the mecha-
 nisms that give rise to observed occupancy patterns. Based on evidence that EFN-bearing cacti
 357 are nearly always ant-occupied (Miller, 2014), we assume that ant partners competitively exclude
 one another from EFN-bearing cacti and that competition is zero-sum. This means that, in scenar-
 ios that remove species from the partner community, remaining species gain access to plants that
 360 the removed species would have tended. In Appendix C, we present results under an alternative
 assumption, that ant visitation is limited by factors other than availability of cactus EFN (e.g.,
 nesting sites or off-plant dietary resources), such that when a species is removed from the partner
 363 community, the plants it would have occupied remain vacant.

Temporal stochasticity experiments

Under the portfolio effect hypothesis, partner diversity may confer a fitness advantage when the
 366 benefits of alternative partners are not perfectly synchronized across temporal environmental vari-

ation, yielding an advantage of a diverse “portfolio” of partners when the environment fluctuates. Our statistical estimation of ant-specific year random effects in the vital rates allows for this possibility. We constructed two versions of the stochastic, multi-state IPMs that allowed us to test this hypothesis. The baseline, ‘non-synchronous’ model described above included ant effects that could vary uniquely across time. In this scenario, good years and bad years can differ between ant states, according to the parameter estimates for the random effects (τ_a). We quantified from the fitted random effects how tightly inter-annual variation was correlated between ant states for each vital rate. The ‘synchronous’ (S) version included ant effects that were forced to be the same across species. To synchronize ant states, we averaged the ant-specific year random effects, thus ensuring that plants in all ant states fluctuated synchronously in response to temporal environmental variation. This second, “synchronized” version of the model effectively turns off any portfolio effect, holding all else equal. Both scenarios of temporal stochasticity, non-synchronized and synchronized, were run for all eight ant partner scenarios described above. We indicate stochastic fitness under the synchronized scenario with $\lambda_S^{C(s)}$, for example.

Statistical inference on fitness consequences of partner identity and diversity

The range of models we created generated many outputs; we focus our inference on the following specific contrasts. First, to determine whether ant occupancy and partner diversity are beneficial, we calculated a posterior distribution of λ_S for each of four partner richness levels (zero, one, two, three), averaging over composition scenarios within each level. If cactus fitness increases with partner richness, this would be interpreted as evidence for benefits of partner diversity. Second, to determine whether each partner, in isolation, confers a fitness advantage and to rank alternative partners, we contrasted the fitness of each single partner scenario (*C. opuntiae* only, *L. apiculatum* only, Other only) against vacancy (zero partners). Third, to determine whether any benefits of diversity are due to the sampling effect or complementarity, we contrasted the fitness of multi-partner scenarios against the single best partner scenario. If the best multi-partner scenario exceeds the fitness associated with the best single partner, this would be interpreted as evidence of comple-

mentarity, a true benefit of diversity *per se*. Alternatively, the sampling effect hypothesis predicts that no multi-partner scenario yields higher plant fitness than the best single partner. It is also possible that multi-partner scenarios yield lower fitness than the single best partner, which would be consistent with an opportunity cost of diversity. Fourth, to quantify any contribution of the portfolio effect, we contrasted λ_S of the full (four-state) scenario to vacancy (as a measure of how much benefit is received from all partners being present) for synchronized and non-synchronized responses to temporal stochasticity. If the portfolio effect confers a benefit of diversity, the fitness advantage of having all vs. no partners should be greater when temporal fluctuations are not synchronized across ant states.

We base our statistical inferences on the posterior probability distributions of the contrasts described above. For example, the contrast of *C. opuntiae* (λ_S^C) with vacancy (λ_S^V) yields a posterior distribution of the difference $\Delta\lambda_S^{C-V}$. We can quantify from this distribution our certainty in the mutualistic effect of *C. opuntiae*, given the data, as $Pr(\Delta\lambda_S^{C-V} > 0)$. We apply similar logic to other contrasts described above. We interpret contrast probabilities ≥ 0.95 as statistically significant differences.

Results

What are the effects of association with alternative partners on vital rates and how do these effects fluctuate across years?

Over the 20-year data set, we found that ant partners influenced demographic performance of cactus hosts, and different ant partners had contrasting demographic effects across host vital rates. Ant-tended plants had a growth advantage over vacant plants, especially at smaller sizes (Figure 1). Over the size range for which plants were likely ant-tended (larger than the minimum observed size of $0.8 \log(cm^3)$; solid lines in Figure 1) there were modest differences between partner species, with the greatest growth advantage associated with *C. opuntiae* followed by *L. apiculatum* and

417 then Other ants. At large sizes the growth trajectories of ant-tended and vacant plants were nearly indistinguishable. For all ant states, growth was left-skewed, with transitions to sizes below the mean were more common than sizes above the mean.

420 Similarly, for plants which were large enough to have ant visitors, visitation enhanced cactus survival (Figure 2). Mean survival rates ranged from 7.7% to 99.9%, with the smallest plants the most vulnerable to mortality. *C. opuntiae*-occupied plants had a survival advantage over other ant-tended plants, particularly at smaller sizes, consistent with the positive effects on growth. At larger sizes, plants in any state of ant occupancy had a survival advantage over vacant plants. Plants that were smaller than $-2 \log(cm^3)$ were predicted to experience a survival advantage from vacancy, but 426 apparent benefits of occupancy were based on extrapolated survival estimates of ant-tended plants. Plants in this size range were never observed to be ant-tended and, because the IPM preserves the size-dependence of vacancy, benefits of vacancy at small sizes are never realized in the IPM.

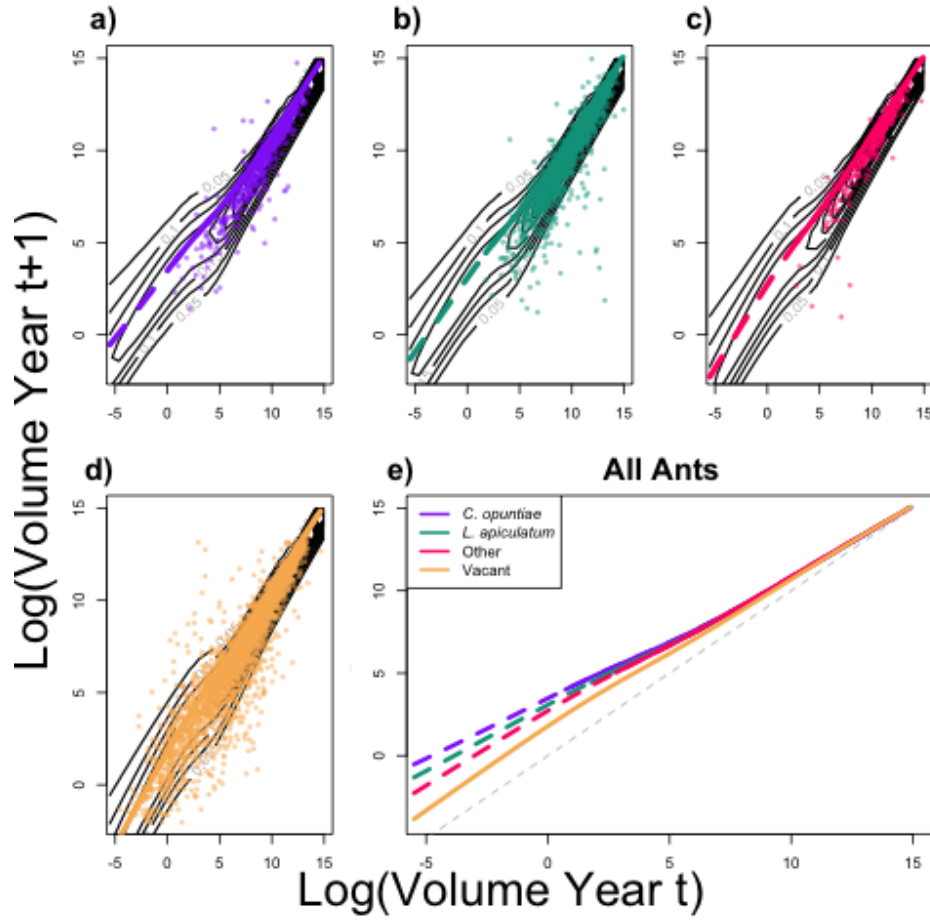


Figure 1: The next predicted size of cholla based on previous size with each individual ant partner. The colored lines (seen in all panels) are the next mean predicted size of cholla. The points (seen in panels a-d) are the observed data which informs these estimates. The black contour lines (seen in a-d) appear at 5% increments showing where 5%, 10%, etc. of the data is expected to fall. Panel e shows all mean estimates from panels a-d on one graph. Dashed colored lines indicate extrapolation while solid lines indicate estimations within the range of observed data. The vacant (yellow) line is solid across the entire range because we have observed vacant plants at all sizes. The other lines must be extrapolated to cover the entire size range of the plants as we do not observe ants on small plants often. The gray dashed line (in panel e only) shows the line where the next predicted size is the same as the previous (aka there is no growth on this line and below this line is shrinkage).

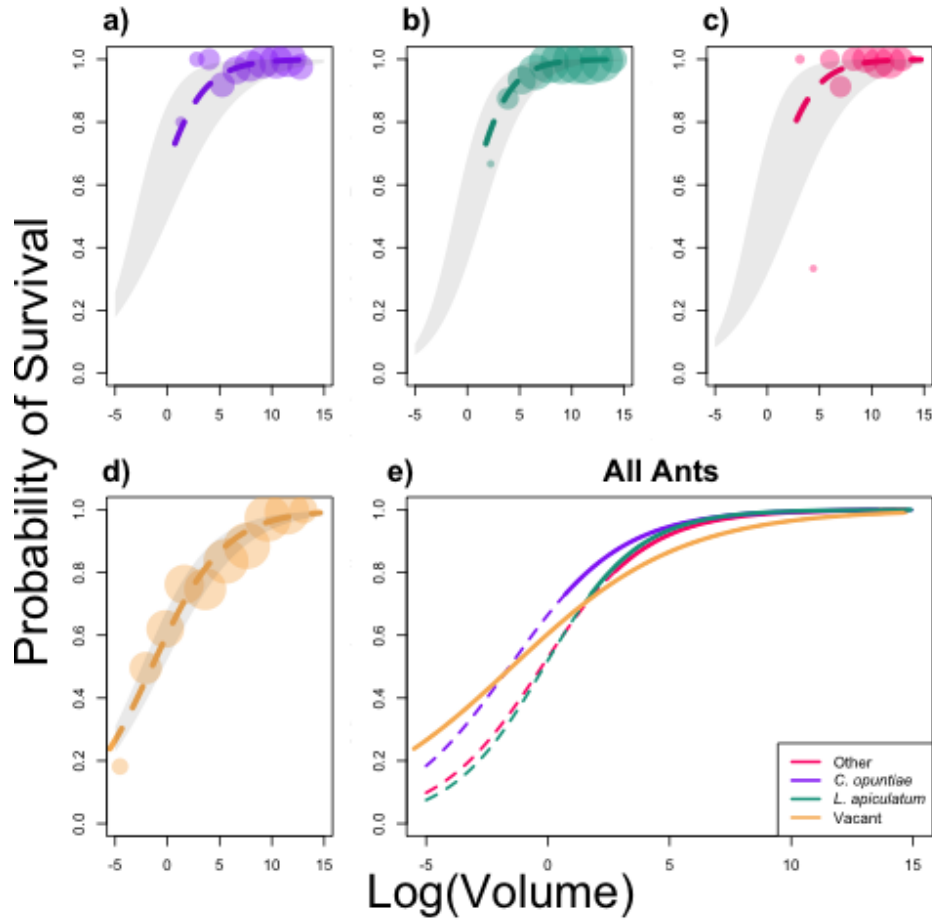


Figure 2: The estimated survival rates based on the size of the cactus with each individual ant partner. The solid colored lines (shown on all panels) indicate the mean estimated survival rates. The dashed lines (shown in panel e) indicate extrapolations beyond existing data (where we estimated survival for plants tended by ants where we had never seen a tended cactus of that size). The grey area around the solid lines (shown in panels a-d) show the 90% confidence interval for the estimates. The colored dots are the real data binned by size to show how our estimates align with real survival observations. A larger circle means we had more data on survival of plants of this size with this partner. The mean estimates from panels a-d are combined on one graph in panel e.

429 We found evidence that ant tending was associated with increased floral viability rates and that ant identity can influence the strength of viability benefits. We observed mean viability rates

of flowers between 40% and 92% (Figure 3). *L. apiculatum*-tended plants had the highest mean
432 viability rate (86.1% [95% credible interval: 77.6–92.4%]), and there were similar viability rates for
vacant (60.0% [44.3–75.0%]), Other-tended plants (at 60.6% [43.7–75.5%]), and *C. opuntiae* tended
plants (57.1% [40.6–72%]). Furthermore, *C. opuntiae*-tended plants had fewer seeds per fruit
435 (115.0[79.5–165.5] seeds) than vacant (147.2[114.1–189.9] seeds) or *L. apiculatum*-tended plants
(142.4[100.7–200.2] seeds) (Figure S1).

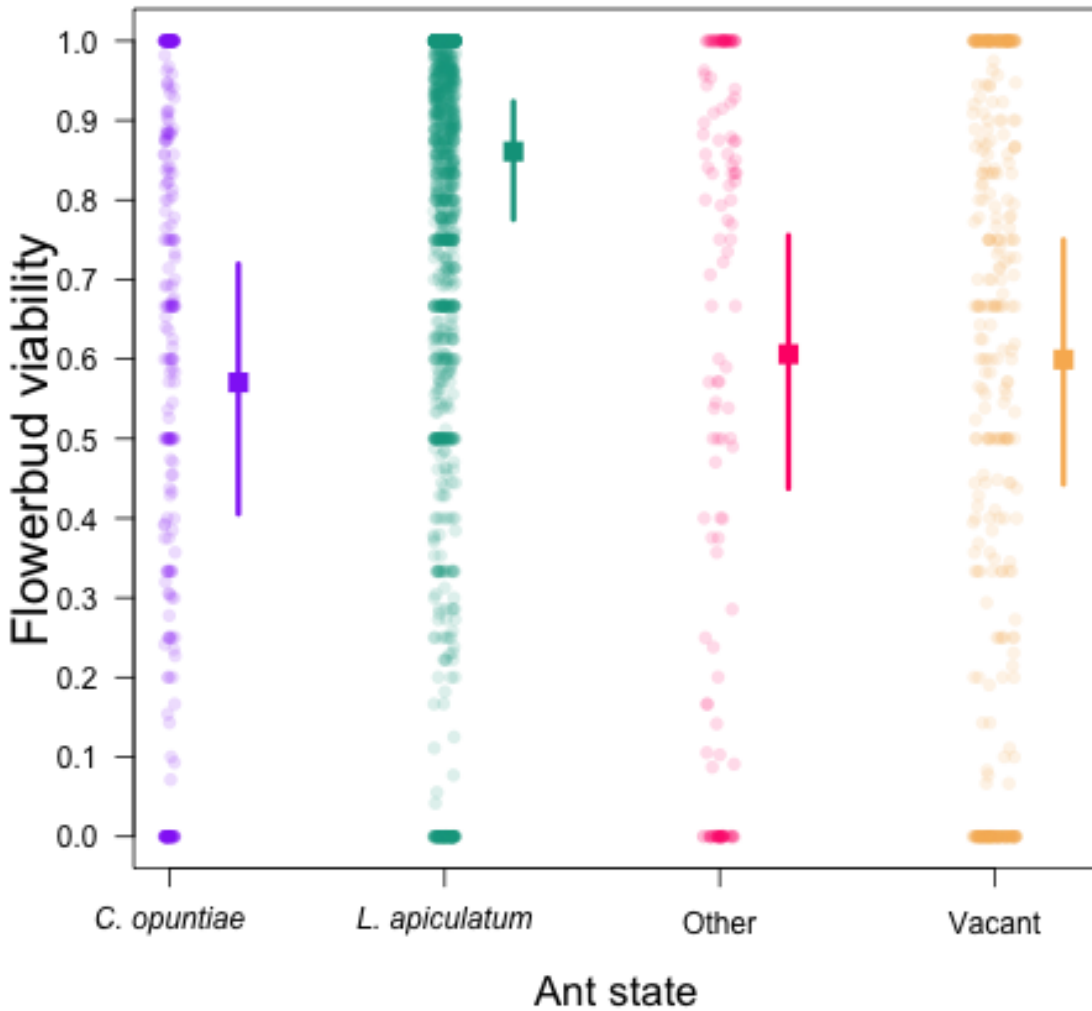


Figure 3: The estimated distributions of floral viability rates compared to observed distributions of floral viability rates of cholla based on ant partner identity. The clouds of points show the observed data, with the color representing the ant state. The solid lines indicate 95% of the estimated viability posterior distribution.

The results of all other vital rates as well as the posterior predictive checks are all included in

the Appendix A : Figures S1 – 4.

We found that the annually varying effects of ants on vital rates differed. We analyzed the correlation coefficients of all models that included ant state as a predictor and found that, across vital rates, the annually varying effects on growth were the most correlated between ant states (mean pairwise correlation: 0.63) and effects on survival were the least correlated (0.36) (Figure 4). For growth, annually varying random effects of Other ants were the least synchronous with other ants states (correlation coefficients ranging from 0.28 to 0.39) while the effects of vacancy were the most synchronous with other states (correlation coefficients ranging from 0.39 to 0.81). For survival, effects of *C. opuntiae* were the least synchronous with other ant states (correlation coefficients ranging from 0 to 0.15) while the effects of vacancy were the most synchronous (correlation coefficients ranging from 0.13 to 2.6). For floral viability, effects of *L. apiculatum* were the least synchronous with other ant states (correlation coefficients ranging from 0.06 to 0.44) while effects of vacancy were the most synchronous (correlation coefficients ranging from 0.44 to 0.62). The variation in synchrony across ant states and vital rates indicates there is potential for portfolio effect, as many of the ant effects revealed low synchronicity, particularly in survival.

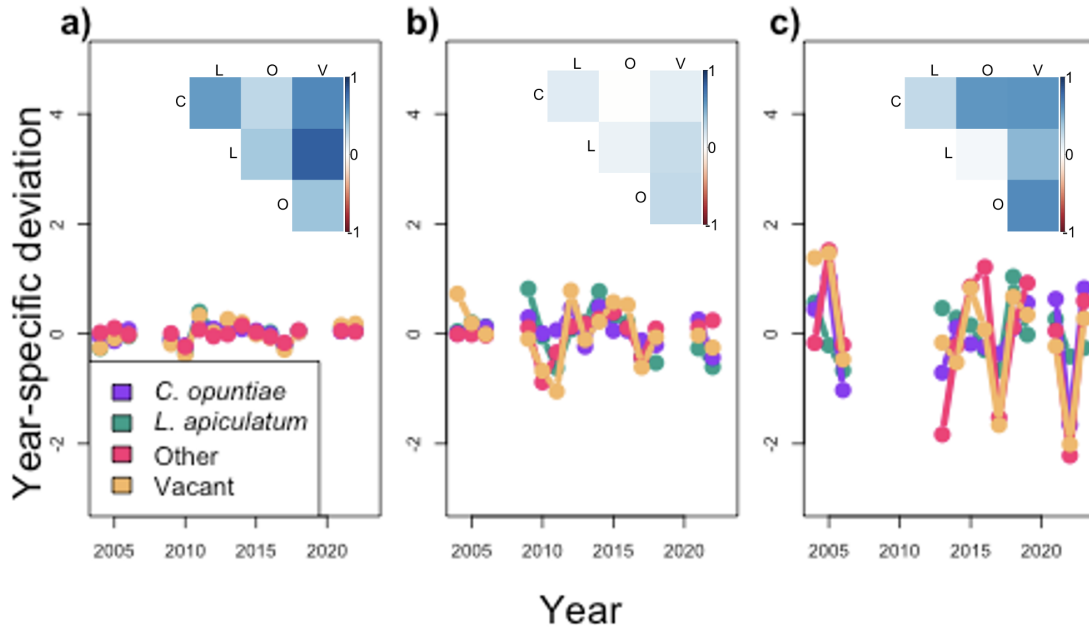


Figure 4: The colored lines in each panel show the mean effect of each ant partner on a vital rate across years (blank periods indicate a lack of transition data). The blue squares in the upper corner of each panel show the correlation matrix of how similar ant-specific annual responses are for each vital rate. Panel a) shows the effects on the estimated next size, b) the effects on estimated survival, and c) the effects on floral viability of cacti across every year of our study. These values are estimated from the fitted random effects of ant and year in our models. Each point represents the mean of the random effect of the identified model, ant, and year (e.g. the lowest dot in panel b) represents the mean effect of vacancy on survival rates in year 2011).

What are the frequency and direction of partner turnover across the plant life cycle?

We found that 55% of individual plants surveyed in the long-term data experienced at least one ant state transition on average, with very distinct size-dependence and directional patterns (Figure 5). Vacancy was the most likely ant state of small plants ($\leq 10 \log(\text{cm}^3)$). Even when small plants

were ant-tended at the start of the transition year, they were most likely to transition back to
459 vacancy (Figure 5b-d). The probability of becoming ant-tended increases with size, though it was
not equally likely to be tended by all partners. For large plants that were initially vacant or tended
by *L. apiculatum* or Other ants, *L. apiculatum* was the most likely next partner, suggesting that
462 this species is able to colonize plants that were previously vacant or occupied by Other ants, and
effectively retain plants that it previously occupied. *C. opuntiae* were also able to retain plants
they previously occupied, but not as well as *L. apiculatum*: for plants that begin the transition
465 year with *C. opuntiae*, the probability that those plants remain occupied by *C. opuntiae* at the end
of the transition year is only slightly greater than the probability of take-over by *L. apiculatum*,
while take-over in the other direction is extremely rare. It is also notable that transitions away
468 from the initial state of *L. apiculatum* were almost always transitions to vacancy (Figure 5d), while
transitions away from the initial states of *C. opuntiae* and Other were often transitions to other
ants. This suggests a competitive hierarchy whereby *L. apiculatum* may abandon low-value plants
471 with little nectar production but is almost never displaced from high-value plants.

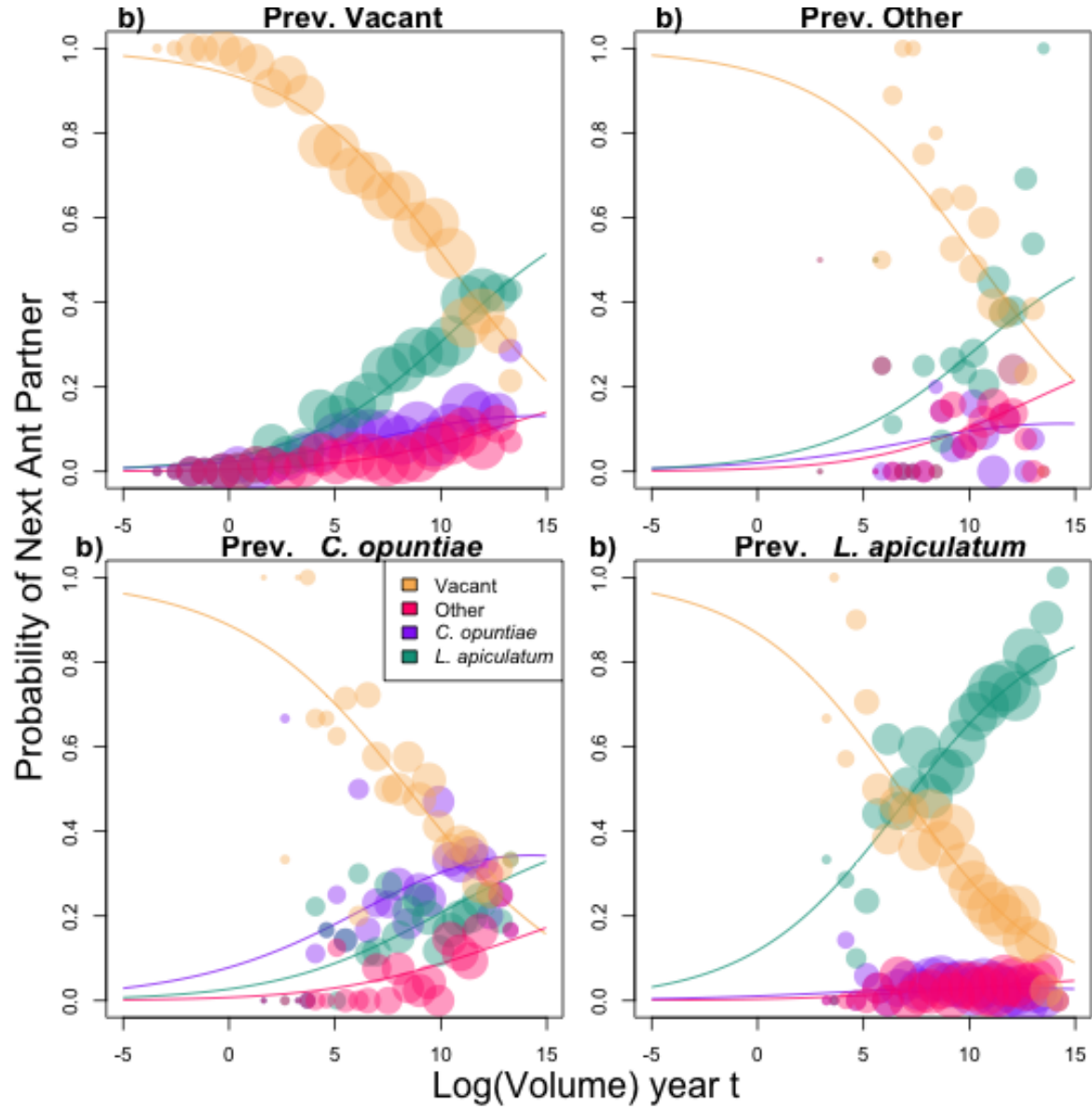


Figure 5: The probability of being tended by each ant partner or vacant based on the size of the plant. Each panel shows these probabilities for a different previous ant state. The solid lines represent the mean probability of being tended by a specific partner. The colored points are the real data binned by size to show how our estimates align with real visitation observations. A larger circle means we had more data on visitation of plants of this size with this previous partner.

What is the net effect of partner diversity on plant fitness, and what mechanism(s) explain(s) this effect?

By integrating vital rate results, temporal fluctuations, and ant transition dynamics into the stochastic, multi-state IPM we can evaluate the fitness implications of different scenarios of partner diversity and identity. First, there was strong evidence that ant visitation had mutualistic fitness effects on plant partners. The lowest mean stochastic fitness was λ_S^0 , the fitness of the cholla with no partners (Figure 6a). Across all 1+ partner scenarios, we found that the λ_S^{1+} posterior distributions were greater than λ_S^0 with nearly 100% certainty. This indicates that ant visitation elevates fitness no matter the identity or number of partners. Furthermore, we found no benefits of partner diversity, with the fitness of cholla 1, 2, and 3 partners roughly equivalent (Figure 6 a). The 1- and 2-partner scenarios were statistically consistent but there was a modest reduction in fitness in the 3-partner scenario compared to 2 partners ($Pr(\Delta\lambda_S^{2-3} = 0.95)$).

Partner identity and composition were not strongly consequential for plant fitness. Among the one-partner scenarios, there was no strong evidence for any single best partner species (Figure 6a). While *L. apiculatum*-tended plants had the highest mean fitness, it was not significantly higher than *C. opuntiae*- ($Pr(\Delta\lambda_S^{L-C} = 0.7)$) or Other- ($Pr(\Delta\lambda_S^{L-O} = 0.69)$) tended plants. Furthermore, the fitness of *L. apiculatum*-tended plants was consistent with all 2-partner scenarios ($Pr(\Delta\lambda_S)$ ranged from 0.37 to 0.83). However, as above, there was evidence for an opportunity cost of diversity wherein fitness of the three-partner scenario was lower than any of the 2-partner scenarios ($Pr(\Delta\lambda_S) > 0.95$). The lack of diversity benefits are not driven by the high overall frequency of *L. apiculatum*. Using the simulations where all ants had equal frequencies across sizes (further explained and analyzed in Appendix C), we found the same fitness patterns as in the competitive exclusion model described above.

We found no evidence that the portfolio effect generated positive effects of partner diversity, as effects of partner richness and composition were highly consistent between the baseline model and the model version that synchronized all ant states (Figure 6). The effect of all ant partners can be

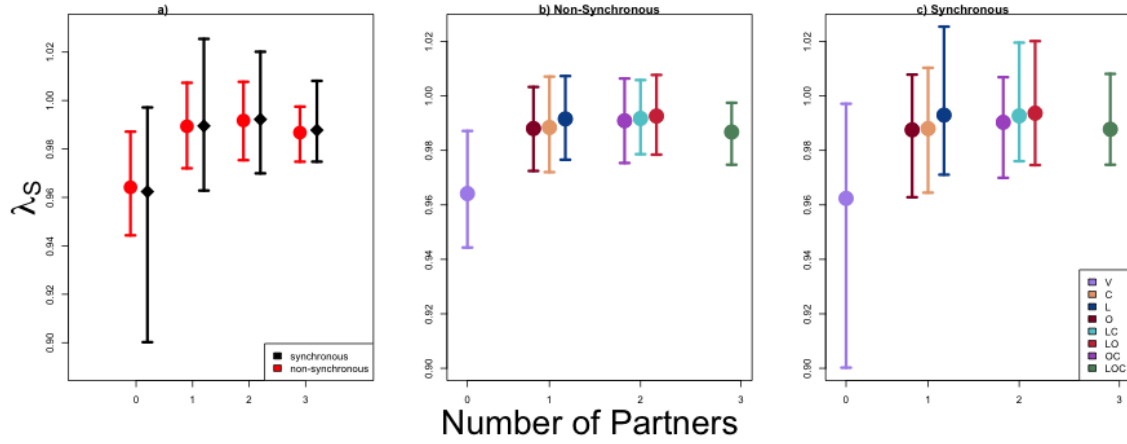


Figure 6: Panel a) the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated λ values. Panels b-c) the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

measured as $\lambda_S^3 - \lambda_S^0$. We have high confidence that this contrast is positive, and equally so for the synchronous and non-synchronous scenarios (Figure S5).

Discussion

Mutualisms commonly involve multiple partners but the ecological consequences of partner diversity remain poorly understood. Here we show that while alternative partners may be ecologically different, their fitness effects on a shared mutualist can be effectively equivalent and interchangeable. The results of our hierarchical models revealed that different ant partners had subtly different effects on vital rates, with *C. opuntiae* tended plants experiencing slight advantages in growth and survival when small, and *L. apiculatum*-tended plants experiencing floral viability advantages. The results of our synchronous and non-synchronous stochastic IPM revealed that all scenarios which included any partners resulted in the highest possible fitness for tree cholla, suggesting that while ant visitation is beneficial, partner identity and diversity are inconsequential in this system. These results highlight that partners can differ while still resulting in the same fitness benefits for a focal mutualist.

The field of multi-species mutualisms is rich with many studies highlighting unique results. Very few studies have quantified the effects of partner diversity on focal mutualist fitness. The vast majority of multi-species mutualism studies have found that partner diversity is beneficial through complementarity (Yang et al., 2024; Hernandez et al., 2020; Larimer et al., 2014; Afkhami and Stinchcombe, 2016; Fehling, 2022; Gustafson and Casper, 2005; McKeon et al., 2012; Palmer et al., 2010; Afkhami et al., 2021; Stachowicz and Whitlatch, 2005) or portfolio effect (Stevens et al., 2024; Thibaut et al., 2012; Rogers et al., 2014; Lázaro et al., 2022). Three studies have reported costs of diversity. Keller et al. 2018 found that having multiple partners resulted in equivalent fitness to no partners and worse fitness than having a single partner. Barrett et al. 2015 found that having multiple partners is worse than having a single partner, but better than having none. Bruna et al. 2014 found that having multiple partners resulted in worse fitness than

the best partner but had no vacancy scenario to compare. Our study revealed modest evidence of opportunity costs imposed by the full diversity scenario. The combination of a slight cost in a full diversity scenario compared to partial diversity scenarios is not unique (as addressed above), but the finding that the full diversity scenario is still better than vacancy does appear to be a unique finding. One study has previously reported a neutral effect of diversity (Ford et al., 2015), finding that vacant plants have the same fitness as plants guarded by any combination of partners. Most of the studies which quantitatively compare the fitness of individual partnered hosts with multi-partnered hosts used individual vital rates to estimate a fitness of the hosts. These measures ranged from growth rates, survival, biomass, to photosynthetic rates. Very few studies within the literature have used integrative fitness estimate approaches which consider the impacts of partners across the entire life cycle (Palmer et al., 2010; Stevens et al., 2024; Ford et al., 2015; Thibaut et al., 2012; Rogers et al., 2014; Lázaro et al., 2022). Across these papers, only three considered sequential partnering, where diversity of interactions accumulate across time (Palmer et al., 2010; Ford et al., 2015; Bruna et al., 2014). There are not enough studies calculating and comparing the fitness of hosts with individual partners to host fitness with multiple partners in sequential systems. Not only are our results entirely unique in terms of benefits and neutrality, but our study adds to the literature studying sequential multi-species mutualisms using integrative methods considering the entire life cycle. No other study has reported that there are clear benefits from interacting with any partner, combination with little importance tied to partner identity is unique. The finding that the optimal scenario includes multiple partners, but there is a slight cost to the full diversity scenario is also unique.¹ Our study offers contrasting results to other studies in sequentially partnered systems, indicating that we need more studies to quantitatively determine the fitness effects of partner diversity in sequentially partnered systems.

Due to the sequential nature of our system, partner turnover is an important aspect of the dynamics we observed. Partner turnover is likely a significant driver behind the fitness we see within Cholla populations, however the processes which drive the actual turnover frequency and

¹*this may not be the best way to describe this*

directions remain a mystery in this system. In the literature, it is clear that the frequency of partner turnover can have big effects on the fitness of the focal mutualist (Fiala et al., 1994; Horvitz and Schemske, 1986; Oliveira et al., 1999; Sachs et al., 2004). The direction of partner turnover is also important when the identity of partners impacts the quality of benefits received (Fonseca et al., 2003; Alonso, 1998; Dejean et al., 2008; Noe and Hammerstein, 1994). In our system we found that there are distinct patterns to the directions and frequency of partner turnover, but we still don't understand the mechanisms driving these turnovers. In the model presented in this paper we assumed that competitive exclusion was the driving mechanism behind the patterns we see. This assumption would indicate that ants are competing for each individual cacti, likely based on EFN quality and composition (Heil et al., 2004, 2010), and that the more common an ant is the better a competitor it is. In our simulations this means that when one ant is removed the ants which are left occupy the newly vacated cacti so that the proportion of vacant cacti never changes. We included this assumption because we believe it to be the most biologically reasonable based on the personal observations of multiple researchers. However, it is possible this assumption is incorrect, so we created two alternative simulations to determine how robust our results were to various assumptions about partner turnover drivers. The first alternative we tested is a model that assumes all ants have the same competitive ability rather than the differences in competitive ability we assume cause observed variety in ant frequency. In this hypothetical simulation all ants have equal probability of occupying any given cactus, but the level of vacancy across the population remains consistent with observed data. This simulation was designed to test if the overwhelming frequency of *L. apiculatum* may be overriding any potential advantages of other partners. The results of this simulation (Figure S8) were effectively the same as the competitive exclusion model presented in the body of this paper. The second alternative we tested was a model that assumed rather than competition driving ant occupations, it was the population size of each ant species alone. In this hypothetical simulation when one ant species is removed from the model the cacti are left vacant. We believe this is an unlikely scenario as there appear to be vast populations

of the ant species in the area.² This simulation was designed to test if the fitness was more
sensitive to changes in vacancy across the population than the frequency of any particular ant
species. The results of this simulation revealed that when vacancy increased sampling effect is at
play (Figure S7). This result highlights that the primary driver behind the results we presented
are ant presence and that it is only in cases of unrealistically high vacancy that there are benefits
to partner diversity. Further information about both of these models can be found in Appendix C.
Together, these results indicate that our assertions that the fitness of the cacti are more sensitive
to changes in vacancy than partner identity are robust across multiple assumptions.

We speculate that the reason our results are more sensitive to changes in vacancy than any
individual partner is due to the vital rate sensitivity structure of this population. Like other long-
lived, iteroparous species (Franco and Silvertown, 2004), tree cholla fitness is most sensitive to the
growth and survival of established individuals (Miller et al., 2009; Elder and Miller, 2016), which
are virtually guaranteed many years of reproductive opportunities once they reach a size that is
protected from mortality. Differences between alternative partners were most pronounced either
in reproductive rates, which contribute relatively weakly to fitness (because a long reproductive
lifespan overrides individual reproductive bouts), or in growth and survival at small sizes, where
mortality risk was high and the probability of ant tending was low. At larger sizes with low
mortality risk, ant-tended plants had a modest survival advantage over vacant plants regardless of
partner identity, and that result dominates our integrative measures of fitness because of the high
sensitivity of λ established plant survival. The very real demographic differences between partner
species in other vital rates and at other sizes do not register nearly as strongly in the currency of
fitness, as seen in other studies (Ford et al., 2015).

Our work explicitly incorporated temporal environmental stochasticity, which raises the opportunity for portfolio effect as a mechanism of diversity benefits. Yet, we find no evidence of portfolio effect within our system. Herbivory is an important driver of the fitness of cholla in this system, as herbivores directly impact the growth, survival, and reproductive efforts of the cacti (Miller et al.,

² *This is just personal observation so it might be best to remove this?*

2009) (Figure 6). We have not explicitly incorporated herbivory but rather assume that herbivory can be captured in the effects they have on growth, survival, and reproductive effort observations.

603 As herbivore populations fluctuate across time, the vital rate level benefits offered by different ant partners also fluctuate across time (Figure 4). When partners exhibit different reactions to varying environments, interacting with multiple partners can lead to more consistent benefits across time (Batstone, 2018). We did not directly measure the reactions of ant partners to temporal environmental stochasticity, rather we measure how the effects of each partner on cholla demography fluctuate through time, and whether those fluctuations are correlated. The vital rate results led us to believe that portfolio effect may be at play due to the differences in ant responses to annual variation. However, there is no stabilization of fitness across temporal heterogeneity and therefore no benefits of diversity due to the portfolio effect in our system. Other portfolio effect studies (Lázaro et al., 2022; Tornos-Estupiña et al., 2023; Thibaut et al., 2012; Rogers et al., 2014) found that asynchrony was linked to weak evidence of portfolio effect. One study Dallas and Kramer (2022) found that while portfolio effect was easy to show in theoretical models, it is often very weak or nonexistent in empirical data across many systems. This indicates that it may be very difficult to detect, disguised by different mechanisms, or uncommon in nature. ³

This paper shows the importance of long-term datasets in investigating species interactions and calls for further use of long-term data. Previously studies have analyzed how partner identity and partner turnover impact focal mutualist fitness (Fonseca et al., 2003; Dejean et al., 2008; Noe and Hammerstein, 1994; Barrett et al., 2015; Bruna et al., 2014; Trøjelsgaard et al., 2015). Separate studies have analyzed how inter-annual variability impacts focal mutualists (Alonso, 1998; Alarcón et al., 2008; Ollerton et al., 2006; Horvitz and Schemske, 1990; Lázaro et al., 2022). The long term data set we used gave us the unique ability to consider the combined effects of partner identity, partner turnover, and temporal stochasticity. By piecing together complete life cycle information from long-term data, we gain a more nuanced understanding of the fitness consequences of specific demographic effects. For example, our previous study suggested that *C. opuntiae* has overall

³*not sure if I need a better transition here*

627 parasitic fitness effects because activity of this species within tree cholla flowers can deter pollinators
and reduce seed set (Ohm and Miller, 2014). Yet, the more complete analysis presented here, which
accounts for reduced seed set alongside other demographic advantages, indicates that this species
630 is clearly a mutualist and nearly as strong a mutualist as *L. apiculatum*.

As with any study, there are limitations to consider when interpreting the results reported here.
These results highlighted in the paper are based on observational data regarding ant effects on plant
633 demography rather than experimental data, meaning we are able to determine correlations but not
causation. However, we have previously conducted experimental manipulations which revealed
that ant presence has direct impacts on plant demography in our system (Miller, 2007). This
636 combination of observational data backed up by experimental results gives us greater confidence in
our causal interpretations than if we had only observational data. Specifically, we are confident in
claiming that ant presence causes increased performance in cactus vital rates and the fitness signals
639 we are detecting are directly related to ant partner presence. One further expansion could include
nectar analysis in conjunction with ant interactions to look not only at the ant partner impacts on
the plant demography, but also at the ability of plants to attract specific partners through nectar
642 composition shifts.

This study highlights that while partners within a mutualistic guild can be ecologically differ-
ent, they may still be interchangeable in their net effects on fitness of a shared mutualist. The
645 individual estimation of each vital rate and explicit inclusion of ant partners in this IPM allowed us
to determine that partner diversity is not beneficial in this system and may highlight circumstances
under which we would expect similar results in other systems. The individual vital rate estimation
648 approach allowed us to determine that while different ant partners did impact each vital rate differ-
ently, the largest differences were estimated to occur at sizes when ants would not be present in real
life or in vital rates that do not have a significant impact on the overall fitness. Considering partner
651 impacts on vital rates across the entire life cycle of the plant allows more power in determining the
importance of these differences. Based on our results we would expect to see that partner diversity
does not matter when partner impacts occur during low sensitivity life stages.

Acknowledgments

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Data and Code Availability

The data that support the findings of this study are openly available as a data package on the Environmental Data Initiative website, package id: knb-lter-sev.323.1

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Tables

Appendix A: Additional Methods and Parameters

867 In addition to the models described in the body of the paper, we fit several other models using data from previous studies. These models are described below.

Seeds Per Fruit. With data from Miller et al. (2006), we fit a model for the number of seeds produced by every fruit on a cholla ($\kappa(a')$) in year $t + 1$ based on the ant partner a' in year $t + 1$. We fit this model to seed data y^κ using a Negative Binomial distribution and the log link function:

$$y^\kappa \sim \text{NegativeBinomial}(\hat{\kappa}, \hat{\phi})$$

$$\hat{\phi} = \beta_0^\phi$$

The data used for this model did not include data on ants in the “other” category, so we used the data from vacant plants to parameterize seeds per flower for plants with “other” ants in the IPM.

We found that vacant plants produced the most mean seeds (147.2 per fruit), followed by *L. apiculatum* tended plants (142.4 per fruit), and finally, *C. opuntiae* tended plants (115.0 per fruit) (Figure 1).

Recruit Size Distribution. We fit this model to recruit size data y^η using a Normal distribution with the identity link function:

$$y^\eta \sim N(\hat{\eta}, \hat{\sigma})$$

where $\hat{\sigma}$ is estimated with a non-informative prior.

We found that the mean size of recruits is $\log(-2.097)m^3$ with an interquartile range from $\log(-2.173)m^3$ to $\log(-1.712)m^3$ (Figure 2).

Germination. With germination data Miller (2007), we fit two models for the probability of germinating from the first year seedbank (γ_1) or the second year seedbank (γ_2) in year $t + 1$, with

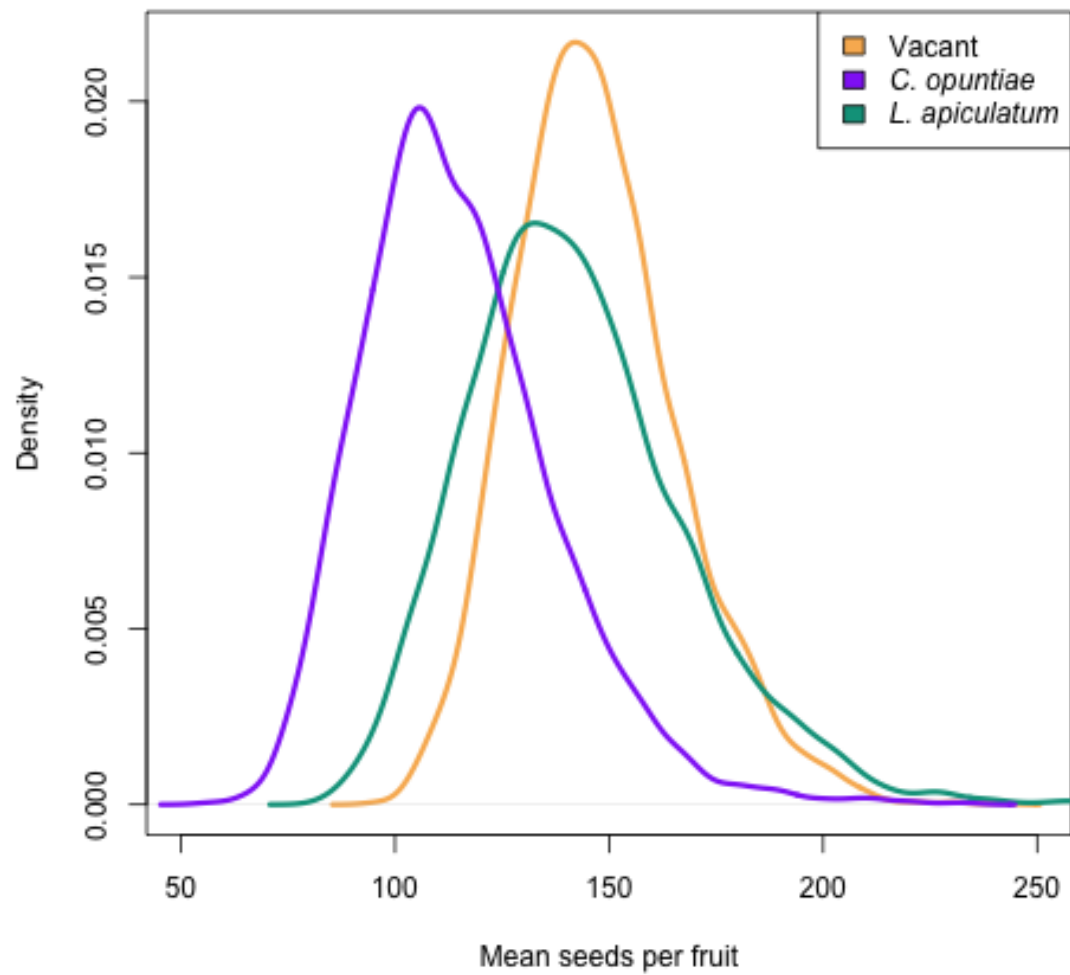


Figure 1: The estimated posterior distribution of seeds per fruit on plants tended by *C. opuntiae* (purple), *L. apiculatum* (Teal), and Vacant plants (Yellow).

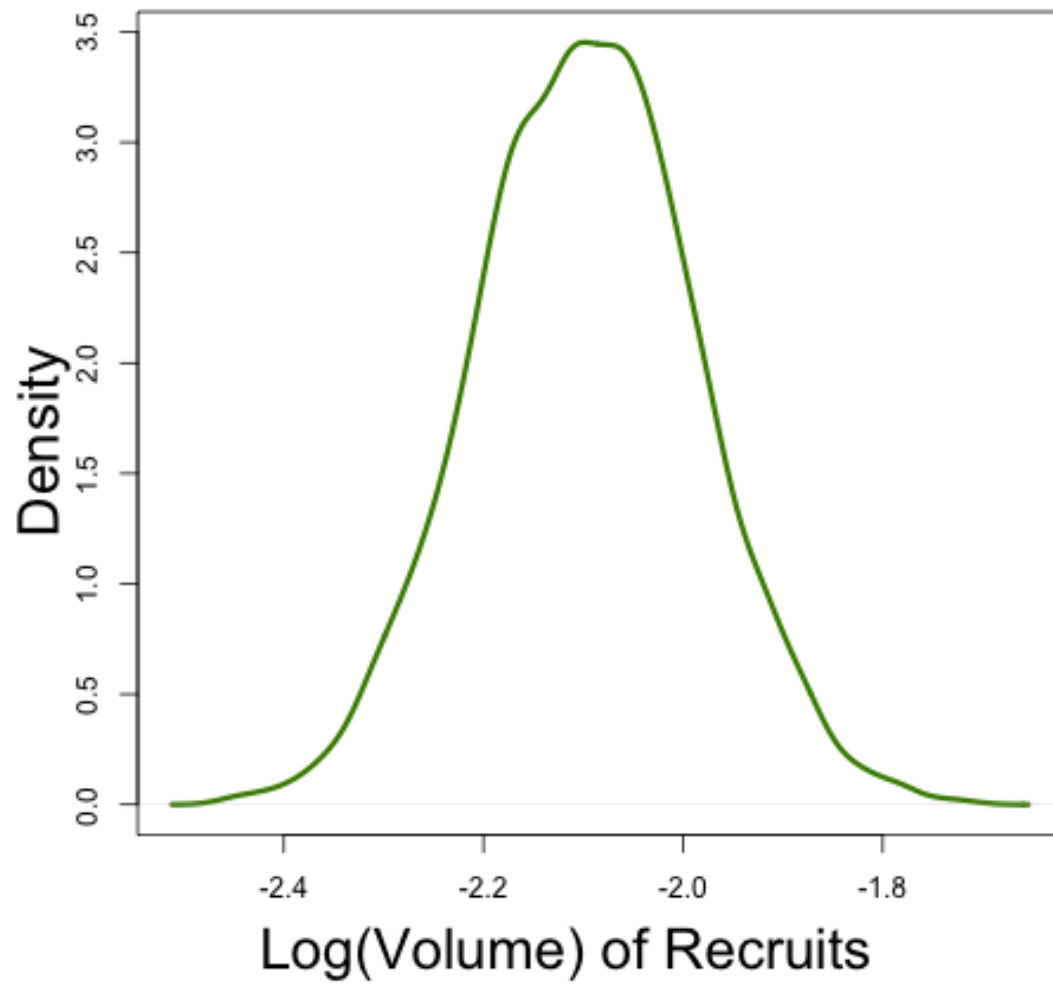


Figure 2: The estimated posterior distribution of recruit sizes as the $\log(volume)$ of cholla.

no fixed or random effects. These models were fit to germination data $y^{\gamma_1}, y^{\gamma_2}$ using the binomial distribution with logit link functions:

$$y^{\gamma_1} \sim \text{Binomial}(\hat{\gamma}_1)$$

$$y^{\gamma_2} \sim \text{Binomial}(\hat{\gamma}_2)$$

We found that the mean germination rates for seeds in the seedbank for one year is 0%, with an interquartile range of 0% and 1%. We found that the mean germination rates for seeds in the seedbank for a second year is 0%, with an interquartile range of 0% to 0.4% (Figure 3). Seeds are more likely to germinate in their first year in the seedbank, but most seeds will never germinate.

Pre-Census Survival. With recruit census data Miller et al. (2006), we fit a model for the probability of a seedling (which germinates in early Fall) surviving to when we census in May (δ) of year $t + 1$ (accounting for missed mortality events), with fixed effects of the previous size x and random effects of the transect m . We fit this model to pre-census survival data y^δ using a Bernoulli distribution with a logit link function:

$$y^\delta \sim \text{Bern}(\hat{\delta})$$

where $m \sim N(0, \sigma_{\text{transect}}^2)$ is the random effect of transect where the recruited individual was analyzed for survival.

We found that plants have a 16.2% probability of surviving from germination to the next census. Our model estimated this very well, expecting a 16.3% probability (Figure 4).

Portfolio Effect.

Appendix B: Observed Herbivory Data

Herbivory is an important driver in this population and shapes the range and demography of cholla. Herbivore presence has been shown to negatively impact growth and fecundity of cholla populations (Miller et al., 2009). Ant visitors are believed to offer defensive benefits to the plants they tend

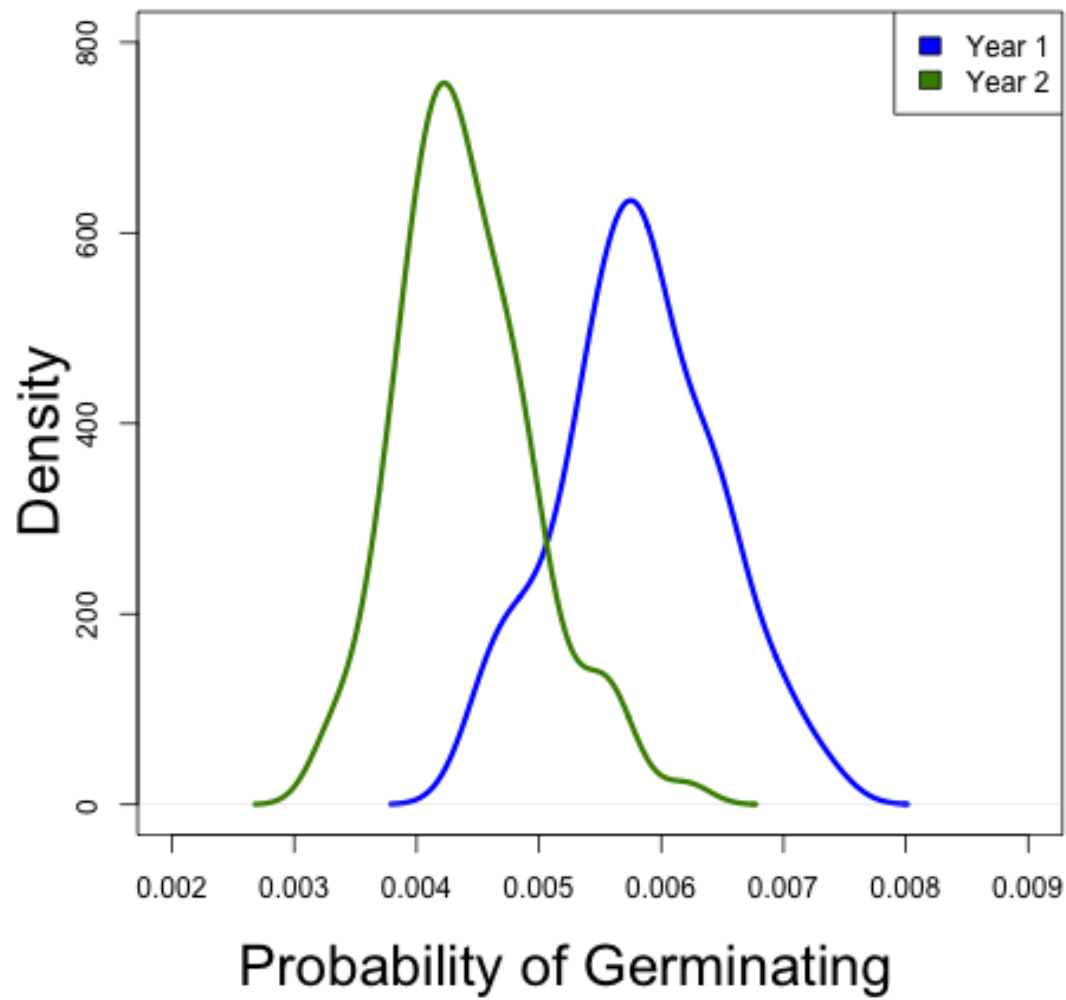


Figure 3: The estimated posterior distributions for germination probability of seeds in the first year seedbank (blue) and seeds in the second year seedbank (green).

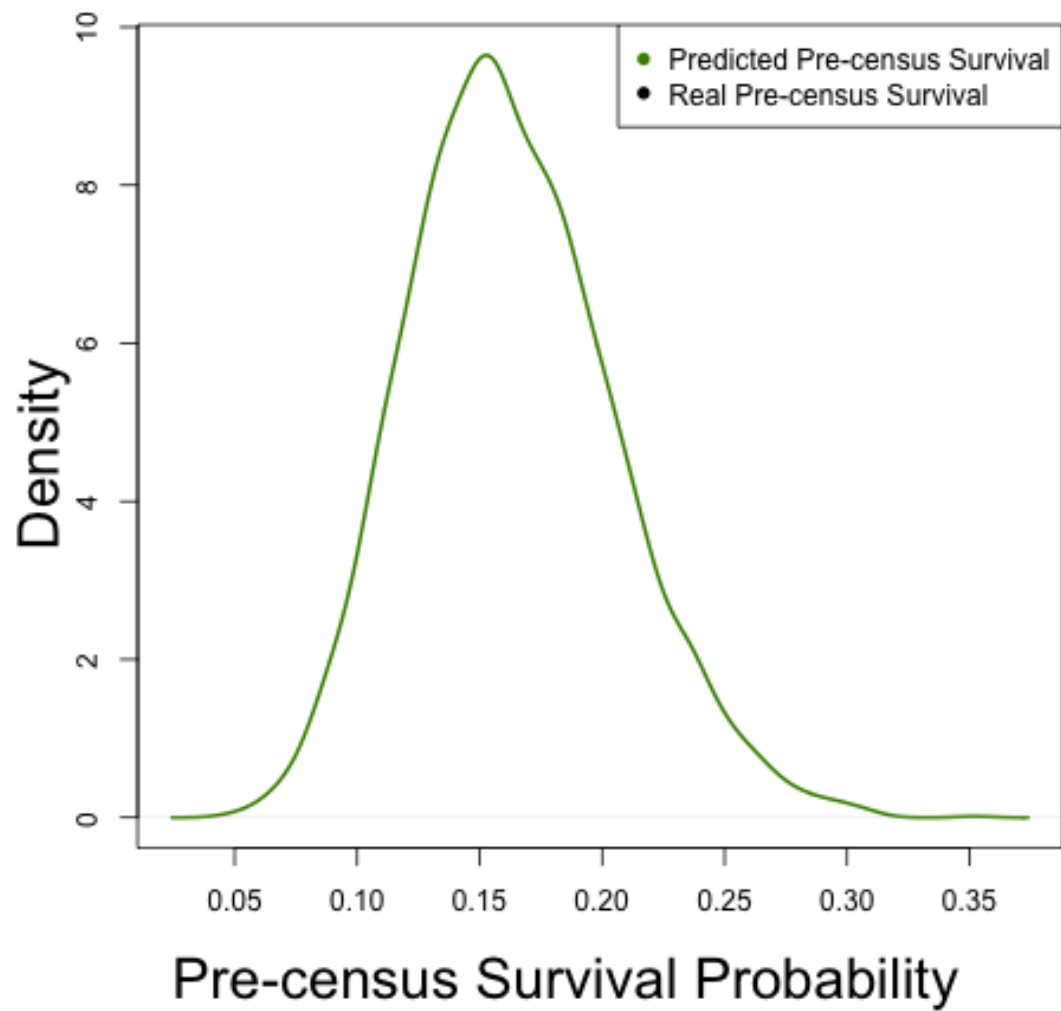


Figure 4: The distribution of the probability of plants which germinated between censuses surviving to the next census to be counted. The green line shows the estimated posterior distribution of pre-census survival.

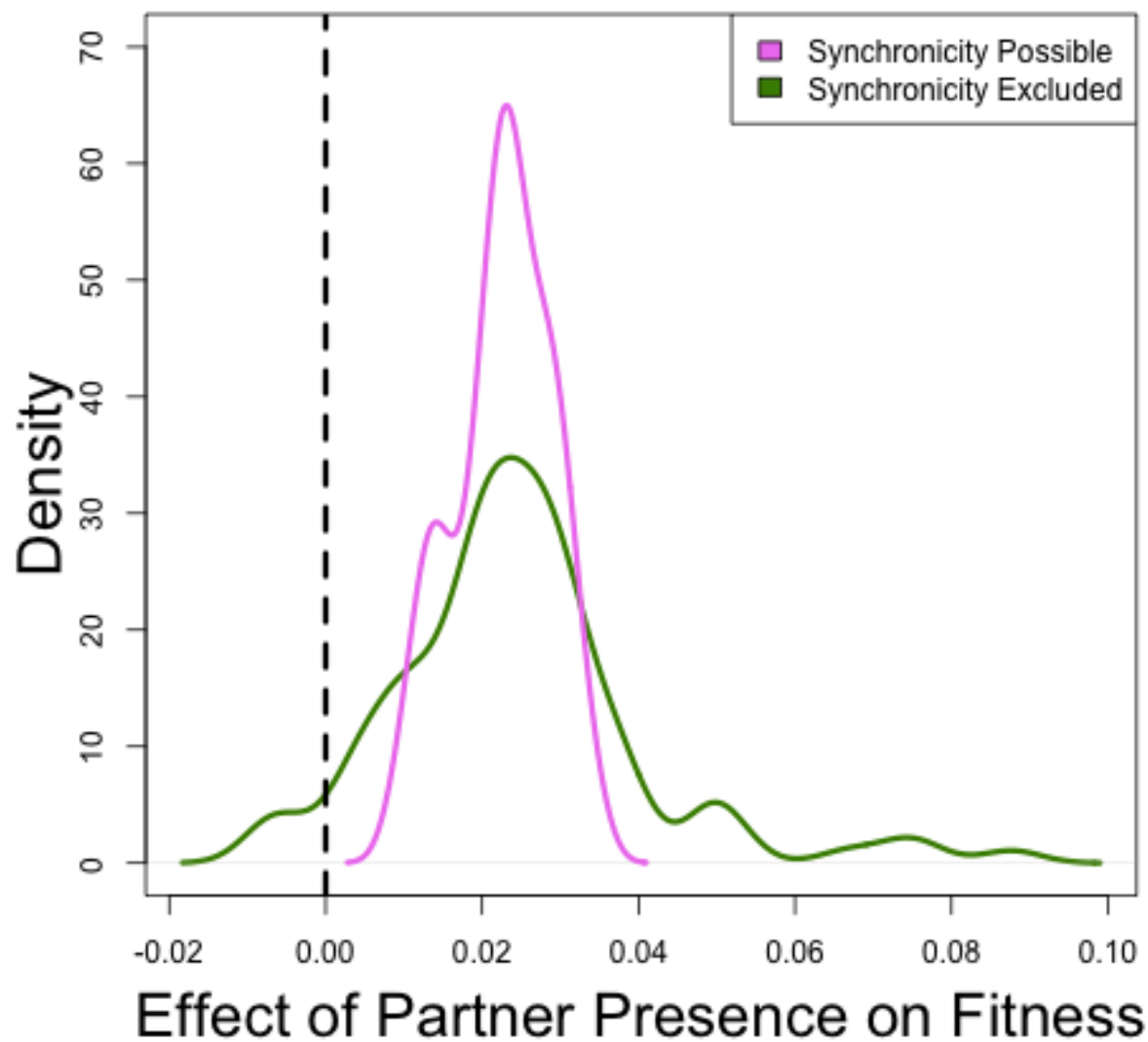


Figure 5: The distribution of $\lambda_{NS,All} - \lambda_{NS,Vacant}$ in pink and $\lambda_{S,All} - \lambda_{S,Vacant}$ in green. The vertical dashed line shows where the effect of partners on the fitness of the cholla is 0 (to the left the partners have a negative effect, to the right the partners are beneficial).

in this system, leading to the hypothesis that ant presence would be correlated with reduced herbivory. Herbivory data was collected during censuses any time herbivores were identified on a plant. This involved noting the type and quantity of herbivores observed. This data has been taken consistently since 2017, so the analysis below considers 6 years of data. We considered only plants which were reproducing, as they were likely to produce the highest quantity of EFN (Miller, 2014). The proportion of reproducing plants that experienced herbivory was calculated for each ant state separately. Analysis showed that ant presence is correlated with lower herbivore visitation. 40% of vacant cacti experienced herbivory. Plants tended by Other ants experienced similar, though lower, levels of herbivory on reproducing plants, with herbivores detected on 37.5% of plants. Herbivores were detected on 25% of plants tended by *C. opuntiae* ants and on 11% of plants tended by *L. apiculatum* ants. These results indicate that ant presence is correlated with lower levels of herbivory and that partner identity has an impact on the level of herbivory. They also indicate that the partner correlated with the lowest levels of herbivory is *L. apiculatum* ants. These findings are consistent with literature findings which show that *L. apiculatum* ants are the most aggressive (therefore the most effective against herbivores), but differ from previous findings that *C. opuntiae* may not offer defensive benefits (Miller, 2007).

Appendix C: Alternative Ant Transition Simulations

In addition to the competitive exclusion model defined and analyzed in the body of the paper, we simulated results from several other potential models. We chose to include competitive exclusion as our primary results in the paper because we believe it to be the most biologically realistic. However, in building and testing of alternative models we found that the method of ants occupying plants significantly impacts the fitness of the population. We tested two alternative transition models, one called the frequency based model and one called the equal likelihood model.

Frequency Based Model. The first alternative hypothesis we tested was what we called the frequency based model. In this model rather than the proportion of vacant cacti being maintained, the

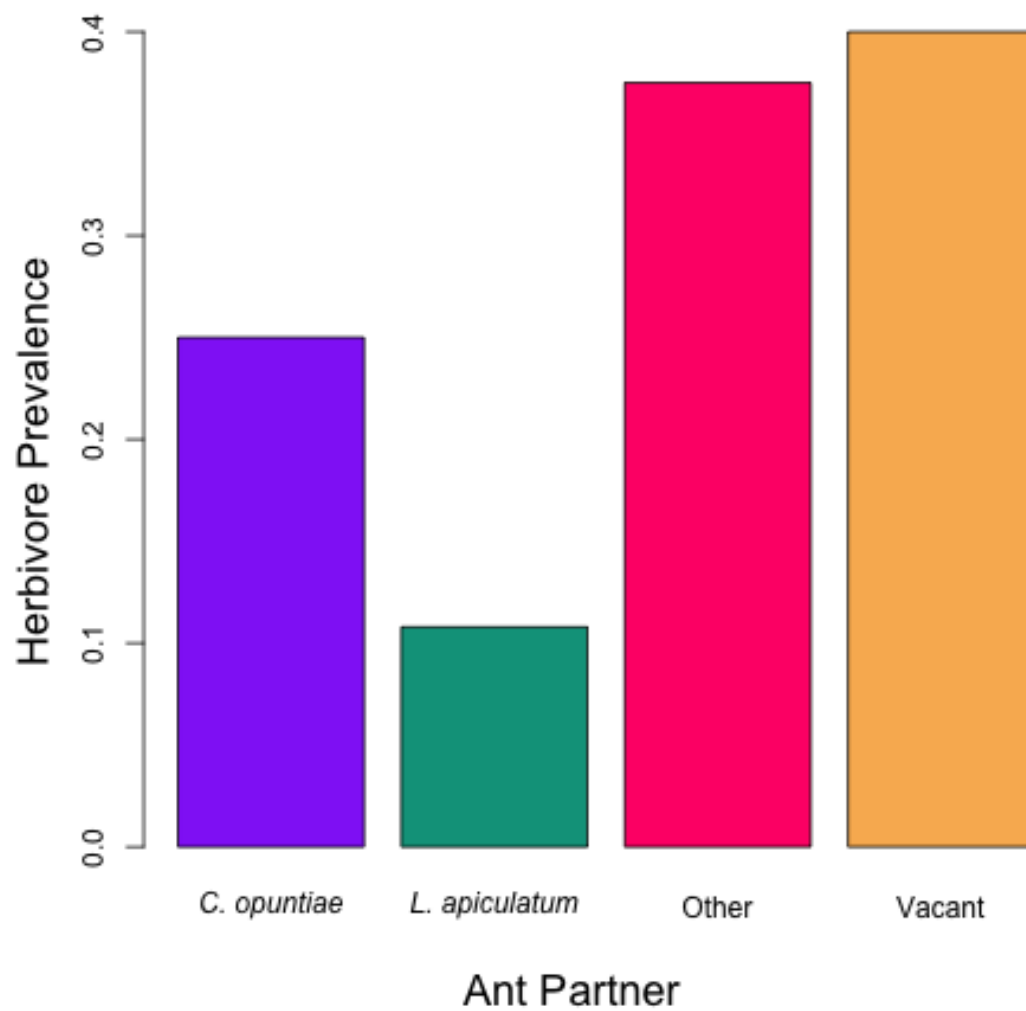


Figure 6: The proportion of reproducing plants which are visited by herbivores. Each bar represents the subset of the cacti population in a different ant state.

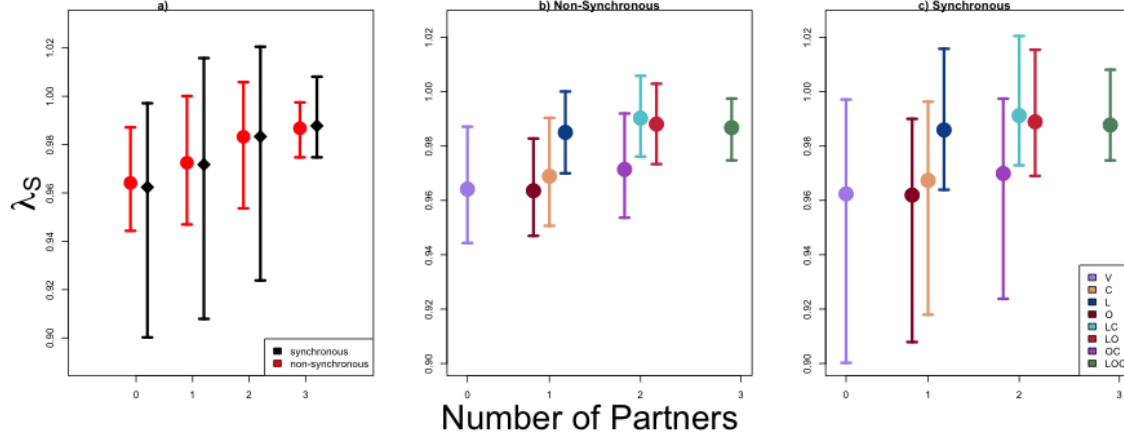


Figure 7: Panel a) the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated λ values. Panels b-c) the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

proportion of cacti occupied by each species is maintained and when one is removed it is replaced with vacancy. This version of the model assumes that the frequency of each ant we see is reflective of the real frequency of populations rather than some other mechanism. With this model we found very clear evidence of Sampling Effect in the system. When only *C. opuntiae*, Other ants, or both ants are present, there is very little difference in the fitness of the cacti from when no partners are present. Only when *L. apiculatum* ants are present do we see an increase in the fitness of the focal mutualist (Figure 7a). In this simulation, the more partners that are present the higher the fitness of the focal mutualist is, confirming that partner diversity would be beneficial through sampling effect if this transition model were correct. (Figure 7b).

Equal Likelihood Model. The second alternative hypothesis we tested was what we called the equal likelihood model. In this model we preserved the observed pattern of size-dependent vacancy/occupancy, but occupancy was manipulated to be equally likely for all partner identities.

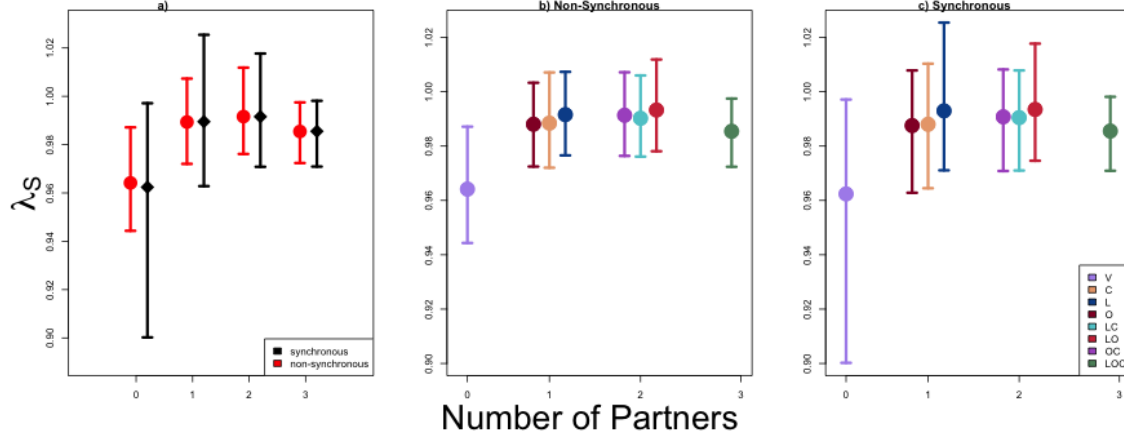


Figure 8: Panel a) the mean estimated λ_{NS} and λ_S for different numbers of partners (0-3) for the synchronous IPM (black circle) and the non-synchronous IPM (red diamond). The lines show the posterior distribution spread of the estimated λ values. Panels b-c) the mean estimated λ_{NS} and λ_S respectively, for each simulated combination of ant partner as the filled in circles. The lines show the posterior distribution spread of estimated λ values. The letters in the legend correspond to what ant partners are present (V = Vacant, C = *C. opuntiae*, L = *L. apiculatum*, O = other).

This was designed to remove the effect overwhelming numbers of *L. apiculatum* ants may have. Despite very different proportions, we found very similar outcomes to the competitive exclusion model analyzed in the paper. All ants are beneficial, but having more than one is not necessarily any better than having an individual species as a partner (Figure 8b). Partner presence is beneficial, but neither identity nor number of partners appears to be important (Figure 8a).

Appendix D: Posterior Checks and Model Validation

For each model fitted, we conducted two tests to determine if the fit was acceptable to use in our IPM. First, we checked the convergence of each parameter. Below we show the convergence of all β terms listed in the Statistical Modeling subsection of Methods. Second, we checked the posterior fit, comparing the estimated values of each model to the y values of the actual data. We show these posterior checks below, split by ant partner where relevant. Each model was run with 10,000

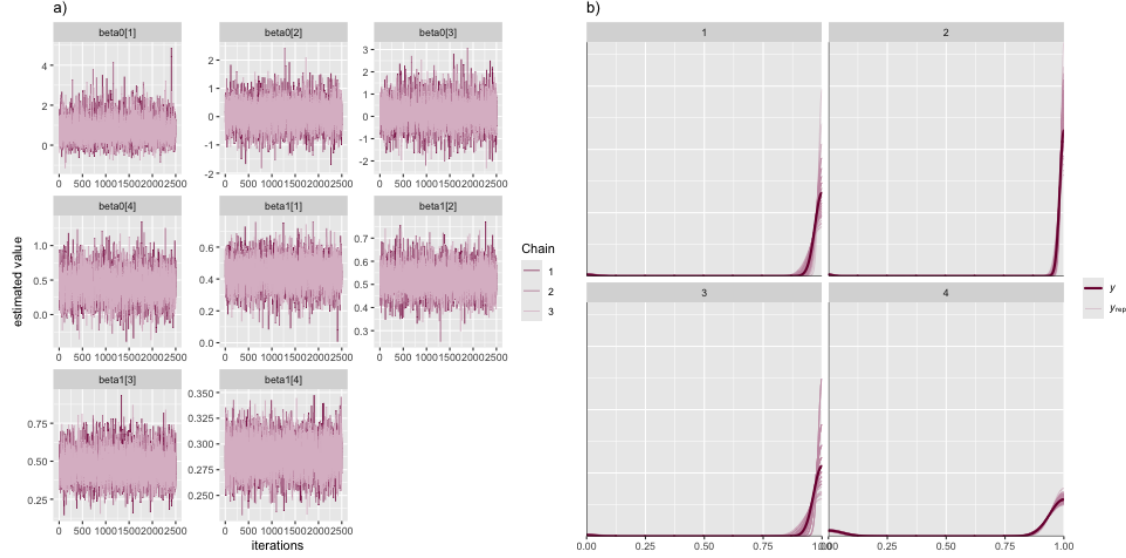


Figure 9: The a) posterior convergence of the parameters estimated by the survival model and the b) posterior distribution of survival estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean survival distribution (black line) of the real data.

total iterations per chain, 5,000 warmups, and thinning set to 2, leaving us with a total of 2,500
 939 iterations post warmup.

Statistical Models – Results

Below are the results reported of all statistical models not described in the main body of the text.

942 *Reproduction Model.* The probability of a plant reproducing in a given year is highly size dependent. The mean probability of reproducing remains at about 0% until the plant reaches a medium size, after which the mean probability of reproducing increases steadily before reaching about 100%
 945 at large sizes.

Seeds Per Flower Model. Each viable flower on a plant produces between 97 and 257 seeds. This number is affected by the ant partner present, as shown in previous work (Ohm and Miller, 2014).

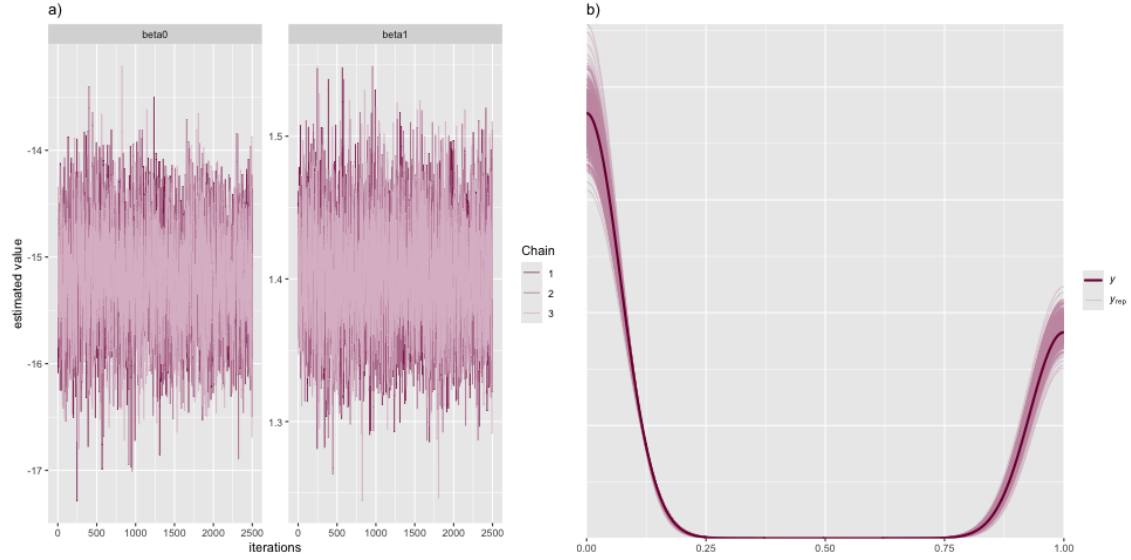


Figure 10: The a) posterior convergence of the parameters estimated by the reproduction model and the b) posterior distribution of reproductive status estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean reproductive status distribution (black line) of the real data.

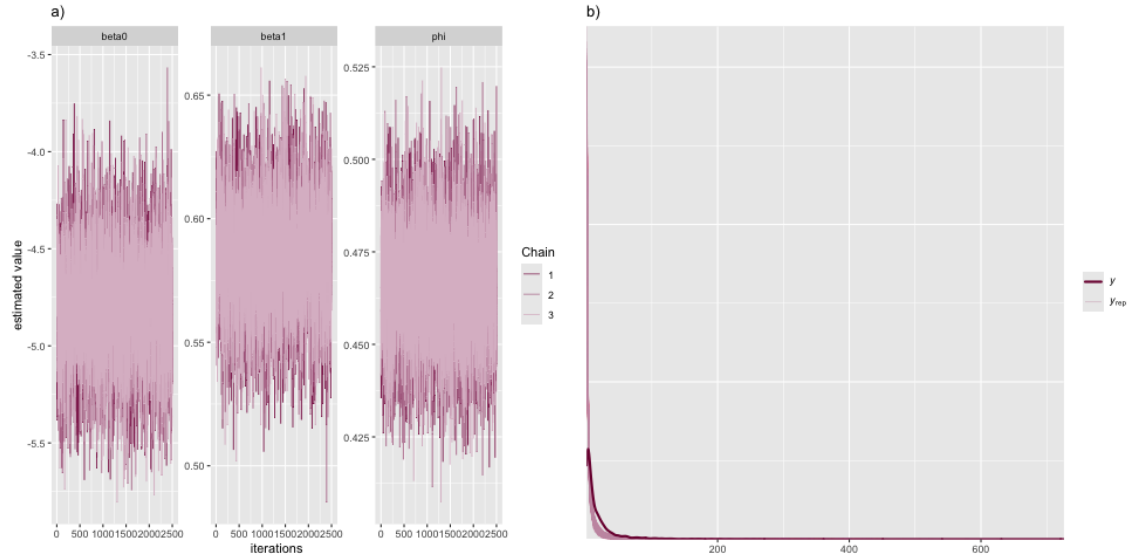


Figure 11: The a) posterior convergence of the parameters estimated by the number of flowers model and the b) posterior distribution of the number of flowers estimated (pink lines) compared to the mean distribution of observed flowers (black line).

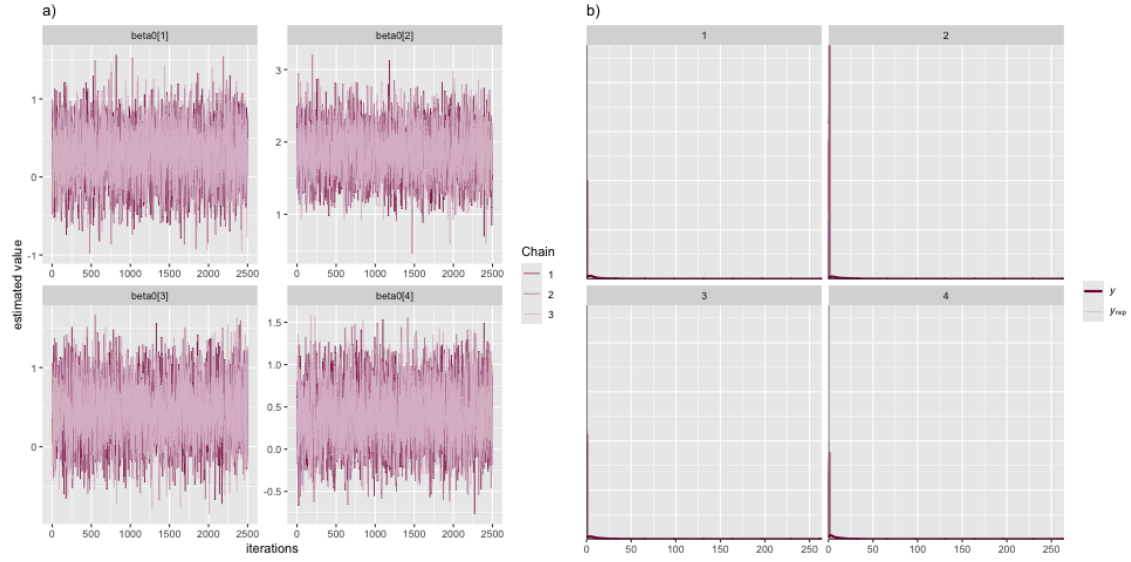


Figure 12: The a) posterior convergence of the parameters estimated by the viability model and the b) posterior distributions of floral viability estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = other, 4 = vacant) compared to the mean floral viability distribution (black line) of the real data.

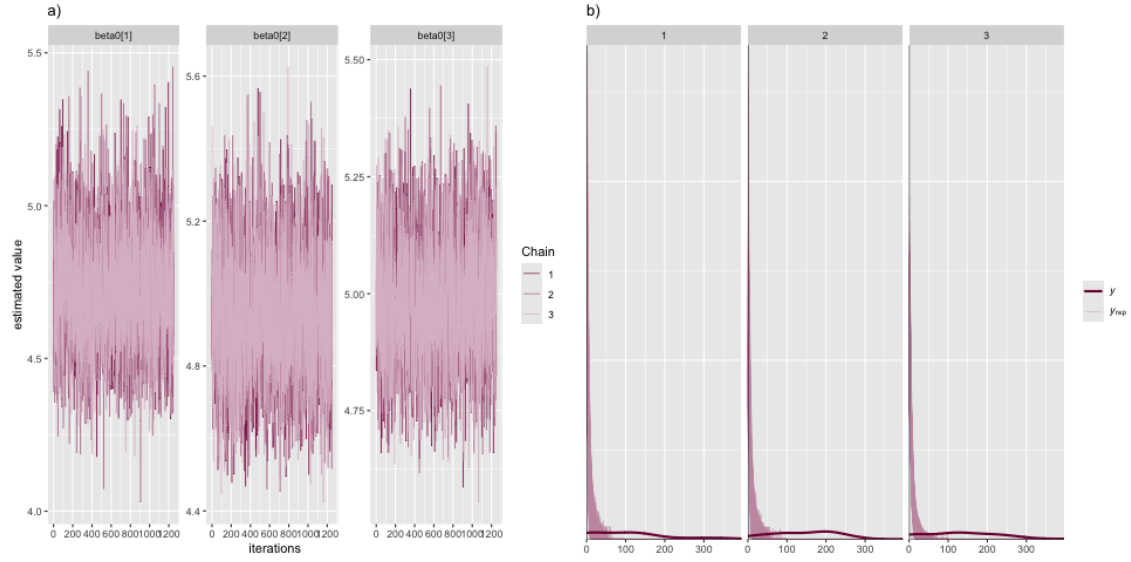


Figure 13: The a) posterior convergence of the parameters estimated by the seeds per fruit model and the b) posterior distributions of seeds per fruit estimates (pink lines) for each ant species (1 = *C. opuntiae*, 2 = *L. apiculatum*, 3 = vacant) compared to the mean seeds per fruit distribution (black line) of the real data.

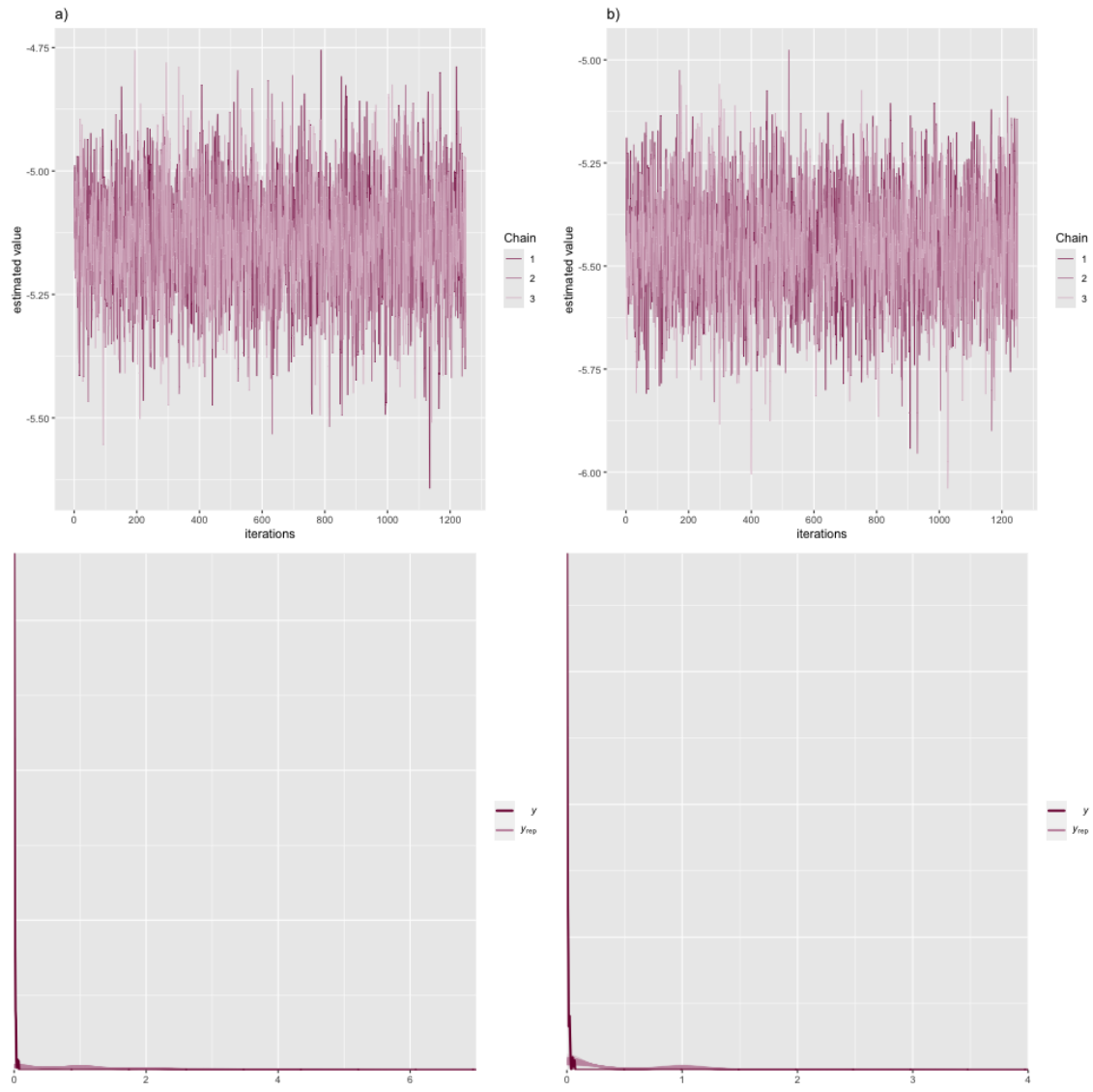


Figure 14: The a-b) posterior convergence of the parameters estimated by the germination from year one seedbank and germination from year two seedbank models respectively. The c-d) posterior distributions of floral viability estimates (pink lines) compared to the mean germination distribution (black line) of the real data for first year germinants and second year germinants respectively.

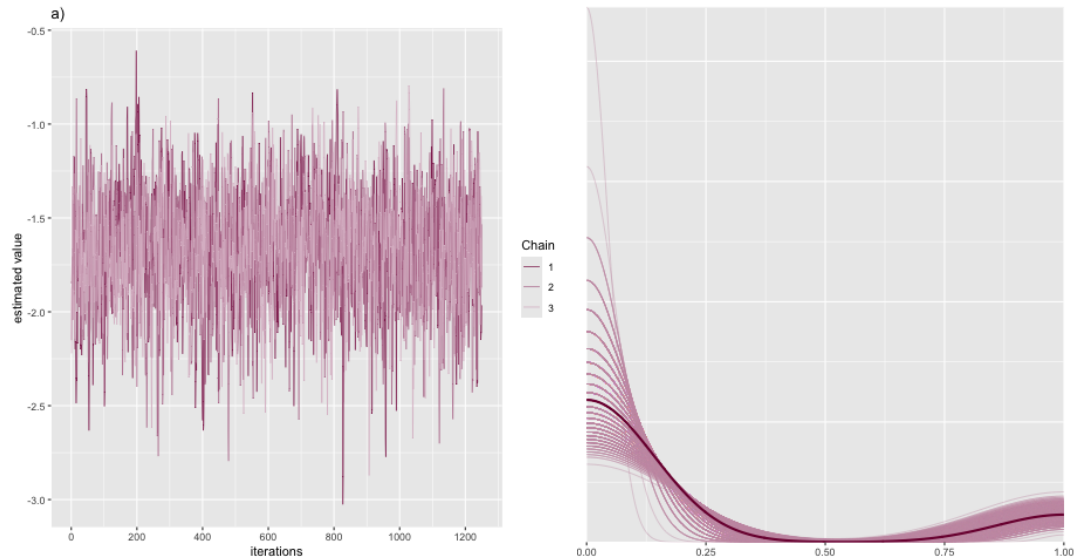


Figure 15: The a) posterior convergence of the parameters estimated by the pre-census survival model and the b) posterior distribution of the pre-census survival estimated (pink lines) compared to the mean distribution of observed pre-census survival (black line).

C. opuntiae tended plants produce a mean of 115 seeds per flower. *L. apiculatum* tended plants produce a mean of 143 seeds per flower. Vacant plants produce a mean of 148 seeds per flower. Comparison between posterior distributions revealed that *C. opuntiae* tended plants produced fewer seeds per flower than *L. apiculatum* tended plants and vacant plants 80% and 87% of the time. Vacant plants produced more seeds per flower than *L. apiculatum* tended plants only 57% of the time. We are confident that *C. opuntiae* tended plants produce the fewest seeds per flowers.

Pre-census Survival Model. Pre-census seed survival rates fall between 0% and 95% with the mean pre-census seed survival at 18%.

Germination Model. Seeds have a significantly higher probability of germinating in year one than in year two. Seeds in year one experience germination rates between 50% and 100% with a mean of 62% germination. Seeds in year two experience germination rates between 50% and 98% with a mean of 58% germination.

960 New recruits are expected to be between the sizes of 0.11 cm^3 and 0.38 cm^3 with a mean size of 0.20 cm^3 .