STAT 400, Aryn Harmon 1

Midterm 1 Material 1.1

Probability is a real-valued function: 1. $\mathbb{P}(S) = 1$; 2. $\mathbb{P}(A) \geq 0$; 3. If A_1, A_2 are mutually exclusive events, $\mathbb{P}(A_1 \cup A_2) = \mathbb{P}(A_1) +$ $\mathbb{P}(A_2)$ and so on.

Inclusion-Exclusion: $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$. Conditional Probability: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$ given $\mathbb{P}(B) > 0$. $\mathbb{P}(A|B) \neq \mathbb{P}(A)$ unless A and B are independent. Also see Bayes's

Theorem. Probability of a string of unions given B is equal to the sum of the individual conditional probabilities.

Multiplication Rule: Probability of two events both occurring: $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$ or $\mathbb{P}(A \cap B) = \mathbb{P}(B) \cdot \mathbb{P}(A|B)$ (one is easier than the other).

Bayes's Rule: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A) \cdot \mathbb{P}(B|A)}{\mathbb{P}(B)}$. $\mathbb{P}(A)$ is the *prior proba*bility of A. $\mathbb{P}(A|B)$ is the posterior probability of A given that B occurred. Use to invert probabilities.

Bayes's Rule 2: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A) \cdot \mathbb{P}(B|A)}{\mathbb{P}(A) \cdot \mathbb{P}(B|A) + \mathbb{P}(A') \cdot \mathbb{P}(B|A')}$.

Bayes's Rule Full: Given some partition of $S: A_1 \cup \cdots \cup A_k = S$. $\mathbb{P}(A_i|B) = \frac{\mathbb{P}(A_i) \cdot \mathbb{P}(B|A_i)}{\sum_{m=1}^k \mathbb{P}(A_m) \cdot \mathbb{P}(B|A_m)}.$

Independence: $\mathbb{P}(A|B) = \mathbb{P}(A)$, $\mathbb{P}(B|A) = \mathbb{P}(B)$, and $\mathbb{P}(A \cap B)$ $(B) = \mathbb{P}(A) \cdot \mathbb{P}(B).$

Pairwise Independence: All of the following must hold: $\mathbb{P}(A \cap$ $(B) = \mathbb{P}(A) \cdot \mathbb{P}(B), \ \mathbb{P}(A \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(C), \ \text{and} \ \mathbb{P}(B \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(C)$ $\mathbb{P}(B) \cdot \mathbb{P}(C)$.

Mutual Independence: A, B, and C must be pairwise independent and in addition: $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C)$.

Number of ways to take r objects from n candidates:

Ordered Sample Unordered Sample

W. R.
$$n^r$$
 $\binom{n+r-1}{r}$ W/o. R. nP_r $\binom{n}{r}$

 ${}^{n}P_{k} = \frac{n!}{(n-k)!}$ - permutation

 $\binom{n}{k} = {}^nC_k = \frac{n!}{k!(n-k)!}$ - combination

Binomial Theorem: $(a+b)^n = \sum_{r=0}^n \binom{n}{r} a^r b^{n-r}$. $\binom{n}{n-r} = \binom{n}{r} a^r b^{n-r}$

Some magic formulae: 1. $\sum_{r=0}^{n} \binom{n}{r} = 2^{n}$. 2. $\sum_{r=0}^{n} (-1)^{r} \binom{n}{r} = 0$. 3. $\sum_{r=0}^{n} {n \choose r} p^r (1-p)^{n-r} = 1$.

Pascal's Equation: $\binom{n}{r} = \binom{n-1}{r} + \binom{n-1}{r-1}$. Hypergeometric Distribution: $f(x) = \mathbb{P}(X = x) = x$

 $\frac{\binom{N_1}{x} \cdot \binom{N_2}{n-x}}{\binom{N}{n}}$, where $x \leq n, x \leq N_1, n-x \leq N_2$ and $N = N_1 + N_2$. $\mathbb{E}(X) = n \frac{N_1}{N}$ and $Var(X) = n \cdot \frac{N_1}{N} \cdot \frac{N_2}{N} \cdot \frac{N-n}{N-1}$. Example: Urn

model: N_1 red balls, N_2 blue balls, draw n balls from $N_1 + N_2$ balls, then look at the probability that there are x red balls in the selected n balls.

Mean: $\mu = \mathbb{E}(X) = \sum_{x \in S} x f(x)$. Variance: $\sigma^2 = Var(X) = \mathbb{E}(X - \mu)^2 = \mathbb{E}(X^2) - \mu^2 = \mathbb{E}(X^2)$

 $\sum_{i=1}^{k} x_i^2 f(x_i) - \mu^2$. Standard Deviation: σ .

r-th Moment: $\mathbb{E}(|X|^r) = \sum_{x \in S} |x|^r f(x) < \infty$; is the moment about the origin.

r-th Moment about b: $\mathbb{E}((X-b)^r) = \sum_{x \in S} (x-b)^r f(x)$; is the moment about b. Facts: μ is the first moment of X about the origin. σ^2 is the second moment of X about μ .

Example of Variance Properties: $Var(aX + b) = a^2 \cdot Var(X)$. Bernoulli Distribution: A random experiment is called a set of Bernoulli trials if each trial has only two outcomes, has a constant p, and each trial is independent. f(x) = p if x = 1, 1 - p if x = 0, with $0 . <math>\mathbb{E}(X) = p$ and Var(X) = p(1 - p).

Binomial Distribution: Let X be the number of successes in n independent Bernoulli trials with p. Then $X \sim b(n, p)$.

 $f(x) = \mathbb{P}(X = x) = \binom{n}{x} p^{x} (1-p)^{n-x}$. $\mathbb{E}(X) = np$ and Var(X) = np(1-p).

A note: binomial and hypergeometric are similar, but binomial has replacement and one category, and hypergeometric has two categories and no replacement.

Cumulative Distribution Function: $F(x) = \mathbb{P}(X \le x), x \in$ $(-\infty,\infty)$. For discrete random variables: $f(x) = \mathbb{P}(X=x) =$ $\mathbb{P}(X \le x) - \mathbb{P}(X \le x - 1) = F(x) - F(x - 1).$

Geometric Distribution: $f(x) = p(1-p)^{x-1}, x = 1, 2, 3, \dots$ X represents the draw in which the first success is drawn. f(x) is the probability of getting a success in the x-th draw.

Negative Binomial Distribution: $f(x) = {x-1 \choose r-1} p^r (1-p)^{x-r}$ and $x = r, r+1, r+2, \ldots X$ is the number of trials it takes to get r successes. f(x) is the probability that the r-th success occurs on the x-th trial.

1.2Midterm 2 Material

Moment Generating Function: $M(t) = \mathbb{E}(e^{tX}) =$ $\sum_{x \in S} e^{tX} f(x)$ if f(x) is the p.d.f. of some distribution and $t \in V_h(0)$ is finite. Theorem: $\mathbb{E}(X^r) = M^{(r)}(0)$, so $\mu = \mathbb{E}(X) =$ M'(0) and $\sigma^2 = \mathbb{E}(X^2) - [\mathbb{E}(X)]^2 = M''(0) - [M'(0)]^2$. To calculate the m.g.f. and p.d.f. of some random variable given only its moments, use the Taylor series expansion centered at zero. $M(t) = M(0) + M'(0) \left(\frac{t}{1!}\right) + M''(0) \left(\frac{t^2}{2!}\right) + \cdots =$

 $1 + \mathbb{E}(X)\left(\frac{t}{1!}\right) + \mathbb{E}(X^2)\left(\frac{t^2}{2!}\right) + \dots$

Poisson distribution: Definition: Poisson process counts the number of events occurring in a fixed time/space given a rate λ . Let X_t be the number events which occur in t unit time intervals. $\mathbb{P}(X_t = x) = \frac{(\lambda t)^x}{x!} e^{-\lambda t}. \ \mu = \lambda, \sigma^2 = \lambda.$

Poisson Approximation of Binomial Distribution: $X \sim b(n,p)$ and n is large while p is small, then X can be approximated as $X \sim poi(\lambda) s.t. \lambda = np$.

Mean: $\mu = \int_{-\infty}^{\infty} x f(x) dx$.

Variance: $\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$.

M.G.F.: $M(t) = \mathbb{E}(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f(x) dx, t \in (-h, h).$

Percentile: $p = \int_{-\infty}^{\pi_p} f(x) dx = F(\pi_p)$.

Uniform Distribution: $f(x) = \frac{1}{b-a}, a \le x \le b.$ $X \sim U(a,b)$, then $\mathbb{E}(X) = \frac{a+b}{2}$, $\sigma^2 = \frac{(b-a)^2}{12}$, and $M(t) = \mathbb{E}(e^{tX}) = \frac{e^{tb} - e^{ta}}{t(b-a)}, t \neq 0 \text{ and } M(0) = 1.$

Exponential Distribution: This describes the waiting time between events in a Poisson process with rate λ . Let $\theta = \frac{1}{\lambda}$. Then if $X \sim Exp(\theta)$, $f(x) = \frac{1}{\theta}e^{-\frac{x}{\theta}}$, $x \ge 0$ and 0 otherwise. $\mu = \theta$, $\sigma^2 = \theta^2$, and $M(t) = (1 - \theta t)^{-1}, t < \frac{1}{4}$.

Memoryless Property of the Exponential Distribution: What happened in the past does not matter now. Only the present can determine the future. $\mathbb{P}(X > a + b \mid X > a) = \mathbb{P}(X > b), \forall a, b \geq 0.$ Gamma Function: $\Gamma(x) = \int_0^\infty y^{x-1} e^{-y} dy, \ x > 0.$

Gamma Distribution: f(t) represents the waiting time unitl the α -th occurrence of some Poisson process with rate $f(t) = \frac{\lambda^{\alpha} t^{\alpha-1}}{(\alpha-1)!} e^{-\lambda t}, t \geq 0.$ Then let $\theta = \lambda^{-1}$, then $f(t) = \frac{1}{\Gamma(\alpha)} \theta^{\alpha} t^{\alpha-1} e^{\frac{-t}{\theta}}, t \geq 0, \alpha \in \mathbb{R}. \ \mu = \alpha \theta, \ \sigma^2 = \alpha \theta^2, \text{ and}$ $M(t) = (1 - \theta t)^{-\alpha}$. If $\alpha = 1$, gamma $(1, \theta) = \text{Exp}(\theta)$.

 χ^2 Distribution: If $X \sim \operatorname{gamma}(\alpha, \theta)$, and $\theta = 2$ and $\alpha = \frac{r}{2}$, where r is a positive integer, then X is a χ^2 distribution with degree of freedom r. $X \sim \chi^2(r)$. $\mu = r$, $\sigma^2 = 2r$, and $M(t) = (1-2t)^{\frac{-r}{2}}, t < \frac{1}{2}, e^2 = \chi^2(2).$

Normal Distribution: Bell curve! $f(x) = \frac{1}{\sqrt{2\pi}\sigma}e^{-\frac{(x-\mu)^2}{2\sigma^2}}, x \in$ \mathbb{R} . $X \sim N(\mu, \sigma^2)$. $X \sim N(0, 1)$ is the standard normal distribution. $\mu = \mu$, $\sigma^2 = \sigma^2$, and $M(t) = e^{\mu t + \frac{\sigma^2 t^2}{2}}$. Standardization: $\epsilon = k\sigma$: $\mathbb{P}(|X - \mu| \ge \epsilon) \le \frac{1}{k^2}$. $X \sim N(\mu, \sigma^2)$, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Normal Square Distribution: Let $X \sim N(\mu, \sigma^2)$ and $Z = \frac{X - \mu}{\sigma}$. Then the random variable $V = Z^2 \sim \chi^2(1)$.

Cauchy Distribution: (Why do we even need this disctribution?) $f(x) = \frac{1}{\pi(1+x^2)}, x \in \mathbb{R}$. Symmetric about zero, so median is zero, but μ is undefined because the tail of the p.d.f. is too heavy (i.e. each integral of the distribution does not converge). $c.d.f. = F(x) = \frac{1}{\pi}\arctan(x) + \frac{1}{2}, x \in \mathbb{R}.$

Joint Probability Density Function: $\mathbb{P}((X,Y) \in A) =$ $\int \int f(x,y)dxdy$.

Marginal Probability Density Function: $f_x(x)$ $\int_{-\infty}^{\infty} f(x,y)dy$ and the other way around for y.

Mathematical Expectation: $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f(x,y) dx dy.$

Independent Random Variables: Two events A and B are independent iff $f(x,y) = f_x(x) \cdot f_y(y) \, \forall x,y$. This works for both the p.d.f. and the c.d.f.

Trinomial Distribution: This is an extension of the binomial distribution into two dimensions. $f(x_1, x_2) = \mathbb{P}(X_1 = x_1, X_2 =$ $x_2) = \frac{n!}{x_1! x_2! (n - x_1 - x_2)!} p_1^{x_1} p_2^{x_2} (1 - p_1 - p_2)^{n - x_1 - x_2} \text{ for } x_1, x_2 \ge 0,$ $x_1 + x_2 \leq n$.

Covariance: $Cov(X,Y) = \mathbb{E}[(X-\mu_x)(Y-\mu_y)], \mu_x = \mathbb{E}(X), \mu_y = \mathbb{E}(X)$ $\mathbb{E}(Y)$. Useful identity: $Cov(X,Y) = \mathbb{E}(XY) - \mathbb{X}\mathbb{E}(Y)$.

Properties of Covariance: Zero-angle: Cov(X, X) = Var(X). Symmetry: Cov(X,Y) = Cov(Y,X). Cov(aX + b,Y) $a \cdot Cov(X,Y)$. Cov(X + Y,W) = Cov(X,W) + Cov(Y,W). $Cov(aX + bY, cX + dY) = (ac) \cdot Var(X) + (bd) \cdot$ $Var(Y) + (ad + bc) \cdot Cov(X, Y).$ Var(aX + bY) $a^2 \cdot Var(X) + 2ab \cdot Cov(X, Y) + b^2 \cdot Var(Y).$

Correlation Coefficient: $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \cdot \sigma_Y} = \frac{Cov(X,Y)}{\sqrt{Var(X)} \cdot \sqrt{Var(Y)}}$.

Uncorrelated: $\rho_{XY} = 0 \Leftrightarrow \sigma_{XY} = 0$. If X and Y are independent, then $\rho_{XY} = 0$, but the implication **does not** go the other way.

I.I.D.: If every marginal probability density function is equal and independent, then all of the random variables are independent and identically distributed (i.i.d.).

Combinations: If Y is a linear combination of independent random variables with constant multiples, then the expectation of Y is simply the sum of the expectations of random variables. The variance of Y is the sum of the variances of the random variables with their constant multiples squared. The moment of a linear combination of random variables is the product of their individual moments.

Summary of Combinations: Suppose X and Y are independent. Then,

- bern(p) + bern(p) = b(2, p)
- $b(n_1, p) + b(n_2, p) = b(n_1 + n_2, p)$
- geo(p) + geo(p) = negbin(2, p)
- $negbin(r_1, p) + negbin(r_2, p) = negbin(r_1 + r_2, p)$
- $poi(\lambda_1) + poi(\lambda_2) = poi(\lambda_1 + \lambda_2)$
- $exp(\theta) + exp(\theta) = gamma(2, \theta)$
- $gamma(\alpha_1, \theta) + gamma(\alpha_2, \theta) = gamma(\alpha_1 + \alpha_2, \theta)$
- $N(\mu_1, \sigma_1^2) + N(\mu_2, \sigma_2^2) = N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

1.3 Final Material

Everything taught in the class is on the final.

Chebyshev's Inequality: If $k \ge 1$: $\mathbb{P}(|X - \mu| \ge k\sigma) \le \frac{\sigma^2}{\epsilon^2}$. Let

Central Limit Theorem (CLT): $\frac{\sqrt{n}(\overline{X}-\mu)}{2} \to Z$, as $n \to \infty$. Normal Approximation to Binomial: If np > 5 and n(1 $p \ge 5$, then

 $N(\mu, \sigma^2)$ b(n,p)Mean np μ Variance np(1-p)

Normal Approximation to Poisson:

 $Poisson(\lambda)$ $N(\mu, \sigma^2)$ Mean σ^2 Variance

Jensen's Inequality: If $g: \mathbb{R} \to \mathbb{R}$ is a convex function and X is a random variable, then $\mathbb{E}[g(X)] \geq g(\mathbb{E}(X))$.

Maximum Likelihood Estimator: Maximize the likelihood function subject to the parameter space.

- 1. Construct likelihood function, $F(x;\theta) = \prod_{i=1}^{n} (f(x_i))$.
- 2. If applicable, take the log of this to make things easier to differentiate and to change the product into a sum. $l(x;\theta)$.
- 3. Differentiate to get $\frac{dl(x;\theta)}{d\theta}$
- 4. Assume convexity (ha) and find optimal θ by solving $\frac{dl(x;\theta)}{d\theta} =$

Sample Mean: $\frac{1}{n} \sum_{i=1}^{n} X_i \equiv \overline{X}$.

Sample Variance: For z distribution: $\frac{1}{n} \sum_{i=1}^{n} (X_i - \overline{X})$, and for t distribution: $\frac{1}{n-1} \sum_{i=1}^{n} (X_i - \overline{X}).$

Method of Moments Estimator: Construct a system of equations using moments and sample data. For example, equate theoretical mean with sample mean, and theoretical variance with sample variance, then solve. <><>maybe put an example here<><>><

Confidence Intervals for Means:

Confidence Intervals for Means:

Known Variance

Two Side
$$\left[\overline{X} - z_{\alpha} \frac{\sigma}{2} \frac{\sigma}{\sqrt{n}}, \overline{X} + z_{\alpha} \frac{\sigma}{2} \frac{\sigma}{\sqrt{n}}\right] = \left[\overline{X} - t_{\frac{\alpha}{2}} (n-1) \frac{S}{\sqrt{n}}, \overline{X} + t_{\frac{\alpha}{2}} (n-1) \frac{S}{\sqrt{n}}\right]$$

Lower $\left[\overline{X} - z_{\alpha} \frac{\sigma}{\sqrt{n}}, \infty\right)$ you can

Upper $\left(-\infty, \overline{X} + z_{\alpha} \frac{\sigma}{\sqrt{n}}\right]$ figure it out

 $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X})$ be the sample