

1 STAT 400, Aryn Harmon

1.1 Midterm 1 Material

Probability is a real-valued function: 1. $\mathbb{P}(S) = 1$; 2. $\mathbb{P}(A) \geq 0$; 3. If A_1, A_2 are mutually exclusive events, $\mathbb{P}(A_1 \cup A_2) = \mathbb{P}(A_1) + \mathbb{P}(A_2)$ and so on.

Inclusion-Exclusion: $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$.

Conditional Probability: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$ given $\mathbb{P}(B) > 0$. $\mathbb{P}(A|B) \neq \mathbb{P}(A)$ unless A and B are independent. Also see Bayes's Theorem. Probability of a string of unions given B is equal to the sum of the individual conditional probabilities.

Multiplication Rule: Probability of two events both occurring: $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$ or $\mathbb{P}(A \cap B) = \mathbb{P}(B) \cdot \mathbb{P}(A|B)$ (one is easier than the other).

Bayes's Rule: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A) \cdot \mathbb{P}(B|A)}{\mathbb{P}(B)}$. $\mathbb{P}(A)$ is the *prior probability* of A . $\mathbb{P}(A|B)$ is the *posterior probability* of A given that B occurred. Use to invert probabilities.

Bayes's Rule 2: $\mathbb{P}(A|B) = \frac{\mathbb{P}(A) \cdot \mathbb{P}(B|A)}{\mathbb{P}(A) \cdot \mathbb{P}(B|A) + \mathbb{P}(A') \cdot \mathbb{P}(B|A')}$.

Bayes's Rule Full: Given some partition of S : $A_1 \cup \dots \cup A_k = S$. $\mathbb{P}(A_i|B) = \frac{\mathbb{P}(A_i) \cdot \mathbb{P}(B|A_i)}{\sum_{m=1}^k \mathbb{P}(A_m) \cdot \mathbb{P}(B|A_m)}$.

Independence: $\mathbb{P}(A|B) = \mathbb{P}(A)$, $\mathbb{P}(B|A) = \mathbb{P}(B)$, and $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$.

Pairwise Independence: All of the following must hold: $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$, $\mathbb{P}(A \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(C)$, and $\mathbb{P}(B \cap C) = \mathbb{P}(B) \cdot \mathbb{P}(C)$.

Mutual Independence: A , B , and C must be pairwise independent and in addition: $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C)$.

Number of ways to take r objects from n candidates:

	Ordered Sample	Unordered Sample
W. R.	n^r	$\binom{n+r-1}{r}$
W/o. R.	nPr	$\binom{n}{r}$
	$nP_k = \frac{n!}{(n-k)!}$ - permutation	
	$\binom{n}{k} = nC_k = \frac{n!}{k!(n-k)!}$ - combination	

Binomial Theorem: $(a+b)^n = \sum_{r=0}^n \binom{n}{r} a^r b^{n-r}$. $\binom{n}{n-r} = \binom{n}{r}$, $\binom{n}{r}, \binom{n}{0} = \binom{n}{n} = 1$, $\binom{n}{1} = \binom{n}{n-1} = n$.

Some magic formulae: 1. $\sum_{r=0}^n \binom{n}{r} = 2^n$. 2. $\sum_{r=0}^n (-1)^r \binom{n}{r} = 0$. 3. $\sum_{r=0}^n \binom{n}{r} p^r (1-p)^{n-r} = 1$.

Pascal's Equation: $\binom{n}{r} = \binom{n-1}{r} + \binom{n-1}{r-1}$.

Hypergeometric Distribution: $f(x) = \mathbb{P}(X = x) = \frac{\binom{N_1}{x} \cdot \binom{N_2}{n-x}}{\binom{N}{n}}$, where $x \leq n$, $x \leq N_1$, $n-x \leq N_2$ and $N = N_1 + N_2$.

$\mathbb{E}(X) = n \frac{N_1}{N}$ and $Var(X) = n \cdot \frac{N_1}{N} \cdot \frac{N_2}{N} \cdot \frac{N-n}{N-1}$. Example: Urn model: N_1 red balls, N_2 blue balls, draw n balls from $N_1 + N_2$ balls, then look at the probability that there are x red balls in the selected n balls.

Mean: $\mu = \mathbb{E}(X) = \sum_{x \in S} x f(x)$.

Variance: $\sigma^2 = Var(X) = \mathbb{E}(X - \mu)^2 = \mathbb{E}(X^2) - \mu^2 = \sum_{i=1}^k x_i^2 f(x_i) - \mu^2$. **Standard Deviation:** σ .

r -th Moment: $\mathbb{E}(|X|^r) = \sum_{x \in S} |x|^r f(x) < \infty$; is the moment about the origin.

r -th Moment about b : $\mathbb{E}((X-b)^r) = \sum_{x \in S} (x-b)^r f(x)$; is the moment about b . Facts: μ is the first moment of X about the origin. σ^2 is the second moment of X about μ .

Example of Variance Properties: $Var(aX + b) = a^2 \cdot Var(X)$.

Bernoulli Distribution: A random experiment is called a set of Bernoulli trials if each trial has only two outcomes, has a constant p , and each trial is independent. $f(x) = p$ if $x = 1$, $1-p$ if $x = 0$, with $0 \leq p \leq 1$. $\mathbb{E}(X) = p$ and $Var(X) = p(1-p)$.

Binomial Distribution: Let X be the number of successes in n independent Bernoulli trials with p . Then $X \sim b(n, p)$.

$f(x) = \mathbb{P}(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$. $\mathbb{E}(X) = np$ and $Var(X) = np(1-p)$.

A note: binomial and hypergeometric are similar, but binomial has replacement and one category, and hypergeometric has two categories and no replacement.

Cumulative Distribution Function: $F(x) = \mathbb{P}(X \leq x)$, $x \in (-\infty, \infty)$. For discrete random variables: $f(x) = \mathbb{P}(X = x) = \mathbb{P}(X \leq x) - \mathbb{P}(X \leq x-1) = F(x) - F(x-1)$.

Geometric Distribution: $f(x) = p(1-p)^{x-1}$, $x = 1, 2, 3, \dots$. X represents the draw in which the first success is drawn. $f(x)$ is the probability of getting a success in the x -th draw.

Negative Binomial Distribution: $f(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}$ and $x = r, r+1, r+2, \dots$. X is the number of trials it takes to get r successes. $f(x)$ is the probability that the r -th success occurs on the x -th trial.

1.2 Midterm 2 Material

Moment Generating Function: $M(t) = \mathbb{E}(e^{tX}) = \sum_{x \in S} e^{tx} f(x)$ if $f(x)$ is the p.d.f. of some distribution and $t \in V_h(0)$ is finite. *Theorem:* $\mathbb{E}(X^r) = M^{(r)}(0)$, so $\mu = \mathbb{E}(X) = M'(0)$ and $\sigma^2 = \mathbb{E}(X^2) - [\mathbb{E}(X)]^2 = M''(0) - [M'(0)]^2$. To calculate the m.g.f. and p.d.f. of some random variable given only its moments, use the Taylor series expansion centered at zero. $M(t) = M(0) + M'(0) \left(\frac{t}{1!}\right) + M''(0) \left(\frac{t^2}{2!}\right) + \dots = 1 + \mathbb{E}(X) \left(\frac{t}{1!}\right) + \mathbb{E}(X^2) \left(\frac{t^2}{2!}\right) + \dots$

Poisson distribution: *Definition:* Poisson process counts the number of events occurring in a fixed time/space given a rate λ . Let X_t be the number events which occur in t unit time intervals. $\mathbb{P}(X_t = x) = \frac{(\lambda t)^x}{x!} e^{-\lambda t}$. $\mu = \lambda$, $\sigma^2 = \lambda$.

Poisson Approximation of Binomial Distribution: If $X \sim b(n, p)$ and n is large while p is small, then X can be approximated as $\hat{X} \sim poi(\lambda)$ s.t. $\lambda = np$.

Mean: $\mu = \int_{-\infty}^{\infty} x f(x) dx$.

Variance: $\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$.

M.G.F.: $M(t) = \mathbb{E}(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$, $t \in (-h, h)$.

Percentile: $p = \int_{-\infty}^{\pi_p} f(x) dx = F(\pi_p)$.

Uniform Distribution: $f(x) = \frac{1}{b-a}$, $a \leq x \leq b$. If $X \sim U(a, b)$, then $\mathbb{E}(X) = \frac{a+b}{2}$, $\sigma^2 = \frac{(b-a)^2}{12}$, and $M(t) = \mathbb{E}(e^{tX}) = \frac{e^{tb} - e^{ta}}{t(b-a)}$, $t \neq 0$ and $M(0) = 1$.

Exponential Distribution: This describes the waiting time between events in a Poisson process with rate λ . Let $\theta = \frac{1}{\lambda}$. Then if $X \sim Exp(\theta)$, $f(x) = \frac{1}{\theta} e^{-\frac{x}{\theta}}$, $x \geq 0$ and 0 otherwise. $\mu = \theta$, $\sigma^2 = \theta^2$, and $M(t) = (1 - \theta t)^{-1}$, $t < \frac{1}{\theta}$.

Memoryless Property of the Exponential Distribution: What happened in the past does not matter now. Only the present can determine the future. $\mathbb{P}(X > a+b | X > a) = \mathbb{P}(X > b)$, $\forall a, b \geq 0$.

Gamma Function: $\Gamma(x) = \int_0^{\infty} y^{x-1} e^{-y} dy$, $x > 0$.

Gamma Distribution: $f(t)$ represents the waiting time until the α -th occurrence of some Poisson process with rate λ . $f(t) = \frac{\lambda^\alpha t^{\alpha-1}}{(\alpha-1)!} e^{-\lambda t}$, $t \geq 0$. Then let $\theta = \lambda^{-1}$, then $f(t) = \frac{1}{\Gamma(\alpha)} \theta^\alpha t^{\alpha-1} e^{-\frac{t}{\theta}}$, $t \geq 0$, $\alpha \in \mathbb{R}$. $\mu = \alpha\theta$, $\sigma^2 = \alpha\theta^2$, and $M(t) = (1 - \theta t)^{-\alpha}$. If $\alpha = 1$, $\text{gamma}(1, \theta) = \text{Exp}(\theta)$.

χ^2 Distribution: If $X \sim \text{gamma}(\alpha, \theta)$, and $\theta = 2$ and $\alpha = \frac{r}{2}$, where r is a positive integer, then X is a χ^2 distribution with degree of freedom r . $X \sim \chi^2(r)$. $\mu = r$, $\sigma^2 = 2r$, and $M(t) = (1 - 2t)^{-\frac{r}{2}}$, $t < \frac{1}{2}$. $e^2 = \chi^2(2)$.

Normal Distribution: Bell curve! $f(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$, $x \in \mathbb{R}$. $X \sim N(\mu, \sigma^2)$. $X \sim N(0, 1)$ is the standard normal distribu-

tion. $\mu = \mu$, $\sigma^2 = \sigma^2$, and $M(t) = e^{\mu t + \frac{\sigma^2 t^2}{2}}$. Standardization: $X \sim N(\mu, \sigma^2)$, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Normal Square Distribution: Let $X \sim N(\mu, \sigma^2)$ and $Z = \frac{X - \mu}{\sigma}$. Then the random variable $V = Z^2 \sim \chi^2(1)$.

Cauchy Distribution: (Why do we even need this distribution?) $f(x) = \frac{1}{\pi(1+x^2)}$, $x \in \mathbb{R}$. Symmetric about zero, so median is zero, but μ is undefined because the tail of the p.d.f. is too heavy (i.e. each integral of the distribution does not converge). c.d.f. $= F(x) = \frac{1}{\pi} \arctan(x) + \frac{1}{2}$, $x \in \mathbb{R}$.

Joint Probability Density Function: $\mathbb{P}((X, Y) \in A) = \int \int f(x, y) dx dy$.

Marginal Probability Density Function: $f_x(x) = \int_{-\infty}^{\infty} f(x, y) dy$ and the other way around for y .

Mathematical Expectation: $\mathbb{E}[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$.

Independent Random Variables: Two events A and B are independent iff $f(x, y) = f_x(x) \cdot f_y(y) \forall x, y$. This works for both the p.d.f. and the c.d.f.

Trinomial Distribution: This is an extension of the binomial distribution into two dimensions. $f(x_1, x_2) = \mathbb{P}(X_1 = x_1, X_2 = x_2) = \frac{n!}{x_1! x_2! (n - x_1 - x_2)!} p_1^{x_1} p_2^{x_2} (1 - p_1 - p_2)^{n - x_1 - x_2}$ for $x_1, x_2 \geq 0$, $x_1 + x_2 \leq n$.

Covariance: $Cov(X, Y) = \mathbb{E}[(X - \mu_x)(Y - \mu_y)]$, $\mu_x = \mathbb{E}(X)$, $\mu_y = \mathbb{E}(Y)$. Useful identity: $Cov(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y)$.

Properties of Covariance: Zero-angle: $Cov(X, X) = Var(X)$. Symmetry: $Cov(X, Y) = Cov(Y, X)$. $Cov(aX + b, Y) = a \cdot Cov(X, Y)$. $Cov(X + Y, W) = Cov(X, W) + Cov(Y, W)$. $Cov(aX + bY, cX + dY) = (ac) \cdot Var(X) + (bd) \cdot Var(Y) + (ad + bc) \cdot Cov(X, Y)$. $Var(aX + bY) = a^2 \cdot Var(X) + 2ab \cdot Cov(X, Y) + b^2 \cdot Var(Y)$.

Correlation Coefficient: $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \cdot \sigma_Y} = \frac{Cov(X, Y)}{\sqrt{Var(X)} \cdot \sqrt{Var(Y)}}$.

Uncorrelated: $\rho_{XY} = 0 \Leftrightarrow \sigma_{XY} = 0$. If X and Y are independent, then $\rho_{XY} = 0$, but the implication **does not** go the other way.

I.I.D.: If every marginal probability density function is equal and independent, then all of the random variables are independent and identically distributed (i.i.d.).

Combinations: If Y is a linear combination of independent random variables with constant multiples, then the expectation of Y is simply the sum of the expectations of random variables. The variance of Y is the sum of the variances of the random variables with their constant multiples squared. The moment of a linear combination of random variables is the product of their individual moments.

Summary of Combinations: Suppose X and Y are independent. Then,

- $bern(p) + bern(p) = b(2, p)$
- $b(n_1, p) + b(n_2, p) = b(n_1 + n_2, p)$
- $geo(p) + geo(p) = negbin(2, p)$
- $negbin(r_1, p) + negbin(r_2, p) = negbin(r_1 + r_2, p)$
- $poi(\lambda_1) + poi(\lambda_2) = poi(\lambda_1 + \lambda_2)$
- $exp(\theta) + exp(\theta) = gamma(2, \theta)$
- $gamma(\alpha_1, \theta) + gamma(\alpha_2, \theta) = gamma(\alpha_1 + \alpha_2, \theta)$
- $N(\mu_1, \sigma_1^2) + N(\mu_2, \sigma_2^2) = N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

1.3 Final Material

Everything taught in the class is on the final.

Chebyshev's Inequality: If $k \geq 1$: $\mathbb{P}(|X - \mu| \geq k\sigma) \leq \frac{\sigma^2}{k^2}$. Let

$\epsilon = k\sigma$: $\mathbb{P}(|X - \mu| \geq \epsilon) \leq \frac{1}{k^2}$.

Central Limit Theorem (CLT): $\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \rightarrow Z$, as $n \rightarrow \infty$.

Normal Approximation to Binomial: If $np \geq 5$ and $n(1 - p) \geq 5$, then

	$b(n, p)$	$N(\mu, \sigma^2)$
Mean	np	μ
Variance	$np(1 - p)$	σ^2

Normal Approximation to Poisson:

	$Poisson(\lambda)$	$N(\mu, \sigma^2)$
Mean	λ	μ
Variance	λ	σ^2

Jensen's Inequality: If $g : \mathbb{R} \rightarrow \mathbb{R}$ is a convex function and X is a random variable, then $\mathbb{E}[g(X)] \geq g(\mathbb{E}(X))$.

Maximum Likelihood Estimator: Maximize the likelihood function subject to the parameter space.

1. Construct likelihood function, $F(x; \theta) = \Pi_{i=1}^n(f(x_i))$.
2. If applicable, take the log of this to make things easier to differentiate and to change the product into a sum. $l(x; \theta)$.
3. Differentiate to get $\frac{dl(x; \theta)}{d\theta}$.
4. Assume convexity (ha) and find optimal θ by solving $\frac{dl(x; \theta)}{d\theta} = 0$.

Sample Mean: $\frac{1}{n} \sum_{i=1}^n X_i \equiv \bar{X}$.

Sample Variance: For z distribution: $\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$, and for t distribution: $\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$.

Method of Moments Estimator: Construct a system of equations using moments and sample data. For example, equate theoretical mean with sample mean, and theoretical variance with sample variance, then solve. <><><maybe put an example here>><><>

Helpful Distribution Values:

α	$z_{1-\alpha}$
0.2	0.840
0.1	1.280
0.05	1.645
0.025	1.960

Confidence Intervals for Means:

	Known Variance	Unknown Variance
Two Side	$\left[\bar{X} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \right]$	$\left[\bar{X} - t_{\frac{\alpha}{2}} (n-1) \frac{S}{\sqrt{n}}, \bar{X} + t_{\frac{\alpha}{2}} (n-1) \frac{S}{\sqrt{n}} \right]$
Lower	$\left[\bar{X} - z_{\alpha} \frac{\sigma}{\sqrt{n}}, \infty \right)$	you can
Upper	$(-\infty, \bar{X} + z_{\alpha} \frac{\sigma}{\sqrt{n}}]$	figure it out

$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ be the sample variance.

Confidence Interval for Variance: $\left[\frac{(n-1)S^2}{b}, \frac{(n-1)S^2}{a} \right]$, $a = \chi_{1-\frac{\alpha}{2}}^2(n-1)$, $b = \chi_{\frac{\alpha}{2}}^2(n-1)$.

Confidence Interval for Standard Deviation: Use a and b from above, and just square root the expressions for the variance C.I.

Confidence Interval for Proportions (and Wilcox Interval for large n): Let $\hat{p} = \frac{Y}{n}$. Then $\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$. Lower bound: $\left[\hat{p} - z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, 1 \right]$. Upper bound: $\left[0, \hat{p} + z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right]$.