The bootstrap revisited

Packages for this section

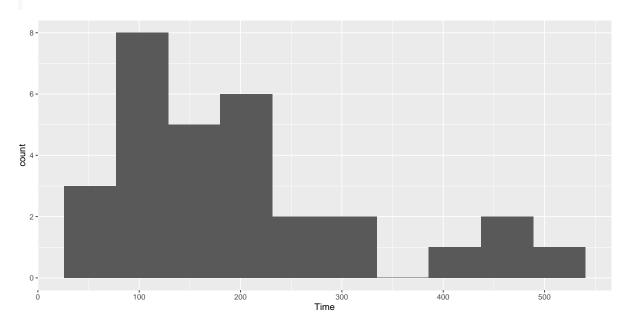
library(tidyverse)
library(bootstrap)

Source: Hesterberg et al

Is my sampling distribution normal enough?

• Recall IRS data (used as a motivation for the sign test) :

ggplot(irs, aes(x=Time))+geom_histogram(bins=10)



 \bullet t procedure for the mean would not be a good idea because the distribution is skewed.

What actually matters

- It's not the distribution of the *data* that has to be approx normal (for a t procedure).
- What matters is the sampling distribution of the sample mean.
- If the sample size is large enough, the sampling distribution will be normal enough even if the data distribution is not.
 - This is why we had to consider the sample size as well as the shape.
- But how do we know whether this is the case or not? We only have *one* sample.

The (nonparametric) bootstrap

- Typically, our sample will be reasonably representative of the population.
- Idea: pretend the sample is the population, and sample from it with replacement.
- Calculate test statistic, and repeat many times.
- This gives an idea of how our statistic might vary in repeated samples: that is, its sampling distribution.
- Called the **bootstrap distribution** of the test statistic.
- If the bootstrap distribution is approx normal, infer that the true sampling distribution also approx normal, therefore inference about the mean such as t is good enough.
- If not, we should be more careful.

Why it works

- We typically estimate population parameters by using the corresponding sample thing: eg. estimate population mean using sample mean.
- This called **plug-in principle**.
- The fraction of sample values less than a value x called the **empirical distribution** function (as a function of x).
- By plug-in principle, the empirical distribution function is an estimate of the population CDF.
- In this sense, the sample *is* an estimate of the population, and so sampling from it is an estimate of sampling from the population.

Bootstrapping the IRS data

• Sampling with replacement is done like this (the default sample size is as long as the original data):

```
boot <- sample(irs$Time, replace=T)
mean(boot)</pre>
```

[1] 196.4

• That's one bootstrapped mean. We need a whole bunch.

A whole bunch

• Use the same idea as for simulating power:

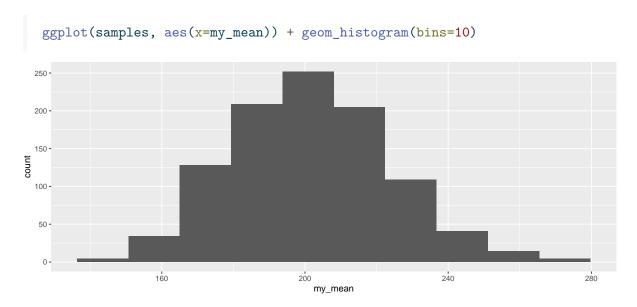
```
tibble(sim = 1:1000) %>%
    rowwise() %>%
    mutate(boot_sample = list(sample(irs$Time, replace = TRUE)))
# A tibble: 1,000 x 2
# Rowwise:
    sim boot_sample
  <int> <list>
     1 <dbl [30]>
      2 <dbl [30]>
      3 <dbl [30]>
     4 <dbl [30]>
 5
     5 <dbl [30]>
      6 <dbl [30]>
 6
 7
      7 <dbl [30]>
     8 <dbl [30]>
8
9
      9 <dbl [30]>
     10 <dbl [30]>
10
# i 990 more rows
```

Get the mean of each of those

```
tibble(sim = 1:1000) %>%
  rowwise() %>%
  mutate(boot_sample = list(sample(irs$Time, replace = TRUE))) %>%
  mutate(my_mean = mean(boot_sample)) -> samples
samples
```

```
# A tibble: 1,000 x 3
# Rowwise:
     sim boot_sample my_mean
   <int> <list>
                        <dbl>
       1 <dbl [30]>
                         196
1
2
                         202.
       2 <dbl [30]>
3
       3 <dbl [30]>
                         263.
4
       4 <dbl [30]>
                         173.
5
       5 <dbl [30]>
                         204.
6
       6 <dbl [30]>
                         197.
7
                         210.
       7 <dbl [30]>
8
       8 <dbl [30]>
                         160.
9
                         198.
       9 <dbl [30]>
10
      10 <dbl [30]>
                         178.
# i 990 more rows
```

Sampling distribution of sample mean

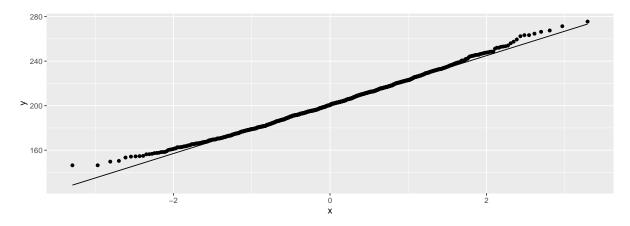


• Is that a slightly long right tail?

Normal quantile plot

might be better than a histogram:

```
ggplot(samples, aes(sample = my_mean)) +
stat_qq()+stat_qq_line()
```



• a very very slight right-skewness, but very close to normal.

Confidence interval from the bootstrap distribution

There are two ways (at least):

• percentile bootstrap interval: take the 2.5 and 97.5 percentiles (to get the middle 95%). This is easy, but not always the best:

```
(b_p=quantile(samples$my_mean, c(0.025, 0.975)))

2.5% 97.5%
162.5775 246.9092
```

• bootstrap t: use the SD of the bootstrapped sampling distribution as the SE of the estimator of the mean and make a t interval:

```
n <- length(irs$Time)
t_star <- qt(0.975, n-1)
b_t <- with(samples, mean(my_mean)+c(-1, 1)*t_star*sd(my_mean))
b_t</pre>
```

[1] 156.5070 246.4032

Comparing

 \bullet get ordinary t interval:

```
my_names=c("LCL", "UCL")
o_t <- t.test(irs$Time)$conf.int</pre>
```

• Compare the 2 bootstrap intervals with the ordinary t-interval:

```
tibble(limit=my_names, o_t, b_t, b_p)
```

- The bootstrap t and the ordinary t are very close
- The percentile bootstrap interval is noticeably shorter (common) and higher (skewness).

Which to prefer?

- If the intervals agree, then they are all good.
- If they disagree, they are all bad!
- In that case, use BCA interval (over).

Bias correction and acceleration

- this from "An introduction to the bootstrap", by Brad Efron and Robert J. Tibshirani.
- there is way of correcting the CI for skewness in the bootstrap distribution, called the BCa method
- complicated (see the Efron and Tibshirani book), but implemented in bootstrap package.

Run this on the IRS data:

```
bca=bcanon(irs$Time, 1000, mean)
bca$confpoints

alpha bca point
[1,] 0.025 161.8333
[2,] 0.050 168.0667
```

```
[3,] 0.100 174.8333

[4,] 0.160 180.7667

[5,] 0.840 224.1333

[6,] 0.900 232.3000

[7,] 0.950 241.9333

[8,] 0.975 253.7333
```

use 2.5% and 97.5% points for CI

```
bca$confpoints %>% as_tibble() %>%
  filter(alpha %in% c(0.025, 0.975)) %>%
  pull(`bca point`) -> b_bca
b_bca
```

[1] 161.8333 253.7333

Comparing

- The BCA interval says that the mean should be estimated even higher than the bootstrap percentile interval does.
- The BCA interval is the one to trust.

Bootstrapping the correlation

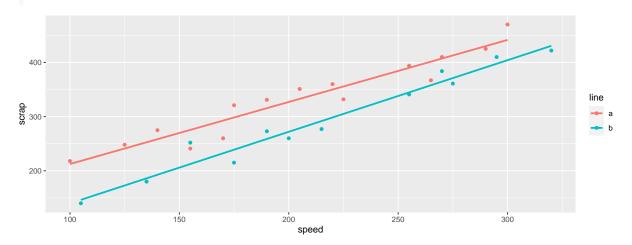
Recall the soap data:

```
url <- "http://ritsokiguess.site/datafiles/soap.txt"
soap <- read_delim(url," ")
soap</pre>
```

```
# A tibble: 27 \times 4
    case scrap speed line
   <dbl> <dbl> <dbl> <chr>
 1
        1
            218
                  100 a
 2
       2
            248
                   125 a
 3
       3
            360
                   220 a
 4
       4
            351
                  205 a
 5
       5
                  300 a
            470
 6
       6
            394
                  255 a
 7
       7
            332
                  225 a
 8
       8
            321
                  175 a
 9
       9
            410
                  270 a
10
            260
                  170 a
       10
# i 17 more rows
```

Scatterplot

```
ggplot(soap, aes(x=speed, y=scrap, colour=line))+
geom_point()+geom_smooth(method="lm", se=F)
```



Comments

- Line B produces less scrap for any given speed.
- For line B, estimate the correlation between speed and scrap (with a confidence interval.)

Extract the line B data; standard correlation test

```
soap %>% filter(line=="b") -> line_b
with(line_b, cor.test(speed, scrap))

Pearson's product-moment correlation

data: speed and scrap
t = 15.829, df = 10, p-value = 2.083e-08
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:
    0.9302445    0.9947166
sample estimates:
        cor
    0.9806224
```

Bootstrapping a correlation 1/2

- This illustrates a different technique: we need to keep the x and y values together.
- Sample *rows* of the data frame rather than individual values of speed and scrap:

```
line_b %>% sample_frac(replace=T)
# A tibble: 12 x 4
   case scrap speed line
  <dbl> <dbl> <chr>
   24 252 155 b
   22 260 200 b
    16 140 105 b
25 422 320 b
    16 140 105 b
5
   19 341 255 b
7
    19 341 255 b
    19 341 255 b
17 277 215 b
10 16 140 105 b
11 20 215 175 b
12 18 384 270 b
```

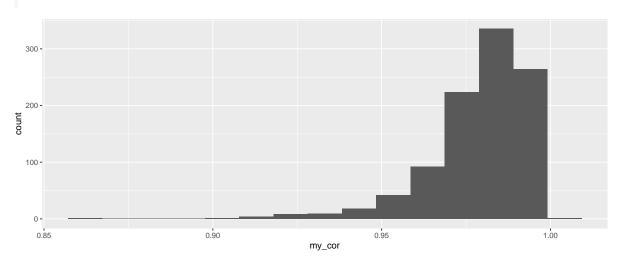
Bootstrapping a correlation 2/2

1000 times:

```
tibble(sim = 1:1000) %>%
  rowwise() %>%
  mutate(boot_df = list(sample_frac(line_b, replace = TRUE))) %>%
  mutate(my_cor = with(boot_df, cor(speed, scrap))) -> cors
```

A picture of this

```
ggplot(cors, aes(x=my_cor))+geom_histogram(bins=15)
```



Comments and next steps

- This is very left-skewed.
- Bootstrap percentile interval is:

```
(b_p=quantile(cors$my_cor, c(0.025, 0.975)))
```

2.5% 97.5% 0.9415748 0.9962462

• We probably need the BCA interval instead.

Getting the BCA interval 1/2

• To use bcanon, write a function that takes a vector of row numbers and returns the correlation between speed and scrap for those rows:

• That looks about right.

Getting the BCA interval 2/2

- Inputs to bcanon are now:
 - row numbers (1 through 12 in our case: 12 rows in line_b)
 - number of bootstrap samples
 - the function we just wrote
 - the data frame:

```
points=bcanon(1:12, 1000, theta, line_b)$confpoints
points %>% as_tibble() %>%
  filter(alpha %in% c(0.025, 0.975)) %>%
  pull(`bca point`) -> b_bca
b_bca
```

[1] 0.9314334 0.9947799

Comparing the results

```
tibble(limit=my_names, o_c, b_p, b_bca)

# A tibble: 2 x 4
  limit o_c b_p b_bca
  <chr> <dbl> <dbl> <dbl> 1 LCL 0.930 0.942 0.931
2 UCL 0.995 0.996 0.995
```

- The bootstrap percentile interval doesn't go down far enough.
- The BCA interval seems to do a better job in capturing the skewness of the distribution.
- The ordinary confidence interval for the correlation is very similar to the BCA one, and thus seems to be trustworthy here even though the correlation has a very skewed distribution. (cor.test uses the Fisher z transformation which "spreads out" correlations close to 1).

The z-transformed bootstrapped correlations

```
cors %>%
  mutate(z = 0.5 * log((1+my_cor)/(1-my_cor))) %>%
  ggplot(aes(sample=z)) + stat_qq() + stat_qq_line()
```

