Regression revisited

Regression

- \blacktriangleright Use regression when one variable is an outcome (response, y).
- See if/how response depends on other variable(s), explanatory, $x_1, x_2,$
- Can have one or more than one explanatory variable, but always one response.
- Assumes a straight-line relationship between response and explanatory.
- Ask:
 - is there a relationship between y and x's, and if so, which ones?
 - what does the relationship look like?

Packages

```
library(MASS) # for Box-Cox, later
library(tidyverse)
library(broom)
library(marginaleffects)
library(conflicted)
conflict_prefer("select", "dplyr")
```

A regression with one x

13 children, measure average total sleep time (ATST, mins) and age (years) for each. See if ATST depends on age. Data in sleep.txt, ATST then age. Read in data:

my_url <- "http://ritsokiguess.site/datafiles/sleep.txt"

```
my_url <- "http://ritsokiguess.site/datafiles/sleep.txt"
sleep <- read_delim(my_url, " ")</pre>
```

Check data

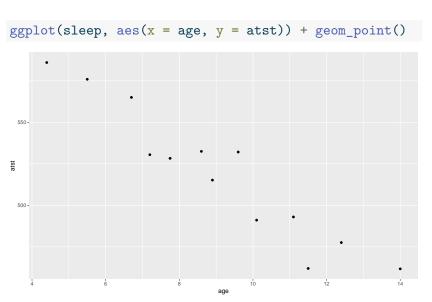
summary(sleep)

```
atst age
Min. :461.8 Min. : 4.400
1st Qu.:491.1 1st Qu.: 7.200
Median :528.3 Median : 8.900
Mean :519.3 Mean : 9.058
3rd Qu.:532.5 3rd Qu.:11.100
Max. :586.0 Max. :14.000
```

sleep

```
# A tibble: 13 x 2
    atst age
    <dbl> <dbl>
1 586 4.4
2 462. 14
3 491. 10.1
4 565 6.7
```

The scatterplot



Correlation

Measures how well a straight line fits the data:

```
with(sleep, cor(atst, age))
```

[1] -0.9515469

- ▶ 1 is perfect upward trend, -1 is perfect downward trend, 0 is no trend.
- This one close to perfect downward trend.
- Can do correlations of all pairs of variables:

cor(sleep)

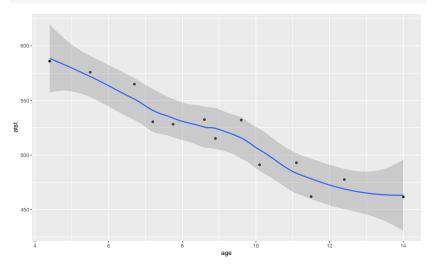
```
atst age
atst 1.0000000 -0.9515469
age -0.9515469 1.0000000
```

Lowess curve

- Sometimes nice to guide the eye: is the trend straight, or not?
- ▶ Idea: lowess curve. "Locally weighted least squares", not affected by outliers, not constrained to be linear.
- Lowess is a guide: even if straight line appropriate, may wiggle/bend a little. Looking for serious problems with linearity.
- ► Add lowess curve to plot using geom_smooth:

Plot with lowess curve

```
ggplot(sleep, aes(x = age, y = atst)) + geom_point() +
  geom_smooth()
```



The regression

Scatterplot shows no obvious curve, and a pretty clear downward trend. So we can run the regression:

```
sleep.1 <- lm(atst ~ age, data = sleep)</pre>
```

The output

```
summary(sleep.1)
```

```
Call:
lm(formula = atst ~ age, data = sleep)
Residuals:
   Min
          1Q Median
                           3Q
                                 Max
-23.011 -9.365 2.372 6.770 20.411
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 646.483 12.918 50.05 2.49e-14 ***
        -14.041 1.368 -10.26 5.70e-07 ***
age
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Conclusions

- ➤ The relationship appears to be a straight line, with a downward trend.
- F-tests for model as a whole and t-test for slope (same) both confirm this (P-value $5.7 \times 10^{-7} = 0.00000057$).
- Slope is -14, so a 1-year increase in age goes with a 14-minute decrease in ATST on average.
- ▶ R-squared is correlation squared (when one x anyway), between 0 and 1 (1 good, 0 bad).
- ▶ Here R-squared is 0.9054, pleasantly high.

Doing things with the regression output

- Output from regression (and eg. *t*-test) is all right to look at, but hard to extract and re-use information from.
- Package broom extracts info from model output in way that can be used in pipe (later):

```
tidy(sleep.1)
```

also one-line summary of model:

```
glance(sleep.1)
# A tibble: 1 x 12
```

r.squared adj.r.squared sigma statistic p.value

d:

Broom part 2

sleep.1 %>% augment(sleep)

```
# A tibble: 13 x 8
   atst age .fitted .resid .hat .sigma .cooksd
  <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <
1 586 4.4 585. 1.30 0.312 13.8 0.00320
2 462. 14 450. 11.8 0.341 13.0 0.319
3 491. 10.1 505. -13.6 0.0887 13.0 0.0568
4
   565 6.7 552. 12.6 0.137 13.1 0.0844
5 462 11.5 485. -23.0 0.141 11.3 0.294
  532. 9.6 512. 20.4 0.0801 12.0 0.114
6
  478. 12.4 472. 5.23 0.198 13.7 0.0243
  515. 8.9 522. -6.32 0.0772 13.6 0.0105
8
9
  493 11.1 491. 2.37 0.122
                                13.8 0.00258
10
  528. 7.75 538. -9.37 0.0954
                                13.4 0.0296
11 576. 5.5 569. 6.64 0.214 13.6 0.0441
12 532. 8.6 526. 6.77 0.0792 13.6 0.0124
13 530. 7.2 545. -14.9 0.114 12.9 0.0933
```

CI for mean response and prediction intervals

Once useful regression exists, use it for prediction:

- To get a single number for prediction at a given x, substitute into regression equation, eg. age 10: predicted ATST is 646.48-14.04(10)=506 minutes.
- ► To express uncertainty of this prediction:
- ► CI for mean response expresses uncertainty about mean ATST for all children aged 10, based on data.
- Prediction interval expresses uncertainty about predicted ATST for a new child aged 10 whose ATST not known. More uncertain.
- Also do above for a child aged 5.

The marginal effects package 1/2

To get predictions for specific values, set up a dataframe with those values first:

```
new <- datagrid(model = sleep.1, age = c(10, 5))
new</pre>
```

```
atst age
1 519.3038 10
2 519.3038 5
```

Any variables in the dataframe that you don't specify are set to their mean values (quantitative) or most common category (categorical).

The marginal effects package 2/2

Then feed into newdata in predictions. This contains a lot of columns, so you probably want only to display the ones you care about:

```
cbind(predictions(sleep.1, newdata = new)) %>%
  select(estimate, conf.low, conf.high, age)
```

```
estimate conf.low conf.high age
1 576.2781 563.2588 589.2974 5
2 506.0729 498.4899 513.6558 10
```

The confidence limits are a 95% confidence interval for the mean response at that age.

Prediction intervals

These are obtained (instead) with predict as below. Use the same dataframe new as before:

```
pp <- predict(sleep.1, new, interval = "p")
pp
      fit lwr
                      upr
1 506.0729 475.8982 536.2475
2 576.2781 543.8474 608.7088
cbind(new, pp) %>% select(-atst)
      fit lwr
 age
                           upr
1 10 506.0729 475.8982 536.2475
2
   5 576.2781 543.8474 608.7088
```

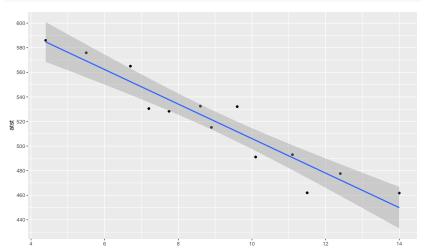
Comments

- Age 10 closer to centre of data, so intervals are both narrower than those for age 5.
- Prediction intervals bigger than CI for mean (additional uncertainty).
- ▶ Technical note: output from predict is R matrix, not data frame, so Tidyverse bind_cols does not work. Use base R cbind.

That grey envelope

Marks confidence interval for mean for all x:

```
ggplot(sleep, aes(x = age, y = atst)) + geom_point() +
  geom_smooth(method = "lm") +
  scale_y_continuous(breaks = seq(420, 600, 20))
```



Diagnostics

How to tell whether a straight-line regression is appropriate?

- ▶ Before: check scatterplot for straight trend.
- After: plot *residuals* (observed minus predicted response) against predicted values. Aim: a plot with no pattern.

Residual plot

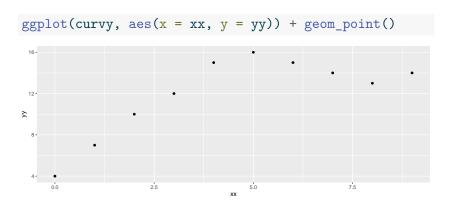
Not much pattern here — regression appropriate.

An inappropriate regression

Different data:

```
my_url <- "http://ritsokiguess.site/datafiles/curvy.txt"
curvy <- read_delim(my_url, " ")</pre>
```

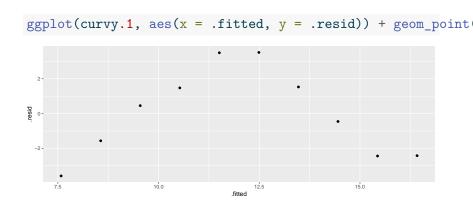
Scatterplot



Regression line, anyway

```
curvy.1 \leftarrow lm(yy \sim xx, data = curvy)
summary(curvy.1)
Call:
lm(formula = vv ~ xx, data = curvy)
Residuals:
  Min 10 Median 30 Max
-3.582 - 2.204  0.000  1.514  3.509
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 7.5818 1.5616 4.855 0.00126 **
            xx
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 2.657 on 8 degrees of freedom
Multiple R-squared: 0.5848, Adjusted R-squared: 0.5329
F-statistic: 11.27 on 1 and 8 DF, p-value: 0.009984
```

Residual plot



No good: fixing it up

- Residual plot has curve: middle residuals positive, high and low ones negative. Bad.
- Fitting a curve would be better. Try this:

```
curvy.2 <- lm(yy \sim xx + I(xx^2), data = curvy)
```

- Adding xx-squared term, to allow for curve.
- Another way to do same thing: specify how model *changes*:

```
curvy.2a <- update(curvy.1, . ~ . + I(xx^2))</pre>
```

Regression 2

tidy(curvy.2)

```
# A tibble: 3 x 5
 term estimate std.error statistic p.value
          <dbl>
 <chr>
                       <dbl>
                               <dbl> <dbl>
1 (Intercept) 3.9 0.773 5.04 0.00149
          3.74 0.400 9.36 0.0000331
2 xx
3 \text{ I}(xx^2) -0.307 0.0428 -7.17 0.000182
glance(curvy.2) #
# A tibble: 1 x 12
 r.squared adj.r.squared sigma statistic p.value
                                              df
     <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <
     0.950
                 0.936 0.983 66.8 0.0000275
# i 6 more variables: logLik <dbl>, AIC <dbl>, BIC <dbl>,
   deviance <dbl>, df.residual <int>, nobs <int>
#
```

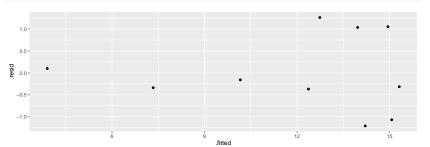
Comments

- xx-squared term definitely significant (P-value 0.000182), so need this curve to describe relationship.
- Adding squared term has made R-squared go up from 0.5848 to 0.9502: great improvement.
- ► This is a definite curve!

The residual plot now

No problems any more:

 $ggplot(curvy.2, aes(x = .fitted, y = .resid)) + geom_point$



Another way to handle curves

- Above, saw that changing x (adding x^2) was a way of handling curved relationships.
- ightharpoonup Another way: change y (transformation).
- ightharpoonup Can guess how to change y, or might be theory:
- **•** example: relationship $y = ae^{bx}$ (exponential growth):
- \blacktriangleright take logs to get $\ln y = \ln a + bx$.
- lacktriangle Taking logs has made relationship linear ($\ln y$ as response).
- Or, estimate transformation, using Box-Cox method.

Box-Cox

- Install package MASS via install.packages("MASS") (only need to do once)
- Every R session you want to use something in MASS, type library(MASS)

Some made-up data

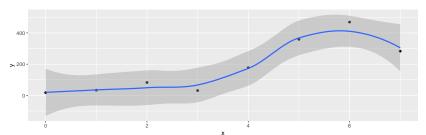
```
my_url <- "http://ritsokiguess.site/datafiles/madeup2.csv"
madeup <- read_csv(my_url)
madeup</pre>
```

```
# A tibble: 8 x 3
  ...1 x y
 <dbl> <dbl> <dbl>
         0 17.9
    2 1 33.6
  3 2 82.7
3
  4 3 31.2
4
5
  5 4 177.
6
    6
         5 359.
         6 469.
8
    8
         7 283.
```

Seems to be faster-than-linear growth, maybe exponential growth.

Scatterplot: faster than linear growth

```
ggplot(madeup, aes(x = x, y = y)) + geom_point() +
  geom_smooth()
```

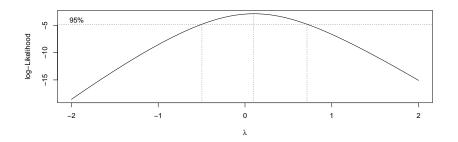


Running Box-Cox

- ▶ library(MASS) first.
- ► Feed boxcox a model formula with a squiggle in it, such as you would use for lm.
- Output: a graph (next page):

```
boxcox(y ~ x, data = madeup)
```

The Box-Cox output



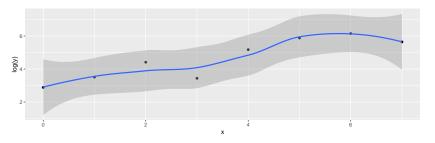
Comments

- λ (lambda) is the power by which you should transform y to get the relationship straight (straighter). Power 0 is "take logs"
- Middle dotted line marks best single value of λ (here about 0.1).
- Outer dotted lines mark 95% CI for λ , here -0.3 to 0.7, approx. (Rather uncertain about best transformation.)
- Any power transformation within the CI supported by data. In this case, log $(\lambda=0)$ and square root $(\lambda=0.5)$ good, but no transformation $(\lambda=1)$ not.
- Pick a "round-number" value of λ like 2,1,0.5,0,-0.5,-1. Here 0 and 0.5 good values to pick.

Did transformation straighten things?

▶ Plot transformed y against x. Here, log:

```
ggplot(madeup, aes(x = x, y = log(y))) + geom_point() +
   geom_smooth()
```



Looks much straighter.

Regression with transformed \boldsymbol{y}

```
madeup.1 \leftarrow lm(log(y) \sim x, data = madeup)
glance(madeup.1)
# A tibble: 1 x 12
  r.squared adj.r.squared sigma statistic p.value
                                                   df
      <dbl>
                 <dbl> <dbl>
                                  <dbl> <dbl> <dbl> <
 0.811
                   0.779 0.588 25.7 0.00228
# i 6 more variables: logLik <dbl>, AIC <dbl>, BIC <dbl>,
#
   deviance <dbl>, df.residual <int>, nobs <int>
tidy(madeup.1)
```

R-squared now decently high.

Multiple regression

- \blacktriangleright What if more than one x? Extra issues:
 - Now one intercept and a slope for each x: how to interpret?
 - Which x-variables actually help to predict y?
 - Different interpretations of "global" *F*-test and individual *t*-tests.
 - R-squared no longer correlation squared, but still interpreted as "higher better".
 - In 1m line, add extra xs after ~.
 - Interpretation not so easy (and other problems that can occur).

Multiple regression example

Study of women and visits to health professionals, and how the number of visits might be related to other variables:

timedrs: number of visits to health professionals (over course of study)

phyheal: number of physical health problems

menheal: number of mental health problems

stress: result of questionnaire about number and type of life

changes

timedrs response, others explanatory.

The data

```
my_url <-
   "http://ritsokiguess.site/datafiles/regressx.txt"
visits <- read_delim(my_url, " ")</pre>
```

Check data

visits

```
# A tibble: 465 x 5
   subjno timedrs phyheal menheal stress
    <dbl>
             <dbl> <dbl> <dbl> <dbl> <
                                   8
                                        265
                                   6
                                        415
 3
                          3
                                         92
                13
                                        241
 5
        5
                15
                          3
                                   6
                                         86
 6
                 3
                          5
                                   5
                                        247
                          5
                                   6
                                         13
 8
                                   5
                                          12
 9
                          5
                                   4
                                        269
10
       10
                          3
                                        391
# i 455 more rows
```

Fit multiple regression

```
visits.1 <- lm(timedrs ~ phyheal + menheal + stress,
   data = visits)
summary(visits.1)</pre>
```

```
Call:
lm(formula = timedrs ~ phyheal + menheal + stress, data = '
```

```
Min 1Q Median 3Q Max -14.792 -4.353 -1.815 0.902 65.886
```

Coefficients:

Residuals:

Estimate Std. Error t value Pr(>|t|)

The slopes

- ▶ Model as a whole strongly significant even though R-sq not very big (lots of data). At least one of the x's predicts timedrs.
- ➤ The physical health and stress variables definitely help to predict the number of visits, but with those in the model we don't need menheal. However, look at prediction of timedrs from menheal by itself:

Just menheal visits.2 <- lm(timedr</pre>

```
visits.2 <- lm(timedrs ~ menheal, data = visits)</pre>
summary(visits.2)
Call:
lm(formula = timedrs ~ menheal, data = visits)
Residuals:
   Min
        1Q Median 3Q
                                 Max
-13.826 -5.150 -2.818 1.177 72.513
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.8159 0.8702 4.385 1.44e-05 ***
menheal 0.6672 0.1173 5.688 2.28e-08 ***
```

Signif. codes:

0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

menheal by itself

- menheal by itself does significantly help to predict timedrs.
- ▶ But the R-sq is much less (6.5% vs. 22%).
- So other two variables do a better job of prediction.
- With those variables in the regression (phyheal and stress), don't need menheal as well.

Investigating via correlation

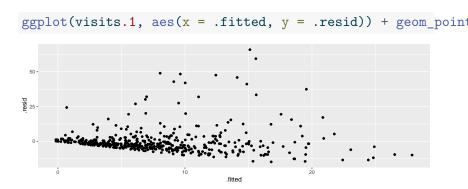
Leave out first column (subjno):

```
visits %>% select(-subjno) %>% cor()
```

```
timedrs phyheal menheal stress
timedrs 1.0000000 0.4395293 0.2555703 0.2865951
phyheal 0.4395293 1.0000000 0.5049464 0.3055517
menheal 0.2555703 0.5049464 1.0000000 0.3697911
stress 0.2865951 0.3055517 0.3697911 1.0000000
```

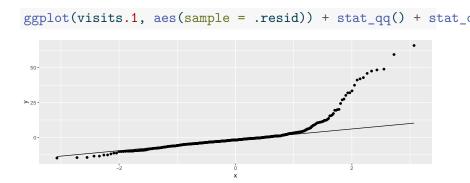
- phyheal most strongly correlated with timedrs.
- Not much to choose between other two.
- But menheal has higher correlation with phyheal, so not as much to add to prediction as stress.
- ▶ Goes to show things more complicated in multiple regression.

Residual plot (from timedrs on all)



Apparently random. But...

Normal quantile plot of residuals

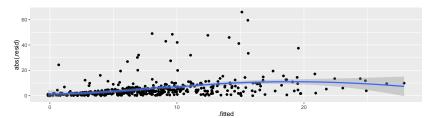


Not normal at all; upper tail is way too long.

Absolute residuals

Is there trend in *size* of residuals (fan-out)? Plot *absolute value* of residual against fitted value:

```
ggplot(visits.1, aes(x = .fitted, y = abs(.resid))) +
  geom_point() + geom_smooth()
```



Comments

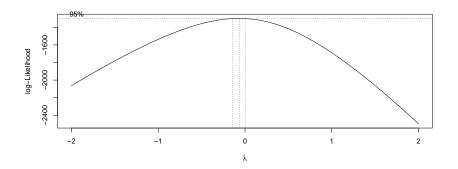
- On the normal quantile plot:
 - highest (most positive) residuals are way too high
 - distribution of residuals skewed to right (not normal at all)
- On plot of absolute residuals:
 - size of residuals getting bigger as fitted values increase
 - predictions getting more variable as fitted values increase
 - that is, predictions getting less accurate as fitted values increase, but predictions should be equally accurate all way along.
- ▶ Both indicate problems with regression, of kind that transformation of response often fixes: that is, predict function of response timedrs instead of timedrs itself.

Box-Cox transformations

- ▶ Taking log of timedrs and having it work: lucky guess. How to find good transformation?
- ► Box-Cox again.
- ► Extra problem: some of timedrs values are 0, but Box-Cox expects all +. Note response for boxcox:

```
boxcox(timedrs + 1 ~ phyheal + menheal + stress, data = vis
```

Try 1



Comments on try 1

- ightharpoonup Best: λ just less than zero.
- ► Hard to see scale.
- Focus on λ in (-0.3, 0.1):

```
my.lambda <- seq(-0.3, 0.1, 0.01) my.lambda
```

```
[1] -0.30 -0.29 -0.28 -0.27 -0.26 -0.25 -0.24 -0.23 -0.22

[10] -0.21 -0.20 -0.19 -0.18 -0.17 -0.16 -0.15 -0.14 -0.13

[19] -0.12 -0.11 -0.10 -0.09 -0.08 -0.07 -0.06 -0.05 -0.04

[28] -0.03 -0.02 -0.01 0.00 0.01 0.02 0.03 0.04 0.05

[37] 0.06 0.07 0.08 0.09 0.10
```

Try 2

```
boxcox(timedrs + 1 ~ phyheal + menheal + stress,
  lambda = my.lambda,
  data = visits
)
```

Comments

- ▶ Best: λ just about -0.07.
- \triangleright CI for λ about (-0.14, 0.01).
- ▶ Only nearby round number: $\lambda = 0$, log transformation.

Fixing the problems

- Try regression again, with transformed response instead of original one.
- ▶ Then check residual plot to see that it is OK now.

```
visits.3 <- lm(log(timedrs + 1) ~ phyheal + menheal + stres
data = visits
)</pre>
```

- timedrs+1 because some timedrs values 0, can't take log of 0.
- ➤ Won't usually need to worry about this, but when response could be zero/negative, fix that before transformation.

Output

```
summary(visits.3)
Call:
lm(formula = log(timedrs + 1) ~ phyheal + menheal + stress, data = visits)
Residuals:
    Min 10 Median
                              30
                                      Max
-1.95865 -0.44076 -0.02331 0.42304 2.36797
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.3903862 0.0882908 4.422 1.22e-05 ***
phyheal
          0.2019361 0.0173624 11.631 < 2e-16 ***
menheal 0.0071442 0.0101335 0.705 0.481
stress 0.0013158 0.0002837 4.638 4.58e-06 ***
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.7625 on 461 degrees of freedom
Multiple R-squared: 0.3682, Adjusted R-squared: 0.3641
F-statistic: 89.56 on 3 and 461 DF. p-value: < 2.2e-16
```

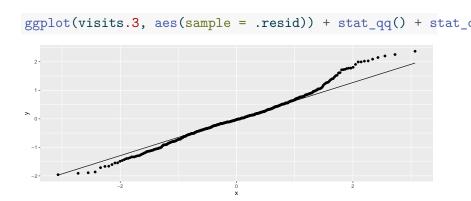
Comments

- Model as a whole strongly significant again
- ▶ R-sq higher than before (37% vs. 22%) suggesting things more linear now
- Same conclusion re menheal: can take out of regression.
- Should look at residual plots (next pages). Have we fixed problems?

Residuals against fitted values

```
ggplot(visits.3, aes(x = .fitted, y = .resid)) +
 geom_point()
                             .fitted
```

Normal quantile plot of residuals



Absolute residuals against fitted

```
ggplot(visits.3, aes(x = .fitted, y = abs(.resid))) +
  geom_point() + geom_smooth()
 2.0 -
abs(.resid)
 0.5 -
 0.0 -
                                   .fitted
```

Comments

- Residuals vs. fitted looks a lot more random.
- Normal quantile plot looks a lot more normal (though still a little right-skewness)
- Absolute residuals: not so much trend (though still some).
- Not perfect, but much improved.

Testing more than one x at once

- ▶ The *t*-tests test only whether one variable could be taken out of the regression you're looking at.
- ➤ To test significance of more than one variable at once, fit model with and without variables
 - then use anova to compare fit of models:

Results of tests

```
anova(visits.6, visits.5)
```

Analysis of Variance Table

```
Model 1: log(timedrs + 1) ~ stress

Model 2: log(timedrs + 1) ~ phyheal + menheal + stress

Res.Df RSS Df Sum of Sq F Pr(>F)

1 463 371.47

2 461 268.01 2 103.46 88.984 < 2.2e-16 ***

---

Signif. codes:
0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

- ▶ Models don't fit equally well, so bigger one fits better.
- Or "taking both variables out makes the fit worse, so don't do it".
- Taking out those x's is a mistake. Or putting them in is a good idea.

The punting data

Data set punting.txt contains 4 variables for 13 right-footed football kickers (punters): left leg and right leg strength (lbs), distance punted (ft), another variable called "fred". Predict punting distance from other variables:

left	right	punt	fred
170	170	162.50	171
130	140	144.0	136
170	180	174.50	174
160	160	163.50	161
150	170	192.0	159
150	150	171.75	151
180	170	162.0	174
110	110	104.83	111
110	120	105.67	114
120	130	117.58	126
140	120	140.25	129
130	140	150.17	136
150	160	165.17	154

Reading in

Separated by *multiple spaces* with *columns lined up*:

```
my_url <- "http://ritsokiguess.site/datafiles/punting.txt"
punting <- read_table(my_url)</pre>
```

The data

punting

```
A tibble: 13 x 4
    left right punt
                         fred
   <dbl> <dbl> <dbl> <dbl> <dbl>
     170
            170
                  162.
                           171
 1
 2
     130
            140
                  144
                           136
 3
     170
            180
                  174.
                           174
 4
     160
            160
                  164.
                           161
 5
     150
            170
                  192
                           159
 6
     150
            150
                  172.
                           151
 7
            170
                  162
     180
                           174
8
            110
     110
                  105.
                           111
9
     110
            120
                  106.
                           114
10
     120
            130
                  118.
                           126
11
     140
            120
                  140.
                           129
12
     130
            140
                  150.
                           136
13
     150
            160
                  165.
                           154
```

```
Regression and output
   punting.1 <- lm(punt ~ left + right + fred, data = punting)
   glance(punting.1)
   # A tibble: 1 x 12
     r.squared adj.r.squared sigma statistic p.value
                                                      df
         <dbl>
                  <dbl> <dbl> <dbl> <dbl> <dbl> <
                      0.704 14.7 10.5 0.00267
        0.778
                                                       3
   # i 6 more variables: logLik <dbl>, AIC <dbl>, BIC <dbl>,
      deviance <dbl>, df.residual <int>, nobs <int>
   tidy(punting.1)
```

```
# A tibble: 4 \times 5
```

estimate std.error statistic p.value term <chr> <dbl> <dbl> <dbl> <dbl>

1 (Intercept) -4.69 29.1 -0.161 0.876

0.268 2.11 0.127 0.902 2 left

-0.267 4.23 -0.0632 0.951

1.05 2.15 0.490 0.636 3 right

4 fred

Comments

- Overall regression strongly significant, R-sq high.
- \blacktriangleright None of the x's significant! Why?
- ▶ t-tests only say that you could take any one of the x's out without damaging the fit; doesn't matter which one.
- Explanation: look at correlations.

The correlations

cor(punting)

```
left right punt fred
left 1.0000000 0.8957224 0.8117368 0.9722632
right 0.8957224 1.0000000 0.8805469 0.9728784
punt 0.8117368 0.8805469 1.0000000 0.8679507
fred 0.9722632 0.9728784 0.8679507 1.0000000
```

- ▶ All correlations are high: x's with punt (good) and with each other (bad, at least confusing).
- ▶ What to do? Probably do just as well to pick one variable, say right since kickers are right-footed.

```
Just right
   punting.2 <- lm(punt ~ right, data = punting)</pre>
   summary(punting.2)
   Call:
   lm(formula = punt ~ right, data = punting)
   Residuals:
       Min 1Q Median
                                 3Q
                                         Max
   -15.7576 -11.0611 0.3656 7.8890 19.0423
   Coefficients:
              Estimate Std. Error t value Pr(>|t|)
   (Intercept) -3.6930 25.2649 -0.146 0.886
   right 1.0427 0.1692 6.162 7.09e-05 ***
   Signif. codes:
   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Comparing R-squareds

```
summary(punting.1)$r.squared
```

[1] 0.7781401

summary(punting.2)\$r.squared

[1] 0.7753629

Basically no difference. In regression (over), right significant:

Regression results

tidy(punting.2)

But...

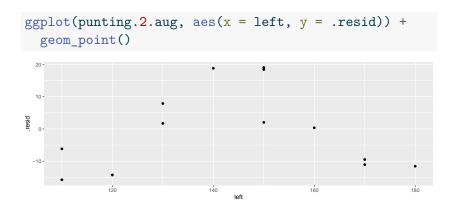
- ▶ Maybe we got the *form* of the relationship with left wrong.
- Check: plot residuals from previous regression (without left) against left.
- Residuals here are "punting distance adjusted for right leg strength".
- If there is some kind of relationship with left, we should include in model.
- Plot of residuals against original variable: augment from broom.

Augmenting punting.2

punting.2 %>% augment(punting) -> punting.2.aug
punting.2.aug

```
A tibble: 13 x 10
    left right punt
                       fred .fitted
                                      .resid
                                                 .hat
                                                      .sigma
   <dbl> <dbl> <dbl> <dbl> <dbl>
                               <dbl>
                                        <dbl>
                                                dbl>
                                                       <dbl>
 1
     170
            170
                 162.
                         171
                                174. -11.1
                                               0.157
                                                         13.5
 2
     130
            140
                 144
                         136
                                142.
                                        1.72
                                               0.0864
                                                         14.0
 3
     170
            180
                 174.
                         174
                                184.
                                       -9.49
                                               0.244
                                                         13.6
 4
                 164.
                                        0.366
     160
            160
                         161
                                163.
                                               0.101
                                                         14.0
 5
     150
            170
                 192
                         159
                                174.
                                       18.4
                                               0.157
                                                         12.5
 6
     150
            150
                 172.
                         151
                                153.
                                       19.0
                                               0.0778
                                                         12.5
 7
     180
            170
                 162
                         174
                                174. -11.6
                                               0.157
                                                         13.4
 8
     110
            110
                 105.
                         111
                                 111.
                                      -6.17
                                               0.305
                                                         13.8
 9
     110
            120
                 106.
                         114
                                121. -15.8
                                               0.2
                                                         12.9
10
     120
            130
                 118.
                         126
                                132. -14.3
                                                         13.1
                                               0.127
11
     140
            120
                 140.
                         129
                                121.
                                       18.8
                                               0.2
                                                         12.3
12
     130
            140
                 150.
                         136
                                 142.
                                        7.89
                                               0.0864
                                                         13.8
```

Residuals against left



Comments

- There is a *curved* relationship with left.
- ▶ We should add left-squared to the regression (and therefore put left back in when we do that):

```
punting.3 <- lm(punt ~ left + I(left^2) + right,
  data = punting
)</pre>
```

Regression with left-squared

```
summary(punting.3)
Call:
lm(formula = punt ~ left + I(left^2) + right, data = punting)
Residuals:
    Min 1Q Median
                              30
                                     Max
-11.3777 -5.3599 0.0459 4.5088 13.2669
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -4.623e+02 9.902e+01 -4.669 0.00117 **
left 6.888e+00 1.462e+00 4.710 0.00110 **
I(left^2) -2.302e-02 4.927e-03 -4.672 0.00117 **
right 7.396e-01 2.292e-01 3.227 0.01038 *
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 7.931 on 9 degrees of freedom
Multiple R-squared: 0.9352, Adjusted R-squared: 0.9136
F-statistic: 43.3 on 3 and 9 DF, p-value: 1.13e-05
```

Comments

- ► This was definitely a good idea (R-squared has clearly increased).
- We would never have seen it without plotting residuals from punting.2 (without left) against left.
- Negative slope for leftsq means that increased left-leg strength only increases punting distance up to a point: beyond that, it decreases again.