Multivariate analysis of variance (MANOVA)

# Multivariate analysis of variance

- Standard ANOVA has just one response variable.
- What if you have more than one response?
- Try an ANOVA on each response separately.
- But might miss some kinds of interesting dependence between the responses that distinguish the groups.

# **Packages**

```
library(car) # may need to install first
library(tidyverse)
library(MVTests) # also may need to install
```

### Small example

- Measure yield and seed weight of plants grown under 2 conditions: low and high amounts of fertilizer.
- Data (fertilizer, yield, seed weight):

```
url <- "http://ritsokiguess.site/datafiles/manova1.txt"</pre>
hilo <- read delim(url, " ")
```

```
## Rows: 8 Columns: 3
## -- Column specification -----
```

## Delimiter: " "

## chr (1): fertilizer ## dbl (2): yield, weight

## ## i Use `spec()` to retrieve the full column specification for

 2 responses, yield and seed weight. Multivariate analysis of variance (MANOVA)

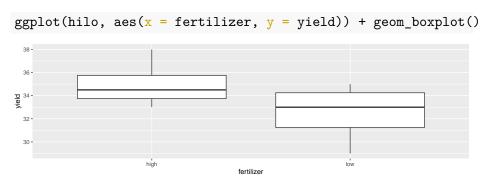
## i Specify the column types or set `show\_col\_types = FALSE`

### The data

#### hilo

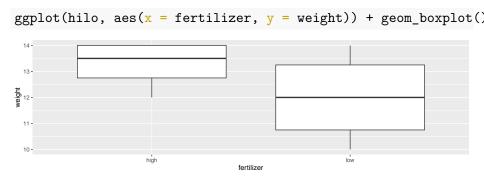
fertilizer	yield	weight	
low	34	10	
low	29	14	
low	35	11	
low	32	13	
high	33	14	
high	38	12	
high	34	13	
high	35	14	

# Boxplot for yield for each fertilizer group



Yields overlap for fertilizer groups.

### Boxplot for weight for each fertilizer group



Weights overlap for fertilizer groups.

# ANOVAs for yield and weight

```
hilo.y <- aov(yield ~ fertilizer, data = hilo)
summary(hilo.y)

## Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer 1 12.5 12.500 2.143 0.194

## Residuals 6 35.0 5.833

hilo.w <- aov(weight ~ fertilizer, data = hilo)
summary(hilo.w)
```

```
## Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer 1 3.125 3.125 1.471 0.271
## Residuals 6 12.750 2.125
```

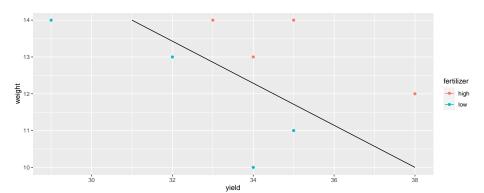
Neither response depends significantly on fertilizer. But...

# Plotting both responses at once

- Have two response variables (not more), so can plot the response variables against each other, labelling points by which fertilizer group they're from.
- ullet First, create data frame with points (31,14) and (38,10) (why? Later):

• Then plot data as points, and add line through points in d:

# The plot



#### Comments

- Graph construction:
  - Joining points in d by line.
  - geom\_line inherits colour from aes in ggplot.
  - Data frame d has no fertilizer (previous colour), so have to unset.
- Results:
  - High-fertilizer plants have both yield and weight high.
  - True even though no sig difference in yield or weight individually.
  - Drew line separating highs from lows on plot.

#### MANOVA finds multivariate differences

• Is difference found by diagonal line significant? MANOVA finds out.

```
response <- with(hilo, cbind(yield, weight))
hilo.1 <- manova(response ~ fertilizer, data = hilo)
summary(hilo.1)</pre>
```

```
## Df Pillai approx F num Df den Df Pr(>F)
## fertilizer 1 0.80154 10.097 2 5 0.01755 *
## Residuals 6
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

 Yes! Difference between groups is diagonally, not just up/down (weight) or left-right (yield). The yield-weight combination matters.

### Strategy

- Create new response variable by gluing together columns of responses, using cbind.
- Use manova with new response, looks like 1m otherwise.
- With more than 2 responses, cannot draw graph. What then?
- If MANOVA test significant, cannot use Tukey. What then?
- Use discriminant analysis (of which more later).

### Another way to do MANOVA

# using Manova from package car: hilo.2.lm <- lm(response ~ fertilizer, data = hilo)

```
hilo.2 <- Manova(hilo.2.lm)
summary(hilo.2)
##
## Type II MANOVA Tests:
##
## Sum of squares and products for error:
        vield weight
## yield 35 -18.00
## weight -18 12.75
## Term: fertilizer
## Sum of squares and products for the hypothesis:
         yield weight
## yield 12.50 6.250
## weight 6.25 3.125
## Multivariate Tests: fertilizer
                   Df test stat approx F num Df den Df Pr(>F)
                   1 0.801542 10.09714 2 5 0.017546 *
## Pillai
## Wilks 1 0.198458 10.09714 2 5 0.017546 *
## Hotelling-Lawley 1 4.038855 10.09714 2 5 0.017546 *
                    1 4.038855 10.09714 2
## Roy
                                                      5 0.017546 *
```

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#### Comments

- Same result as small-m manova.
- Manova will also do repeated measures, coming up later.

# **Assumptions**

- normality of each response variable within each treatment group
  - this is actually multivariate normality, with correlations
- equal spreads: each response variable has same variances and correlations (with other response variables) within each treatment group. Here:
  - yield has same spread for low and high fertilizer
  - weight has same spread for low and high fertilizer
  - correlation between yield and weight is same for low and high fertilizer
- ullet test equal spread using Box's M test
  - a certain amount of unequalness is OK, so only a concern if P-value from M-test is very small (eg. less than 0.001).

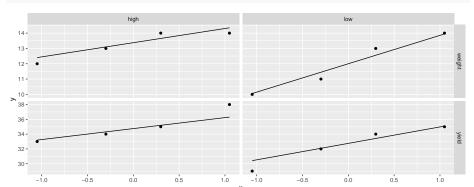
# Assumptions for yield-weight data

For normal quantile plots, need "extra-long" with all the data values in one column:

There are only four observations per response variable - treatment group combination, so graphs are not very informative (over):

# The plots





#### Box M test

- Make sure package MVTests loaded first.
- inputs:

##

- the response matrix (or, equivalently, the response-variable columns from your dataframe)
- the column with the grouping variable in it (most easily gotten with \$).

```
hilo %>% select(yield, weight) -> numeric_values summary(BoxM(numeric_values, hilo$fertilizer))
```

```
##
## Chi-Squared Value = 1.002964 , df = 3 and p-value: 0.801
```

No problem at all with unequal spreads.

Box's M Test

### Another example: peanuts

- Three different varieties of peanuts (mysteriously, 5, 6 and 8) planted in two different locations.
- Three response variables: y, smk and w.

```
u <- "http://ritsokiguess.site/datafiles/peanuts.txt"
peanuts.orig <- read_delim(u, " ")</pre>
```

```
## Rows: 12 Columns: 6
```

```
## -- Column specification -----
```

```
## Delimiter: " "
```

```
##
```

```
## i Use `spec()` to retrieve the full column specification fo
## i Specify the column types or set `show col types = FALSE`
```

### The data

#### peanuts.orig

obs	location	variety	у	smk	w
1	1	5	195.3	153.1	51.4
2	1	5	194.3	167.7	53.7
3	2	5	189.7	139.5	55.5
4	2	5	180.4	121.1	44.4
5	1	6	203.0	156.8	49.8
6	1	6	195.9	166.0	45.8
7	2	6	202.7	166.1	60.4
8	2	6	197.6	161.8	54.1
9	1	8	193.5	164.5	57.8
10	1	8	187.0	165.1	58.6
11	2	8	201.5	166.8	65.0
12	2	8	200.0	173.8	67.2

# Setup for analysis

## [6,] 195.9 166.0 45.8

```
peanuts.orig %>%
  mutate(
    location = factor(location),
    variety = factor(variety)
  ) -> peanuts
response <- with(peanuts, cbind(y, smk, w))
head(response)
##
                smk w
## [1.] 195.3 153.1 51.4
   [2,] 194.3 167.7 53.7
## [3,] 189.7 139.5 55.5
## [4,] 180.4 121.1 44.4
## [5,] 203.0 156.8 49.8
```

# Analysis (using Manova)

##

```
peanuts.2 <- Manova(peanuts.1)
summary(peanuts.2)
##
## Type II MANOVA Tests:
## Sum of squares and products for error:
                 smk
      104.205 49.365 76.480
## smk 49.365 352.105 121.995
       76 480 121 995 94 835
## Term: location
## Sum of squares and products for the hypothesis:
        0.7008333 -10.6575
                              7.129167
## smk -10.6575000 162.0675 -108.412500
        7 1291667 -108 4125 72 520833
##
## Multivariate Tests: location
                  Df test stat approx F num Df den Df Pr(>F)
## Pillai
                 1 0.893484 11.18432
                                            3 4 0.020502 *
                   1 0.106516 11.18432 3 4 0.020502 *
## Wilks
## Hotelling-Lawley 1 8.388243 11.18432 3 4 0.020502 *
## Rov
                   1 8.388243 11.18432
                                                   4 0.020502 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

peanuts.1 <- lm(response ~ location \* variety, data = peanuts)</pre>

#### Comments

- Interaction not quite significant, but main effects are.
- Combined response variable (y,smk,w) definitely depends on location and on variety
- Weak dependence of (y,smk,w) on the location-variety combination.
- Understanding that dependence beyond our scope right now.

### Normal quantile plots

Setup for normal quantile plots

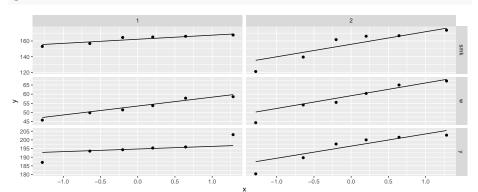
Make the plots (save to show on next pages)

```
ggplot(d, aes(sample = xvals)) + stat_qq() +
   stat_qq_line() +
   facet_grid(xname ~ location, scales = "free") -> g1

ggplot(d, aes(sample = xvals)) + stat_qq() +
   stat_qq_line() +
   facet_grid(xname ~ variety, scales = "free") -> g2
```

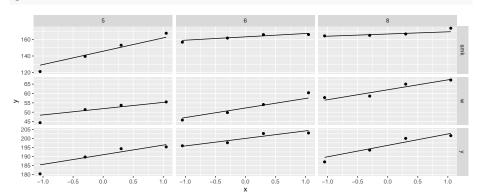
# The plot for location





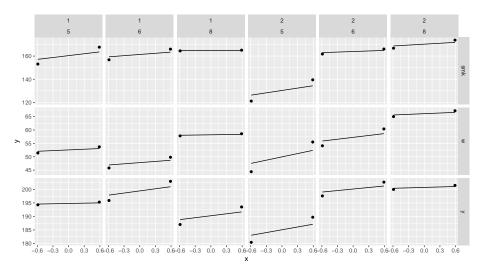
# The plot for variety

g2



### Alternative: show both these plots at once:

# The combined plot



#### Comments

- this time there are only six observations per location and four per variety, so normality is still difficult to be confident about
- ullet y at location 1 seems to be the worst for normality (long tails / outliers), and maybe y at location 2 is skewed left, but the others are not bad
- there is some evidence of unequal spread (slopes of lines), but is it bad enough to worry about? (Box M-test, over).

#### Box's M tests

##

• One for location, one for variety:

Box's M Test

```
summary(BoxM(response, peanuts$location))

## Box's M Test
##

## Chi-Squared Value = 12.47797 , df = 6 and p-value: 0.0521
summary(BoxM(response, peanuts$variety))
```

```
##
## Chi-Squared Value = 10.56304 , df = 12 and p-value: 0.567
```

 Neither of these P-values is low enough to worry about. (Remember, the P-value here has to be *really* small to indicate a problem.)