

Open problem - Better privacy guarantees for larger groups

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1 Defintion and Problem

Definition 1 (Group-wise zero-concentrated differential privacy). *Assume possible datasets consist of records from domain U , and U can be partitioned into k fixed, disjoint groups U_1, \dots, U_k . Let $v, \xi : \mathcal{D} \rightarrow \mathbb{R}^k$ be two functions associating a dataset to a vector of privacy budgets (one per group). We say a mechanism \mathcal{M} satisfies v, ξ -group-wise zero-concentrated differential privacy (zCDP) if, for any two datasets D, D' differing in the addition or removal of a record in U_i , and for all $\alpha > 1$, we have:*

$$\begin{aligned} D_\alpha(\mathcal{M}(D) \parallel \mathcal{M}(D')) &\leq \alpha \cdot v(D)_i + \xi(D)_i \\ D_\alpha(\mathcal{M}(D') \parallel \mathcal{M}(D)) &\leq \alpha \cdot v(D)_i + \xi(D)_i \end{aligned}$$

where D_α is the Rényi divergence of order α .

Problem: Let $r \in (0, 1]$ be an acceptable level of relative error, and k be the number of distinct, mutually-exclusive partitions of domain X . Given a dataset D , let $x(D)$ be a vector containing the count of records in each partition. The objective is to find a mechanism \mathcal{M} which takes in r, k and D , and outputs $\hat{x}(D)$ such that $\mathbb{E}[|x(D)_i - \hat{x}(D)_i|] < r \cdot x(D)_i$ for all i , and satisfies v -group-wise zCDP where $v(D)_i$ is as small as possible for all i . The privacy guarantee $v(D)_i$ should only depend on $x(D)_i$, and should be non-increasing with $x(D)_i$.

2 An Example Algorithm

Algorithm 1. Adding data-dependent noise as a post-processing step.

Require: A dataset D where each data point belongs to one of k groups, a privacy parameter ρ , and a relative error rate r .

1. Let $\sigma^2 = 1/(2\rho)$
2. For $i = 1$ to k do:
3. Let x_i be the number of people in D in group i
4. Sample $X_i \sim \mathcal{N}(x_i, \sigma^2)$

5. Sample $Y_i \sim \mathcal{N}(X_i, (rX_i)^2)$
6. end for
7. return Y_1, \dots, Y_k

2.1 Accuracy Analysis:

Computing the expectation, we have

$$\begin{aligned}\mathbb{E}[Y_i] &= \mathbb{E}[\mathbb{E}[Y_i | X_i]] \\ &= \mathbb{E}[X_i] = x_i\end{aligned}$$

And for the variance, we get

$$\begin{aligned}\mathbb{V}[Y_i] &= \mathbb{E}[\mathbb{V}[Y_i | X_i]] + \mathbb{V}[\mathbb{E}[Y_i | X_i]] \\ &= \mathbb{E}[r^2 X_i^2] + \mathbb{V}[X_i] \\ &= r^2(x_i^2 + \sigma^2) + \sigma^2 = r^2 x_i^2 + \sigma^2(1 + r^2)\end{aligned}$$

Combining both gives that

$$\begin{aligned}\mathbb{E}[|Y_i - x_i|] &\leq \sqrt{\mathbb{V}[Y_i]} \\ &= \sqrt{r^2 x_i^2 + \sigma^2(1 + r^2)} \\ &\leq r x_i + \sigma \sqrt{1 + r^2}\end{aligned}$$

2.2 Privacy Analysis:

Attempt 1: We can rewrite $Y_i = X_i Z$ as the product of two i.i.d Gaussian random variables, where $X_i \sim \mathcal{N}(x_i, \sigma^2)$ and $Z \sim \mathcal{N}(1, r^2)$.

Using Theorem 2.1 from [CYIK16], we get the exact PDF distribution of Y_i given by:

$$\begin{aligned}f_{Y_i}(y) &= \exp\left(-\frac{x_i^2}{2\sigma^2} - \frac{1}{2r^2}\right) \times \sum_{n=0}^{\infty} \sum_{m=0}^{2n} \binom{2n}{m} \frac{x_i^m y^{2n-m} |y|^{m-n}}{\pi(2n)! \sigma^{n+m+1} r^{3n-m+1}} K_{m-n}\left(\frac{|y|}{\sigma r}\right) \\ &\triangleq \exp\left(-\frac{x_i^2}{2\sigma^2} - \frac{1}{2r^2}\right) \times h(x_i, \sigma, r, y)\end{aligned}$$

where K_ν denotes the modified Bessel function of the second kind and order ν .

For a neighbouring dataset, $Y'_i = X'_i Z$ where $X'_i \sim \mathcal{N}(x_i \pm 1, \sigma^2)$ and $Z \sim \mathcal{N}(1, r^2)$.

The PDF distribution of Y'_i is then given by:

$$f_{Y'_i}(y) = \exp\left(-\frac{(x_i \pm 1)^2}{2\sigma^2} - \frac{1}{2r^2}\right) \times h(x_i \pm 1, \sigma, r, y)$$

Finally, we get that

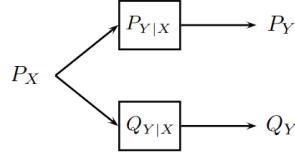
$$\begin{aligned} D_\alpha(Y_i \| Y'_i) &= \frac{1}{\alpha - 1} \log \left(\int_{-\infty}^{\infty} \left(\frac{f_{Y_i}(y)}{f_{Y'_i}(y)} \right)^\alpha f_{Y'_i}(y) dy \right) \\ &= \frac{\alpha}{\alpha - 1} \frac{1 \pm 2x_i}{2\sigma^2} + \frac{1}{\alpha - 1} \log \left(\int_{-\infty}^{\infty} \left(\frac{h(x_i, \sigma, r, y)}{h(x_i \pm 1, \sigma, r, y)} \right)^\alpha f_{Y'_i}(y) dy \right) \end{aligned}$$

TO DO: Compute an upper bound of the second term.

Attempt 2: We have $X_i \sim \mathcal{N}(x_i, \sigma^2)$ and $Y_i | X_i \sim \mathcal{N}(X_i, r^2 X_i^2)$.

On the other hand, for a neighbouring dataset, $X'_i \sim \mathcal{N}(x_i \pm 1, \sigma^2)$ and $Y'_i | X'_i \sim \mathcal{N}(X'_i, r^2 X'^2_i)$.

We recall the following theorem:



If $P_X \xrightarrow{P_{Y|X}} P_Y$ and $P_X \xrightarrow{Q_{Y|X}} Q_Y$, then

$$D_f(P_Y \| Q_Y) \leq \mathbb{E}_{X \sim P_X} [D_f(P_{Y|X} \| Q_{Y|X})].$$

where D_f is an f -divergence.

The idea is to use the theorem, where the input is the iid Gaussian pair $Z_i \triangleq (X_i, X'_i)$, the first channel is $P_{Y|Z_i} = \mathcal{N}(X_i, r^2 X_i^2)$, the second channel is $Q_{Y|Z_i} = \mathcal{N}(X'_i, r^2 X'^2_i)$. The marginals are then resp Y_i and Y'_i . Applying the Theorem gives that

$$D_f(Y_i \| Y'_i) \leq \mathbb{E}_{(X_i, X'_i)} [D_f(\mathcal{N}(X_i, r^2 X_i^2) \| \mathcal{N}(X'_i, r^2 X'^2_i))].$$

Unfortunately, the Renyi divergence is not directly an f -divergence (maybe a link could be found to apply a version of this result).

For now, let us look at just the KL ($\alpha = 1$), we get that

$$\begin{aligned} D_1(Y_i \| Y'_i) &\leq \mathbb{E}_{(X_i, X'_i)} \left[D_1(\mathcal{N}(X_i, r^2 X_i^2) \| \mathcal{N}(X'_i, r^2 X'^2_i)) \right] \\ &= \mathbb{E}_{(X_i, X'_i)} \left[2 \log \left(\frac{X'_i}{X_i} \right) + \frac{r^2 X_i^2 + (X_i - X'_i)^2}{2r^2 X'^2_i} - \frac{1}{2} \right] \\ &\leq 2 \log \left(\mathbb{E} \left[\frac{X'_i}{X_i} \right] \right) + \frac{1}{2} \left(1 + \frac{1}{r^2} \right) \mathbb{E} \left[\frac{X_i^2}{X'^2_i} \right] - \frac{1}{r^2} \mathbb{E} \left[\frac{X_i}{X'_i} \right] + \frac{1}{2} \left(\frac{1}{r^2} - 1 \right) \end{aligned}$$

which reduces to computing the expectation of the quotient of two iid Gaussian variables.

Unfortunately, these expectations do not exist as Ordinary integrals but only in a Principal Value sense (linguisticturn comment).

Plugging the formulas of the expectations in the Principal Value sense gives

$$\mathbb{E} \left[\frac{X'_i}{X_i} \right] = \mathbb{E} [X'_i] \mathbb{E} \left[\frac{1}{X_i} \right] = (x_i \pm 1) \frac{\sqrt{2}}{\sigma} F \left(\frac{x_i}{\sqrt{2}\sigma} \right)$$

$$\mathbb{E} \left[\frac{X_i}{X'_i} \right] = \mathbb{E} [X_i] \mathbb{E} \left[\frac{1}{X'_i} \right] = (x_i) \frac{\sqrt{2}}{\sigma} F \left(\frac{x_i \pm 1}{\sqrt{2}\sigma} \right)$$

$$\mathbb{E} \left[\frac{X_i^2}{X_i'^2} \right] = \mathbb{E} [X_i^2] \mathbb{E} \left[\frac{1}{X_i'^2} \right] = (\sigma^2 + x_i^2) \frac{1}{2\sigma^2} \left(\frac{\sqrt{2}(x_i \pm 1)}{\sigma} F \left(\frac{x_i \pm 1}{\sqrt{2}\sigma} \right) - 1 \right)$$

where F is the Dawson function $F(x) = e^{-x^2} \int_0^x e^{t^2} dt$.

Plugging everything in the upper bound gives an upper bound on the KL.

TO DO: Generalise to α Renyi divergence, and see the dependence on x_i (is the upper bound on the KL decreasing in x_i).

3 Another (simpler) algorithm

Algorithm 2. Adding noise directly using $x(D)_i$.

Require: A dataset D where each data point belongs to one of k groups, a privacy parameter ρ , and a relative error rate r .

1. Let $\sigma^2 = 1/(2\rho)$
2. For $i = 1$ to k do:
3. Let x_i be the number of people in D in group i
4. Sample $X_i \sim \mathcal{N}(x_i, (rx_i)^2)$
6. end for
7. return X_1, \dots, X_k

3.1 Accuracy Analysis

We have directly that

$$\begin{aligned} \mathbb{E} [|Y_i - x_i|] &\leq \sqrt{\mathbb{V}[Y_i]} \\ &= \sqrt{r^2 x_i^2} = r x_i \end{aligned}$$

3.2 Privacy Analysis

Let $X_i \sim \mathcal{N}(x_i, (rx_i)^2)$ and for a neighbouring dataset $X'_i \sim \mathcal{N}(x_i + 1, r^2 (x_i + 1)^2)$, using the formula for renyi divergence between Gaussian random variables, we get that

$$\begin{aligned}
D_\alpha(X_i, X'_i) &= 2 \log \left(\frac{x_i + 1}{x_i} \right) + \frac{1}{2(\alpha - 1)} \log \left(\frac{x_i^2 + 2x_i + 1}{x_i^2 + \alpha(2x_i + 1)} \right) + \frac{1}{2} \frac{\alpha}{r^2 x_i^2 + \alpha r^2 (2x_i + 1)} \\
&\leq \frac{1}{2} \frac{\alpha}{r^2 x_i^2} + 2 \log \left(1 + \frac{1}{x_i} \right)
\end{aligned}$$

using that $\alpha > 1$ and $r, x_i > 0$

This means that Algorithm 2 verifies v, ξ -group-wise approximate zCDP, where

$$v(D)_i = \frac{1}{2r^2 x(D)_i^2}$$

and

$$\xi(D)_i = 2 \log \left(1 + \frac{1}{x(D)_i} \right)$$

Indeed the privacy budgets are non-increasing functions of $x(D)_i$.

Comment. In the original post, one reads "Of course, directly using $x(D)_i$ to determine the scale of the noise for group i leads to a privacy loss which is data dependent, similarly to e.g. PATE [PAEGT17], and as such should be treated as a protected value."

However, any attempt that tries to first estimate $x(D)_i$ and then use the estimated (noisy) counts to add a variance (like in Algorithm 1) will too have a privacy loss that depends on $x(D)_i$ eventually. Thus the comment is not very clear to me.

References

- [CYIK16] Guolong Cui, Xianxiang Yu, Salvatore Iommelli, and Lingjiang Kong. Exact distribution for the product of two correlated gaussian random variables. *IEEE Signal Processing Letters*, 23(11):1662–1666, 2016.