

1 Proof of ACMMD expression in the synthetic setting

Recall that we have

$$\text{ACMMD}^2 = \text{MMD}(\mathbb{P}_X \mathbb{P}_Y, \mathbb{P}_X Q)^2 = \int k(p, p') (T_{11} + T_{22} - 2T_{12}) d\mathbb{P}_X(p) d\mathbb{P}_X(p')$$

where

$$T_{12} = \int k_y(y, y') p(y|p) q(y'|p') dy dy'$$

and T_{22} and T_{11} are defined similarly. For a sequence y , we define the function len given by $\text{len}(y) := \min \{i \in \mathbb{N} | y_i = \text{STOP}\}$, which intuitively returns the length of the sequence.

Computing T_{ij} As we will see, a lot of the computations are agnostic to whether we are computing T_{11}, T_{22} or T_{12} . Note that the kernel on Y writes as a product

$$k_y(y, y') = e^{-\lambda_y d_H(y, y')} = e^{-\lambda_y \sum_{i=0}^{\infty} \delta(y_i \neq y'_i)} = \prod_{i=0}^{\infty} e^{-\lambda_y \delta(y_i \neq y'_i)} = \prod_{i=0}^{\max(\text{len}(y), \text{len}(y'))} e^{-\lambda_y \delta(y_i \neq y'_i)}$$

let us define the following events

$$\begin{aligned} F(m) &:= \left\{ \min(\text{len}(y), \text{len}(y')) = m \right\} \\ G(m, \delta m) &:= \left\{ \max(\text{len}(y), \text{len}(y')) = m + \delta m \right\} \end{aligned}$$

which we further break down as

$$\begin{aligned} F_1(m) &= \{\text{len}(y) = m\} \cap \{\text{len}(y') > m\} \\ F_2(m) &= \{\text{len}(y) > m\} \cap \{\text{len}(y') = m\} \\ F_3(m) &= \{\text{len}(y) = m\} \cap \{\text{len}(y') = m\} \\ \implies F(m) &= F_1(m) \cup F_2(m) \cup F_3(m) \end{aligned}$$

For which the following probabilities hold:

$$\begin{aligned} P(F_1(m)) &= P(\text{len}(y) = m) \times P(\text{len}(y') > m) = ((2p)^m \times (1 - 2p)) \times (2p')^{m+1} \\ P(F_2(m)) &= P(\text{len}(y') = m) \times P(\text{len}(y) > m) = ((2p')^m \times (1 - 2p')) \times (2p)^{m+1} \\ P(F_3(m)) &= P(\text{len}(y') = m) \times P(\text{len}(y) = m) = ((2p')^m \times (1 - 2p')) \times ((2p)^m \times (1 - 2p)) \\ P(G(m, \delta m) | F_1(m)) &= P(\text{len}(y') = m + \delta m | \text{len}(y) = m, \text{len}(y') > m) = (2p')^{\delta m - 1} \times (1 - 2p') \delta_{(\delta m \geq 1)} \\ P(G(m, \delta m) | F_2(m)) &= P(\text{len}(y) = m + \delta m | \text{len}(y') = m, \text{len}(y) > m) = (2p)^{\delta m - 1} \times (1 - 2p) \delta_{(\delta m \geq 1)} \\ P(G(m, \delta m) | F_3(m)) &= \delta(\delta m = 0) \end{aligned}$$

Let us note

$$E(m, \delta m, i) := F_i(m) \cap G(m, \delta m)$$

We have that $E(m, \delta m, i) \cap E(m', \delta m', j) = \emptyset$ if $(m, \delta m, i) \neq (m', \delta m', j)$.

$$\Omega = \bigcup_{m=0}^{+\infty} \bigcup_{i=1}^3 \bigcup_{\delta m=0}^{+\infty} E(m, \delta m)$$

Using the law of total probability, we have that Thus, using the law of total probability:

$$\begin{aligned}
T_{ij}(p, p') &= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \sum_{\delta m=0}^{+\infty} \mathbb{P}(E_i(m, \delta m, i)) \mathbb{E}(e^{-\lambda d_H(y, y')} | E(m, \delta m, i)) \\
&= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \sum_{\delta m=0}^{+\infty} \mathbb{P}(F_i(m) \cap G(m, \delta m)) \mathbb{E}(e^{-\lambda d_H(y, y')} | E(m, \delta m, i)) \\
&= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \mathbb{P}(F_i(m)) \sum_{\delta m=0}^{+\infty} P(G(m, \delta m) | F_i(m)) \mathbb{E}(e^{-\lambda d_H(y, y')} | E(m, \delta m, i)) \\
&= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \mathbb{P}(F_i(m)) \mathbb{E}(e^{-\lambda d_H(y, y'_m)} | F_i(m)) \\
&\quad \times \sum_{\delta m=0}^{+\infty} P(G(m, \delta m) | F_i(m)) \mathbb{E}(e^{-\lambda d_H(y_{m+1:m+\max(\delta m, 1)}, y'_{m+1:m+\max(\delta m, 1)})} | E(m, \delta m, i), p, p') \\
&= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \mathbb{P}(F_i(m)) \mathbb{E}(e^{-\lambda d_H(y, y'_m)} | F_i(m)) \times \sum_{\delta m=0}^{+\infty} P(G(m, \delta m) | F_i(m)) e^{-\lambda(\max(0, \delta m-1) + \delta(m>0))} \\
&= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \mathbb{P}(F_i(m)) \mathbb{E}(e^{-\lambda d_H(y, y'_m)} | F_i(m)) \left(\prod_{i=1}^{\max(m-1, 1)} \mathbb{E}(e^{-\lambda \delta(y_i \neq y'_i)} | F_i(m)) \right)^{\delta(m \geq 2)} \\
&\quad \times \sum_{\delta m=0}^{+\infty} P(G(m, \delta m) | F_i(m)) e^{-\lambda(\max(0, \delta m-1) + \delta(m>0))}
\end{aligned}$$

where we break down the factorized hamming distance over the sequence into the sum of the hamming distances over each coordinate, and made use of the fact that

$$d_H(y_{m:m+\delta m}, y'_{m:m+\delta m}) = \max(0, \delta m - 1) + \delta(m > 0)$$

conditioned on $F_i(m)$ and $G(m, \delta m)$. The disjunction of cases is necessary in order to not count the term 0^{th} term twice in the event when $m = 0$. This representation is convenient since whenever $m \geq 2$, for any $1 \leq i \leq m - 1$,

$$P(\delta(y_i, y'_i) = 1 | F_i(m)) = \frac{(pp') + (pp')}{(p+p) \times (p'+p')} = \frac{1}{2} = P(\delta(y_i, y'_i) = 0 | F_i(m))$$

meaning we have

$$\begin{aligned}
T_{ij}(p, p') &= \sum_{m=0}^{+\infty} \sum_{i=1}^3 \mathbb{E}(e^{-\lambda \delta(y_0 \neq y'_0)} | F_i(m)) \mathbb{P}(F_i(m)) \left(\frac{1 + e^{-\lambda}}{2} \right)^{\max(m-1, 0)} \\
&\quad \times \sum_{\delta m=0}^{+\infty} P(G(m, \delta m) | F_i(m)) e^{-\lambda(\max(0, \delta m-1) + \delta(m>0))}
\end{aligned}$$

Inserting the relevant event probabilities into the expression for T_{ij} , we have

$$\begin{aligned}
T_{ij}(p, p') &= \sum_{m=0}^{+\infty} \left(\frac{1+e^{-\lambda}}{2} \right)^{\max(m-1, 0)} \\
&\times \left(\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_1(m)) (2p)^m \times (1-2p)(2p')^{m+1} \times (1-2p') e^{-\lambda\delta(m>0)} \sum_{\delta m=1}^{+\infty} e^{-\lambda(\delta m-1)} (2p')^{\delta m-1} \right. \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_2(m)) (2p')^m \times (1-2p')(2p)^{m+1} \times (1-2p) e^{-\lambda\delta(m>0)} \sum_{\delta m=1}^{+\infty} e^{-\lambda(\delta m-1)} (2p)^{\delta m-1} \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_3(m)) (2p')^m \times (1-2p')(2p)^m (1-2p) \Big) \\
&= \sum_{m=0}^{+\infty} \left(\frac{1+e^{-\lambda}}{2} \right)^{\max(m-1, 0)} \\
&\times \left(\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_1(m)) (2p)^m \times (1-2p)(2p')^{m+1} \times (1-2p') e^{-\lambda\delta(m>0)} \sum_{\delta m=0}^{+\infty} e^{-\lambda\delta m} (2p')^{\delta m} \right. \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_2(m)) (2p')^m \times (1-2p')(2p)^{m+1} \times (1-2p) e^{-\lambda\delta(m>0)} \sum_{\delta m=0}^{+\infty} e^{-\lambda\delta m} (2p)^{\delta m} \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_3(m)) (2p')^m \times (1-2p')(2p)^m (1-2p) \Big) \\
&= \sum_{m=0}^{+\infty} \left(\frac{1+e^{-\lambda}}{2} \right)^{\max(m-1, 0)} \\
&\times \left(\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_1(m)) (2p)^m \times (1-2p)(2p')^{m+1} \times (1-2p') \times \frac{e^{-\lambda\delta(m>0)}}{1-2p'e^{-\lambda}} \right. \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_2(m)) (2p')^m \times (1-2p')(2p)^{m+1} \times (1-2p) \times \frac{e^{-\lambda\delta(m>0)}}{1-2pe^{-\lambda}} \\
&+ \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_3(m)) (2p')^m \times (1-2p')(2p)^m (1-2p) \Big)
\end{aligned}$$

Now, some simplification arise when $m \geq 1$. Indeed, in that case, $\mathbb{E}(e^{-\lambda\delta(y_0, y'_0)} | F_i(m))$ is independent of i . Noting $T_{ij}^1(p, p')$ the sum of the terms for $m \geq 1$, we thus have

$$\begin{aligned}
T_{ij}^1(p, p') &= \sum_{m=1}^{+\infty} \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F(m)) \left(\frac{1+e^{-\lambda}}{2} \right)^{m-1} \\
&\times \left((2p)^m \times (1-2p)(2p')^{m+1} \times (1-2p') \times \frac{e^{-\lambda}}{1-2p'e^{-\lambda}} \right. \\
&+ (2p')^m \times (1-2p')(2p)^{m+1} \times (1-2p) \times \frac{e^{-\lambda}}{1-2pe^{-\lambda}} \\
&+ (2p')^m \times (1-2p')(2p)^m (1-2p) \Big)
\end{aligned}$$

Noting A_{ij} the term $\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F(m))$, which is constant for all $m \geq 1$

$$\begin{aligned}
T_{ij}^1(p, p') &= A_{ij} \sum_{m=1}^{+\infty} \left(\frac{1+e^{-\lambda}}{2} \right)^{m-1} \\
&\quad \times \left((2p)^m \times (1-2p)(2p')^{m+1} \times (1-2p') \times \frac{e^{-\lambda}}{1-2p'e^{-\lambda}} \right. \\
&\quad + (2p')^m \times (1-2p')(2p)^{m+1} \times (1-2p) \times \frac{e^{-\lambda}}{1-2pe^{-\lambda}} \\
&\quad \left. + (2p')^m \times (1-2p')(2p)^m(1-2p) \right) \\
&= A_{ij}(1-2p)(1-2p')4pp' \left(\frac{2p'e^{-\lambda}}{1-2p'e^{-\lambda}} + \frac{2pe^{-\lambda}}{1-2pe^{-\lambda}} + 1 \right) \sum_{m=0}^{+\infty} (4pp'(1+e^{-\lambda})/2)^m \\
&= A_{ij} \times \frac{(1-2p)(1-2p')4pp'}{1-4pp'(1+e^{-\lambda})/2} \left(\frac{2p'e^{-\lambda}}{1-2p'e^{-\lambda}} + \frac{2pe^{-\lambda}}{1-2pe^{-\lambda}} + 1 \right) \\
&= C \times A_{ij}
\end{aligned}$$

where

$$C(p, p') = \frac{(1-2p)(1-2p')4pp'}{1-4pp'(1+e^{-\lambda})/2} \left(\frac{2p'e^{-\lambda}}{1-2p'e^{-\lambda}} + \frac{2pe^{-\lambda}}{1-2pe^{-\lambda}} + 1 \right)$$

is a constant does not depend on i, j . We compute the $m = 0$ sum, noted $T_{ij}^0(p, p')$. We have

$$\begin{aligned}
T_{ij}^0(p, p') &= \left(\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_1(0)) \times (1-2p)(2p') \times (1-2p') \times \frac{1}{1-2p'e^{-\lambda}} \right. \\
&\quad + \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_2(0)) \times (1-2p')(2p) \times (1-2p) \times \frac{1}{1-2pe^{-\lambda}} \\
&\quad \left. + \mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_3(0)) \times (1-2p')(1-2p) \right)
\end{aligned}$$

And we need to compute the terms $\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_i(0))$ individually.

i=1, i=2 For $i = 1$, we must have $y_0 \neq y'_0$, since $y_0 = \text{STOP}$, and $\text{len}(y') > 0$. Thus, $\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_1(0)) = e^{-\lambda}$. Similarly, $\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_2(0)) = e^{-\lambda}$.

i=3 In that case, we must have $y_0 = y'_0 = \text{STOP}$, since $\text{len}(y) = \text{len}(y') = 0$. Thus, $\mathbb{E}(e^{-\lambda\delta(y_0 \neq y'_0)} | F_3(0)) = 1$.

Putting this together, we have

$$T_{ij}^0(p, p') = (1-2p)(1-2p') \left(\frac{2p'e^{-\lambda}}{1-2p'e^{-\lambda}} + \frac{2pe^{-\lambda}}{1-2pe^{-\lambda}} + 1 \right)$$

With that notation, we have:

$$\begin{aligned} \text{ACMMD}(\mathbb{P}_\downarrow, \mathbb{Q}_\downarrow) &= \int k(p, p') C(p, p') (A_{11} + A_{22} - 2A_{12}) d\mathbb{P}_X(p) d\mathbb{P}_X(p') \\ &\quad + \int k(p, p') (T_{11}^0(p, p') + T_{22}^0(p, p') - 2T_{12}^0(p, p')) d\mathbb{P}_X(p) d\mathbb{P}_X(p') \\ &= \int k(p, p') C(p, p') (A_{11} + A_{22} - 2A_{12}) d\mathbb{P}_X(p) d\mathbb{P}_X(p') \end{aligned}$$

since T_{ij}^0 does not depend on i, j . We can narrow the variation down even further: by noting $p_{ij}^A = P(\delta(y_i \neq y'_i) = 0 | F(m))$ (resp $p_{ij}^B = P(\delta(y_i \neq y'_i) = 0 | F(0))$), since $\mathbb{E}(e^{-\lambda\epsilon}) = p(\epsilon = 0)(1 - e^{-\lambda}) + e^{-\lambda}$ if ϵ is a Bernoulli random variable,

$$\text{ACMMD}(\mathbb{P}_\downarrow, \mathbb{Q}_\downarrow) = \int k(p, p') C(p, p') (1 - e^{-\lambda}) (p_{11}^A + p_{22}^A - 2p_{12}^A) d\mathbb{P}_X(p) d\mathbb{P}_X(p')$$

We now compute the probabilities p_{ij}^A for $i, j \in \{1, 2\}$. In every case, such p_{ij}^A can be written as:

$$p_{ij}^A = \frac{P(y_0 = y'_0 = A) + P(y_0 = y'_0 = B)}{P(\{y_0 \in \{A, B\}\} \cap \{y'_0 \in \{A, B\}\})} = \frac{P(y_0 = y'_0 = A) + P(y_0 = y'_0 = B)}{4pp'}$$

and we have

$$\begin{aligned} p_{11}^A &= \frac{pp' + pp'}{4pp'} = \frac{1}{2} \\ p_{22}^A &= \frac{(p + \Delta p)(p' + \Delta p) + (p - \Delta p)(p' - \Delta p)}{4pp'} = \frac{2pp' + 2\Delta p^2}{4pp'} \\ p_{12}^A &= \frac{(p)(p' + \Delta p) + (p)(p' - \Delta p)}{4pp'} = \frac{1}{2} \\ \implies p_{11}^A + p_{22}^A - 2p_{12}^A &= \frac{2pp' + 2\Delta p^2}{4pp'} - \frac{1}{2} = \frac{2\Delta p^2}{4pp'} \end{aligned}$$

Putting it together We thus have

$$\text{ACMMD}(\mathbb{P}_\downarrow, \mathbb{Q}_\downarrow) = \int C(p, p') k(p, p') (1 - e^{-\lambda}) \frac{2\Delta p^2}{4pp'} d\mathbb{P}_X(p) d\mathbb{P}_X(p')$$

Recalling that

$$C(p, p') = \frac{(1 - 2p)(1 - 2p')4pp'}{1 - 4pp'(1 + e^{-\lambda})/2} \left(\frac{2p'e^{-\lambda}}{1 - 2p'e^{-\lambda}} + \frac{2pe^{-\lambda}}{1 - 2pe^{-\lambda}} + 1 \right)$$

yields the desired result.