#### Homework 4

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### 1 Convex Functions

### log-sum-exponential function convexity

*Proof.* To prove that the log-sum-exp function is convex, we will use the second-order condition of convexity, which states that a function is convex if and only if its Hessian matrix is positive semi-definite.

The log-sum-exp function is defined as:

$$f(\mathbf{x}) = \log \left( \sum_{i=1}^{n} \exp(x_i) \right)$$

where **x** =  $(x_1, x_2, ..., x_n)^{\top}$ .

Let's compute the first and second derivatives of  $f(\mathbf{x})$  with respect to  $x_i$ :

First derivative:

$$\frac{\partial f(\mathbf{x})}{\partial x_i} = \frac{\exp(x_i)}{\sum_{j=1}^n \exp(x_j)}$$

Second derivative:

$$\frac{\partial^2 f(\mathbf{x})}{\partial x_i \partial x_j} = \frac{\exp(x_i) \exp(x_j)}{\left(\sum_{k=1}^n \exp(x_k)\right)^2} - \delta_{ij} \frac{\exp(x_i)}{\sum_{k=1}^n \exp(x_k)}$$

where  $\delta_{ij}$  is the Kronecker delta, which is equal to 1 if i=j and 0 otherwise.

Let's examine the Hessian matrix,  $H \in \mathbb{R}^{n \times n}$ , where  $H_{ij} = \frac{\partial^2 f(\mathbf{x})}{\partial x_i \partial x_j}$ , which is the second derivative of the log-sum-exp function. We want to show that H is positive semi-definite.

Let  $\mathbf{v} \in \mathbb{R}^n$  be an arbitrary non-zero vector. We have to show that  $\mathbf{v}^\top H \mathbf{v} \ge 0$  for all  $\mathbf{v}$ .

$$\mathbf{v}^{\top} H \mathbf{v} = \sum_{i=1}^{n} \sum_{j=1}^{n} \nu_{i} \left( \frac{\exp(x_{i}) \exp(x_{j})}{\left(\sum_{k=1}^{n} \exp(x_{k})\right)^{2}} - \delta_{ij} \frac{\exp(x_{i})}{\sum_{k=1}^{n} \exp(x_{k})} \right) \nu_{j}$$

Let's define:

$$A = \sum_{i=1}^{n} v_i \frac{\exp(x_i)}{\sum_{k=1}^{n} \exp(x_k)}$$

Thus, we can rewrite the expression as:

$$\mathbf{v}^{\mathsf{T}}H\mathbf{v} = A^2 - B$$

where:

$$B = \sum_{i=1}^{n} v_i^2 \frac{\exp(x_i)}{\sum_{k=1}^{n} \exp(x_k)}$$

Using the Cauchy-Schwarz inequality, we have:

$$\left(\sum_{i=1}^{n} v_{i} \frac{\exp(x_{i})}{\sum_{k=1}^{n} \exp(x_{k})}\right)^{2} \leq \left(\sum_{i=1}^{n} v_{i}^{2}\right) \left(\sum_{i=1}^{n} \left(\frac{\exp(x_{i})}{\sum_{k=1}^{n} \exp(x_{k})}\right)^{2}\right)$$

We can divide it on both sides of the inequality by  $\sum_{i=1}^{n} v_i^2$ , obtaining:

$$\sum_{i=1}^{n} \frac{\left(\frac{\exp(x_{i})}{\sum_{k=1}^{n} \exp(x_{k})}\right)^{2}}{\sum_{i=1}^{n} \nu_{i}} \leq \sum_{i=1}^{n} \left(\frac{\exp(x_{i})}{\sum_{k=1}^{n} \exp(x_{k})}\right)^{2}$$

We can multiply  $\sum_{i=1}^{n} v_i^2 > 0$  on both sides of the inequality, obtaining:

$$A^2 < B$$

Now, we have shown that  $A^2 \leq B$ , which means that  $\mathbf{v}^\top H \mathbf{v} = A^2 - B \geq 0$ . Therefore, the Hessian matrix H is positive semi-definite, and the log-sum-exp function  $f(\mathbf{x})$  is convex.

# the objective of logistic regression for binary classification convexity

*Proof.* To prove that the objective function of logistic regression for binary classification  $f(w) = \frac{1}{n} \sum_{i=1}^{n} \log(1 + \exp(-y_i w^{\mathsf{T}} x_i)) + \frac{\lambda}{2} \|w\|_2^2$  is convex, we will first prove the convexity of  $h(s) = \log(1 + \exp(s))$ .

First let's compute the first and second derivatives of h(s).

First derivative:

$$\frac{dh(s)}{ds} = \frac{1}{1 + \exp(s)} \cdot \exp(s)$$

Second derivative:

$$\frac{d^2h(s)}{ds^2} = \frac{d}{ds} \left( \frac{\exp(s)}{1 + \exp(s)} \right) = \frac{\exp(s)(1 + \exp(s)) - \exp^2(s)}{(1 + \exp(s))^2} = \frac{\exp(s)}{(1 + \exp(s))^2}$$

Notice that the denominator is  $(1 + \exp(s))^2$  is always positive for all  $s \in \mathbb{R}$ , so we only need to analyze the numerator  $\exp(s)$ . Since the exponential function  $\exp(s)$  is always positive, so the second derivative is non-negative for all real values of s, which means that the function  $h(s) = \log(1 + \exp(s))$  is convex.

By composition with affine function rule that preserves convexity,

$$\log(1 + \exp(-y_i w^{\mathsf{T}} x_i))$$

is convex. So

$$\frac{1}{n} \sum_{i=1}^{n} \log(1 + \exp(-y_i w^{\mathsf{T}} x_i))$$

is convex based on the non-negative weighted sum rule and non-negative multiple rules that preserve convexity.

 $\|w\|_2^2$  is convex from the slide (least-squares objective), and  $\frac{\lambda}{2}\|w\|_2^2$  is convex based on nonnegative multiple rule that preserves convexity.

So

$$\frac{1}{n} \sum_{i=1}^{n} \log(1 + \exp(-y_i w^{\mathsf{T}} x_i)) + \frac{\lambda}{2} \|w\|_2^2$$

is convex based on non-negative weighted sum rule that preserves convexity.

Therefore, the objective function of logistic regression for binary classification f(w) is convex.

# the objective of support vector machine convexity

*Proof.* To prove that the objective function  $f(w) = \frac{1}{n} \sum_{i=1}^{n} \max(0, 1 - y_i w^{\mathsf{T}} x_i) + \frac{\lambda}{2} \|w\|_2^2$  of the support vector machine is convex, we will first establish the convexity of  $q_i(s) = \max(0, 1 - y_i s^{\mathsf{T}} x_i)$ . Let  $g(s) = y_i s^{\mathsf{T}} x_i$  and h(s) = 0. Both g(s) and h(s) are affine functions with respect to s, so they are convex (affine functions are both convex and concave).

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Now  $q_i(s) = \max(g(s), h(s))$ . Based on the pointwise maximum rule that preserves convexity, the pointwise maximum of convex functions is convex, thus  $q_i(s)$  is convex.

We just showed  $\max(0, 1 - y_i w^{\mathsf{T}} x_i)$  is convex, and

$$\frac{1}{n}\sum_{i=1}^{n} \max(0, 1 - y_i w^{\mathsf{T}} x_i)$$

is convex based on the non-negative weighted sum rule and non-negative multiple rules that preserve convexity.

In a previous proof, we have shown  $\frac{\lambda}{2} ||w||_2^2$  is convex. So

$$\frac{1}{n} \sum_{i=1}^{n} \max(0, 1 - y_i w^{\mathsf{T}} x_i) + \frac{\lambda}{2} \|w\|_2^2$$

is convex based on the non-negative weighted sum rule that preserves convexity.

Therefore, the objective function of the support vector machine f(w) is convex.

# 2 the constrained version of Ridge Regression

## Does strong duality hold?

Yes.

*Proof.* To analyze whether strong duality holds for the constrained ridge regression problem, we can check if the problem satisfies Slater's constraint qualification (SCQ). The ridge regression problem is a convex optimization problem, so if it satisfies the SCQ, then strong duality holds. The constrained ridge regression problem can be written as:

$$\min_{w \in \mathbb{R}^d} \frac{1}{2} \|\Phi w - y\|_2^2 \quad \text{s.t.} \quad \|w\|_2^2 \le s$$

Here, the objective function is convex since it is a quadratic function with positive coefficients, and the constraint is also convex as it represents a Euclidean ball. Slater's constraint qualification requires the existence of a strictly feasible point, i.e., a point that strictly satisfies the inequality constraint. In this case, we can choose w = 0, which satisfies the constraint as  $\|0\|_2^2 = 0 < s$ , assuming s > 0. Thus, the SCQ is satisfied, and strong duality holds.

Now derive the KKT conditions for the optimal solution  $w^*$ . The Lagrangian of the problem is given by:

$$\mathcal{L}(w, v) = \frac{1}{2} \|\Phi w - y\|_2^2 + v \left(\frac{1}{2} \|w\|_2^2 - s\right)$$

where  $v \ge 0$  is the Lagrange multiplier. The KKT conditions consist of the following:

1. The gradient of Lagrangian with respect to  $\Phi^{\top}(\Phi w - \gamma)$  vanishes:

$$\nabla_{w} \mathcal{L}(w, v) = \Phi^{\top} (\Phi w - y) + v w = 0$$

2. Primal constraints:

$$||w||_{2}^{2} \leq s$$

3. Dual constraints:

$$v \ge 0$$

4. Complementary slackness:

$$v\left(\frac{1}{2}\|w\|_2^2 - s\right) = 0$$

These KKT conditions characterize the optimal solution  $w^*$  of the constrained ridge regression problem. Solving these conditions can help find the optimal  $w^*$  and the corresponding Lagrange multiplier v.

#### Does a close-formed solution exist?

For the constrained ridge regression problem, there isn't a closed-form solution like in the unconstrained ridge regression case. However, we can use an algorithm to compute the optimal solution.

- 1. Initialize point and step size
- 2. Calculate the gradient of the objective function at the current point.
- 3. Update the point using the gradient.
- 4. Project the updated point onto the constraint set.
- 5. If the algorithm has converged, stop. Otherwise, continue for another iteration and go back to step 2.

# 3 The equivalence between Max Entropy Model and the Logistic Regression

We want to show that the Maximum Entropy Model is equivalent to the multi-class logistic regression model.

*Proof.* To show that the Maximum Entropy Model with feature function  $f_j(x_i) = [x_i]_j$  is equivalent to the multi-class logistic regression model without regularization.

The optimization problem for the Maximum Entropy Model is

$$\max_{p(y|x_i)} - \sum_{i=1}^{n} \sum_{y=1}^{K} p(y|x_i) \ln p(y|x_i)$$
s.t. 
$$\sum_{y=1}^{K} p(y|x_i) = 1$$

$$\frac{1}{n} \sum_{i=1}^{n} \delta(y, y_i) [x_i]_j = \frac{1}{n} \sum_{i=1}^{n} p(y|x_i) [x_i]_j, \quad j = 1, ..., d, \quad y = 1, ..., K$$

where  $\delta(y, y_i)$  is equal to 1 if  $y_i = y$ , and 0 otherwise.

We use the Lagrangian dual theory to solve the constrained optimization problem:

$$\mathcal{L}(p,\lambda,\mu) = -\sum_{i=1}^{n} \sum_{y=1}^{K} p(y|x_i) \ln p(y|x_i) + \sum_{i=1}^{n} \lambda_i \left( \sum_{y=1}^{K} p(y|x_i) - 1 \right) + \sum_{y=1}^{K} \sum_{j=1}^{d} \mu_{j,y} \left( \frac{1}{n} \sum_{i=1}^{n} \delta(y,y_i) [x_i]_j - \frac{1}{n} \sum_{i=1}^{n} p(y|x_i) [x_i]_j \right)$$

Taking the derivative with respect to  $p(y|x_i)$  and setting it to zero, we get:

$$\frac{\partial \mathcal{L}}{\partial p(y|x_i)} = -\ln p(y|x_i) - 1 + \lambda_i + \frac{\mu_{j,y}[x_i]_j}{n} = 0$$

Then solving for  $p(y|x_i)$ :

$$p(y|x_i) = \exp\left(\lambda_i + \frac{\mu_{j,y}[x_i]_j}{n} - 1\right)$$

To satisfy the constraint  $\sum_{y=1}^{K} p(y|x_i) = 1$ , we can normalize the probabilities:

$$p(y|x_i) = \frac{\exp\left(\lambda_i + \frac{\mu_{j,y}[x_i]_j}{n} - 1\right)}{\sum_{y'=1}^K \exp\left(\lambda_i + \frac{\mu_{j,y'}[x_i]_j}{n} - 1\right)}$$

Comparing this expression of the multi-class logistic regression model:

$$p(y|x_i) = \frac{\exp(w_y^{\mathsf{T}} x_i)}{\sum_{y'=1}^K \exp(w_{y'}^{\mathsf{T}} x_i)}$$

We can see that both expressions have the same form, where the correspondence between the parameters is:

$$\frac{\mu_{j,y}[x_i]_j}{n} = w_y^{\mathsf{T}} x_i$$

Therefore, we have shown that the Maximum Entropy Model with the given feature function is equivalent to the multi-class logistic regression model without regularization.  $\Box$