

# Regularized regression

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#### Basic idea

- 1. Fit a regression model
- 2. Penalize (or shrink) large coefficients

#### **Pros:**

- Can help with the bias/variance tradeoff
- · Can help with model selection

#### Cons:

- May be computationally demanding on large data sets
- · Does not perform as well as random forests and boosting

#### A motivating example

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \epsilon$$

where  $X_1$  and  $X_2$  are nearly perfectly correlated (co-linear). You can approximate this model by:

$$Y = \beta_0 + (\beta_1 + \beta_2)X_1 + \epsilon$$

#### The result is:

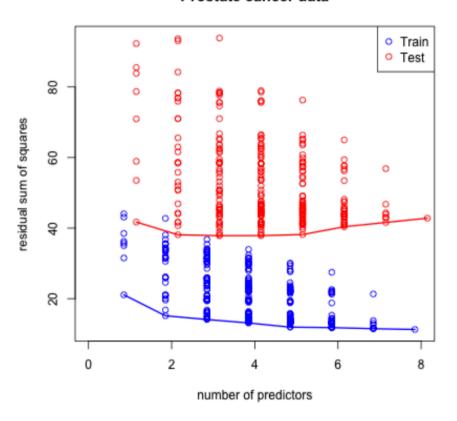
- You will get a good estimate of Y
- The estimate (of Y) will be biased
- · We may reduce variance in the estimate

#### **Prostate cancer**

```
library(ElemStatLearn); data(prostate)
str(prostate)
```

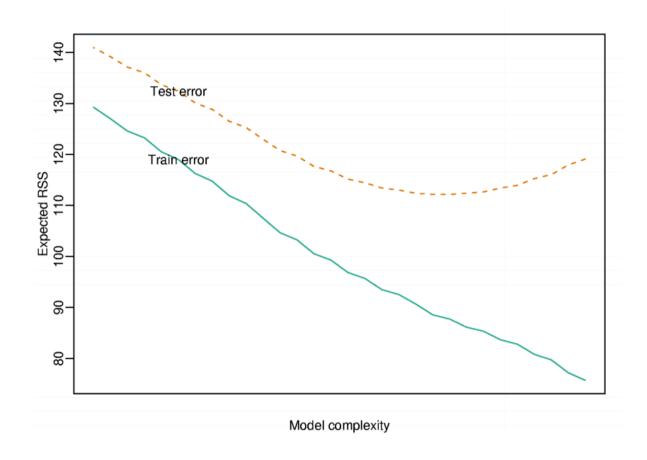
#### **Subset selection**

#### Prostate cancer data



Code here

## Most common pattern



http://www.biostat.jhsph.edu/~ririzarr/Teaching/649/

### Model selection approach: split samples

- No method better when data/computation time permits it
- · Approach
  - 1. Divide data into training/test/validation
  - 2. Treat validation as test data, train all competing models on the train data and pick the best one on validation.
  - 3. To appropriately assess performance on new data apply to test set
  - 4. You may re-split and reperform steps 1-3
- Two common problems
  - Limited data
  - Computational complexity

## Decomposing expected prediction error

Assume  $Y_i = f(X_i) + \epsilon_i$ 

$$EPE(\lambda) = E\left[\left\{Y - \hat{f}_{\lambda}(X)\right\}^{2}\right]$$

Suppose  $\hat{f}_{\lambda}$  is the estimate from the training data and look at a new data point  $X = x^*$ 

$$E\Big[\{Y - \hat{f}_{\lambda}(x^*)\}^2\Big] = \sigma^2 + \{E[\hat{f}_{\lambda}(x^*)] - f(x^*)\}^2 + var[\hat{f}_{\lambda}(x_0)]$$

= Irreducible error + Bias<sup>2</sup> + Variance

http://www.biostat.jhsph.edu/~ririzarr/Teaching/649/

http://www.cbcb.umd.edu/~hcorrada/PracticalML/

#### Another issue for high-dimensional data

```
small = prostate[1:5,]
lm(lpsa ~ .,data =small)
```

```
Call:
lm(formula = lpsa ~ ., data = small)
Coefficients:
(Intercept)
                  lcavol
                              lweight
                                                            lbph
                                                                                       lcp
                                               age
                                                                          svi
     9.6061
                  0.1390
                              -0.7914
                                            0.0952
                                                              NA
                                                                           NA
                                                                                        NA
                            trainTRUE
    gleason
                   pgg45
    -2.0871
                      NA
                                   NA
```

# Hard thresholding

- Model  $Y = f(X) + \epsilon$
- Set  $\hat{f}_{\lambda}(x) = x'\beta$
- Constrain only  $\lambda$  coefficients to be nonzero.
- Selection problem is after chosing  $\lambda$  figure out which  $p \lambda$  coefficients to make nonzero

### Regularization for regression

If the  $\beta_i$ 's are unconstrained:

- · They can explode
- And hence are susceptible to very high variance

To control variance, we might regularize/shrink the coefficients.

$$PRSS(\beta) = \sum_{j=1}^{n} (Y_j - \sum_{i=1}^{m} \beta_{1i} X_{ij})^2 + P(\lambda; \beta)$$

where *PRSS* is a penalized form of the sum of squares. Things that are commonly looked for

- Penalty reduces complexity
- Penalty reduces variance
- Penalty respects structure of the problem

# Ridge regression

Solve:

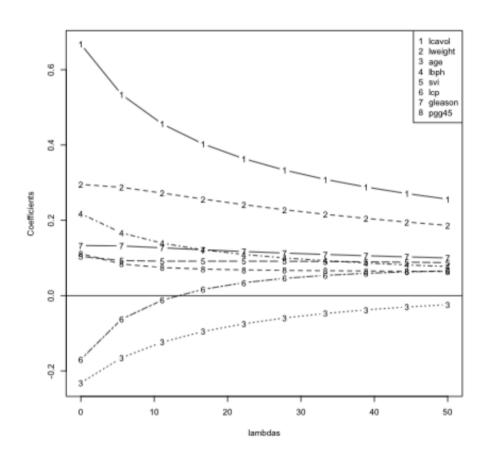
$$\sum_{i=1}^{N} \left( y_i - \beta_0 + \sum_{j=1}^{p} x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2$$

equivalent to solving

 $\sum_{i=1}^{N} \left( y_i - \beta_0 + \sum_{j=1}^{p} x_{ij} \beta_j \right)^2 \text{ subject to } \sum_{j=1}^{p} \beta_j^2 \leq s \text{ where } s \text{ is inversely proportional to } \lambda$ 

Inclusion of  $\lambda$  makes the problem non-singular even if  $X^TX$  is not invertible.

## Ridge coefficient paths



### Tuning parameter $\lambda$

- $\lambda$  controls the size of the coefficients
- $\lambda$  controls the amount of {\bf regularization}
- As  $\lambda \to 0$  we obtain the least square solution
- As  $\lambda \to \infty$  we have  $\hat{\beta}_{\lambda=\infty}^{ridge} = 0$

#### Lasso

$$\sum_{i=1}^{N} \left( y_i - \beta_0 + \sum_{j=1}^{p} x_{ij} \beta_j \right)^2 \text{ subject to } \sum_{j=1}^{p} |\beta_j| \le s$$

also has a lagrangian form

$$\sum_{i=1}^{N} \left( y_i - \beta_0 + \sum_{j=1}^{p} x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$$

For orthonormal design matrices (not the norm!) this has a closed form solution

$$\hat{\beta}_j = sign(\hat{\beta}_j^0)(|\hat{\beta}_j^0 - \gamma)^+$$

but not in general.

# Notes and further reading

- Hector Corrada Bravo's Practical Machine Learning lecture notes
- Hector's penalized regression reading list
- · Elements of Statistical Learning
- In caret methods are:
  - ridge
  - lasso
  - relaxo