



$$Var[X_*|x] = r \frac{1-p}{p^2} = \hat{\theta}_{MSE} \cdot \frac{1}{p} = \frac{\beta+n+1}{\beta+n} \hat{\theta}_{MSE} \in [1, 2]$$

Thus, the posterior predictive distribution has variance up to 2x the Poisson (i.e. more spread out or less sure of where the realization will be).

curlyF: one  $N(\theta, \sigma^2) = P(X|\theta, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{1}{2\sigma^2}(X-\theta)^2}$   
 $\vec{\theta} = \begin{bmatrix} \theta \\ \sigma^2 \end{bmatrix}$

let  $\sigma^2$  be fixed / known in advance.

$$P(X|\theta, \sigma^2) \propto e^{-\frac{1}{2\sigma^2}(X-\theta)^2} = e^{-\frac{1}{2\sigma^2}(X^2 - 2X\theta + \theta^2)}$$

$$= e^{-\frac{X^2}{2\sigma^2}} e^{\frac{X\theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}} \propto e^{-\frac{X^2}{2\sigma^2}} e^{\frac{X\theta}{\sigma^2}} = e^{ax - bx^2}, \quad a = \frac{\theta}{\sigma^2}, \quad b = \frac{1}{2\sigma^2}$$

$$P(\theta|X, \sigma^2) \propto e^{-\frac{1}{2\sigma^2}(X-\theta)^2} = e^{-\frac{X^2}{2\sigma^2}} e^{\frac{X\theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}}$$

$$\propto e^{\frac{X\theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}} = e^{a\theta - b\theta^2}, \quad a = \frac{X}{\sigma^2}, \quad b = \frac{1}{2\sigma^2}$$

$$E[\theta] = \frac{a}{2b} = \frac{X}{2}, \quad Var[\theta] = \sigma^2 = \frac{1}{2b}$$

$$E[X] = \theta = \frac{a}{2b} = \frac{\frac{\theta}{\sigma^2}}{2 \cdot \frac{1}{2\sigma^2}} = \theta$$

$$Var[X] = \sigma^2 = \frac{1}{2b} = \frac{1}{2 \cdot \frac{1}{2\sigma^2}} = \sigma^2$$

$$F: \overset{iid}{\sim} N(\theta, \sigma^2) \quad X_1, \dots, X_n$$

$$P(X|\theta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{1}{2\sigma^2}(X_i - \theta)^2} = (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum (X_i - \theta)^2}$$

$$\sum (X_i - \theta)^2 = \sum X_i^2 - 2X_i\theta + \theta^2 = \sum X_i^2 - 2n\bar{X}\theta + n\theta^2$$

$$= (2\pi\sigma^2)^{-n/2} e^{-\frac{\sum X_i^2}{2\sigma^2}} e^{\frac{n\bar{X}\theta}{\sigma^2}} e^{-\frac{n\theta^2}{2\sigma^2}} = \mathcal{L}(\theta; X, \sigma^2)$$

$$\propto e^{-\frac{\sum X_i^2}{2\sigma^2}} e^{\frac{n\bar{X}\theta}{\sigma^2}}$$

$$P(\theta|X, \sigma^2) \propto e^{\frac{n\bar{X}\theta}{\sigma^2}} e^{-\frac{n\theta^2}{2\sigma^2}} = e^{a\theta - b\theta^2}, \quad a = \frac{n\bar{X}}{\sigma^2}, \quad b = \frac{n}{2\sigma^2} \Rightarrow 2b = \frac{n}{\sigma^2}$$

$$\propto N\left(\frac{a}{2b}, \frac{1}{2b}\right) = N\left(\frac{\frac{n\bar{X}}{\sigma^2}}{\frac{n}{\sigma^2}}, \frac{1}{\frac{n}{\sigma^2}}\right) = N\left(\bar{X}, \frac{\sigma^2}{n}\right)$$

All we've done thus far is probability theory and we seemingly just made random computations for fun. Now we'll do Bayesian...

$$F: X_1, \dots, X_n \overset{iid}{\sim} N(\theta, \sigma^2) \text{ with } \sigma^2 \text{ known. Let's find posterior.}$$

$$P(\theta|X, \sigma^2) \propto P(X|\theta, \sigma^2) P(\theta|\sigma^2)$$

$$= (2\pi\sigma^2)^{-n/2} e^{-\frac{\sum X_i^2}{2\sigma^2}} e^{\frac{n\bar{X}\theta}{\sigma^2}} e^{-\frac{n\theta^2}{2\sigma^2}} P(\theta|\sigma^2)$$

$$\propto e^{\frac{n\bar{X}\theta}{\sigma^2}} e^{-\frac{n\theta^2}{2\sigma^2}} P(\theta|\sigma^2)$$

$$\propto e^{a\theta - b\theta^2} \left\{ e^{\alpha\theta - \beta\theta^2} \right\} \Rightarrow P(\theta|\sigma^2) = N\left(\frac{\alpha}{2\beta}, \frac{1}{2\beta}\right)$$

$$= e^{(a+\alpha)\theta - (b+\beta)\theta^2}$$

$$\propto N\left(\frac{a+\alpha}{2(b+\beta)}, \frac{1}{2(b+\beta)}\right)$$

$$= N\left(\frac{\frac{n\bar{X}}{\sigma^2} + \alpha}{\frac{n}{\sigma^2} + 2\beta}, \frac{1}{\frac{n}{\sigma^2} + 2\beta}\right) = N\left(\frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}, \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}\right)$$

Traditionally...  
 $\alpha = \frac{\mu_0}{\tau^2}, \beta = \frac{1}{2\tau^2}$   
 $P(\theta|\sigma^2) = N(\mu_0, \tau^2)$

The normal-normal conjugate model (where sigsq is assumed fixed).

$$P(\theta|\sigma^2) = N(\mu_0, \tau^2) \Rightarrow P(\theta|X, \sigma^2) = N\left(\frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}, \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}\right)$$

Point estimation

$$\hat{\theta}_{MSE} = E[\theta|X, \sigma^2] = \frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

$$\hat{\theta}_{MAE} = Med[\theta|X, \sigma^2] = \frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

$$\hat{\theta}_{MAP} = Mode[\theta|X, \sigma^2] = \frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

$$\hat{\theta}_{MSE} = \hat{\theta}_{MAE} = \hat{\theta}_{MAP}$$

Credible Regions

$$CR_{\theta, 1-\alpha_0} = \left[ q_{norm}\left(\frac{\alpha_0}{2}, \frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}, \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}\right), q_{norm}\left(1 - \frac{\alpha_0}{2}, \frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}, \frac{n}{\sigma^2} + \frac{1}{\tau^2}\right) \right]$$

Hypothesis Tests

$$H_1: \theta < \theta_0 \Rightarrow H_0: \theta \geq \theta_0$$

$$p_{val} = P(H_0|X, \sigma^2) = \int_{\theta_0}^{\infty} P(\theta|X, \sigma^2) d\theta = 1 - p_{norm}(\theta_0, \frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}, \frac{n}{\sigma^2} + \frac{1}{\tau^2})$$

Let's calculate the MLE (this was on HW1)

$$\mathcal{L}(\theta; X, \sigma^2) = (2\pi\sigma^2)^{-n/2} e^{-\frac{\sum X_i^2}{2\sigma^2} + \frac{n\bar{X}\theta}{\sigma^2} - \frac{n\theta^2}{2\sigma^2}}$$

$$\mathcal{L}(\theta; X, \sigma^2) = \ln(\mathcal{L}) = -\frac{\sum X_i^2}{2\sigma^2} + \frac{n\bar{X}\theta}{\sigma^2} - \frac{n\theta^2}{2\sigma^2}$$

$$\mathcal{L}'(\theta; X, \sigma^2) = \frac{n\bar{X}}{\sigma^2} - \frac{n\theta}{\sigma^2} \stackrel{set}{=} 0 \Rightarrow n\bar{X} = n\theta \Rightarrow \hat{\theta}_{MLE} = \bar{X}$$