

Assume Jeffrey's Prior $P(\theta, \sigma^2) \propto \frac{1}{\sigma^2}$

$$\begin{aligned}
 P(x|x) &\propto \int_0^\infty \int_{\mathbb{R}} k(x|\theta, \sigma^2) k(\theta, \sigma^2|x) d\theta d\sigma^2 \\
 &= \int_0^\infty \int_{\mathbb{R}} \frac{1}{\sqrt{\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\theta)^2} (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)s^2/2}{\sigma^2}} e^{-\frac{h}{2\sigma^2}(\theta-\bar{x})^2} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)s^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{1}{2\sigma^2} \left(\underbrace{x^2 - 2x\theta + \theta^2}_{x^2 - 2x\theta + (n+1)\theta^2 - 2n\theta\bar{x} + n\bar{x}^2} \right)} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)s^2/2 + x^2/n + n\bar{x}^2/n}{\sigma^2}} \int_{\mathbb{R}} e^{\frac{x^2 + n\bar{x}^2}{\sigma^2} \theta - \frac{n+1}{2\sigma^2} \theta^2} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} \underbrace{\int_{\mathbb{R}} \frac{\sqrt{\pi}}{\frac{n+1}{2\sigma^2}} e^{\frac{(x^2 + n\bar{x}^2)^2}{\sigma^2} / \frac{n+1}{2\sigma^2}} d\theta}_{\beta} d\sigma^2 \quad \left(\int_{\mathbb{R}} e^{a\theta - b\theta^2} d\theta = \sqrt{\frac{\pi}{b}} e^{a^2/4b} \right) \\
 &= \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)s^2/2 + x^2/n + n\bar{x}^2/n - (x^2 + n\bar{x}^2)^2 / (2(n+1))}{\sigma^2}} (\sigma^2)^{\frac{1}{2}} \sqrt{\frac{\pi}{n+1}} d\sigma^2 \\
 &\propto \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{b}{\sigma^2}} d\sigma^2 = \Gamma(\alpha) \tau^{-\alpha} = \Gamma\left(\frac{n}{2}\right) \tau^{-\alpha} \propto \tau^{-\alpha}
 \end{aligned}$$

$$= \left(\frac{(n-1)s^2}{2} + \frac{x^2}{n} + \frac{n\bar{x}^2}{n} - \frac{(x^2 + n\bar{x}^2)^2}{2(n+1)} \right)^{-\frac{n}{2}} = \left(a x^2 + b x + c \right)^{-\frac{n}{2}}$$

$$a = \frac{1}{2} - \frac{1}{2(n+2)} = \frac{1}{2} \left(1 - \frac{1}{n+1} \right) = \frac{1}{2} \frac{n}{n+1}$$

$$b = -\frac{2n\bar{x}}{2(n+2)} = -\frac{n\bar{x}}{n+1}, \quad c = \frac{(n-1)s^2}{2} + \frac{n\bar{x}^2}{2} - \frac{n^2\bar{x}^2}{2(n+2)} = \frac{1}{2} \left((n-1)s^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1} \right)$$

$$= \left(\frac{1}{a} \right)^{n/2} \left(\frac{1}{a} \right)^{-n/2} \left(a x^2 + b x + c \right)^{-n/2} = \left(\frac{1}{a} \right)^{-n/2} \left(x^2 + \frac{b}{a} x + \frac{c}{a} \right)^{-n/2}$$

$$\propto \left(x^2 + \frac{b}{a} x + \frac{c}{a} \right)^{-n/2} \propto \left(\left(x + \frac{b}{2a} \right)^2 + \frac{c}{a} - \frac{b^2}{4a^2} \right)^{-n/2} \left(\frac{1}{a} - \frac{b^2}{4a^2} \right)^{-n/2}$$

$$\propto \left(1 + \frac{\left(x + \frac{b}{2a} \right)^2}{\frac{c}{a} - \frac{b^2}{4a^2}} \right)^{-\frac{(n-1)+1}{2}} = \left(1 + \frac{1}{\frac{n-1}{2}} \frac{\left(x + \frac{b}{2a} \right)^2}{\left(\frac{c}{a} - \frac{b^2}{4a^2} \right) / \frac{1}{n-1}} \right)^{-\frac{(n-1)+1}{2}}$$

$$\propto T_n \left(\mu, s_0^2 \right) = T_{n-1} \left(\bar{x}, \frac{n+1}{n} s^2 \right) \stackrel{\text{large}}{\approx} N(\bar{x}, s^2)$$

$$-\frac{b}{2a} = \frac{\frac{n\bar{x}}{n+1}}{\frac{1}{n+1}} = \bar{x}$$

$$\frac{c}{a} = \frac{\frac{1}{2} \left((n-1)s^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1} \right)}{\frac{1}{2} \frac{n}{n+1}} = \frac{(n-1)(n+1)}{n} s^2 + \frac{n\bar{x}^2 + \bar{x}^2}{(n+1)} - \frac{n\bar{x}^2}{n+1}$$

$$-\frac{b^2}{4a^2} = -\left(\frac{b}{2a} \right)^2 = -\left(\bar{x} \right)^2 \Rightarrow \frac{c}{a} - \frac{b^2}{4a^2} = \frac{(n-1)(n+1)}{n} s^2$$

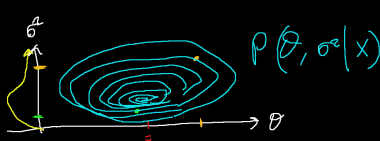
$$s_0^2 = \frac{\frac{(n-1)(n+1)}{n} s^2}{\frac{n}{n+1}} = \frac{n+1}{n} s^2$$

Under Jeffrey's prior...

$$P(\theta, \sigma^2 | x) = P(\theta | \bar{x}, \sigma^2) P(\sigma^2 | x) \quad \text{By Def of cond probability}$$

$$\text{Norm Inv Gamma}(\cdot, \cdot, \cdot) = N(\bar{x}, \frac{\sigma^2}{n}) \cdot \text{InvGamma}\left(\frac{n-1}{2}, \frac{(n-1)s^2}{2}\right)$$

You can think of a normal inverse gamma as first sampling from an InverseGamma((n-1)/2, (n-1)s^2/2) to get a sigsq value and then you use that value of sigsq to draw a theta from N(xbar, sigsq/n) and return the two-dimensional point [theta sigsq].



This can also be done the other way...

$$P(\theta, \sigma^2 | x) = \underbrace{P(\sigma^2 | x, \theta)}_{\text{InvGamma}\left(\frac{n}{2}, \frac{h_0 \sigma_0^2}{2}\right)} \underbrace{P(\theta | x)}_{T_{n-1}\left(\bar{x}, \frac{s^2}{2}\right)}$$

If we decompose the first way, we draw theta from N(xbar, sigsq/n) and thus sigsq must be known. What if we break this by instead of using the Jeffrey's prior, use

$$P(\theta) = N(\mu_0, \tau^2) \quad \text{and} \quad P(\sigma^2) = \text{InvGamma}\left(\frac{h_0}{2}, \frac{h_0 \sigma_0^2}{2}\right)$$

These were the two priors we began with when we started investigating the normal likelihood model. However, it's important to note we are not allowing $\tau^2 = \sigma^2 / h_0$

What happens? The two priors are disconnected completely.

$$P(\theta, \sigma^2) = P(\theta) P(\sigma^2) \quad \text{not} \quad P(\theta | \sigma^2) P(\sigma^2)$$

Let's derive the posterior under this two-dimensional prior.

$$\begin{aligned}
 P(\theta, \sigma^2 | x) &\propto P(x | \theta, \sigma^2) P(\theta, \sigma^2) = P(x | \theta, \sigma^2) P(\theta) P(\sigma^2) \\
 &\propto k(x | \theta, \sigma^2) k(\theta) k(\sigma^2) \\
 &= (\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + h_0\sigma_0^2 + n\bar{x}^2)} e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2} e^{-\frac{h_0}{2\sigma^2}(\theta - \mu_0)^2} \\
 &= (\sigma^2)^{-\frac{n}{2} - \frac{h_0}{2} - 1} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + h_0\sigma_0^2 + n\bar{x}^2)} \underbrace{\int_{\mathbb{R}} \frac{\sqrt{\pi}}{b} e^{x^2/4b}}_{\propto N\left(\frac{h_0\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2}, \frac{1}{\frac{h_0}{\sigma^2} + \frac{1}{\tau^2}}\right)} \\
 &\propto P(\theta | x, \sigma^2) (\sigma^2)^{-\frac{n+h_0}{2}-1} e^{-\frac{A}{2\sigma^2} \left(\frac{n}{2\sigma^2} + \frac{1}{\tau^2} \right)^{-\frac{1}{2}} e^{\frac{\left(\frac{h_0\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2} \right)^2}{\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2} \right)}}} \\
 &\quad k(\sigma^2 | x) = (\sigma^2)^{-\alpha-1} e^{-\frac{\beta}{\sigma^2}} \left(\frac{\gamma}{\sigma^2} + \delta \right)^{-\frac{1}{2}} e^{\frac{\left(\frac{\gamma}{\sigma^2} + \delta \right)^2}{\left(\frac{\gamma}{\sigma^2} + \delta \right)}}
 \end{aligned}$$

Is this kernel $k(\text{sigsq} | x)$ proportional to any distribution you know? NO!!! This means we can't sample from it using the table you've seen. Get a bigger table? No... this is not a known distribution! So we're in trouble... because we can't sample from $P(\text{sigsq} | x)$ thus we can't sample from the posterior.

We need a general way to sample from kernels of unknown distributions.