

Assume Jeffrey's Prior $P(\theta, \sigma^2) \propto \frac{1}{\sigma^2}$

$$\begin{aligned}
 P(x|x) &\propto \int_0^\infty \int_{\mathbb{R}} k(x|\theta, \sigma^2) k(\theta, \sigma^2|x) d\theta d\sigma^2 \\
 &= \int_0^\infty \int_{\mathbb{R}} \frac{1}{\sqrt{\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\theta)^2} (\sigma^2)^{-\frac{1}{2}-1} e^{-\frac{(1-1)\sigma^2/2}{\sigma^2}} e^{-\frac{1}{2\sigma^2}(\theta-\bar{x})^2} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{1}{2}} (\sigma^2)^{-\frac{1}{2}-1} e^{-\frac{(1-1)\sigma^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{1}{2\sigma^2}(\underbrace{(x-\theta)^2 + n(\theta-\bar{x})^2}_{x^2 - 2x\theta + (1+1)\theta^2 - 2n\theta\bar{x} + n\bar{x}^2})} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{1}{2}-1} e^{-\frac{(1-1)\sigma^2/2 + x^2/2 + n\bar{x}^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{\frac{x^2 + n\bar{x}^2}{\sigma^2}\theta - \frac{1}{2\sigma^2}\theta^2} d\theta d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{1}{2}-1} \underbrace{e^{-\frac{(1-1)\sigma^2/2 + x^2/2 + n\bar{x}^2/2}{\sigma^2}}}_{\beta} \underbrace{\int_{\mathbb{R}} e^{\frac{x^2 + n\bar{x}^2}{\sigma^2}\theta - \frac{1}{2\sigma^2}\theta^2} d\theta}_{\frac{(x^2 + n\bar{x}^2)^2}{2\sigma^2(1+1)}} d\sigma^2 \\
 &= \int_0^\infty (\sigma^2)^{-\frac{1}{2}-1} e^{-\frac{(1-1)\sigma^2/2 + x^2/2 + n\bar{x}^2/2 - (x^2 + n\bar{x}^2)^2/(2(1+1))}{\sigma^2}} (\sigma^2)^{\frac{1}{2}} \sqrt{\frac{2\pi}{n+1}} d\sigma^2 \\
 &\propto \int_0^\infty (\sigma^2)^{-\frac{1}{2}-1} e^{-\frac{b}{\sigma^2}} d\sigma^2 = \Gamma(\alpha) \beta^{-\alpha} = \Gamma\left(\frac{1}{2}\right) \beta^{-\alpha} \propto \beta^{-\alpha} \\
 &= \left(\frac{(1-1)\sigma^2}{2} + \frac{x^2}{2} + \frac{n\bar{x}^2}{2} - \frac{(x^2 + n\bar{x}^2)^2}{2(n+1)}\right)^{-\frac{1}{2}} = \left(a x^2 + b x + c\right)^{-\frac{1}{2}} \\
 a &= \frac{1}{2} - \frac{1}{2(n+1)} = \frac{1}{2}\left(1 - \frac{1}{n+1}\right) = \frac{1}{2} \frac{n}{n+1} \\
 b &= -\frac{2n\bar{x}}{2(n+1)} = -\frac{n\bar{x}}{n+1}, \quad c = \frac{(1-1)\sigma^2}{2} + \frac{n\bar{x}^2}{2} - \frac{n^2\bar{x}^2}{2(n+1)} = \frac{1}{2}\left((1-1)\sigma^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1}\right)
 \end{aligned}$$

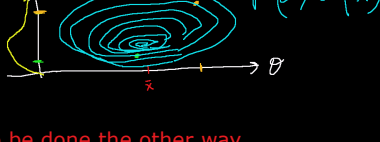
$$\begin{aligned}
 &= \left(\frac{1}{n}\right)^{1/2} \left(\frac{1}{n}\right)^{-1/2} \left(n x^2 + b x + c\right)^{-1/2} = \left(\frac{1}{n}\right)^{-1/2} \left(x^2 + \frac{b}{n} x + \frac{c}{n}\right)^{-1/2} \\
 &\propto \left(x^2 + \frac{b}{n} x + \frac{c}{n}\right)^{-1/2} \propto \left(\left(x^2 + \frac{b}{n} x + \frac{c}{n}\right)^2 + \frac{c}{n} - \frac{b^2}{4n^2}\right)^{-1/2} \left(\frac{1}{\frac{c}{n} - \frac{b^2}{4n^2}}\right)^{-1/2} \\
 &\propto \left(1 + \frac{\left(x^2 + \frac{b}{n} x + \frac{c}{n}\right)^2}{\frac{c}{n} - \frac{b^2}{4n^2}}\right)^{-1/2} = \left(1 + \frac{1}{\frac{c}{n} - \frac{b^2}{4n^2}} \left(x^2 + \frac{b}{n} x + \frac{c}{n}\right)^2\right)^{-1/2} \\
 &\propto T_n\left(\mu, s_0^2\right) = T_{n-1}\left(\bar{x}, \frac{n+1}{n} s^2\right) \stackrel{\text{large } n}{\approx} N(\bar{x}, s^2) \\
 -\frac{b}{2n} &= \frac{\frac{n\bar{x}}{n+1}}{\frac{n}{n+1}} = \bar{x} \\
 \frac{c}{n} &= \frac{\frac{1}{2}\left((1-1)\sigma^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1}\right)}{\frac{n}{n+1}} = \frac{(1-1)\sigma^2}{n} s^2 + (1+1)\bar{x}^2 - \frac{n\bar{x}^2}{n+1} \\
 -\frac{b^2}{4n^2} &= -\left(\frac{b}{2n}\right)^2 = -\left(\frac{-b}{2n}\right)^2 = -\bar{x}^2 \Rightarrow \frac{c}{n} - \frac{b^2}{4n^2} = \frac{(1-1)\sigma^2}{n} s^2 \\
 s_0^2 &= \frac{(1-1)\sigma^2}{n} s^2 = \frac{n+1}{n} s^2
 \end{aligned}$$

Under Jeffrey's prior...

$$P(\theta, \sigma^2|x) = P(\theta|\bar{x}, \sigma^2) P(\sigma^2|x) \quad \text{By Def of cond. probability}$$

$$\text{Norm Inv Gamma}(\cdot, \cdot, \cdot, \cdot) = N(\bar{x}, \frac{\sigma^2}{n}) \cdot \text{InvGamma}\left(\frac{n-1}{2}, \frac{(n-1)s^2}{2}\right)$$

You can think of a normal inverse gamma as first sampling from an `InverseGamma((n-1)/2, (n-1)s^2/2)` to get a `sigsq` value and then you use that value of `sigsq` to draw a `theta` from `N(xbar, sigsq/n)` and return the two-dimensional point `[theta, sigsq]`.



This can also be done the other way...

$$P(\theta, \sigma^2|x) = \underbrace{P(\sigma^2|x, \theta)}_{\text{InvGamma}\left(\frac{n}{2}, \frac{n\hat{\sigma}_{MLE}^2}{2}\right)} \underbrace{P(\theta|x)}_{T_{n-1}\left(\bar{x}, \frac{s^2}{2}\right)}$$

If we decompose the first way, we draw `theta` from `N(xbar, sigsq/n)` and thus `sigsq` must be known. What if we break this by instead of using the Jeffrey's prior, use

$$P(\theta) = N(\mu_0, \tau^2) \quad \text{and} \quad P(\sigma^2) = \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0\sigma_0^2}{2}\right)$$

These were the two priors we began with when we started investigating the normal likelihood model. However, it's important to note we are not allowing $\tau^2 = \sigma^2 / n_0$

What happens? The two priors are disconnected completely.

$$P(\theta, \sigma^2) = P(\theta) P(\sigma^2) \quad \text{not} \quad P(\theta|\sigma^2) P(\sigma^2)$$

Let's derive the posterior under this two-dimensional prior.

$$\begin{aligned}
 P(\theta, \sigma^2|x) &\propto P(x|\theta, \sigma^2) P(\theta, \sigma^2) = P(x|\theta, \sigma^2) P(\theta) P(\sigma^2) \\
 &\propto k(x|\theta, \sigma^2) k(\theta) k(\sigma^2) \\
 &= (\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2}((1-1)\sigma^2 + n(\bar{x}-\theta)^2)} e^{-\frac{1}{2\tau^2}(\theta-\mu_0)^2} e^{-\frac{n_0}{2}\left(\frac{\sigma^2}{\sigma_0^2}\right)} \\
 &= (\sigma^2)^{-\frac{n}{2}-\frac{n_0}{2}-1} e^{-\frac{1}{2\sigma^2}((1-1)\sigma^2 + n_0\sigma_0^2 + n\bar{x}^2)} e^{\left(\frac{n\bar{x}}{\sigma^2} + \frac{n_0}{\tau^2}\right)\theta - \left(\frac{n}{2\sigma^2} + \frac{1}{\tau^2}\right)\theta^2}
 \end{aligned}$$