	This project explores weekly matchup data from my fantasy baseball league for the 2016-2019 seasons. Parts of this analysis are motivated by Peadar Coyle's Rugby Analysis case study, and PyMC3's GLM: Logistic Regression example. The data was collected using the Yahoo Fantasy Baseball API and is available in the full_data.csv file. The league consists of 14 teams and each season usually consists of 21 weekly matchups. Scoring is based on 18 statistical categories. Weekly "scores" are based on the head-to-head matchup in each category. Category wins losses and ties are compiled each week and throughout the season to determine the league standings. The 2016-2019 seasons were chosen because the categories were consistent throughout those four seasons. They consist of the following counting (aggregate) categories as identified in the data: Batting: b_BB: Walks; b_HR: Home Runs; b_K: Strikeouts; b_NSB: Net Stolen Bases; b_R: Runs; b_RBI: Runs Batted	S,				
	In. Pitching: p_BB: Walks; p_IP: Innings Pitched; p_K: Strikeouts; p_L: Losses; p_NSVH: Net Saves and Holds; p_QS: Quality Starts; p_W: Wins. The rest are rate categories: Batting: b_AVGr: Batting Average; b_OBPr: On-Base Percentage; b_SLGr: Slugging Percentage. Pitching: p_ERAr: Earned-run Average; p_WHIPr: Walks plus Hits per Innings Pitched. As noted above, the season standings are not influenced by the overall weekly winner, and consist of an aggregate score. For examples					
	<pre>import varnings warnings.filterwarnings('ignore') # Import data directories, or load file if in the same directory # full data.csv will be included with this notebook try:</pre>					
[2]:						
[3]: [4]:	<pre># Prvot the data df_pivoted = df.pivot(index=['id'], columns='category', values = 'value') df_pivoted['won_week'] = df.groupby('id')['won_week'].mean()</pre>					
[4]:	<pre>mask = np.zeros_like(corr, dtype=np.bool_) mask[np.triu_indices_from(mask)] = True f, ax = plt.subplots(figsize=(8, 6)) cmap = sns.diverging_palette(10, 220, as_cmap=True) sns.heatmap(corr, mask=mask, cmap=cmap, vmax=0.3, linewidths=0.5, cbar_kws={"shrink": 0.5}, ax=ax) </pre> <pre> **AxesSubplot:xlabel='category', ylabel='category'> **BB-bHR-bK-bK-category', dtype=np.bool_) b</pre>					
	b_NSB - b_OBPr - b_R - b_R - b_RBI - c_D -					
	P_W - P_WHIPP - Won_week - P - P - P - P - P - P - P - P - P -					
[5]: [5]:	mp.absolute(Corr['won_week']).sort_values(ascending=raise)					
[6]:	p_QS					
	<pre>df_batting = df_pivoted.loc[:, [c for c in df_pivoted.columns if 'b_' in c]] df_pitching = df_pivoted.loc[:, [c for c in df_pivoted.columns if 'p_' in c]] counts = df_pivoted.loc[:, [c for c in df_pivoted.columns if not 'r' in c]] rates = df_pivoted.loc[:, [c for c in df_pivoted.columns if 'r' in c]] batting_counts = counts.loc[:, [c for c in counts.columns if 'b_' in c]] batting_rates = rates.loc[:, [c for c in rates.columns if 'b_' in c]] pitching_counts = counts.loc[:, [c for c in counts.columns if 'p_' in c]] df_batting_rates = df_batting.loc[:, [c for c in df_batting.columns if 'r' in c]] df_batting_rates = df_batting.loc[:, [c for c in df_batting.columns if not 'r' in c]] df_pitching_counts = df_pitching.loc[:, [c for c in df_pitching.columns if not 'r' in c]] df pitching rates = df pitching.loc[:, [c for c in df_pitching.columns if 'r' in c]]</pre>					
[7]: [8]:	<pre># sns.pairplot(df_batting)> Too large for printing # A few pitching categories for further exploration. sns.pairplot(df_pitching.iloc[:, :5]) <seaborn.axisgrid.pairgrid 0x2030be4bb50="" at=""></seaborn.axisgrid.pairgrid></pre>					
	50 40 40 20 10 10 10 10 10 10 10 10 10 10 10 10 10					
	120 100 100 100 100 100 100 100 100 100	• • • • • •				
	80	•				
	60 40 20 40 40 40 40 40 40 40 40 40 40 40 40 40	1				
	p_BB p_ERAr p_IP p_K p_K Not surprisingly, the rate categories broadly correlate with Innings Pitched. The losses category has some interesting points in it, as well, as there is a clear ERA threshold as losses increase. The strong correlation between innings pitched and strikeouts is logical, but what's notable is the lesser correlation with walks. It's also worth noting that a few of these categories may be multimodal.	10				
LO]:	<pre># The strongest category-won_week correlation sns.distplot(df_pivoted['b_OBPr']) <axessubplot:xlabel='b_obpr', ylabel="Density"> 14 12</axessubplot:xlabel='b_obpr',></pre>					
	10 -					
l1]: l1]:	<pre>sns.distplot(df_pivoted['b_R'])</pre>					
	0.05 0.04 0.02 0.01 0.00 10 20 30 40 50 60 70 80 b_R					
12]: 12]:	<pre>sns.distplot(df_pivoted['p_WHIPr'])</pre>					
	2.0 1.5 0.0 0.5 0.0 0.5 0.0 0.5 0.0 0.5 0.0 0.5 0.0 0.5 0.5					
L3]:	<pre>p_WHIPr appears to be normally distributed, but may be multimodal. sns.stripplot(x="won_week", y="b_OBPr", data=df_pivoted, jitter=True)</pre>					
	0.40 - 0.35 - 0.30 - 0.25 - False True won_week					
L4]:	<pre>sns.stripplot(x="won_week", y="b_R", data=df_pivoted, jitter=True) <axessubplot:xlabel='won_week', ylabel="b_R"> 70 -</axessubplot:xlabel='won_week',></pre>					
5]:	40 - 30 - 20 - True won_week					
L5]:	Sils.Stripprot(x-woll_week , y- p_whiri , data-dr_prvoted, jitter-irue)					
	Again, because of the large number of categories, we're seeing that there isn't a lot we can infer from these categories about how the contribute to weekly wins, though p_WHIPr shows that, like ERA, there appears to be a bit of a threshold where a WHIP above 1.75 crarely, if ever, results in a win. This may be a good category for our regression models.	-				
	<pre>import pynics as pm import arviz as az # This models won_week as the likelihood product of bernoulli n trials # By default the priors on the regressors are very vague at N(0, 1.0E-16) with pm.Model() as base_model:</pre>					
	<pre>pm.glm.GLM.from_formula("won_week ~ b_OBPr + p_WHIPr + b_R", df_pivoted, family=pm.glm.families.Binomial()) base_trace = pm.sample(init='adapt_diag') Auto-assigning NUTS sampler Initializing NUTS using adapt_diag Multiprocess sampling (2 chains in 2 jobs) NUTS: [b_R, p_WHIPr, b_OBPr, Intercept] 100.00% [4000/4000 01:02<00:00 Sampling 2 chains, 0 divergences] Sampling 2 chains for 1_000 tune and 1_000 draw iterations (2_000 + 2_000 draws total) took 72 seconds.</pre>					
L7]:	az.plot_trace(base_trace, compact=True)					
	Intercept -4 -6 -8 -7 -6 -5 -4 -3 0 200 400 600 800 -8 -0 25 -0 -0 -0 -0 -0 -0 -0 -0 -0 -0 -0 -0 -0	11/2/2/2014 11/2				
	15.0 17.5 20.0 22.5 25.0 27.5 30.0 32.5 15 0 200 400 600 800 p_WHIPr	11 (1) (1) (1) (1) (1) (1) (1) (1) (1) (
	The model converges, the effective sample size is smaller than the size of the data we provided the model, and the r_hat values are not stated as $\frac{1}{200}$ and $\frac{1}{400}$ and $\frac{1}{600}$ are not size of the data we provided the model, and the r_hat values are not size of the data we provided the model.	ot 1				
	high. Also of note, the credible intervals for each regressor do not include zero, which is an issue that other categories will present in upcoming models. az.summary (base_trace)	, to				
	<pre>p_WHIPr -3.341 0.354 -3.961 -2.655 0.010 0.007 1187.0 1104.0 1.00 b_R 0.043 0.008 0.028 0.058 0.000 0.000 1405.0 924.0 1.00 Next we'll look at how the probability of winning the week changes with On-base Percentage values. Since that category corresponds other hitting metrics, and it showed a correlation with weekly wins, this will show us how the model predicts the probability of a win o changes b_OBPr in conjunction with different values of p_WHIPr and b_R.</pre> def lm_full(trace, OBP, WHIP, R): shape = np.broadcast(OBP, WHIP, R).shape					
	<pre>shape = np.broadcast(OBP, WHIP, R).shape x_norm = np.asarray([np.broadcast_to(x, shape) for x in [OBP, WHIP, R]]) return 1 / (</pre>					
20]:	# These values were interred from the original dataset for the purpose of demonstrating the model's pred # values of the regressors lm = lambda x, samples: lm_full(samples, x, 1.52, 26) lm2 = lambda x, samples: lm_full(samples, x, 1.24, 35.5) lm3 = lambda x, samples: lm_full(samples, x, 1.02, 48)	llict				
	<pre>base_trace, eval=np.linspace(0, .75, 1000), lm=lm, samples=500, color="blue", alpha=0.15) pm.plot_posterior_predictive_glm(base_trace, eval=np.linspace(0, .75, 1000), lm=lm2, samples=500, color="green", alpha=0.15) pm.plot_posterior_predictive_glm(base_trace, eval=np.linspace(0, .75, 1000), lm=lm3, samples=500, color="red", alpha=0.15) import matplotlib.lines as mlines blue_line = mlines.Line2D(["lm"], [], color="b", label="10th percentile WHIP, R") green_line = mlines.Line2D(["lm2"], [], color="g", label="Mean values")</pre>					
	<pre>green_line = mlines.Line2D(["lm2"], [], color="g", label="Mean values") red_line = mlines.Line2D(["lm3"], [], color="r", label="90th percentile WHIP, R") plt.legend(handles=[blue_line, green_line, red_line], loc="lower right") plt.show() C:\Users\mandelson\Miniconda3\envs\mypm3env2\lib\site-packages\pymc3\plots\posteriorplot.py:59: Deprecati ing: The `plot_posterior_predictive_glm` function will migrate to Arviz in a future release. Keep up to date with `ArviZ <https: arviz="" arviz-devs.github.io=""></https:>`_ for future updates. warnings.warn(C:\Users\mandelson\Miniconda3\envs\mypm3env2\lib\site-packages\pymc3\plots\posteriorplot.py:59: Deprecati ing: The `plot_posterior_predictive_glm` function will migrate to Arviz in a future release. Keep up to date with `ArviZ <https: arviz="" arviz-devs.github.io=""></https:>`_ for future updates. warnings.warn(</pre>	ionW				
		LonW				
	lt's clear from this data that with better values for each category, the model's posterior distribution predicts a higher probability of wir the week. It's also interesting to note that the better values of WHIP and R offset lower OBP values to the point that a weekly win is higher probable right around the mean value for OBP.					
	probable right around the mean value for OBP. Other model runs were less informative, and are not worth including here. The same goes for attempts to model scaled data. One model that I did want to compare was the full dataset in a logistic regression model:					
23]:	Auto-assigning NUTS sampler Initializing NUTS using adapt_diag Multiprocess sampling (2 chains in 2 jobs) NUTS: [sd, p_BB, b_K, p_NSVH, p_L, b_NSB, p_K, p_IP, b_HR, p_W, p_QS, b_BB, b_RBI, p_ERAr, p_WHIPr, b_R, r, b_AVGr, b_OBPr, Intercept] 100.00% [4000/4000 02:28<00:00 Sampling 2 chains, 0 divergences] Sampling 2 chains for 1_000 tune and 1_000 draw iterations (2_000 + 2_000 draws total) took 160 seconds.	b_S				
23]:						
	p_WHIPr -0.130 0.122 -0.369 0.094 0.004 0.003 1109.0 1219.0 1.00 p_ERAr -0.041 0.019 -0.077 -0.007 0.001 0.000 1322.0 1377.0 1.00 b_RBI 0.003 0.002 -0.002 0.007 0.000 0.000 1993.0 1659.0 1.00 b_BB 0.001 0.005 -0.008 0.009 0.000 0.000 759.0 1155.0 1.00 p_QS 0.028 0.011 0.009 0.050 0.000 0.000 1674.0 1625.0 1.00 p_W 0.012 0.009 -0.006 0.027 0.000 0.000 1635.0 1415.0 1.00 b_HR 0.003 0.009 -0.013 0.019 0.000 0.000 1385.0 1293.0 1.00 p_IP 0.004 0.002 -0.001 0.008 0.000 0.000 1100.0 1268.0 1.00					
	p_IP 0.004 0.002 -0.001 0.008 0.000 0.000 1100.0 1268.0 1.00 p_K 0.000 0.001 -0.003 0.003 0.000 0.000 1824.0 1665.0 1.00 b_NSB 0.024 0.005 0.014 0.033 0.000 0.000 2609.0 1268.0 1.01 p_L -0.014 0.009 -0.031 0.003 0.000 0.000 1624.0 1414.0 1.00 p_NSVH 0.003 0.005 -0.007 0.013 0.000 0.000 2455.0 1385.0 1.00 b_K -0.005 0.001 -0.007 -0.002 0.000 0.000 1612.0 1533.0 1.00 p_BB -0.009 0.003 -0.015 -0.004 0.000 0.000 1417.0 1397.0 1.00 sd 0.433 0.009 0.416 0.448 0.000 0.000 2035.0 1382.0 1.00					
	Some of these coefficients have slightly higher r_hat values, or low effective sample size values. A number of them contain zero within confidence interval, though I think with more distinct priors that could be avoided.	the				
	<axessubplot:title={'center':'b_obpr'}>, <axessubplot:title={'center':'b_avgr'}>,</axessubplot:title={'center':'b_avgr'}></axessubplot:title={'center':'b_obpr'}>					
24]:	<pre><axessubplot:title={'center':'b_obpr'}>,</axessubplot:title={'center':'b_obpr'}></pre>					
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