Logistic Regression example with many independent variables:

Health study investigated epidemic outbreak of a disease spread by mosquitoes. 196 individuals from two sectors of a city were selected and the following variables were recorded:

```
Response variable Y=1 if disease was present, 0 otherwise P = Pr(Y=1) probability of disease
Age
Socioeconomic status (upper, middle, lower)

D<sub>lower</sub> =1 for status = lower (dummy variable)
D<sub>middle</sub> =1 for status = middle (dummy variable)
Sector of the city (sector 1, sector 2)
D<sub>sector2</sub> = 1 for sector =2 (dummy variable)
```

Summary of analysis

> # FIT FULL LOGISTIC MODEL

The full logistic regression model to predict probability of disease p=Pr(disease=1) is as follows:

> # logistic regression model fitted using qlm() function with family=binomial

 $log(p/(1-p)=-2.293+0.027 \ age + 0.045 \ middle + 0.253 \ lower +1.243 \ sector2 +e$ Note that the dummy variables for status are not significant.

```
> full <- glm(disease~age +st.m+st.l+sec.2, data=myd, family=binomial())</pre>
> summary(full) # display results
Deviance Residuals:
  Min 1Q Median 3Q
                              Max
-1.6576 -0.8295 -0.5652 1.0092 2.0842
Coefficients:
         Estimate Std. Error z value Pr(>|z|)
0.026991 0.008675 3.111 0.001862 **
         0.044609 0.432490 0.103 0.917849
st.m
        st.l
sec.2
        1.243630 0.352271 3.530 0.000415 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 236.33 on 195 degrees of freedom
Residual deviance: 211.22 on 191 degrees of freedom
AIC: 221.22
Number of Fisher Scoring iterations: 3
> # APPLY VARIABLE SELECTION PROCEDURES
> # Run backward selection procedure for variable selection
> # step() function works for both lm and glm objects
> step.mod=step(full, direction=c("backward"))
Start: AIC=221.22
disease ~ age + st.m + st.l + sec.2
       Df Deviance AIC
- st.m 1 211.23 219.23
      1 211.61 219.61
- st.l
           211.22 221.22
<none>
- age 1 221.26 229.26
- sec.2 1 224.22 232.22
Step: AIC=219.23
disease ~ age + st.l + sec.2
       Df Deviance AIC
      1 211.64 217.64
- st.l
<none>
           211.23 219.23
          221.30 227.30
- age 1
- sec.2 1 224.22 230.22
Step: AIC=217.64
disease ~ age + sec.2
                                     Selected model
       Df Deviance AIC
          211.64 217.64
<none>
                                      Variables that were
       1 221.60 225.60
- age
                                      removed
- sec.2 1 224.32 228.32
> #fit selected model
> fit <- glm(disease~age +sec.2, data=myd, family=binomial())</pre>
> summary(fit) # display results
glm(formula = disease ~ age + sec.2, family = binomial(), data = myd)
Deviance Residuals:
   Min 1Q Median
                            3Q
                                    Max
-1.6839 -0.8199 -0.5607 1.0093 2.0275
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
```

Number of Fisher Scoring iterations: 3

The selected logistic regression model M1 for p=Pr(disease=1) is log(p/(1-p)=-2.16+0.027 age +1.181 sector 2+e since age and sector 2 have a positive parameter, they both have a positive association with p.

> #compute goodness of fit test (likelihood ratio test)

```
> library(lmtest) #in lmtest package
> lrtest(fit)
Likelihood ratio test

Model 1: disease ~ age + sec.2
Model 2: disease ~ 1
    #Df LogLik Df Chisq Pr(>Chisq)
1    3 -105.82
2    1 -118.17 -2 24.69   4.351e-06 ***
---
Signif. codes: 0 \***' 0.001 \**' 0.01
```

Goodness of fit test:

Ho: null model b1=b2=0

H1: Model M1 with b1 and b2 \neq 0

LR statistic = 24.69 with chi-square distribution with 2 DF.

P-value is almost zero (4.351e-06). So we can conclude that null hypothesis can be rejected and current model is better than the null model.

> confint(fit) # 95% CI for the coefficients

```
Waiting for profiling to be done...
                  2.5 %
                             97.5 %
(Intercept) -2.86990940 -1.51605906
            0.01010532 0.04421365
            0.52854584 1.85407936
sec.2
> exp(coef(fit)) # compute exp(coefficients) to analyze change in odds for
changes in X
(Intercept)
                              sec.2
                    age
             1.0271756
 0.1153644
                          3.2598900
> exp(confint(fit)) # 95% CI for exp(coefficients), that is change in odds
Waiting for profiling to be done...
                 2.5 %
                          97.5 %
(Intercept) 0.05670406 0.2195755
age
            1.01015655 1.0452056
sec.2
           1.69646359 6.3858165
```

The selected logistic regression model M1 for p=Pr(disease=1) is log(p/(1-p)=-2.16+0.027 age +1.181 sector 2+e since age and sector 2 have a positive parameter, they both have a positive association with p.

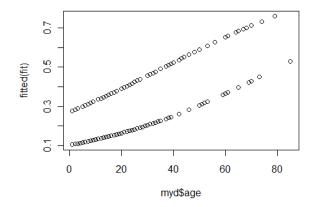
The coefficient beta for a certain variable X represents the change in $\log(\text{odds})$ for any 1-unit increase in X. Exp(beta) is the estimated odds ratio for a unit increase of X, and is equal to the odds at X+1 divided by the odds at X: odds_ratio=odds(x+1)/odds(x).

Thus, exp(beta) represents the change (increase or decrease) in odds for any 1-unit increase in X. If exp(beta)>1, the odds increase, if exp(beta)<1, the odds decrease (this is explained in week 7 slides). INTERPRETATION OF MODEL 1 PARAMETERS:

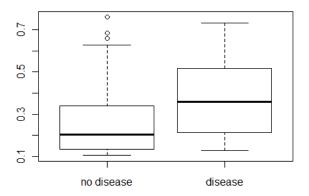
AGE: for people living in the same sector, any additional year of age increases the average odds of disease by 2.7%, and we are 95% confidence that the average increase is between 1% and 4.5% SECTOR2: for people with the same age, if they live in sector 2 the odds of disease are 326% the odds of disease in sector 1, or in other words, living in sector 2 increases the odds of disease by 226%. Also, we are 95% confidence that the average increase is between 69% and 538%.

```
> predict(fit, type="response") # predicted values
```

- > residuals(fit, type="deviance") # residuals
- > #plot of predicted probabilities vs age
- > plot(myd\$age, fitted(fit))

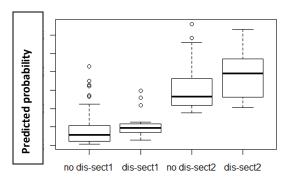


- > #boxplot of predicted probabilities by disease
- > # useful visualization for classification purposes
- > boxplot(fitted(fit)~myd\$disease, names=c("no disease", "disease"))



- > #boxplot of predicted probabilities by task success and sector
- > # useful visualization for classification purposes
- > boxplot(fitted(fit)~myd\$disease*sec.2, names=c("no dis-sect1", "dis-

sect1", "no dis-sect2", "dis-sect2"))



The boxplot shows the distribution of predicted probabilities for each observed patient by disease (whether sick or not) and sector (1 or 2). The predicted probabilities for people in sector 1 are all below 0.3 (indicating low chance of disease in sector 1) and the predicted probabilities for sector 2 are all above 0.3 (indicating higher chance of disease in sector 2).

- > #compute predicted probability of contracting disease for a 20 year old
- > # individual living in the second sector of the city
- > newd = data.frame(age=c(20), sec.2=c(1))
- > predict(fit,newdata=newd,type="response", se.fit=T) #type="response" for probabilities

\$fit 1 0.3913341

\$se.fit 1 0.05817779

\$residual.scale
[1] 1

The predicted probability for a 20 year old individual living in sector 2 of the city is computed as $\hat{p}=0.39$ with a standard error of 0.058. So we can say that the 95% prediction interval is

0.39+/- 1.96*0.058 or (0.27, 0.50) (using the asymptotic normality of predictions)

Note that the predicted value of p can be computed manually from the model:

$$\log(\frac{\hat{p}}{1-\hat{p}}) = -2.16 + 0.027 * 20 + 1.181 * 1 = -0.439 \text{ and therefore}$$

$$\hat{p} = \exp(-0.439) / (1 + \exp(-0.439) = 0.39$$