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UNIVERSITY OF NEW SOUTH WALES

SCHOOL OF MATHEMATICS AND STATISTICS

Assignment 2

Measure Theory

Author:
Adam J. Gray

Student Number:
3329798

1**2****2.1**

Suppose μ and ν are σ -finite *positive* measures on (Ω, \mathcal{F}) . Then suppose that $\mu \ll \nu$ and $\nu \ll \mu$. Then for $A \in \mathcal{F}$ we have that $\mu(A) = 0 \Rightarrow \nu(A) = 0$ and $\nu(A) = 0 \Rightarrow \mu(A)$. That is to say that $\nu(A) = 0 \Leftrightarrow \mu(A) = 0$. That is to say that ν and μ have the same null sets. This argument is symmetric so it is clear that the reverse implication also holds.

We now wish to show that there is an \mathcal{F} -measurable function g that satisfies $0 < g(\omega) < +\infty$ at each $\omega \in \Omega$ and is such that $\nu(A) = \int_A g d\mu$ for all $A \in \mathcal{F}$.

3**3.1**

Firstly note that the definition of the characteristic function is

$$\hat{\mu}_X(u) = \mathbb{E}[\exp(i\langle X, u \rangle)] \quad (1)$$

and so for the random vector cX with $c \in \mathbb{R}$ we have that

$$\hat{\mu}_{cX}(u) = \mathbb{E}[\exp(i\langle cX, u \rangle)] \quad (2)$$

$$= \mathbb{E}[\exp(i\langle X, cu \rangle)] \quad (3)$$

$$= \hat{\mu}_X(cu) \quad (4)$$

3.2**3.3**

Let $d = 1$ and let μ have the Lebesgue density,

$$f(x) = \frac{C}{(1+x^2)\log(e+x^2)}, \quad x \in \mathbb{R}. \quad (5)$$

We wish to show that $E[X]$ is not defined but $\hat{\mu}(u)$ is differentiable at 0. Firstly we show that $E[X]$ is not defined.

$$E[X] = \int_{-\infty}^{\infty} \frac{Cx}{(1+x^2)\log(e+x^2)} dx \quad (6)$$

$$= \int_{-\infty}^{\infty} \frac{C}{(x^{-1}+x)\log(e+x^2)} \quad (7)$$

note that

$$\frac{C}{(x^{-1} + x) \log(e + x^2)} \sim \frac{C}{2x \log(x)} \quad (8)$$

and

$$\int_a^\infty \frac{C}{2x \log(x)} dx, \text{ and } \int_{-\infty}^b \frac{C}{2x \log(x)} dx \quad (9)$$

do not converge so $\mathbb{E}[X]$ does not exist.

We now just have to show that $\hat{m}u(u)$ is differentiable at $u = 0$.

$$\hat{\mu}(u) = \int_{-\infty}^\infty \frac{e^{iux} C}{(1 + x^2) \log(e + x^2)} dx \quad (10)$$

and

$$\frac{d}{du} \hat{\mu}(u) \Big|_{u=0} = \int_{-\infty}^\infty \frac{ixe^{iux} C}{(1 + x^2) \log(e + x^2)} dx \quad (11)$$

4

Let μ be the binomial distribution with n trials and probability of success p , that is $\mu = \text{Bin}(n, p)$, and let ν be the Poisson distribution with mean $\lambda > 0$.

4.1

We wish to verify that $\hat{\mu}(u) = (1 - p + pe^{iu})^n$. Because the binomial distribution is just the convolution of identical independent Bernoulli distributions then we just have to verify that $(1 - p + pe^{iu})$ is the characteristic function for *Bernoulli*(p).

If ν is the Bernoulli measure then

$$\mathbb{E}[\exp(iuX)] = \sum_{k \in \{0,1\}} e^{iuk} \nu_X(k) \quad (12)$$

$$= pe^{iu} + (1 - p). \quad (13)$$

Then by repeated application of the convolution theorem we get that $\hat{\mu}(u) = (1 - p + pe^{iu})^n$.

4.2

We wish to verify that $\hat{\nu}(u) = \exp(\lambda(e^{iu} - 1))$. The probability mass function of the Poisson distribution is

$$\frac{\lambda^k}{k!} e^{-\lambda} \quad (14)$$

and thus

$$\mathbb{E}[\exp(iuX)] = \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} e^{-\lambda} e^{iuk} \quad (15)$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} (e^{iu})^k \quad (16)$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^{iu})^k}{k!} \quad (17)$$

$$= e^{-\lambda} e^{\lambda e^{iu}} \quad (18)$$

$$= e^{\lambda(e^{iu}-1)} \quad (19)$$