Spatial Data and Analysis in R

A PRISM Workshop

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Outline

Introduction

Spatial Data Prep

Spatial Autocorrelation

Regression

Spatial Regression

Discussion



Why Are We Here?

► Tobler's first law of geography:



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- ► "Everything is related to everything else, but near things are more related than distant things"



Why Are We Here?

- ► Tobler's first law of geography:
- "Everything is related to everything else, but near things are more related than distant things"
- We want to quantify how the spatial relationship between our observations affect our inferences



A Caveat

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- ▶ I will be introducing *spatial statistics* with a touch of *GIS*
- ▶ I will not be discussing GIS in depth, nor will I discuss remote sensing at all



What are Spatial Data?

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- Information (attributes) associated with a location
- Many kinds of spatial data: Points, Lines, Polygons, Raster data
- Today, we are working with polygon data



Prepping our data

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Prepping our data

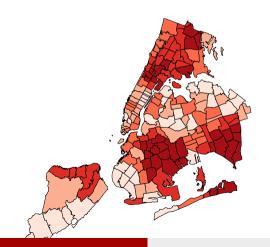
- ► Spatial data come in *shapefiles* which are really mini-databases
- ORDBMS Linking spatial and attribute data
- ▶ Six parts, all combine to create a map to represent data



Loading Our data

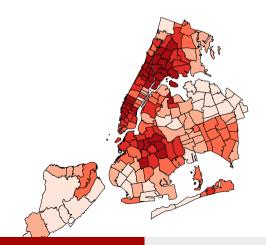


Percent Black





AIDs Rate per 1000 people



Measuring Spatial Autocorrelation

What is spatial autocorrelation?

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- What is spatial autocorrelation?
- ▶ Observations with more similar values tend to occur more closely together



Measuring Spatial Autocorrelation

- What is spatial autocorrelation?
- ▶ Observations with more similar values tend to occur more closely together
- Most common test: Moran's I



I's formula is:



Formula for Moran's I

$$I = \frac{n}{\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}} \frac{\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}(x_i - \bar{x})(x_j - \bar{x})}{\sum_{i=1}^{n} (x_i - \bar{x})^2},$$

where w_{ij} is the weight between observation i and j



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- ▶ It allows us to measure the effect neighboring observations *j* have on our observation of interest i
- ► Can be specified in a variety of ways, the simplest of which is binary ("contiguity"): 1 if observations share a boundary, 0 if they do not
- ▶ The default in R is "row standardized," where $w_{ij} = \frac{i}{\sum j}$



Creating a weights matrix in R

Running the Moran's I

Moran I statistic	Expectation	Variance	p-value
0.68	-0.01	0.00	0.00



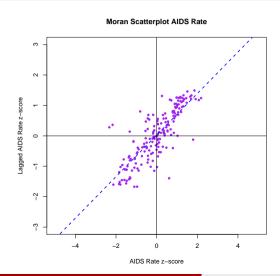
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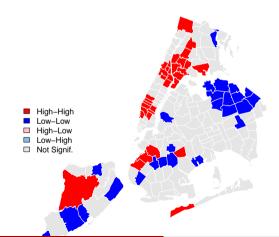


- ▶ We can calculate this using a *Local Indicator of Spatial Autocorrelation* (LISA)
- Measure how similar a value is compared to neighboring values
- ▶ While the Moran's I detects clustering, the LISA detects *clusters*





LISA Map AIDS Rate; weights: Q1



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- Addtionally, with spatial autocorrelation, our coefficients may be biased
- ▶ Two ways of handling this: Spatial Error Models, and Spatial Autoregressive models



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- ▶ The full model: $y_i = x_i \beta + \sum_{j=1}^n w_{ij} e_j + \varepsilon_i$





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- ▶ Instead of lagging the error term, lag y, the DV
- $y_i = x_i \beta + \sum_{j=1}^n w_{ij} y_j + \varepsilon_i$
- ► SAR vs. SEM



Plain OLS

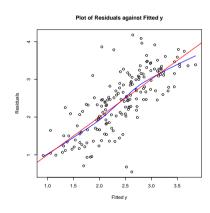
	Model 1		
(Intercept)	-3.05***		
	(0.80)		
PctWht	-0.01***		
	(0.00)		
PctHisp	0.02***		
	(0.00)		
Gini	7.97***		
	(0.71)		
PctHSEd	0.02**		
	(0.01)		
PctFemHH	0.01		
	(0.01)		
R ²	0.57		
Adj. R ²	0.55		
Num. obs.	174		
RMSE	0.54		
*** p < 0.001, ** p < 0.01, * p < 0.05			

Table: Statistical models





Did we model out our autocorrelation?



Did we model out our spatial autocorrelation?

Moran I statistic	Expectation	Variance	p-value
0.63	-0.01	0.00	0.00

No! This statistic didn't change much from before we ran our regression: .68 vs .63



Where is the residual autocorrelation?





Can we Model this Spatial Autocorrelation

One option is to use fixed effects



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- ▶ But this removes something interesting the spatial relationship



Can we Model this Spatial Autocorrelation

- One option is to use fixed effects
- ▶ But this removes something interesting the spatial relationship
- ▶ Another option is to explicitly model the spatial relationship



Spatial Autoregressive Model

	Model 1
(Intercept)	-1.52*
	(0.63)
PctWht	-0.01***
	(0.00)
PctHisp	0.01**
	(0.00)
Gini	4.48***
	(0.64)
PctHSEd	0.01
	(0.01)
PctFemHH	0.00
	(0.01)
ρ	0.54***
	(0.06)
Num. obs.	174
Parameters	8
Log Likelihood	-101.93
AIC (Linear model)	285.72
AIC (Spatial model)	219.85
LR test: statistic	67.86
LR test: p-value	0.00

A A I I I I

 $^{^{***}\}rho < 0.001,\,^{**}\rho < 0.01,\,^{*}\rho < 0.05$







Spatial Error Model

	Model 1
(Intercept)	-0.28
	(0.88)
PctWht	-0.01**
	(0.00)
PctHisp	0.01**
	(0.00)
Gini	4.25***
	(0.90)
PctHSEd	0.01
	(0.01)
PctFemHH	0.01
	(0.01)
λ	0.66***
	(0.06)
Num. obs.	174
Parameters	8
Log Likelihood	-107.85
AIC (Linear model)	285.72
AIC (Spatial model)	231.70
LR test: statistic	56.01
LR test: p-value	0.00
***p < 0.001, **p < 0.01, *p	o < 0.05

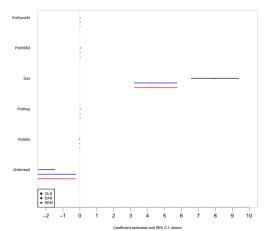
Table: Statistical models





Comparing Findings - OLS

Comparing Regression Results



Did These Resolve our Spatial Autocorrelation?

Moran I statistic	Expectation	Variance	p-value
-0.10	-0.01	0.00	0.08

YES!



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$$E(y_i|y_{-i}) = x_i\beta + \sum_{j=1}^{n} + c_{ij}[y_j - x_i\beta]$$

$$\operatorname{var}(y_i|y_{-i}) = \sigma_i^2$$





$$\sum_{Y} = (I - C)^{-1} \sum_{C}$$

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$$\sum_{Y} = (I - C)^{-1} \sum_{C}$$

$$\sum_{CAR} = \sigma^2 (I - C)^{-1} V_C$$

$$=\sigma^2 V_{CAR}$$



SAR vs. CAR

► SAR:

$$y_i \sim N(0, (I-W)^{-1}D^{\sim}(I-W)'^{-1})$$

 $\sigma_{SAR} = \sigma^2(I-W)^{-1}V_{\epsilon}(I-W')^{-1}$

SAR vs. CAR

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► CAR:

$$y_i \sim N(0, (I-C)^{-1}D)$$
$$\sum_{CAR} = \sigma^2 (I-C)^{-1} V_C$$



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- ▶ One final check for autocorrelation in your residuals

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- ► Spatial/Spatio-temporal scan statistics



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➤ Yuri Zhukov's Spatial Workshop: http://www.people.fas.harvard.edu/~zhukov/spatial.html



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References



