

# Lab 4

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## Packages

```
library(car)
library(broom)
```

## Homes Data

```
homes <- read.table("./house.txt", header=TRUE)
```

## Fit multiple linear regression model

```
m123 <- lm(Price ~ Improve + Area + Land, x=TRUE, data=homes)
summary(m123)

##
## Call:
## lm(formula = Price ~ Improve + Area + Land, data = homes, x = TRUE)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -14.856  -2.897   1.797   2.783  16.246
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  229.5069    97.9279   2.344  0.03234 *
## Improve       0.7932     0.2232   3.553  0.00265 **
## Area        13.3934     6.6878   2.003  0.06246 .
## Land         1.0104     0.6735   1.500  0.15299
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7.979 on 16 degrees of freedom
## Multiple R-squared:  0.8959, Adjusted R-squared:  0.8763
## F-statistic: 45.88 on 3 and 16 DF, p-value: 4.397e-08
```

## Correlations and variance inflation factors

```
X <- m123$x
cor(X)
```

```
## Warning in cor(X): the standard deviation is zero
```

```
##           (Intercept)  Improve      Area      Land
## (Intercept)           1         NA         NA         NA
## Improve              NA 1.0000000 0.7881460 0.7866654
## Area                 NA 0.7881460 1.0000000 0.7328524
## Land                 NA 0.7866654 0.7328524 1.0000000
```

We get a bunch of NA values for the rows and columns involving (Intercept) because our design matrix has an entire column of ones in the (Intercept) column.

```
head(X)
```

```
## (Intercept) Improve Area Land
## 1           1  94.967 2.573 162.0
## 2           1  77.860 1.628 165.0
## 3           1  81.439 1.826 165.5
## 4           1  89.592 1.965 166.0
## 5           1 122.827 2.914 174.0
## 6           1  77.317 1.612 164.5
```

As such, the resulting sample standard deviation is zero. Therefore it may be of interest to use `cor()` on the data frame rather than the design matrix for “prettier” output.

```
cor(homes)
```

```
##           Price  Improve      Area      Land
## Price  1.0000000 0.9156607 0.8489815 0.8295975
## Improve 0.9156607 1.0000000 0.7881460 0.7866654
## Area    0.8489815 0.7881460 1.0000000 0.7328524
## Land    0.8295975 0.7866654 0.7328524 1.0000000
```

```
vif(m123)
```

```
## Improve      Area      Land
## 3.516128 2.895052 2.877341
```

From *Module 4.8*, variance inflation factors measure how the correlations among the predictor variables affect the variances of the least squares estimators. If the value of the VIF is greater than 10, then there is evidence of severe multicollinearity.

Since none of the VIF values are greater than 10 here, there is insufficient evidence of severe multicollinearity.

## Partial $F$ -tests

```
anova(m123)
```

```
## Analysis of Variance Table
##
## Response: Price
##           Df Sum Sq Mean Sq  F value    Pr(>F)
## Improve    1  8202.5   8202.5  128.8261 4.589e-09 ***
## Area        1   418.5    418.5   6.5736  0.02081 *
## Land        1   143.3    143.3   2.2511  0.15299
## Residuals  16  1018.7     63.7
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The Sum Sq column represents the sequential increase of SSR (or equivalently, the sequential decrease in SSE) as the predictors **Improve**, **Area**, and **Land** are included into the model that came before it.

### Build all the models we will use

```
m1 <- lm(Price ~ Improve, data=homes)
m2 <- lm(Price ~ Area, data=homes)
m13 <- lm(Price ~ Improve + Land, data=homes)
summary(m1)

##
## Call:
## lm(formula = Price ~ Improve, data = homes)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -22.7199  -3.3395  -0.6258   5.0580  18.8687
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  376.3830    12.1101  31.080 < 2e-16 ***
## Improve       1.3513     0.1398   9.665 1.51e-08 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 9.371 on 18 degrees of freedom
## Multiple R-squared:  0.8384, Adjusted R-squared:  0.8295
## F-statistic: 93.41 on 1 and 18 DF,  p-value: 1.505e-08
summary(m2)

##
## Call:
## lm(formula = Price ~ Area, data = homes)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -23.2919  -4.7552   0.2277   9.1076  25.4814
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  405.485    12.939  31.338 < 2e-16 ***
## Area         41.364     6.068   6.816 2.21e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.32 on 18 degrees of freedom
## Multiple R-squared:  0.7208, Adjusted R-squared:  0.7053
## F-statistic: 46.46 on 1 and 18 DF,  p-value: 2.213e-06
summary(m13)

##
## Call:
## lm(formula = Price ~ Improve + Land, data = homes)
##
## Residuals:
```

```
## -18.8960 -1.3223 -0.0745 2.5029 20.9228
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 172.1617   101.6027   1.694 0.108417
## Improve      1.0184     0.2092   4.868 0.000145 ***
## Land         1.4109     0.6977   2.022 0.059171 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.657 on 17 degrees of freedom
## Multiple R-squared:  0.8698, Adjusted R-squared:  0.8544
## F-statistic: 56.77 on 2 and 17 DF,  p-value: 2.987e-08
```

### Partial $F$ -tests: round 1

```
anova(m1)
```

```
## Analysis of Variance Table
##
## Response: Price
##           Df Sum Sq Mean Sq F value    Pr(>F)
## Improve    1 8202.5  8202.5    93.41 1.505e-08 ***
## Residuals 18 1580.6    87.8
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(m1, m123)
```

```
## Analysis of Variance Table
##
## Model 1: Price ~ Improve
## Model 2: Price ~ Improve + Area + Land
##   Res.Df    RSS Df Sum of Sq    F Pr(>F)
## 1      18 1580.6
## 2      16 1018.7  2    561.88 4.4123 0.02978 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The associated hypotheses for the partial  $F$ -test are:

$$H_0 : \beta_2 = \beta_3 = 0 \quad \text{vs} \quad H_A : \text{At least one of } \beta_2, \beta_3 \text{ non-zero}$$

The  $p$ -value (0.02978) is less than 0.05 so we reject the null hypothesis. We conclude that at least one of  $\beta_2, \beta_3$  is non-zero.

## Partial $F$ -tests: round 2

```
anova(m2)

## Analysis of Variance Table
##
## Response: Price
##           Df Sum Sq Mean Sq F value    Pr(>F)
## Area         1 7051.4   7051.4   46.463 2.213e-06 ***
## Residuals    18 2731.8    151.8
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

anova(m2, m123)
```

```
## Analysis of Variance Table
##
## Model 1: Price ~ Area
## Model 2: Price ~ Improve + Area + Land
##   Res.Df    RSS Df Sum of Sq    F    Pr(>F)
## 1       18 2731.8
## 2       16 1018.7  2      1713 13.452 0.0003741 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The associated hypotheses for the partial  $F$ -test are:

$$H_0 : \beta_1 = \beta_3 = 0 \quad \text{vs} \quad H_A : \text{At least one of } \beta_1, \beta_3 \text{ non-zero}$$

The  $p$ -value (0.0003741) is less than 0.05 so we reject the null hypothesis. We conclude that at least one of  $\beta_1, \beta_3$  is non-zero.

### Partial $F$ -tests: round 3

```
anova(m13)
```

```
## Analysis of Variance Table
##
## Response: Price
##          Df Sum Sq Mean Sq  F value    Pr(>F)
## Improve    1 8202.5   8202.5 109.4437 7.96e-09 ***
## Land        1   306.5    306.5   4.0897 0.05917 .
## Residuals 17 1274.1     74.9
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(m13, m123)
```

```
## Analysis of Variance Table
##
## Model 1: Price ~ Improve + Land
## Model 2: Price ~ Improve + Area + Land
##   Res.Df    RSS Df Sum of Sq    F Pr(>F)
## 1      17 1274.1
## 2      16 1018.7  1    255.37 4.0107 0.06246 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The associated hypotheses for the partial  $F$ -test are:

$$H_0 : \beta_2 = 0 \quad \text{vs} \quad H_A : \beta_2 \neq 0$$

The  $p$ -value (0.06246) is greater than 0.05 so we fail to reject the null hypothesis. There is insufficient evidence to support the claim that  $\beta_2$  is non-zero. In other words, we can safely drop **Area** from the model.

## Brand preference data

```
muffin <- read.table("./brandpref.txt", header=TRUE)
```

## Fit multiple linear regression model

```
m12 <- lm(Liking ~ Moisture + Sweetness, x=TRUE, data=muffin)
summary(m12)
```

```
##
## Call:
## lm(formula = Liking ~ Moisture + Sweetness, data = muffin, x = TRUE)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.400 -1.762  0.025  1.587  4.200
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   37.6500     2.9961  12.566 1.20e-08 ***
## Moisture       4.4250     0.3011  14.695 1.78e-09 ***
## Sweetness     4.3750     0.6733   6.498 2.01e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.693 on 13 degrees of freedom
## Multiple R-squared:  0.9521, Adjusted R-squared:  0.9447
## F-statistic: 129.1 on 2 and 13 DF,  p-value: 2.658e-09
```

## Correlations and variance inflation factors

```
X <- m12$x
head(X)
```

```
##      (Intercept) Moisture Sweetness
## 1              1         4         2
## 2              1         4         4
## 3              1         4         2
## 4              1         4         4
## 5              1         6         2
## 6              1         6         4
```

```
cor(X)
```

```
## Warning in cor(X): the standard deviation is zero
```

```
##              (Intercept) Moisture Sweetness
## (Intercept)           1      NA      NA
## Moisture             NA       1       0
## Sweetness            NA       0       1
```

```
cor(muffin)
```

```
##              Liking  Moisture Sweetness
## Liking      1.0000000 0.8923929 0.3945807
## Moisture    0.8923929 1.0000000 0.0000000
## Sweetness   0.3945807 0.0000000 1.0000000
```

```
vif(m12)
```

```
## Moisture Sweetness
##      1      1
```

Interestingly, the correlation between **Moisture** and **Sweetness** is zero. The resulting variance inflation factors are all one. This means that the correlation between **Moisture** and **Sweetness** is not affecting the variance of the least squares estimates (scaling by a factor of 1). Since none of the variance inflation factors are greater than 10, there is no evidence of severe multicollinearity (but we already knew that).

## Brand preference data with centred predictors

```
muffin <- transform(muffin,
                     Moistcen=Moisture - mean(Moisture),
                     Sweetcen=Sweetness - mean(Sweetness))
```

## Fit multiple linear regression model with centred predictors

```
m12c <- lm(Liking ~ Moistcen + Sweetcen, x=TRUE, data=muffin)
summary(m12c)

##
## Call:
## lm(formula = Liking ~ Moistcen + Sweetcen, data = muffin, x = TRUE)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.400 -1.762  0.025  1.587  4.200
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   81.7500     0.6733  121.413  < 2e-16 ***
## Moistcen       4.4250     0.3011   14.695 1.78e-09 ***
## Sweetcen       4.3750     0.6733    6.498 2.01e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.693 on 13 degrees of freedom
## Multiple R-squared:  0.9521, Adjusted R-squared:  0.9447
## F-statistic: 129.1 on 2 and 13 DF,  p-value: 2.658e-09
```

## Correlations and variance inflation factors

```
with(muffin, cor(Moistcen, Sweetcen))
```

```
## [1] 0
```

```
Xc <- m12c$x
cor(Xc)
```

```
## Warning in cor(Xc): the standard deviation is zero
```

```
##              (Intercept) Moistcen Sweetcen
## (Intercept)           1         NA         NA
## Moistcen             NA          1          0
## Sweetcen             NA          0          1
```



```
vif(m12c)
```

```
## Moistcen Sweetcen  
##          1          1
```

The correlation of `Moistcen` and `Sweetcen` is zero. The variance inflation factors are 1, i.e. the variances of the least squares estimators are not being inflated.

## Comparisons of centred and non-centred models

```
options(pillar.sigfig=4)
```

For the code below, we will be creating lists of tibbles (tibbles are special data frames). There is some oddness that will occur due to truncating/rounding when printing these lists of tibbles, making some numbers seem unequal when they are actually equal. We will change the number of significant digits (only for tibbles) to 4 (default is 3) as a workaround.

### Fit the rest of the models

```
m1 <- lm(Liking ~ Moisture, data=muffin)  
m1c <- lm(Liking ~ Moistcen, data=muffin)  
m2 <- lm(Liking ~ Sweetness, data=muffin)  
m2c <- lm(Liking ~ Sweetcen, data=muffin)
```

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## Coefficient comparisons

```
(all_coefs <- list(m12=tidy(m12), m12c=tidy(m12c),  
                  m1=tidy(m1), m1c=tidy(m1c),  
                  m2=tidy(m2), m2c=tidy(m2c)))
```

```
## $m12  
## # A tibble: 3 x 5  
##   term      estimate std.error statistic  p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  37.65      2.996      12.57  1.200e-8  
## 2 Moisture     4.425      0.3011     14.70  1.778e-9  
## 3 Sweetness    4.375      0.6733      6.498  2.011e-5  
##  
## $m12c  
## # A tibble: 3 x 5  
##   term      estimate std.error statistic  p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  81.75      0.6733     121.4  3.017e-21  
## 2 Moistcen     4.425      0.3011     14.70  1.778e- 9  
## 3 Sweetcen     4.375      0.6733      6.498  2.011e- 5  
##  
## $m1  
## # A tibble: 2 x 5  
##   term      estimate std.error statistic    p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  50.78      4.395      11.55  0.00000001519  
## 2 Moisture     4.425      0.5980      7.399  0.000003356  
##  
## $m1c  
## # A tibble: 2 x 5  
##   term      estimate std.error statistic  p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  81.75      1.337      61.13  2.115e-18  
## 2 Moistcen     4.425      0.5980      7.399  3.356e- 6  
##  
## $m2  
## # A tibble: 2 x 5  
##   term      estimate std.error statistic    p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  68.62      8.610      7.970  0.000001431  
## 2 Sweetness    4.375      2.723      1.607  0.1304  
##  
## $m2c  
## # A tibble: 2 x 5  
##   term      estimate std.error statistic  p.value  
##   <chr>      <dbl>     <dbl>     <dbl>    <dbl>  
## 1 (Intercept)  81.75      2.723      30.02  4.127e-14  
## 2 Sweetcen     4.375      2.723      1.607  1.304e- 1
```

## ANOVA comparisons

```
(all_anova <- list(m12=tidy(anova(m12)), m12c=tidy(anova(m12c)),  
                  m1=tidy(anova(m1)), m1c=tidy(anova(m1c)),  
                  m2=tidy(anova(m2)), m2c=tidy(anova(m2c))))
```

```
## $m12  
## # A tibble: 3 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Moisture     1 1566.  1566.     215.9  1.778e-9  
## 2 Sweetness     1  306.2  306.2     42.22  2.011e-5  
## 3 Residuals    13   94.3   7.254      NA      NA  
##  
## $m12c  
## # A tibble: 3 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Moistcen     1 1566.  1566.     215.9  1.778e-9  
## 2 Sweetcen     1  306.2  306.2     42.22  2.011e-5  
## 3 Residuals    13   94.3   7.254      NA      NA  
##  
## $m1  
## # A tibble: 2 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Moisture     1 1566.  1566.     54.75  0.000003356  
## 2 Residuals    14 400.5   28.61      NA      NA  
##  
## $m1c  
## # A tibble: 2 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Moistcen     1 1566.  1566.     54.75  0.000003356  
## 2 Residuals    14 400.5   28.61      NA      NA  
##  
## $m2  
## # A tibble: 2 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Sweetness     1  306.2  306.2     2.582  0.1304  
## 2 Residuals    14 1661.  118.6      NA      NA  
##  
## $m2c  
## # A tibble: 2 x 6  
##   term      df  sumsq  meansq statistic    p.value  
##   <chr>   <int> <dbl>   <dbl>     <dbl>    <dbl>  
## 1 Sweetcen     1  306.2  306.2     2.582  0.1304  
## 2 Residuals    14 1661.  118.6      NA      NA
```