# **Macroeconometrics**

# Lecture 4 Bayesian Estimation

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**Useful distributions** 

Likelihood function

**Prior distribution** 

Posterior distribution

### Compulsory reading:

Woźniak (2021) Posterior derivations for a simple linear regression model, Lecture notes

Greenberg (2008) Chapter 4: Prior Distributions, Introduction to Bayesian Econometrics

#### Materials:

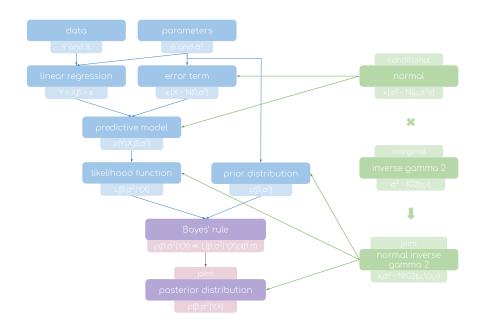
An R file L3 grphs.R for the reproduction of graphs Selected quotations from Sims (2012) — Bayesian vs. frequentist approach

## Objectives.

- ► To learn the basics of Bayesian inference
- ► To familiarise with essential probability distributions
- ► To derive Bayesian estimation results for a simple model

### Learning outcomes.

- ► Using basic Bayesian terminology
- ► Identifying parameters of transformed distributions
- ► Recognising functional form of kernels of basic distributions





Bayes' rule

$$p(\theta|Y) = \frac{p(Y|\theta)p(\theta)}{p(Y)}$$

 $p(\theta|Y)$  - posterior distribution of parameters  $\theta$  given data Y  $p(Y|\theta)$  - sampling distribution of data Y given parameters  $\theta$   $p(\theta)$  - prior distribution of parameters  $\theta$  p(Y) - marginal data density

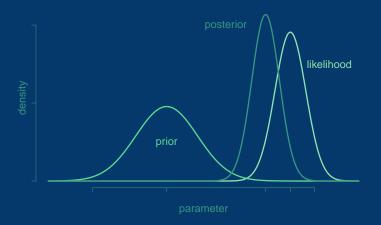
$$p(\theta|Y) \propto L(\theta|Y) p(\theta)$$

### Learning mechanism.

The prior information about the parameters is updated by the information contained in the data and represented by the likelihood function resulting in the posterior distribution.

### Likelihood principle.

All the information about the parameters of the model included in the data is captured by the likelihood function.



Joint distribution of data and parameters.

$$p(Y|\theta)p(\theta) = p(Y,\theta) = p(\theta|Y)p(Y)$$

The joint distribution of data and parameters is decomposed into:

**Inputs:** likelihood function  $p(Y|\theta)$  and prior distribution  $p(\theta)$ 

**Outputs:** posterior distribution  $p(\theta|Y)$  and marginal data density p(Y)

# **Useful** distributions

## Multivariate normal distribution

Let an  $N \times 1$  real-valued random vector X follow a multivariate normal distribution:

$$X \sim \mathcal{N}_N(\mu, \Sigma)$$

with the mean vector  $\mu$  and the covariance matrix  $\Sigma$ .

pdf.

$$\mathcal{N}_{N}(\mu, \Sigma) = (2\pi)^{-\frac{N}{2}} \det(\Sigma)^{-\frac{1}{2}} \exp\left\{-\frac{1}{2}(X - \mu)' \Sigma^{-1}(X - \mu)\right\}$$

Moments.

$$\mathbb{E}(X) = \mu$$
, and  $Var(X) = \Sigma$ 

# Inverse gamma 2 distribution

Let a positive real-valued scalar random variable x follow an inverse gamma 2 distribution:

$$x \sim \mathcal{IG}2(s, \nu)$$

with the shape parameter  $\nu > 0$  and the scale parameter s > 0.

pdf.

$$\mathcal{IG}2(s,\nu) = \Gamma\left(\frac{\nu}{2}\right)^{-1} \left(\frac{s}{2}\right)^{\frac{\nu}{2}} \times \frac{\nu+2}{2} \exp\left\{-\frac{1}{2}\frac{s}{x}\right\}$$

Moments.

$$\mathbb{E}(x) = \frac{s}{\nu - 2}, \text{ for } \nu > 2, \quad Var(x) = \frac{2}{\nu - 4} \left[ \mathbb{E}(x) \right]^2, \text{ for } \nu > 4$$

$$mode = \frac{s}{\nu + 2}$$

# Normal inverse gamma 2 distribution

$$p(X|\sigma^2) = \mathcal{N}_N(\mu, \sigma^2 \Sigma)$$
$$p(\sigma^2) = \mathcal{I}\mathcal{G}2(s, \nu)$$

Then,  $(X, \sigma^2)$  follow a normal inverse gamma 2 distribution:

$$p(X, \sigma^2) = p(X|\sigma^2)p(\sigma^2) = \mathcal{NIG2}_N(\mu, \Sigma, s, \nu)$$

pdf.

$$\mathcal{NIG2}(\mu, \Sigma, s, \nu) = c_{nig2}^{-1} \left(\sigma^{2}\right)^{-\frac{\nu+N+2}{2}} \exp\left\{-\frac{1}{2} \frac{1}{\sigma^{2}} \left[s + (X - \mu)' \Sigma^{-1} (X - \mu)\right]\right\}$$
$$c_{nig2} = \Gamma\left(\frac{\nu}{2}\right) \left(\frac{s}{2}\right)^{-\frac{\nu}{2}} (2\pi)^{\frac{N}{2}} \det(\Sigma)^{\frac{1}{2}}$$

#### Moments.

$$\mathbb{E}(X) = \mu, \text{ for } \nu > 1, \ Var(X) = \frac{s}{\nu - 2} \Sigma, \text{ for } \nu > 2$$

$$\mathbb{E}(\sigma^2) = \frac{s}{\nu - 2}, \text{ for } \nu > 2, \ Var(\sigma^2) = \frac{2}{\nu - 4} \left[\mathbb{E}(\sigma^2)\right]^2, \text{ for } \nu > 4$$

# Normal inverse gamma 2 distribution

Kernel of the  $\mathcal{NIG}2$  distribution.

$$\mathcal{NIG2}\left(\mu, \Sigma, s, \nu\right) \propto \left(\sigma^2\right)^{-\frac{\nu+N+2}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}(X-\mu)'\Sigma^{-1}(X-\mu)\right\} \exp\left\{-\frac{1}{2}\frac{s}{\sigma^2}\right\}$$

# Normal inverse gamma 2 distribution

## Generating random numbers from the $\mathcal{NIG}2$ distribution.

$$p(\beta, \sigma^{2}) = p(\beta|\sigma^{2})p(\sigma^{2})$$
$$p(\beta|\sigma^{2}) = \mathcal{N}(\mu, \sigma^{2}\Sigma)$$
$$p(\sigma^{2}) = \mathcal{IG}2(s, \nu)$$

## To draw S draws from the $\mathcal{NIG}2$ distribution...

- **Step 1:** Draw independently *S* draws from the  $\mathcal{IG}2(s, \nu)$ . Collect these draws in sequence  $\{\sigma^{2(s)}\}_{s=1}^{S}$
- **Step 2:** For each  $\sigma^{2(s)}$  sample a corresponding draw of  $\beta^{(s)}$  from  $\mathcal{N}\left(\mu,\sigma^{2(s)}\Sigma\right)$

**Return:**  $\left\{\beta^{(s)}, \sigma^{2(s)}\right\}_{s=1}^{S}$  as draws from the target distribution.

## Likelihood function

## A simple linear regression model.

$$Y = \beta X + E$$

$$E|X \sim \mathcal{N}\left(\mathbf{0}_{T}, \sigma^{2}I_{T}\right)$$

$$\downarrow$$

$$Y|X \sim \mathcal{N}\left(\beta X, \sigma^{2}I_{T}\right)$$

#### The likelihood function.

$$L(\theta|Y,X) = (2\pi)^{-\frac{T}{2}} \left(\sigma^2\right)^{-\frac{T}{2}} \exp\left\{-\frac{1}{2} \frac{1}{\sigma^2} (Y - \beta X)'(Y - \beta X)\right\}$$

## Likelihood function

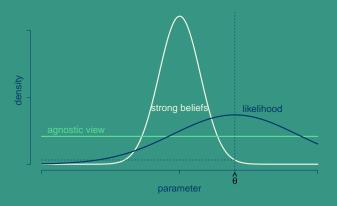
## The likelihood function as the $\mathcal{NIG}2$ distribution.

$$\begin{split} L\left(\theta|Y,X\right) &\propto \left(\sigma^2\right)^{-\frac{T}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}(Y-\beta X)'(Y-\beta X)\right\} \\ &= \left(\sigma^2\right)^{-\frac{T}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}(Y-\hat{\beta}X+\hat{\beta}X-\beta X)'(Y-\hat{\beta}X+\hat{\beta}X-\beta X)\right\} \\ &= \left(\sigma^2\right)^{-\frac{T}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}\left[(\beta-\hat{\beta})'X'X(\beta-\hat{\beta})+(Y-\hat{\beta}X)'(Y-\hat{\beta}X)\right]\right\} \\ &= \left(\sigma^2\right)^{-\frac{T-3+1+2}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}(\beta-\hat{\beta})'X'X(\beta-\hat{\beta})\right\} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^2}(Y-\hat{\beta}X)'(Y-\hat{\beta}X)\right\} \end{split}$$

### The result.

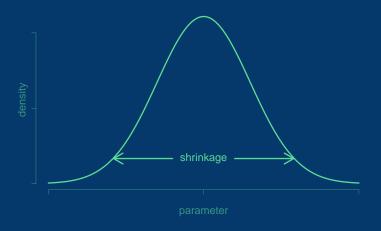
$$L(\theta|Y,X) = \mathcal{N}\mathcal{I}\mathcal{G}^{2}(\mu = \hat{\beta}, \Sigma = (X'X)^{-1}, s = (Y - \hat{\beta}X)'(Y - \hat{\beta}X), \nu = T - 3)$$
  
where  $N = 1$ .

## Prior distribution



A prior distribution formalizes researcher's beliefs regarding the parameters of the model before seeing the data.

## Prior distribution



# Natural-conjugate prior distribution

A natural-conjugate prior distribution is of the same form as the distribution of the parameters implied by the likelihood function.

$$p(\beta, \sigma^{2}) = p(\beta|\sigma^{2}) p(\sigma^{2})$$
$$p(\beta|\sigma^{2}) = \mathcal{N}(\underline{\beta}, \sigma^{2}\underline{\sigma}_{\beta}^{2})$$
$$p(\sigma^{2}) = \mathcal{I}\mathcal{G}2(\underline{s}, \underline{\nu})$$

Then,  $(\beta, \sigma^2)$  follow a priori a normal inverse gamma 2 distribution:

$$p\left(\beta,\sigma^{2}\right) = \mathcal{NIG2}_{N}\left(\underline{\beta},\underline{\sigma_{\beta}^{2}},\underline{s},\underline{\nu}\right)$$

pdf.

$$\mathcal{NIG2}_{(N=1)}\left(\underline{\beta},\underline{\sigma}_{\beta}^{2},\underline{s},\underline{\nu}\right) \propto \left(\sigma^{2}\right)^{-\frac{\nu+3}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^{2}}\frac{1}{\underline{\sigma}_{\beta}^{2}}(\beta-\underline{\beta})'(\beta-\underline{\beta})\right\} \exp\left\{-\frac{1}{2}\frac{\underline{s}}{\sigma^{2}}\right\}$$

$$\begin{split} p\left(\beta, \sigma^2 | Y, X\right) &\propto L\left(Y | X, \beta, \sigma^2\right) p\left(\beta, \sigma^2\right) \\ &= L\left(Y | X, \beta, \sigma^2\right) p\left(\beta | \sigma^2\right) p\left(\sigma^2\right) \end{split}$$

### Kernel of posterior distribution.

$$\begin{split} \rho\left(\beta,\sigma^{2}|Y,X\right) &\propto \left(\sigma^{2}\right)^{-\frac{T}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^{2}}(\beta-\hat{\beta})'X'X(\beta-\hat{\beta})\right\} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^{2}}(Y-\hat{\beta}X)'(Y-\hat{\beta}X)\right\} \\ &\times \left(\sigma^{2}\right)^{-\frac{\nu+3}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^{2}}\frac{1}{\sigma^{2}}(\beta-\underline{\beta})'(\beta-\underline{\beta})\right\} \exp\left\{-\frac{1}{2}\frac{\underline{s}}{\sigma^{2}}\right\} \end{split}$$

### Kernel of posterior distribution.

$$p(\beta, \sigma^{2}|Y, X) \propto (\sigma^{2})^{-\frac{\nu+T+3}{2}} \exp\left\{-\frac{1}{2} \frac{1}{\sigma^{2}} \left[\frac{1}{\underline{\sigma}_{\beta}^{2}} (\beta - \underline{\beta})'(\beta - \underline{\beta}) + (\beta - \hat{\beta})'X'X(\beta - \hat{\beta}) + \underline{s} + (Y - \hat{\beta}X)'(Y - \hat{\beta}X)\right]\right\}$$

After derivations, the expression in the square parentheses can be shown to have the following form:

$$\begin{split} \frac{1}{\underline{\sigma_{\beta}^{2}}}(\beta-\underline{\beta})'(\beta-\underline{\beta}) + (\beta-\hat{\beta})'X'X(\beta-\hat{\beta}) + \underline{s} + (Y-\hat{\beta}X)'(Y-\hat{\beta}X) \\ &= \overline{\sigma_{\beta}^{-2}}(\beta-\overline{\beta})'(\beta-\overline{\beta}) + \underline{s} + \underline{\beta}^{2}\underline{\sigma_{\beta}^{-2}} - \overline{\beta}^{2}\overline{\sigma_{\beta}^{-2}} + Y'Y \end{split}$$

where expressions for  $\overline{\beta}$  and  $\overline{\sigma}_{\beta}^2$  are given on the next slides.

After plugging in the expression, the kernel of the posterior distribution takes the form of:

$$p(\beta, \sigma^{2}|Y, X) \propto (\sigma^{2})^{-\frac{\overline{\nu}+3}{2}} \exp\left\{-\frac{1}{2}\frac{1}{\sigma^{2}}\frac{1}{\overline{\sigma}_{\beta}^{2}}(\beta - \overline{\beta})'(\beta - \overline{\beta})\right\} \exp\left\{-\frac{1}{2}\frac{\overline{s}}{\sigma^{2}}\right\}$$

in which we recognize the kernel of the normal inverse gamma 2 distribution.

$$\begin{split} \rho\left(\beta,\sigma^{2}|Y,X\right) &= \mathcal{N}\mathcal{I}\mathcal{G}2_{(N=1)}\left(\overline{\beta},\overline{\sigma}_{\beta}^{2},\overline{s},\overline{\nu}\right) \\ \overline{\sigma}_{\beta}^{2} &= \left(\underline{\sigma}_{\beta}^{-2} + X'X\right)^{-1} \\ \overline{\beta} &= \overline{\sigma}_{\beta}^{2}\left(\underline{\sigma}_{\beta}^{-2}\underline{\beta} + X'Y\right) \\ \overline{s} &= \underline{s} + \underline{\sigma}_{\beta}^{-2}\underline{\beta}^{2} - \overline{\sigma}_{\beta}^{-2}\overline{\beta}^{2} + Y'Y \\ \overline{\nu} &= \underline{\nu} + T \end{split}$$

## The posterior mean of $\beta$ .

$$\overline{\beta} = \overline{\sigma}_{\beta}^{2} \left( \underline{\sigma}_{\beta}^{-2} \underline{\beta} + X'Y \right)$$

$$= \overline{\sigma}_{\beta}^{2} \underline{\sigma}_{\beta}^{-2} \underline{\beta} + \overline{\sigma}_{\beta}^{2} X'X(X'X)^{-1} X'Y$$

$$= \frac{\underline{\sigma}_{\beta}^{-2}}{\underline{\sigma}_{\beta}^{-2} + X'X} \underline{\beta} + \frac{X'X}{\underline{\sigma}_{\beta}^{-2} + X'X} \hat{\beta}$$

$$= \underline{\omega}\underline{\beta} + (1 - \underline{\omega})\hat{\beta}$$

The posterior mean of  $\beta$  is the weighted average between the prior mean  $\beta$  and the MLE  $\hat{\beta}$ .

$$\overline{\beta} = \frac{\underline{\sigma}_{\beta}^{-2}}{\underline{\sigma}_{\beta}^{-2} + X'X} \underline{\beta} + \frac{X'X}{\underline{\sigma}_{\beta}^{-2} + X'X} \hat{\beta}$$

$$= \frac{\underline{\sigma}_{\beta}^{-2}}{\underline{\sigma}_{\beta}^{-2} + X'X} \underline{\beta} + \frac{\underline{X'X}}{\underline{\sigma}_{\beta}^{-2} + X'X} \hat{\beta}$$

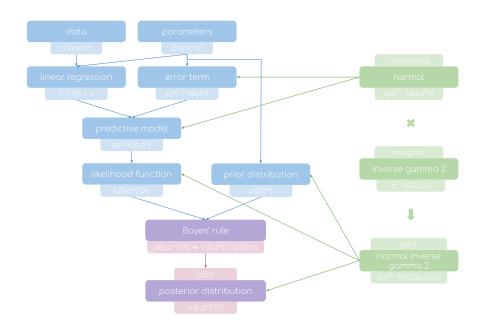
#### Limits.

 $\lim_{T\to\infty} \frac{\underline{\sigma}_{\beta}^{-2}}{T} = 0$  — as  $\underline{\sigma}_{\beta}^{-2}$  is a constant

 $\lim_{T \to \infty} \frac{X'X}{T} = \sigma_X^2 - \sigma_X^2$  is the second non-central moment of X

The posterior mean of  $\beta$  when  $T \to \infty$ .

$$\lim_{T \to \infty} \overline{\beta} = \hat{\beta}$$



## Bayesian estimation

### For a linear Gaussian regression:

**Likelihood function** has a form of a Normal inverse gamma 2 distribution for the parameters of the model

**Normal inverse gamma 2 distribution** for the parameters is the naturally-conjugate prior distribution leading to...

Normal inverse gamma 2 posterior distribution

**Asymptotically** Bayesian estimation converges to the MLE