

$$y_i = \beta_1 + \beta_2 X_{2i} + \beta_3 X_{3i} + \varepsilon_i$$

$$(a) \quad H_0: \beta_2 = 0$$

$$H_a: \beta_2 \neq 0$$

$$t = \frac{\hat{\beta}_2 - 0}{\text{se}(\hat{\beta}_2)} \sim t_{n-k}$$

\nearrow $\# \text{ obs}$ \nearrow $\# \text{ est. coef.}$

$$(b) \quad H_0: \beta_2 = 1$$

$$t = \frac{\hat{\beta}_2 - 1}{\text{se}(\hat{\beta}_2)} \sim t_{n-k}$$

$$(c) \quad H_0: \beta_2 = \beta_3 = 0$$

$$H_a: \beta_2 \neq 0 \vee \beta_3 \neq 0$$

$$F = \frac{R^2 / 2}{1 - R^2 / (n - 3)} \sim F(2, n - 3)$$

F-test for linear restrictions

$$H_0: \beta_4 = \beta_7 = \beta_8 = 0 \Rightarrow RSS_R$$

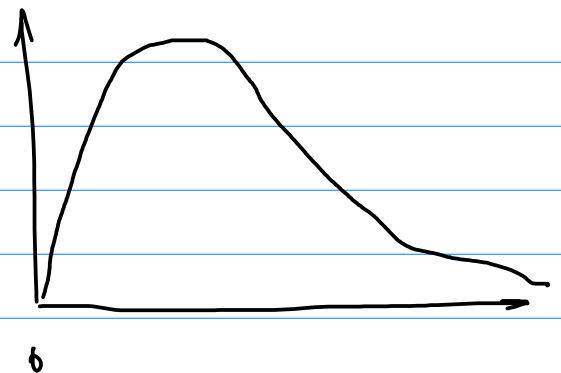
$$H_a: \text{one of restrictions} \Rightarrow RSS_{UR}$$

isn't satisfied, # of linear restrictions

$$F = \frac{(RSS_R - RSS_{UR}) / q}{RSS_{UR} / (n - k)} \sim F(q, n - k)$$

est. coef. in UR model

$$RSS_R > RSS_{UR}$$



F-test for goodness of fit

$$H_0: \beta_2 = \dots = \beta_k = 0 \quad y_i^R = \beta_1 + \epsilon_i$$

$$H_a: \exists i \beta_i \neq 0$$

$$F = \frac{(TSS - RSS) / (k - 1)}{RSS / (n - k)} = \frac{ESS / (k - 1)}{RSS / (n - k)} = \frac{R^2 / (k - 1)}{(1 - R^2) / (n - k)} \sim F(k - 1, n - k)$$

$$d) \quad H_0: \beta_2 = \beta_3 \quad \theta = \beta_2 - \beta_3 = 0$$

$$\beta_3 = \beta_2 - \theta$$

$$UR: y_i = \beta_1 + \beta_2 X_{2i} + \beta_3 X_{3i} + \varepsilon_i$$

$$R: y_i = \beta_1 + \beta_2 (X_{2i} + X_{3i}) + \varepsilon_i$$

$$t: y_i = \beta_1 + \beta_2 (X_{2i} + X_{3i}) - \theta X_{3i} + \varepsilon_i$$

$$F = \frac{(RSS_R - RSS_{UR}) / q}{RSS_{UR} / (n - k)} \sim F(q, n - k)$$

$$= \frac{(RSS_R / TSS - RSS_{UR} / TSS) / q}{RSS_{UR} / TSS / (n - k)}$$

$$= \frac{(1 - R^2_R - 1 + R^2_{UR}) / q}{(1 - R^2_{UR}) / (n - k)} = \frac{R^2_{UR} - R^2_R / q}{1 - R^2_{UR} / (n - k)}$$

(e) $H_0: \beta_2 + \beta_3 = 1$

UR: $y_i = \beta_1 + \beta_2 x_{2i} + \beta_3 x_{3i} + \varepsilon_i$

R: $y_i = \beta_1 + \beta_2 x_{2i} + (1 - \beta_2) x_{3i} + \varepsilon_i$

$y_i - x_{3i} = \beta_1 + \beta_2 (x_{2i} - x_{3i}) + \varepsilon_i$

\Rightarrow F-test using RSS

(not R^2 !)

$$TSS_R \neq TSS_{UR}$$

PE?

$$y_i = \beta_1 + \beta_2 x_i + \varepsilon_i \quad i = \overline{1, n}$$

$$\hat{y}_{n+1} = \hat{\beta}_1 + \hat{\beta}_2 x_{n+1}$$

$$E(\hat{y}_{n+1}) = E(\hat{\beta}_1 + \hat{\beta}_2 x_{n+1}) = \beta_1 + \beta_2 x_{n+1}$$

$$E(y_{n+1}) = E(\beta_1 + \beta_2 x_{n+1} + \varepsilon_i) = \beta_1 + \beta_2 x_{n+1} + 0$$

$$\Rightarrow E(y_{n+1}) = E(\hat{y}_{n+1})$$

$$\text{Var}(\hat{y}_{n+1} - y_{n+1}) = E(\hat{y}_{n+1} - y_{n+1})^2 =$$

$$= E(\hat{\beta}_1 + \hat{\beta}_2 \cdot x_{n+1} - \beta_1 - \beta_2 x_{n+1} - \varepsilon_{n+1})^2 =$$

$$= E((\hat{\beta}_1 - \beta_1) + (\hat{\beta}_2 - \beta_2) x_{n+1} - \varepsilon_{n+1})^2 =$$

$$= E(\hat{\beta}_1 - \beta_1)^2 + x_{n+1} E(\hat{\beta}_2 - \beta_2)^2 + E(\varepsilon_{n+1})^2 +$$

$$2 x_{n+1} E((\hat{\beta}_1 - \beta_1)(\hat{\beta}_2 - \beta_2)) - 2 E((\hat{\beta}_1 - \beta_1) \cdot \varepsilon_{n+1}) -$$

$$- 2 x_{n+1} E((\hat{\beta}_2 - \beta_2) \varepsilon_{n+1}) =$$

$$= \text{var}(\hat{\beta}_1) + x_{n+1} \cdot \text{var}(\hat{\beta}_2) + \sigma^2 +$$

$$2 x_{n+1} \text{cov}(\hat{\beta}_1, \hat{\beta}_2) - 0 - 0 =$$

$$= \frac{\sigma^2 \sum x_i^2}{n \sum (x_i - \bar{x})^2} + x_{n+1} \frac{\sigma^2}{\sum (x_i - \bar{x})^2} + \sigma^2 + 2 x_{n+1} \frac{-\bar{x} \sigma^2}{\sum (x_i - \bar{x})^2} =$$

$$= \underbrace{\sigma^2}_{S^2} \left(1 + \frac{1}{n} + \frac{(X_{n+1} - \bar{X})^2}{\sum (X_i - \bar{X})^2} \right)$$

$$\sigma^2(\beta) = \sigma^2 (X'X)^{-1} = \begin{bmatrix} \frac{\sigma^2 \sum X_i^2}{n \sum (X_i - \bar{X})^2} & \frac{-\bar{X} \sigma^2}{\sum (X_i - \bar{X})^2} \\ \frac{-\bar{X} \sigma^2}{\sum (X_i - \bar{X})^2} & \frac{\sigma^2}{\sum (X_i - \bar{X})^2} \end{bmatrix}$$

$$X'X = \begin{bmatrix} n & \sum X_i \\ \sum X_i & \sum X_i^2 \end{bmatrix}$$

Multicollinearity

- perfect multicollinearity

↳ one regressor is a lin. com.

of others

$$\hat{\beta} = (X'X)^{-1}X'y$$

X - not full rank

$X'X$ - cannot be inverted

$$\begin{pmatrix} 1 & 1 & 0 & X_1 \\ \vdots & \vdots & \vdots & \vdots \\ \vdots & 1 & 0 & \vdots \\ \vdots & 0 & 1 & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 1 & X_n \end{pmatrix}$$

$$I = n + 1$$

dummy variable trap (example of perfect MC)

multicollinearity:

Consequences:

- $SE(\hat{\beta}_i)$ are inflated
- t -stat. are decreased
- instability of estimates

(change in sign)

