Imbalanced Classification. Multiclass Classification

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Outline

- Imbalanced Classification
- Multi-Class Classification: General Strategies and Particular Models
- Multi-Class Classification with Naive Bayes Classifier
- 4 Appendix. Materials for Additional Study: Nonparametric Estimation

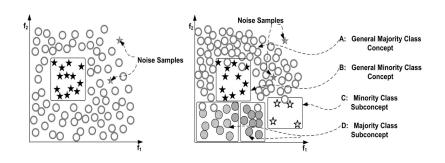
Imbalanced Classification

2 Multi-Class Classification: General Strategies and Particular Models

Multi-Class Classification with Naive Bayes Classifier

Appendix. Materials for Additional Study: Nonparametric Estimation

- Between-class imbalance (relative imbalance)
- Relative imbalance vs. imbalance due to rare instances or "absolute rarity"
- Data complexity vs. imbalanced data vs. small sample size



- Binary classification: often dataset has "natural" imbalance
- Minor class (of prime interest) vs. major class: e.g. classification of "cancerous" vs. "healthy" mammography image
- Standard classifiers (SVM, kNN, log. reg., etc.): classes are equally important ⇒ results are biased towards the major class
- Poor prediction of minor class while the average quality can be good:
 - target events occurs in 1% of all cases
 - classifier always gives a 'no-event' answer
 - it is wrong just 1% of all cases

- Approaches to increase importance of the minor class:
 - Adapt a probability threshold for classifiers,
 - Modify a loss function, e.g., by assigning more weight to the minor class error.
 - Resample a dataset in order to soften or remove class imbalance
- We focus on resampling: convenient, allows to use standard classifiers
- The main aim:
 - review and compare main resampling methods,
 - compare strategies of resampling amount (i.e., how many observations to add or drop) selection,
 - explore their influence on quality of classification

- Dataset $S = \{(\mathbf{x}_i, y_i)\}_{i=1}^n$, where $\mathbf{x}_i \in \mathbb{R}^d$, $y_i \in \{-1, +1\}$
- $C_{+1}(S) = \{(\mathbf{x}_i, y_i) \in S \mid y_i = +1\}$ is a major class,
- $C_{-1}(S)=\{(\mathbf{x}_i,y_i)\in S\mid y_i=-1\}$ is a minor class, i.e. $|C_{+1}(S)|>|C_{-1}(S)|$
- Imbalance ratio $IR(S) = \frac{|C_{-1}(S)|}{|C_{+1}(S)|}$, $IR(S) \leq 1$

Learning a classifier

- ullet Learn a classifier using imbalanced training sample S
- ullet The dataset S is resampled using a method r:
 - some observations in S are dropped, or
 - some new synthetic observations are added to S
- The result of resampling is a dataset r(S) with IR(r(S)) > IR(S)
- Standard classification model f is learned on r(S) to construct a classifier $f_{r(S)}: \mathbb{R}^d \to \{-1, +1\}$

- Performance is determined by a predefined classifier quality metrics $Q(f_{S_{train}}, S_{test})$ (e.g. AUC under Precision-Recall curve):
 - input classifier $f_{S_{train}}$,
 - testing dataset S_{test} ,
 - the higher value is the better
- ullet M-fold cross-validation is used to estimate Q^{CV} on S

Resampling method r:

- Takes input:
 - dataset S;
 - resampling multiplier m > 1 for resulting imbalance ratio $\overline{IR(r(S))} = m \cdot \overline{IR(S)}$;
 - · additional parameters, specific for the method
- Add synthesized objects to the minor class (<u>oversampling</u>), or drop objects from the major class (<u>undersampling</u>), or both
- Outputs resampled dataset r(S) with imbalance ratio $IR(r(S)) = m \cdot IR(S)$

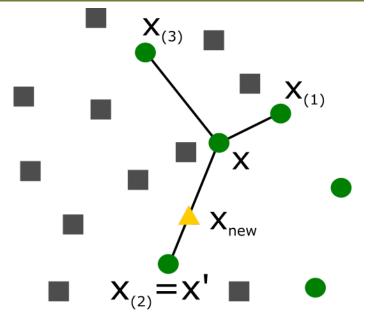
Random Oversampling (ROS)

- ROS, also known as bootstrap oversampling
- No additional input parameters
- ullet It adds to the minor class new $(m-1)|C_{-1}(S)|$ objects
- \bullet Each of objects is drawn from uniform distribution on ${\cal C}_{-1}(S)$

Random Undersampling (RUS)

- No additional input parameters
- It chooses random subset of $C_{+1}(S)$ with $|C_{+1}(S)| \frac{m-1}{m}$ elements and drops it from the dataset
- All subsets of $C_{+1}(S)$ have equal probabilities to be chosen

Synthetic Minority Oversampling Technique (SMOTE)



Synthetic Minority Oversampling Technique (SMOTE)

- Input parameter: k (number of neighbors)
- Oversampling, it adds to the minor class new synthesized objects
- Initialize: $S_{new} := \emptyset$. Repeat $(m-1)|C_{-1}(S)|$ times:
 - Select randomly $\mathbf{x}_i \in C_{-1}(S)$
 - ② Find k minor class NN of \mathbf{x}_i , randomly select \mathbf{x}'_i from them
 - lacksquare Select randomly \mathbf{x}_{new} on the segment connecting \mathbf{x}_i and \mathbf{x}_i'
- \bullet Add objects to the dataset: $r(S) = S \cup S_{new}$

- Artificial pool of data with ~ 1000 datasets
- Artificial datasets were drawn from a Gaussian mixture distribution
- Each of two classes is represented as a Gaussian mixture with not more than 3 components
- Number of features varies from 6 to 40, size of dataset from 200 to 1000, IR from 0.05 to 0.35

- Real pool of data with ~ 100 datasets
- Different areas: biology, medicine, engineering, sociology
- \bullet All features are numeric or binary, their number varies from 3 to 1000
- \bullet Size of dataset varies from 200 to 1000, IR from 0.02 to 0.75

Setup of Experiments

- For each dataset we varied classifier model, resampling method and resampling multiplier
- ullet We used Bootstrap, RUS and SMOTE with k=5
- ullet We varied resampling multiplier m from 1.25 to 10.0
- ullet We used Decision Trees, k-Nearest Neighbors, and Logistic Regression with ℓ_1 regularization
- Optimal parameters of a classifier were selected by cross-validation

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Setup of Experiments

- ullet Accuracy measure = Area under precision-recall curve Q_{PRC}
- \bullet We performed 10-fold cross-validation and calculated Q_{PRC}^{CV} average of Q_{PRC}

- Two strategies of resampling multiplier selection:
 - equalizing strategy, $\underline{\mathsf{EqS}}$: select multiplier providing balanced classes ($\mathit{IR}=1$) in resulting dataset
 - CV-search, CVS: select optimal multiplier (i.e., providing maximum of $\overline{Q^{CV}}$) by cross-validation
- The equalizing strategy seems to be reasonable as it removes class imbalance which we initially tried to tackle. It is quick and widely used
- CV-search may provide better quality but it is more time-consuming

- $\{r_1,\ldots,r_n\}$ the set of considered methods (e.g. resampling methods)
- $\{S_1, \dots S_T\}$ the set of tasks (datasets),
- q_{ti} the quality of the method i on the dataset t,
- $p_i(\beta)$ is a fraction of tasks, on which the method i is worse than the best one not more than β times:

$$p_i(\beta) = \frac{1}{T} \left| \left\{ t : q_{ti} \ge \frac{1}{\beta} \max_i q_{ti} \right\} \right|, \ \beta \ge 1$$

Dolan-More Curves

- $p_i(1)$ is a fraction of datasets where the method i is the best
- A graph of $p_i(\beta)$ on β is called Dolan-More curve for the method i
- The higher the curve, the better the method!

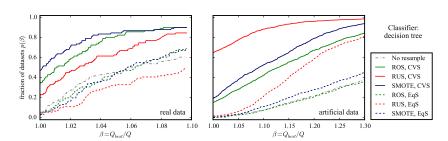


Figure – Dolan-More curves for metric Q_{PRC}^{CV}

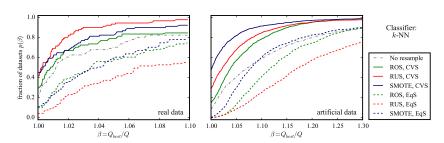


Figure – Dolan-More curves for metric Q_{PRC}^{CV}

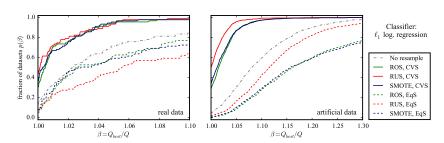


Figure – Dolan-More curves for metric Q_{PRC}^{CV}

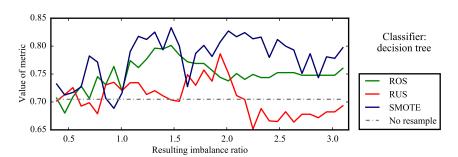


Figure – Value of Q_{PRC}^{CV} vs. resulting value of $I\!R$ for dataset "Delft pump 1x3"

- Resampling with CV-search of multiplier provides better results, especially for Decision trees and Logistic regression
- The equalizing strategy (EqS) shows much lower quality, especially in case of k-Nearest neighbors and Logistic regression
- There is no method that would always outperform the others
- Classification without resampling is the best choice in some cases. E.g., for Logistic regression it is about 15% of real datasets and 5% of artificial
- Resampling improves classification of imbalanced datasets in most cases if a method and a multiplier are selected properly

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Imbalanced Classification

2 Multi-Class Classification: General Strategies and Particular Models

Multi-Class Classification with Naive Bayes Classifier

4 Appendix. Materials for Additional Study: Nonparametric Estimation

• **Training sample**: i.i.d. generated by D

$$S = \{(\mathbf{x}_i, y_i)\}_{i=1}^m \in X^m \times Y^m, \ y = f(\mathbf{x})$$

- mono-label case: Card(Y) = K
- multi-label case: $Y = \{-1, +1\}^K$
- **Problem**: find classifier $h: X \to Y$ swith small generalization error

 - $\begin{array}{l} \text{ mono-label case: } R_D(h) = \mathbb{E}_{\mathbf{x} \sim D} \left[\mathbf{1}_{h(\mathbf{x}) \neq f(\mathbf{x})} \right] \\ \text{ multi-label case: } R_D(h) = \mathbb{E}_{\mathbf{x} \sim D} \left[\frac{1}{K} \sum_{j=1}^K \mathbf{1}_{[h(\mathbf{x})]_j \neq [f(\mathbf{x})]_j} \right] \end{array}$
 - Remark: Precision/Recall with Micro/Macro averaging!

Comments

- Usually $K \leq 100$
- ullet If $K\gg 1$ then some other methods are used, e.g. ranking
- ullet Big values of K increases computational burden
- In general, classes are not balanced

Technique

— for each class $k \in Y$ learn a binary classifier

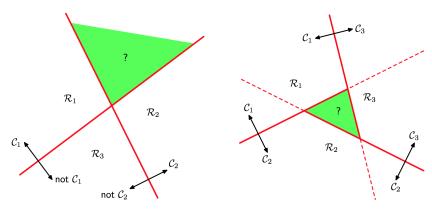
$$h_k(\mathbf{x}) = \operatorname{sign}(f_k(\mathbf{x}))$$

- combine binary classifiers via voting, e.g. majority voting

$$h: \mathbf{x} \to \arg\max_{k \in Y} f_k(\mathbf{x})$$

Comments

- calibration: classifiers scores are not comparable
- simple and frequently used in practice, computational advantages in some cases



Consider the use of K-1 classifiers each of which solves a two-class problem of separating points in a particular class from points not in that class. This approach leads to regions of input space that are ambiguously classified

Technique

- for each pair $(k, k') \in Y$, $k \neq k'$ learn a binary classifier $h_{k,k'} : X \to \{0,1\}$
- combine binary classifiers via majority vote

$$h(\mathbf{x}) = \arg \max_{k' \in Y} |\{k : h_{k,k'}(\mathbf{x}) = 1\}|$$

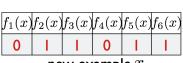
Comments

- computational complexity: train K(K-1)/2 binary classifiers
- overfitting: size of a training sample can be small for a given pair of classifiers

ullet 8 classes, codes of length 6

codes

	I	2	3	4	5	6
	0	0	0	ı	0	0
2	ı	0	0	0	0	0
3	0	I	I	0	-	0
4	I	I	0	0	0	0
5	I	I	0	0	Ι	0
6	0	0	ı	ı	0	ı
7	0	0	I	0	0	0
8	0	ı	0	1	0	0



 $\mathsf{new}\,\,\mathsf{example}\,x$

 Assign L-long binary code word to each class, i.e. represent each class as

$$\mathbb{C} = [\mathbb{C}_{k,j}] \in \{0,1\}^{[1,K] \times [1,L]}$$

- Learn a binary classifier $f_j: X \to \{0,1\}$ for each column. Example $\mathbf x$ in class k is labeled with $\mathbb C_{k,j}$
- Classifier output:

$$\mathbf{f}(\mathbf{x}) = (f_1(\mathbf{x}), \dots, f_L(\mathbf{x})),$$

Final classifier

$$h: \mathbf{x} \to \arg\min_{k \in Y} d_{\text{Hamming}} \left(\mathbb{C}_{k,\cdot,\cdot}, \mathbf{f}(\mathbf{x}) \right)$$

- One-vs-all approach is the most widely used
- No clear empirical evidence of the superiority of other approaches
- Large structured multi-class problems are often treated as ranking problems
- Above we considered how to reduce multi-class classification to the binary case.
- Also we can incorporate a multi-class structure explicitly into a classification algorithm, see e.g. multi-class logistic regression or multi-class SVM (below)

Particular Model: Multi-class Logistic Regression

- ullet Linear Classifier in a multi-class case, i.e. |Y|>1
- ullet Probability of an object to belong to some class y is equal to

$$p(y|\mathbf{x}; \mathbf{w}) = \frac{\exp(\mathbf{w}_y^{\top} \mathbf{x})}{\sum_{z \in Y} \exp(\mathbf{w}_z^{\top} \mathbf{x})} = \text{SoftMax}_{y \in Y}(\mathbf{w}_y^{\top} \mathbf{x})$$

Regularized logistic regression

$$\overline{L}(\mathbf{w}) = \frac{1}{m} \sum_{i=1}^{m} \log p(y_i | \mathbf{x}_i, \mathbf{w}) - \frac{\lambda}{2} \sum_{y \in Y} \|\mathbf{w}_y\|^2 \to \max_{\mathbf{w}}$$

Optimization problem

$$\min_{\mathbf{w}, \xi} \frac{1}{2} \sum_{k=1}^{K} \|\mathbf{w}_k\|^2 + C \sum_{i=1}^{m} \xi_i$$

$$s.t. \ \mathbf{w}_{y_i} \mathbf{x}_i^{\top} - \mathbf{w}_k \mathbf{x}_i^{\top} + \delta_{y_i, k} \ge 1 - \xi_i, \ \xi_i \ge 0$$

$$(i, k) \in [1, m] \times Y$$

Decision function:

$$h: \mathbf{x} \to \arg\max_{k \in Y} (\mathbf{w}_k \cdot \mathbf{x}^\top) = \arg\max_{k \in Y} \left(\sum_{i=1}^m \alpha_{i,k} (\mathbf{x}_i \cdot \mathbf{x}^\top) \right),$$

where $\{\alpha_{i,k}\}_{i=1}^m$, $k \in Y$ are dual variables

ullet Complex constraints, $m \cdot K$ size

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- Naive Bayes (NB) is a conditional probability model
- Given an object to be classified, represented by features $\mathbf{x} = (x_1, \dots, x_d)$, Bayes classifier assigns probabilities

$$p(y|x_1,\ldots,x_d)$$

for each of K possible classes $y \in Y = \{1, \dots, K\}$

The conditional probability can be decomposed as

$$p(y|\mathbf{x}) = \frac{p(y)p(\mathbf{x}|y)}{p(\mathbf{x})} \sim p(y)p(\mathbf{x}|y), \ y \in Y$$

The maximum a posteriori (MAP) decision rule

$$\widehat{y} = \arg\max_{y \in \{1, \dots, K\}} p(y) p(\mathbf{x}|y)$$

• How to model $p(\mathbf{x}|y)$?

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• Let us decompose $p(y, \mathbf{x})$. Using the chain rule we get that

$$p(y, x_1, \dots, x_d) = p(x_1, \dots, x_d, y)$$

$$= p(x_1 | x_2, \dots, x_d, y) \cdot (x_2, \dots, x_d, y)$$

$$= p(x_1 | x_2, \dots, x_d, y) \cdot p(x_2 | x_3, \dots, x_d, y) \cdot p(x_3, \dots, x_d, y) = = \dots =$$

$$= p(x_1 | x_2, \dots, x_d, y) \cdot \dots \cdot p(x_{d-1} | x_d, y) \cdot p(x_d | y) \cdot p(y)$$

- \bullet Usually $d\gg 1$ or a feature can take on a large number of values \Leftrightarrow we assume independence of features
- The "naive" conditional independence: assume that each feature x_j is conditionally independent of every other feature x_s given the category y, i.e.

$$p(x_j|x_{j+1},\ldots,x_d,y)=p(x_j|y)$$

The joint model

$$p(y|\mathbf{x}) \sim p(y) \prod_{j=1}^{d} p(x_j|y), y \in Y$$

Normalized model

$$p(y|\mathbf{x}) = \frac{1}{Z}p(y)\prod_{j=1}^{d}p(x_{j}|y), y \in Y$$

where
$$Z = p(\mathbf{x}) = \sum_{y \in Y} p(y) p(\mathbf{x}|y)$$

• The maximum a posteriori (MAP) decision rule

$$\widehat{y} = \arg \max_{y \in \{1, \dots, K\}} p(y) \prod_{j=1}^{d} p(x_j|y)$$

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- How to define individual models for $p(x_j|y)$?
- We can use any parametric distribution model for $p(x_i|y)!$
- ullet E.g. if x_j is continuous, we can use a Gaussian distribution, i.e.

$$p(x_j = v|y = k) = (2\pi\sigma_{jk}^2)^{-1/2}e^{-(v-\mu_{jk})^2/2\sigma_{jk}^2}$$

• Parameters $(\mu_{jk}, \sigma^2_{jk})$ are estimated using a subsample S_{jk} of the initial sample $S = \{(\mathbf{x}_i, y_i)\}_{i=1}^m$

$$S_{jk} = \{x_{ij} : (\mathbf{x}_i, y_i) \in S, y_i = k, i = 1, \dots, m\}$$

- We can use various discrete distributions in case some features are categorical
- Multinomial NB: feature x_j counts the number of times event j was observed in a particular instance
- If J is a set of indexes of categorical features, then we can model their joint multimomial distribution as

$$p(x_j, j \in J | y = k) = \frac{(\sum_{j \in J} x_j)!}{\prod_{j \in J} x_j!} \prod_{j \in J} p_{kj}^{x_j}$$

In log-scale we get a linear classifier

$$\log p(y = k | x_j, j \in J) \sim \log p(y = k) + \sum_{j \in J} x_j \cdot \log p_{kj} = b + \mathbf{w}_k \mathbf{x}_J^\top,$$

with
$$\mathbf{w}_k = (\log p_{kj}, j \in J)$$
 and $\mathbf{x}_J = (x_j, j \in J)$

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• Analogously we can define Bernoulli NB: x_j is a boolean expressing the occurrence or absence of the jth term from the vocabulary

$$p(x_j, j \in J | y = k) = \prod_{j \in J} p_{kj}^{x_j} (1 - p_{kj})^{(1 - x_j)}$$

- In order to learn NB:
 - Define models for $p(\mathbf{x}|y)$
 - Estimate parameters of these models using corresponding subsampls, extracted w.r.t. values of \boldsymbol{y}
 - Use MAP to predict \widehat{y}

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Nonparametric Estimation: Problem Statement

- $S = {\mathbf{x}_1, \dots, \mathbf{x}_m} \sim F$ is a given sample, $\mathbf{x} \in \mathbb{R}^1$
- \bullet $F(\mathbf{x})$ is an absolutely continuous CDF with an unknown density $p(\mathbf{x})$
- We estimate $p(\mathbf{x})$ in the point \mathbf{x} , i.e. construct $\widehat{p}_m(\mathbf{x}) = \widehat{p}_m(\mathbf{x}|S)$
- Typical parametric assumption

$$p \in \{p(\mathbf{x}; \boldsymbol{\theta}), \boldsymbol{\theta} \in \Theta\}, \Theta \subset \mathbb{R}^p$$

• Now we do not use such assumption \Rightarrow Nonparametric Estimation

ullet We estimate $\widehat{p}_m(\mathbf{x}) \, \forall \mathbf{x} \in \mathbb{R}^1$

Definition

Mean Integrated Squared Error:

$$MISE(\widehat{p}_m, p) = \mathbb{E}_p \left[\int_{\mathbb{R}} (\widehat{p}_m(\mathbf{x}) - p(\mathbf{x}))^2 d\mathbf{x} \right]$$

Histogram

• We consider interval $\{\mathbf{x}_1, \dots \mathbf{x}_m\} \in [a, b)$ and divide it into N equal bins Δ_i of size $h = \frac{b-a}{N}$:

$$\Delta_i = [a+ih, a+(i+1)h), i = 0, 1, \dots, N-1$$

ullet Let u_i be a number of data points, belonging to Δ_i

Definition

$$\widehat{p}_m(\mathbf{x}) = \begin{cases} \frac{\nu_0}{mh}, & \mathbf{x} \in \Delta_0, \\ \dots & \\ \frac{\nu_{N-1}}{mh}, & \mathbf{x} \in \Delta_{N-1}; \end{cases} = \frac{1}{mh} \sum_{i=0}^{N-1} \nu_i 1\{\mathbf{x} \in \Delta_i\}$$

• For $\mathbf{x} \in \Delta_i$ and small h:

$$\mathbb{E}_{p}\widehat{p}_{m}(\mathbf{x}) = \frac{\mathbb{E}_{p}\nu_{j}}{mh} = \frac{\int_{\Delta_{j}} p(\mathbf{z})d\mathbf{z}}{h} \approx \frac{p(\mathbf{x})h}{h} = p(\mathbf{x})$$

At the same time

$$\operatorname{Var}(\widehat{p}_m(\mathbf{x})) = \frac{\int_{\Delta_j} p(\mathbf{z}) d\mathbf{z} \left(1 - \int_{\Delta_j} p(\mathbf{z}) d\mathbf{z}\right)}{mh^2}$$

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Smoothing Selection: MISE

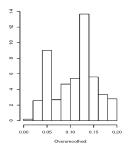
We can prove that mean integrated squared error

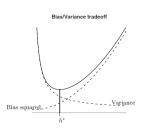
$$MISE(\widehat{p}_m, p) \approx \left(\int_{\mathbb{R}} [p'(\mathbf{x})]^2 d\mathbf{x} \right) \frac{h^2}{12} + \frac{1}{mh}$$

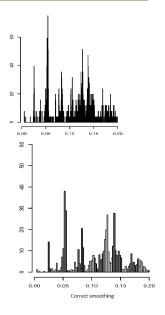
- ullet The bigger h we use the bigger bias and smaller variance we get, and vice versa: Bias-Variance Tradeoff
- ullet Too big h= oversmoothing, too small h= undersmoothing
- ullet Thus for a histogram with optimal h, we get that

$$MISE = O(m^{-\frac{2}{3}})$$

Smoothing Selection: Example







KDE allows to get smoother estimate (compared to histogram based ones) with faster convergence rates

Definition

Kernel is a function K, such that

$$K(\mathbf{x}) \ge 0, \int_{\mathbb{R}} K(\mathbf{x}) d\mathbf{x} = 1, \int_{\mathbb{R}} \mathbf{x} K(\mathbf{x}) d\mathbf{x} = 0, \sigma_K^2 \equiv \int_{\mathbb{R}} \mathbf{x}^2 K(\mathbf{x}) d\mathbf{x} < \infty$$

Examples

- $\blacktriangleleft K(x) = \frac{1}{2}\mathbb{I}\{|x| < 1\}$ rectangular kernel
- $\blacktriangleleft K(x) = (1-|x|)\mathbb{I}\{|x|<1\} \text{ -- triangle kernel }$
- $\blacktriangleleft K(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$ Gaussian kernel
- $K(x) = \frac{3}{4}(1-x^2)$ $\mathbb{I}\{|x|<1\}$ Epanechnikov kernel

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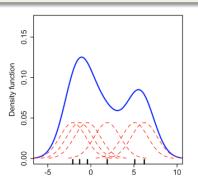
Kernel Density Estimation: Definition

Definition

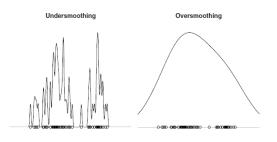
KDE has the form

$$\widehat{p}_m(\mathbf{x}) = \frac{1}{mh} \sum_{i=1}^m K\left(\frac{\mathbf{x} - \mathbf{x}_i}{h}\right),$$

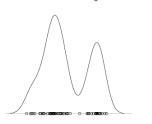
h is a kernel width



Kernel width selection: Example







- Multidimensional case, i.e. $\mathbf{x} = (x_1, \dots, x_d) \in \mathbb{R}^d$ is a point in \mathbb{R}^d .
- Thus i—th observation is an d-dimensional vector:

$$\mathbf{x}_i = (x_{i1}, \dots x_{id})$$

- Let $h=(h_1,\ldots,h_d)$ be a vector of kernel widths
- Then

$$\widehat{p}_m(\mathbf{x}) = \frac{1}{mh_1 \cdot \ldots \cdot h_d} \sum_{i=1}^m \left[\prod_{j=1}^d K\left(\frac{x_j - x_{ij}}{h_j}\right) \right]$$

Curse of Dimensionality

- Optimal rate of convergence is $O(m^{-\frac{4}{4+d}})$: if d increases convergence rate decreases
- Let us consider a table with values of m necessary to get the mean squared estimation error in $\mathbf{x}_0 = 0$ less than 0.1 depending on d for a multidimensional normal density and optimal kernel width:

d	1	2	3	4	5	6	7	8	9
m	4	19	67	223	768	2790	10700	43700	187000

ullet Here d is a dimension, m is a sample size

- Let us consider m observations: $S = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_m, y_m)\},$ generated from a joint density $p(\mathbf{x}, y)$
- These observations are generated by the model

$$y_i = f(\mathbf{x}_i) + \varepsilon_i,$$

where ε_i is an i.i.d white noise, $\mathbb{E}\varepsilon_i = 0$, $\mathbb{V}(\varepsilon_i) = \sigma^2$

We should estimate a regression function

$$f(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = \int_{\mathbb{R}} y p(y|\mathbf{x}) dy = \frac{\int_{\mathbb{R}} y p(\mathbf{x}, y) dy}{\int_{\mathbb{R}} p(\mathbf{x}, y) dy} = \frac{\int_{\mathbb{R}} y p(\mathbf{x}, y) dy}{p(\mathbf{x})}$$

Definition

Let K denote a kernel, and

- ullet $\widehat{p}_m(\mathbf{x})$ be a kernel density estimate, obtained using $\{\mathbf{x}_1,\ldots,\mathbf{x}_m\}$,
- $\widehat{p}_m(\mathbf{x}, y)$ be a kernel density estimate, obtained using $\{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_m, y_m)\}.$

If $\widehat{p}_m(\mathbf{x}) \neq 0$, then

$$\widehat{f}_m^{NW}(\mathbf{x}) = \frac{\int_{\mathbb{R}} y \widehat{p}_m(\mathbf{x}, y) dy}{\widehat{p}_m(\mathbf{x})}$$

We can notice that Nadaraya-Watson estimate can be used also in case when $\{\mathbf{x}_1,\dots,\mathbf{x}_m\}$ are some fixed and deterministic values, e.g. $\mathbf{x}_i=\frac{i}{m}$

We use Nadaraya-Watson estimate for $f(\mathbf{x})$:

Definition

Nadaraya-Watson estimate has the form

$$\widehat{f}_m^{NW}(\mathbf{x}) = \sum_{i=1}^m w_i(\mathbf{x}) y_i,$$

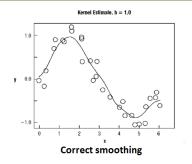
where

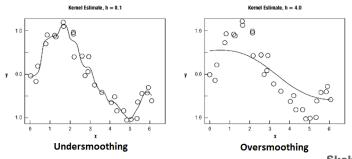
$$w_i = \frac{K(\frac{\mathbf{x} - \mathbf{x}_i}{h})}{\sum_{j=1}^m K(\frac{\mathbf{x} - \mathbf{x}_j}{h})},$$

and K is a some kernel function

Thus, the estimate is a weighted sum of y_i , where any point, close to x, has a big weight

Nonparametric regression: Example





Definition

Bias: bias(\mathbf{x}_0) = $\mathbb{E}_p \widehat{p}_m(\mathbf{x}_0) - p(\mathbf{x}_0)$

We get the following decomposition

Lemma

$$MSE(\widehat{p}_m, p, \mathbf{x}_0) = bias^2(\mathbf{x}_0) + \mathbb{V}_p(\widehat{p}_m(\mathbf{x}_0)) =$$

$$= [\mathbb{E}_p \widehat{p}_m(\mathbf{x}_0) - p(\mathbf{x}_0)]^2 + \mathbb{E}_p[\widehat{p}_m(\mathbf{x}_0) - \mathbb{E}_p \widehat{p}_m(\mathbf{x}_0)]^2$$

Lemma

$$MISE(\widehat{p}_m, p) = \int_{\mathbb{R}} bias^2(\mathbf{x}) d\mathbf{x} + \int_{\mathbb{R}} \mathbb{V}_p(\widehat{p}_m(\mathbf{x})) dx$$

We will use these statements when constructing optimal density estimates

KDE: MISE

A shape of K influence quality of estimate not so significant compared to a value of \boldsymbol{h}

Theorem

$$MISE(\widehat{p}_m, p) \approx \frac{1}{4} \sigma_K^4 h^4 \int_{\mathbb{R}} (p''(\mathbf{x}))^2 d\mathbf{x} + \frac{1}{mh} \int_{\mathbb{R}} (K(\mathbf{x}))^2 d\mathbf{x}$$

For an optimal kernel width we get that $MISE(\widehat{p}_m,p)=O\left(m^{-\frac{4}{5}}\right)$

- Again we can not calculate h^* in practice, since it depends on unknown values of $r(\mathbf{x}), p(\mathbf{x})$
- Then we should minimize the risk estimate w.r.t. h

$$\widehat{\mathcal{J}}(h) = \sum_{i=1}^{m} (y_i - \widehat{r}_{(-i)}^{NW}(\mathbf{x}_i))^2,$$

where $\hat{r}_{(-i)}^{NW}$ is a Nadaraya-Watson estimate, constructed using the sample, from which the observation (\mathbf{x}_i,y_i) is excluded

Theorem

$$\widehat{\mathcal{J}}(h) = \sum_{i=1}^{m} \left(y_i - \widehat{r}^{NW}(\mathbf{x}_i) \right)^2 \frac{1}{\left(1 - \frac{K(0)}{\sum_{j=1}^{m} K\left(\frac{\mathbf{x}_i - \mathbf{x}_j}{h} \right)} \right)^2}$$

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