

Bayesian Interval Estimation

CHAPTER 4. BAYESIAN ESTIMATION AND LOSS FUNCTIONS

Practical Question

A researcher collects standardized exam score differences from a random sample of students. The data are obtained from the R built-in dataset `sleep`, using the variable `extra`. The dataset is as given below



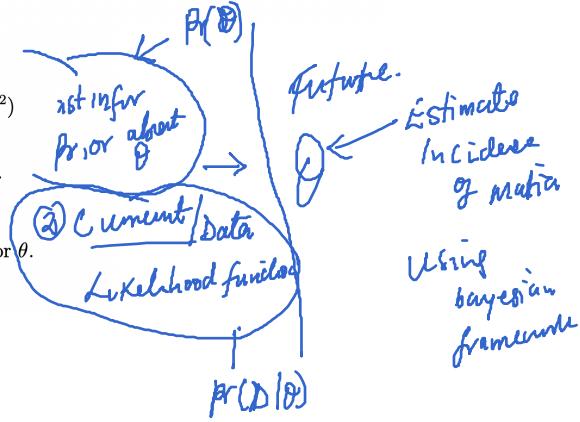
Assume the Bayesian model:

- Likelihood:
- Prior:

$$X_i | \theta \sim N(\theta, \sigma^2)$$

$$\theta \sim N(0, 100).$$

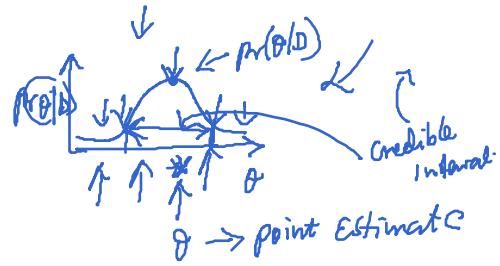
- Derive the posterior distribution of $\theta | x$.
- Construct a 95% equal-tailed credible interval for θ .
- Interpret the credible interval.



$$\text{Prior} \rightarrow \text{Normal } (0, 100)$$

$$\text{Likelihood} \rightarrow \text{Normal } (\theta, \sigma^2)$$

$$\text{Posterior} \rightarrow \text{Normal}$$



$$\theta | x \sim N(\mu_n, \sigma_n^2)$$

$$\sigma_n^2 = \left(\frac{n}{\sigma^2} + \frac{1}{T^2} \right)^{-1} \quad \mu_n = \sigma_n^2 \left(\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{T^2} \right)$$

$$\mu_0 = 0, T^2 = 100, \bar{x} = \frac{0.7 + -1.6 + \dots + 3.4}{20} = 1.54$$

$$n = 20$$

$$\sigma^2 = S^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 = 4.54$$

$$\sigma_n^2 = \left(\frac{20}{4.54} + \frac{1}{100} \right)^{-1} = 0.2265$$

$$\mu_n = 0.2265 \left(\frac{20 \times 1.54}{4.54} + \frac{0}{100} \right) = 1.54$$

ID	Group	Extra
1	1	0.7
2	1	-1.6
3	1	-0.2
4	1	-1.2
5	1	-0.1
6	1	3.4
7	1	3.7
8	1	0.8
9	1	0.0
10	1	2.0
1	2	1.9
2	2	0.8
3	2	1.1
4	2	0.1
5	2	-0.1
6	2	4.4
7	2	5.5
8	2	1.6
9	2	4.6
10	2	3.4

$$\frac{3.4596}{\text{sum}} \Rightarrow \frac{\sum (x_i - \bar{x})^2}{\text{sum}}$$

$$\theta | x \sim N(1.54, 0.2265)$$

(b) 95% equal tail interval

$$\text{Interval} \quad [\bar{x}_n - 1.96\delta_n, \bar{x}_n + 1.96\delta_n].$$

$$\delta_n = \sqrt{\delta_n^2} = \sqrt{0.2265} = 0.476.$$

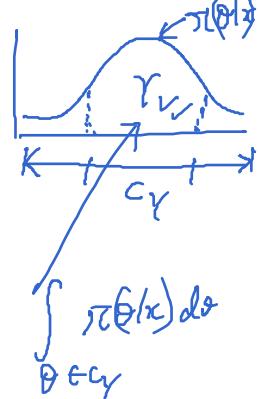
$$\begin{aligned} CI &= [1.54 - 1.96(0.476), 1.54 + 1.96(0.476)] \\ &= [0.61, 2.47] \end{aligned}$$

(c) Interpretation: There is 95% posterior probability that the true mean lies between 0.61 and 2.47 given the observed data & prior info.

Highest Posterior Density Interval (HPD) / Highest Density Interval (HDI)

Let $\pi(\theta|x)$ be the posterior density of θ . A set C_γ is called γ -level HPD credible set if

$$\Pr(\theta \in C_\gamma | x) = \int_{\theta \in C_\gamma} \pi(\theta|x) d\theta = \gamma.$$



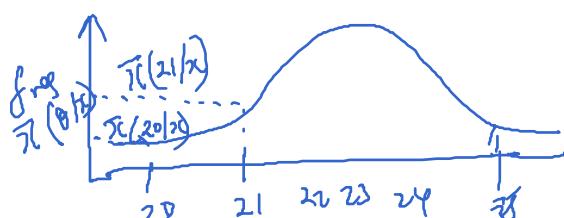
the highest density condition:

$$\text{for any } \theta_1 \in C_\gamma \text{ and } \theta_2 \notin C_\gamma \quad \int_{\theta_1}^{\theta_2} \pi(\theta|x) d\theta = 0$$

$$\underline{\pi}(\theta_1|x) \geq \underline{\pi}(\theta_2|x) =$$

$$\theta = \{20, 21, 22, 23, 24, 25\}$$

$$C_\gamma = \{21, 22, 23, 24\}$$



$$\begin{aligned} \theta_1 &\\ \theta_2 &\\ \underline{\pi}(\theta_1|x) &\geq \underline{\pi}(\theta_2|x) \end{aligned}$$

$$\underline{\pi}(21|x) + \underline{\pi}(22|x) + \underline{\pi}(23|x) + \underline{\pi}(24|x) = \gamma \quad \leftarrow \text{HDI}$$

$$C_\gamma = [21, 24] \leftarrow \text{HDI}$$

$$C_{90} = [21, 24]$$

The Equivalent definition:

$$\text{HPD} \quad C_\gamma = \{ \theta : \pi(\theta|x) \geq K_\gamma \},$$

where K is the constant chosen s.t

$$\int_{\pi(\theta|x) \geq K_\gamma} \pi(\theta|x) d\theta = \gamma.$$

HPD consist of all parameter θ values whose posterior density $\pi(\theta|x)$ is above certain cut off K and the limit γ is chosen so that exactly γ of different mass is included.

$$\theta = [20, 21, 22, 23, 24, 25]$$

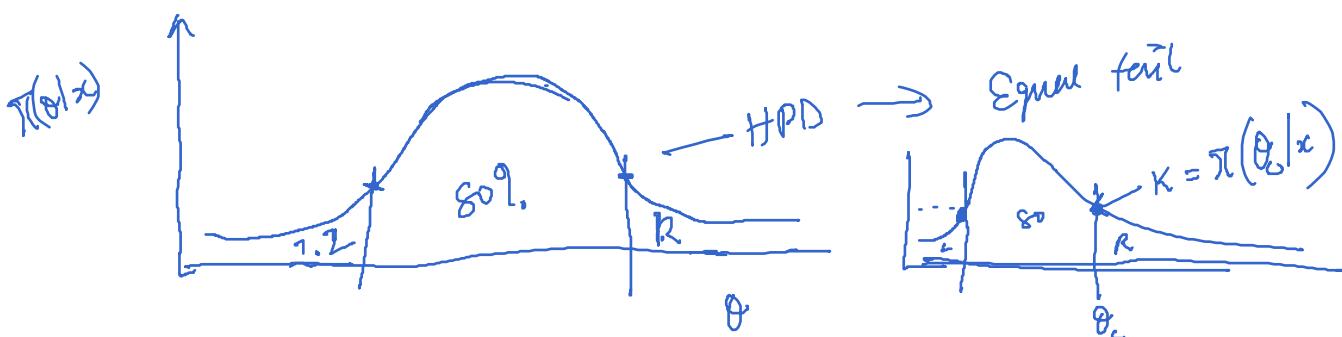
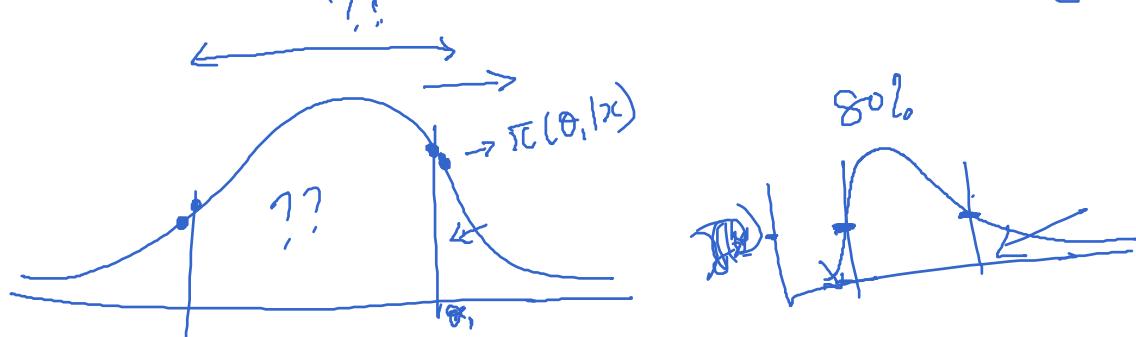
$\pi(\theta|x) = \sum \pi(20|x), \pi(21|x) = \dots, \pi(24|x), \pi(25|x)$

$\pi(20|x) < K$ $\gamma \approx \text{proportion}$ $\pi(25|x) < K$

$K = 0.05$

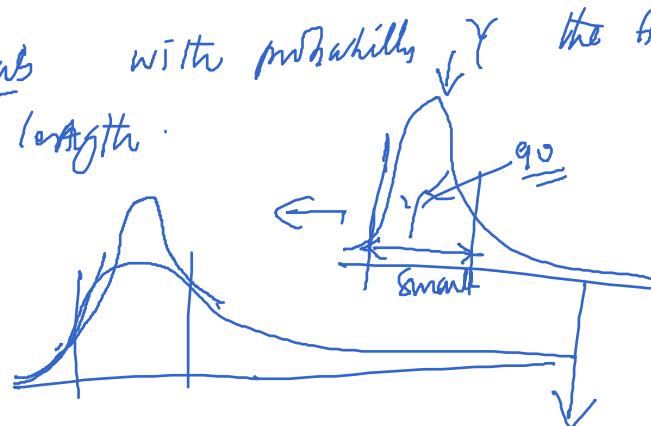
$$\gamma = \frac{4}{6} \times 100 = 66\% = [$$

$$66\% \text{credible interval} = [21, 24].$$



Among all Credible intervals with probability γ the HPD interval has the smallest length.

$$(a) \theta|x \sim N(\mu, \sigma^2) \quad \gamma = 0.95 \\ \text{HPD} = \text{Equal tail} \quad \text{HPD} = [\mu - 1.96\sigma, \mu + 1.96\sigma] \\ C_{0.95} = \left[\mu - q_{0.95}, \mu + q_{0.95} \right]$$

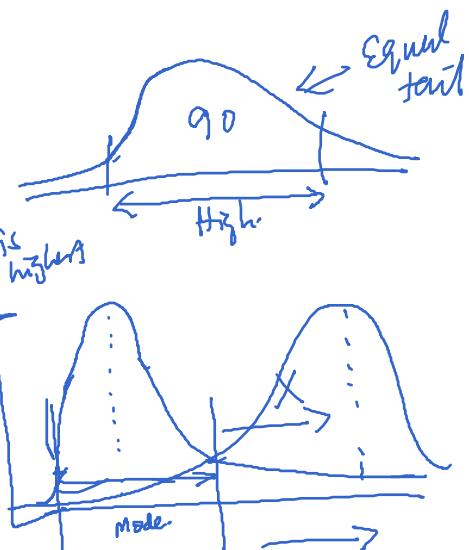


(b) When posterior distribution is not normal (Type I error)

① Equal tail is not shortest

② HPD moves towards the mode

HPD — Contains the mode
always contain (MAP Estimate)



③ HPD will always have an equal tails

Example :-

Suppose the posterior distⁿ of θ is

$$\theta|x \sim \text{Gamma}(\alpha=4, \beta=1)$$

cdf of Gamma

$$\pi(\theta|x) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta} \quad \theta > 0$$

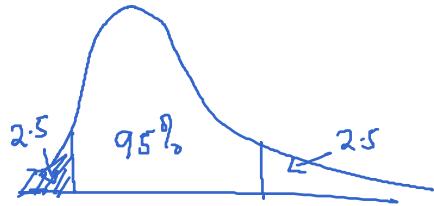
$$= \frac{1}{\Gamma(4)} \theta^3 e^{-\theta} \quad \theta > 0$$

$$\Gamma(4) = 3\Gamma(3) = 3 \times 2 \Gamma(2) = 3 \times 2 \times 1 \Gamma(1)^{-1} = 6$$

$$= \frac{1}{6} \theta^3 e^{-\theta}$$

$$F(\theta) = \underline{\Pr(\theta \leq \theta/x)} = \int_0^\theta \pi(t/x) dt.$$

Equal tail : 95%.



$$95\% CI \rightarrow [a, b]$$

$$\underline{\Pr(\theta < a)} = 0.025, \quad \underline{\Pr(\theta > b)} = 0.025$$

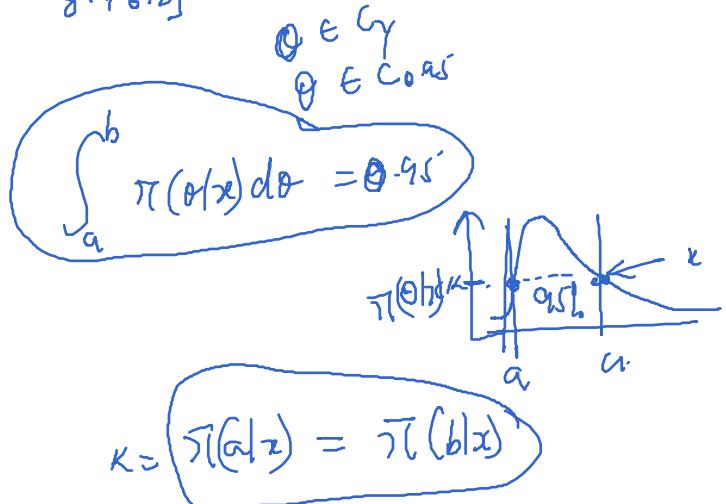
$$F(a) = 0.025$$

$$\downarrow \quad F(b) = \Pr(\theta < b) = 0.975$$

$$[a, b] = [\theta_{0.025}, \theta_{0.975}] \\ = [1.089, 8.767]$$

$$C_{95\%} = [a, b]$$

HPD 95% Credible Interval
 $\gamma = 95\%$



$$\pi(\theta|x) = \frac{1}{6} \theta^3 e^{-\theta}$$

$$\frac{1}{6} a^3 e^{-a} = \frac{1}{6} b^3 e^{-b}$$

$$a^3 e^{-a} = b^3 e^{-b}$$

$$3 \ln a - a = 3 \ln b - b \quad \dots \textcircled{1}$$



$$\int_a^b \pi(\theta|x) d\theta$$

Also

$$\Pr(a < \theta < b | x) = 0.95$$

$$F(b) - F(a) = 0.95 \dots \textcircled{ii}$$

$$\begin{aligned}
 a &= 0.7125 \\
 b &= 7.9483 \\
 [a, b] &= [0.7125, 7.9483] \quad \text{HPD} \\
 &= [1.089, 8.7675] \quad \text{equal tail} \\
 &\quad \leftarrow \text{long data}
 \end{aligned}$$

Example 2

Monthly snakebite cases recorded in a county over one year are:

$$y = (3, 5, 4, 6, 2, 7, 5, 4, 6, 5, 3, 4).$$

Assume the Bayesian model:

- Likelihood:

$$Y_i | \theta \sim \text{Poisson}(\theta).$$

- Prior:

$$\theta \sim \text{Gamma}(1, 1),$$

where the second parameter is the rate.

- Derive the posterior distribution of $\theta | y$.
- Compute the 95% equal-tailed credible interval.
- Obtain the 95% HPD credible interval.
- Compare the two intervals.

Solution : Likelihood $y | \theta \sim \text{pois}(\theta)$

$$\begin{aligned}
 L(\theta | y) &= \prod_{i=1}^n \frac{\theta^{y_i} e^{-\theta}}{y_i!} \\
 &= \frac{\theta^{\sum y_i} e^{-\theta}}{\prod y_i!} \rightarrow \frac{1}{\prod y_i!} \theta^{\sum y_i} e^{-\theta}
 \end{aligned}$$

$$L(y | \theta) \propto \boxed{\theta^{\sum y_i} e^{-\theta}}$$

Posterior $\pi(\theta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta}$

$$\pi(\theta) \propto \boxed{\theta^{\alpha-1} e^{-\beta\theta}}$$

$$\pi(\theta | y) \propto \theta^{\alpha + \sum y_i - 1} e^{-\frac{(\beta + \sum y_i)}{\beta} \theta}$$

$$\boxed{\theta^{\alpha+1} e^{-\beta\theta}}$$

$$\sum y_i = 54 \quad n = 12$$

$$\alpha_x = 1 + 54$$

$$\beta_x = 1 + 12 = 13$$

$\pi(\theta|y) \sim \text{Gamma}(55, 13)$

(b) 95% equal tails

$$(a, b) = [0.025, 0.975]$$

$$= (3.09, 5.88) \leftarrow$$

(c) $P(a < \theta < b)$

$$P_{0.025}^{0.975} = \{3.21, 5.72\} \leftarrow \text{shorter length}$$

