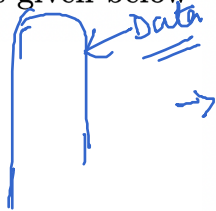


Bayesian Interval Estimation

CHAPTER 4. BAYESIAN ESTIMATION AND LOSS FUNCTIONS

Practical Question

A researcher collects standardized exam score differences from a random sample of students. The data are obtained from the R built-in dataset `sleep`, using the variable `extra`. The dataset is as given below



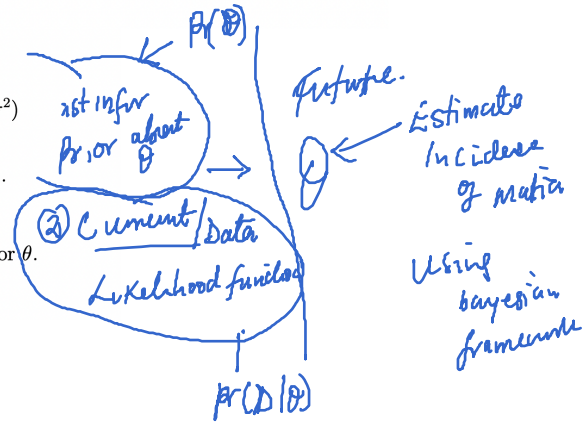
Assume the Bayesian model:

- Likelihood:
- Prior:

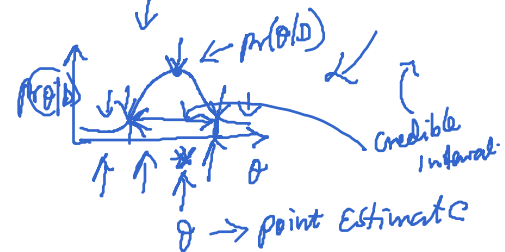
$$X_i | \theta \sim \mathcal{N}(\theta, \sigma^2)$$

$$\theta \sim \mathcal{N}(0, 100).$$

- Derive the posterior distribution of $\theta | x$.
- Construct a 95% equal-tailed credible interval for θ .
- Interpret the credible interval.



Prior \rightarrow Normal (μ_0, τ^2)
 Likelihood \rightarrow Normal (θ, σ^2)
 Posterior \rightarrow Normal.



θ/x

$$\theta/x \sim \mathcal{N}(\mu_n, \sigma_n^2)$$

$$\sigma_n^2 = \left(\frac{n}{\sigma^2} + \frac{1}{\tau^2} \right)^{-1} \quad \mu_n = \sigma_n^2 \left(\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2} \right)$$

$$\mu_0 = 0, \tau^2 = 100, \quad \bar{x} = \frac{0.7 + -1.6 + \dots + 3.4}{20} = 1.54$$

$$n = 20$$

$$\sigma^2 = s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 = 4.54$$

$$\sigma_n^2 = \left(\frac{20}{4.54} + \frac{1}{100} \right)^{-1} = 0.2265$$

$$\mu_n = 0.2265 \left(\frac{20 \times 1.54}{4.54} + \frac{0}{100} \right) = 1.54$$

ID	Group	Extra	$(x_i - \bar{x})^2$
1	1	0.7	0.7056
2	1	-1.6	
3	1	-0.2	
4	1	-1.2	
5	1	-0.1	
6	1	3.4	
7	1	3.7	
8	1	0.8	
9	1	0.0	
10	1	2.0	
1	2	1.9	
2	2	0.8	
3	2	1.1	
4	2	0.1	
5	2	-0.1	
6	2	4.4	
7	2	5.5	
8	2	1.6	
9	2	4.6	
10	2	3.4	

3.4596

sum $\Rightarrow \sum (x_i - \bar{x})^2 \Rightarrow$

$$\theta/x \sim \mathcal{N}(1.54, 0.2265)$$

(b) 95% equal tail interval

$$\text{Interval } [\mu_n - 1.96\delta_n, \mu + 1.96\delta_n].$$

$$\delta_n = \sqrt{\sigma_n^2} = \sqrt{0.2265} = 0.476.$$

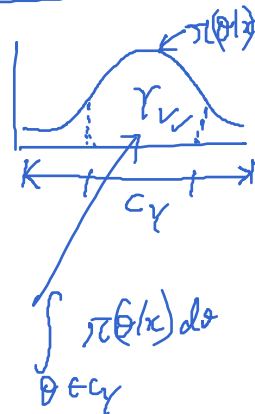
$$CI = [1.54 - 1.96(0.476), 1.54 + 1.96(0.476)] \\ = [0.61, 2.47]$$

(c) Interpretation: There is 95% posterior probability that the true mean lies between 0.61 and 2.47 given the observed data & prior info

Highest Posterior Density Interval (HPD) / Highest Density Interval (HDI)

Let $\pi(\theta|x)$ be the posterior density of θ
A set C_γ is called γ -level HPD credible set if

$$Pr(\theta \in C_\gamma | x) = \int_{\theta \in C_\gamma} \pi(\theta|x) d\theta = \gamma.$$



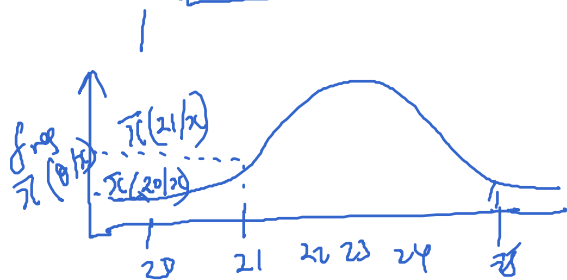
the highest density condition:

$$\text{for any } \theta_1 \in C_\gamma \text{ and } \theta_2 \notin C_\gamma \quad \int_{\theta_1} \pi(\theta|x) d\theta \geq \int_{\theta_2} \pi(\theta|x) d\theta$$

$$\pi(\theta_1|x) \geq \pi(\theta_2|x)$$

$$\theta = \{20, 21, 22, 23, 24, 25\}$$

$$C_\gamma = \{21, 22, 23, 24\}$$



$$\pi(21|x) \geq \pi(20|x)$$

$$\pi(21|x) + \pi(22|x) + \pi(23|x) + \pi(24|x) = \gamma \leftarrow \text{HPD}$$

$$C_\gamma = [21, 24] \leftarrow \text{HPD}$$

$$C_{90} = [21, 24]$$

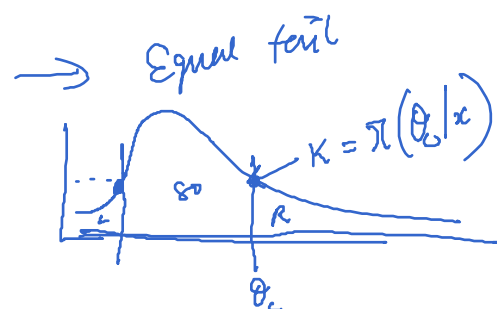
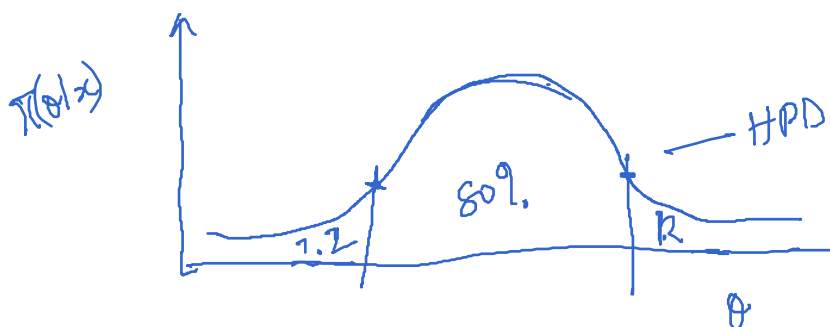
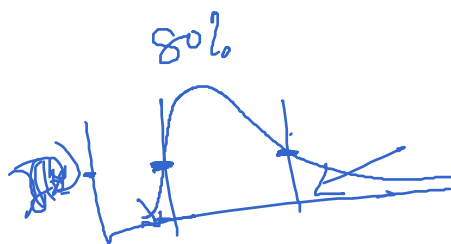
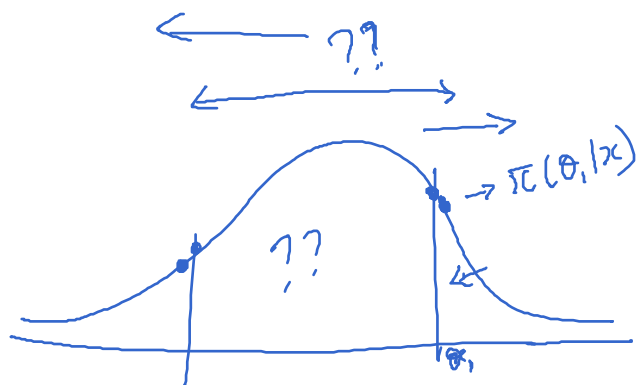
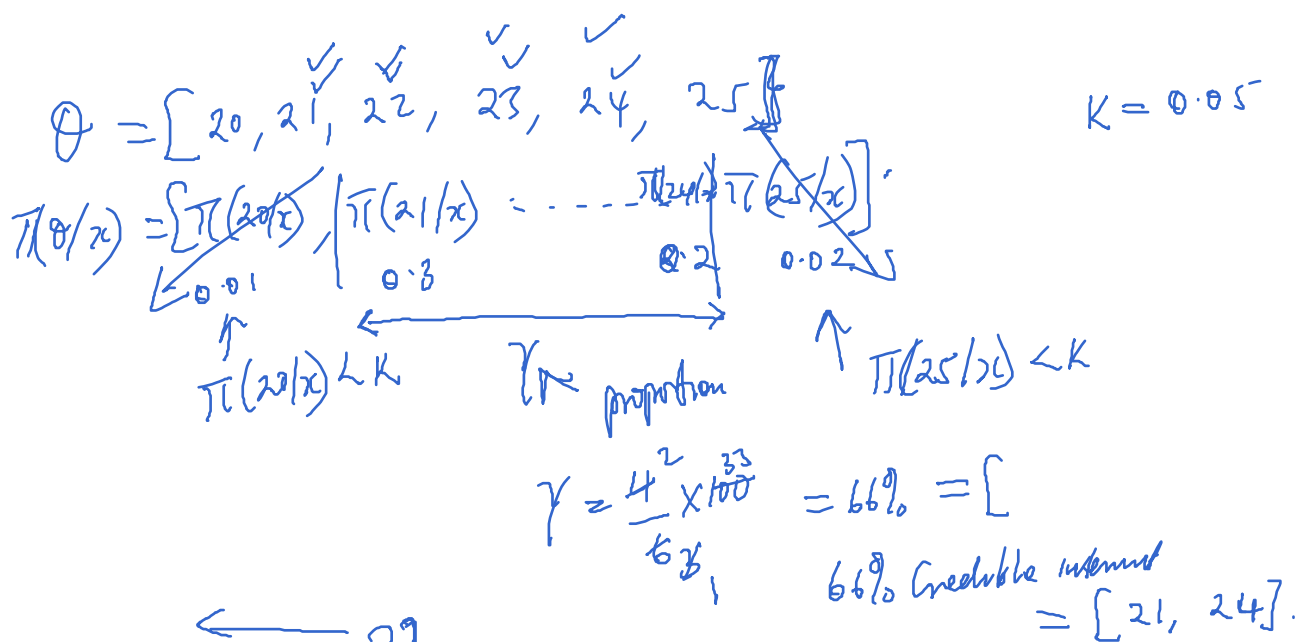
The Equivalent Definition

HPD $C_\gamma = \{ \theta : \pi(\theta|x) \geq K_\gamma \}$,

where K is the constant chosen s.t

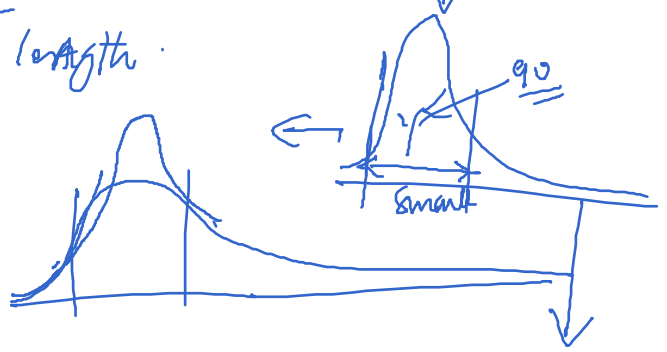
$$\int_{\pi(\theta|x) \geq K_\gamma} \pi(\theta|x) d\theta = \gamma.$$

HPD consist of all parameter θ values whose posterior density $\pi(\theta|x)$ is above certain cut off K and the cut off is chosen so that exactly γ of posterior mass is included.



Among all Credible intervals with probability γ the HPD interval has the smallest length.

(a) $\theta/x \sim N(\mu, \sigma^2)$ $\gamma = 0.95$
 HPD = Equal tail $\gamma = 0.95$
 $C_{0.95} = [\mu - 1.96\sigma, \mu + 1.96\sigma]$ \uparrow 95% CI

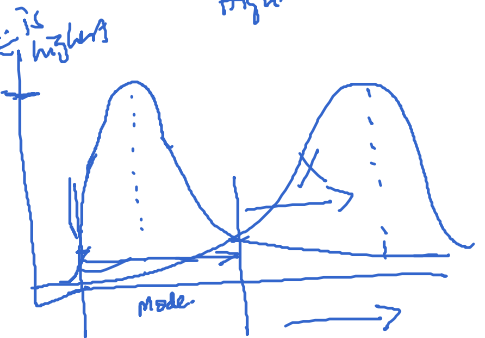
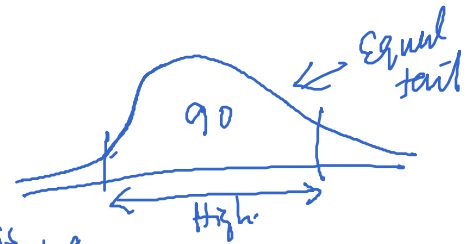


(b) When posterior is not normal
 (Typical) Gamma distⁿ / Beta / Lognormal

- (1) Equal tail is not shortest
- (2) HPD moves towards the mode

HPD — Contains the mode, it always contains (MAP estimate)

(3) HPD will always have an equal tails
 $\hat{\theta}_{MAP} = \arg \max_{\theta} \pi(\theta/x)$



Example :-

Suppose the posterior distⁿ of θ is

$\theta/x \sim \text{gamma}(\alpha=4, \beta=1)$

cdf of gamma

$\pi(\theta/x) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta} \quad \theta > 0$

$= \frac{1}{\Gamma(4)} \theta^3 e^{-\theta} \quad \theta > 0$

$$\Gamma(4) = 3\Gamma(3) = 3 \times 2\Gamma(2) = 3 \times 2 \times 1\Gamma(1) \stackrel{!}{=} 6$$

$$= \frac{1}{6} \theta^3 e^{-\theta}$$

$$F(\theta) = \Pr(\theta \leq \theta/x) = \int_0^\theta \pi(t/x) dt.$$

Equal tail 95%

95% CI $\rightarrow [a, b]$

$$\Pr(\theta < a) = 0.025, \quad \Pr(\theta > b) = 0.025$$

$$F(a) = 0.025$$

$$\Downarrow F(b) = \Pr(\theta < b) = 0.975$$

$$[a, b] = [\theta_{0.025}, \theta_{0.975}]$$

$$= [1.089, 8.7673]$$

$$C_{0.95} = [a, b]$$

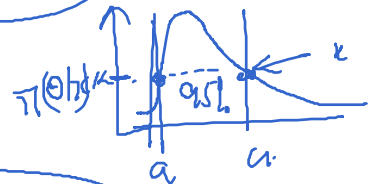
HPD

95% Credible Interval
 $\gamma = 95\%$

$$\int_a^b \pi(\theta/x) d\theta = 0.95$$

$$\theta \in C_\gamma$$

$$\theta \in C_{0.95}$$



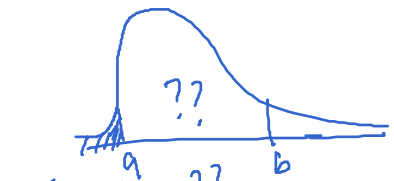
$$\kappa = \pi(a|x) = \pi(b|x)$$

$$\pi(\theta/x) = \frac{1}{6} \theta^3 e^{-\theta}$$

$$\frac{1}{6} a^3 e^{-a} = \frac{1}{6} b^3 e^{-b}$$

$$a^3 e^{-a} = b^3 e^{-b}$$

$$3 \ln a - a = 3 \ln b - b \quad \dots (i)$$



$$\int_a^b \pi(\theta/x) d\theta$$

Also

$$\Pr(a < \theta < b | x) = 0.95$$

$$F(b) - F(a) = 0.95 \quad \dots (ii)$$

$$\begin{aligned}
 a &= 0.7125 \\
 b &= 7.9483 \leftarrow \text{Short dist} \\
 [a, b] &= [0.7125, 7.9483] \leftarrow \text{HPD} \\
 &= [1.089, 8.767] \leftarrow \text{Equal tail.} \\
 &\quad \quad \quad \leftarrow \text{long dist.}
 \end{aligned}$$

Example 2

Monthly snakebite cases recorded in a county over one year are:

$$y = (3, 5, 4, 6, 2, 7, 5, 4, 6, 5, 3, 4).$$

Assume the Bayesian model:

- Likelihood:

$$Y_i | \theta \sim \text{Poisson}(\theta).$$

- Prior:

$$\theta \sim \text{Gamma}(1, 1),$$

where the second parameter is the rate.

- Derive the posterior distribution of $\theta | y$.
- Compute the 95% equal-tailed credible interval.
- Obtain the 95% HPD credible interval.
- Compare the two intervals.

Solution:

Likelihood

$$\begin{aligned}
 y | \theta &\sim \text{pois}(\theta) \\
 L(y | \theta) &= \prod_{i=1}^n \frac{\theta^{y_i} e^{-\theta}}{y_i!}
 \end{aligned}$$

$$\begin{aligned}
 &= \frac{\theta^{\sum y_i} e^{-n\theta}}{\prod y_i!} \Rightarrow \frac{1}{\prod y_i!} \theta^{\sum y_i} e^{-n\theta} \\
 L(y | \theta) &\propto \boxed{\theta^{\sum y_i} e^{-n\theta}}
 \end{aligned}$$

Prior

$$\pi(\theta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta}$$

$$\pi(\theta) \propto \boxed{\theta^{\alpha-1} e^{-\beta\theta}}$$

$$\begin{aligned}
 \pi(\theta | y) &\propto \frac{\theta^{\sum y_i + \alpha - 1} e^{-(\beta + n)\theta}}{\Gamma(\sum y_i + \alpha) \beta^{\sum y_i + \alpha}} \\
 &\propto \boxed{\theta^{\alpha_* - 1} e^{-\beta_* \theta}}
 \end{aligned}$$

$$\sum y_i = 54 \quad n = 12$$

$$\alpha_x = 1 + 54$$

$$\beta_x = 1 + 12 = 13$$

$$\pi(\theta|y) \sim \text{Gamma}(55, 13)$$

(b) 95% equal tails

$$(a, b) = [\theta_{0.025}, \theta_{0.975}]$$

$$= (3.09, 5.88) \leftarrow$$

(c) $\text{pr}(a \leq \theta \leq b)$

$$\text{HPD}_{0.95} = [3.21, 5.72] \leftarrow \text{shorter length}$$

