

How to Sample Unbounded Monotone Density Functions

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This short note shows a preprocessing algorithm to generate a random number in $[0, 1]$ from a distribution whose density function—

- is continuous in the interval $[0, 1]$,
- is monotonically decreasing in $[0, 1]$, and
- has an unbounded peak at 0.

The trick here is to sample the peak in such a way that the result is either forced to be 0 or forced to belong to the bounded part of the density function. This algorithm does not require the area under the curve of the density in $[0, 1]$ to be 1; in other words, this algorithm works even if the density function is known up to a normalizing constant. The algorithm is as follows.

1. Set i to 1.
2. Calculate the cumulative probability of the interval $[0, 2^{-i}]$ and that of $[0, 2^{-(i-1)}]$, call them p and t , respectively.
3. With probability p/t , add 1 to i and go to step 2. (Alternatively, if i is equal to or higher than the desired number of fractional bits in the result, return 0 instead of adding 1 and going to step 2.)
4. At this point, the density function at $[2^{-i}, 2^{-(i-1)})$ is bounded from above, so sample a random number in this interval using any appropriate algorithm, including rejection sampling. Because the density is monotonically decreasing, the peak of the density at this interval is located at 2^{-i} , so that rejection sampling becomes trivial.

It is relatively straightforward to adapt this algorithm for monotonically increasing density functions with the unbounded peak at 1, or to density functions with a different domain than $[0, 1]$.

This algorithm is similar to the “inversion-rejection” algorithm mentioned in section 4.4 of chapter 7 of Devroye’s *Non-Uniform Random Variate Generation* (1986)⁽¹⁾. I was unaware of that algorithm at the time I started writing this article. The difference here is that it assumes the whole distribution (including its density function and cumulative distribution function) is supported on the interval $[0, 1]$, while the algorithm presented in this article doesn’t make that assumption (e.g., the interval $[0, 1]$ can cover only part of the density’s support).

By the way, this algorithm arose while trying to devise an algorithm that can generate an integer power of a uniform random number, with arbitrary precision, without actually calculating that power (a naïve calculation that is merely an approximation and usually introduces bias); for more information, see my other article on [partially-sampled random numbers](#). Even so, the algorithm I have come up with in this note may be of independent interest.

In the case of powers of a uniform $[0, 1]$ random number X , namely X^n , the ratio p/t in this algorithm has a very simple form, namely $(1/2)^{1/n}$, which is possible to simulate using a so-called *Bernoulli factory* algorithm without actually having to calculate this ratio. Note that this formula is the same regardless of i . This is found by taking the density function $f(x) = x^{1/n}/(x * n)$ and finding the appropriate p/t ratios by integrating f over the two intervals mentioned in step 2 of the algorithm.

Notes

⁽¹⁾ Devroye, L., [*Non-Uniform Random Variate Generation*](#), 1986.

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