

Q5)

A) I will try to compute the posterior using the multiplication of the prior of the mean with the likelihood and then complete the square in the exponent. $p(\theta|x) = p(\theta) \cdot p(x|\theta)$; putting in the necessary data from the problem statement we have something like the following:

$$N(\mu, \sigma^2) \cdot N(\mu_0, \sigma_0^2) = c \cdot \exp(- (x_1 - \mu)^2 / 2\sigma^2 - (\mu - \mu_0)^2 / 2\sigma_0^2)$$

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② $P(\mu) = \frac{1}{\sqrt{2\pi}\sigma_0} \exp\left(\frac{-(\mu - \mu_0)^2}{2\sigma_0^2}\right)$

$P(x|\mu) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(\frac{-(x - \mu)^2}{2\sigma^2}\right)$

$P(\mu|x) = \frac{P(\mu) P(x|\mu)}{\int_{-\infty}^{\infty} P(\mu) P(x|\mu) d\mu} = \frac{P(\mu) P(x|\mu)}{P(x)}$

$\propto P(\mu) P(x|\mu)$

$= \frac{1}{2\pi\sqrt{\sigma_0^2\sigma^2}} \exp\left\{-\frac{\mu^2 + 2\mu\mu_0 + \mu_0^2}{2\sigma_0^2} - \frac{x^2 - 2\mu x + \mu^2}{2\sigma^2}\right\}$

$= (\text{constant}) \exp\left\{-\frac{\mu^2\sigma^2 + 2\mu\mu_0\sigma^2 + \mu_0^2\sigma^2 - \sigma^2x^2 + 2\mu\sigma^2x - \mu^2\sigma^2}{2\sigma_0^2\sigma^2}\right\}$

$\propto \exp\left\{-\frac{\mu^2(\sigma^2 + \sigma_0^2) + 2\mu(\mu_0\sigma^2 + \sigma_0^2x) - (\mu_0^2\sigma^2 + \sigma_0^2x^2)}{2\sigma_0^2\sigma^2}\right\}$

$\propto \exp\left\{-\frac{\mu^2 + 2\mu\left(\frac{\mu_0\sigma^2 + \sigma_0^2x}{\sigma^2 + \sigma_0^2}\right) - \left(\frac{\mu_0\sigma^2 + \sigma_0^2x}{\sigma^2 + \sigma_0^2}\right)^2}{2\left(\frac{\sigma_0^2\sigma^2}{\sigma^2 + \sigma_0^2}\right)}\right\}$

$\propto \exp\left\{-\frac{\mu_0^2\sigma^2 + \sigma_0^2x^2}{2\sigma_0^2\sigma^2}\right\}$

$\propto \exp\left\{-\frac{\left(\mu - \left(\frac{\mu_0\sigma^2 + \sigma_0^2x}{\sigma^2 + \sigma_0^2}\right)\right)^2}{2\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma^2}\right)}\right\}$

μ_1

σ_1^2

B)

$$\textcircled{b} \quad P(x|x_1) = \int_{\theta} P(x, \theta | x_1) d\theta = \int_{\theta} P(x | x_1, \theta) P(\theta | x_1) d\theta$$

$$= \int_{\theta} P(x | \theta) P(\theta | x_1) d\theta$$

$$\int_{\theta} P(x | \theta) P(\theta | x_1) d\theta = \int_{\theta} P(x = x_1 - \theta + \theta | \theta) P(\theta | x_1) d\theta$$

$$= \int_{\theta} P(x = x_1 + \theta | \theta) P(\theta | x_1) d\theta = \int_{\theta} P(x - \theta = x_1 | \theta) P(\theta | x_1) d\theta$$

$$= \int_{\theta} P(x - \theta | \theta) P(\theta | x_1) d\theta$$

→ convolution of two functions.

$$\Rightarrow x | x_1 \sim N(\mu_1, \sigma_x^2 + \tau_1^2)$$

given

calculated in part

(a)

calculated in part (a)

$$(e) \quad \mu \sim N(\mu_0, \sigma_0^2) \quad x_i | \mu \sim N(\mu, \sigma^2)$$

$$P(\mu | \underline{x}) \propto P(\mu) P(\underline{x} | \mu) = P(\mu) P(x_1 | \mu) P(x_2 | \mu) \dots P(x_n | \mu)$$

$$= \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp\left\{-\frac{(\mu - \mu_0)^2}{2\sigma_0^2}\right\} \times \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x_i - \mu)^2}{2\sigma^2}\right\}$$

$$= \frac{1}{(2\pi)^{n+1/2} \sqrt{\sigma_0^2 \sigma^{2n}}} \exp\left\{-\frac{\mu^2 + 2\mu\mu_0 - \mu^2}{2\sigma_0^2} - \sum_{i=1}^n \frac{x_i^2 - 2\mu x_i + \mu^2}{2\sigma^2}\right\}$$

$$\propto \exp\left\{-\frac{\mu^2(\sigma^2 + n\sigma_0^2) + 2\mu(\mu_0\sigma^2 + \sigma_0^2 x_1 + \dots + \sigma_0^2 x_n) - (\mu_0^2\sigma^2 + \dots + \sigma_0^2 x_n^2)}{2\sigma_0^2 \sigma^2}\right\}$$

$$\propto \exp\left\{-\frac{\mu^2 + 2\mu\left(\frac{\mu_0\sigma^2 + \sum \sigma_0^2 x_i}{\sigma^2 + n\sigma_0^2}\right) - \left(\frac{\mu_0\sigma^2 + \sum \sigma_0^2 x_i}{\sigma^2 + n\sigma_0^2}\right)^2}{2 \frac{\sigma_0^2 \sigma^2}{\sigma^2 + n\sigma_0^2}}\right\}$$

\times

$$\exp\left\{-\frac{\mu^2\sigma^2 + \sum \sigma_0^2 x_i^2}{2\sigma_0^2 \sigma^2}\right\}$$

$$\mu_1 = \sigma_1^2 \left(\frac{\mu_0}{\sigma_0^2} + \frac{\bar{x}}{\sigma^2/n} \right)$$

\Rightarrow sample mean

$$\propto \exp\left\{-\frac{\left(\mu - \frac{\mu_0\sigma^2 + \sum \sigma_0^2 x_i}{\sigma^2 + n\sigma_0^2}\right)^2}{2 \frac{\sigma_0^2 \sigma^2}{\sigma^2 + n\sigma_0^2}}\right\}$$

$$\sigma_1^2 = \left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)^{-1}$$

$$p(\tilde{x} | x) = \frac{1}{(\sqrt{2\pi})^m} \det \left(\sigma^2 I + T_n^2 i i^T \right)^{-1/2} \cdot \exp \left(-\frac{1}{2} (\tilde{x} - \mu_{ni})^T (\sigma^2 I + T_n^2 i i^T)^{-1} (\tilde{x} - \mu_{ni}) \right)$$

$$\tilde{x} = \begin{bmatrix} x_{n+1} \\ \vdots \\ x_{n+m} \end{bmatrix}$$

I is an $m \times m$ matrix.

μ_{ni} (posterior mean of μ)

$\sigma^2 I + T_n^2 i i^T$ (T_n = posterior variance of μ)
 and the whole term is the covariance matrix.

extending the same logic as the case where $m=1$

↳ we can be a little inventive too —, we can consider the sufficient statistics of the data that is $T(x) = \bar{x}$ = sample mean of $\{x_1, \dots, x_n\}$

can convert the whole problem into the univariate case. again.