

	EXAMINATION QUESTIONS/SOLUTIONS 2009-10	Course Comp245																																
Question 1.		Marks & seen/unseen																																
Parts	<div>(i) (e). Using Bayes Theorem, $P(\overline{B} A) = \frac{P(A \cap \overline{B})}{P(A)} = \frac{P(A \overline{B})P(\overline{B})}{P(A)} = \frac{(1 - P(\overline{A} \overline{B}))(1 - P(B))}{P(A)}$$= \frac{(1 - 0.8)(1 - 0.3)}{0.2} = 0.7.$</div> <div>(ii) (d). From the standard normal table, $\Phi(1.282) \approx 0.9$ and so $\Phi(-1.282) \approx 0.1$. It follows that $\frac{90 - 100}{\sigma} \approx -1.282 \implies \sigma \approx \frac{10}{1.282} \approx 7.8.$</div> <div>(iii) (e).</div> <div>(iv) (f). For $X_i \sim \text{Exponential}(\lambda)$, $f(x_i) = \lambda \exp(-\lambda x_i)$. So, $\ell(\lambda) = \sum_{i=1}^3 \log f(x_i) = 3 \log(\lambda) - \lambda \sum_{i=1}^3 x_i$$\frac{d\ell(\lambda)}{d\lambda} = \frac{3}{\lambda} - \sum_{i=1}^3 x_i$$\implies \hat{\lambda} = \frac{3}{\sum_{i=1}^3 x_i} = \frac{3}{0.1 + 0.5 + 0.9} = \frac{3}{1.5} = 2.$</div> <div>(v) (b). <table><tr><td></td><td></td><td>1.9GHz</td><td>2.1GHz</td><td></td></tr><tr><td rowspan="3">Observed:</td><td>SATA</td><td>26</td><td>27</td><td>53</td></tr><tr><td>SSD</td><td>48</td><td>81</td><td>129</td></tr><tr><td></td><td>74</td><td>108</td><td>182</td></tr><tr><td></td><td></td><td>1.9GHz</td><td>2.1GHz</td><td></td></tr><tr><td rowspan="2">Expected:</td><td>SATA</td><td>21.55</td><td>31.45</td><td></td></tr><tr><td>SSD</td><td>52.45</td><td>76.55</td><td></td></tr></table> Test statistic $X^2=2.1853$.</div>			1.9GHz	2.1GHz		Observed:	SATA	26	27	53	SSD	48	81	129		74	108	182			1.9GHz	2.1GHz		Expected:	SATA	21.55	31.45		SSD	52.45	76.55		<div>seen ↓</div> <div>unseen ↓</div> <div>seen ↓</div> <div>Each 4 marks</div>
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Question 2.		Marks & seen/unseen
Parts	<p>(i) (a) $X \sim \text{Geometric}(0.2)$. From the formula sheet, $E(X) = 1/0.2 = 5$.</p> <p>(b)</p> $P(X \leq 2) = P(X = 1) + p(X = 2) = 0.2 + 0.2 * (1 - 0.2) = 0.36.$ <p>(c) [A one-sided hypothesis test for a geometric parameter higher than 0.2 would consider the left tail of the distribution, so 0.36 is in fact the p-value for this test.] There is very little evidence to reject the null hypothesis of 80% next day deliveries.</p> <p>(ii) (a) $Y \sim \text{Binomial}(20, 0.2)$. From the formula sheet, $E(Y) = 20 \times 0.2 = 4$.</p> <p>(b)</p> $P(Y \leq 5) = \sum_{i=0}^5 P(Y = i) = \sum_{i=0}^5 \binom{20}{i} 0.2^i \times 0.8^{20-i} \approx 0.804.$ <p>(c) [A one-sided hypothesis test for a binomial probability parameter higher than 0.2 would consider the right tail of the distribution. The p-value for the test is:]</p> $P(Y \geq 6) = 1 - P(Y \leq 5) \approx 0.196.$ <p>So there is more evidence against the null hypothesis than before, but insufficient to reject at any typical significance level.</p> <p>(iii) Let Z be the number of parcels delivered on time, then $Z \sim \text{Binomial}(1000, 0.8)$. So $E(Z) = 1000 \times 0.8 = 800$ and $\text{Var}(Z) = 1000 \times 0.8 \times (1 - 0.8) = 160$, and hence by the CLT approximately $Z \sim N(800, 160)$.</p> $P(Z \leq 774) \approx \Phi\left(\frac{774-800}{\sqrt{160}}\right) \approx \Phi(-2.055) \approx 0.020.$ <p>So with this larger sample there is now much stronger evidence that 80% is not being achieved.</p>	<div>seen ↓</div> <div>4 marks</div> <div>3 marks</div> <div>unseen ↓</div> <div>1 marks</div> <div>seen ↓</div> <div>4 marks</div> <div>3 marks</div> <div>unseen ↓</div> <div>1 marks</div> <div>seen ↓</div> <div>4 marks</div>
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Question 3.		Marks & seen/unseen
Parts	<p>(i) Properties of a pdf:</p> <p>I. $f(x) \geq 0, \forall x \in \mathbb{R};$</p> <p>II. $\int_{x=-\infty}^{\infty} f(x)dx = 1.$</p> <p>From the second point, we have</p> $1 = \int_{x=-\infty}^{\infty} f(x)dx = \int_{x=1}^{\infty} \frac{c}{x^3}dx = \frac{-c}{2x^2} \Big _{x=1}^{\infty} = \frac{c}{2}.$ <p>So $c = 2.$</p> <p>(ii)</p> $E(X) = \int_{x=-\infty}^{\infty} xf(x)dx = \int_{x=1}^{\infty} \frac{2}{x^2}dx = \frac{-2}{x} \Big _{x=1}^{\infty} = 2.$ <p>(iii) To find the median, we first work out the cdf:</p> $F(x) = \int_{u=-\infty}^x f(u)du = \int_{u=1}^x \frac{2}{u^3}du = \frac{-1}{u^2} \Big _{u=1}^x = 1 - \frac{1}{x^2}.$ <p>The median is then x satisfying $F(x) = 1/2.$</p> $\frac{1}{2} = 1 - \frac{1}{x^2} \iff \frac{1}{x^2} = \frac{1}{2} \implies x = \sqrt{2}.$ <p>So the median is $\sqrt{2}.$</p> <p>(iv) We have seen the median is less than the mean, suggestive of right skew. This is consistent with the shape of the density function $f(x)$, which is decreasing in x on $[1, \infty).$</p>	<div>seen ↓</div> <div>unseen ↓</div> <div>5 marks</div> <div>5 marks</div> <div>7 marks</div> <div>3 marks</div>
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Question 4.		Marks & seen/unseen
Parts	<p>(i) (a) $\text{bias}(T) = E[T \theta] - \theta$.</p> <p>(b) Obvious unbiased estimators for mean and variance are $\bar{x} = 1001.625$ and $s_{n-1}^2 = 117.982$ respectively.</p> <p>(c) Let X be the weight of a randomly chosen bag of potatoes, so approximately $X \sim N(1001.625, 117.982)$. Then $P(X \leq 990) \approx \Phi\left(\frac{990 - 1001.625}{\sqrt{117.982}}\right) \approx \Phi(-1.070) \approx 0.14$.</p> <p>(ii) (a) Let $z = \frac{\bar{x} - \mu}{\sqrt{\sigma^2/n}} = \frac{1001.625 - 1000}{\sqrt{100/8}} = 0.460$. From the formula sheet, the rejection region of a two-sided normal test at the 1% level is $(-2.576, 2.576)$, so there is insufficient to reject the null hypothesis of the mean being 1kg.</p> <p>(b) If each bag weight X_i is i.i.d. $N(1000, 100)$ then the sum S_5 of the weights of five independent bags is $N(5000, 500)$. So $P(S_5 > 4990) = 1 - \Phi\left(\frac{4990 - 5000}{\sqrt{500}}\right) \approx 1 - \Phi(-0.447) \approx 0.673$.</p>	<div>seen ↓</div> <p>1 marks</p> <p>6 marks</p> <div>unseen ↓</div> <p>3 marks</p> <div>seen ↓</div> <p>7 marks</p> <p>3 marks</p>
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