Inference part 1 STAT-S520

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► These slides complement material from ISI Chapter 9

So far, we have a random sample $X_1, X_2, \dots, X_n \stackrel{i.i.d}{\sim} \mathbb{P}$ with $EX_i = \mu$ and $VarX_i = \sigma^2$ for $i = 1, \dots, n$ and

$$Y = \sum_{i=1}^{n} X_{i}$$

and

$$\bar{X}_n = \sum_{i=1}^n \frac{X_i}{n} = \frac{1}{n} \sum_{i=1}^n X_i$$

▶ If we know \mathbb{P} , we know everything about Y and \bar{X}_n .

The inference problem (in terms of μ)

- ightharpoonup Statistical inference describes any procedure for extracting information about a probability distribution (population) from an observed sample. Let's focus, for example, in the population mean, μ
- We want to determine whether or not a claim about μ is correct or estimate as accurately as possible the value of μ

Types of statistical inference

- ▶ Point estimation
- Hypothesis testing
- ► Set estimation

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- ► I would like to know what time the entire population of STAT-S 520 students go to sleep, on average. I randomly select a sample of students and find the sample average of sleeping times.

Point Estimation: Terminology

- Estimand: The value we are trying to estimate
- Estimator: The method or process used to estimate a parameter
- Estimate: The observed value obtained using the method

Point Estimation

Point estimation is a method used to obtained a single value as estimator or estimate

Properties of Estimators

► **Unbiasedness** An estimator is unbiased if its expected value is equal to the estimand

Consistency An estimator is consistent if it converges in probability to the estimand

► The plug-in estimate for the variance is

$$\hat{\sigma}^2 = \sum_{i=1}^n \frac{(x_i - \bar{x}_n)^2}{n}$$

The plug-in estimator is biased:

$$E\left|\frac{1}{n}\sum_{i=1}^{n}(X_i-\bar{X}_n)^2\right|=\frac{n-1}{n}\sigma^2$$

▶ The sample variance is the estimate

$$s_n^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_n)^2$$

with corresponding estimator

$$S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

Example 2 Simulation

#> [1] 4.991494

```
mu = 0; sigma2 = 25; sigma = sqrt(sigma2); n = 60
plug.sd = function(x) sqrt(1/n*sum((x - mean(x))^2))
set.seed(1003)
vec_plug_sd =replicate(10^4, plug.sd(rnorm(n, mu, sigma)))
mean(vec plug sd)
#> [1] 4.949723
set.seed(1003)
vec_sample_sd =replicate(10^4, sd(rnorm(n, mu, sigma)))
mean(vec sample sd)
```

Hypothesis Testing

- A claim about μ is made (null hypothesis) about the mean of the population
- A random sample from this population is obtained and used to obtain the sample mean \bar{x}_n
- Assuming the claim about μ is correct, we determine how likely is to observe a value as extreme as or more extreme than \bar{x}_n
 - If, getting \bar{x}_n is very unlikely, we reject the claim (reject the null hypothesis)

Hypotheses

- The null hypothesis, H_0 , is what we initially assume to be true. It's a statement about μ .
 - For procedural reasons, we **always** include the equal sign in the statement under H_0
- The alternative hypothesis, H_1 , is what we would conclude (about μ) if we were to reject H_0 .

Test Statistic

The statistic is the method we use to assess the claim made under the null hypothesis. Many statistics exist, the one we encounter often looks like this:

$$\textit{statistic} = \frac{\textit{estimator} - \textit{parameter under } \textit{H}_0}{\textit{standard error}}$$

Hypothesis about μ when σ is known

When the hypothesis is made about μ , and σ is known, the test statistic is given by

$$Z = \frac{\bar{X}_n - \mu_0}{\sigma / \sqrt{n}}$$

and due to the CLT, $Z \sim N(0,1)$.

Hypothesis about μ when σ is unknown

When σ is unknown (as in most real-life problems), the test statistic is given by

$$T = \frac{\bar{X}_n - \mu_0}{S_n / \sqrt{n}}$$

where $S_n = \sqrt{\frac{\sum (X_i - \bar{X}_n)^2}{n-1}}$ is the sample standard deviation, another estimator. Given the added uncertainty of S_n , T is no longer normal, but T follows a T-distribution with n-1 degrees of freedom and we write $T \sim T_{n-1}$

Observed test statistic

Once you collect a (random) sample of n observations, the sample observed is $\overrightarrow{x} = (x_1, \dots, x_n)$. If σ is known, the observed test statistic is

$$z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}$$

and if σ is unknown, the observed test statistic is

$$t = \frac{\bar{x} - \mu_0}{s / \sqrt{n}}$$

where $s = \sqrt{\frac{\sum (x_i - \bar{x})^2}{n-1}}$ is the sample standard deviation.

Significance Probability (*p*-value)

- ▶ The p-value is the probability of observing a test statistic that is as extreme or more extreme than the one observed in our data, when we assume that H_0 is true.
- ▶ The p-value depends on the hypotheses statements made

Conclusion and Interpretation of results

- We want to reject H_0 if the observed estimate is highly unlikely to have been obtained by chance
 - ▶ The smaller the *p*-value, the more evidence to reject H_0 .
- While how small p-value needs to be is somewhat arbitrary, it should be guided by how important is not to make a mistake by rejecting H_0 when it is actually true (Type I Error) or by failing to reject H_0 when it is actually false (Type II Error).

Reasonable ranges for the *p*-value

To guide you in the decision process, some reasonable values (although still arbitrary) can be:

- ▶ If p-value > 0.1, do not reject H_0 .
- ▶ If p-value < 0.001, reject H_0 .
- ▶ If $0.001 \le p$ -value ≤ 0.1 , decide based on your own perception of this uncertainty (different people may make different decisions).
- ▶ Alternatively, come up with a significance level, α , such that if p-value $\leq \alpha$, we reject H_0 , and if p-value $> \alpha$, we fail to reject H_0 .
 - The choice of α should be set before collecting and/or observing the random sample.