Probabilistic ML IV

CS771: Introduction to Machine Learning

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Recap of Last Lecture

Some nice continuous distributions

Uniform distribution – support is within an internal – no partiality within that!

Gaussian distribution – support is entire \mathbb{R} – concentrates around the mean.

Concentrates strongly if variance small, weakly if variance large

Sum of two independent Gaussians is Gaussian, scaled Gaussian is a Gaussian

Tail Rule for Gaussians – deviation of more than $5\sigma < 0.00004$

Laplacian distribution – cousin of Gaussian – concentrates mode strongly

Probabilistic Regression

MLE with Gaussian Likelihood gives us least squares loss function

MLE with Laplacian Likelihood gives us absolute loss function

Probabilistic Regularization via Priors

Posterior, Maximum a Posteriori Estimator (MAP)

MAP can give us regularized or even constrained optimization problems

Be careful not to have strong priors (uninformed strong opinions are bad in life too (3)) to ML

Random vectors can be thought of as simply a collection of random variables arranged in an array $\mathbf{X} = [X_1, X_2, \dots, X_d]^\mathsf{T}$

No restriction on the random variables being independent or uncorrelated

PMF/PDF of **X** is simply the joint PMF/PDF of $\{X_1, X_2, ..., X_d\}$

Can talk about marginal/conditional prob among $X_1, ... X_d$ Think of $X_1, X_2, ..., X_d$ as just a bunch of r.v.s $\mathbb{P}[X_2, X_3 \mid X_1, X_4, X_5]$

Since PMF/PDF of **X** is simply a joint PMF/PDF, all probability laws we learnt earlier continue to hold if we apply them correctly

Chain Rule, Sum Rule, Product Rule, Bayes Rule Conditional/marginal variants of all these rules

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Expectation of a random variable is simply another vector (of same dim) of the expectations of the individual random variables

$$\mathbb{E}\mathbf{X} = [\mathbb{E}X_1, \mathbb{E}X_2, \dots \mathbb{E}X_d]^{\mathsf{T}}$$

Linearity of expectation continues to hold: if X, Y any two vector r.v. (not necessarily independent, then $\mathbb{E}[X + Y] = \mathbb{E}X + \mathbb{E}Y$

Scaling Rule: If $c \in \mathbb{R}$ is a constant then $\mathbb{E}[c \cdot \mathbf{X}] = c \cdot \mathbb{E}\mathbf{X}$

Dot Product Rule: If $\mathbf{a} \in \mathbb{R}^d$ is a constant vector, then $\mathbb{E}[\mathbf{a}^\mathsf{T}\mathbf{X}] = \mathbf{a}^\mathsf{T}\mathbb{E}\mathbf{X}$

Proof:
$$\mathbb{E}[\mathbf{a}^{\mathsf{T}}\mathbf{X}] = \mathbb{E}[\sum_{i=1}^{d} a_i X_i] = \sum_{i=1}^{d} \mathbb{E}[a_i X_i] = \sum_{i=1}^{d} a_i \cdot \mathbb{E}[X_i] = \mathbf{a}^{\mathsf{T}} \mathbb{E}\mathbf{X}$$

Matrix Product Rule: If $A \in \mathbb{R}^{n \times d}$ is a constant matrix then $\mathbb{E}[A\mathbf{X}] = A\mathbb{E}\mathbf{X}$

Proof: Use Dot Product Rule n times



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Mode easy to define: \underset{X_1,...,X_d}{\operatorname{max}} \mathbb{P}[X_1,...,X_d]
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Median not easy to define – no unique definition

Definition 1: $med(\mathbf{X}) = [med(X_1), med(X_2), ..., med(X_d)]^{\mathsf{T}}$

Definition 2: minimizer of absolute distance (in this case L1 norm)

$$med(\mathbf{X}) = arg \min_{\mathbf{v} \in \mathbb{R}^d} \mathbb{E}[\|\mathbf{X} - \mathbf{v}\|_2]$$

Note: even here we still have $\mathbb{E}[\mathbf{X}] = \arg\min_{\mathbf{v} \in \mathbb{R}^d} \mathbb{E}[\|\mathbf{X} - \mathbf{v}\|_2^2]$

Proof: $\mathbb{E}[\|\mathbf{X} - \mathbf{v}\|_2^2] = \mathbb{E}[\|\mathbf{X}\|_2^2] + \mathbb{E}[\|\mathbf{v}\|_2^2] - 2 \cdot \mathbf{v}^{\mathsf{T}} \mathbb{E}[\mathbf{X}]$

Taking derivative w.r.t ${f v}$ and using first order optimality does the trick



Since random vectors are a bunch of real valued r.v.s, to specify the variance of this collection, need to have all pairwise covariances

Covariance

$$\operatorname{Cov}(\mathbf{X}) = \begin{bmatrix} \mathbb{V}X_1 & \operatorname{Cov}(X_1, X_2) & \dots & \operatorname{Cov}(X_1, X_d) \\ \operatorname{Cov}(X_2, X_1) & \mathbb{V}X_1 & \dots & \operatorname{Cov}(X_2, X_d) \\ \vdots & \vdots & \ddots & \vdots \\ \operatorname{Cov}(X_d, X_1) & \operatorname{Cov}(X_d, X_2) & \dots & \mathbb{V}X_d \end{bmatrix}$$

Another cute formula

$$Cov(\mathbf{X}) = \mathbb{E}[(\mathbf{X} - \mathbf{\mu})(\mathbf{X} - \mathbf{\mu})^{\mathsf{T}}] = \mathbb{E}[\mathbf{X}\mathbf{X}^{\mathsf{T}}] - \mathbf{\mu}\mathbf{\mu}^{\mathsf{T}}$$
, where $\mathbf{\mu} = \mathbb{E}\mathbf{X}$
 $Cov(c \cdot \mathbf{X}) = c^2 \cdot Cov(\mathbf{X})$

Since random vectors are a bunch of variance of this collection, need to have all pairwise covariances

If \mathbf{X} is a vector, isn't $\mathbf{X}\mathbf{X}^{\mathsf{T}}$ a matrix? What does $\mathbb{E}[\mathbf{X}\mathbf{X}^{\mathsf{T}}]$ even mean?

Covariance

Just as a random vector is a collection of random variables arranged as a 1D array, a random matrix $Cov(X_2, X_1)$ is a collection of r.v.s arranged as a 2D array!

Note that the (i,i)-th entry of the matrix $(\mathbf{X} - \mathbf{\mu})(\mathbf{X} - \mathbf{\mu})^{\mathsf{T}}$ is simply Another CI $\mathbb{E}[(X_i - \mu_i)(X_j - \mu_j)]$. Thus, (i,j)-th entry of $\mathbb{E}[(X - \mu)(X - \mu)^T]$ is $\mathbb{E}[(X_i - \mu_i)(X_j - \mu_j)] = \text{Cov}(X_i, X_j)$

$$Cov(X) = \mathbb{E}[(X - \mu)(X - \mu)'] = \mathbb{E}[XX'] - \mu\mu'$$
, where μ

$$Cov(c \cdot \mathbf{X}) = c^2 \cdot Cov(\mathbf{X})$$



Useful Operations on Vector R.V.

If $X \in \mathbb{R}^m$, $Y \in \mathbb{R}^n$ are two random vectors (not necessarily independent), then

$$\begin{aligned} &\text{Cov}(\textbf{X},\textbf{Y}) = \mathbb{E}[(\textbf{X} - \boldsymbol{\mu}_{\textbf{X}})(\textbf{Y} - \boldsymbol{\mu}_{\textbf{Y}})^{\top}] = \mathbb{E}[\textbf{X}\textbf{Y}^{\top}] - \boldsymbol{\mu}_{\textbf{X}}\boldsymbol{\mu}_{\textbf{Y}}^{\top} \in \mathbb{R}^{m \times n} \\ &\text{where } \boldsymbol{\mu}_{\textbf{X}} = \mathbb{E}\textbf{X} \text{ and } \boldsymbol{\mu}_{\textbf{Y}} = \mathbb{E}\textbf{Y}, \text{Cov}(\textbf{X},\textbf{Y}) \end{aligned}$$

Dot Product Rule: If $\mathbf{a} \in \mathbb{R}^d$ is a constant vector, then $\mathbb{V}[\mathbf{a}^\mathsf{T}\mathbf{X}] = \mathbf{a}^\mathsf{T}\mathrm{Cov}[\mathbf{X}]\mathbf{a}$

$$\begin{aligned} \textit{Proof:} \, \mathbb{V}[a^{\top}X] &= \mathbb{E}[(a^{\top}X)^2] - (a^{\top}\mu_X)^2 = \mathbb{E}[a^{\top}XX^{\top}a] - a^{\top}\mu_X\mu_X^{\top}a \\ &= a^{\top}\mathbb{E}[XX^{\top}]a - a^{\top}\mu_X\mu_X^{\top}a = a^{\top}\big(\mathbb{E}[XX^{\top}] - \mu_X\mu_X^{\top}\big)a = a^{\top}\text{Cov}[X]a \end{aligned}$$

Matrix Product Rule: If $A \in \mathbb{R}^{n \times d}$ is a constant matrix then $Cov[A\mathbf{X}] = ACov[\mathbf{X}]A^{\mathsf{T}} \in \mathbb{R}^{n \times n}$

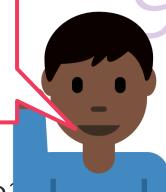
Proof: Try arguing similarly as the dot product rule



Useful Operations on

If $\mathbf{X} \in \mathbb{R}^m$, $\mathbf{Y} \in \mathbb{R}^n$ are two random independent), then

Can you prove that the covariance matrix of any random vector is always a PSD matrix?



$$Cov(\mathbf{X}, \mathbf{Y}) = \mathbb{E}[(\mathbf{X} - \mathbf{\mu}_{\mathbf{X}})(\mathbf{Y} - \mathbf{\mu}_{\mathbf{Y}})^{\mathsf{T}}] = \mathbb{E}[\mathbf{X}\mathbf{Y}^{\mathsf{T}}] - \mathbf{\mu}_{\mathbf{X}}\mathbf{\mu}_{\mathbf{Y}}^{\mathsf{T}} \in \mathbb{R}^{\mathsf{T}}$$

where $\mu_X = \mathbb{E} X$ and $\mu_Y = \mathbb{E} Y$, Cov(X, Y)

Dot Product Rule: If $\mathbf{a} \in \mathbb{R}^d$ is a constant vector, then $\mathbb{V}[\mathbf{a}^\mathsf{T}\mathbf{X}] = \mathbf{a}^\mathsf{T}\mathrm{Cov}[\mathbf{X}]\mathbf{a}$

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Matrix Product Rule: If $A \in \mathbb{R}^{n \times d}$ is a constant matrix then $Cov[A\mathbf{X}] = ACov[\mathbf{X}]A^{\mathsf{T}} \in \mathbb{R}^{n \times n}$

Proof: Try arguing similarly as the dot product rule



Gaussian Random Vector

As in the scalar case, the *multivariate* Gaussian requires just the mean $\mu \in \mathbb{R}^d$ and the covariance $\Sigma \in \mathbb{R}^{d \times d}$ to be specified $\mathcal{N}(\mu, \Sigma)$

$$\mathbb{P}[\mathbf{x} \mid \boldsymbol{\mu}, \boldsymbol{\Sigma}] = \frac{1}{\sqrt{(2\pi)^d |\boldsymbol{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

Special case $\mathbf{\mu}=\mathbf{0}$ and $\mathbf{\Sigma}=I_d$ called *standard Gaussian/Normal dist*

$$\mathbb{P}[\mathbf{x} \mid \mathbf{0}, I_d] = \frac{1}{\sqrt{(2\pi)^d}} \exp\left(-\frac{1}{2} \|\mathbf{x}\|_2^2\right) = \prod_{i=1}^d \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} x_i^2\right)$$

However, $\frac{1}{\sqrt{2\pi}}\exp\left(-\frac{1}{2}x_i^2\right)$ is simply $\mathcal{N}(0,1)$ i.e. we indeed have

$$\mathbb{P}[x_1, ..., x_d \mid \mathbf{0}, I] = \prod_{i=1}^d \mathbb{P}[x_i \mid 0, 1]$$

All d coordinates of a standard Gaussian r.vec. are independent!