

Delayed dichromatism as a convenient tool to disentangle the effects of survival and productivity on the population dynamics in waterfowl

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Abstract

Monitoring the number of individuals is by far the most popular strategy for studying the environmental factors that determine population dynamics and for measuring the effectiveness of management actions aimed at population recovery, control, or eradication. Unfortunately, monitoring numbers is inefficient in identifying the mechanisms underlying demographic processes and, in particular, in assessing the extent to which population growth rate is influenced by changes in adult survival rather than variations in reproductive parameters. The usual method for determining adult survival and productivity relies on capturing and marking, which has three major drawbacks. Firstly, the selectivity of capture methods can lead to a distorted representation of the underlying population structure. Secondly, the capture and release of animals is quite invasive, causing stress and posing an additional threat to the often already endangered species. Thirdly, release may simply be impossible, especially for invasive alien species that represent a threat to native taxa. In many waterfowl species, sexual dichromatism is observed in adults, while immatures of both sexes display a plumage pattern similar to that of adult females. Using two non-native populations of Ruddy duck, a species that exhibits this plumage pattern, we present a non-invasive method to estimate the respective contribution of adult survival and productivity (also called recruitment rate) to the population growth rate. We exemplify how changes in the apparent proportion of males in counts (due to female-like immature males gradually acquiring adult male plumage) can be exploited to split the population growth rate into adult survival and recruitment rates, which contributes significantly to understanding the demographic impact of two different eradication strategies.

Keywords: age ratio - alien species - demography - fecundity - harvesting - intrinsic growth rate - *Oxyura jamaicensis* - reproductive success - sex ratio - vital rate

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1 Introduction

Assessing the growth rate of populations is a first step towards a better understanding of the factors underlying the dynamics of natural populations (e.g. Niel & Lebreton, 2005). It is also crucial for measuring the effectiveness of management actions that are taken to restore, control, or eradicate populations (Shea & NCEAS Working Group on Population Management, 1998). Among the approaches available to managers to achieve these goals, those that rely on monitoring the number of individuals (counts) are by far the most popular (Rintala et al., 2022). In many cases, these methods allow the investigation of environmental factors underlying changes in population size, and thus help in the implementation of mitigating actions (Faillettaz et al., 2019). Unfortunately, relying solely on the monitoring of population abundance generally hinders the fundamental understanding of the demographic mechanisms underlying changes in population growth rates. More specifically, on the basis of counts alone, it is almost impossible to assess the extent to which population growth rate is affected by changes in adult survival rather than to variations in young production (Austin et al., 2000). To assess the relative sensitivity of population growth rate to factors affecting adult survival or reproductive parameters, demographers usually rely on the monitoring of individuals by capture-mark-recapture (Lebreton et al., 1992). In practise, however, capture and release of individuals cannot always be relied upon, as the legal status of the species precludes any additional disturbance (e.g. critically endangered species) or prohibits the release of live individuals (e.g. invasive species or pests). Even though capture-mark-recapture methods are efficient for assessing demographic parameters, they have potential drawbacks, such as their invasiveness, which can affect behaviour and thus survival or reproductive success, and their affordability when time and money are scarce. Genetic monitoring is a non-invasive alternative to capture-mark-recapture, but it is costly and requires large samples (Caniglia et al., 2011).

Most of the time, therefore, managers make the best of a bad job by using counts as a viable option for tracking changes in population size and then population growth rate to assess the relevance of management actions. Distinguishing between immatures and adults in counts allows estimation of adult survival rate, i.e. the proportion of breeders that have survived for one year, and productivity/recruitment rate, i.e. the number of immatures produced per breeder that reach sexual maturity, which are, by definition, recruited in

27 the breeding population (Smith et al., 2001). This additional effort makes it possible to assess the relative
28 influence of each of these components on the population growth rate. Alternatively, the assessment of age
29 structure in hunting bags has been used to infer the role of decreasing reproductive success in population
30 declines in a number of game species, including ducks and geese, but suffers from intractable biases (Fox &
31 Cristensen, 2018). Of course, the latter approach is not appropriate for protected/endangered species.

32 Here we exemplify how, in dimorphic species, delayed sexual maturity of males can be used to estimate adult
33 survival and recruitment rates by distinguishing male-like and female-like individuals in winter counts. We
34 used two non-native European populations of Ruddy duck (*Oxyura jamaicensis*) as a study model. Like
35 many other duck species, Ruddy duck is dimorphic: newborn males look like females until the prenuptial
36 moult, which takes place in late winter. As a result of the late prenuptial moult of immatures, the apparent
37 proportions of males increases over the course of the wintering season. These changes in the apparent
38 proportions of males during this period are therefore directly related to the proportion of immatures in the
39 populations and thus to the reproductive success of the previous breeding season. The “apparent sex ratio”
40 method developed in this study was used to estimate adult survival and recruitment rates and evaluate the
41 effects of two different eradication strategies used in Great Britain and France, respectively.

42 2 Materials & methods

43 The Ruddy duck is a stiff-tailed duck native to the American continent. Starting from seven individuals
44 acclimatised at the Slimbridge Wetland Centre in the 1940s (Gutiérrez-Expósito et al., 2020), a feral popu-
45 lation began to establish in Great Britain and the first attempts at reproduction in the wild were observed
46 in the 1960s (Figure 1). This feral population grew rapidly and spread across the country until it reached
47 more than 5,000 individuals in the early 2000s. In the late 1980s, a feral population also began to establish
48 on the continent, particularly in France, supposedly due to the arrival of British-born individuals. However,
49 in contrast to observations in Great Britain, Ruddy ducks have not spread very much in France. Most
50 observations and breeding attempts have been concentrated in the north-west of the country. In France,
51 almost no Ruddy ducks were observed in winter outside the lake of Grand Lieu (47.09°N, 1.67°W), which

52 greatly facilitated the monitoring of this population. Ruddy ducks are considered a major threat to the
 53 endangered native White-headed duck (*Oxyura leucocephala*) in the south-western Mediterranean, as they
 54 hybridise and thus pose an increased risk of genetic pollution and genetic assimilation of the latter by the
 55 former (Muñoz-Fuentes et al., 2007). To reduce the risk of genetic pollution of the White-headed duck by
 56 the Ruddy duck, eradication measures were implemented in both Great Britain and France in the late 1990s
 57 (Gutiérrez-Expósito et al., 2020), followed by a European Ruddy duck eradication plan in 1999 (Hughes et
 58 al., 1999) (Figure 1).

59 The trends and status of the White-headed duck and Ruddy duck populations used to be determined ex-
 60clusively by counts. It was therefore difficult to assess the factors impairing the recovery of the former and
 61 the effectiveness of the eradication programme for the latter. In particular, counts alone cannot be used to
 62 assess the relative impact of changes in adult survival and recruitment rates on population growth rates,
 63 which is a prerequisite for identifying limiting factors.

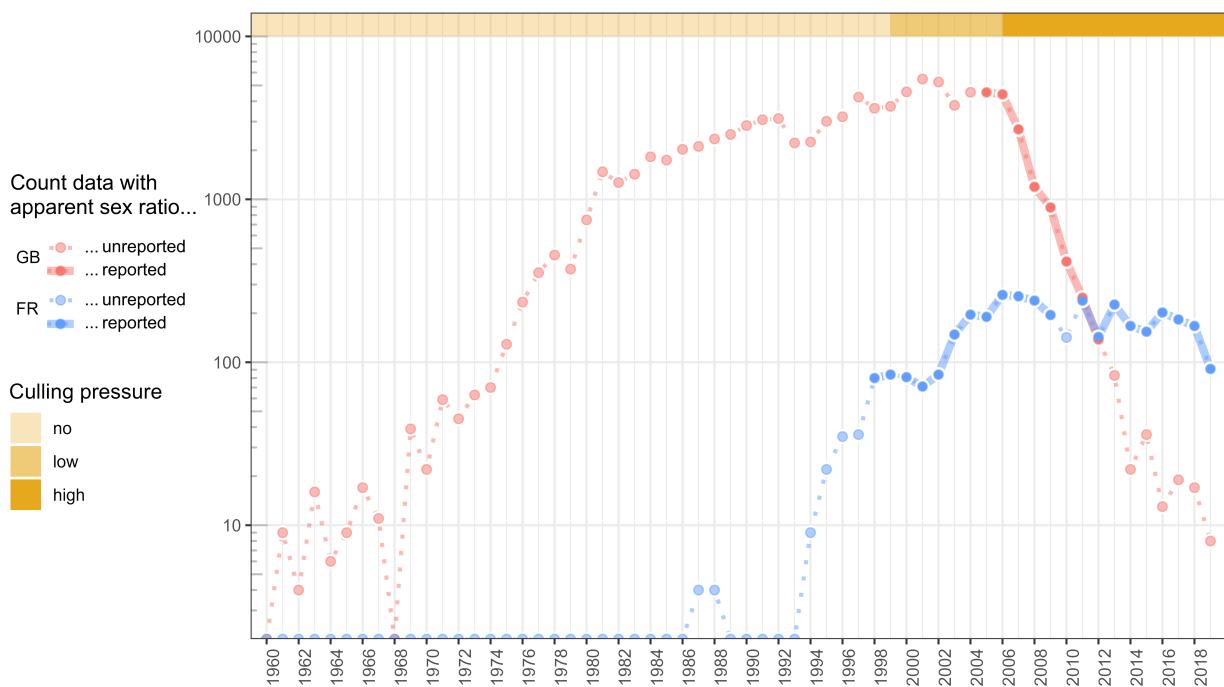


Figure 1: Temporal changes in the number of Ruddy ducks (logarithmic scale) in Great Britain (GB) and France (FR) from 1960 to 2019, with references to periods when apparent sex ratios were reported

64 Like many other duck species, both White-headed duck and Ruddy duck exhibit delayed dichromatism,

65 i.e. young males acquire the colourful plumage typical of their species and can therefore be distinguished from
66 females during the pre-breeding period, typically from mid-winter in the earliest individuals (Baldassarre,
67 2014). Delayed dichromatism generally explains the discrepancies in estimates of apparent proportions
68 of males counted in winter compared to proportions of males counted in spring or males identified from
69 individuals culled in winter (Figures 2 & 3).

70 Interestingly, delayed dichromatism allows the estimation of adult survival and recruitment/productivity by
71 monitoring the seasonal evolution of the apparent sex ratio and making few and reasonable assumptions (see
72 next section).



Figure 2: A flock of Ruddy ducks observed in winter, including ten female-like individuals with whitish striped cheek (some of which are immature males), four male-like individuals with white cheek and black cap (all adult males), three unidentified individuals (1st, 4th, and 10th from left) © Jay McGowan - 3 February 2013 - Tompkins, New York, United States

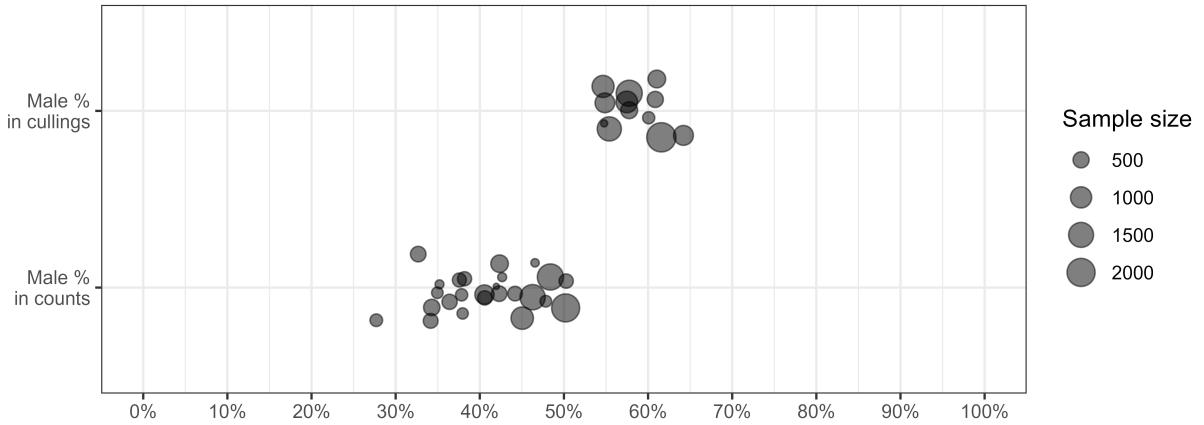


Figure 3: Proportions of males estimated by internal examination of individuals culled in a same year under the eradication programmes (Male % in cullings), and proportions of individuals showing male-like plumage counted in winter (Male % in counts); discrepancies between the two estimates are due to delayed dichromatism (immature males looking like females before moulting); we used these discrepancies to estimate survival and productivity/recruitment

73 We monitored the demography of Ruddy duck populations in both Great Britain (GB) and France (FR) by
 74 conducting extensive counts on the wintering grounds between 1 December and 31 January. In addition,
 75 the apparent sex ratio could be determined using winter counts that distinguished between female-like and
 76 male-like individuals. Such counts were carried out in Great Britain from 2006 to 2012, which corresponded
 77 to a period of sharp population decline (Figure 1). In France, count surveys taking into account apparent
 78 sex were conducted in 1999, 2001-2009, and 2012-2019. The French population grew rapidly in the first years
 79 of monitoring and then stabilised from 2006 onwards as a result of high culling pressure. In both countries,
 80 culling was carried out under strict official control so that the exact number of individuals culled was known.
 81 Although the age ratio of birds culled before 2009 could not be determined with certainty in France, it was
 82 still possible to roughly categorise the culling pressure in both countries into three categories: “no culling”
 83 before 1999, “low culling” between 1999 and 2005 and “high culling” from 2006 onwards (see Figure 8 in
 84 Appendix section).

85 **2.1 Inferring adult survival and recruitment rates from changes in population
86 size and proportion of immatures**

87 Our method exploits the difference between the sex ratio estimated from counts before young males acquire
88 their male-like plumage pattern (and are therefore mistaken for females) and the sex ratio among adults
89 (which is stable over time), which can be estimated from either spring or summer counts or from hunting bags.
90 In the present study, we estimated the adult sex-ratio from culled individuals using the presence/absence of
91 the bursa of Fabricius as a criterion for distinguishing adults from immatures (see hereafter). In France, it
92 was not possible to obtain sufficiently reliable estimates of the sex ratio due to the small sample size. In
93 Great Britain, adult sex ratios did not differ from month to month, so data from culled adults collected
94 throughout the whole biological cycle were pooled together each year. The proportion of males also did not
95 differ statistically between years in which more than 500 adults were culled. This is consistent with previous
96 findings showing that the adult sex ratio in ducks is usually very stable in the short term (Wood et al.,
97 2021) but can fluctuate in the long term. Therefore, we pooled the data from all adults culled in the control
98 programme to estimate the adult sex ratio, see Equation (1).

99 Assuming the additive property of the binomial distribution, the proportion of immatures in a population can
100 be derived from the proportion of males among adults (alternatively the adult sex ratio) and the proportion
101 of male-like individuals in the different counts in winter (assuming that all these individuals are adult males),
102 see Equation (2). By combining the population growth rate derived from the interannual variation of an
103 index of abundance with the proportion of immatures, adult survival and recruitment rates can be derived,
104 see Equation (3) and Figure 4. If the absolute value of the population size is known, the number of adults
105 and the number of recruits can be estimated, see Equation (4). We called this approach the “apparent sex
106 ratio” method.

Variable name	Description
<i>Observation (input)</i>	
AM	Total number of adult males culled as part of the eradication programme
AF	Total number of adult females culled as part of the eradication programme
$CML_{i,t}$	Number of male-like individuals counted in population i in year t
$C_{i,t}$	Number of individuals counted in population i in year t
$N_{i,t}$	Size of the population i in year t (maximum number of individuals counted in the wintering population)

Variable name	Description
<i>Parameter (output)</i>	
$p(m a)$	Proportion of males among adults (or probability of being a male knowing it is an adult)
$p(a \cap m)_{i,t}$	Proportion of adult males in population i in year t
$p(a)_{i,t}$	Proportion of adults in population i in year t
$p(re)_{i,t}$	Proportion of immatures/recruits in population i in year t
$s_{i,t}$	Adult survival rate, i.e. proportion of individuals in year $t - 1$ still alive in year t
$r_{i,t}$	Recruitment rate, i.e. number of recruits in population i in year t per adult in year $t - 1$
$S_{i,t}$	Number of adults in population i in year t
$R_{i,t}$	Number of recruits in population i in year t

$$p(m|a) \sim \text{Beta}(AM, AF) \quad (1)$$

$$\begin{aligned} CML_{i,t} &\sim \text{Binom}(p(a \cap m)_{i,t}, C_{i,t}) \\ &\sim \text{Binom}(p(m|a).p(a)_{i,t}, C_{i,t}) \\ &\sim \text{Binom}(p(m|a).(1 - p(re)_{i,t}), C_{i,t}) \end{aligned} \quad (2)$$

$$\begin{aligned} s_{i,t} &= \frac{N_{i,t}}{N_{i,t-1}} \cdot (1 - p(re)_{i,t}) \\ r_{i,t} &= \frac{N_{i,t}}{N_{i,t-1}} \cdot p(re)_{i,t} \end{aligned} \quad (3)$$

$$\begin{aligned} S_{i,t} &= N_{i,t} \cdot (1 - p(re)_{i,t}) \\ R_{i,t} &= N_{i,t} \cdot p(re)_{i,t} \end{aligned} \quad (4)$$

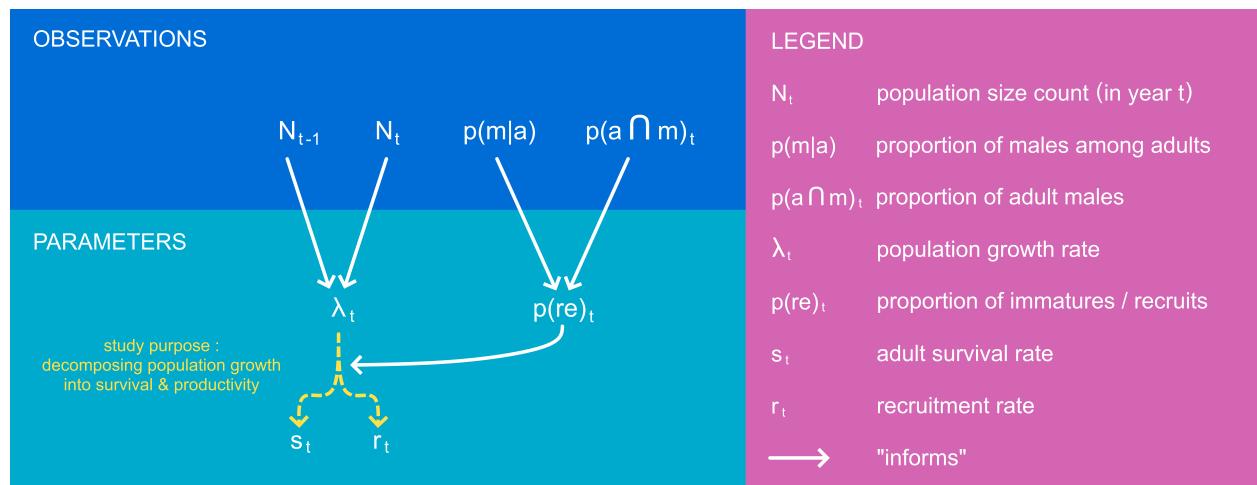


Figure 4: Overview of the “apparent sex ratio” method used to disentangle the population growth into survival (adult survival rate) and productivity (recruitment rate)

¹⁰⁷ **2.2 Validating the “apparent sex ratio” method**

¹⁰⁸ The relevance of the method was assessed by testing the likelihood of the estimates of both vital rates. For
¹⁰⁹ the adult survival rate, we checked whether the values were contained in the interval [0; 1], and compared
¹¹⁰ our estimates with those in the literature Nichols et al. (1997). Validation of the recruitment rate was
¹¹¹ more difficult because if the lower limit is 0, the upper limit is a product of all maximum values of its
¹¹² individual components, i.e. nesting rate, nesting success, clutch-size, hatching success, pre-fledging survival,
¹¹³ and fledgling survival. Although some of these reproductive parameters may be known for some duck species,
¹¹⁴ the last one is never mentioned (e.g. Baldassarre, 2014).

¹¹⁵ To overcome this problem, we developed an indirect approach to estimate the maximum expected recruitment
¹¹⁶ rate without exploitation and assuming that the level of population density does not affect the components
¹¹⁷ of the recruitment process. Estimated values above this maximum recruitment rate would reflect a partial
¹¹⁸ detection of adult males. On the other hand, null or negative recruitment rates would indicate that immature
¹¹⁹ males acquired male plumage prior to the survey, or that the assumption about the stability of the adult sex
¹²⁰ ratio was violated.

¹²¹ In practise, the maximum recruitment rate was estimated from the difference between the maximum growth
¹²² rate and the maximum survival rate of adults, see Equation (5). This relationship resulted from a simple
¹²³ consideration for a closed population: the population size in year t is equal to the number of adults that
¹²⁴ survived the whole year $t - 1$ plus the offspring produced in year $t - 1$ that survived to the reproductive
¹²⁵ period of year t , i.e. the recruitment in year t (Flint, 2015). In an open population, adult survival and
¹²⁶ recruitment rates are confounded with adult and recruit migrations, respectively, but this does not change
¹²⁷ the equation. This relationship becomes more complex when a species with delayed maturity is considered,
¹²⁸ see Robertson (2008).

Variable name	Description
N_t	Number of adults in year t
D_t	Number of adults dead during year t
R_t	Number of recruits in year t
s_t	Adult survival rate, i.e. proportion of individuals in year $t - 1$ still alive in year t
r_t	Recruitment rate, i.e. number of recruits in year t produced per individual in year $t - 1$
λ_t	Growth rate of the population between year $t - 1$ and year t

$$\begin{aligned}
N_t &= N_{t-1} - D_{t-1} + R_t \\
N_t &= N_{t-1} - (1 - s_t).N_{t-1} + r_t.N_{t-1} \\
N_t &= s_t.N_{t-1} + r_t.N_{t-1} \\
\frac{N_t}{N_{t-1}} &= s_t + r_t \\
\lambda_t &= s_t + r_t \\
r_t &= \lambda_t - s_t
\end{aligned} \tag{5}$$

129 The maximum adult survival rate was defined as a uniform distribution on the interval [0.7; 1]. The lower
 130 limit corresponded to the upper range of survival rates observed in waterfowl species of similar weight, and
 131 the upper limit was set to 1, as survival rates of long-lived waterfowl species can be very high (Buxton et
 132 al., 2004; Koons et al., 2014; Krementz et al., 1997; Nichols et al., 1992, 1997).
 133 For both populations, the maximum growth rates occurred during their geographic expansion phase, which
 134 preceded the start of the eradication programmes (Figure 1). To estimate robust maximum population
 135 growth rates for both populations, we smoothed the annual growth rates over a consistent time period by
 136 using a linear regression on the logarithm scale, see Equation (6). We discarded the data for Great Britain
 137 before 1972 as the estimates were likely noisy when the population was low (below 50 individuals) (Figure
 138 1). After reaching the threshold of 1,000 individuals, the population growth in Great Britain (GB) showed
 139 a strong inflection, although culling had not yet started (Figure 1). This observation suggests that beyond
 140 1,000 individuals, a negative density-dependent process probably took place and led us to consider only the
 141 first sequence of the time series to infer the maximum growth rate in Great Britain, i.e. 1972-1981. For the
 142 French (FR) population, the sequence without culling pressure covered the period 1994-1999.

Variable name	Class	Description
$N0_{i,J}$	Parameter	Intercept of the regression model
$\lambda_{i,J}$	Parameter	Average population growth rate over a restricted time interval J for a population i (in year^{-1})
$t \in J$	Index	Year index within the time interval J
$\sigma_{i,J}$	Parameter	Standard deviation of the regression model

$$\log(N_{i,t}) \sim \text{Norm}(N0_{i,J} + \log(\lambda_{i,J}).t, \sigma_{i,J}) \tag{6}$$

143 Estimating vital rates within an acceptable range, as defined above, would not prove that our modelling

144 approach correctly reflects the actual interannual variability. To determine this, we compared the proportions
 145 of immatures and the vital rates derived from counts with those derived from culling, i.e. from individuals
 146 culled as part of the eradication programme. As we restricted the dataset to years with more than 100
 147 individuals culled in winter for sake of precision, the analysis only covered five years of the GB population
 148 time series. The proportions of immatures were estimated from the culled individuals by checking the
 149 presence of the bursa of Fabricius, which is only present in immature individuals (Hochbaum, 1942); see
 150 Equation (7). Adult survival and recruitment rates were then derived from the proportion of immatures and
 151 the Equation (3).

Variable name	Class	Description
$SI_{i,t}$	Observation	Number of immatures culled in the wintering population i in year t
$S_{i,t}$	Observation	Number of individuals culled in the wintering population i in year t

$$SI_{i,t} \sim \text{Binom}(p(re)_{i,t}, S_{i,t}) \quad (7)$$

152 2.3 Assessing the effects of culling strategies

153 During the period when culling pressure was high (i.e. from 2006 onwards), eradication strategies in Great
 154 Britain and France differed. In Great Britain, most individuals were culled in winter (53.4% of adults
 155 were culled before 30 May), whereas in France, most individuals were culled during the breeding season
 156 (81.6% of adults were culled after 30 May). We investigated whether these strategies had different effects on
 157 the populations by comparing the resulting growth rates and the relative contributions of both vital rates
 158 to them (by comparing the average values during the period of high culling pressure with proxies for the
 159 maximum vital rates estimated when both populations reached their maximum growth, see previous section).
 160 In France, a LIFE project was implemented from 2019 to intensify the culling pressure, especially in winter
 161 and spring. We therefore excluded the FR time series from 2019 onwards in order to compare homogeneous
 162 culling strategies.

₁₆₃ **2.4 Statistical framework**

₁₆₄ We used the Bayesian framework to implement all submodels presented in Materials & methods section. The
₁₆₅ Bayesian framework is both straightforward and efficient in propagating error through the parameters. We
₁₆₆ used uninformative priors for all parameters. Since the maximum growth rate is a life history trait expected
₁₆₇ to be stable between populations of the same species, we used an uninformative hierarchical prior for this
₁₆₈ parameter. We generated three chains of length 500,000, with a thinning of 10 to avoid autocorrelation
₁₆₉ in the samples, and discarded the first 2,000 samples as burn-in. We checked the chain convergence using
₁₇₀ the Gelman and Rubin convergence diagnostic ($R < 1.1$, Gelman & Rubin (1992)). The models were fitted
₁₇₁ using NIMBLE (de Valpine et al., 2017) run from R (R Core Team, 2022). The values $\mathbf{X}[\mathbf{Y}; \mathbf{Z}]$ reported in
₁₇₂ Results section are the medians and the corresponding limits of the 95% confidence interval of the posterior
₁₇₃ distributions. We preferred the median to the mean because it is more robust to skewed distributions.

₁₇₄ **3 Results**

₁₇₅ **3.1 Estimating vital rates**

₁₇₆ The “apparent sex ratio” method successfully provided estimates and associated uncertainties for the pro-
₁₇₇ portion of immatures in both wintering populations. Adult survival and recruitment rates, i.e. the two
₁₇₈ components that determine the population growth rate, were also correctly estimated (Figure 5).

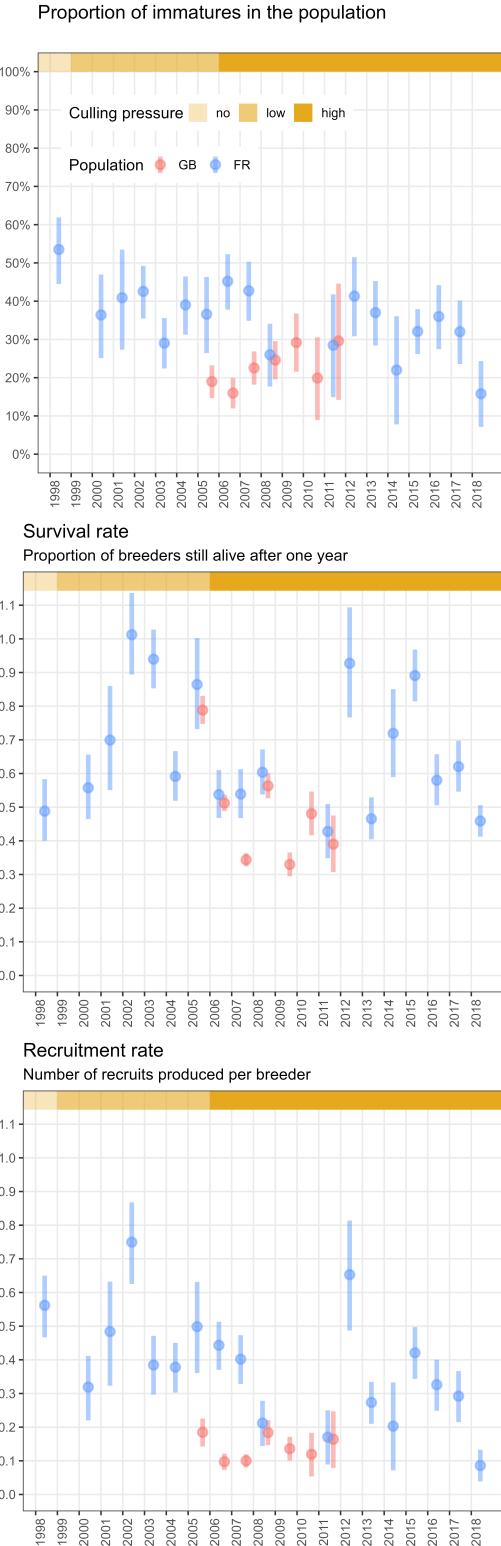


Figure 5: Temporal changes in the proportions of immatures and vital rates following culling pressure in the populations of Great Britain (GB) and France (FR); the vertical bars represent the 95% confidence intervals

179 The proportion of males among adults, which is a prerequisite for inferring the proportion of immatures,
180 was estimated at 0.60 [0.59; 0.61]. The proportion of immatures ranged between 0.16 [0.07; 0.24] and 0.54
181 [0.44; 0.62], depending on the population and year. The lowest values were similar in both populations, but
182 the range of the proportion of immatures in the GB population was much smaller (upper value GB: 0.30
183 [0.14; 0.45] & FR: 0.54 [0.44; 0.62]). The proportion of immatures in the GB population was stable over
184 time, while a slight but significant decrease was observed in the FR population.

185 Adult survival rates ranged between 0.33 [0.29; 0.37] and 1.01 [0.89; 1.14]. None of the estimates were
186 significantly outside the range of a survival rate defined without immigration [0; 1]. No trend in adult
187 survival rate was observed in either population, but the patterns were different: adult survival in GB was
188 lower and showed less interannual variability than in FR.

189 Recruitment rates ranged between 0.09 [0.04; 0.13] and 0.75 [0.63; 0.87]. All estimates were above 0, which
190 is not inconsistent with reality. Furthermore, the maximum recruitment rate was 0.68 [0.36; 0.78] (see 3rd
191 Results section), and no estimate was significantly outside the range of recruitment rates defined without
192 immigration [0; 0.78]. No trend was observed for the GB population, but the recruitment rate decreased for
193 the FR population, although it was more noisy than the proportion of immatures. As with the survival rate,
194 the GB population also showed lower recruitment rate values with less variability than the FR population.

195 Lower values and lower variability of both adult survival and recruitment rates estimated in the GB popu-
196 lation than in the FR population probably explain the different trends of the two Ruddy duck populations:
197 a sharp and constant decline in the GB population versus a slow and variable decline in the FR population
198 (Figure 1). In the FR population, the range of recruitment rates (min/max difference: 0.66) was larger than
199 the range of survival rates (min/max difference: 0.58). Conversely, the range of survival rates (min/max
200 difference: 0.46) in the GB population was much larger than the range of recruitment rates (min/max dif-
201 ference: 0.08). This suggests that the variability in population growth rate in FR was mainly determined
202 by recruitment, whereas in GB it was mainly determined by changes in adult survival.

203 **3.2 Testing the reliability of the estimates**

204 In Great Britain, the proportions of immatures derived from the counts were positively correlated with
205 those derived from the culled individuals (Figure 6). This result therefore supports the hypothesis that the
206 proportion of female-like individuals in the overwintering populations is efficient for the assessment of age
207 structure and thus productivity. The correlation was even stronger when looking at vital rates (Figure 6),
208 but this was expected because the two methods for estimating these vital rates had a common component,
209 namely the population growth rate (see Equation (3)). Interestingly, a 1:1 correlation was not achieved
210 for any of the parameters. The proportion of immatures derived from the counts is always lower than the
211 proportion derived from the culled individuals. The discrepancy between these two estimates is due to
212 the fact that immatures are probably more vulnerable to culling than adults and may therefore be over-
213 represented in the culling samples. As a result, the correlation coefficients for both vital rates also differed.
214 However, the differences between the two approaches do not call into question the ability of the “apparent
215 sex ratio” method to capture interannual variability of the vital rates.

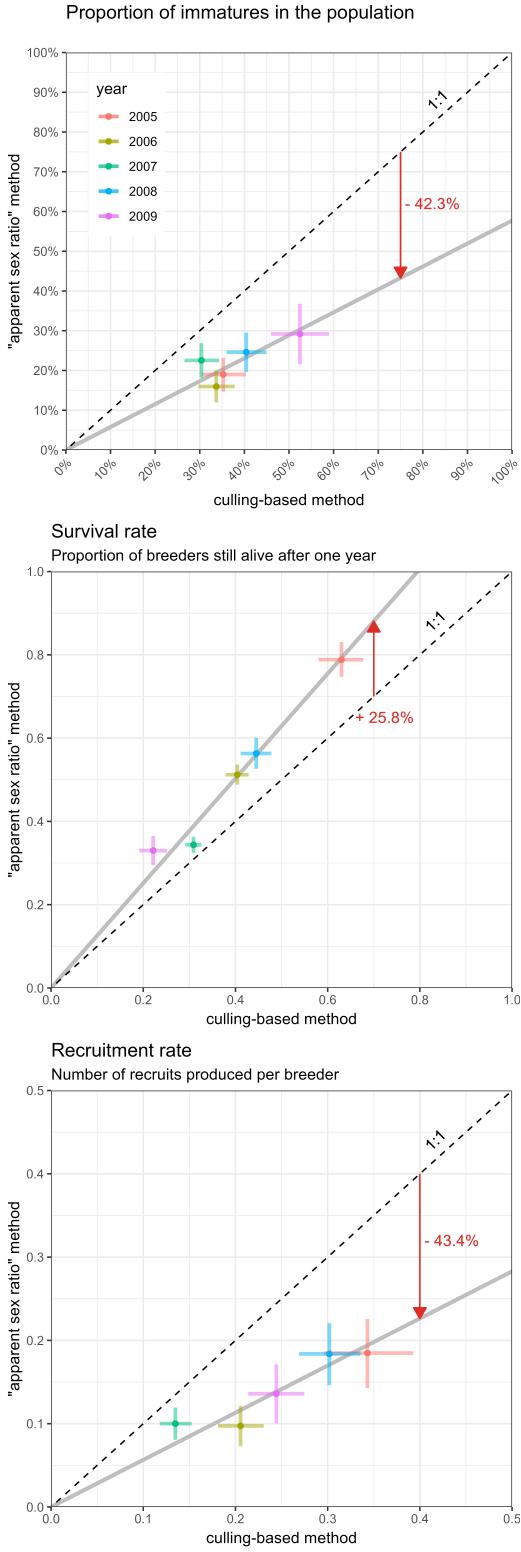


Figure 6: Relationship between parameter estimates obtained from counts (“apparent sex ratio” method) and those obtained from culled individuals (culling-based method); only five years were available for the GB population; the bars represent 95% confidence intervals and the red arrow indicates the direction and average differences between the two methods

216 3.3 Assessing the demographic response to culling strategies

217 When there was no culling pressure, maximum growth rates were very similar for the two populations (Figure
218 7), namely 1.45 [1.36; 1.55] and 1.52 [1.25; 1.88] for GB and FR, respectively. These values corresponded
219 to population increases of 45% and 52% per year, respectively. Assuming a maximum adult survival rate of
220 0.85 [0.7; 1] for both populations, we derived maximum recruitment rates of 0.60 [0.42; 0.78] and 0.68 [0.36;
221 1.06] for GB and FR, respectively.

222 Despite the low number of sex- and age-determined culled individuals, the proxies of culling pressure showed
223 significant changes over time (Figure 8). Therefore, we categorised culling pressure according to this intensity.
224 High culling pressure occurred from 2006 onwards in both populations (although the signal was noisy).
225 However, despite the similar level of culling pressure, the culling strategies of the two populations and their
226 outcomes differed considerably, see 3rd Materials & Methods section. Under high culling pressure, the growth
227 rate decreased to 0.56 [0.53; 0.59] and 0.97 [0.94; 1.00] for GB and FR, corresponding to a decrease of 44%
228 per year and 3% per year, respectively. The GB population declined significantly during the high culling
229 pressure, while the FR population was just stabilised.

230 Adult survival rates under high culling pressure reached 0.44 [0.42; 0.46] and 0.62 [0.59; 0.64] for GB and
231 FR, respectively, and the recruitment rates decreased to 0.13 [0.11; 0.15] and 0.32 [0.29; 0.34], respectively.
232 The resulting decrease in the growth rate of the GB population corresponded to a similar average decrease
233 in adult survival rate (-0.41) and recruitment rate (-0.47), whereas the stabilisation of the FR population
234 was due to a decrease in recruitment rate (-0.36) rather than a decrease in adult survival (-0.23). Overall,
235 the culling strategy used in GB was more effective than that used in FR as it affected both adult survival
236 and recruitment rates.

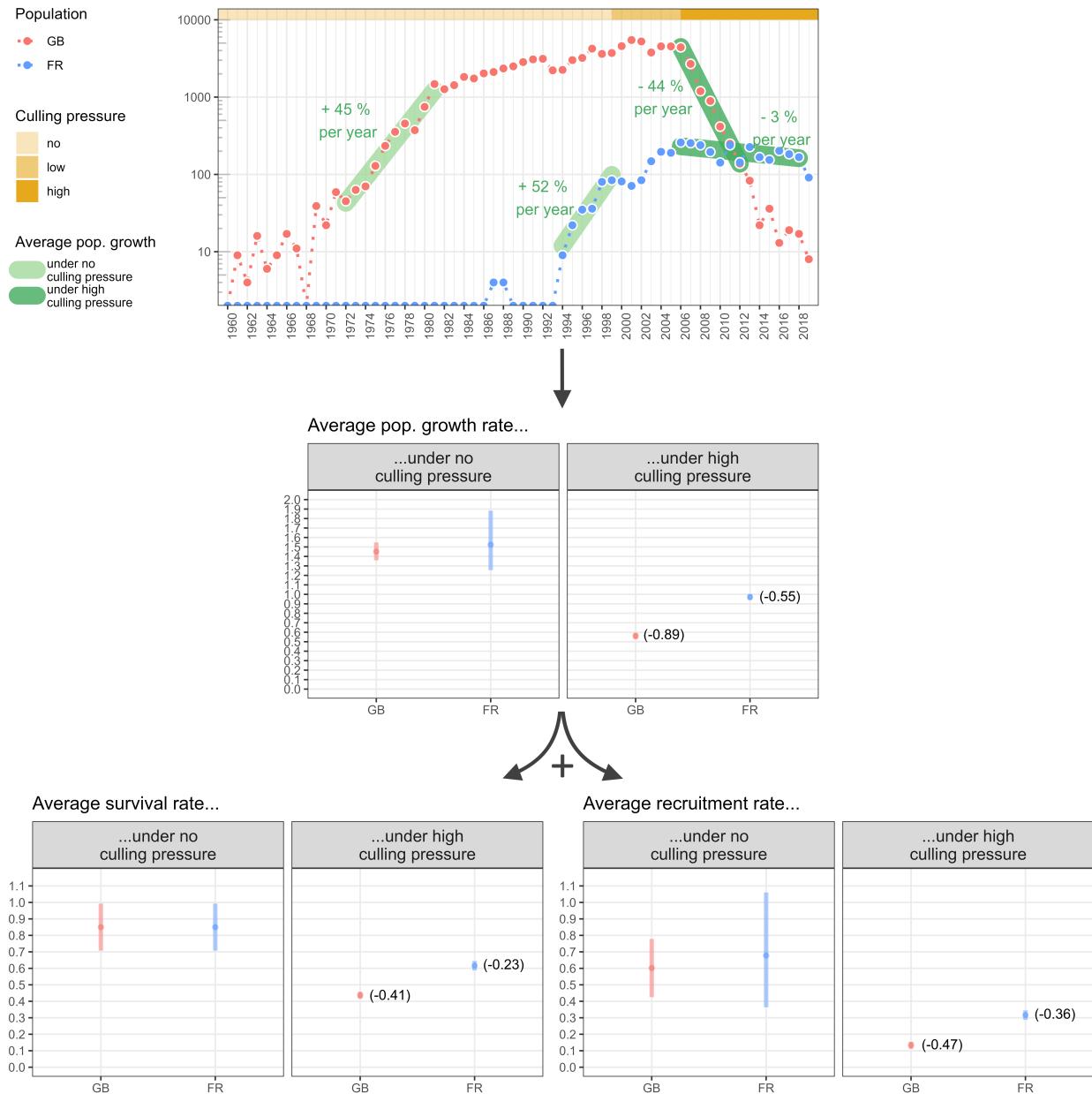


Figure 7: Average effects of culling pressure and different culling strategies (GB versus FR) on population growth rates (derived from counts only), and on adult survival and recruitment rates (derived from counts and reports of apparent sex ratios); the culling pressure for the GB population occurred mainly during the pre-breeding season, while for the FR population it was mainly concentrated during and after the breeding season; adult survival and recruitment rates under high culling pressure were averaged over a period in which consistent culling pressure was observed; the vertical bars show 95% confidence intervals

²³⁷ **4 Discussion**

²³⁸ Disentangling the contribution of vital rates to population growth rate is a key step towards better un-
²³⁹ derstanding a population's response to management actions (Williams et al., 2002). We exploited delayed
²⁴⁰ sexual dichromatism to develop a new approach based on counts to disentangle the growth rate of a popu-
²⁴¹ lation into its two major structural components: adult survival and recruitment rates. The development of
²⁴² the "apparent sex ratio" method is based on a unique data set of two comparable populations of the same
²⁴³ species, the Ruddy duck. Both populations were monitored in a similar way from their introduction to a
²⁴⁴ period of intense culling pressure, but the culling strategies differed markedly. The very similar demographic
²⁴⁵ trajectories and properties of both populations under similar ecological conditions made it possible to assess
²⁴⁶ the response to both culling pressure and culling strategies.

²⁴⁷ **4.1 Monitoring adult survival and recruitment rates**

²⁴⁸ For both populations, vital rate estimates were within the same range, showing some consistency in the results
²⁴⁹ obtained using the "apparent sex ratio" method. The greater variability observed for the FR population
²⁵⁰ was expected as the time series covered a wider range of culling pressure than that of the GB population.
²⁵¹ We found no temporal correlation of demographic parameters between the two populations, suggesting that
²⁵² regional (e.g. weather) rather than large-scale factors (e.g. climate) were predominant. This is consistent
²⁵³ with previous findings showing, for example, that in waterfowl both breeding success and juvenile survival
²⁵⁴ depend on the onset of the laying period, which is closely linked to local weather (e.g. spring temperature,
²⁵⁵ cumulative rainfall, water levels) (Blums & Clark, 2004; Dzus & Clark, 1998; Folliot et al., 2017) and local
²⁵⁶ changes in predation pressure (Jaatinen et al., 2022).

²⁵⁷ The counts used to implement our model took place in mid-winter, at a time when some immature males
²⁵⁸ may be initiating their moult. This could partially violate the assumption that the male-like individuals
²⁵⁹ were exclusively adult males, and thus potentially bias our estimates. Nevertheless, as the survival estimates
²⁶⁰ are in the range of those reported for similar species (Buderman et al., 2023; Buxton et al., 2004; Krementz
²⁶¹ et al., 1997; Nichols et al., 1997), this assumption appears to be reliable. Nevertheless, the time window for

262 the count surveys must be carefully chosen to take into account the time before the immature males start
263 moulting.

264 In some years, the estimated adult survival values were outside the range of expected values, although never
265 significantly above 1. One possible explanation for these discrepancies is a violation of the “closed population”
266 assumption. Although the ringing data suggest that Ruddy ducks do not undertake long-distance seasonal
267 migrations in GB (Henderson, pers. comm.), the possibility that immigration occurred cannot be completely
268 ruled out. The outliers in survival and recruitment rates observed in FR in 2002 and 2012 suggest that there
269 was significant immigration in those years.

270 The theoretical maximum population growth rate is an inherent trait related to the demographic charac-
271 teristics of a species (Dmitriew, 2010; Niel & Lebreton, 2005). This trait is expected to be similar in all
272 Ruddy duck populations around the world. In our study, we observed that the maximum growth rates of
273 the two populations were very close to each other, which was expected as both populations were exposed to
274 similar ecological conditions, i.e. no harvest and similar breeding conditions. In addition, before the eradi-
275 cation programmes started, the two populations had population growth rates much closer to the expected
276 maximum values than any other native duck species in the same areas during the same period. Prior to the
277 eradication programmes, the survival rates of both populations were in the upper range of those observed
278 for other diving duck species (e.g. ~ 0.8 for the Common pochard (Folliot et al., 2020; Nichols et al., 1997)).
279 Predation on diving duck nests was found to decrease during the breeding season in a French fishpond
280 complex (Bourdais et al., 2015). As Ruddy ducks generally breed later than Common pochards and Tufted
281 ducks, both a higher survival rate and greater nesting success could be responsible for the high growth rates
282 of Ruddy duck populations in the absence of culling. The “apparent sex ratio” method provided estimates of
283 adult survival and recruitment rates that were not only consistent with observed increasing population sizes,
284 but also compatible with both declining phases following the implementation of eradication programmes and
285 observed differences in population growth rates under different culling strategies. As expected, higher values
286 for survival and recruitment rates were observed when the culling strategy focussed on the post-breeding
287 season rather than the pre-breeding season.

288 Assuming that culling data provide a good picture of interannual variability in age structure (Fox et al., 2014),

289 the strong correlation between estimates derived from counts and estimates derived from culled individuals
290 demonstrates the ability of the “apparent sex ratio” method presented here to properly capture interannual
291 variation in demographic parameters such as survival and recruitment rates. The fact that the observed
292 correlation was strong despite being based on a short time series strengthens its robustness. In addition,
293 the strength of the relationship between the parameters obtained with the two approaches also indicates
294 that important assumptions of our model were met, such as the constant adult sex ratio. The temporal
295 autocorrelation of the adult sex ratio could be explained by the fact that it includes many age cohorts, which
296 makes it structurally strong. In the long term, there may be substantial fluctuations in the adult sex ratio,
297 but our results suggest that it is not necessary to monitor and update it annually. The proportion of males
298 among adults estimated in the present study (i.e., 0.60) is within the range of values observed in native
299 Ruddy duck populations of North America (0.62 in Bellrose (1980)) and in other duck species (Wood et al.,
300 2021). The “apparent sex ratio” method consistently yielded recruitment rates that were almost a factor of
301 two lower than the method based on culling data, a result consistent with higher vulnerability of immatures
302 to shooting (Fox et al., 2014). In North America, for example, immature waterfowl were found to be 1.3 to
303 2.6 times more vulnerable to hunting than adults (Bellrose, 1980).

304 As no estimates of adult survival were available in the literature (see Buderman et al., 2023), we were
305 unable to properly assess the accuracy of the “apparent sex ratio” method for estimating this parameter.
306 Nevertheless, there is evidence that the method is not highly biased, as explained above. Thus, the “apparent
307 sex ratio” method would be much more efficient for monitoring key demographic parameters than alternative
308 methods such as those based on culling/hunting bag data, which have been shown to be unreliable (Fox et
309 al., 2014), even when focusing on the end of the hunting season to limit potential bias (Fox et al., 2016).

310 4.2 Evaluating eradication strategies

311 Both populations responded to culling pressure with a significant decline in their growth rate, meaning that
312 the increased mortality from culling was not compensated for by an increase in natural survival or improved
313 breeding success. Culling was therefore effective in affecting the demographic trajectory of Ruddy duck
314 in both cases. However, the decline in population growth was much greater in GB, where the population

315 declined steadily and sharply, than in FR, where the increase was halted and the population size remained
316 more or less stable. The greater decline in population growth observed in GB compared to FR, was due
317 to a greater decline in adult survival and recruitment rates. Overall, adult survival and recruitment rates
318 appeared to be significantly affected by culling in both countries. However, culling pressure targeting future
319 breeders (pre-breeding culling strategy applied in GB), appeared to be much more effective in reducing both
320 survival and recruitment rates than culling pressure targeting breeders and newly born individuals during
321 the breeding and rearing season (breeding culling strategy applied in FR).

322 As expected, the pre-breeding culling strategy had the same effect on adult survival and recruitment rates,
323 as a bird killed before the breeding season reduces the breeding population (lower adult survival rate) and
324 prevents the bird from reproducing (lower recruitment rate). When using the breeding culling strategy,
325 the recruitment rate was more affected than the adult survival rate. This was probably due to the fact
326 that immatures are much more susceptible to being shot than adults (Bellrose, 1980; Fox et al., 2014).
327 The compensation hypothesis states that the increased mortality of immatures due to harvesting may be
328 compensated for by a higher natural survival (Cooch et al., 2014), but this hypothesis is not supported by
329 the observations. Overall, there was no evidence of a compensatory mechanism in either context. This could
330 be due to the fact that both populations were far from reaching carrying capacity, meaning that competition
331 for resources was relaxed (Péron, 2013).

332 With more or less comparable culling pressure, the pre-breeding culling strategy proved to be much more
333 effective. However, this does not mean that the breeding culling strategy was not also effective. On the
334 contrary, it proved to be quite efficient in stopping both the population increase and the expansion of the
335 distribution area, and in triggering a significant population decline in a second attempt by slightly increasing
336 the culling pressure. So if winter culling is not possible, which was the case in FR, culling during the breeding
337 season is effective in stopping or reversing the increase in populations of this species. However, the breeding
338 culling strategy is compromised by the highly variable detectability of immatures, as these individuals colonise
339 new ponds. This leads to fluctuations in the recruitment rate and thus jeopardises the long term efficiency
340 of this strategy. This problem does not exist with the pre-breeding strategy, as the Ruddy ducks form large
341 flocks that occupy a limited number of sites in winter (Johnsgard & Carbonell, 1996).

342 The response of the two Ruddy duck populations to culling pressure shows that it is necessary to take the
343 culling period into account in order to make correct predictions of population trajectory. Predicting the
344 effects of harvest pressure on a waterfowl population is then not only a question of the level of harvest,
345 but also the timing of the harvest (Kokko et al., 1998). Unexpectedly, our results suggest that harvesting
346 waterfowl during the breeding season has far less impact on population growth rate than harvesting in
347 winter. Consequently, a target for effective management of harvested waterfowl could be an earlier starting
348 and ending of the hunting season.

349 **4.3 Implications for waterfowl management**

350 An important implication of our study concerns the timing of harvest to limit its impact on populations
351 growth rates. Unexpectedly, our results show that the persistence of waterfowl populations should benefit
352 from earlier openings and closings of the hunting season (see above).

353 Tracking fluctuations of population abundance is a common tool for determining the conservation status of a
354 population (e.g. Folliot et al., 2022). But tracking abundance alone does not provide enough information to
355 assess the underlying mechanisms behind changes in population growth (Williams et al., 2002). This requires
356 “digging deeper” (Austin et al., 2000), for example, by monitoring individuals to assess parameters such as
357 survival and recruitment rates (e.g. Arnold, 2018). However, monitoring individuals is time-consuming
358 (e.g. Souchay & Schaub, 2016) and not always possible for endangered species. The “apparent sex ratio”
359 method makes it possible to circumvent these disadvantages in dichromatic species with delayed sexual
360 maturity of the males. These species include the White headed duck in southern Spain, which is highly
361 endangered and cannot be disturbed through capture-mark-recapture. Thus, by improving census strategies
362 and disaggregating the effects of variations in adult survival and productivity on population growth using
363 the approach presented here, much more invaluable information on the relevance of management strategies
364 could be obtained.

365 Finally, we would like to emphasise that the main aim of our approach was to decompose population growth
366 into its two main components, and not to provide unbiased estimates of adult survival or recruitment rates.

367 However, if necessary, count surveys could be designed to do so, at least in theory, and for species that exhibit
368 observable delayed dichromatism. This trait affects most species of stiff-tailed duck (Johnsgard & Carbonell,
369 1996), as well as numerous other duck species, including Tufted duck, *Aythya fuligula*, Black scoter, *Melanitta*
370 *nigra americana*, Common Goldeneye, *Bucephala clangula americana* (Bellrose, 1980; Johnsgard, 1978).
371 Modifying standard monitoring protocols to distinguish between male-like and female-like individuals is
372 almost costless, but worth the effort as it would greatly increase the efficiency of conservation/management
373 actions (Nichols & Williams, 2006).

374 Appendix

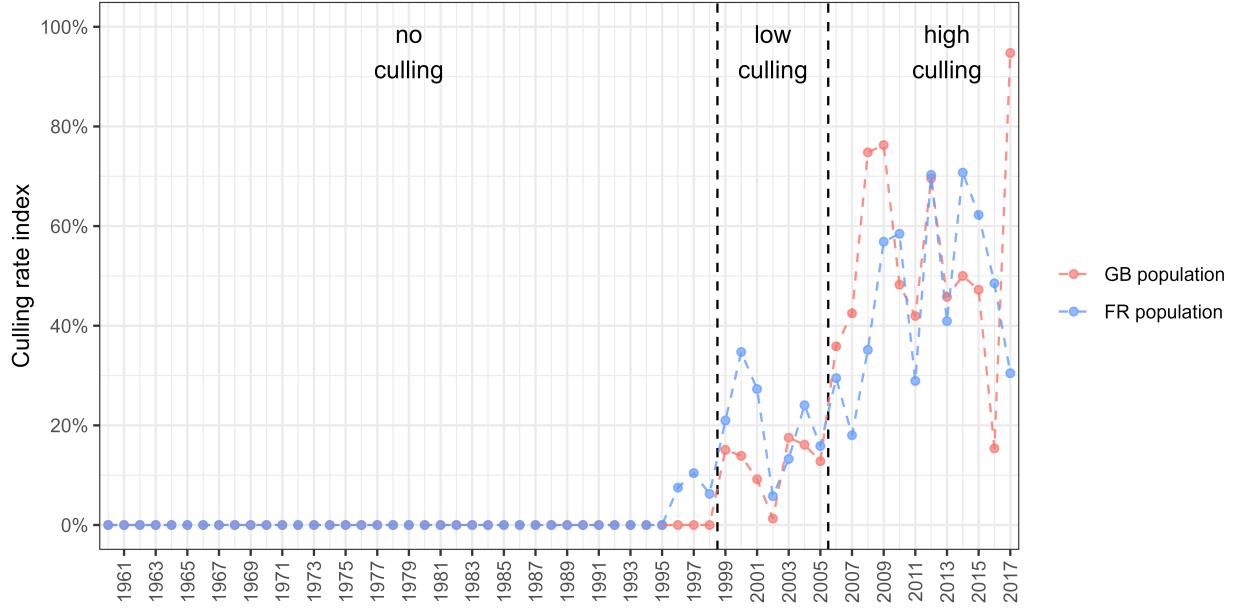


Figure 8: Definition of three periods with different culling pressure using an index based on an estimate of the adult culling rate; as only half of the culled individuals in France were age-determined before 2009, we applied the average age ratio over 2009-2019 to the remaining individuals in order to obtain an estimate of the number of adults in the culling data before 2009; the culling rate increased over time, so we divided the time series into three categories: we defined a “no culling” period before 1999 as the culling rate for both countries was mostly zero and always below 10%, we defined a “low culling” period between 1999 and 2005 as the culling rate for both countries fluctuated around 20%, the culling rate then increased significantly for both countries, therefore we defined a “high culling” period from 2006 onwards

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³⁸⁸ **Conflict of interest disclosure**

³⁸⁹ The authors declare that they comply with the PCI rule of having no financial conflicts of interest in relation
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³⁹¹ **Data, scripts, code, and supplementary information availability**

³⁹² Data and scripts are available online: <https://doi.org/10.5281/zenodo.11471723>; Tableau et al., 2024

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