

# Delayed dichromatism as a convenient tool to disentangle the effects of survival and productivity on the population dynamics in waterfowl

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## Abstract

Monitoring the number of individuals is by far the most popular strategy for studying the environmental factors that determine population dynamics and for measuring the effectiveness of management actions aimed at population recovery, control, or eradication. Unfortunately, monitoring numbers is inefficient in identifying the mechanisms underlying demographic processes and, in particular, in assessing the extent to which population growth rate is influenced by changes in adult survival rather than variations in reproductive parameters. The usual methods for determining adult survival and productivity, such as analysing capture-mark-recapture data and game-hunting bag data, rely on capturing animals, which has three major drawbacks. Firstly, the selectivity of capture methods can lead to a distorted representation of the underlying population structure. Secondly, the capture and release of animals is quite invasive, causing stress and posing an additional threat to the often already endangered species. Thirdly, release may simply be impossible, especially for invasive alien species that represent a threat to native taxa. In many waterfowl species, sexual dichromatism is observed in adults, while immatures of both sexes display a cover-up plumage similar to that of adult females. Using two non-native populations of Ruddy duck, a species that exhibits this pattern, we present a non-invasive method to estimate the respective contribution of adult survival and productivity (also called recruitment rate) to the population growth rate. We exemplify how changes in the apparent proportion of males in counts (due to female-like immature males gradually acquiring adult male plumage) can be exploited to split the population growth rate into adult survival and recruitment rates, which contributes significantly to understanding the demographic impact of two different eradication strategies.

**Keywords:** intrinsic growth rate - vital rate - reproductive success - fecundity - sex ratio - age ratio  
- *Oxyura jamaicensis* - harvesting

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## **1 Introduction**

Assessing the growth rate of populations is a first step towards a better understanding of the factors underlying the dynamics of natural populations (e.g. Niel & Lebreton, 2005). It is also crucial for measuring the effectiveness of management actions that are taken to recover, control, or eradicate populations (Shea & NCEAS Working Group on Population Management, 1998). Among the approaches available to managers to achieve these goals, those that rely on monitoring the number of individuals (counts) are by far the most popular (Rintala et al., 2022). In many cases, these methods allow the investigation of environmental factors underlying changes in population size, and thus help in the implementation of mitigating actions (Faillettaz et al., 2019). Unfortunately, relying solely on the monitoring of population abundance generally hinders the fundamental understanding of the demographic mechanisms underlying changes in population growth rates. More specifically, on the basis of counts alone, it is almost impossible to assess the extent to which population growth rate is affected by changes in adult survival rather than to variations in productivity (Austin et al., 2000), which is defined as the product of all reproductive parameters. To assess the relative sensitivity of population growth rate to factors affecting adult survival or reproductive parameters, demographers usually rely on the monitoring of individuals by capture-mark-recapture (Lebreton et al., 1992). In practise, however, capture and release of individuals cannot always be relied upon, as the legal status of the species precludes any additional disturbance (e.g. critically endangered species) or prohibits the release of live individuals (e.g. invasive species or pests). Even though capture-mark-recapture methods are efficient for assessing demographic parameters, they have potential drawbacks, such as their invasiveness, which can affect behaviour and thus survival or reproductive success, and their affordability when time and money are scarce. Genetic monitoring is a non-invasive alternative to capture-mark-recapture, but it is costly and requires large samples (Caniglia et al., 2011).

Most of the time, therefore, managers make the best of a bad job by using counts as a viable option for tracking changes in population size and then population growth rate to assess the relevance of management actions. Distinguishing between immatures and adults in counts allows estimation of adult survival rate, i.e. the proportion of breeders that have survived for one year, and productivity/recruitment rate, i.e. the number of immatures produced per breeder that reach sexual maturity, which are, by definition, recruited in the breeding population. This additional effort makes it possible to assess the relative influence of each of these components on the population growth rate. Alternatively, the assessment of age structure in hunting bags has been used to infer the role of decreasing reproductive success in population declines in a number of game species, including ducks and geese, but suffers from intractable biases (Fox & Cristensen, 2018). Of course, the latter approach is not appropriate for protected/endangered species.

Here we exemplify how, in dimorphic species, delayed sexual maturity of males can be used to estimate adult survival and recruitment rates by distinguishing male-like and female-like individuals in repeated winter counts. We used two non-native European populations of Ruddy duck (*Oxyura jamaicensis*) as a study model. Like many other duck species, Ruddy duck is dimorphic: newborn males look like females until the prenuptial moult, which takes place in late winter. As a result of the late prenuptial moult of immatures, the apparent proportion of males increases over the course of the wintering season. These changes in the apparent proportion of males during this period are therefore directly related to the proportion of immatures in the populations and thus to the reproductive success of the previous breeding season. The “apparent sex ratio” method developed in this study was used to estimate adult survival and recruitment rates and evaluate the effects of two different eradication strategies used in Great Britain and France, respectively.

## **2 Materials & methods**

The Ruddy duck is a stiff-tailed duck native to the American continent. Starting from seven individuals acclimatised at the Slimbridge Wetland Centre in the 1940s (Gutiérrez-Expósito et al., 2020), a feral population began to establish in Great Britain and the first attempts at reproduction in the wild were observed in the 1960s (Figure 1). This feral population grew rapidly and spread across the country until it reached more than 5,000 individuals in the early 2000s. In the late 1980s, a feral population also began to establish

49 on the continent, particularly in France, supposedly due to the arrival of British-born individuals. However,  
 50 in contrast to observations in Great Britain, Ruddy ducks have not spread very much in France. Most  
 51 observations and breeding attempts have been concentrated in the north-west of the country. In France,  
 52 almost no Ruddy ducks were observed in winter outside the lake of Grand Lieu ( $47.09^{\circ}\text{N}$ ,  $1.67^{\circ}\text{W}$ ), which  
 53 greatly facilitated the monitoring of this population. Ruddy ducks are considered a major threat to the  
 54 endangered native White-headed duck (*Oxyura leucocephala*) in the south-western Mediterranean, as they  
 55 hybridise and thus pose an increased risk of genetic pollution and genetic assimilation of the latter by the  
 56 former (Muñoz-Fuentes et al., 2007). To reduce the risk of genetic pollution of the White-headed duck by  
 57 the Ruddy duck, eradication measures were implemented in both Great Britain and France in the late 1990s  
 58 (Gutiérrez-Expósito et al., 2020), followed by a European Ruddy duck eradication plan in 1999 (Hughes et  
 59 al., 1999) (Figure 1).

60 The trends and status of the White-headed duck and Ruddy duck populations used to be determined ex-  
 61clusively by counts. It was therefore not possible to assess the factors impairing the recovery of the former  
 62 and the effectiveness of the eradication programme for the latter. In particular, counts alone cannot be used  
 63 to assess the relative impact of changes in adult survival and recruitment rates on population growth rates,  
 64 which is a prerequisite for identifying limiting factors.

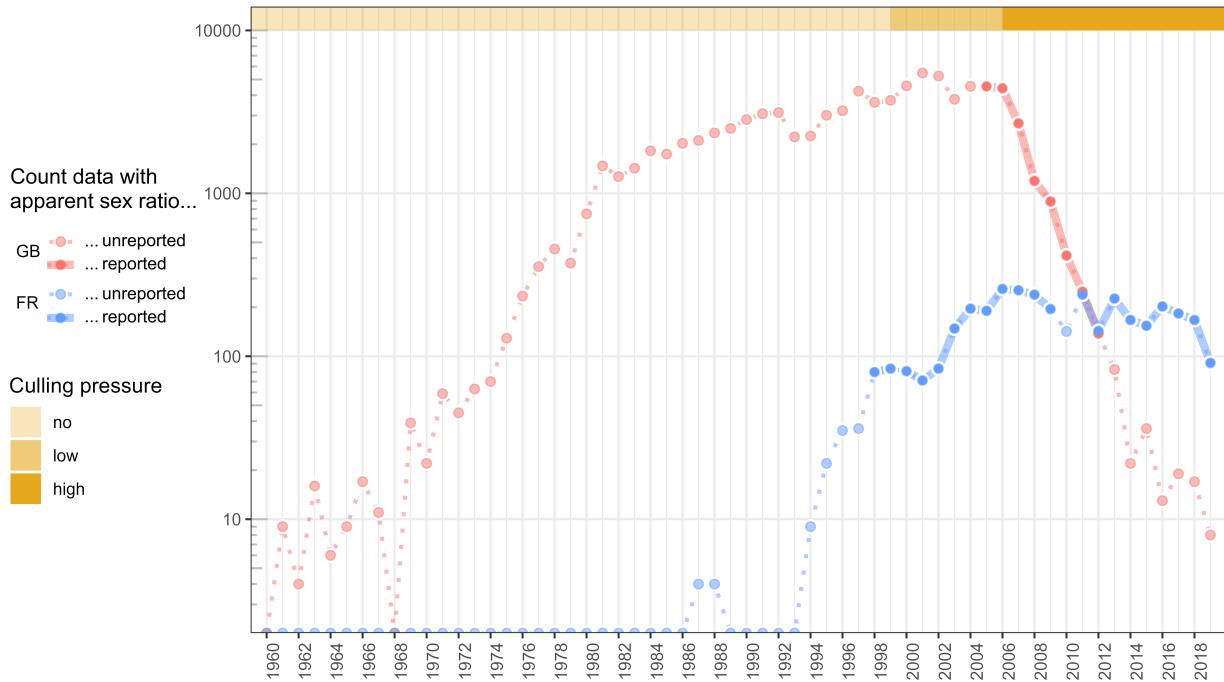


Figure 1: Temporal changes in the number of Ruddy ducks (logarithmic scale) in Great Britain (GB) and France (FR) from 1960 to 2019, with references to periods when apparent sex ratios were reported

65 Like many other duck species, both White-headed duck and Ruddy duck exhibit delayed dichromatism,  
 66 i.e. young males acquire the colourful plumage typical of their species and can therefore be distinguished from  
 67 females during the pre-breeding period, typically from mid-winter in the earliest individuals (Baldassarre,  
 68 2014). Delayed dichromatism generally explains the discrepancies in estimates of apparent proportions  
 69 of males counted in winter compared to proportions of males counted in spring or males identified from  
 70 individuals culled in winter (Figures 2 & 3).

71 Interestingly, delayed dichromatism allows the estimation of adult survival and recruitment/productivity by  
 72 observing the seasonal evolution of the apparent sex ratio and making few and reasonable assumptions (see  
 73 next section).



Figure 2: A flock of Ruddy ducks observed in winter, including ten female-like individuals with whitish striped cheek (some of which are immature males), four male-like individuals with white cheek and black cap (all adult males), three unidentified individuals (1<sup>st</sup>, 4<sup>th</sup>, and 10<sup>th</sup> from left) © Jay McGowan - 3 February 2013 - Tompkins, New York, United States

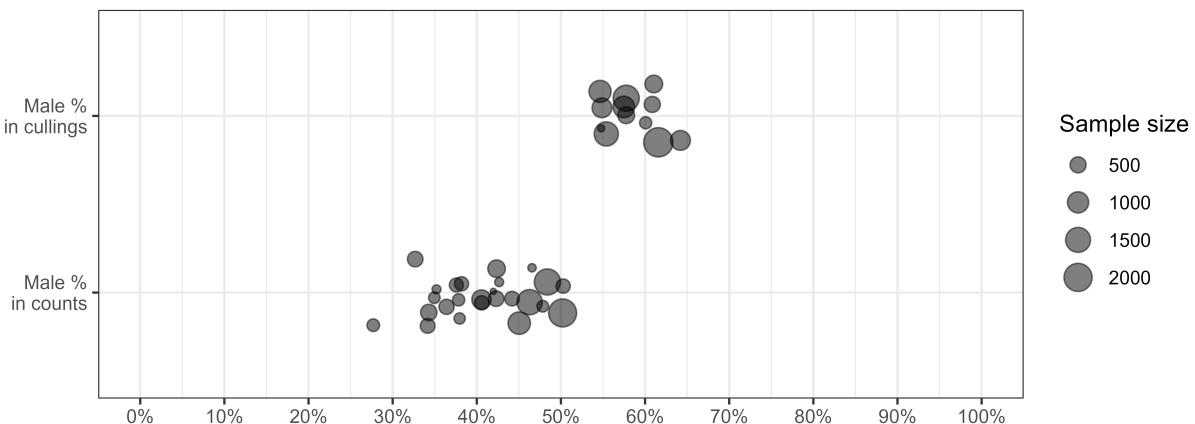


Figure 3: Proportions of males estimated by internal examination of individuals culled in a population under the eradication programme over one year (Male % in cullings), and proportions of individuals showing male-like plumage in a population counted in winter (Male % in counts); discrepancies between the two estimates are due to delayed dichromatism (immature males looking like females before moulting); we used these discrepancies to estimate survival and productivity/recruitment

<sup>74</sup> We monitored the demography of Ruddy duck populations in both Great Britain (GB) and France (FR) by  
<sup>75</sup> conducting extensive counts on the wintering grounds between 1 December and 31 January. In addition,  
<sup>76</sup> the apparent sex ratio could be determined using winter counts that distinguished between female-like and  
<sup>77</sup> male-like individuals. Such counts were carried out in Great Britain from 2006 to 2012, which corresponded  
<sup>78</sup> to a period of sharp population decline (Figure 1). In France, count surveys taking into account apparent  
<sup>79</sup> sex were conducted in 1999, 2001-2009, and 2012-2019. The French population grew rapidly in the first years  
<sup>80</sup> of monitoring and then stabilised from 2006 onwards as a result of high culling pressure. In both countries,  
<sup>81</sup> culling was carried out under strict official control so that the exact number of individuals culled was known.  
<sup>82</sup> Although the age ratio of birds culled before 2009 could not be determined with certainty in France, it was  
<sup>83</sup> still possible to roughly categorise the culling pressure in both countries into three categories: “no culling”  
<sup>84</sup> before 1999, “low culling” between 1999 and 2005 and “high culling” from 2006 onwards (see Figure 8 in  
<sup>85</sup> Supplementary Material).

86 **2.1 Inferring adult survival and recruitment rates from changes in population  
87 size and proportion of immatures**

88 The proportion of immature males is the difference between the proportion of female-like individuals between  
89 winter and spring. The proportion of all immature individuals is twice as high if we assume an even  
90 sex ratio at birth (Bellrose et al., 1961; Blums & Mednis, 1996) and comparable mortality rates between  
91 immature females and immature males (Ramula et al., 2018; Wood et al., 2021). However, spring counts  
92 suffer from intractable problems, such as the different detectability of females and males during this season.  
93 Fortunately, the proportion of immatures can also be estimated from winter counts alone, assuming a known  
94 and unchanged adult sex ratio along the monitoring period (Devineau et al., 2010).

95 The adult sex ratio is therefore a prerequisite for estimating the proportion of immatures. This ratio could be  
96 derived from culling data as we determined the sex of each adult culled as part of the eradication programmes  
97 in both Great Britain and France. Nevertheless, the low number of individuals culled in France prevented  
98 from getting robust adult sex ratio. In Great Britain, adult sex ratios did not differ from month to month,  
99 so data from culled adults collected throughout the whole biological cycle were pooled together each year.  
100 The proportion of males also did not differ statistically between years in which more than 500 adults were  
101 culled. This is consistent with previous findings showing that the adult sex ratio in ducks is usually very  
102 stable in the short term (Wood et al., 2021) but can fluctuate in the long term. Therefore, we pooled the  
103 data from all adults culled in the control programme to estimate the adult sex ratio, see Equation (1).

104 Assuming the additive property of the binomial distribution, the proportion of immatures in a population can  
105 be derived from the proportion of males among adults (alternatively the adult sex ratio) and the proportion  
106 of male-like individuals in the different counts in winter (assuming that all these individuals are adult males),  
107 see Equation (2). By combining the population growth rate derived from the interannual variation of an  
108 index of abundance with the proportion of immatures, adult survival and recruitment rates can be derived,  
109 see Equation (3) and Figure 4. If the absolute value of the population size is known, the number of adults  
110 and the number of recruits can be estimated, see Equation (4). We called this approach the “apparent sex  
111 ratio” method.

Variable name	Description
<i>Observation (input)</i>	
$AM$	Total number of adult males culled as part of the eradication programme
$AF$	Total number of adult females culled as part of the eradication programme
$CML_{i,t}$	Number of male-like individuals counted in population $i$ in year $t$
$C_{i,t}$	Number of individuals counted in population $i$ in year $t$
$N_{i,t}$	Size of the population $i$ in year $t$ (maximum number of individuals counted in the wintering population)
<i>Parameter (output)</i>	
$p(m a)$	Proportion of males in adults (or probability of being a male knowing it is an adult)
$p(a \cap m)_{i,t}$	Proportion of adult males in population $i$ in year $t$
$p(a)_{i,t}$	Proportion of adults in population $i$ in year $t$
$p(re)_{i,t}$	Proportion of immatures/recruits in population $i$ in year $t$
$s_{i,t}$	Adult survival rate, i.e. proportion of individuals in year $t - 1$ still alive in year $t$
$r_{i,t}$	Recruitment rate, i.e. number of recruits in population $i$ in year $t$ per individual in year $t - 1$
$S_{i,t}$	Number of adults in population $i$ in year $t$
$R_{i,t}$	Number of recruits in population $i$ in year $t$

$$p(m|a) \sim \text{Beta}(AM, AF) \quad (1)$$

$$\begin{aligned}
CML_{i,t} &\sim \text{Binom}(p(a \cap m)_{i,t}, C_{i,t}) \\
&\sim \text{Binom}(p(m|a).p(a)_{i,t}, C_{i,t}) \\
&\sim \text{Binom}(p(m|a).(1 - p(re)_{i,t}), C_{i,t})
\end{aligned} \tag{2}$$

$$\begin{aligned}
s_{i,t} &= \frac{N_{i,t}}{N_{i,t-1}} \cdot (1 - p(re)_{i,t}) \\
r_{i,t} &= \frac{N_{i,t}}{N_{i,t-1}} \cdot p(re)_{i,t}
\end{aligned} \tag{3}$$

$$\begin{aligned}
S_{i,t} &= N_{i,t} \cdot (1 - p(re)_{i,t}) \\
R_{i,t} &= N_{i,t} \cdot p(re)_{i,t}
\end{aligned} \tag{4}$$

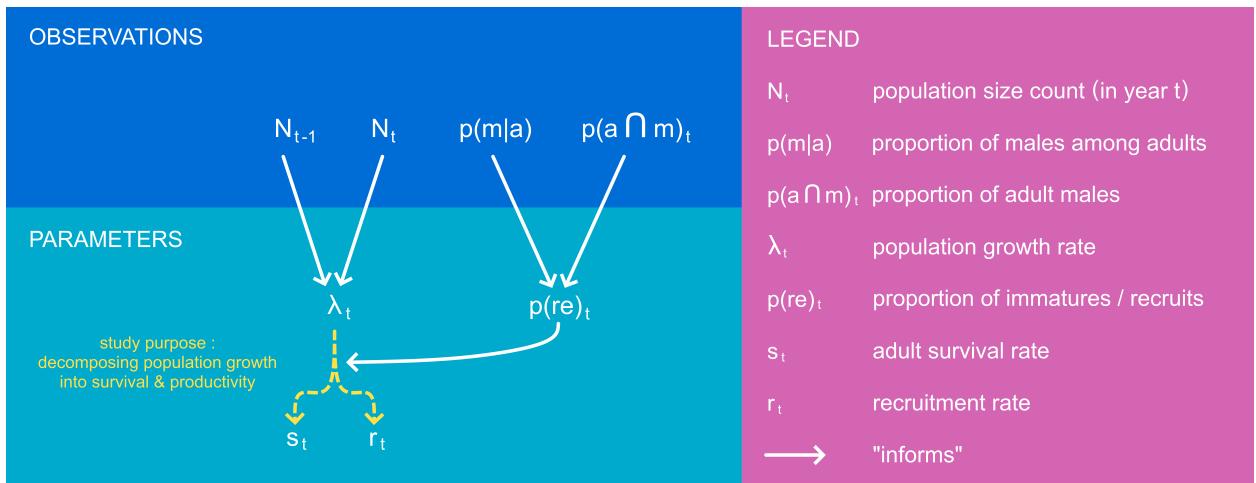


Figure 4: Overview of the “apparent sex ratio” method used to disentangle the population growth into survival (adult survival rate) and productivity (recruitment rate)

## 112 2.2 Validating the “apparent sex ratio” method

- 113 The relevance of the method was assessed by testing the likelihood of the estimates of both vital rates. For  
 114 the adult survival rate, we checked whether the values were contained in the interval [0; 1], and compared  
 115 our estimates with those in the literature Nichols et al. (1997). Validation of the recruitment rate was  
 116 more difficult because if the lower limit is 0, the upper limit is a product of all maximum values of its  
 117 individual components, i.e. nesting rate, nesting success, clutch-size, hatching success, pre-fledging survival,  
 118 and fledgling survival. These reproductive parameters may be known for some duck species, but the last one  
 119 is never mentioned (e.g. Baldassarre, 2014).
- 120 To overcome this problem, we developed an indirect approach to estimate the maximum expected recruitment  
 121 rate without exploitation and assuming that the level of population density does not affect the components  
 122 of the recruitment process. Estimated values above this maximum recruitment rate would reflect a partial  
 123 detection of adult males. On the other hand, null or negative recruitment rates would indicate that immature  
 124 males acquired male plumage prior to the survey, or that the assumption about the stability of the adult sex  
 125 ratio was violated.

126 In practise, the maximum recruitment rate was estimated from the difference between the maximum growth  
 127 rate and the maximum survival rate of adults, see Equation (5). This relationship resulted from a simple  
 128 consideration for a closed population: the population size in year  $t$  is equal to the number of adults that  
 129 survived the whole year  $t - 1$  plus the offspring produced in year  $t - 1$  that survived to the reproductive  
 130 period of year  $t$ , i.e. the recruitment in year  $t$  (Flint, 2015). In an open population, adult survival and  
 131 recruitment rates are confounded with adult and recruit migrations, respectively, but this does not change  
 132 the equation. This relationship becomes more complex when a species with delayed maturity is considered,  
 133 see Robertson (2008).

Variable name	Description
$N_t$	Number of adults in year $t$
$D_t$	Number of adults dead during year $t$
$R_t$	Number of recruits in year $t$
$s_t$	Adult survival rate, i.e. proportion of individuals in year $t - 1$ still alive in year $t$
$r_t$	Recruitment rate, i.e. number of recruits in year $t$ produced per individual in year $t - 1$
$\lambda_t$	Growth rate of the population between year $t - 1$ and year $t$

$$\begin{aligned}
 N_t &= N_{t-1} - D_{t-1} + R_t \\
 N_t &= N_{t-1} - (1 - s_t).N_{t-1} + r_t.N_{t-1} \\
 N_t &= s_t.N_{t-1} + r_t.N_{t-1} \\
 \frac{N_t}{N_{t-1}} &= s_t + r_t \\
 \lambda_t &= s_t + r_t \\
 r_t &= \lambda_t - s_t
 \end{aligned} \tag{5}$$

134 The maximum adult survival rate was defined as a uniform distribution on the interval [0.7; 1]. The lower  
 135 limit corresponded to the upper range of survival rates observed in waterfowl species of similar weight, and  
 136 the upper limit was set to 1, as survival rates of long-lived waterfowl species are very high (Buxton et al.,  
 137 2004; Koons et al., 2014; Krementz et al., 1997; Nichols et al., 1992, 1997).

138 For both populations, the maximum growth rates occurred during their geographic expansion phase, which  
 139 preceded the start of the eradication programmes (Figure 1). To estimate robust maximum population  
 140 growth rates for both populations, we smoothed the annual growth rates over a consistent time period by  
 141 using a linear regression on the logarithm scale, see Equation (6). We discarded the data for Great Britain  
 142 before 1972 as the estimates were likely noisy when the population was low (below 50 individuals) (Figure  
 143 1). After reaching the threshold of 1,000 individuals, the population growth in Great Britain (GB) showed a  
 144 strong inflection, although culling had not yet started (Figure 1). This observation suggests that a negative  
 145 density-dependent process could occur beyond 1,000 individuals and led us to consider only the first sequence  
 146 of the time series to infer the maximum growth rate in Great Britain, i.e. 1972-1981. For the French (FR)  
 147 population, the sequence without culling pressure covered the period 1994-1999.

Variable name	Class	Description
$N0_{i,J}$	Parameter	Intercept of the regression model
$\lambda_{i,J}$	Parameter	Average population growth rate over a restricted time interval $J$ for a population $i$ (in $\text{year}^{-1}$ )
$t \in J$	Index	Year index within the time interval $J$
$\sigma_{i,J}$	Parameter	Standard deviation of the regression model

$$\log(N_{i,t}) \sim \text{Norm}(N0_{i,J} + \log(\lambda_{i,J}).t, \sigma_{i,J}) \tag{6}$$

148 Estimating vital rates within an acceptable range, as defined above, would not prove that our modelling  
 149 approach correctly reflects the actual interannual variability. To determine this, we compared the proportions  
 150 of immatures and the vital rates derived from counts with those derived from culling, i.e. from individuals  
 151 culled as part of the eradication programme. As we restricted the dataset to years with more than 100  
 152 individuals culled in winter for sake of precision, the analysis only covered five years of the GB population  
 153 time series. The proportions of immatures were estimated from the culled individuals by checking the  
 154 presence of the bursa of Fabricius, which is only present in immature individuals (Hochbaum, 1942); see  
 155 Equation (7). Adult survival and recruitment rates were then derived from the proportion of immatures and  
 156 the Equation (3).

Variable name	Class	Description
$SI_{i,t}$	Observation	Number of immatures culled in the wintering population $i$ in year $t$
$S_{i,t}$	Observation	Number of individuals culled in the wintering population $i$ in year $t$

$$SI_{i,t} \sim \text{Binom}(p(re)_{i,t}, S_{i,t}) \quad (7)$$

## 157 2.3 Assessing the effects of culling strategies

158 During the period when culling pressure was high (i.e. from 2006 onwards), eradication strategies in Great  
 159 Britain and France differed. In Great Britain, most individuals were culled in winter (53.4% of adults were  
 160 culled before 30 May), whereas in France, most individuals were culled during the breeding season (81.6%  
 161 of adults were culled after 30 May). We investigated whether these strategies had different effects on the  
 162 populations by comparing the resulting growth rates and the relative contributions of both vital rates to them  
 163 (by comparing the average values during the period of high culling pressure with proxies for the maximum  
 164 vital rates estimated when both populations reached their maximum growth, see Section 2.2). In France, a  
 165 LIFE project was implemented from 2019 to intensify the culling pressure, especially in winter and spring.  
 166 We therefore excluded the FR time series from 2019 onwards in order to compare homogeneous culling  
 167 strategies.

## 168 2.4 Statistical framework

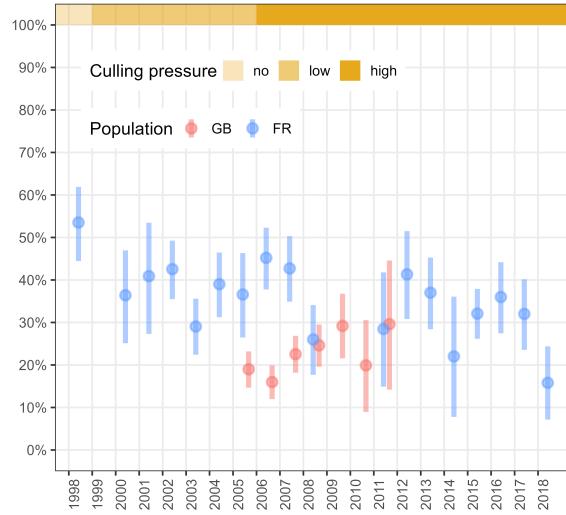
169 We used the Bayesian framework to implement all submodels presented in Section 2. The Bayesian framework  
 170 is both straightforward and efficient in propagating error through the parameters. We used uninformative  
 171 priors for all parameters. Since the maximum growth rate is a life history trait expected to be stable  
 172 between populations of the same species, we used an uninformative hierarchical prior for this parameter.  
 173 We generated three chains of length 500,000, with a thinning of 10 to avoid autocorrelation in the samples,  
 174 and discarded the first 2,000 samples as burn-in. We checked the chain convergence using the Gelman and  
 175 Rubin convergence diagnostic ( $R < 1.1$ , Gelman & Rubin (1992)). The models were fitted using NIMBLE  
 176 (de Valpine et al., 2017) run from R (R Core Team, 2022). The values  $\mathbf{X}[\mathbf{Y}; \mathbf{Z}]$  reported in Section 3 are  
 177 the medians and the corresponding limits of the 95% confidence interval of the posterior distributions. We  
 178 preferred the median to the mean because it is more robust to skewed distributions.

## 179 3 Results

### 180 3.1 Estimating vital rates

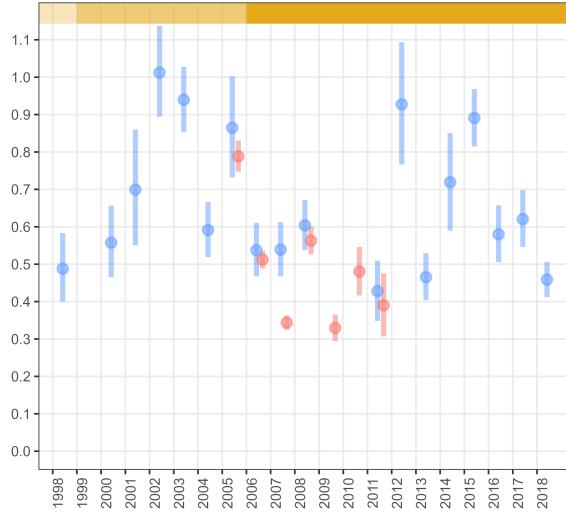
181 The “apparent sex ratio” method successfully provided estimates and associated uncertainties for the pro-  
 182 portion of immatures in both wintering populations. Adult survival and recruitment rates, i.e. the two  
 183 components that determine the population growth rate, were also correctly estimated (Figure 5).

### Proportion of immatures in the population



### Survival rate

#### Proportion of breeders still alive after one year



### Recruitment rate

#### Number of recruits produced per breeder

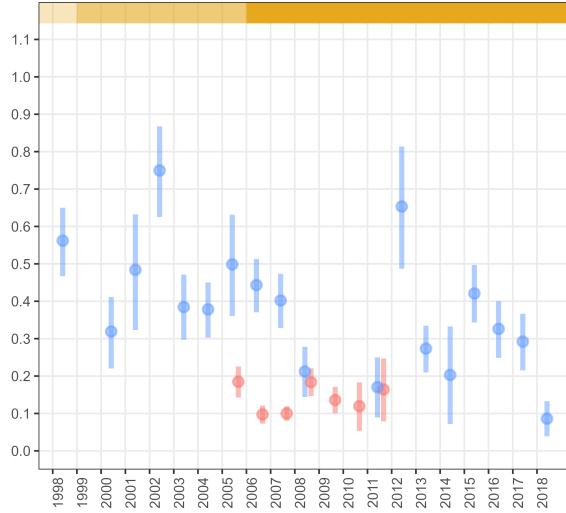


Figure 5: Temporal changes in the proportions of immatures and vital rates following culling pressure on populations of Ruddy ducks in Great Britain (GB) and France (FR); the vertical bars represent the 95% confidence intervals

184 The proportion of males among adults, which is a prerequisite for inferring the proportion of immatures,  
185 was estimated at 0.60 [0.59; 0.61]. The proportion of immatures ranged between 0.16 [0.07; 0.24] and 0.54  
186 [0.44; 0.62], depending on the population and year. The lowest values were similar in both populations, but  
187 the range of the proportion of immatures in the GB population was much smaller (upper value GB: 0.30  
188 [0.14; 0.45] & FR: 0.54 [0.44; 0.62]). The proportion of immatures in the GB population was stable over  
189 time, while a slight but significant decrease was observed in the FR population.

190 Adult survival rates ranged between 0.33 [0.29; 0.37] and 1.01 [0.89; 1.14]. None of the estimates were  
191 significantly outside the range of a survival rate defined without immigration [0; 1]. No trend in adult  
192 survival rate was observed in either population, but the patterns were different: adult survival in GB was  
193 lower and showed less interannual variability than in FR.

194 Recruitment rates ranged between 0.09 [0.04; 0.13] and 0.75 [0.63; 0.87]. All estimates were above 0, which is  
195 not inconsistent with reality. Furthermore, the maximum recruitment rate was 0.68 [0.36; 0.78] (see Section  
196 3.3), and no estimate was significantly outside the range of recruitment rates defined without immigration  
197 [0; 0.78]. No trend was observed for the GB population, but the recruitment rate decreased for the FR  
198 population, although it was more noisy than the proportion of immatures. As with the survival rate, the  
199 GB population also showed lower recruitment rate values with less variability than the FR population.

200 Lower values and lower variability of both adult survival and recruitment rates estimated in the GB popula-  
201 tion than in the FR population probably explain the different trends of the two Ruddy duck populations:  
202 a sharp and constant decline in the GB population versus a slow and variable decline in the FR population  
203 (Figure 1). In the FR population, the range of recruitment rates (min/max difference: 0.66) was larger than  
204 the range of survival rates (min/max difference: 0.58). Conversely, the range of survival rates (min/max  
205 difference: 0.46) in the GB population was much larger than the range of recruitment rates (min/max dif-  
206 ference: 0.08). This suggests that the variability in population growth rate in FR was mainly determined  
207 by recruitment, whereas in GB it was mainly determined by changes in adult survival.

### 208 3.2 Testing the reliability of the estimates

209 In Great Britain, the proportions of immatures derived from the counts were positively correlated with  
210 those derived from the culled individuals (Figure 6). This result therefore supports the hypothesis that  
211 the proportion of female-like individuals in the overwintering populations is efficient for the assessment of  
212 age structure and thus productivity. The correlation was even stronger when looking at vital rates (Figure  
213 6), but this was to be expected because the two methods for estimating these vital rates had a common  
214 component, namely the population growth rate (see Equation (3)). Interestingly, a 1:1 correlation was not  
215 achieved for any of the parameters. The proportion of immatures derived from the counts is always lower  
216 than the proportion derived from the culled individuals. The discrepancy between these two estimates is  
217 due to the fact that immatures are probably more vulnerable to culling than adults and may therefore be  
218 over-represented in the culling samples. As a result, the correlation coefficients for both vital rates also  
219 differed. However, the differences between the two approaches do not call into question the ability of the  
220 “apparent sex ratio” method to capture interannual variability of the vital rates.

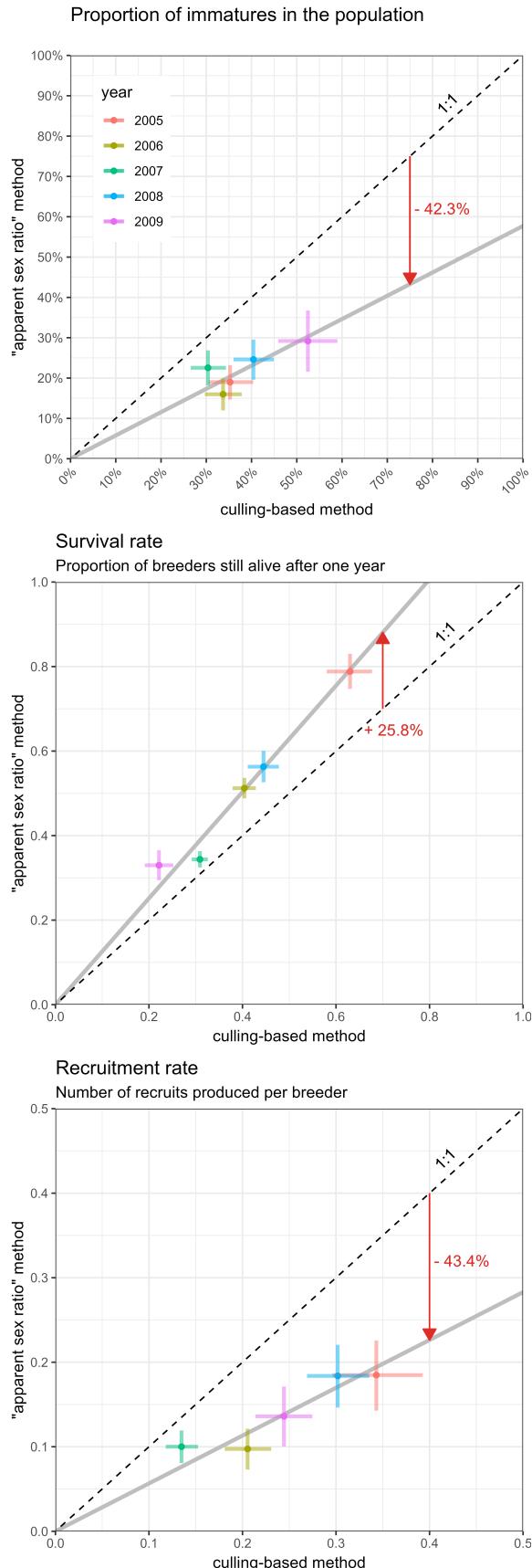


Figure 6: Relationship between parameter estimates obtained from counts (“apparent sex ratio” method) and those obtained from culled individuals (culling-based method); only five years were available for the GB population; the bars represent 95% confidence intervals and the red arrow indicates the direction and average differences between the two methods

### **221 3.3 Assessing the demographic response to culling strategies**

222 When there was no culling pressure, maximum growth rates were very similar for the two populations (Figure  
223 7), namely 1.45 [1.36; 1.55] and 1.52 [1.25; 1.88] for GB and FR, respectively. These values corresponded to  
224 an increase in the population of 45% and 52% per year, respectively. Assuming a maximum adult survival  
225 rate of 0.85 [0.7; 1] for both populations, we derived maximum recruitment rates of 0.60 [0.42; 0.78] and 0.68  
226 [0.36; 1.06] for GB and FR, respectively.

227 Despite the low number of sex- and age-determined culled individuals, the proxies of culling pressure showed  
228 significant changes over time (Figure 8). Therefore, we categorised culling pressure according to this intensity.  
229 High culling pressure occurred from 2006 onwards in both populations (although the signal was noisy).  
230 However, despite the similar level of culling pressure, the culling strategies of the two populations differed  
231 considerably, see Section 2.3. Under high culling pressure, the growth rate decreased to 0.56 [0.53; 0.59] and  
232 0.97 [0.94; 1.00] for GB and FR, corresponding to a decrease of 44% per year and 3% per year, respectively.  
233 This result shows that the GB population declined significantly during the high culling pressure, while the  
234 FR population was just stabilised.

235 Adult survival rates under high culling pressure reached 0.44 [0.42; 0.46] and 0.62 [0.59; 0.64] for GB and  
236 FR, respectively, and the recruitment rates decreased to 0.13 [0.11; 0.15] and 0.32 [0.29; 0.34], respectively.  
237 The resulting decrease in the growth rate of the GB population corresponded to a similar average decrease  
238 in adult survival rate (-0.41) and recruitment rate (-0.47), whereas the stabilisation of the FR population  
239 was due to a decrease in recruitment rate (-0.36) rather than a decrease in adult survival (-0.23). Overall,  
240 the culling strategy used in GB was more effective than that used in FR as it affected both adult survival  
241 and recruitment rates.

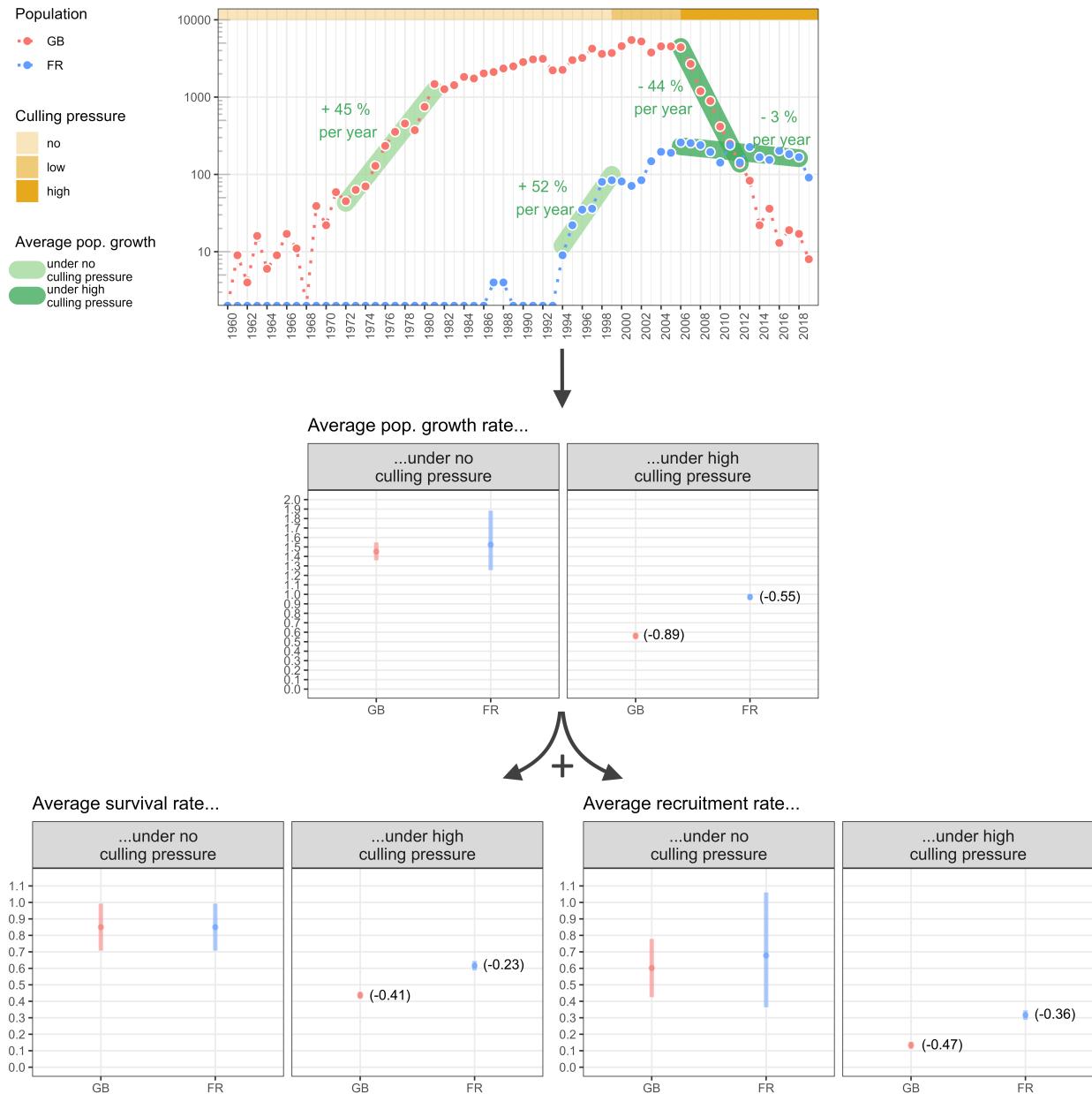


Figure 7: Average effects of culling pressure and different culling strategies (GB versus FR) on population growth rates (derived from counts only), and on adult survival and recruitment rates (derived from counts and reports of apparent sex ratios); the culling pressure for the GB population occurred mainly during the pre-breeding season, while for the FR population it was mainly concentrated during and after the breeding season (see 2.3 for details); adult survival and recruitment rates under high culling pressure were averaged over a period in which consistent culling pressure was observed; the vertical bars show 95% confidence intervals

## 242 4 Discussion

243 Disentangling the contribution of vital rates to population growth rate is a key step towards better un-  
 244 derstanding a population's response to management actions (Williams et al., 2002). We exploited delayed

245 sexual dichromatism to develop a new approach based on counts to disentangle the growth rate of a popu-  
246 lation into its two major structural components: adult survival and recruitment rates. The development of  
247 the “apparent sex ratio” method is based on a unique data set of two comparable populations of the same  
248 species, the Ruddy duck. Both populations were monitored in a similar way from their introduction to a  
249 period of intense culling pressure, but the culling strategies differed markedly. The very similar demographic  
250 trajectories and properties of both populations under similar ecological conditions made it possible to assess  
251 the response to both culling pressure and culling strategies.

## 252 4.1 Monitoring adult survival and recruitment rates

253 For both populations, vital rate estimates were within the same range, showing some consistency in the results  
254 obtained using the “apparent sex ratio” method. The greater variability observed for the FR population was  
255 to be expected as the time series covered a wider range of culling pressure than that of the GB population.  
256 We found no temporal correlation of demographic parameters between the two populations, suggesting that  
257 regional (e.g. weather) rather than large-scale factors (e.g. climate) were predominant. This is consistent  
258 with previous findings showing, for example, that in waterfowl both breeding success and juvenile survival  
259 depend on the onset of the laying period, which is closely linked to local weather (e.g. spring temperature,  
260 cumulative rainfall, water levels) (Blums & Clark, 2004; Dzus & Clark, 1998; Folliot et al., 2017) and local  
261 changes in predation pressure (Jaatinen et al., 2022).

262 The counts used to implement our model took place in mid-winter, at a time when some immature males  
263 may be initiating their moult. This could partially violate the assumption that the male-like individuals  
264 were exclusively adult males, and thus potentially bias our estimates. Nevertheless, as the survival estimates  
265 are in the range of those reported for similar species (Buderman et al., 2023; Buxton et al., 2004; Krementz  
266 et al., 1997; Nichols et al., 1997), this assumption appears to be reliable. Therefore, the time window for  
267 the count surveys must be carefully chosen to take into account the time before the immature males start  
268 moulting.

269 In some years, the estimated adult survival values were outside the range of expected values, although never  
270 significantly above 1. One possible explanation for these discrepancies is a violation of the “closed population”  
271 assumption. Although the ringing data suggest that Ruddy ducks do not undertake long-distance seasonal  
272 migrations in GB (Henderson, pers. comm.), the possibility that immigration occurred cannot be completely  
273 ruled out. The outliers in survival and recruitment rates observed in FR in 2002 and 2012 suggest that there  
274 was significant immigration in those years.

275 The theoretical maximum population growth rate is an inherent trait related to the demographic charac-  
276 teristics of a species (Dmitriew, 2010; Niel & Lebreton, 2005). This trait is expected to be similar in all  
277 Ruddy duck populations around the world. In our study, we observed that the maximum growth rates  
278 of the two populations were very close to each other, which was to be expected as both populations were  
279 exposed to similar ecological conditions, i.e. no harvest and similar breeding conditions. In addition, before  
280 the eradication programmes started, the two populations had population growth rates much closer to the  
281 expected maximum values than any other native duck species in the same areas during the same period.  
282 Prior to the eradication programmes, the survival rates of both populations were in the upper range of those  
283 observed for other diving duck species (e.g. ~ 0.8 for the Common pochard (Folliot et al., 2020; Nichols et  
284 al., 1997)). Predation on diving duck nests was found to decrease during the breeding season in a French  
285 fishpond complex (Bourdais et al., 2015). As Ruddy ducks generally breed later than Common pochards  
286 and Tufted ducks, both a higher survival rate and greater nesting success could be responsible for the high  
287 growth rates of Ruddy duck populations in the absence of culling. The “apparent sex ratio” method not  
288 only provided estimates of adult survival and recruitment rates that were consistent with observed increasing  
289 population sizes, but also provided estimates that were consistent with both declining phases following  
290 the implementation of eradication programmes and observed differences in population growth rates under  
291 different culling strategies. As expected, higher values for survival and recruitment rates were observed when  
292 the culling strategy focussed on the post-breeding season rather than the pre-breeding season.

293 Assuming that culling data provide a good picture of interannual variability in age structure (Fox et al., 2014),  
294 the strong correlation between estimates derived from counts and estimates derived from culled individuals

295 demonstrates the ability of the “apparent sex ratio” method presented here to properly capture interannual  
296 variation in demographic parameters such as survival and recruitment rates. The fact that the observed  
297 correlation was strong despite being based on a short time series strengthens its robustness. In addition,  
298 the strength of the relationship between the parameters obtained with the two approaches also indicates  
299 that important assumptions of our model were met, such as the constant adult sex ratio. The temporal  
300 autocorrelation of the adult sex ratio could be explained by the fact that it includes many age cohorts, which  
301 makes it structurally strong. In the long term, there may be substantial fluctuations in the adult sex ratio,  
302 but our results suggest that it is not necessary to monitor and update it annually. The proportion of males  
303 among adults estimated in the present study (i.e., 0.60) is within the range of values observed in native  
304 Ruddy duck populations of North America (0.62 in Bellrose (1980)) and in other duck species (Wood et al.,  
305 2021). The “apparent sex ratio” method consistently yielded recruitment rates that were almost a factor of  
306 two lower than the method based on culling data, a result consistent with higher vulnerability of immatures  
307 to shooting (Fox et al., 2014). In North America, for example, immature waterfowl were found to be 1.3 to  
308 2.6 times more vulnerable to hunting than adults (Bellrose, 1980).

309 As no estimates of adult survival were available in the literature (see Buderman et al., 2023), we were  
310 unable to properly assess the accuracy of the “apparent sex ratio” method for estimating this parameter.  
311 Nevertheless, there is evidence that the method is not highly biased, as described above. Thus, the “apparent  
312 sex ratio” method would be much more efficient for monitoring key demographic parameters than alternative  
313 methods such as those based on culling/hunting bag data, which have been shown to be unreliable (Fox et  
314 al., 2014), even when focussing on the end of the hunting season to limit potential bias (Fox et al., 2016).

## 315 4.2 Evaluating eradication strategies

316 Both populations responded to culling pressure with a significant decline in their growth rate, meaning that  
317 the increased mortality from culling was not compensated for by an increase in natural survival or improved  
318 breeding success. Culling was therefore effective in affecting the demographic trajectory of Ruddy duck  
319 in both cases. However, the decline in population growth was much greater in GB, where the population  
320 declined steadily and sharply, than in FR, where the increase was halted and the population size remained  
321 more or less stable. The greater decline in population growth observed in GB compared to FR, was due  
322 to a greater decline in adult survival and recruitment rates. Overall, adult survival and recruitment rates  
323 appeared to be significantly affected by culling in both countries. However, culling pressure targeting future  
324 breeders (pre-breeding culling strategy applied in GB), appeared to be much more effective in reducing both  
325 survival and recruitment rates than culling pressure targeting breeders and newly born individuals during  
326 the breeding and rearing season (breeding culling strategy applied in FR).

327 As expected, the pre-breeding culling strategy had the same effect on adult survival and recruitment rates,  
328 as a bird killed before the breeding season reduces the breeding population (lower adult survival rate) and  
329 prevents the bird from reproducing (lower recruitment rate). When using the breeding culling strategy,  
330 the recruitment rate was more affected than the adult survival rate. This was probably due to the fact  
331 that immatures are much more susceptible to being shot than adults (Bellrose, 1980; Fox et al., 2014).  
332 The compensation hypothesis states that the increased mortality of immatures due to harvesting may be  
333 compensated for by a higher natural survival (Cooch et al., 2014), but this hypothesis is not supported by  
334 the observations. Overall, there was no evidence of a compensatory mechanism in either context. This could  
335 be due to the fact that both populations were far from reaching carrying capacity, meaning that competition  
336 for resources was relaxed (Péron, 2013).

337 With more or less comparable culling pressure, the pre-breeding culling strategy proved to be much more  
338 effective. However, this does not mean that the breeding culling strategy was not also effective. On the  
339 contrary, it proved to be quite efficient in stopping both the population increase and the expansion of the  
340 distribution area, and in triggering a significant population decline in a second attempt by slightly increasing  
341 the culling pressure. So if winter culling is not possible, which was the case in FR, culling during the breeding  
342 season is effective in stopping or reversing the increase in populations of this species. However, the breeding  
343 culling strategy is compromised by the highly variable detectability of immatures, as these individuals colonise  
344 new ponds. This leads to fluctuations in the recruitment rate and thus jeopardises the long term efficiency

345 of this strategy. This problem does not exist with the pre-breeding strategy, as the Ruddy ducks congregate  
346 at a limited number of sites in winter (Johnsgard & Carbonell, 1996).

347 The response of the two Ruddy duck populations to culling pressure shows that it is necessary to take the  
348 culling period into account in order to make correct predictions of population trajectory. Predicting the  
349 effects of harvest pressure on a waterfowl population is then not only a question of the level of harvest,  
350 but also the timing of the harvest (Kokko et al., 1998). Unexpectedly, our results suggest that harvesting  
351 waterfowl during the breeding season has far less impact on population growth rate than harvesting in winter.  
352 Consequently, a target for effective management of harvested waterfowl could be an earlier start and end of  
353 the hunting season.

### 354 4.3 Implications for waterfowl management

355 An important implication of our study concerns the timing of harvest to limit its impact on populations  
356 growth rates. Unexpectedly, our results show that the persistence of waterfowl populations should benefit  
357 from earlier openings and closings of the hunting season (see above).

358 Tracking fluctuations of population abundance is a common tool for determining the conservation status of a  
359 population (e.g. Folliot et al., 2022). But tracking abundance alone does not provide enough information to  
360 assess the underlying mechanisms behind changes in population growth (Williams et al., 2002). This requires  
361 “digging deeper” (Austin et al., 2000), for example, by monitoring individuals to assess parameters such as  
362 survival and recruitment rates (e.g. Arnold, 2018). However, monitoring individuals is time-consuming  
363 (e.g. Souchay & Schaub, 2016) and not always possible for endangered species. The “apparent sex ratio”  
364 method makes it possible to circumvent these disadvantages in dichromatic species with delayed sexual  
365 maturity of the males. These species include the White headed duck in southern Spain, which is highly  
366 endangered and cannot be disturbed through capture-mark-recapture. Thus, by improving census strategies  
367 and disaggregating the effects of variations in adult survival and productivity on population growth using  
368 the approach presented here, much more invaluable information on the relevance of management strategies  
369 could be obtained.

370 Finally, we would like to emphasise that the main aim of our approach was to decompose population growth  
371 into its two main components, and not to provide unbiased estimates of adult survival or recruitment rates.  
372 However, if necessary, count surveys could be designed to do so, at least in theory, and for species that exhibit  
373 observable delayed dichromatism. This trait affects most species of stiff-tailed duck (Johnsgard & Carbonell,  
374 1996), as well as numerous other duck species, including Tufted duck, *Aythya fuligula*, Black scoter, *Melanitta*  
375 *nigra americana*, Common Goldeneye, *Bucephala clangula americana* (Bellrose, 1980; Johnsgard, 1978).  
376 Modifying standard monitoring protocols to distinguish between male-like and female-like individuals is  
377 almost costless, but worth the effort as it would greatly increase the efficiency of conservation/management  
378 actions (Nichols & Williams, 2006).

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## A Supplementary Materials

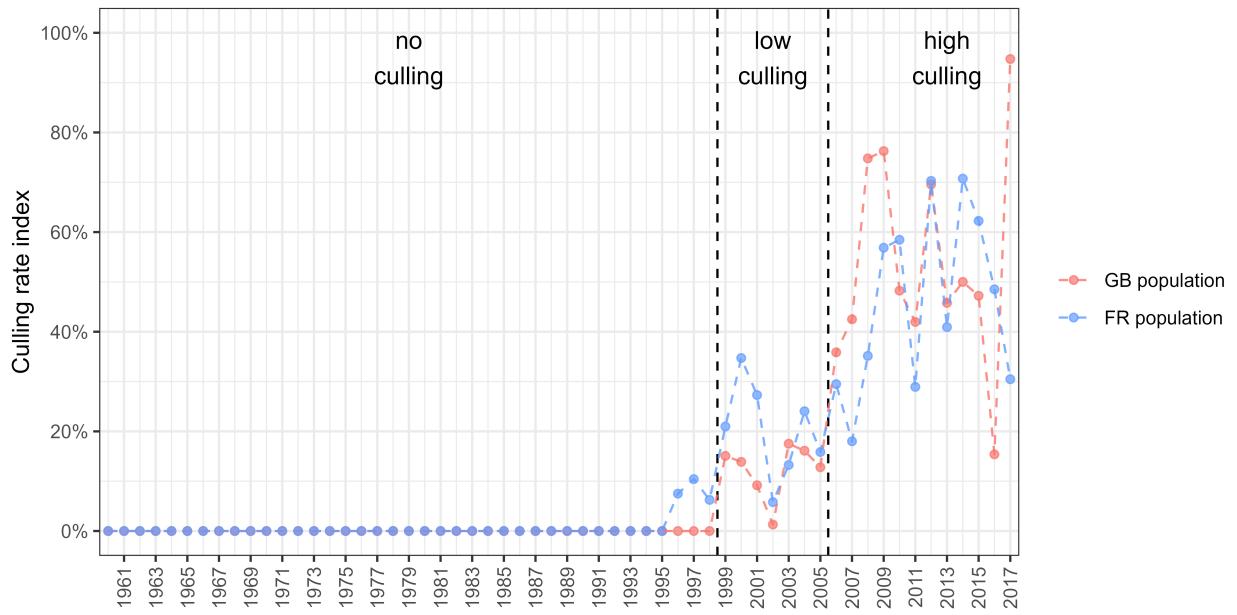


Figure 8: Definition of three periods with different culling pressure using an index based on an estimate of the adult culling rate; as only half of the culled individuals in France were age-determined before 2009, we applied the average age ratio over 2009-2019 to the remaining individuals in order to obtain an estimate of the number of adults in the culling data before 2009; the culling rate increased over time, so we divided the time series into three categories: we defined a “no culling” period before 1999 as the culling rate for both countries was mostly zero and always below 10%, we defined a “low culling” period between 1999 and 2005 as the culling rate for both countries fluctuated around 20%, the culling rate then increased significantly for both countries, therefore we defined a “high culling” period from 2006 onwards