

**Predictors of Height in Romanian Infants 6-23 Months Old:
Findings from a National Representative Sample**

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Background. The research explored the association between infants' height and various demographic factors in Romania, a country where such critical information has been lacking.

Methods. This study was conducted on a nationally representative sample and used a family physicians database to determine a sample of 1,532 children (713 girls and 819 boys) 6 to 23 months of age ($M = 14.26$; $SD = 5.15$). Infants' height-for-age z -scores (HAZ) were calculated using the World Health Organization's computing algorithm. A multiple regression analysis was conducted to investigate whether certain risk factors, such as infant mother's age, location, marital status, socioeconomic status (SES), as well as infant's term status at birth, age, anemia, minimum dietary diversity (MDD), and birth order, could significantly predict the HAZ.

Results. The study identified several significant predictors of height. Specifically, lower HAZ was associated with rural living, preterm birth, age 18-23 months, unmarried mothers, anemia, lack of MDD, and being third or later born in the family. In contrast, higher HAZ was associated with medium or high maternal socioeconomic status and older maternal age.

Conclusions. The study underscores the importance of addressing these significant risk factors through distinct interventions to improve height outcomes in at-risk Romanian populations.

Keywords: Romanian Infants, Height (Height-For-Age z -Scores), Public Health

Introduction

In public health domains, height-for-age z-scores (HAZ) are deemed a pivotal metric. These scores, which compare an infant's height to a reference population, are instrumental in detecting growth deviations such as stunting or accelerated growth. Notably, stunting often signals chronic malnutrition, with significant repercussions on an infant's cognitive and physical development. A myriad of factors, from genetic makeup and physical condition to nutrition, socio-economic factors, influence this growth pattern.^{1,2,3,4,5}

Romania, with its distinct socio-cultural and economic landscape, remains underrepresented in this discourse. While global insights offer a comprehensive perspective, regional nuances often demand specialized studies. This literature void prompts an essential query: What drives the height of Romanian infants aged 6 to 23 months? This paper sought to unravel this question.

Drawing from a nationally representative sample, our study endeavored to pinpoint the primary factors influencing the height of Romanian infants in the specified age bracket. Beyond augmenting existing knowledge, our aim was to furnish actionable insights for targeted interventions benefiting at-risk groups, thereby aiding Romania's policymakers and healthcare practitioners. Our objective was to investigate the relationship between infants' height and various demographic parameters, including the mother's location, marital status, socioeconomic status, age, and infant-specific factors such as being born at term, age, anemia condition, birth order, and dietary patterns, specifically minimum dietary diversity (MDD).

Methods

Participants

This study utilized data from an extensive cross-sectional survey carried out by Stativa et

al.⁶ in the latter half of 2010. The overarching study, titled “Assessment of the Current Situation of Breastfeeding and Nutrition Practices for Children from Birth to Two Years Old,” aimed to enhance the access, quality, and efficiency of medical services for pregnant women and infants up to two years of age. The first segment of the study assessed the national iron supplementation program, while the second delved into breastfeeding and nutritional practices among infants, identifying factors that influence these practices. The insights garnered from this research informed the creation and execution of policies and programs to bolster the nutritional status of infants nationwide. Prior to the study’s commencement, all participating parents (specifically the children’s mothers) provided written consent to the National Institute for Mother and Child Health “Alessandrescu-Rusescu” in Bucharest, Romania. Data from this primary study was also employed in a subsequent article that scrutinized the feeding patterns of infants aged 6-23 months.⁶

For this secondary analysis, our sample comprised 1,532 children (713 girls and 819 boys) aged between 6 to 23 months ($M = 14.26$; $SD = 5.15$).

Procedure

The study employed a stratified multistage probabilistic sampling strategy. The population was segmented based on geographical regions, encompassing the eight macro-development regions in Romania, and further categorized by residence (urban vs. rural). Within these strata, participants were grouped into age brackets: 0-3 months, 3-6 months, 6-11 months, 11-13 months, 13-18 months, 18-23 months, and 23-25 months. This approach aimed to account for potential regional variations in infant nutrition practices based on age. From each region, two counties were chosen. Within these counties, two urban and three rural areas were randomly selected, using the 2009 Statistical Yearbook of Romania as a reference. For the Bucharest-Ilfov

region, two sectors (districts) from six, and an additional urban locality outside Bucharest were chosen, along with two rural areas.

Medical inspectors and nurses from the County Directorates of Public Health collaborated with research teams to aid in the selection of family doctors from both urban and rural localities. In urban areas, 4-5 family doctors were selected. In contrast, for rural areas, while considering the existing family doctors, a maximum of 2 were recruited for the study. Family doctors were instructed to randomly select a certain number of children from all age groups defined, with the reference age being the number of months completed until the date of the visit of the research teams. Doctors were instructed to select and reserve children in case they did not show up. For children selected in the sample, nurses telephoned families about the research and asked for their verbal consent to participate. Written consent was obtained at the scheduled presentation at the family doctor's office. The research teams verified on site their compliance with the requirements for randomized selection and noted, if necessary, that the child was not randomly selected and was excluded from the data analysis.

Interviews were conducted with mothers who brought the children chosen for the sample. A subset of the 17-member team, specifically 2-3 trained and seasoned researchers, carried out these interviews. Following the interviews, these researchers took blood samples from the children and recorded their height and weight measurements. At the family doctors' clinics, mothers were inquired about their prenatal consultation experiences and their use of iron prophylaxis during the pregnancy under investigation. Additional inquiries covered topics such as the mother's initiation, understanding, and practices related to breastfeeding; child feeding habits; maternal smoking history; and the use of iron and vitamin D prophylaxis.^{7,8} The data was collected between July and November 2010.

The study encompassed 2,117 children spanning 16 counties, effectively covering all eight regions of the nation. Of this cohort, 51% (1,080) were from urban settings, while 49% (1,037) hailed from rural locales. Age-wise, 1,002 children (47.3%) were between 0-12 months, and 1,115 (52.7%) were older than a year.

Variables

The dependent variable used in this study was child HAZ, calculated using the World Health Organization's algorithm.⁸

Several demographic variables related to mothers served as independent variables in our study. These included the mother's age (in years) and categorical variables such as location (urban vs. rural), marital status (categorized as married vs. others, with "others" encompassing cohabitation, divorced/separated/widow, and unmarried). SES of the mother was also considered, segmented into low, medium, or high. SES determination was based on the ownership of 12 items, ranging from basic household appliances to luxury goods. Ownership of up to four items was classified as low SES, five to eight items as medium SES, and ownership of nine or more items indicated high SES.

Infant characteristics were also factored in as categorical independent variables. These encompassed the infant's age (grouped as 6-11 months, 12-17 months, or 18-23 months), term status at birth (categorized as 37-40 weeks [at term] vs. 27-36 weeks [preterm]), birth order (firstborn, second-born, or third and subsequent), and hemoglobin level (no anemia vs. anemia). For hemoglobin determination, capillary blood samples were taken from presumed healthy infants using a finger-prick method. Hemoglobin concentration was measured with a portable hemoglobinometer (HemoCue), and parents were subsequently informed about their child's iron status. Hemoglobin concentration was categorized into two levels: no anemia (>11.0 g/dl) and

anemia (<11 g/dl).

Furthermore, the study incorporated the World Health Organization's^{9,10} minimum dietary diversity (MDD) for analysis. MDD represents the proportion of children aged 6–23 months who consumed food from four or more of the seven designated food groups the previous day. These groups included (1) grains, roots, and tubers; (2) legumes and nuts; (3) dairy products (e.g., milk, yogurt, cheese); (4) flesh foods (e.g., meat, fish, poultry, and liver/organ meats); (5) eggs; (6) vitamin-A rich fruits and vegetables; and (7) other fruits and vegetables. MDD was dichotomized, with category 1 indicating adherence to the complementary feeding indicator and category 2 indicating non-adherence.

Analyses

All analyses, including accuracy of data coding and entry and the statistical assumptions of the tests, were conducted using SPSS¹¹ and JASP¹². The analyses were conducted with all available data (pairwise deletion) and no imputation was conducted due to the small percentage of missing cases. There are several compelling reasons to choose this technique. Primarily, it excels in harnessing the power of in-depth data analytics, guaranteeing the most effective use of available information. Especially in scenarios with sparse missing data across multiple variables, pairwise deletion becomes invaluable, minimizing drastic reductions in sample size. When data is missing entirely at random, pairwise deletion produces parameter estimates that are closely aligned with those obtained from a complete dataset analysis.^{13,14} A multiple regression (forced entry method) was used to determine the best predictors of children's HAZ. The forced entry method in multiple regression is favored for its theoretical justification, simplicity, reduced overfitting risk, model stability, and potential for greater generalizability. Its usage hinges on the researcher's theoretical foundation and data characteristics.¹⁵

Results

Table 1 presents the characteristics of the sampled infants and their mothers. Of the 1,532 infants, 819 (53.5%) were male, and 593 (38.7%) were aged between 6-11 months. A majority, 789 (51.5%), lived in urban areas with their mothers. Most mothers identified as Romanian (84.3%), were married (81.4%), and fell into the medium socio-economic status bracket (48.4%). Additionally, over half of the children were first-born (54.8%) and were born at term (89.9%). Notably, 49.9% of the infants had anemia, 76% met the criteria for minimum dietary diversity (MDD), and 54.8% were the eldest in their families. The multivariate linear regression (see Table 2) demonstrated that the independent variables had a significant negative and positive impact on the dependent variable. The overall result of the regression (R^2) indicated that 9% of the variance in the data was explained by the predictor variables [$F(9,1522) = 16.125, p < 0.001$]. A relatively low R^2 value suggests that a significant proportion of the variability in the dependent variable remains unexplained by the current model. This indicates that the model may not encompass all relevant predictors. Factors such as genetic makeup, specific medical conditions, or other pertinent socio-economic conditions might play a crucial role in influencing HAZ.

On average, infants living in rural areas had significantly lower HAZ ($B = -0.30, p = 0.001$) compared to those in urban areas. Infants not born at term had significantly lower HAZ ($B = -0.65, p < 0.001$) compared to those born at term. Infants aged 18-23 months had significantly lower HAZ ($B = -0.54, p < 0.001$) compared to those aged 6-11 months. Infants with unmarried mothers had significantly lower HAZ ($B = -0.24, p = 0.032$) compared to those with married mothers. Infants with anemia had significantly lower HAZ ($B = -0.29, p < 0.001$) compared to those without anemia. Infants without MDD had significantly lower HAZ ($B = -0.23, p = 0.015$)

compared to those with MDD. Infants who were third-born or later had significantly lower HAZ ($B = -0.54, p < 0.001$) compared to firstborn children.

Conversely, children whose mothers had a medium SES had significantly higher HAZ ($B = 0.31, p = 0.013$) compared to those with mothers of low SES, and this was even higher for children whose mothers had a high SES ($B = 0.41, p = 0.003$). Children with older mothers had significantly higher HAZ ($B = 0.02, p = 0.018$) compared to those with younger mothers.

Discussion

Main Finding of This Study

While the study reaffirms certain established knowledge, it offers nuanced insights into the determinants of infant HAZ within the Romanian demographic. The research highlighted that Romanian infants' HAZ is influenced by a myriad of factors. Specifically, residing in rural areas, being born preterm, being aged 8-23 months, having unmarried mothers, suffering from anemia, lacking sufficient MDD, and being third-born or later were all associated with decreased HAZ. Conversely, infants with mothers from a higher SES background or of advanced age exhibited enhanced HAZ.

What Is Already Known on This Topic

Geographical disparity. The relationship between geographical location and height growth is complex and not fully understood. Therefore, geographical disparity should be interpreted with caution, as some studies support^{16,17} and others reject^{18,19} the claim that infants raised in urban areas have greater exposure to resources that increase height growth compared to infants raised in rural areas.

Born at Term. Literature on prematurely born infants presents mixed findings. While many exhibit low birth weight and reduced HAZ, they often undergo rapid growth in their first

year, aligning with or even surpassing the growth rates of full-term infants.²⁰ Gong et al.²¹ suggest preterm infants often surpass the growth rates of full-term infants and Barreto and colleagues²² indicate that, despite starting with a lower birth size, many preterm infants experience a growth spurt between months 6-23, reaching typical sizes. Yet, Finken et al.²³ contend that preterm children largely develop at standard rates.

Infants' Age. Several studies indicate that infants aged 18-23 months are more prone to stunting in height growth compared to their younger counterparts. Specifically, one Marriot and colleagues¹ highlighted that over half of the children in the 18–23 months age group exhibited signs of stunting. A study of Burundian²⁴ children under two found that 58% of children aged five and below are affected by stunting. Older infants, specifically 12-17 and 18-23 months, were at higher risk, as were infants with low birth weight, poor nutritional assessments, and mothers with low education levels. Additionally, 79.3% of stunted infants were between the ages of 12-23 months. A growing body of evidence suggests that stunting can be attributed to a variety of factors, including the rapid growth spurt that occurs during this age, the introduction of complementary foods, and an increased risk of infection.^{2,3,4}

Marital status. Reurings et al.²⁵ found that maternal marital status is closely tied to HAZ. Specifically, Amadu and colleagues²⁶ suggests that children in single-mother households using clean cooking fuel, and those in married households using unclean fuel, face higher stunting risks than children in married households with clean fuel. Female children in households with married, educated parents, non-pregnant maternal figures, and water-sealed toilets are less malnutrition-prone.²⁷ Infants with unmarried mothers often show more stunting than those with married mothers.^{28,29} Yet, Young and Declercq³⁰ showed children from two-parent households generally have better height growth, regardless of marital status. While maternal marital status

can influence growth, it is not the primary factor, especially for unmarried mothers, with education often being a key determinant.³¹

Iron deficiency. Results consistent with other studies have been reported.^{32,33} Stativa and colleagues³⁴ conducted a study using a nationally representative sample in Romania, evaluating 1,532 infants aged between six months and two years. They discovered that 46% of these infants were anemic. The high prevalence of anemia among Romanian children has become a significant concern for specialists in the country, especially considering its potential long-term developmental consequences.

Feeding patterns. Diet diversity significantly predicts stunting in developing infants. Ahmad et al.³⁵ revealed that access to a diverse diet reduces stunting risk by 83%. In research conducted by Khamis and colleagues³⁶ and involving 2,960 children, it was found that 74% did not meet the MDD criteria, with 31% experiencing stunting, 6% wasting, and 14% being underweight. However, some studies^{37,38} found no significant impact of MDD on infant height growth or stunting. Complementary feeding is recommended around five months, but only 66.4% of infants in northern Ghana received it on time³⁸. Assisted feeding practices had a negative impact on growth, especially in the 6-11 months and 12-17 months age ranges³⁹. Although infants aged 18-23 months may experience stunted growth, their mortality rates remain largely unaffected⁴⁰.

Birth order. Firstborns typically exhibit greater height, with the HAZ disparity doubling for children born third or later when siblings are spaced less than three years apart.⁴¹ Yet, this difference diminishes with a three-year or larger gap. Families surpassing the perceived ideal size risk adverse growth effects in children.⁴² In contrast, in a study where weight, HAZ, and hemoglobin levels were assessed, no correlation was found between birth order and ideal family

size.

Socio-economic status. Devakumar et al.⁴³ indicated a positive link between HAZ and weight-for-age z-scores concerning stunting and underweight infants, particularly when considering SES. However, an infant mother's SES can sometimes yield contrasting effects.⁴⁴ Specifically, mothers with lower SES often prioritize breastfeeding and infant care, unlike higher SES mothers who may be career focused. Contrarily, Palanichamy and Solanki⁴⁵ suggest that low SES affects weight more than height, challenging the prevalent belief of its impact on stunting.

Age of mothers. Yu et al.⁴⁶ as well as Wemakor et al.⁴⁷ links low HAZ to young maternal ages. Yet, when considering factors like maternal height and SES, these findings vary. Stunted growth rates in children are reported as 9% in Asia, 14% in Africa, and 10% in Latin America, regardless of maternal age.⁴⁶ Children over 24 months with younger mothers consistently show poor growth. Infants of mothers under 20 are eight times more prone to stunting, three times to wasting, and 13 times to being underweight.^{46,48} Conversely, infants of mothers over 40 have low HAZ at birth, but this doesn't persist.^{49,50}

What This Study Adds

This study uniquely centers on Romania, a country often overlooked in global HAZ research. Given Romania's unique socio-cultural and economic context, prior research missed specific regional HAZ determinants. Our sampling strategy ensured national representation, encompassing all eight of Romania's macro-development regions and differentiating urban from rural areas.

Limitations Of This Study

While HAZ is commonly used to describe child growth, its suitability for longitudinal

growth tracking has been questioned due to its reliance on cross-sectional standard deviations. Authors⁵⁵ suggest using HAZ differences for linear growth measurement. This study incorporates socio-economic status but uses a non-validated categorization based on owned possessions. While suitable for Romania, the lack of validation may limit broader applicability. Future studies should validate this approach for Romania. The study's cross-sectional design precludes inferring causal relationships between variables.

Conclusion

This study delves into HAZ determinants for Romanian children, a topic previously under-researched. It enhances our understanding of factors leading to stunting. The findings underscore the need for age-specific interventions, promoting maternal education, improving nutrition, and addressing socio-economic disparities. By targeting these areas, we can reduce stunting, benefiting children's health in Romania and similar settings.

References

1. Marriott BP, White AJ, Hadden L, Davies JC, Wallingford JC. How well are infant and young child World Health Organization (WHO) feeding indicators associated with growth outcomes? An example from Cambodia. *Matern & Child Nutr.* 2010;6(4):358-373. doi:<https://doi.org/10.1111/j.1740-8709.2009.00217.x>
2. Aryastami, N.K., Shankar, A., Kusumawardani, N. et al. Low birth weight was the most dominant predictor associated with stunting among children aged 12–23 months in Indonesia. *BMC Nutr* 3, 16 (2017). <https://doi.org/10.1186/s40795-017-0130-x>
3. Gebreyohanes M, Dessie A. Prevalence of stunting and its associated factors among children 6-59 months of age in pastoralist community, Northeast Ethiopia: A community-based cross-sectional study. *PLoS One.* 2022 Feb 3;17(2):e0256722. doi: 10.1371/journal.pone.0256722. PMID: 35113874; PMCID: PMC8812981.
4. Nshimiyiryo, A., Hedt-Gauthier, B., Mutaganzwa, C. et al. Risk factors for stunting among children under five years: a cross-sectional population-based study in Rwanda using the 2015 Demographic and Health Survey. *BMC Public Health* 19, 175 (2019). <https://doi.org/10.1186/s12889-019-6504-z>
5. Phuphaibul R, Kongsaktrakul C, Phusamon S, Peasue N, Mosuwan L, Choprapawon C. Socioeconomic determinants of infant growth: The Perspective Cohort Study of Thai Children. *Jpn J Nurs Sci.* 2014;11(1):16-22. doi: 10.1111/j.1742-7924.2012.00225.x. Epub 2012 Aug 30. PMID: 24460598.
6. Stativa E, Stoicescu S (eds). Evaluarea situatiei curente a alaptarii si a practicilor de nutritie a copiilor de la nastere la 2 ani (The Assessment of the Current Situation of Breastfeeding and Nutrition Practices for Children from Birth to Two Years Old). Bucharest, Romania: Institute for Mother and Child Care ‘Alfred Rusescu’, Ministry of

- Health, *UNICEF*. 2011.
7. Stativa E, Rus AV, Stoicescu SM, et al. Feeding practices patterns in Romanian infants 6-23 months old: Findings from a national representative sample. *Health, Sports & Rehabilitation Medicine*. 2021;22:75–81. doi: 10.26659/pm3.2021.22.2.75
 8. de Onis M, Onyango AW, Borghi E, Garza C, Yang H; WHO Multicentre Growth Reference Study Group. Comparison of the World Health Organization (WHO) Child Growth Standards and the National Center for Health Statistics/WHO international growth reference: implications for child health programmes. *Public Health Nutr*. 2006;9(7):942-947. doi:10.1017/phn20062005
 9. WHO, UNICEF, IFPRI, UCDavis, FANTA, AED, USAID. Indicators for assessing infant and young child feeding practices. Part 1: Definitions. Geneva, *World Health Organization*. 2008. http://whqlibdoc.who.int/publications/2008/9789241596664_eng.pdf (13 January 2014, date last accessed).
 10. WHO, UNICEF/IFPRI, UCDavis, FANTA, AED, USAID (2010). Indicators for assessing infant and young child feeding practices. Part 2: Measurement. Geneva, *World Health Organization*. http://whqlibdoc.who.int/publications/2010/9789241599290_eng.pdf (13 January 2014, date last accessed).
 11. IBM Corp. *IBM SPSS Statistics for Windows (Version 19.0)*. Armonk, NY: IBM Corp, 2010.
 12. JASP Team. *JASP (Version 0.17.2)*[Computer software], 2023

13. Kang H. The prevention and handling of the missing data. *Korean J Anesthesiol.* 2013 May;64(5):402-6. doi: 10.4097/kjae.2013.64.5.402. Epub 2013 May 24. PMID: 23741561; PMCID: PMC3668100.
14. Shi D, Lee T, Fairchild AJ, Maydeu-Olivares A. Fitting Ordinal Factor Analysis Models With Missing Data: A Comparison Between Pairwise Deletion and Multiple Imputation. *Educ Psychol Meas.* 2020 Feb;80(1):41-66. doi: 10.1177/0013164419845039. Epub 2019 Apr 26. PMID: 31933492; PMCID: PMC6943991.
15. Tabachnick BG, Fidell LS. Using multivariate statistics. Boston: *Pearson*; 2013.
16. Nguyen HT, Eriksson B, Nguyen LT, et al. Physical growth during the first year of life. A longitudinal study in rural and urban areas of Hanoi, Vietnam. *BMC Pediatr.* 2012;12:26. Published 2012 Mar 12. doi:10.1186/1471-2431-12-26
17. Vaktskjold A, Văn Trí D, Phi DN, Sandanger T. Infant growth disparity in the Khanh Hoa province in Vietnam: a follow-up study. *BMC Pediatr.* 2010;10:62. Published 2010 Aug 23. doi:10.1186/1471-2431-10-62
18. Demilew YM, Tafere TE, Abitew DB. Infant and young child feeding practice among mothers with 0-24 months old children in Slum areas of Bahir Dar City, Ethiopia. *Int Breastfeed J.* 2017;12:26. Published 2017 Jun 14. doi:10.1186/s13006-017-0117-x
19. Sharma N, Gupta M, Aggarwal AK, Gorle M. Effectiveness of a culturally appropriate nutrition educational intervention delivered through health services to improve growth and complementary feeding of infants: A quasi-experimental study from Chandigarh, India. *PLoS One.* 2020;15(3):e0229755. Published 2020 Mar 17. doi:10.1371/journal.pone.0229755

20. Kang L, Wang H, He C, et al. Postnatal growth in preterm infants during the first year of life: A population-based cohort study in China. *PLoS One*. 2019;14(4):e0213762.
Published 2019 Apr 11. doi:10.1371/journal.pone.0213762
21. Gong YH, Ji CY, Shan JP. A longitudinal study on the catch-up growth of preterm and term infants of low, appropriate, and high birth weight. *Asia Pac J Public Health*. 2015;27(2):NP1421-NP1431. doi:10.1177/1010539513489129
22. Barreto GMS, Balbo SL, Rover MS, Toso BRGO, Oliveira HR de, Viera CS. Growth and biochemical markers of preterm newborns up to six months of corrected age. *J Hum Dev*. 2018;28(1):18. doi:https://doi.org/10.7322/jhgd.138687
23. Finken MJ, Dekker FW, de Zegher F, Wit JM; Dutch Project on Preterm and Small-for-Gestational-Age-19 Collaborative Study Group. Long-term height gain of prematurely born children with neonatal growth restraint: parallellism with the growth pattern of short children born small for gestational age. *Pediatrics*. 2006;118(2):640-643.
24. Nkurunziza S, Meessen B, Van Geertruyden JP, Korachais C. Determinants of stunting and severe stunting among Burundian children aged 6-23 months: evidence from a national cross-sectional household survey, 2014. *BMC Pediatr*. 2017;17(1):176.
Published 2017 Jul 25. doi:10.1186/s12887-017-0929-2
25. Reurings M, Vossenaar M, Doak CM, Solomons NW. Stunting rates in infants and toddlers born in metropolitan Quetzaltenango, Guatemala. *Nutrition*. 2013;29(4):655-660. doi:10.1016/j.nut.2012.12.012
26. Amadu I, Seidu AA, Duku E, et al. The Joint Effect of Maternal Marital Status and Type of Household Cooking Fuel on Child Nutritional Status in Sub-Saharan Africa: Analysis

- of Cross-Sectional Surveys on Children from 31 Countries. *Nutrients*. 2021;13(5):1541.
Published 2021 May 3. doi:10.3390/nu13051541
27. Ramirez MAR, Ducay AJ. Determinants of normal nutrition among 0-59-month-old Filipino children living in low-income households. *Nutr Health*. 2021;27(4):423-434.
doi:10.1177/0260106021992669
 28. Bamanikar AA, Shah S, Aboudi D, et al. Impact of paternal presence and parental social-demographic characteristics on birth outcomes. *J Perinat Med*. 2021;49(9):1154-1162.
Published 2021 Aug 5. doi:10.1515/jpm-2021-0078
 29. Liczbińska G, Králík M. The strong impact of maternal marital status on birth body size before and during the Second World War in Poznań district, Poland. *Am J Hum Biol*. 2022;34(5):e23707. doi:10.1002/ajhb.23707
 30. Young RL, Declercq E. Implications of subdividing marital status: are unmarried mothers with partners different from unmarried mothers without partners? An exploratory analysis. *Matern Child Health J*. 2010;14(2):209-214. doi:10.1007/s10995-009-0450-9
 31. Ngandu CB, Momberg D, Magan A, Norris SA, Said-Mohamed R. Association Between Household and Maternal Socioeconomic Factors with Birth Outcomes in the Democratic Republic of Congo and South Africa: A Comparative Study. *Matern Child Health J*. 2021;25(8):1296-1304. doi:10.1007/s10995-021-03147-x
 32. Pasricha SR, Hayes E, Kalumba K, Biggs BA. Effect of daily iron supplementation on health in children aged 4-23 months: a systematic review and meta-analysis of randomised controlled trials [published correction appears in Lancet Glob Health. 2014 Mar 2(3):e144]. *Lancet Glob Health*. 2013;1(2):e77-e86. doi:10.1016/S2214-109X(13)70046-9

33. Smuts CM, Dhansay MA, Faber M, et al. Efficacy of multiple micronutrient supplementation for improving anemia, micronutrient status, and growth in South African infants. *J Nutr*. 2005;135(3):653S-659S. doi:10.1093/jn/135.3.653S
34. Stativa E, Rus AV, Stanescu A, Pennings JS, Parris SR, Wenyika R. Prevalence and predictors of anaemia in Romanian infants 6-23 months old. *J Public Health (Oxf)*. 2016;38(3):e272-e281. doi:10.1093/pubmed/fdv145
35. Ahmad I, Khalique N, Khalil S, Urfi, Maroof M. Dietary Diversity and Stunting among Infants and Young Children: A Cross-sectional Study in Aligarh. *Indian J Community Med*. 2018;43(1):34-36. doi:10.4103/ijcm.IJCM_382_16
36. Khamis AG, Mwanri AW, Ntwenya JE, Kreppel K. The influence of dietary diversity on the nutritional status of children between 6 and 23 months of age in Tanzania. *BMC Pediatr*. 2019;19(1):518. Published 2019 Dec 28. doi:10.1186/s12887-019-1897-5
37. Reinbott A, Kuchenbecker J, Herrmann J, et al. A child feeding index is superior to WHO IYCF indicators in explaining length-for-age Z-scores of young children in rural Cambodia. *Paediatr Int Child Health*. 2015;35(2):124-134. doi:10.1179/2046905514Y.0000000155
38. Anin SK, Saaka M, Fischer F, Kraemer A. Association between Infant and Young Child Feeding (IYCF) Indicators and the Nutritional Status of Children (6-23 Months) in Northern Ghana. *Nutrients*. 2020;12(9):2565. Published 2020 Aug 24. doi:10.3390/nu12092565
39. Briaux J, Fortin S, Kameli Y, et al. Dissimilarities across age groups in the associations between complementary feeding practices and child growth: Evidence from rural Togo. *Matern Child Nutr*. 2019;15(4):e12843. doi:10.1111/mcn.12843

40. Hussaini KS, Ritenour D, Coonrod DV. Interpregnancy intervals and the risk for infant mortality: a case control study of Arizona infants 2003-2007. *Matern Child Health J.* 2013;17(4):646-653. doi:10.1007/s10995-012-1041-8
41. Dhingra S, Pingali PL. Effects of short birth spacing on birth-order differences in child stunting: Evidence from India. *Proc Natl Acad Sci U S A.* 2021;118(8):e2017834118. doi:10.1073/pnas.2017834118
42. Costa ME, Trumble B, Kaplan H, Gurven MD. Child nutritional status among births exceeding ideal family size in a high fertility population. *Matern Child Nutr.* 2018;14(4):e12625. doi:10.1111/mcn.12625
43. Devakumar D, Kular D, Shrestha BP, et al. Socioeconomic determinants of growth in a longitudinal study in Nepal. *Matern Child Nutr.* 2018;14(1):e12462. doi:10.1111/mcn.12462
44. Brush G, Harrison GA, Zumrawi FY. A path analysis of some determinants of infant growth in Khartoum. *Ann Hum Biol.* 1993;20(4):381-387. doi:10.1080/03014469300002782
45. Palanichamy M, Solanki MJ. Infant and Child Feeding Index and nutritional status of children aged 6 to 24 months in a Metropolitan city. *J Family Med Prim Care.* 2021;10(1):175-181. doi:10.4103/jfmprc.jfmprc_1023_20
46. Yu SH, Mason J, Crum J, Cappa C, Hotchkiss DR. Differential effects of young maternal age on child growth. *Glob Health Action.* 2016;9:31171. Published 2016 Nov 15. doi:10.3402/gha.v9.31171

47. Wemakor A, Garti H, Azongo T, Garti H, Atosona A. Young maternal age is a risk factor for child undernutrition in Tamale Metropolis, Ghana. *BMC Res Notes*. 2018;11(1):877. Published 2018 Dec 10. doi:10.1186/s13104-018-3980-7
48. Nguyen PH, Sanghvi T, Tran LM, et al. The nutrition and health risks faced by pregnant adolescents: Insights from a cross-sectional study in Bangladesh. *PLoS One*. 2017;12(6):e0178878. Published 2017 Jun 8. doi:10.1371/journal.pone.0178878
49. Braggion A, Favre G, Lepigeon K, Sichitiu J, Baud D, Desseauve D. Advanced Maternal Age Among Nulliparous at Term and Risk of Unscheduled Cesarean Delivery [published online ahead of print, 2023 Apr 14]. *Am J Obstet Gynecol MFM*. 2023;100972. doi:10.1016/j.ajogmf.2023.100972
50. Workicho A, Belachew T, Argaw A, et al. Adolescent pregnancy and linear growth of infants: a birth cohort study in rural Ethiopia. *Nutr J*. 2019;18(1):22. Published 2019 Apr 2. doi:10.1186/s12937-019-0448-0