

Task B.1

Let $\{x_t\}_{t \in \{1, \dots, T\}}$ be the returns of the Apple stocks.

Assuming they follow a $AR(1)$ - T - $GARCH(2,2)$ the data generating process is

$$x_{t+1} = \hat{x}_{t+1} + \cancel{\sigma_{t+1}} \varepsilon_{t+1}$$

where we model $\hat{x}_{t+1} := \phi_0 + \phi_1 x_t$, ^{i.e.} the conditional mean of x_{t+1} $E[x_{t+1} | \mathcal{F}_t]$ as an $AR(1)$ -process. \mathcal{F}_t is the information set available at time t .

We model the conditional variance of ε_{t+1} , $V(\varepsilon_{t+1} | \mathcal{F}_t)$ with the T - $GARCH(2,2)$, i.e.

$\varepsilon_{t+1} = \hat{\sigma}_{t+1} \eta_{t+1}$ where η_{t+1} ^{follows} a white noise process, i.e. $WNP(0,1)$ with variance of 1 and mean 0 and ~~and~~.

$$\begin{aligned} \hat{\sigma}_{t+1}^2 &= V(\varepsilon_{t+1} | \mathcal{F}_t) = \\ &\alpha_0 + (\alpha_1 + \theta_1 \mathbb{1}_{(\varepsilon_t < 0)}) \cdot \varepsilon_t^2 + (\alpha_2 + \theta_2 \mathbb{1}_{(\varepsilon_{t-1} < 0)}) \varepsilon_{t-1}^2 \quad (*) \\ &\quad + \beta_1 \hat{\sigma}_t^2 + \beta_2 \hat{\sigma}_{t-1}^2. \end{aligned}$$

$\mathbb{1}$ is the indicator function: $\mathbb{1}_{(\varepsilon_t < 0)} = \begin{cases} 1 & \varepsilon_t < 0, \\ 0 & \text{sonst.} \end{cases}$

We can assume that the mean cond. mean of ε_{t+1} is 0, because of the $AR(1)$ term \hat{x}_{t+1} .

Now for the joint log-likelihood function:

We need to assume some distribution of η . Let f be the probability density function of η . It must be a WMP(0,1).

The following describes the log-likelihood function $l(\lambda|X)$ where $X = \{x_t\}_{t \in \{1, \dots, T\}}$ and λ is the parameter vector

which we want to optimize with a maximum likelihood estimation (MLE). The parameter vector λ consists of

$(\phi_0, \phi_1, \theta_1, \theta_2, \alpha_1, \alpha_2, \beta_1, \beta_2, \hat{\sigma}_2^2, \hat{\sigma}_3^2)$ *these are needed because of (*)*

$$\hat{x}_t = \phi_0 + \phi_1 \{x_t\}_{t \in \{1, \dots, T-1\}}$$

so in vector notation $\hat{X} = \phi_0 + \phi_1 X_{[1:T-1]}$

$\begin{matrix} \text{vector} & & \text{vector} & & \text{vector} \\ (T-1 \times 1) & & (T-1 \times 1) & & (T-1 \times 1) \end{matrix}$

We can now obtain $\varepsilon = \hat{X} - X_{[1:T-1]} = \{\varepsilon_t\}_{t \in \{2, \dots, T\}}$

For the T-GARCH we need a for loop as well as initial estimates for $\hat{\sigma}_2^2, \hat{\sigma}_3^2$. (*)

for $t = 4, \dots, T$ (the first two epsilons are available ^{only} at $t=2, t=3$ and they are required for the calculation)

$$\hat{\sigma}_{t+1}^2 = \dots \text{ see } (*)$$

Then $l(\lambda|X) = \sum_{t=4}^{T \log} f\left(\frac{x_t - \hat{x}_t}{\hat{\sigma}_t}\right)$ is the log likelihood function.

We can then maximize $l(\lambda|X)$ or minimize $-l(\lambda|X)$ w.r.t.

λ using e.g. gradient descent or simulated annealing or some other algorithm for finding minima. We can restrict search

space by imposing restrictions on the search space, e.g.

finite unconditional mean $\phi_1 \leq 1$ or positive unconditional

mean: $\phi_1 > 0$ & $\phi_0 > 0$. This increases the speed of convergence.

We can obtain the time series of return variance of apple via $\{\hat{\sigma}_t^2\}_{t \in \{2, \dots, T\}} = \{\varepsilon_t^2\}_{t \in \{2, \dots, T\}}$ since

$$E[\varepsilon_t] = 0, V(\varepsilon_t) = E[\varepsilon_t^2] - E[\varepsilon_t]^2 = E[\varepsilon_t^2].$$

If our model is a good fit, then $\hat{\sigma}_t$ and σ_t should not deviate a lot. Furthermore we can check if we ~~decorrelated the variance~~ $\{\varepsilon_t^2 - \hat{\sigma}_t^2\}_{t \in \{2, \dots, T\}}$ still has

any serial correlation with e.g. Ljung-Box test. If we modeled the heteroscedasticity correctly there should be none.

2.2. Task B.2

Let $\{x_t\}_{t \in \{1, \dots, T\}} = \begin{bmatrix} x_1 \\ \vdots \\ x_T \end{bmatrix}$ be the returns. When it follows a Gaussian ARMA(1,1)-GARCH(1,1) model the DGP is

$$x_t = \hat{x}_t + \varepsilon_t \quad \text{where}$$

$$\hat{x}_t = \phi_0 + \phi_1 x_{t-1} + \theta_1 \varepsilon_{t-1} = \mathbb{E}[x_t | \mathcal{F}_{t-1}] \quad \text{and}$$

$$\varepsilon_t = \hat{\sigma}_t \eta_t \quad \text{where } \eta \sim \mathcal{N}(0, 1) \text{ and}$$

$$\hat{\sigma}_t^2 = \alpha_0 + \alpha_1 \varepsilon_{t-1}^2 + \beta_1 \hat{\sigma}_{t-1}^2 = \mathbb{V}(\varepsilon_t | \mathcal{F}_{t-1}).$$

ε_{t-1}^2 is the realized variance at time $t-1$: σ_{t-1}^2 .

It follows that $x_t | \mathcal{F}_{t-1} \sim \mathcal{N}(\hat{x}_t, \hat{\sigma}_t^2)$ so the joint log-likelihood can be written as

$$\ell(\lambda | X) = \sum_{t=3}^T \log f(x_t | \lambda, \mathcal{F}_{t-1}) + \log f(x_1 | \lambda) + \log f(x_2 | \lambda).$$

where $f(x_t | \lambda, \mathcal{F}_{t-1}) = \frac{1}{\sqrt{2\pi} \hat{\sigma}_t} \cdot e^{-\frac{(x_t - \hat{x}_t)^2}{2\hat{\sigma}_t^2}}$ is the probability

density function of x_t . The first two observations are special because no ~~conditional~~ conditional variance estimation & cond. mean can be calculated for these observations, because we lack the preceding data necessary*. We can either include $\hat{\sigma}_2^2, \varepsilon_1$ in the parameter vector λ , so that they will be estimated or use the unconditional ~~mean~~ values,

$$\text{i.e. } \hat{\sigma}_2^2 = \mathbb{V}(\varepsilon_t) = \frac{\alpha_0}{1 - \alpha_1 - \beta_1} \quad \text{and} \quad \mathbb{E}[\varepsilon] = 0 = \varepsilon_1.$$

This way we can "salvage" x_2 and use it for the MLE. We can ignore and not use $f(x_1)$ and $f(x_2)$ in the $\ell(\lambda)$ function at all, if we have enough data, however we need estimates for ε_1 and $\hat{\sigma}_2^2$, because of their recursive nature/definition.

will not use x_1 , because we would have to estimate α_0 and ε_0 etc. too many additional parameters

* namely ε_1 and $\hat{\sigma}_1^2$

Conditional return density 2-step ahead. For this

we need $E[x_{t+2} | \mathcal{F}_t]$ and $V(x_{t+2} | \mathcal{F}_t)$.

First the conditional mean:

$$\begin{aligned} E[x_{t+2} | \mathcal{F}_t] &= E[E[x_{t+2} | \mathcal{F}_{t+1}] | \mathcal{F}_t] \quad (\text{law of iterated expectations}) \\ &= E[\phi_0 + \phi_1 x_{t+1} + \theta_1 \varepsilon_{t+1} + \varepsilon_{t+2} | \mathcal{F}_t] \\ &= E[\phi_0 + \phi_1 (\phi_0 + \phi_1 x_t + \theta \varepsilon_t + \varepsilon_{t+1}) + \theta_1 \varepsilon_{t+1} + \varepsilon_{t+2} | \mathcal{F}_t] \\ &\quad \text{with } E[\varepsilon_{t+2} | \mathcal{F}_t] = 0 \\ &= \phi_0 + \phi_1 \phi_0 + \phi_1^2 x_t + \theta \varepsilon_t + 0 + \theta_1 \cdot 0 + 0 = \phi_0(1 + \phi_1) + \phi_1^2 x_t + \theta \varepsilon_t \end{aligned}$$

because x_t and ε_t are known in \mathcal{F}_t and $E[\varepsilon_{t+h} | \mathcal{F}_t] = 0 \forall h$.

Now the conditional variance:

$$\begin{aligned} V[x_{t+2} | \mathcal{F}_t] &= E[(x_{t+2} - E[x_{t+2} | \mathcal{F}_t])^2] \quad \text{with the law} \\ &\quad \text{of iterated exp. and the calculations from above we} \\ &\quad \text{arrive at } V[x_{t+2} | \mathcal{F}_t] = E[(\theta_1 + \phi_1) \varepsilon_{t+1} + \varepsilon_{t+2}]^2 \\ &\quad \text{because the variance of everything known at time } t \text{ is zero.} \end{aligned}$$

Since $\varepsilon_{t+1} \perp \varepsilon_{t+2}$

follows from above &
law of iterated expect.

We model the conditional variance with GARCH, so

$$\begin{aligned} V_t(x_{t+2} | \mathcal{F}_t) &= E_t[(x_{t+2} - E_t[x_{t+2} | \mathcal{F}_t])^2] = E_t[(\phi_1 \varepsilon_{t+1} + \theta \varepsilon_{t+1} + \varepsilon_{t+2})^2] \\ &= E_t[\alpha_0 + \alpha_1 \varepsilon_{t+1}^2 + \beta_1 \hat{\sigma}_{t+1}^2 | \mathcal{F}_t] \quad (\text{GARCH}) \\ &= E_t[\alpha_0 + \alpha_1 \varepsilon_{t+1}^2 + \beta_1 (\alpha_0 + \alpha_1 \varepsilon_t^2 + \beta_1 \hat{\sigma}_t^2)] \\ &= \alpha_0 + \alpha_1 \alpha_0 + \alpha_1^2 \varepsilon_t^2 + \alpha_1 \beta_1 \hat{\sigma}_t^2 + \beta_1 \alpha_0 + \beta_1 \alpha_1 \varepsilon_t^2 + \beta_1^3 \hat{\sigma}_t^2 \end{aligned}$$

$$\text{because } E_t[\varepsilon_{t+1}^2 | \mathcal{F}_t] = V(\varepsilon_{t+1} | \mathcal{F}_t) = \alpha_0 + \alpha_1 \varepsilon_t^2 + \beta_1 \hat{\sigma}_t^2 \quad (\text{GARCH})$$

$$\begin{aligned} &= E_t[\alpha_0 + \alpha_1 \varepsilon_{t+1}^2 + 2\phi_1 \theta \varepsilon_{t+1}^2 + \theta^2 \varepsilon_{t+1}^2 + \varepsilon_{t+2}^2] \quad \text{because } \varepsilon_{t+1} \perp \varepsilon_{t+2} \\ &\quad \text{with } \boxed{\varepsilon_{t+1} \perp \varepsilon_{t+2}} \\ &= E_t[E_t[\varepsilon_{t+2}^2 | \mathcal{F}_{t+1}]] = E_t[\alpha_0 + \alpha_1 \varepsilon_{t+1}^2 + \beta_1 \hat{\sigma}_{t+1}^2] \\ &\quad \text{with } E_t[\hat{\sigma}_{t+1}^2] = E_t[\varepsilon_{t+1}^2] \end{aligned}$$

$$\Rightarrow E_t[\varepsilon_{t+2}^2] = E_t[\alpha_0 + (\alpha_1 + \beta_1)\varepsilon_{t+1}^2]$$

since $E_t[\varepsilon_{t+1}] = \alpha_0 + \alpha_1\varepsilon_t^2 + \beta_1\hat{\varepsilon}_t^2$ the equation becomes:

$$\begin{aligned} V_t(x_{t+2}) &= \phi_1^2(\alpha_0 + \alpha_1\varepsilon_t^2 + \beta_1\hat{\varepsilon}_t^2) + 2\phi_1\theta_1(\alpha_0 + \alpha_1\varepsilon_t^2 + \beta_1\hat{\varepsilon}_t^2) \\ &\quad + \theta^2(\alpha_0 + \alpha_1\varepsilon_t^2 + \beta_1\hat{\varepsilon}_t^2) + \alpha_0 + \alpha_1\varepsilon_t^2 + \beta_1\hat{\varepsilon}_t^2 \\ &= (\phi_1^2 + 2\phi_1\theta_1 + \theta^2 + 1)(\alpha_0 + \alpha_1\varepsilon_t^2 + \hat{\varepsilon}_t^2) \end{aligned}$$

Having determined $V_t(x_{t+2})$ and $E_t(x_{t+2})$ we can write the conditional pdf density ~~as~~ (2-step ahead) as

$$f(x_{t+2} | \mathcal{F}_t) = \frac{1}{\sqrt{\pi^2 \Sigma}} \cdot e^{-\frac{(x_{t+2} - \mu)^2}{2\Sigma}}$$

Note: For the calculation of the conditional variance we can neglect η , because its variance is always assumed to be 1.