

# The ghost of unbalanced species distribution data in geographical model predictions

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## **ABSTRACT**

Unbalanced samples are considered a drawback in predictive modelling of species' potential habitats, and a prevalence of 0.5 has been extensively recommended. We argue that unbalanced species distribution data are not such a problem from a statistical point of view, and that good models can be obtained provided that the right predictors and cut-off to convert probabilities into presence/absence are chosen. The effects of unbalanced prevalence should not be confused with those of low-quality data affected by false absences, low sample size, or unrepresentativeness of the environmental and spatial gradient. Finally, we point out the necessity of greater research effort aimed at improving both the quality of training data sets, and the processes of validating and testing of models.

## Keywords

Species distribution models, unbalanced samples, predictive reliability, threshold cut-off.

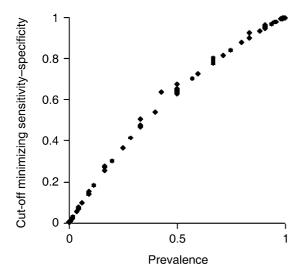
### INTRODUCTION

Species distribution modelling is now in wide use to develop analytical and prediction tools for ecology and conservation biology (Guisan & Zimmermann, 2000; Guisan & Thuiller, 2005), to locate previously unknown populations of rare and endangered species (Raxworthy et al., 2003; Guisan et al., 2006), to study the effect of climate warming on species distribution (Peterson, 2003; Thuiller et al., 2005a), to assess the possible impact of biological invasions (Rouget et al., 2004; Thuiller et al., 2005b), and to aid management in taking decisions (Schadt et al., 2002; Barbosa et al., 2003; Russell et al., 2004; Chefaoui et al., 2005). Quantified species-environment relationships, obtained through the development of a mathematical function linking species distribution information (usually presence/ absence) to environmental predictors, are used to map decimal fraction probabilities. These probabilities are usually taken as probabilities of presence and, so, as a measure of habitat adequacy. However, probability values are highly dependent on the relative proportion of each event in the sample, being biased toward the highest number of either presences or absences, where they differ. This inherent and unavoidable bias has long been recognized by statisticians under the name of the unbalanced sample effect (Hosmer & Lemeshow, 1989). This has some important consequences for the prediction of species distributions using models and has generated confused debate in the ecological literature that is not yet resolved.

## Statistical effects of unbalanced samples

The influence of prevalence on the performance of model predictions has repeatedly been judged to be of major importance (Vaughan & Ormerod, 2003; McPherson et al., 2004), leading to the supposition that the more unbalanced the samples, the less reliable the model predictions. In principle, there is no reason why the rarest events should necessarily be badly predicted, provided that models fit the data well (Cramer, 1999). Good fits can be obtained when good predictors are used and the dependent variable reflects all environmental variability. However, even in such circumstances, mean estimated probabilities of each event will be biased as a consequence of prevalence. This bias could be especially noticeable in the case of models that do not fit the data well (Cramer, 1999), typical of those derived from field studies where the most adequate predictors are usually unknown. This interaction between model fit and prevalence bias is a question that deserves further attention.

The apparently negative effect of prevalence on prediction reliability is mediated by the cut-off value selected to convert decimal fraction probabilities to a binary variable. This cut-off should be selected appropriately to account for unbalanced samples in the conversion of the decimal fraction probabilities to presence/absence, and to evaluate the model correctly when such measures as sensitivity, specificity, or the Kappa statistic, derived from a confusion matrix, are used (Fielding & Bell, 1997). As this conversion will determine model output, it will condition the



**Figure 1** Relationship between the cut-off that minimized the sensitivity–specificity difference and prevalence, using data from a simulated species and from randomly resampling different training data sets varying in prevalence. Data were modelled using logistic regressions.

cases assigned to each of the four categories of the matrix (true and false predicted presence, true and false predicted absences). The intuitively appealing 0.5 cut-off (e.g. Li et al., 1997; Berg et al., 2004; Meggs et al., 2004) makes no sense, as each model has its own characteristics related to prevalence and fit. For example, in the case of rare species data, a 0.5 cut-off would convert presences to absences and would yield a false sensitivity value (true predicted presences) of zero in the most extreme case. In a recently published paper (Liu et al., 2005), the optimum cut-off is sought through comparison of numerous criteria. Therein, the fixed 0.5 cut-off, or the widely used one that maximizes the Kappa value, was found to be among those that produced the worst results. The best presence/absence models were derived from cut-offs that maximize the sum, or minimize the difference, between sensitivity and specificity (true predicted absences), among others. Interestingly, cut-offs selected by these two criteria are highly and positively correlated with prevalence. Figure 1 shows the relationship between prevalence and the cut-off which minimizes the difference between sensitivity and specificity (see also Fig. 5 in Manel et al., 2001), using data from a simulated species and randomly resampling different training data sets varying in prevalence. Data were modelled using logistic regressions. These results suggest that the prevalence value itself could be used as a cut-off (Liu et al., 2005), as formerly recognized and suggested by statisticians (Cramer, 1999).

# **CONFOUNDING FACTORS**

Prevalence is a characteristic of the data that may sometimes correlate with species ecology, such as marginality, rarity, or specialization; these species are generally those of higher conservation concern. Bearing this in mind, caution must be exercised to avoid confusion between the effects of these biological

attributes and their associated data problems and those of prevalence.

When threshold-independent accuracy measures, such as area under the Receiver Operating Characteristic curve (Swets, 1988), are used to validate predictive models, confusing results have been obtained, as in some cases low prevalence values are related to high AUC scores, while the inverse has been found in other studies (see, e.g. Brotons et al., 2004; Luoto et al., 2005). McPherson et al. (2004) found best AUC scores with prevalence values around 0.5. But, if as pointed out before, there is no sound reason for models to perform poorly with unbalanced samples, what do these results mean? Effects of poor quality data can be misunderstood as false prevalence effects. For example, performance of species distribution models could depend on the sampling size of each event (independently of their relative size) and on the representativeness of the training data (i.e. presences and absences must be evenly distributed across the environmental and geographical gradient; a low sample size for an event implies poor representativeness), independently of prevalence. Additionally, the inclusion of false absences is surely a confounding factor present in many data sets whose effect will interact with prevalence. Poor quality data are usually associated with rare species, as presences are usually scarce and absences are prone to contain a high proportion of false data.

Thus, the true effect of prevalence is probably negligible when building predictive distribution models, and its 'ghostly' effect is due to other puzzling factors. To avoid this supposed unbalanced-sample problem, some authors recommend resampling the training data to balance presences and absences (McPherson et al., 2004; Liu et al., 2005). However, in the case of reliable training data, resampling would yield only a loss of information, mainly in rare species with scarce reliable data, and should be avoided.

## **RESCALING PROBABILITIES**

Finally, fitted probabilities from probability maps published in research papers, if considered indicative of habitat suitability, could be misleading. While potential probability may range from 0 to 1, probabilities that do not surpass a minimum value due to low prevalence could erroneously be interpreted as low, even for well-established populations. Although it could seem paradoxical, a low value of fitted probability may be assigned to a known presence event (Pontius & Batchu, 2003), given that an underrepresented event is less likely to occur in any sampling universe. To adjust the representativeness of the obtained probabilities adequately, favourability functions, such as the one proposed by Real et al. (2006), should be used, whose outputs are independent of prevalence due to the elimination of the random probability element. These favourability functions can be considered to be rescaling functions, as they convert logistic probabilities (P) into favourability values (F), assigning a value of F = 0.5 to the predictor conditions for which P = prevalence (Real et al., 2006). Interestingly, whereas P-values for different species are not comparable site to site because of the prevalence bias, this is not the case for F-values, which are directly equivalent (Real et al., 2006).

### **COROLLARY**

In conclusion, low prevalence is a property of low probability events, not a problem to be solved. Its effects on predictive tools are well known and, once accounted for, rare events should be accurately predicted if predictors are powerful and training data are reliable (especially absences) and neither spatially nor environmentally biased. These considerations are of special relevance in conservation biology, as low prevalence is usually a property of data from endangered species. Greater research effort aimed at improving both the quality of training data sets (Vaughan & Ormerod, 2003) and the processes of validating, and testing of models (Vaughan & Ormerod, 2005) should be made.

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