

To find the equations for params, we take log likelihood of the distribution:

$$LL = \sum_{n=1}^N \left\{ -\frac{1}{2} \log 2\pi\sigma_u^2 - \frac{1}{2\sigma_u^2} \left( x_{u,n} - \left( \sum_{c \in \pi_u} w_{uc} x_{uc,n} + w_{u0} \right) \right)^2 \right\} \quad (1)$$

We'll proceed to find the parameters for each node independently, and so dropping the  $u$  subscript from here on. Further, to avoid confusion, let's call the output node  $y$ , and the parents of  $y$  (which will our inputs) as  $x_c$  for  $c^{\text{th}}$  parent/input.  $y_n$  is the  $n^{\text{th}}$  output observation for the node in consideration, and  $x_{c,n}$  is the  $n^{\text{th}}$  observation for the  $c^{\text{th}}$  input.

$$LL = \sum_{n=1}^N \left\{ -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \left( y_n - \left( \sum_{c \in \pi} w_c x_{c,n} + w_0 \right) \right)^2 \right\} \quad (2)$$

Since we will be working with matrices, let's also define the required matrices for these variables:

$$W = \begin{bmatrix} w_0 \\ w_1 \\ w_2 \\ \vdots \\ w_c \end{bmatrix}, \dim(W) = (c+1) \times 1 \text{ is our weights matrix including } w_0. Y = \begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix}, \dim(Y) = n \times 1 \text{ where } n \text{ is the}$$

total number of samples.

$$X = \begin{bmatrix} 1 & x_{1,1} & x_{2,1} & x_{3,1} & \dots & x_{c,1} \\ 1 & x_{1,2} & x_{2,2} & x_{3,2} & \dots & x_{c,2} \\ 1 & x_{1,3} & x_{2,3} & x_{3,3} & \dots & x_{c,3} \\ \vdots & & & & & \\ 1 & x_{1,n} & x_{2,n} & x_{3,n} & \dots & x_{c,n} \end{bmatrix}, \dim(X) = n \times (c+1), \text{ and } x_{i,j} \text{ is } j^{\text{th}} \text{ observation sample for } i^{\text{th}} \text{ input}$$

variable/parent. For notational convenience, we define  $x_{0,i} = 1$  for all  $i$ .

Going back to our  $LL$  equation, for MLE, we take partial derivatives *w.r.t.* our weight parameters

$\{w_0, w_1, w_2, \dots, w_c\}$  and setting them to 0 one by one. That yields  $(c+1)$  equations. For  $w_i$ , the equation is:

$$\frac{\partial(LL)}{\partial w_i} = \sum_{n=1}^N (y_n - (w_0 x_{0,n} + w_1 x_{1,n} + \dots + w_c x_{c,n})) x_{i,n} = 0$$

Solving this further, and moving the summations in:

$$w_0 \sum_{n=1}^N x_{0,n} x_{i,n} + w_1 \sum_{n=1}^N x_{1,n} x_{i,n} + w_2 \sum_{n=1}^N x_{2,n} x_{i,n} + w_3 \sum_{n=1}^N x_{3,n} x_{i,n} + \dots + w_c \sum_{n=1}^N x_{c,n} x_{i,n} = \sum_{n=1}^N y_n x_{i,n}$$

We have  $(c+1)$  such equations for each  $w_i$  with  $i \in [0, c]$ . This is a system of  $(c+1)$  equations in  $(c+1)$  variables which can be uniquely solved. With the above matrix definitions, we can write the combined matrix replacement for the above equations as:

$$(X^T \times X) \times W = X^T \times Y$$

which is what we have implemented in code.

To solve for variance  $\sigma^2$ , we again take partial derivative of  $LL$  *w.r.t*  $\sigma^2$  and set it to 0, which after some simplification gives:

$$\sigma^2 = \frac{1}{N} \left( \sum_{n=1}^N (y_n - (w_0 x_{0,n} + w_1 x_{1,n} + w_2 x_{2,n} \dots w_c x_{c,n}))^2 \right)$$

which can be rewritten in matrix form as:

$$\sigma^2 = \frac{1}{N} [(XW - Y)^T \times (XW - Y)]$$