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Accounting for missing data in autoregressive models of ecological time series

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Key words: Time series; missing data; autoregressive models; forecasting; simulations

1 Abstract

2 Long time series are valuable and increasingly available tools for understanding ecological
3 systems. Unfortunately, sustaining data collection efforts over time is challenging, and gaps in
4 time series data can complicate data analysis. While several methods exist to account for missing
5 data, we know of no established best practice for handling missing data in time series models.
6 Using simulated and empirical time series of primary productivity in a river and great tit
7 population size, we compared the performance of six missing data approaches across scenarios
8 with varying amounts and types of missing data. In each scenario, we fit statistical models across
9 all approaches and assessed their ability to recover simulation parameters and forecast future
10 dynamics. When data were missing completely at random, our results indicate that researchers
11 have multiple methods available to them. Parameters were recovered well even with as high as
12 50% missingness. Conversely, no method can adequately account for data missing not at random
13 (e.g., temperature sensors fail more often on hot days) by default. Our study emphasizes that
14 identifying the mechanism of missingness is the most critical component of handling missing
15 data, at which point there may be multiple suitable analysis options one could implement.

16 Introduction

17 Long-term time series contribute to our understanding of many ecological phenomena, from
18 the impacts of species diversity on predator-prey dynamics to patterns of nutrient cycling in
19 ecosystems (Hughes et al., 2017; Likens et al., 1970; Sinclair et al., 2003). There has been a
20 concerted effort in recent years to facilitate the collection of long-term ecological datasets (e.g.,
21 the U.S. National Science Foundation Long Term Ecological Research Program), as well as to
22 re-analyze historical long-term datasets with modern methods (Adler et al., 2009; Buma et al.,
23 2017). However, even the most rigorously maintained long-term datasets are likely to have

24 missing observations, due in part to unpredictable barriers to data collection that are difficult if
25 not impossible to overcome (Łopucki et al., 2022; Nakagawa & Freckleton, 2008). Challenges
26 include faulty sensors (Hossie et al., 2021), inaccessibility of field sites due to safety concerns or
27 travel restrictions, funding lapses, and human error during data entry or collection. It is
28 impossible to retroactively collect observations to fill gaps in time series, instead requiring
29 statistical methods that account for missingness. Doing so is critical, as missing data can have
30 cascading negative effects on subsequent analyses including reduced statistical power (Kang,
31 2013; Moritz & Bartz-Beielstein, 2017) and possibly biased estimation of parameters, leading to
32 both inaccurate and imprecise conclusions (Aleryani et al., 2018; Junger & Ponce de Leon, 2015;
33 Kim et al., 2018).

34 Challenges in accounting for missing values are compounded when dealing with time series
35 data. For example, ecological time series are often autoregressive, meaning that the value of a
36 data point depends in part on the values of previous observations. Thus, a single missing
37 observation could lead to the effective deletion of multiple data points in downstream analyses.
38 Further, many ecological time series consist of discrete data such as Poisson-distributed counts or
39 binomial presence-absence data, which precludes the use of primary time series approaches (e.g.
40 ARIMA models). As such, approaches that directly impute missing values or are model-based
41 must be adjusted to account for time series that do not conform to classical models that assume
42 Gaussian error distributions.

43 An additional challenge arises via the underlying mechanism driving missingness in time
44 series, i.e. the statistical relationship between observations and the probability that a given
45 observation goes missing. Missing data can be Missing Completely at Random (MCAR), Missing
46 at Random (MAR), or Missing Not at Random (MNAR) (Nakagawa, 2015; Rubin, 1976,
47 Appendix S1: Fig. S1). Data are MCAR when the probability of a missing observation is

48 independent of both the numerical value of the observation itself, as well as any external
49 information available (e.g., the value of a covariate) (Horton & Kleinman, 2007; Nakagawa,
50 2015; D. A. Newman, 2014). Data are MAR when the probability of missingness is not related to
51 the numerical value of the missing observation, but the probability that an observation goes
52 missing can be explained by one or more observed predictor variables (e.g., a buried soil
53 temperature sensor malfunctions when soil moisture is high, causing wetter study sites to have
54 more missing soil temperature observations) (Ellington et al., 2015; Nakagawa, 2015;
55 D. A. Newman, 2014). Data are MNAR when the probability of missingness depends on the
56 value of either the missing observation (e.g., a sensor cannot record values above a threshold) or
57 an unmeasured predictor (e.g., high stream flow may be associated with high turbidity, but also
58 leads to missing observations in both datasets because of sensor disruption—missing turbidity
59 values are dependent on the missing stream flow values). Data that are MAR can often be treated
60 like data that are MCAR after conditioning on the known variables driving missingness
61 (Nakagawa, 2015; Nakagawa & Freckleton, 2011). MCAR and MNAR represent the two
62 extremes of missing data in terms of the statistical complications they pose, and as such our
63 analyses focus on these two missingness types.

64 Multiple methods exist for dealing with missing data (Fig. 1). Broadly speaking, these can be
65 categorized into data processing approaches that occur prior to statistical analysis or model-based
66 approaches that directly incorporate a model for missingness into the analysis. Data processing
67 approaches are often simple and intuitive, including dropping all missing time points or imputing
68 the data to fill them in before analysis. Even with more advanced imputation techniques (Kang,
69 2013; Nakagawa, 2015; Nakagawa & Freckleton, 2011; Onkelinx et al., 2017; Rubin, 1988, 1996)
70 the general workflow is still straightforward: repair the missingness in the data, then analyze it as
71 you would a complete dataset. By contrast, model-based approaches involve specifying a model

72 for missingness as part of your statistical analysis where missing values are treated as parameters
73 to be estimated (Kalman, 1960; Kang, 2013; Kong et al., 1994; Li et al., 2019; Nadjafi &
74 Gholami, 2022). Many approaches in both categories have been applied to ecological data
75 (K. Newman et al., 2023; Soldaat et al., 2007), but which method to choose and how their relative
76 performance depends on time series attributes remains unclear. This is particularly true when
77 confronted with time series that have different amounts and types of missing data or that have
78 discrete values or non-Gaussian error distributions.

79 Here, we evaluate the performance of six different approaches for dealing with missing data in
80 autoregressive models of ecological time series. We compare approaches using both real and
81 simulated datasets with both Gaussian and Poisson error distributions. We artificially introduce
82 different mechanisms and amounts of missingness and then quantify the performance of different
83 missing data approaches in terms of both parameter recovery (simulated datasets) and forecasting
84 accuracy (empirical datasets). Though no single method emerged as the overall best approach, we
85 provide a detailed discussion of the relative merits of each with recommendations for when and
86 how to use them in different contexts. Thus, we hope our results will be a resource for ecologists
87 and environmental scientists in search of robust, reproducible methods for analyzing time series
88 with missing data.

89 **Methods**

90 We first discuss the motivating case studies that we used to simulate time series with either
91 Gaussian-distributed or Poisson-distributed error and their two corresponding empirical
92 examples. We next overview creating data sets with different mechanisms, amounts, and
93 autocorrelation in missingness, and finally describe how we compare performance of the six
94 approaches for accounting for missingness.

95 **Real-valued time series:** Sensors that collect daily or hourly readings of environmental data have

96 become ubiquitous in environmental research, and are an ideal example of data typically modeled
 97 with a Gaussian error distribution. We thus refer to these time series as “real-valued” moving
 98 forward. Despite the prevalence of sensor data, these types of data are highly prone to
 99 missingness (Chen et al., 2013). We evaluated the impact of missingness using both simulated
 100 and empirical data that represent daily measures of environmental and response variables from a
 101 sensor. We simulated and analyzed such real-valued time series using a first-order auto-regressive
 102 (AR(1)) error model with explanatory covariates, such that:

$$Y_t = \mathbf{x}'_t \beta + \phi(Y_{t-1} - \mathbf{x}'_{t-1} \beta) + \varepsilon_t \quad (1a)$$

$$\varepsilon_t \sim \mathcal{N}(0, \sigma^2) \quad (1b)$$

103 where the data (Y_t) depend on the modeled effects of covariates (\mathbf{x}'_t), their associated parameters
 104 (β), and on an autoregressive error term (ϕ) which determines the influence of the previous time
 105 point on the current value of the time series. After accounting for the autoregressive structure,
 106 residual error (ε_t) at each time point is normally distributed with a variance of σ^2 . We used Eq. 1a
 107 to simulate 1000 datasets, each with two covariates and 365 observations, representing a year of
 108 daily sensor data.

109 We further applied Eq. 1a as our statistical model to an empirical dataset consisting of
 110 estimates of daily river gross primary productivity (GPP; units: $g O_2 m^{-2} d^{-1}$) across three years
 111 in the Au Sable river in Michigan, USA (Appling et al., 2018). We note that while this times
 112 series is missing GPP estimates for 50 days, it still provides 1046 non-missing values, and indeed
 113 we were unable to find long term empirical GPP time series without any missing values. In fitting
 114 Eq. 1a to these data, we used estimates of incoming light ($\mu\text{mol } m^{-2} s^{-1}$) and flow ($m^3 s^{-1}$) as
 115 covariates, since they have previously been identified as primary drivers of GPP (Bernhardt et al.,
 116 2022).

117 **Time series of counts:** Time series with integer-valued error distributions (subsequently referred
118 to as “time series of counts”), such as annual censuses of population size, are also very common
119 in ecology. Approaches for dealing with missing data in these types of time series are not as well
120 developed as they are for data with continuous values that can be modeled using Gaussian error
121 structure. For these reasons, we additionally evaluated the impact of missing data on estimator
122 performance using both simulated and empirical data that represent annual counts of individuals
123 in a population. To simulate and analyze a time series of counts, we used a stochastic Ricker
124 population model (Ricker, 1954) such that:

$$\eta_{t+1} = N_t e^{(r - \alpha N_t)} \quad (2a)$$

$$N_{t+1} \sim f(\eta_{t+1}) \quad (2b)$$

125 where η_{t+1} represents the expected population size at time $t + 1$, r is the intrinsic,
126 density-independent growth rate of the population, and $\alpha > 0$ is the intraspecific competitive
127 effect that induces negative density dependence. Realized population size, N_{t+1} , is a random
128 draw from the distribution $f()$ with mean η_{t+1} . Throughout our simulations and subsequent
129 analyses, we set $f()$ to be a Poisson distribution with rate parameter η_{t+1} . We used Eq. 2a to
130 simulate 1000 datasets of 60 observations each, representing 60 years of annual census data.

131 Paralleling our above approach, we also fit Eq. 2a to an empirical dataset consisting of a
132 59-year sequence of annual counts of great tit (*Parus major*) broods in the Wytham Woods in
133 Oxford, UK (Ben Sheldon, personal communication, <https://wythamtits.com>). These census data
134 are continuous from 1960 - 2018 with no missing values.

135 **Introducing missingness into time series:** To assess how missing data approaches perform
136 across differing amounts and types of missing data, we systematically removed observations from
137 both the empirical and simulated time series to create “missing datasets” (Appendix S1: Fig. S1).

138 Datasets differed in both the amount of missingness and mechanism of missingness, where we
139 examined data missing completely at random (MCAR) and missing not at random (MNAR), as
140 described below. These scenarios are meant to represent processes that could commonly occur in
141 ecological data collection but are not intended to be exhaustive.

142 For all simulated and empirical time series, we created MCAR datasets with varying
143 proportions of missing data and degrees of autocorrelation in missingness (Appendix S1: Fig. S1
144 B–E) by viewing a time series as a Markov-modulated Bernoulli process where the response
145 variable could have two states: missing or not missing (Edwards, 1960; Gharib et al., 2014). We
146 used a transition matrix to stochastically introduce increasing levels of missingness and different
147 degrees of autocorrelation in missingness (where missing values were temporally clumped or
148 evenly distributed across the time series) into our time series (see **Appendix S1: Introducing**
149 **Missingness**). For each simulated time series, we created 150 MCAR datasets by combining 15
150 levels of missing data (from 5–75% by increments of 5%) with 10 levels of autocorrelation (from
151 0 to ~ 0.9 by increments of ~ 0.1). For both empirical time series, we created 450 MCAR
152 datasets, representing 50 instances drawn from each combination of three levels of missingness
153 ($20 \pm 5\%$, $40 \pm 5\%$, and $60 \pm 5\%$) with three levels of autocorrelation (0.25 ± 0.05 , 0.5 ± 0.05 , and
154 0.75 ± 0.05).

155 We created real-valued MNAR datasets by removing observations at both the high and low
156 tails of the distribution of data (Appendix S1: Fig. S1 F,G). For each of the simulated real-valued
157 time series we created 15 MNAR datasets with missingness increasing from 5 to 75% by
158 increments of 5%. We also created 15 total MNAR datasets from the empirical GPP time series,
159 resulting in five MNAR datasets within each of three binned levels of missingness: $20 \pm 5\%$,
160 $40 \pm 5\%$ and $60 \pm 5\%$ (see **Appendix S1: Introducing Missingness**). We did not create time
161 series of counts with MNAR data because that has no real-world parallel: in population count

162 data, a population's size is unlikely to affect the probability that a sampling event occurs that year.

163 It is important to note that while very low population size may lead to false zeros, observation

164 error is a problem that is distinct from missing data.

165 **Comparing missing data approaches:** We evaluated several, previously published approaches

166 for accounting for missing data: simple and complete data deletion (Nakagawa & Freckleton,

167 2011), multiple imputation (MI) (Rubin, 1988), the Kalman filter (KF) (Kalman, 1960), the

168 expectation maximization (EM) algorithm (Kang, 2013), and Bayesian data augmentation (DA)

169 (Kong et al., 1994) (Fig. 1). Briefly, both types of data deletion and MI are data processing

170 approaches, meaning we applied them *before* fitting the corresponding statistical models. In

171 simple data deletion, only the missing values are removed, and the resulting time series is

172 compressed, violating the assumption of equal spacing between observations. In contrast,

173 complete data deletion removes the missing value and any subsequent observation(s) it would

174 predict from the response variable (note that subsequent observations are retained as predictors of

175 future observations). MI systematically fills in missing observations with imputed values, creating

176 multiple versions of the dataset to be fit and averaged over. The KF, EM algorithm, and Bayesian

177 DA are all model-based approaches and thus are implemented *simultaneously* with fitting the

178 relevant statistical model. The KF derives the likelihood of a time series with missing

179 observations by following a two-step procedure: forecasting a future state given an initial state,

180 then using the next time point to update the forecast using Bayes' theorem. The KF assumes a

181 Gaussian error distribution, so we only used this method with our real-valued time series. The

182 EM algorithm is conceptually similar to the KF but can be used for count data. As such, we only

183 employed the EM algorithm with our time series of counts to replace the KF. Finally, Bayesian

184 DA uses a model-based framework to simultaneously estimate both the standard model

185 parameters and the value of any missing data points which are treated as additional parameters.

186 See **Appendix S1: Missing data approaches** for more detailed descriptions of all these methods.

187 To assess the impacts of missing data approach and amount and type of missingness on

188 parameter recovery, we applied each missing data approach to our MCAR and MNAR datasets

189 from *simulated* time series. For our real-valued time series, we used the missing data approaches

190 to fit Eq. 1a and assessed the recovery of the ϕ and β parameters. For our time series of counts,

191 we fit Eq. 2a and assessed the recovery of the r and α parameters. For each parameter and model

192 fit, we calculated three metrics: relative error (e_s), absolute relative error ($|e_s|$), and whether or not

193 the estimated 95% Confidence Interval (CI) contained the true parameter value used for the

194 simulation (i.e., coverage). Relative error was calculated as $e_s = \frac{\hat{\theta}_s - \theta_s}{\theta_s}$ where $\hat{\theta}_s$ is the estimated

195 parameter value and θ_s is the true value for simulation s . Using relativized values facilitated

196 comparison across different simulated datasets and among different parameters within the same

197 model. We use both relative and absolute relative error as relative error measures bias in the

198 parameter estimate (positive values indicate an overestimate and vice versa) while the absolute

199 relative error facilitates comparisons between methods in the magnitude of error. We use

200 coverage to assess how well methods accurately convey the uncertainty associated with parameter

201 estimates across different missingness levels and mechanisms.

202 To aggregate these metrics across our simulations, we grouped results according to data type

203 (real-valued or count), missing data approach, missingness type (MCAR or MNAR), proportion

204 of missing data (rounded to the nearest 10%), and three levels of autocorrelation ($0.25 \pm 0.05\%$,

205 0.5 ± 0.05 , and 0.75 ± 0.05). Within these groups, we calculated the median relative error, median

206 absolute relative error, and coverage. We reported the median of the error distributions (rather

207 than mean) to reduce the influence of rare, extremely large error values. If the models and missing

208 data approaches worked perfectly, the median relative error would be 0 (indicating no bias), the

209 absolute relative error would be low (indicating low magnitudes of error), and the coverage would

210 be 0.95 (indicating the 95% CIs accurately capture the associated uncertainty in a given parameter
211 estimate). Finally, we note that, since we relativized the estimates and because the two β
212 parameters in Eq. 1a are of the same class (i.e. regression coefficients), we pooled their results.

213 Next, to quantify the impacts of missing data approach and amount and type of missingness on
214 forecasting accuracy, we applied each missing data approach to our MCAR and MNAR datasets
215 generated from the two *empirical* time series. To this end, we first split each empirical time series
216 into training and testing sets such that models were fit to the training set and those fits were
217 subsequently used to forecast the unobserved testing set. The Au Sable GPP time series
218 (real-valued) used the first two years of daily GPP estimates (66% of the total time series) as the
219 training set with the remaining year (33%) used for the testing set. In the great tit data (time
220 series of counts), we used the first 49 years (83%) for the training set and the last 10 years of data
221 (17%) as the testing set. Then for each model forecast we calculated the root mean squared error
222 (RMSE) of the forecasted values compared to the true observations. As described above, we
223 aggregated results from each time series according to the missing data approach used, type of
224 missingness, proportion of missing data, and autocorrelation in missing data. Within each
225 grouping, we then calculated the mean RMSE value.

226 **Results**

227 **Parameter recovery in models of real-valued time series:** For the real-valued MCAR data, all
228 missing data approaches performed similarly well in recovering the covariate β parameters, with
229 median error and median absolute error both remaining below 10% and coverage very near 95%
230 across all levels of missing data. In contrast, estimates of the autoregressive parameter, ϕ , were
231 more affected by missing data, and some missing data approaches performed better than others
232 (Fig. 2). Notably, the DA and KF methods had the strongest overall performance, with low error
233 and accurate coverage of ϕ with as much as 60% missing data. All other approaches lost both

accuracy and coverage as the amount of missing data increased. The two data deletion methods, simple and complete-case, performed nearly identically in their ability to recover ϕ . Both these methods had increasing median error and median absolute error as the amount of missingness increased, while coverage dropped as low as $\sim 70\%$ (Fig. 2). MI had the worst ability to recover the ϕ parameter across every proportion of missing data. Coverage dropped sharply with more than 10% missing data, while median error became more negative (eventually underestimating ϕ by $\sim 50\%$) and median absolute error steadily increased. The combination of large magnitude errors and low coverage seen in the lowest-performing approaches demonstrates they could be especially misleading.

The degree of autocorrelation in MCAR data also affected parameter recovery, but only for the two data deletion methods and only for ϕ (Fig. 3; see Appendix S1: Fig. S2 for the effect of autocorrelation on β estimates). Specifically, higher levels of autocorrelation led to more accurate estimates of ϕ when either complete case or simple data deletion were used, while other approaches were consistently accurate. For this reason, we primarily present results from moderate autocorrelation MCAR datasets in the main text. However the corresponding results for low and high autocorrelation can be found in the supplement (Appendix S1: Fig. S2).

Unsurprisingly, misleading parameter estimates were an even greater concern when fitting models to MNAR data. In contrast to our results for MCAR data when data are MNAR, our analysis indicates these misleading estimates occur for both β and ϕ parameters across all missing data approaches, with increasing effects as missingness increases (Fig. 2). All missing data approaches underestimated ϕ by $\sim 75\%$ and overestimate β by $\sim 40\%$ at the highest proportion of missingness (60%). Coverage for β plummeted from $\sim 75\%$ with low missing data to below 50% with 10% missing data. Once 40% or more of data were missing, coverage of the β parameters approached 0%, indicating that almost none of the CIs included the true values for

258 these parameters (Fig. 2).

259 **Parameter recovery in models of count time series:** When fitting the Eq. 2a population model
260 to simulated time series of count data with observations that were MCAR, the simple data
261 deletion approach did not recover the density-independent growth rate (r) or intraspecific
262 competitive effect (α) as accurately as other approaches, particularly as the proportion of missing
263 data increased (Fig. 4). Among the remaining approaches, EM consistently had low median error
264 and median absolute error. Interestingly, complete data deletion also had low median error,
265 similar to EM, but higher median absolute error than EM. However, this pattern was reversed for
266 DA, which recovered parameters with relatively higher median error than EM, but lower median
267 absolute error (i.e. the estimates were more stable, but slightly more biased). Finally, MI
268 recovered parameters with higher median error and median absolute error, though was still much
269 more accurate and precise when compared to simple data deletion. In addition to generating
270 higher errors, both simple data deletion and MI had substantially reduced coverage of both
271 parameters. Similar to the results for real-valued time series, autocorrelation in missing data led
272 to more accurate parameter estimates for the simple data deletion method, but had little impact on
273 other methods (Appendix S1: Fig. S3).

274 **Forecasting with missing data:** Examining the effect of missing data on forecasts of real-valued
275 time series revealed several key patterns. First, and unsurprisingly, increasing the proportion of
276 MCAR data resulted in higher forecast RMSE across multiple model runs (Fig. 5B). Second,
277 when faced with MCAR data, all missing data approaches performed similarly in terms of
278 forecast RMSE. This likely reflects the relative importance of environmental covariates versus
279 autoregressive structure in the GPP time series as our simulations showed β estimates to be much
280 less impacted by missing data compared to ϕ estimates (Fig. 2). Finally, when confronted with
281 MNAR data, model forecasts had substantially increased RMSE, even with relatively low levels

282 of missing data. In fact, with only 20% of data MNAR, model forecasts had higher RMSE across
283 all methods than scenarios in which 60% of the data were MCAR.

284 As with the real-valued time series, forecasts for time series of counts with MCAR
285 observations also had increasing RMSE with increasing proportions of missing data. There was
286 also a dramatic rise in the width of the interquartile ranges of RMSE values with increasing
287 proportions of missing data (Fig. 6). We expect this is due to instability of the estimates as the
288 total number of observations decreases. Similar to our forecasts of real-valued time series, all
289 missing data approaches performed broadly similarly when forecasting time series of counts with
290 MCAR data.

291 Discussion

292 Missing data are ubiquitous in ecological studies and can be particularly problematic when
293 fitting autoregressive time series models since missing values violate a key statistical assumption
294 that observations are equally spaced in time. We evaluated parameter recovery and forecasting
295 ability of six previously proposed approaches for dealing with missing data by fitting
296 autoregressive models to simulated and empirical examples of two different types of time series
297 common in ecology. Our results indicate that parameter estimation can be fairly robust to missing
298 data so long as an appropriate method is used to account for the missing data and the data are
299 missing completely at random (MCAR). Thus, missing values in time series do not necessarily
300 have a catastrophic effect on model accuracy, precision, or forecasts. In fact, several methods,
301 such as DA, the KF (for real-valued time series), and EM (for time series of counts), could
302 recover simulation parameters relatively well, even when 50% of observations were MCAR
303 (although we still suggest caution if fitting models to time series with this much missing data).

304 While several missing data approaches performed well for a range of data and missingness
305 types, we identified other approaches that were uniformly unsuitable and that we advocate against

306 using. For example, MI performed poorly, especially when applied to real-valued time series.
307 This is particularly notable given that MI has been shown to work quite well in non-time series
308 data (Graham, 2009) and is commonly recommended for ecology (Ellington et al., 2015;
309 Nakagawa & Freckleton, 2008). Overall, though, the most consistently poor performing method
310 was simple data deletion in which unequal spacing between observations is ignored, violating a
311 key assumption of most time series models. Simple data deletion tended to cause high error and
312 simultaneously low coverage, meaning it produced incorrect estimates, but with relatively narrow
313 confidence intervals. This effect was most pronounced for time series of counts (Fig. 4) but still
314 present for real-valued time series (Fig. 2). We hypothesize the adverse effects of simple data
315 deletion may have been mitigated in our real-valued time series by including two additional
316 covariates whereas our time series of counts contained no covariates. The poor performance of
317 simple data deletion is alarming because it is easy to implement and can be commonly applied in
318 ecological studies, potentially even unknowingly depending on how software handles missing
319 observations. However, based on our results we echo previous advice that researchers should
320 avoid simple data deletion (Łopucki et al., 2022; Nakagawa & Freckleton, 2011; Shoari & Dubé,
321 2018).

322 Before selecting a particular approach for dealing with missing data, we argue it is critical to
323 first identify the mechanism leading to missing data. We found that when data were missing not at
324 random (MNAR), none of the methods we tested provided robust parameter estimates or
325 forecasts. Unfortunately, many ecological time series may contain data that are MNAR (Bowler
326 et al., 2025), but this may be difficult or even impossible to verify (Nakagawa, 2015). However, if
327 the processes behind MNAR data can be identified and modeled, the missing data can be treated
328 as MCAR conditional on the missingness process, which then makes applying common missing
329 data approaches more viable (Nakagawa, 2015; D. A. Newman, 2014). With MCAR data, our

330 results suggest that within both real-valued time series and time series of counts, there are several
331 methods that tend to perform consistently well. The best suited approach will depend on the type
332 of time series being investigated and the goals of each particular analysis. For example, if
333 researchers wish to estimate the value of missing data points in a time series, then DA would be
334 the best choice. If a clear quantification of uncertainty in parameter estimates or forecasts is
335 important, researchers should avoid EM which does not directly provide standard errors of the
336 estimates (Figs. 1, 4).

337 Another key consideration when choosing among missing data approaches is the
338 autoregressive nature of the time series and model. For example, if forecasting real-valued time
339 series with covariates, the choice among missing data approaches may not matter as all were able
340 to accurately recover the β parameters in simulations (Fig. 2), leading to similar accuracy in
341 forecasts (Fig. 5). However, if the autoregressive model is itself of interest, our results suggest
342 options for adequately dealing with missing data become more limited. Only two approaches had
343 similarly strong recovery of the autoregressive parameter ϕ : KF and DA. In contrast, both data
344 deletion methods struggled to accurately recover ϕ except at high levels of autocorrelation in
345 missingness (Fig. 3). This is likely because high autocorrelation in missing values also preserved
346 longer stretches of intact data, allowing easier estimation of the autoregressive structure in these
347 instances. In contrast to our model for real-valued time series, our model for time series of count
348 data contained no additional covariates and estimated current population abundance purely based
349 on abundance in the previous time point (Eq. 2a). Thus, the forecasting performance of all
350 missing data approaches with this model demonstrate the potential importance of autoregression,
351 as RMSE increased much more dramatically with increasing proportions of missing data than
352 compared to forecasts of real-valued time series (Figs. 5, 6). However, it should be noted that
353 previous analyses of the great tit data suggest that more complex models than used here (e.g.,

354 incorporating varied functional relationships or spatial processes) greatly aid in adequately
355 modeling dynamics in this system (Lebreton, 1990).

356 Our results provide important insights into the performance of a variety of methods for dealing
357 with missing data in ecological time series, where our simulations can be viewed as benchmarks
358 demonstrating both best and worst case scenarios, respectively. For example, when working with
359 a specific dataset, several methods can be ‘tuned’ to achieve better or worse performance (e.g., MI
360 might perform better if using more imputed datasets (Honaker & King, 2010), while the KF might
361 perform worse when setting the Kalman gain below one to account for observation error (Kalman,
362 1960)). Further, we chose to explore the effect of missing data in the response variable and
363 assumed no missingness in the covariates. However, ecological datasets often have missing data
364 in the covariates as well as the response, which would complicate the modeling approach. There
365 is much literature on the problem of missing covariates (albeit, not necessarily in the context of
366 time series) that is beyond the scope of this manuscript (Little, 1992). Similarly, we chose an
367 empirical real-valued time series with known, strong effects from environmental covariates
368 (Bernhardt et al., 2022; Hall et al., 2015). In more exploratory scenarios or systems with weaker
369 covariate effects, inaccurate estimates of autoregressive parameters might have larger negative
370 effects on model forecasting ability. Finally, we only explored MCAR and MNAR data,
371 representing the two opposite ends of a spectrum from best to worst case scenarios
372 (D. A. Newman, 2014). In between these are scenarios with data missing at random (MAR),
373 meaning the probability of missingness depends on one or more covariates or known processes
374 (D. A. Newman, 2014). Properly accounting for such missing data requires knowing the
375 underlying mechanism such that the likelihood of missingness can itself be modeled. This may be
376 a tall order in many ecological systems (e.g., McCall et al., 2014; Shoari & Dubé, 2018; Sotto
377 et al., 2011). Though we did not explore MAR data, we argue model performance in this case

378 would fall in between the two extremes of MCAR and MNAR presented here and depend on how
379 well the missingness mechanism is modeled.

380 Collectively, our comparison of missing data approaches provides a guide for dealing with
381 missing data in ecological time series. Our results, especially for the real-valued data, suggest that
382 parameter estimation and forecast accuracy can be robust and even relatively high with MCAR
383 data – especially when using a suitable method for accounting for missingness. However, we
384 acknowledge this is unlikely to hold for all data types or model structures. As a result, we suggest
385 that researchers confronting missing data in time series with other characteristics not considered
386 here (e.g., different model or error structures, different lengths, etc.) run similar simulations to
387 those we performed to decide how to best deal with missing observations. We argue this step
388 should become part of the analysis workflow, particularly in a field-based discipline such as
389 ecology where missing observations are often unavoidable.

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394 **Conflict of Interest Statement**

395 The authors declare no conflicts of interest.

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529 **Figure captions**

530 **Fig 1:** Conceptual figure showing the different approaches to handling missing data considered in
531 this manuscript and the mechanisms for each approach.

532 **Fig 2:** Parameter recovery metrics for real-valued time series, showing the median error of
533 parameter recovery, absolute median error of parameter recovery, and 95% coverage of parameter
534 estimates, resulting from models fit using five missing data approaches (colors) across an
535 increasing proportion of missing data. These models were fit to simulated, real-valued datasets
536 with data missing completely at random (MCAR; left panel) with autocorrelation in missingness
537 >0.3 and <0.6 , and data missing not at random (MNAR; right panel). The coverage panel shows
538 the proportion of model runs where the 95% confidence interval of a parameter estimate includes
539 the true simulation parameter (dotted line at 0.95). Each point in each panel shows the median
540 value of error or coverage across all models fit to simulations that used the same missing data
541 approach, had the same amount of missing data (within a 10% bin), approximately the same
542 amount of autocorrelation in missingness, and the same type of missingness.

543 **Fig 3:** Median error of parameter recovery of ϕ depending on the proportion of missing data and
544 autocorrelation in missingness for each of five missing data approaches, using simulated,
545 real-valued datasets with data missing completely at random (MCAR). Cells in dark blue (closer
546 to 0) show that the median of model estimates of ϕ were close to the actual simulation parameter.
547 See Appendix S1: Figure S2 for a version of this figure expanded to include all parameter
548 recovery metrics (error, absolute error, and coverage) for both ϕ and β parameters.

549 **Fig 4:** Parameter recovery metrics for time series of counts, showing the median error of
550 parameter recovery, absolute median error of parameter recovery, and 95% coverage of parameter
551 estimates, resulting from models fit using five missing data approaches (colors) across an
552 increasing proportion of missing data. These models were fit to simulated time series of count
553 data where observations were missing completely at random (MCAR) with autocorrelation in
554 missingness >0.3 and <0.6 . The coverage panel shows the proportion of model runs where the

555 95% confidence interval of a parameter estimate includes the true simulation parameter (dotted
556 line at 0.95). Each point in each panel shows the median value of error or coverage across all
557 models fit to simulations that used the same missing data approach, had the same amount of
558 missing data (within a 10% bin), approximately the same amount of autocorrelation in
559 missingness, and the same type of missingness.

560 **Fig 5:** (A) Daily estimates of scaled gross primary productivity (GPP) from the Au Sable River
561 from 2012 through 2015 (699 days). Red tick marks on the x-axis indicate days when
562 measurements were missing. The gray box indicates the data (347 days) excluded from model
563 fitting, which were then forecast using the resulting parameterized model. (B) Root mean square
564 error (RMSE) of forecasts made using a model fit to the Au Sable GPP time series shown in panel
565 A. Each point shows the mean RMSE across all forecasts with a given missing data approach
566 (indicated by color), in a given level of proportion of missing data (0.2 ± 0.05 , 0.4 ± 0.05 , $0.6 \pm$
567 0.05), and with a given type of missingness (missing completely at random (MCAR) with
568 autocorrelation of 0.5 ± 0.05 or missing not at random (MNAR)). Error bars indicate the
569 inter-quartile range.

570 **Fig 6:** A) Annual counts of great tit (*Parus major*) broods in the Wytham Woods from 1960 –
571 2018. The gray box indicates the data (10 years) that were excluded from model fitting, which
572 were then forecast using the resulting model. (B) Forecast root mean square error (RMSE) of the
573 great tit dataset with an increasing proportion of missing data using 5 approaches that account for
574 missingness. Each point shows the mean RMSE across all forecasts with a given missing data
575 approach (indicated by color), in a given level of proportion of missing data (0.2 ± 0.05 , $0.4 \pm$
576 0.05 , 0.6 ± 0.05) with all bars shown for autocorrelation of 0.5 ± 0.05 . Error bars indicate
577 inter-quartile range

578 Figures

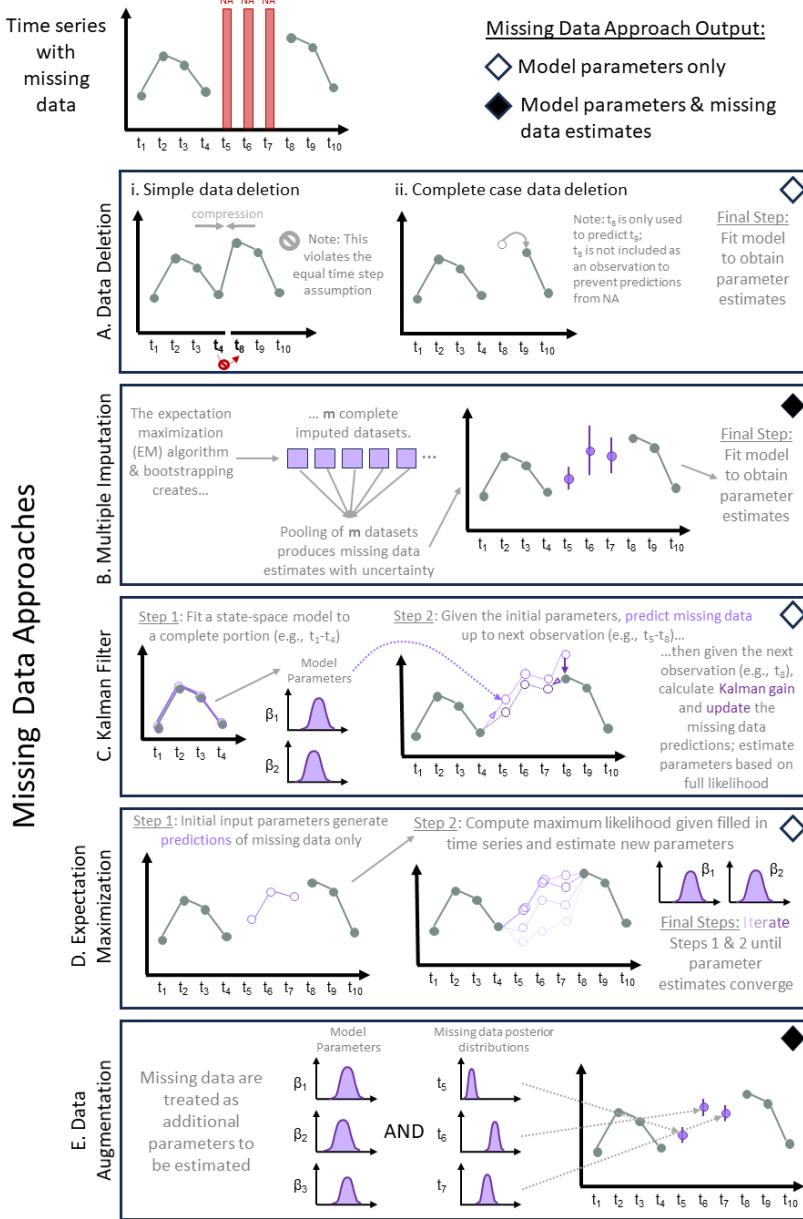


Figure 1:

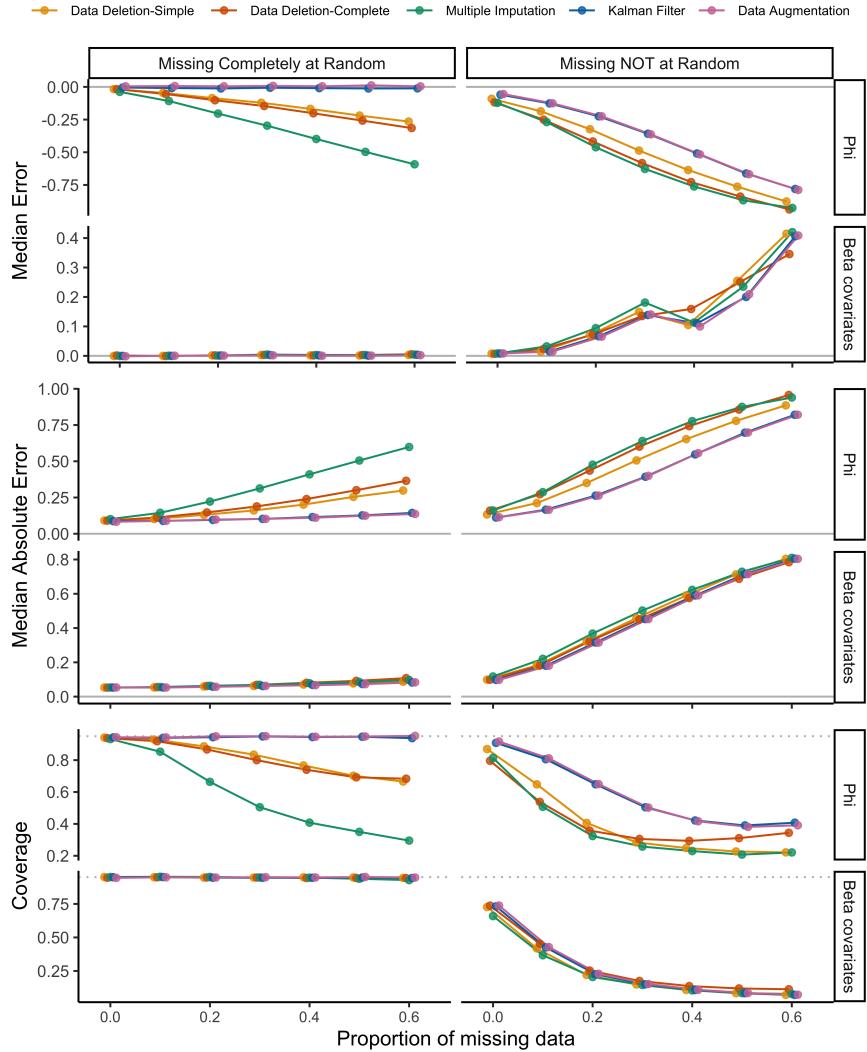


Figure 2:

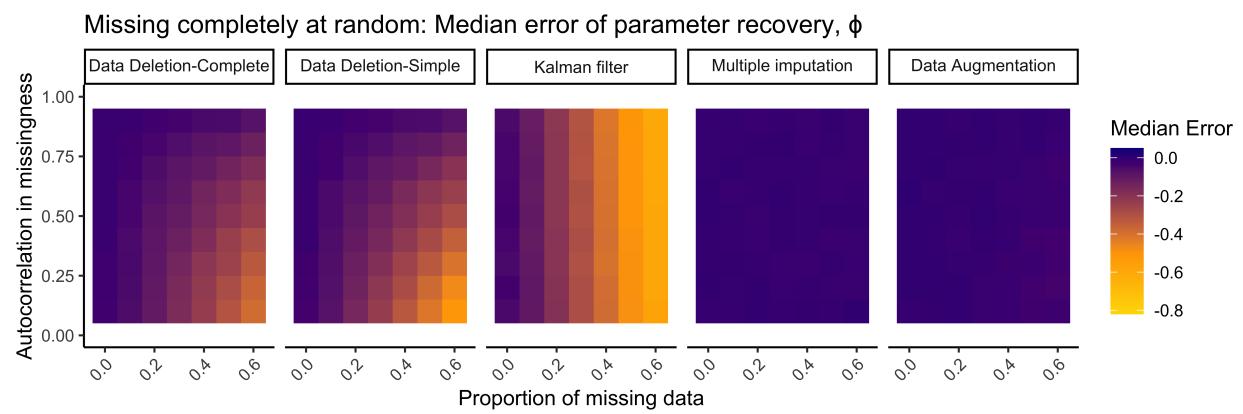


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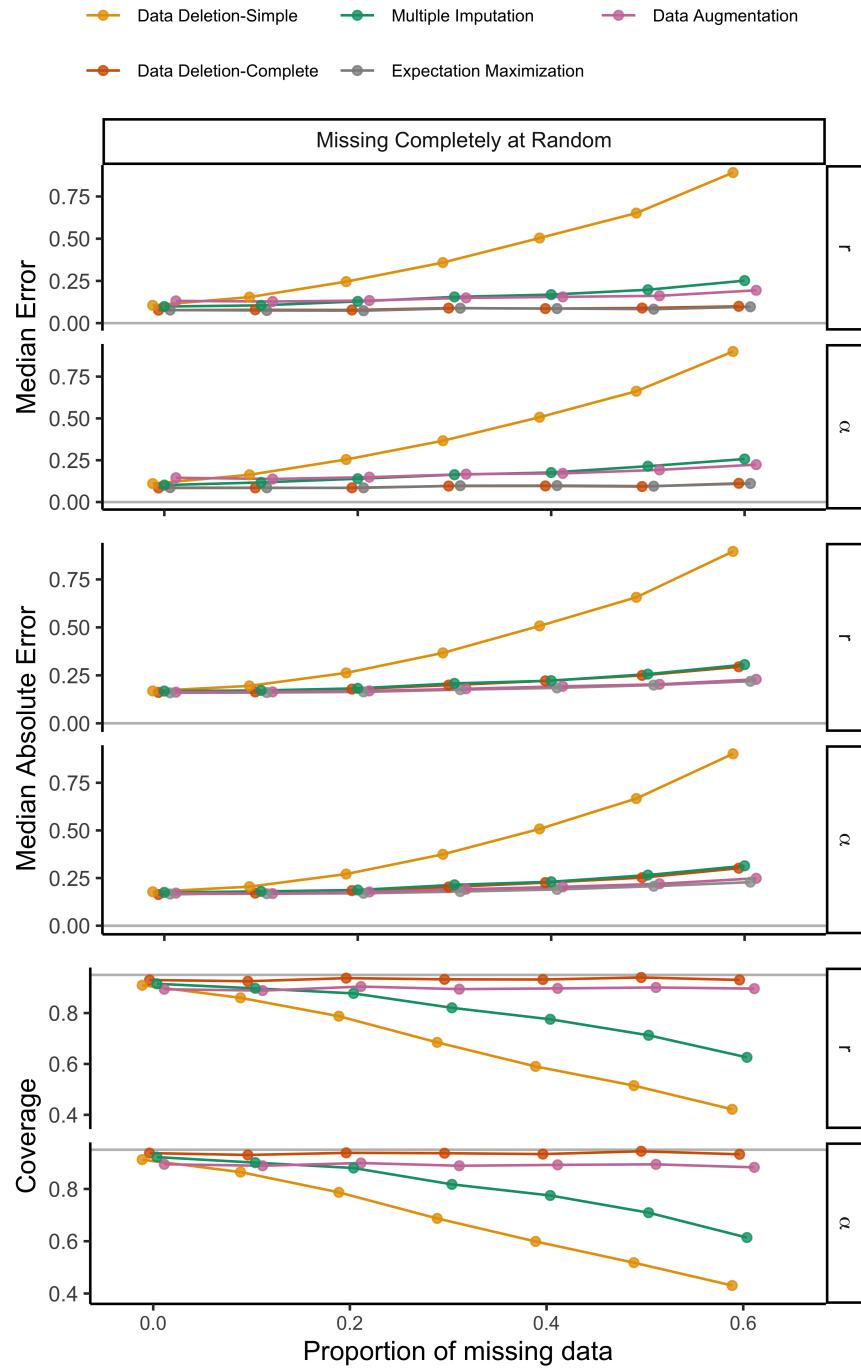


Figure 4:

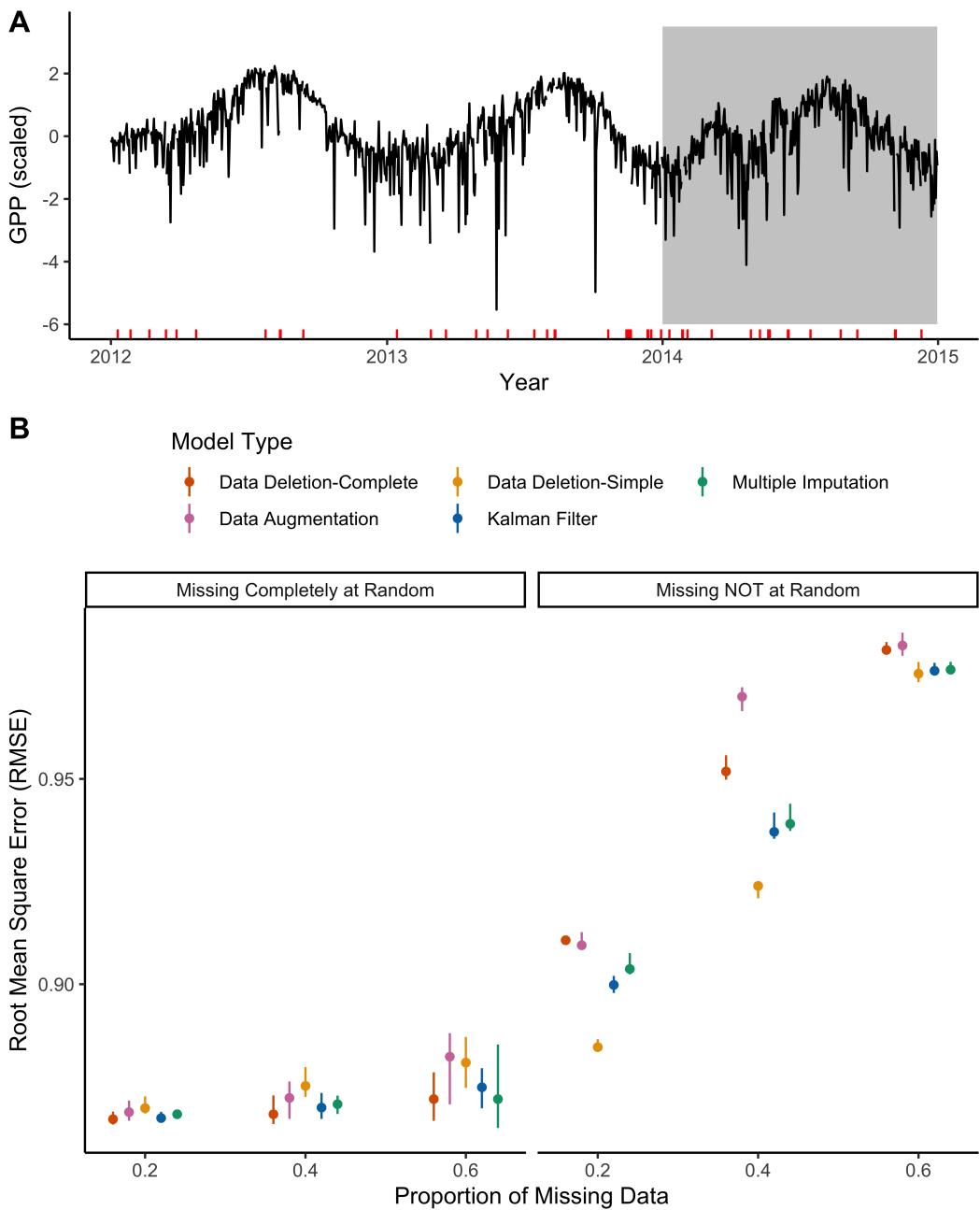


Figure 5:

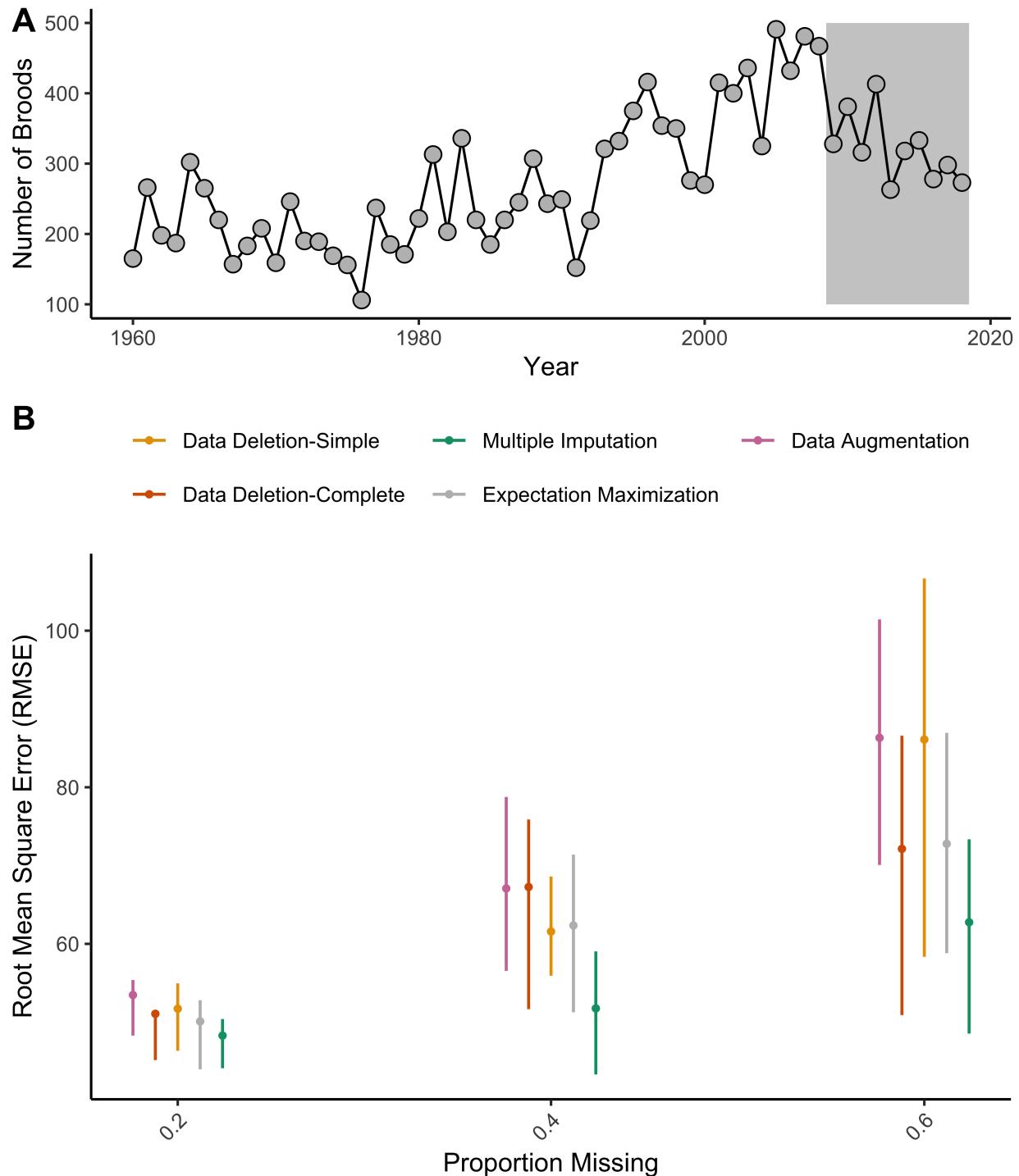


Figure 6: