# Lossless compression and cumulative log loss

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Lossless data compression
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Lossless data compression

The guessing game

Arithmetic coding

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Lossless data compression

The guessing game

Arithmetic coding

The performance of arithmetic coding

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Lossless data compression

The guessing game

Arithmetic coding

The performance of arithmetic coding

Source entropy

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Unbiased prediction

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Lossless data compression

The guessing game

Arithmetic coding

The performance of arithmetic coding

Source entropy

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The guessing game

Arithmetic coding

The performance of arithmetic coding

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- A natural way for describing a distribution.

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  - To decode use the same prediction algorithm

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- Widely used in practice.

Easier notation: represent characters by numbers

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- ► Code = discriminating binary expansion of a point in  $[l_t, u_t)$ .

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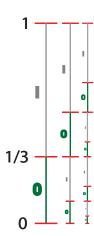
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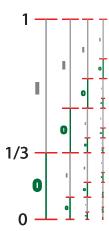
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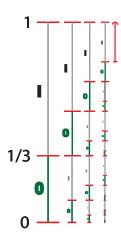
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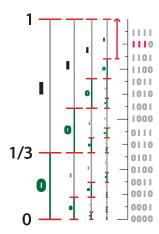
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- ► Holds for all sequences.

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Entropy is the expected value of the cumulative log logs

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The proof of Shannon's lower bound is not trivial (Can be a student lecture).

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- There are other losses with this property, for example, square loss.

Other examples for using log loss

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- If forecaster predicts with the true probabilities then

$$E(\log b_T) = T - H(p_T)$$

and that is the maximal expected value for  $E(\log b_T)$ 



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- Taking logs, we get cumulative log loss.

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  - Good prediction model = model that minimizes the total code length
- Often inappropriate because based on lossless coding. Lossy coding often more appropriate.

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- Goal: Total loss of algorithm minus loss of best predictor should be at most log<sub>2</sub> N

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- ► \[\[ \bigcup\_A^T \] is the code length if \( \bigcup\_A \) is combined with arithmetic coding.

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  - c<sup>t</sup> is revealed.
- ► Goal: minimize regret:

$$-\sum_{t=1}^{T} \log p_{\mathcal{A}}^{t}(c^{t}) + \min_{i=1,\dots,N} \left( -\sum_{t=1}^{T} \log p_{i}^{t}(c^{t}) \right)$$

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#### **EQUALITY** not bound!

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▶ Dividing by T we get  $\frac{L_{T}^{T}}{T} = \min_{i} \frac{L_{T}^{T}}{T} + \frac{\log N}{T}$ 

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- We don't pay a penalty for copies.
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- Can we still get a meaningful bound?

# Bayes Algorithm for biased coins

Replace the initial weight by a density measure

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We need a new lower bound on the final total weight

Bayes using Jeffrey's prior

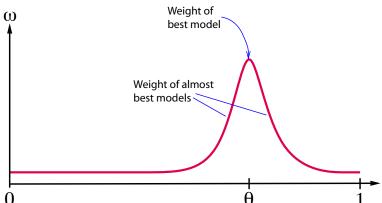
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# Laplace approximation (idea)

► Taylor expansion of  $g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta})$  around  $\theta = \hat{\theta}$ .

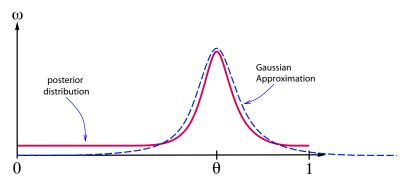
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# Laplace Approximation (details)

$$\int_0^1 w(\theta) e^{T(g(\hat{\theta},\theta)-g(\hat{\theta},\hat{\theta}))} d\theta$$

# Laplace Approximation (details)

$$\int_{0}^{1} w(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^{2}}{d\theta^{2}} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}} + O(T^{-3/2})$$

# Choosing the optimal prior

▶ Choose  $w(\theta)$  to maximize the worst-case final total weight

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▶ Make bound equal for all  $\hat{\theta} \in [0, 1]$  by choosing

$$w^*(\hat{\theta}) = \frac{1}{Z} \sqrt{\frac{\left. \frac{g^2}{d\theta^2} \right|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}{-2\pi}},$$

where **Z** is the normalization factor:

$$Z = \sqrt{rac{1}{2\pi}} \int_0^1 \left. \sqrt{rac{d^2}{d heta^2}} 
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# The bound for the optimal prior

Plugging in we get

$$L_{A} - L_{\min} \leq \ln \int_{0}^{1} w^{*}(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= \ln \left( \sqrt{\frac{2\pi Z}{T}} + O(T^{-3/2}) \right)$$

$$= \frac{1}{2} \ln \frac{T}{2\pi} - \frac{1}{2} \ln Z + O(1/T) .$$

#### Solving for log-loss

The exponent in the integral is

$$g(\hat{ heta}, heta) - g(\hat{ heta}, \hat{ heta}) = \hat{ heta} \ln \frac{\hat{ heta}}{ heta} + (1 - \hat{ heta}) \ln \frac{1 - \hat{ heta}}{1 - heta} = D_{KL}(\hat{ heta}|| heta)$$

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The second derivative

$$\left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} D_{KL}(\hat{\theta}||\theta) = \hat{\theta}(1 - \hat{\theta})$$

Is called the empirical Fisher information

#### Bayes using Jeffrey's prior

# Solving for log-loss

The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

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The optimal prior:

$$w^*(\hat{\theta}) = \frac{1}{\pi \sqrt{\hat{\theta}(1-\hat{\theta})}}$$

Known in general as Jeffrey's prior. And, in this case, the Dirichlet-(1/2, 1/2) prior.

# The cumulative log loss of Bayes using Jeffrey's prior

$$L_A - L_{\min} \le \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} + O(1/T)$$

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► This is called the Trichevsky Trofimov prediction rule.

Bayes using Jeffrey's prior

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Suffers larger regret when  $\hat{\theta}$  is far from 1/2

#### Shtarkov Lower bound

▶ What is the optimal prediction when T is know in advance?

#### ☐ Bayes using Jeffrey's prior

#### Shtarkov Lower bound

What is the optimal prediction when T is know in advance?

$$L_*^T - \min_{\theta} L_{\theta}^T \geq \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} - O(\frac{1}{\sqrt{T}})$$

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The constant C is optimal.

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