Combining infinite sets of experts

Yoav Freund

January 24, 2006

Review

Review

The Universal prediction machine

Review

The Universal prediction machine

The biased coins set of experts

Review

The Universal prediction machine

The biased coins set of experts

Bayes using Jeffrey's prior
Laplace Approximation
Choosing the optimal prior
Kritchevski Trofimov Prediction Rule
Laplace Rule of Succession

Review

The Universal prediction machine

The biased coins set of experts

Bayes using Jeffrey's prior
Laplace Approximation
Choosing the optimal prior
Kritchevski Trofimov Prediction Rule
Laplace Rule of Succession

Shtarkov lower bound for finite horizon

Review

The Universal prediction machine

The biased coins set of experts

Bayes using Jeffrey's prior
Laplace Approximation
Choosing the optimal prior
Kritchevski Trofimov Prediction Rule
Laplace Rule of Succession

Shtarkov lower bound for finite horizon

Generalization to larger sets of distributions

The online Bayes Algorithm

► Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

The online Bayes Algorithm

Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{s=1}^{t-1} p_i^s(c^s)$$

The online Bayes Algorithm

Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{s=1}^{t-1} p_i^s(c^s)$$

Freedom to choose initial weights.

$$w_t^1 \ge 0, \sum_{i=1}^n w_i^1 = 1$$

Review

The online Bayes Algorithm

Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{c=1}^{t-1} p_i^s(c^s)$$

Freedom to choose initial weights.

$$w_t^1 \ge 0, \sum_{i=1}^n w_i^1 = 1$$

Prediction of algorithm A

$$\mathbf{p}_{A}^{t} = \frac{\sum_{i=1}^{N} w_{i}^{t} \mathbf{p}_{i}^{t}}{\sum_{i=1}^{N} w_{i}^{t}}$$

Total weight: $W^t \doteq \sum_{i=1}^N w_i^t$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t}$$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t}$$

Review

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

Total weight:
$$W^t \doteq \sum_{i=1}^{N} w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

$$-\log \frac{W^{t+1}}{W^t} = -\log p_{\mathcal{A}}^t(c^t)$$

Total weight:
$$W^t \doteq \sum_{i=1}^{N} w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log \frac{W^{T+1}}{W^t} = -\sum_{t=1}^{T} \log p_A^t(c^t)$$

Total weight:
$$W^t \doteq \sum_{i=1}^{N} w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log W^{T+1} = -\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log W^{T+1} = -\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

EQUALITY not bound!

▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- ▶ Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1}$$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- ▶ Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1}$$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1} = -\log \sum_{i=1}^N w_i^{T+1}$$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1} = -\log \sum_{i=1}^N w_i^{T+1}$$
$$= -\log \sum_{i=1}^N w_i^1 e^{-L_i^T}$$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_{A}^{T} = -\log W^{T+1} = -\log \sum_{i=1}^{N} w_{i}^{T+1}$$

$$= -\log \sum_{i=1}^{N} w_{i}^{1} e^{-L_{i}^{T}} \le -\log \max_{i} \left(w_{i}^{1} e^{-L_{i}^{T}}\right)$$

- ▶ Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_{A}^{T} = -\log W^{T+1} = -\log \sum_{i=1}^{N} w_{i}^{T+1}$$

$$= -\log \sum_{i=1}^{N} w_{i}^{1} e^{-L_{i}^{T}} \le -\log \max_{i} \left(w_{i}^{1} e^{-L_{i}^{T}}\right)$$

$$= \min_{i} \left(L_{i}^{T} - \log w_{i}^{1}\right)$$

The Universal prediction machine

Standardizing online prediction algorithms

► Fix a universal Turing machine *U*.

The Universal prediction machine

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.

- Fix a universal Turing machine U.
- ▶ An online prediction algorithm *E* is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n*/*m*

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n*/*m*
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - To ensure p has a finite description. Restrict to rational numbers n/m
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.

- ► Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n*/*m*
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.
- ▶ $V(\vec{b}, \vec{X}, t) = 1$ if the program \vec{b} , given \vec{X} as input, halts within t steps and outputs a well-formed prediction. Otherwise $V(\vec{b}, \vec{X}, t) = 0$

Standardizing online prediction algorithms

- Fix a universal Turing machine U.
- ► An online prediction algorithm *E* is a program that
 - given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n*/*m*
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.
- ▶ $V(\vec{b}, \vec{X}, t) = 1$ if the program \vec{b} , given \vec{X} as input, halts within t steps and outputs a well-formed prediction. Otherwise $V(\vec{b}, \vec{X}, t) = 0$
- ▶ $V(\vec{b}, \vec{X}, t)$ is computable (recursively enumerable).

► Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| - \log_2 |\vec{b}|}$.

- Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- ► The total initial weight over all finite binary sequences is one.

- ► Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- ► The total initial weight over all finite binary sequences is one.
- ► Run the Bayes algorithm over "all" prediction algorithms.

- Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- ► The total initial weight over all finite binary sequences is one.
- ► Run the Bayes algorithm over "all" prediction algorithms.
- ▶ technical details: On iteration t, $|\vec{X}| = t$. Use the predictions of programs \vec{b} such that $|\vec{b}| \le t$ and for which $V(\vec{b}, \vec{X}, 2^t) = 1$. Assing the remaining mass the prediction 1/2 (insuring a loss of 1)

▶ Using $L_A \leq \min_i (L_i - \log w_i^1)$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- ► Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ► We get that the loss of the Universal algorithm is at most $2|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ▶ We get that the loss of the Universal algorithm is at most $2|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- More careful analysis can reduce $2|\vec{b}(E)| + \log_2|\vec{b}(E)|$ to $|\vec{b}(E)|$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ▶ We get that the loss of the Universal algorithm is at most $2|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- More careful analysis can reduce $2|\vec{b}(E)| + \log_2 |\vec{b}(E)|$ to $|\vec{b}(E)|$
- Code length is arbitrarily close to the Kolmogorov Complexity of the sequence.

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ▶ We get that the loss of the Universal algorithm is at most $2|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- More careful analysis can reduce $2|\vec{b}(E)| + \log_2 |\vec{b}(E)|$ to $|\vec{b}(E)|$
- Code length is arbitrarily close to the Kolmogorov Complexity of the sequence.
- Ridiculously bad running time.

Bayes coding is better than two part codes

 Simple bound as good as bound for two part codes (MDL) but enables online compression

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- ► Two part code has to point to one of the KN experts $L_A \le \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- ► Two part code has to point to one of the KN experts $L_A \le \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- ▶ If we use Bayes predictor + arithmetic coding we get:

$$L_A = -\log W^{T+1} \le \log K \max_i \frac{1}{NK} e^{-L_i^T} = \log N + \min_i L_i^T$$

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- ► Two part code has to point to one of the KN experts $L_A \le \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- ▶ If we use Bayes predictor + arithmetic coding we get:

$$L_A = -\log W^{T+1} \le \log K \max_i \frac{1}{NK} e^{-L_i^T} = \log N + \min_i L_i^T$$

We don't pay a penalty for copies.

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- ► Two part code has to point to one of the KN experts $L_A \le \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- ▶ If we use Bayes predictor + arithmetic coding we get:

$$L_A = -\log W^{T+1} \le \log K \max_i \frac{1}{NK} e^{-L_i^T} = \log N + \min_i L_i^T$$

- We don't pay a penalty for copies.
- More generally, the regret is smaller if many of the experts perform well.

► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- ▶ Instead, we assign the experts a Density Measure.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- Instead, we assign the experts a Density Measure.
- ▶ $L_A \le \min_i (L_i \log w_i^1)$ is meaningless.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- ▶ Instead, we assign the experts a Density Measure.
- ▶ $L_A \le \min_i (L_i \log w_i^1)$ is meaningless.
- Can we still get a meaningful bound?

Bayes Algorithm for biased coins

► Replace the initial weight by a density measure $w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$

Bayes Algorithm for biased coins

- ► Replace the initial weight by a density measure $w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$
- ► Relationship between final total weight and total log loss remains unchanged:

$$L_{A} = \ln \int_{0}^{1} w(\theta) e^{-L_{\theta}^{T+1}} d\theta$$

Bayes Algorithm for biased coins

- ► Replace the initial weight by a density measure $w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$
- Relationship between final total weight and total log loss remains unchanged:

$$L_A = \ln \int_0^1 w(\theta) e^{-L_{\theta}^{T+1}} d\theta$$

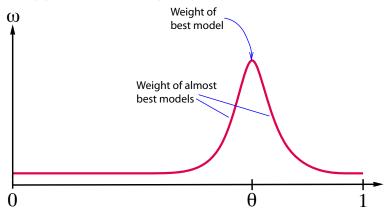
We need a new lower bound on the final total weight

Main Idea

If $\mathbf{w}^t(\theta)$ is large then $\mathbf{w}^t(\theta + \epsilon)$ is also large.

Main Idea

If $w^t(\theta)$ is large then $w^t(\theta + \epsilon)$ is also large.



 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{\mathbf{x}^t = 1; \ 1 \le t \le T\}}{T}$$

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

$$L_A - L_{\min} \le \ln \int_0^1 w(\theta) e^{-L_{\theta}} d\theta - \ln e^{L_{\min}}$$

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{\mathbf{x}^t = \mathbf{1}; \ \mathbf{1} \le t \le T\}}{T}$$

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

$$\begin{array}{lcl} L_A - L_{\min} & \leq & \ln \int_0^1 w(\theta) e^{-L_{\theta}} d\theta - \ln e^{L_{\min}} \\ \\ & = & \ln \int_0^1 w(\theta) e^{-(L_{\theta} - L_{\min})} d\theta \\ \\ pause & = & \ln \int_0^1 w(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta \end{array}$$

Laplace approximation (idea)

► Taylor expansion of $g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.

Laplace Approximation

Laplace approximation (idea)

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.

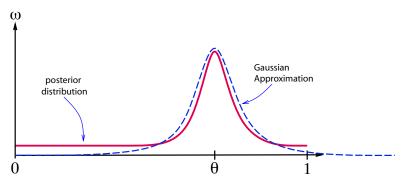
Laplace Approximation

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.



Laplace Approximation (details)

$$\int_0^1 w(\theta) e^{T(g(\hat{\theta},\theta)-g(\hat{\theta},\hat{\theta}))} d\theta$$

Laplace Approximation

Laplace Approximation (details)

$$\int_{0}^{1} w(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^{2}}{d\theta^{2}} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}} + O(T^{-3/2})$$

Choosing the optimal prior

Choosing the optimal prior

▶ Choose $w(\theta)$ to maximize the worst-case final total weight

$$\min_{\hat{\theta}} w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}}$$

Choosing the optimal prior

▶ Choose $w(\theta)$ to maximize the worst-case final total weight

$$\min_{\hat{\theta}} w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}}$$

▶ Make bound equal for all $\hat{\theta} \in [0, 1]$ by choosing

$$w^*(\hat{\theta}) = \frac{1}{Z} \sqrt{\frac{\frac{d^2}{d\theta^2}\Big|_{\theta=\hat{\theta}} (g(\hat{\theta},\theta) - g(\hat{\theta},\hat{\theta}))}{-2\pi}},$$

where **Z** is the normalization factor:

$$Z=\sqrt{rac{1}{2\pi}}\int_0^1\left.\sqrt{rac{d^2}{d heta^2}}
ight|_{ heta=\hat{ heta}}\left(g(\hat{ heta},\hat{ heta})-g(\hat{ heta}, heta)
ight)\;d\hat{ heta}$$

The bound for the optimal prior

Plugging in we get

$$L_{A} - L_{\min} \leq \ln \int_{0}^{1} w^{*}(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= \ln \left(\sqrt{\frac{2\pi Z}{T}} + O(T^{-3/2}) \right)$$

$$= \frac{1}{2} \ln \frac{T}{2\pi} - \frac{1}{2} \ln Z + O(1/T) .$$

Solving for log-loss

The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

Solving for log-loss

The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

The second derivative

$$\left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} D_{KL}(\hat{\theta}||\theta) = \hat{\theta}(1 - \hat{\theta})$$

Is called the empirical Fisher information

Solving for log-loss

► The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

The second derivative

$$\left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} D_{KL}(\hat{\theta}||\theta) = \hat{\theta}(1 - \hat{\theta})$$

Is called the empirical Fisher information

The optimal prior:

$$w^*(\hat{\theta}) = \frac{1}{\pi \sqrt{\hat{\theta}(1-\hat{\theta})}}$$

Known in general as Jeffrey's prior. And, in this case, the Dirichlet-(1/2, 1/2) prior.

Choosing the optimal prior

The cumulative log loss of Bayes using Jeffrey's prior

$$L_A - L_{\min} \le \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} + O(1/T)$$

But what is the prediction rule?

➤ As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.

⁻ Bayes using Jeffrey's prior

Kritchevski Trofimov Prediction Bule

Kritchevski Trofimov Prediction Rule

But what is the prediction rule?

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed t bits, n of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

But what is the prediction rule?

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed t bits, n of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

The posterior average is:

$$\frac{\int_0^1 \theta^{n+1/2} (1-\theta)^{t-n-1/2} d\theta}{\int_0^1 \theta^{n-1/2} (1-\theta)^{t-n-1/2} d\theta} = \frac{n+1/2}{t+1}$$

Bayes using Jeffrey's prior

But what is the prediction rule?

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed t bits, n of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

The posterior average is:

$$\frac{\int_0^1 \theta^{n+1/2} (1-\theta)^{t-n-1/2} d\theta}{\int_0^1 \theta^{n-1/2} (1-\theta)^{t-n-1/2} d\theta} = \frac{n+1/2}{t+1}$$

This is called the Trichevsky Trofimov prediction rule.

Laplace suggested using the uniform prior, which is also a conjugate prior.

Laplace Rule of Succession

Laplace Rule of Succession

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

$$\frac{\int_0^1 \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_0^1 \theta^n (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

Laplace Rule of Succession

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

$$\frac{\int_{0}^{1} \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_{0}^{1} \theta^{n} (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

▶ The bound on the cumulative log loss is worse:

$$L_A - L_{\min} = \ln T + O(1)$$

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

$$\frac{\int_{0}^{1} \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_{0}^{1} \theta^{n} (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

▶ The bound on the cumulative log loss is worse:

$$L_A - L_{\min} = \ln T + O(1)$$

Suffers larger regret when $\hat{\theta}$ is far from 1/2

Shtarkov Lower bound

▶ What is the optimal prediction when *T* is know in advance?

Shtarkov Lower bound

What is the optimal prediction when T is know in advance?

$$L_*^T - \min_{\theta} L_{\theta}^T \geq \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} - O(\frac{1}{\sqrt{T}})$$

► For a distribution over *k* elements (Multinomial) [Xie and Barron]

- ► For a distribution over *k* elements (Multinomial) [Xie and Barron]
- ▶ Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

- ► For a distribution over *k* elements (Multinomial) [Xie and Barron]
- ▶ Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

Bound is

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T + C + o(1)$$

- ► For a distribution over k elements (Multinomial) [Xie and Barron]
- ▶ Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

Bound is

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T + C + o(1)$$

► The constant C is optimal.

Exponential Distributions

► For any set of distributions from the exponential family defined by *k* parameters (Some technical conditions on closure of set??) [Rissanen??]

Exponential Distributions

- ► For any set of distributions from the exponential family defined by *k* parameters (Some technical conditions on closure of set??) [Rissanen??]
- Use Bayes Algorithm with Jeffrey's prior:

$$w^*(\hat{\theta}) = \frac{1}{Z} \frac{1}{\sqrt{\left|\mathbf{H}\left(D_{KL}(\hat{\theta}||\theta)\right)\right|_{\theta=\hat{\theta}}}}$$

H denotes the Hessian.

Exponential Distributions

- ► For any set of distributions from the exponential family defined by *k* parameters (Some technical conditions on closure of set??) [Rissanen??]
- Use Bayes Algorithm with Jeffrey's prior:

$$w^*(\hat{\theta}) = \frac{1}{Z} \frac{1}{\sqrt{\left|\mathbf{H}\left(D_{KL}(\hat{\theta}||\theta)\right)\right|_{\theta=\hat{\theta}}}}$$

H denotes the Hessian.

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T - \ln Z + o(1)$$

General Distributions

Characterize distribution family by metric entropy.

General Distributions

- Characterize distribution family by metric entropy.
- Fixed parameter set usually corresponds to polynomial metrix entropy

$$N(1/\epsilon) = O\left(\frac{1}{\epsilon^d}\right)$$

d is the number of parameters.

General Distributions

- Characterize distribution family by metric entropy.
- Fixed parameter set usually corresponds to polynomial metrix entropy

$$N(1/\epsilon) = O\left(\frac{1}{\epsilon^d}\right)$$

d is the number of parameters.

► [Haussler and Opper] show that the coefficient in front of In *T* is optimal for distribution families where the metric entropy is up to

$$N(1/\epsilon) = O(e^{\epsilon^{-\alpha}})$$

For all $\alpha \leq 5/2$.

next Class

Variable-length markov models - a set of distributions with increasing number of parameters.

next Class

- Variable-length markov models a set of distributions with increasing number of parameters.
- THe context algorithm: An efficient implementation of the Bayes algorithm which achieves close-to-optimal worst case bounds.