

AFFINE DIFFUSION MODELS*

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1 Affine Diffusions

A diffusion process is a Markov process solving the stochastic differential equation

$$dY_t = \mu(Y_t, \theta_0) dt + \sigma(Y_t, \theta_0) dW_t.$$

A discount-rate function $R : D \rightarrow \mathbb{R}$ is an affine function of the state

$$R(Y) = \rho_0 + \rho_1 \cdot Y,$$

for $\rho = (\rho_0, \rho_1) \in \mathbb{R} \times \mathbb{R}^N$. The affine dependence of the drift and diffusion coefficients of Y are determined by coefficients (K, H) defined by:

- $\mu(Y) = K_0 + K_1 Y$, for $K = (K_0, K_1) \in \mathbb{R}^N \times \mathbb{R}^{N \times N}$,
- $[\sigma(Y) \sigma(Y)']_{ij} = [H_0]_{ij} + [H_1]_{ij} \cdot Y$, for $H = (H_0, H_1) \in \mathbb{R}^{N \times N} \times \mathbb{R}^{N \times N \times N}$.

Here

$$[H_1]_{ij} \cdot Y = \sum_{k=1}^N [H_1]_{ijk} Y_k.$$

A characteristic $\chi = (K, H, \rho)$ captures both the distribution of Y as well as the effects of any discounting.

Suppose we have “intraday” observations $Y_{t+\frac{j-1}{n}h, t+\frac{j}{n}h}$ for $j = 1, \dots, n$. These observations can be used to compute realized variance (RV) that approximates true integrated variance better than just squared “daily” observations $Y_{t,t+h}^2$:

$$RV_{t,h} \equiv \frac{1}{h} \sum_{j=1}^n Y_{t+\frac{j-1}{n}h, t+\frac{j}{n}h}^2 \xrightarrow{a.s.} \frac{1}{h} \int_t^{t+h} \sigma^2(Y_s, \theta_0) ds = \frac{1}{h} \int_t^{t+h} d[Y, Y]_s \equiv \mathcal{V}_{t,h}.$$

*Python implementation can be found here: <https://github.com/khrapovs/diffusions>

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2 Geometric Brownian Motion

2.1 The model

Suppose that S_t evolves according to

$$\frac{dS_t}{S_t} = \mu dt + \sigma dW_t.$$

In logs:

$$d \log S_t = \left(\mu - \frac{1}{2} \sigma^2 \right) dt + \sigma dW_t.$$

2.2 AJD representation

Since the model belongs to affine class, we have

$$K_0 = \mu - \frac{1}{2} \sigma^2, \quad K_1 = 0, \quad H_0 = \sigma^2, \quad H_1 = 0.$$

2.3 Exact solution

The solution is

$$r_{t,h} = \frac{1}{h} \log \frac{S_{t+h}}{S_t} = \left(\mu - \frac{1}{2} \sigma^2 \right) + \frac{1}{h} \sigma (W_{t+h} - W_t).$$

In other words,

$$r_{t,h} \sim N \left(\mu - \frac{1}{2} \sigma^2, \sigma^2 \right).$$

2.4 Moments

After integration on the interval $[t, t+H]$:

$$r_{t,H} = \frac{1}{H} \log \frac{S_{t+H}}{S_t} = \mu - \frac{1}{2} \sigma^2 + \frac{\sigma}{\sqrt{H}} \varepsilon_{t+H},$$

where $\varepsilon_t \sim N(0, 1)$. The first conditional moment is

$$E_t[r_{t,H}] = \mu - \frac{1}{2} \sigma^2.$$

The second conditional moment is

$$V_t[r_{t,H}] = E_t[r_{t,H}^2] - (E_t[r_{t,H}])^2 = \sigma^2.$$

Hence, the moment function is

$$g(r_{t,H}; \theta) = \begin{bmatrix} r_{t,H} - \left(\mu - \frac{1}{2} \sigma^2 \right) \\ r_{t,H}^2 - \sigma^2 - \left(\mu - \frac{1}{2} \sigma^2 \right)^2 \end{bmatrix}$$

with

$$E_t [g(r_{t,H}; \theta)] = 0.$$

The derivative of the moment function with respect to model parameters $\theta = [\mu, \sigma]$ is

$$\frac{\partial g}{\partial \theta}(r_{t,H}; \theta) = \begin{bmatrix} -1 & \sigma \\ -2\left(\mu - \frac{1}{2}\sigma^2\right) & -2\sigma + 2\left(\mu - \frac{1}{2}\sigma^2\right)\sigma \end{bmatrix}.$$

2.5 Integrated moments

It is trivial to deduce that the integrated variance is constant

$$\mathcal{V}_{t,H} = \sigma^2.$$

Hence, the moment function is instead

$$g(r_{t,H}, RV_{t,H}; \theta) = X_{t,H} - C = \begin{bmatrix} r_{t,H} \\ RV_{t,H} \\ RV_{t,H}^2 \end{bmatrix} - \begin{bmatrix} \mu - \frac{1}{2}\sigma^2 \\ \sigma^2 \\ \sigma^4 \end{bmatrix}.$$

Here we used the fact that $V_t[\mathcal{V}_{t,H}] = 0$ and $(E_t[\mathcal{V}_{t,H}])^2 = \sigma^4$.

3 Vasicek model

3.1 The model

Consider

$$dr_t = \kappa(\mu - r_t)dt + \sigma dW_t.$$

3.2 AJD representation

Since the model belongs to affine class, we have

$$K_0 = \kappa\mu, \quad K_1 = -\kappa, \quad H_0 = \sigma^2, \quad H_1 = 0.$$

3.3 Exact solution

Using Ito's lemma for $r_t e^{\kappa t}$ we have

$$d(r_t e^{\kappa t}) = (\kappa r_t e^{\kappa t} + \kappa(\mu - r_t) e^{\kappa t}) dt + \sigma e^{\kappa t} dW_t,$$

or

$$d(r_t e^{\kappa t}) = \kappa \mu e^{\kappa t} dt + \sigma e^{\kappa t} dW_t.$$

After integration we have

$$r_{t+h}e^{\kappa(t+h)} = r_te^{\kappa t} + \kappa\mu \int_t^{t+h} e^{\kappa u} du + \sigma \int_t^{t+h} e^{\kappa u} dW_u,$$

or

$$\begin{aligned} r_{t+h} &= r_te^{-\kappa h} + \kappa\mu \int_t^{t+h} e^{-\kappa(t+h-u)} du + \sigma \int_t^{t+h} e^{-\kappa(t+h-u)} dW_u \\ &= r_te^{-\kappa h} + \mu \left(1 - e^{-\kappa h}\right) + \sigma \int_t^{t+h} e^{-\kappa(t+h-u)} dW_u. \end{aligned}$$

This implies that

$$r_{t+h}|r_t \sim N \left(r_te^{-\kappa h} + \mu \left(1 - e^{-\kappa h}\right), \sigma^2 \int_t^{t+h} e^{-2\kappa(t+h-u)} du = \frac{\sigma^2}{2\kappa} \left(1 - e^{-2\kappa h}\right) \right).$$

4 Cox-Ingersoll-Ross model

4.1 The model

Cox, Ingersoll, and Ross (1985; Ecta) developed a theory of the term structure of interest rates in which the instantaneous short-term rate of interest, r , follows the mean reverting diffusion

$$dr_t = \kappa(\mu - r_t) dt + \sigma\sqrt{r_t}dW_t.$$

Feller condition for positivity of the process is $\kappa\mu > \frac{1}{2}\sigma^2$.

4.2 AJD representation

Since the model belongs to affine class, we have

$$K_0 = \kappa\mu, \quad K_1 = -\kappa, \quad H_0 = 0, \quad H_1 = \sigma^2.$$

4.3 The solution

Using Ito's lemma for $r_te^{\kappa t}$ we have

$$r_{t+h} = r_te^{-\kappa h} + \mu \left(1 - e^{-\kappa h}\right) + \sigma \int_t^{t+h} e^{-\kappa(t+h-u)} \sqrt{r_u} dW_u.$$

It can be shown that the transition density is known in closed form:

$$f(r_{t+h}|r_t; \theta) = ce^{-u-v} \left(\frac{u}{v}\right)^{q/2} I_q(2\sqrt{uv}),$$

where

$$c = \frac{2\kappa}{\sigma^2(1 - e^{-\kappa h})}, \quad u = cr_te^{-\kappa h}, \quad v = cr_{t+h}, \quad q = \frac{2\kappa\mu}{\sigma^2} - 1,$$

and I_q is the modified Bessel function of the first kind of order q . Sometimes, it is useful to work with a transformation $x_{t+h} = 2cr_{t+h}$. We can derive that the transition density of x_{t+h} is

$$f(x_{t+h}|x_t; \theta) = f(2cr_{t+h}|2cr_t; \theta) = \frac{1}{2c} f(r_{t+h}|r_t; \theta),$$

which is a noncentral χ^2 with $2q + 2$ degrees of freedom and noncentrality parameter $2u$.

4.4 Moments

This implies that

$$\begin{aligned} E_t[r_{t+h}] &= r_t e^{-\kappa h} + \mu (1 - e^{-\kappa h}), \\ V_t[r_{t+h}] &= \frac{\sigma^2}{\kappa} r_t e^{-\kappa h} (1 - e^{-\kappa h}) + \frac{\sigma^2}{2\kappa} \mu (1 - e^{-\kappa h})^2, \end{aligned}$$

since

$$\begin{aligned} V_t[r_{t+h}] &= \sigma^2 \int_t^{t+h} e^{-2\kappa(t+h-u)} E_t[r_u] du \\ &= \sigma^2 \int_t^{t+h} e^{-2\kappa(t+h-u)} (r_t e^{-\kappa(u-t)} + \mu (1 - e^{-\kappa(u-t)})) du. \end{aligned}$$

5 Heston model

5.1 The model

Consider

$$\begin{aligned} dp_t &= \left(r + \left(\lambda_r - \frac{1}{2} \right) \sigma_t^2 \right) dt + \sigma_t dW_t^r, \\ d\sigma_t^2 &= \kappa (\mu - \sigma_t^2) dt + \eta \sigma_t dW_t^\sigma, \end{aligned}$$

with $p_t = \log S_t$, and $\text{Corr}[dW_s^r, dW_s^\sigma] = \rho$, or in other words $W_t^\sigma = \rho W_t^r + \sqrt{1 - \rho^2} W_t^v$. Also let $R(Y_t) = r$. Feller condition is

$$2\kappa\mu > \eta^2.$$

5.2 Risk-neutral model

Let the log stochastic discount factor (SDF) process $m_t = \log M_t$ be represented by the following SDE:

$$dm_t = -r dt - \zeta_r dW_t^r - \zeta_v dW_t^v,$$

with

$$\begin{bmatrix} \zeta_r \\ \zeta_v \end{bmatrix} = \begin{bmatrix} \lambda_r \sigma_t \\ -\frac{1}{\sqrt{1-\rho^2}} (\rho \lambda_r + \lambda_\sigma) \sigma_t \end{bmatrix}.$$

The risk-neutral innovations are then

$$\begin{aligned} d\tilde{W}_t^r &= dW_t^r + \zeta_r dt, \\ d\tilde{W}_t^v &= dW_t^v + \zeta_v dt. \end{aligned}$$

After the substitution

$$\begin{aligned} dp_t &= \left(r - \frac{1}{2}\sigma_t^2\right) dt + \sigma_t d\tilde{W}_t^r, \\ d\sigma_t^2 &= \tilde{\kappa} \left(\tilde{\mu} - \sigma_t^2\right) dt + \eta \sigma_t d\tilde{W}_t^\sigma, \end{aligned}$$

with $\tilde{W}_t^\sigma = \rho \tilde{W}_t^r + \sqrt{1 - \rho^2} \tilde{W}_t^v$, and the modified parameters are

$$\tilde{\kappa} = \kappa - \lambda_\sigma \eta, \quad \tilde{\mu} = \mu \frac{\kappa}{\tilde{\kappa}}.$$

Feller condition $2\tilde{\kappa}\tilde{\mu} > \eta^2$ is equivalent to the one for physical measure, $2\mu\kappa > \eta^2$. Still, we need $\tilde{\kappa} > 0$ which is equivalent to $\lambda_\sigma < \kappa/\eta$.

5.3 AJD representation

The drift and diffusion functions are

$$\mu(Y_t, \theta_0) = \begin{bmatrix} r \\ \kappa\mu \end{bmatrix} + \begin{bmatrix} 0 & \lambda_r - \frac{1}{2} \\ 0 & -\kappa \end{bmatrix} \begin{bmatrix} p_t \\ \sigma_t^2 \end{bmatrix}, \quad \sigma(Y_t, \theta_0) = \begin{bmatrix} \sigma_t & 0 \\ \eta\rho\sigma_t & \eta\sqrt{1 - \rho^2}\sigma_t \end{bmatrix}.$$

It follows that

$$\sigma(Y_t, \theta_0) \sigma(Y_t, \theta_0)' = \begin{bmatrix} \sigma_t^2 & \eta\rho\sigma_t^2 \\ \eta\rho\sigma_t^2 & \eta^2\sigma_t^2 \end{bmatrix} = \begin{bmatrix} 1 & \eta\rho \\ \eta\rho & \eta^2 \end{bmatrix} \sigma_t^2.$$

So we have

$$\rho_0 = r, \quad \rho_1 = [0, 0],$$

and

$$K_0 = \begin{bmatrix} r \\ \kappa\mu \end{bmatrix}, \quad K_1 = \begin{bmatrix} 0 & \lambda_r - \frac{1}{2} \\ 0 & -\kappa \end{bmatrix},$$

and

$$H_0 = \begin{bmatrix} 0 & 0 \\ 0 & 0 \end{bmatrix}, \quad H_{1,1} = \begin{bmatrix} 0 & 0 \\ 0 & 0 \end{bmatrix}, \quad H_{1,2} = \begin{bmatrix} 1 & \eta\rho \\ \eta\rho & \eta^2 \end{bmatrix}.$$

5.4 Integrated moments

Define average excess return as

$$\tilde{r}_{t,H} \equiv \frac{1}{H} \int_t^{t+H} dp_s - r.$$

Given observations on average excess returns $\tilde{r}_{t,H}$ and realized volatility $RV_{t,H}$ we have the following

four integrated moments (the proof is given in section §A):

$$\begin{aligned}
E_t [(1 - A_h L) \mathcal{V}_{t+h,H}] &= C_1, \\
E_t [(1 - A_h L) (1 - A_h^2 L) \mathcal{V}_{t+2h,H}^2] &= C_2, \\
E_t \left[\tilde{r}_{t,H} - \left(\lambda_r - \frac{1}{2} \right) \mathcal{V}_{t,H} \right] &= C_3 = 0, \\
E_t \left[(1 - A_h L) \left(\tilde{r}_{t+h,H} \mathcal{V}_{t+h,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t+h,H}^2 \right) \right] &= C_4,
\end{aligned}$$

or in matrix notation

$$E_{t-2h} \left[\begin{pmatrix} (1 - A_h L) L & \cdot & \cdot & \cdot \\ \cdot & (1 - A_h L) (1 - A_h^2 L) & \cdot & \cdot \\ \left(\frac{1}{2} - \lambda_r \right) L^2 & \cdot & L^2 & \cdot \\ \cdot & \left(\frac{1}{2} - \lambda_r \right) (1 - A_h L) L & \cdot & (1 - A_h L) L \end{pmatrix} \begin{pmatrix} \mathcal{V}_{t,H} \\ \mathcal{V}_{t,H}^2 \\ \tilde{r}_{t,H} \\ \tilde{r}_{t,H} \mathcal{V}_{t,H} \end{pmatrix} \right] = C.$$

Define the data vector as $X_{t,H} = [\mathcal{V}_{t,H}, \mathcal{V}_{t,H}^2, \tilde{r}_{t,H}, \tilde{r}_{t,H} \mathcal{V}_{t,H}]$. Then, the moment conditions can be written more compactly as

$$E_{t-2h} [A(L) X_{t,H}] = C,$$

with the lag polynomial

$$A(L) = A_0 + A_1 L + A_2 L^2,$$

and

$$A_0 = \begin{pmatrix} \cdot & \cdot & \cdot & \cdot \\ \cdot & 1 & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \end{pmatrix}, \quad A_1 = \begin{pmatrix} 1 & \cdot & \cdot & \cdot \\ \cdot & -A_h (1 + A_h) & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \frac{1}{2} - \lambda_r & \cdot & 1 \end{pmatrix}, \quad A_2 = \begin{pmatrix} -A_h & \cdot & \cdot & \cdot \\ \cdot & A_h^3 & \cdot & \cdot \\ \frac{1}{2} - \lambda_r & \cdot & 1 & \cdot \\ \cdot & \left(\lambda - \frac{1}{2} \right) A_h & \cdot & -A_h \end{pmatrix}.$$

To make this one single matrix product, we can write

$$E_{t-2h} \left[\begin{pmatrix} A_0 & A_1 & A_2 \end{pmatrix} \begin{pmatrix} X_{t,H} \\ L X_{t,H} \\ L^2 X_{t,H} \end{pmatrix} \right] = C,$$

where

$$C = (A_0 + A_1 + A_2) E[X_{t,H}],$$

and

$$\begin{aligned}
E[\mathcal{V}_{t,H}] &= \mu, \\
E[\mathcal{V}_{t,H}^2] &= \frac{\eta^2}{\kappa^2} \frac{c_H}{H} + \mu^2, \\
E[\tilde{r}_{t,H}] &= \left(\lambda_r - \frac{1}{2}\right) \mu, \\
E[\tilde{r}_{t,H} \mathcal{V}_{t,H}] &= \rho \frac{\eta}{\kappa} \frac{c_H}{H} + \left(\lambda_r - \frac{1}{2}\right) E[\mathcal{V}_{t,H}^2].
\end{aligned}$$

6 Central Tendency

6.1 The model

Consider

$$\begin{aligned}
dp_t &= \left(r + \left(\lambda_r - \frac{1}{2}\right) \sigma_t^2\right) dt + \sigma_t dW_t^r, \\
d\sigma_t^2 &= \kappa_\sigma \left(y_t^2 - \sigma_t^2\right) dt + \eta_\sigma \sigma_t dW_t^\sigma, \\
dy_t^2 &= \kappa_y \left(\mu - y_t^2\right) dt + \eta_y v_t dW_t^y,
\end{aligned}$$

with $p_t = \log S_t$, and $\text{Corr}[dW_s^r, dW_s^\sigma] = \rho$, or in other words $W_t^\sigma = \rho W_t^r + \sqrt{1 - \rho^2} W_t^y$. Also let $R(Y_t) = r$.

6.2 Risk-neutral model

Let the log stochastic discount factor (SDF) process $m_t = \log M_t$ be represented by the following SDE:

$$dm_t = -r dt - \zeta_r dW_t^r - \zeta_v dW_t^v - \zeta_y dW_t^y.$$

with

$$\begin{bmatrix} \zeta_r \\ \zeta_v \\ \zeta_y \end{bmatrix} = \begin{bmatrix} \lambda_r \sigma_t \\ -\frac{1}{\sqrt{1-\rho^2}} (\rho \lambda_r + \lambda_\sigma) \sigma_t \\ -\lambda_y y_t \end{bmatrix}.$$

The risk-neutral innovations are then

$$\begin{aligned}
d\tilde{W}_t^r &= dW_t^r + \zeta_r dt, \\
d\tilde{W}_t^v &= dW_t^v + \zeta_v dt, \\
d\tilde{W}_t^y &= dW_t^y + \zeta_y dt.
\end{aligned}$$

After the substitution

$$\begin{aligned} dp_t &= \left(r - \frac{1}{2} \sigma_t^2 \right) dt + \sigma_t d\tilde{W}_t^r, \\ d\sigma_t^2 &= \tilde{\kappa}_\sigma \left(\frac{\kappa_\sigma}{\tilde{\kappa}_\sigma} y_t^2 - \sigma_t^2 \right) dt + \eta_\sigma \sigma_t d\tilde{W}_t^\sigma, \\ dy_t^2 &= \tilde{\kappa}_y \left(\tilde{\mu} - y_t^2 \right) dt + \eta_y y_t d\tilde{W}_t^y, \end{aligned}$$

with $\tilde{W}_t^\sigma = \rho \tilde{W}_t^r + \sqrt{1 - \rho^2} \tilde{W}_t^v$, and the modified parameters are

$$\tilde{\kappa}_\sigma = \kappa_\sigma - \lambda_\sigma \eta_\sigma, \quad \tilde{\kappa}_y = \kappa_y - \lambda_y \eta_y, \quad \tilde{\mu} = \mu \frac{\kappa_y}{\tilde{\kappa}_y}.$$

Feller condition $2\tilde{\kappa}\tilde{\mu} > \eta^2$ is equivalent to the one for physical measure, $2\mu\kappa > \eta^2$. Still, we need $\tilde{\kappa} > 0$ which is equivalent to $\lambda_\sigma < \kappa/\eta$.

Note that the unconditional means of the processes at hand under the physical measure are the same:

$$E^P [\sigma_t^2] = E^P [y_t^2] = \mu,$$

while under the risk-neutral measure, they are different:

$$E^Q [y_t^2] = \tilde{\mu} = \mu \frac{\kappa_y}{\kappa_y - \lambda_y \eta_y} > \mu,$$

and

$$E^Q [\sigma_t^2] = E^Q [y_t^2] \frac{\kappa_\sigma}{\tilde{\kappa}_\sigma} = \mu \frac{\kappa_y}{\kappa_y - \lambda_y \eta_y} \frac{\kappa_\sigma}{\kappa_\sigma - \lambda_\sigma \eta_\sigma} > \tilde{\mu}.$$

Given that all of the parameters above are non-negative, we can write the following inequality:

$$E^P [\sigma_t^2] = E^P [y_t^2] \leq E^Q [y_t^2] \leq E^Q [\sigma_t^2].$$

6.3 AJD representation

The drift and diffusion functions are

$$\mu(Y_t, \theta_0) = \begin{bmatrix} r \\ 0 \\ \kappa_y \mu_y \end{bmatrix} + \begin{bmatrix} 0 & \lambda_r - \frac{1}{2} & 0 \\ 0 & -\kappa_\sigma & \kappa_\sigma \\ 0 & 0 & -\kappa_y \end{bmatrix} \begin{bmatrix} p_t \\ \sigma_t^2 \\ y_t^2 \end{bmatrix}, \quad \sigma(Y_t, \theta_0) = \begin{bmatrix} \sigma_t & 0 & 0 \\ \eta_\sigma \rho \sigma_t & \eta_\sigma \sqrt{1 - \rho^2} \sigma_t & 0 \\ 0 & 0 & \eta_y y_t \end{bmatrix}.$$

It follows that

$$\sigma(Y_t, \theta_0) \sigma(Y_t, \theta_0)' = \begin{bmatrix} \sigma_t^2 & \eta_\sigma \rho \sigma_t^2 & 0 \\ \eta_\sigma \rho \sigma_t^2 & \eta_\sigma^2 \sigma_t^2 & 0 \\ 0 & 0 & \eta_y^2 y_t^2 \end{bmatrix} = \begin{bmatrix} 1 & \eta_\sigma \rho & 0 \\ \eta_\sigma \rho & \eta_\sigma^2 & 0 \\ 0 & 0 & 0 \end{bmatrix} \sigma_t^2 + \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & \eta_y^2 \end{bmatrix} y_t^2.$$

So we have

$$\rho_0 = r, \quad \rho_1 = [0, 0, 0],$$

and

$$K_0 = \begin{bmatrix} r \\ 0 \\ \kappa_y \mu_y \end{bmatrix}, \quad K_1 = \begin{bmatrix} 0 & \lambda_r - \frac{1}{2} & 0 \\ 0 & -\kappa_\sigma & \kappa_\sigma \\ 0 & 0 & -\kappa_y \end{bmatrix},$$

and

$$H_0 = [\mathbf{0}_{3 \times 3}], \quad H_{1,1} = [\mathbf{0}_{3 \times 3}], \quad H_{1,2} = \begin{bmatrix} 1 & \eta_\sigma \rho & 0 \\ \eta_\sigma \rho & \eta_\sigma^2 & 0 \\ 0 & 0 & 0 \end{bmatrix}, \quad H_{1,3} = \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & \eta_y^2 \end{bmatrix}.$$

6.4 Integrated moments

Define average excess return as

$$\tilde{r}_{t,H} \equiv \frac{1}{H} \int_t^{t+H} dp_s - r.$$

Given observations on average excess returns $\tilde{r}_{t,H}$ and realized volatility $RV_{t,H}$ we have the following four integrated moments (the proof is given in section §B):

$$\begin{aligned} E_t [(1 - A_h^y L) (1 - A_h^\sigma L) \mathcal{V}_{t+2h,H}] &= C_1, \\ E_t [(1 - A_{2h}^\sigma L) (1 - A_{2h}^y L) (1 - A_h^\sigma A_h^y L) (1 - A_h^y L) (1 - A_h^\sigma L) \mathcal{V}_{t+5h,H}^2] &= C_2, \\ E_t \left[\tilde{r}_{t,H} - \left(\lambda_r - \frac{1}{2} \right) \mathcal{V}_{t,H} \right] &= C_3 = 0, \\ E_t \left[(1 - A_h^y L) (1 - A_h^\sigma L) \left(\tilde{r}_{t+h,H} \mathcal{V}_{t+h,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t+h,H}^2 \right) \right] &= C_4, \end{aligned}$$

or in matrix notation

$$E_{t-5h} \left[\begin{pmatrix} M_1 L^3 & \cdot & \cdot & \cdot \\ \cdot & M_2 & \cdot & \cdot \\ \left(\frac{1}{2} - \lambda_r \right) L^5 & \cdot & L^5 & \cdot \\ \cdot & \left(\frac{1}{2} - \lambda_r \right) M_1 L^3 & \cdot & M_1 L^3 \end{pmatrix} \begin{pmatrix} \mathcal{V}_{t,H} \\ \mathcal{V}_{t,H}^2 \\ \tilde{r}_{t,H} \\ \tilde{r}_{t,H} \mathcal{V}_{t,H} \end{pmatrix} \right] = C,$$

where

$$\begin{aligned} M_1 &= (1 - A_h^y L) (1 - A_h^\sigma L), \\ M_2 &= M_1 (1 - A_h^\sigma A_h^y L) (1 - A_{2h}^y L) (1 - A_{2h}^\sigma L), \end{aligned}$$

Define the data vector as $X_{t,H} = [\mathcal{V}_{t,H}, \mathcal{V}_{t,H}^2, \tilde{r}_{t,H}, \tilde{r}_{t,H} \mathcal{V}_{t,H}]$. Then, the moment conditions can be written more compactly as

$$E_{t-5h} [A(L) X_{t,H}] = C,$$

with the lag polynomial

$$A(L) = A_0 + A_1L + A_2L^2 + A_3L^3 + A_4L^4 + A_5L^5,$$

and

$$\begin{aligned} A_0 &= \begin{pmatrix} \cdot & \cdot & \cdot & \cdot \\ \cdot & m_{2,0} & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \end{pmatrix}, \quad A_1 = \begin{pmatrix} \cdot & \cdot & \cdot & \cdot \\ \cdot & m_{2,1} & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \end{pmatrix}, \quad A_2 = \begin{pmatrix} \cdot & \cdot & \cdot & \cdot \\ \cdot & m_{2,2} & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \end{pmatrix}, \\ A_3 &= \begin{pmatrix} m_{1,0} & \cdot & \cdot & \cdot \\ \cdot & m_{2,3} & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \frac{1}{2} - \lambda_r & \cdot & m_{1,0} \end{pmatrix}, \quad A_4 = \begin{pmatrix} m_{1,1} & \cdot & \cdot & \cdot \\ \cdot & m_{2,4} & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \left(\frac{1}{2} - \lambda_r\right) m_{1,1} & \cdot & m_{1,1} \end{pmatrix}, \\ A_5 &= \begin{pmatrix} m_{1,2} & \cdot & \cdot & \cdot \\ \cdot & m_{2,5} & \cdot & \cdot \\ \frac{1}{2} - \lambda_r & \cdot & 1 & \cdot \\ \cdot & \left(\frac{1}{2} - \lambda_r\right) m_{1,2} & \cdot & m_{1,2} \end{pmatrix}. \end{aligned}$$

Here

$$m_{1,0} = 1, \quad m_{1,1} = -(A_h^y + A_h^\sigma), \quad m_{1,2} = A_h^y A_h^\sigma,$$

and for the second polynomial the coefficients are found accordingly to the relation between polynomial roots and coefficients¹. To make this one single matrix product, we can write

$$E_{t-2h} \begin{bmatrix} \begin{pmatrix} A_0 & A_1 & A_2 & A_3 & A_4 & A_5 \end{pmatrix} \begin{pmatrix} X_{t,H} \\ LX_{t,H} \\ L^2 X_{t,H} \\ L^3 X_{t,H} \\ L^4 X_{t,H} \\ L^5 X_{t,H} \end{pmatrix} \end{bmatrix} = C,$$

where

$$C = (A_0 + A_1 + A_2 + A_3 + A_4 + A_5) E[X_{t,H}],$$

¹See e.g.: <http://math.stackexchange.com/questions/88917/relation-between-coefficients-and-roots-of-a-polynomial>

and

$$\begin{aligned}
E[\mathcal{V}_{t,H}] &= \mu, \\
E[\mathcal{V}_{t,H}^2] &= \mu^2 + (a_H^\sigma)^2 V[\sigma_t^2] + (b_H^\sigma)^2 V[y_t^2] + V\left[\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds\right], \\
E[\tilde{r}_{t,H}] &= \left(\lambda_r - \frac{1}{2}\right) \mu, \\
E[\tilde{r}_{t,H} \mathcal{V}_{t,H}] &= \rho \frac{\mu \eta_\sigma}{\kappa_\sigma} \frac{1 - a_H^\sigma}{H} + \left(\lambda_r - \frac{1}{2}\right) E[\mathcal{V}_{t,H}^2].
\end{aligned}$$

Appendix

A Integrated moments for Heston model

The spot volatility equation can be integrated on the interval $[t, t+h]$ with h possibly different from H to get

$$\sigma_{t+h}^2 = C_h + A_h \sigma_t^2 + \epsilon_{t,h}^\sigma, \quad (\text{A.1})$$

where I define coefficients and the integrated innovation as

$$A_h = \exp(-\kappa h), \quad C_h = \mu(1 - A_h), \quad \epsilon_{t,h}^\sigma = \eta_\sigma \int_t^{t+h} \sigma_u A_{t+h-u} dW_u^\sigma.$$

The spot volatility first moment is

$$E_t [\sigma_{t+h}^2] = C_h + A_h \sigma_t^2,$$

or, using lag operator L ,

$$E_t [(1 - A_h L) \sigma_{t+h}^2] = (1 - A_h) E [\sigma_t^2] = (1 - A_h) \mu.$$

Given the above definitions it is also not hard to guess that σ_t^4 is autoregressive:

$$E_t [(1 - A_h L) (1 - A_h^2 L) \sigma_{t+2h}^4] = C_1 = (1 - A_h) (1 - A_h^2) E [\sigma_t^4]$$

with

$$E [\sigma_t^4] = V [\sigma_t^2] + (E [\sigma_t^2])^2 = \frac{\mu \eta^2}{2\kappa} + \mu^2.$$

After repeated integration of the volatility equation (A.1) over t as a dummy of the integration on the interval $[t, t+H]$ we obtain

$$\mathcal{V}_{t+h,H} = C_h + A_h \mathcal{V}_{t,H} + \frac{1}{H} \int_0^H \epsilon_{t+s,h}^\sigma ds.$$

Hence, the integrated volatility first moment is

$$E_t [(1 - A_h L) \mathcal{V}_{t+h,H}] = (1 - A_h) E [\mathcal{V}_{t,H}] = (1 - A_h) \mu.$$

We can also integrate the volatility equation (A.1) using h as a dummy of integration on the interval from $[0, H]$:

$$\mathcal{V}_{t,H} = c_H + a_H \sigma_t^2 + \frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds, \quad (\text{A.2})$$

with

$$a_H = \frac{1}{H} \int_0^H A_s ds = \frac{1}{\kappa H} (1 - A_H), \quad c_H = \mu (1 - a_H).$$

The error term may be simplified:

$$\begin{aligned}
\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds &= \frac{1}{H} \int_0^H \left(\eta \int_t^{t+s} \sigma_u A_{t+s-u} dW_u^\sigma \right) ds \\
&= \frac{\eta}{H} \int_t^{t+H} \sigma_u \left(\int_{u-t}^H A_{t+s-u} ds \right) dW_u^\sigma \\
&= \frac{\eta}{H} \int_t^{t+H} \sigma_u \left(\int_0^{t+H-u} A_s ds \right) dW_u^\sigma \\
&= \eta \int_t^{t+H} \sigma_u a_{t+H-u} dW_u^\sigma.
\end{aligned}$$

From this we can guess that

$$E_t \left[(1 - A_h L) \left(1 - A_h^2 L \right) \mathcal{V}_{t+2h,H}^2 \right] = C_2 = (1 - A_h) \left(1 - A_h^2 \right) E \left[\mathcal{V}_{t,H}^2 \right].$$

From the equation (A.2) for $\mathcal{V}_{t,H}$ we can derive

$$\begin{aligned}
V[\mathcal{V}_{t,H}] &= a_H^2 V[\sigma_t^2] + V \left[\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds \right] = a_H^2 \frac{\mu \eta^2}{2\kappa} + V \left[\eta \int_t^{t+H} \sigma_u a_{t+H-u} dW_u^\sigma \right] \\
&= a_H^2 \frac{\mu \eta^2}{2\kappa} + E \left[\eta^2 \int_t^{t+H} \sigma_u^2 a_{t+H-u}^2 du \right] = a_H^2 \frac{\mu \eta^2}{2\kappa} + \mu \eta^2 \int_0^H a_{H-u}^2 du \\
&= a_H^2 \frac{\mu \eta^2}{2\kappa} + \frac{\mu \eta^2}{\kappa^2 H^2} \int_0^H (1 - A_{H-u})^2 du \\
&= (1 - A_H)^2 \frac{\mu \eta^2}{2\kappa^3 H^2} + \frac{\mu \eta^2}{\kappa^2 H^2} \left(H - \frac{2}{\kappa} (1 - A_H) + \frac{1}{2\kappa} (1 - A_{2H}) \right) \\
&= \frac{\mu \eta^2}{\kappa^2 H^2} \left[(1 - 2A_H + A_{2H}) \frac{1}{2\kappa} + H - \frac{2}{\kappa} (1 - A_H) + \frac{1}{2\kappa} (1 - A_{2H}) \right] \\
&= \frac{\mu \eta^2}{\kappa^2 H^2} \left[(1 - A_H) \frac{1}{\kappa} + H - \frac{2}{\kappa} (1 - A_H) \right] \\
&= \frac{\mu \eta^2}{\kappa^2 H^2} \left[H - \frac{1}{\kappa} (1 - A_H) \right] = \frac{\mu \eta^2}{\kappa^2 H} (1 - a_H) = \frac{\eta^2}{\kappa^2} \frac{c_H}{H}.
\end{aligned}$$

Hence,

$$E[\mathcal{V}_{t,H}^2] = V[\mathcal{V}_{t,H}] + (E[\mathcal{V}_{t,H}])^2 = \frac{\eta^2}{\kappa^2} \frac{c_H}{H} + \mu^2.$$

The return equation can be integrated on $[t, t+H]$ to obtain

$$\tilde{r}_{t,h} \equiv \frac{1}{H} \int_t^{t+H} dp_s - r = \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H} + \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r,$$

The return first conditional moment is

$$E_t[\tilde{r}_{t,H}] = \left(\lambda - \frac{1}{2} \right) E_t[\mathcal{V}_{t,H}],$$

or $\frac{\{\kappa_{\sigma}\}\{\kappa_{\sigma}-\kappa_y\}}{\{\kappa_{\sigma}\}\{\kappa_{\sigma}-\kappa_y\}}$

$$E_t \left[\tilde{r}_{t,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H} \right] = C_3 = 0.$$

Now multiply return by integrated volatility and take the expectation:

$$E_t \left[\tilde{r}_{t,H} \mathcal{V}_{t,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H}^2 \right] = E_t \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right].$$

We can guess that

$$E_t \left[(1 - A_h L) \left(\tilde{r}_{t,H} \mathcal{V}_{t,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H}^2 \right) \right] = (1 - A_h) E \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right].$$

The last term is

$$\begin{aligned} E \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right] &= E \left[\frac{1}{H^2} \int_0^H \epsilon_{t,s}^\sigma ds \int_t^{t+H} \sigma_s dW_s^r \right] \\ &= \frac{\eta}{H} E \left[\left(\int_t^{t+H} \sigma_u a_{t+H-u} dW_u^\sigma \right) \left(\int_t^{t+H} \sigma_s dW_s^r \right) \right] \\ &= \rho \frac{\eta}{H} E \left[\int_t^{t+H} \sigma_u^2 a_{t+H-u} du \right] = \rho \frac{\eta}{H} \int_t^{t+H} E \left[\sigma_u^2 \right] a_{t+H-u} du. \\ &= \rho \frac{\mu \eta}{H} \int_0^H a_{H-u} du = \rho \frac{\mu \eta}{H^2} \int_0^H \frac{1}{\kappa} (1 - A_{H-u}) du \\ &= \rho \frac{\mu \eta}{H^2} \left(H - \frac{1}{\kappa} (1 - A_H) \right) = \rho \frac{\mu \eta}{\kappa H} (1 - a_H) = \rho \frac{\eta}{\kappa} \frac{c_H}{H}. \end{aligned}$$

B Integrated moments for Central Tendency model

The volatility processes can be discretized to obtain

$$\begin{aligned} \sigma_{t+h}^2 &= A_h^\sigma \sigma_t^2 + B_h^\sigma y_t^2 + C_h^\sigma + \epsilon_{t,h}^\sigma, \\ y_{t+h}^2 &= A_h^y y_t^2 + C_h^y + \epsilon_{t,h}^y. \end{aligned} \tag{B.1}$$

with

$$A_h^\sigma = \exp(-\kappa_\sigma h), \quad B_h^\sigma = \frac{\kappa_\sigma}{\kappa_\sigma - \kappa_y} (A_h^y - A_h^\sigma), \quad C_h^\sigma = \mu (1 - A_h^\sigma - B_h^\sigma),$$

and

$$A_h^y = \exp(-\kappa_y h), \quad C_h^y = \mu (1 - A_h^y).$$

The same system can be written using lag operators as

$$\begin{aligned} (1 - A_h^\sigma L) \sigma_{t+h}^2 &= B_h^\sigma y_t^2 + C_h^\sigma + \epsilon_{t,h}^\sigma, \\ (1 - A_h^y L) y_{t+h}^2 &= C_h^y + \epsilon_{t,h}^y. \end{aligned}$$

Note that A_h^y and A_h^σ are multiplicative functions of time interval, that is $A_{h_1}^y A_{h_2}^y = A_{h_1+h_2}^y$. The error structure is represented by

$$\begin{aligned}\epsilon_{t,h}^\sigma &= \eta_\sigma \int_t^{t+h} \sigma_u A_{t+h-u}^\sigma dW_u^\sigma + \eta_y \int_t^{t+h} y_u B_{t+h-u}^\sigma dW_u^y, \\ \epsilon_{t,h}^y &= \eta_y \int_t^{t+h} y_u A_{t+h-u}^y dW_u^y.\end{aligned}$$

Clearly, $E_t^P [\epsilon_{t,t+h}^\sigma] = 0$, and $E_t^P [\epsilon_{t,t+h}^y] = 0$. Note that the same processes may be represented as infinite stochastic integrals with respect to Brownian motion only:

$$\begin{aligned}y_t^2 &= \mu + \eta_y \int_{-\infty}^t y_u A_{t-u}^y dW_u^y, \\ \sigma_t^2 &= \mu + \eta_y \int_{-\infty}^t y_u B_{t-u}^\sigma dW_u^y + \eta_\sigma \int_{-\infty}^t \sigma_u A_{t-u}^\sigma dW_u^\sigma.\end{aligned}\tag{B.2}$$

Next I define integrated variance and central tendency as

$$\mathcal{V}_{t,H} \equiv \frac{1}{H} \int_t^{t+H} \sigma_u^2 du, \quad \mathcal{Y}_{t,H} \equiv \frac{1}{H} \int_t^{t+H} y_u^2 du,\tag{B.3}$$

where the first subscripted value denotes the beginning of the time interval, and the second denotes the length of this interval.

In order to move from instantaneous vector (σ_t^2, y_t) to integrated analog $(\mathcal{V}_{t,H}, \mathcal{Y}_{t,H})$, I integrate the linear system in equation (B.1) over t as a dummy of the integration in the interval $[0, H]$ with the following result

$$\begin{aligned}\mathcal{V}_{t+h,H} &= A_h^\sigma \mathcal{V}_{t,H} + B_h^\sigma \mathcal{Y}_{t,H} + C_h^\sigma + \frac{1}{H} \int_0^H \epsilon_{t+s,h}^\sigma ds, \\ \mathcal{Y}_{t+h,H} &= A_h^y \mathcal{Y}_{t,H} + C_h^y + \frac{1}{H} \int_0^H \epsilon_{t+s,h}^y ds.\end{aligned}\tag{B.4}$$

Using the lag operator L and taking the conditional expectation, this system may be written as

$$\begin{aligned}E_t^P [(1 - A_h^\sigma L) \mathcal{V}_{t+h,H}] &= B_h^\sigma E_t^P [\mathcal{Y}_{t,H}] + C_h^\sigma, \\ E_t^P [(1 - A_h^y L) \mathcal{Y}_{t+h,H}] &= C_h^y.\end{aligned}$$

Multiply the first equation by $(1 - A_h^y L)$, shift the time by h , and make a substitution from the second equation to obtain

$$E_t^P [(1 - A_h^y L) (1 - A_h^\sigma L) \mathcal{V}_{t+2h,H}] = C_1,$$

where

$$C_2 = (1 - A_h^y) (1 - A_h^\sigma) E^P [\mathcal{V}_{t,H}] = (1 - A_h^y) (1 - A_h^\sigma) \mu.$$

In equation (B.1) replace h by another time indicator s and integrate from 0 to H which leads to the

following expression for integrated volatility in terms of spot variables

$$\mathcal{V}_{t,H} = c_H^\sigma + a_H^\sigma \sigma_t^2 + b_H^\sigma y_t^2 + \frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds,$$

where I denote

$$\begin{aligned} a_H^\sigma &= \frac{1}{H} \int_0^H A_s^\sigma ds = \frac{1}{\kappa_\sigma H} (1 - A_H^\sigma), \\ b_H^\sigma &= \frac{1}{H} \int_0^H B_s^\sigma ds = \frac{\kappa_\sigma}{\kappa_\sigma - \kappa_y} \frac{1}{H} \int_0^H (A_s^y - A_s^\sigma) ds = \frac{\kappa_\sigma}{\kappa_\sigma - \kappa_y} (a_H^y - a_H^\sigma), \\ c_H^\sigma &= \mu (1 - a_H^\sigma - b_H^\sigma). \end{aligned}$$

Clearly, the second moment of $\mathcal{V}_{t,H}$ will be a function of σ_t^2 , y_t^2 , σ_t^4 , y_t^4 , and $\sigma_t^2 y_t^2$. In order to eliminate the first two, we will need to apply $(1 - A_h^\sigma L)$ and $(1 - A_h^y L)$. The squared central tendency y_t^4 is a function of itself in the past and itself squared, so it can be eliminated using $(1 - A_h^y L) \left(1 - (A_h^y)^2 L\right)$. The square volatility will be a function of all the above and can be eliminated using $\left(1 - (A_h^\sigma)^2 L\right) (1 - A_h^\sigma A_h^y L) \left(1 - (A_h^y)^2 L\right) (1 - A_h^y L) (1 - A_h^\sigma L)$. Hence,

$$E_t^P \left[\left(1 - (A_h^\sigma)^2 L\right) \left(1 - (A_h^y)^2 L\right) (1 - A_h^\sigma A_h^y L) (1 - A_h^y L) (1 - A_h^\sigma L) \mathcal{V}_{t+5h,H}^2 \right] = C_2,$$

where

$$C_3 = \left(1 - (A_h^\sigma)^2\right) (1 - A_h^\sigma A_h^y) \left(1 - (A_h^y)^2\right) (1 - A_h^y) (1 - A_h^\sigma) E^P \left[\mathcal{V}_{t,H}^2 \right].$$

The unconditional variance of integrated volatility is

$$V[\mathcal{V}_{t,H}] = (a_H^\sigma)^2 V[\sigma_t^2] + (b_H^\sigma)^2 V[y_t^2] + V\left[\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds\right].$$

Directly from (B.2) we see that

$$E[y_t^4] = \mu \eta_y^2 \int_{-\infty}^t (A_{t-u}^y)^2 du = \mu \eta_y^2 \int_{-\infty}^t e^{-2\kappa_y(t-u)} du = \frac{\mu \eta_y^2}{2\kappa_y},$$

and

$$\begin{aligned}
E[\sigma_t^4] &= \mu\eta_y^2 \int_{-\infty}^t (B_{t-u}^\sigma)^2 du + \mu\eta_\sigma^2 \int_{-\infty}^t (A_{t-u}^\sigma)^2 du \\
&= \mu\eta_y^2 \left(\frac{\kappa_\sigma}{\kappa_\sigma - \kappa_y} \right)^2 \int_{-\infty}^t \left(A_{2(t-u)}^y - 2A_{t-u}^y A_{t-u}^\sigma + A_{2(t-u)}^\sigma \right) du + \mu\eta_\sigma^2 \int_{-\infty}^t A_{2(t-u)}^\sigma du \\
&= \mu\eta_y^2 \left(\frac{\kappa_\sigma}{\kappa_\sigma - \kappa_y} \right)^2 \left(\frac{1}{2\kappa_y} - \frac{2}{\kappa_y + \kappa_\sigma} + \frac{1}{2\kappa_\sigma} \right) + \frac{\mu\eta_\sigma^2}{2\kappa_\sigma} \\
&= \frac{\mu\eta_y^2}{2\kappa_y} \frac{\kappa_\sigma}{\kappa_y + \kappa_\sigma} + \frac{\mu\eta_\sigma^2}{2\kappa_\sigma} \\
&= E[y_t^4] \frac{\kappa_\sigma}{\kappa_y + \kappa_\sigma} + \frac{\mu\eta_\sigma^2}{2\kappa_\sigma}.
\end{aligned}$$

Rewrite the error:

$$\begin{aligned}
\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds &= \eta_\sigma \frac{1}{H} \int_0^H \int_t^{t+s} \sigma_u A_{t+s-u}^\sigma dW_u^\sigma ds + \eta_y \frac{1}{H} \int_0^H \int_t^{t+s} y_u B_{t+s-u}^\sigma dW_u^y ds \\
&= \eta_\sigma \int_t^{t+H} \sigma_u a_{t+H-u}^\sigma dW_u^\sigma + \eta_y \int_t^{t+H} y_u b_{t+H-u}^\sigma dW_u^y \\
&= \eta_\sigma \int_0^H \sigma_u a_{H-u}^\sigma dW_u^\sigma + \eta_y \int_0^H y_u b_{H-u}^\sigma dW_u^y.
\end{aligned}$$

Hence,

$$V \left[\frac{1}{H} \int_0^H \epsilon_{t,s}^\sigma ds \right] = \mu\eta_\sigma^2 \int_0^H (a_{H-u}^\sigma)^2 du + \mu\eta_y^2 \int_0^H (b_{H-u}^\sigma)^2 du$$

The return equation can be integrated on $[t, t+H]$ to obtain

$$\tilde{r}_{t,h} \equiv \frac{1}{H} \int_t^{t+H} dp_s - r = \left(\lambda_r - \frac{1}{2} \right) \mathcal{V}_{t,H} + \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r,$$

The return first conditional moment is

$$E_t[\tilde{r}_{t,H}] = \left(\lambda - \frac{1}{2} \right) E_t[\mathcal{V}_{t,H}],$$

or

$$E_t \left[\tilde{r}_{t,H} - \left(\lambda_r - \frac{1}{2} \right) \mathcal{V}_{t,H} \right] = C_3 = 0.$$

Now multiply return by integrated volatility and take the expectation:

$$E_t \left[\tilde{r}_{t,H} \mathcal{V}_{t,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H}^2 \right] = E_t \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right].$$

We can guess that

$$E_t \left[(1 - A_h^y L) (1 - A_h^\sigma L) \left(\tilde{r}_{t,H} \mathcal{V}_{t,H} - \left(\lambda - \frac{1}{2} \right) \mathcal{V}_{t,H}^2 \right) \right] = C_4,$$

where

$$C_4 = (1 - A_h^y) (1 - A_h^\sigma) E \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right].$$

The last term is

$$\begin{aligned} E^P \left[\mathcal{V}_{t,H} \frac{1}{H} \int_t^{t+H} \sigma_s dW_s^r \right] &= \frac{1}{H^2} E^P \left[\int_0^H \epsilon_{t,s}^\sigma ds \int_t^{t+H} \sigma_s dW_s^r \right] \\ &= \eta_\sigma \frac{1}{H^2} E^P \left[\left(\int_0^H \int_t^{t+s} \sigma_u A_{t+s-u}^\sigma dW_u^\sigma ds \right) \left(\int_t^{t+H} \sigma_s dW_s^r \right) \right] \\ &= \eta_\sigma \frac{1}{H^2} E^P \left[\left(\int_t^{t+H} \sigma_u \left(\int_{u-t}^h A_{t+s-u}^\sigma ds \right) dW_u^\sigma \right) \left(\int_t^{t+H} \sigma_u dW_u^r \right) \right] \\ &= \rho \frac{\eta_\sigma}{\kappa_\sigma} \frac{1}{H^2} E^P \left[\int_t^{t+H} \sigma_u^2 (1 - A_{t+H-u}^\sigma) du \right] \\ &= \rho \frac{\mu \eta_\sigma}{\kappa_\sigma} \frac{1}{H^2} \int_0^H (1 - A_{H-u}^\sigma) du \\ &= \rho \frac{\mu \eta_\sigma}{\kappa_\sigma} \frac{1}{H^2} \left(H - \frac{1}{\kappa^\sigma} (1 - A_H^\sigma) \right) \\ &= \rho \frac{\mu \eta_\sigma}{\kappa_\sigma} \frac{1 - a_H^\sigma}{H}. \end{aligned}$$