This section will briefly walk through two general concepts, one econometric and one R oriented. We will examine the characteristics of generalized least squares (GLS), and specifically the efficiency gains from a special case of GLS, weighted least squares (WLS). This is the econometric concepts. We will then recreate the graphs from Figures 2.6 and 2.7, roughly, in the notes using ggplot2 a very popular, very powerful graphing package in R. This is part is optional, especially since it is only a very brief treatment of the package — there is a lot more to learn.

Let $x \sim U(0, 2000)$ and $\epsilon \sim N(0, (x/1000)^2)$. The underlying data generating process in (2.102) is $y_i = \alpha + x_i \beta + \epsilon$, where $\alpha = 0.5$ and $\beta = 1.5$. The objective is to plot the simulated sampling distribution of the OLS estimator applied to B = 10,000 draws, each of size n = 1000. First, let's generate the sample data for one draw.

```
n <- 1000
x <- runif(n, min=0, max=2000)
eps <- rnorm(n, 0, sqrt((x/1000)^2))
y <- 0.5 + x*1.5 + eps</pre>
```

Now we can calculate the standard OLS parameter vector $[\hat{\alpha} \ \hat{\beta}]'$ by noting that **X** is just the *x* vector bound to a column of ones. We will only examine $\hat{\beta}$ for this section, rather than both parameters.

```
X <- cbind(1, x)
params <- solve(t(X) %*% X) %*% t(X) %*% y
beta <- params[2]
print(beta)
[1] 1.500036</pre>
```

Let's package this into a function, called $\mathtt{rnd.beta}$, so that we can collect the OLS parameter for an arbitrary number of random samples, noting that n is a constant so we may as well keep it out of the function so that 1000 is not reassigned thousands of times to n.

```
rnd.beta <- function(i) {
    x <- runif(n)
    eps <- rnorm(n, 0, sqrt(x/10))
    y <- 0.5 + x*1.5 + eps
    X <- cbind(1, x)
    params <- solve(t(X) %*% X) %*% t(X) %*% y
    beta <- params[2]
    return(beta)
}</pre>
```

Since there aren't any supplied arguments, the function will return an estimated $\hat{\beta}$ from a different random sample for each call:

```
rnd.beta()
rnd.beta()
[1] 1.465081
[1] 1.487995
```

This is convenient for bootstrapping without loops, but rather applying the function to a list of effective indices. Now replicating the process for B draws is straightforward:

¹This is much more comfortable for me, with a background in functional programming. There is some inherent value, however, in keeping code compact by mapping across indices rather than incrementing an index within a for loop; the code is more readable and less prone to typos.

```
B <- 1000
beta.vec <- sapply(1:B, rnd.beta)
mean(beta.vec)
[1] 1.500485</pre>
```

Alright. Looking good. The average of the simulated sample is much closer to β than any individual call of rnd.beta, suggesting that the distribution of the simulated parameters will be appropriately centered. Now, let's create another, similar function that returns the WLS estimates.

```
rnd.wls.beta <- function(i) {
    x <- runif(n)
    y <- 0.5 + x*1.5 + rnorm(n, 0, sqrt(x/10))
    C <- diag(1/sqrt(x/10))
    y.wt <- C %*% y
    X.wt <- C %*% cbind(1, x)
    param.wls <- solve(t(X.wt) %*% X.wt) %*% t(X.wt) %*% y.wt
    beta <- param.wls[2]
    return(beta)
}
wls.beta.vec <- sapply(1:B, rnd.wls.beta)</pre>
```

We now have two collections of parameter estimates, one based on OLS and another based on WLS. It is straightforward to plot two separate histograms using R's core histogram plotting function hist(). However, we can use this to introduce a more flexible, powerful graphing package called ggplot2.

```
library(ggplot2)
labels <- c(rep("ols", B), rep("wls", B))
data <- data.frame(beta=c(beta.vec, wls.beta.vec), method=labels)
ggplot(data, aes(x=beta, fill=method)) + geom_density(alpha=0.2)</pre>
```

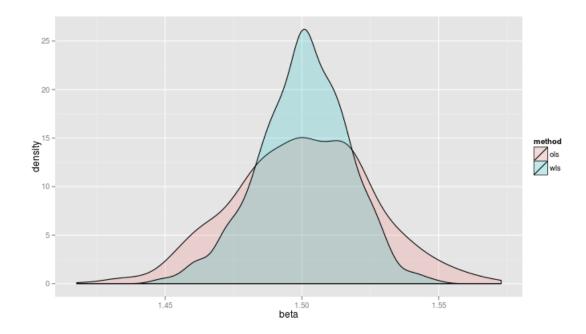


Figure 1: Relative efficiency of WLS