

# PROBABILITY THEORY

## MM 4

### MM 4: Special probability distributions

Topics: Discrete distributions:  
Bernoulli, binomial, geometric, Poisson

# Quizz

Suppose  $X$  has mean 2 and variance 3

- Compute  $\text{Var}(3X)$ 
  - 3
  - 9
  - 27
- Compute  $\text{Var}(3X+4)$ 
  - 27
  - 31
  - 43
- Compute  $E[X^2]$ 
  - 4
  - 7
- Compute  $\text{Var}(X^2)$ 
  - 9
  - Can not be computed from information given

# Quizz

- Can the following be negative:
  - A r.v.  $X$
  - $E[X]$
  - $E[X^2]$
  - $\text{Var}(X)$
  - $\text{Cov}(X, Y)$

# Quiz

- If you have a pdf (pmf), can you find a cdf?
- Is the opposite true?

# The $n$ th moment of r.v. $X$

- The mean value is sometimes called the first moment of  $X$ .
- **Definition.** The  $n$ th moment of  $X$  is defined by

$$E[X^n] = \sum_i x_i^n p(x_i)$$

$$E[X^n] = \int_{-\infty}^{\infty} x^n f(x) dx$$

# Moment generating function

- **Definition.** The **moment generating function** of a r.v.  $X$  is defined for all values  $t$  by

$$\varphi(t) = E[e^{tX}] = \sum_i e^{tx_i} p(x_i)$$

$$\varphi(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

- This function is called moment generating function, because all moments of  $X$  can be obtained by successively differentiating it.
- There is a one-to-one correspondence between the moment generating function and the distribution function of a r.v.: the mgf uniquely determines the distribution and vice versa.

# How to obtain $E[X]$ and $\text{Var}(X)$ from moment generating function?

$$\begin{aligned}\varphi'(t) &= \frac{d}{dt} E[e^{tX}] = E\left[\frac{d}{dt}(e^{tX})\right] = \\ &= E[X e^{tX}]\end{aligned}$$

$$\varphi'(0) = E[X]$$

$$\begin{aligned}\varphi''(t) &= \frac{d}{dt} E[X e^{tX}] = E\left[\frac{d}{dt}(X e^{tX})\right] = \\ &= E[X^2 e^{tX}]\end{aligned}$$

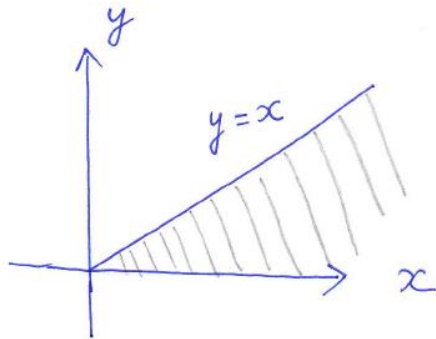
$$\varphi''(0) = E[X^2]$$

$$\text{Var}(X) = \varphi''(0) - \varphi'(0)^2$$

# Example: covariance

joint pdf for  $X$  and  $Y$ :

$$f_{X,Y}(x,y) = \begin{cases} c e^{-x} e^{-y}, & 0 \leq y < x < \infty \\ 0, & \text{otherwise} \end{cases}$$



The constant  $c$  is found from the normalization condition:

$$1 = \int_0^{+\infty} \int_0^x c e^{-x} e^{-y} dy dx = \frac{c}{2} \quad \Rightarrow \quad c=2$$



# Example: covariance (ctnd)

The marginal pdfs:

$$f_X(x) = \int_0^x 2e^{-x}e^{-y} dy = 2e^{-x}(1-e^{-x}), \quad x \in [0, \infty)$$

$$f_Y(y) = \int_0^\infty 2e^{-x}e^{-y} dx = 2e^{-2y}, \quad y \in [0, \infty)$$

$$E[X] = \frac{3}{2}, \quad \text{Var}(X) = \frac{5}{4}$$

$$E[Y] = \frac{1}{2}, \quad \text{Var}(Y) = \frac{1}{4}$$

# Example: covariance (ctnd)

$$\begin{aligned} E[XY] &= \int_0^{\infty} \int_0^x xy \cdot 2e^{-x}e^{-y} dy dx = \\ &= \int_0^{\infty} 2xe^{-x}(1 - e^{-x} - xe^{-x}) dx = 1 \end{aligned}$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 1 - \frac{3}{4} = \frac{1}{4}$$

$$\text{Corr}(X, Y) = \frac{1/4}{\sqrt{5/4} \sqrt{1/4}} = \frac{1}{\sqrt{5}}$$

# What should we learn today?

- A number of r.v.s. arises in many diverse, unrelated applications → learn them
- What are the main facts about these distributions:
  - Their cdf and pdf
  - Mean and variance
  - Moment generating function
- In which situations the special distributions arise and how are they interrelated?

# Bernoulli r.v.

- An experiment with 2 possible outcomes (success or failure) is called a Bernoulli trial.
- The indicator of event  $A$  is called the Bernoulli r.v., since it describes outcome of a Bernoulli trial.

$$E[I] = p$$

$$Var(I) = p(1 - p)$$

- Every Bernoulli trial, regardless of the definition of  $A$ , is equivalent to the tossing of a biased coin.

# Example: indicator of an event

- Example: a r.v.  $I$  is an indicator of the event  $A$
- *The expected value of the indicator of an event is equal to the probability of the event*

$$I = \begin{cases} 1, & \text{if } A \text{ occurs} \\ 0, & \text{if } A \text{ does not occur} \end{cases}$$

$$E[I] = 1 \cdot P(A) + 0 \cdot P(A^c) = P(A)$$

# Example: indicator of an event

$$I = \begin{cases} 1, & \text{if event } A \text{ occurs} \\ 0, & \text{if event } A \text{ does not occur} \end{cases}$$

$$\text{Var}(I) = E[I^2] - E[I]^2 = E[I] - E[I]^2 =$$

$$I^2 = I$$

$$= E[I](1 - E[I]) = P(A)(1 - P(A)) = P(A)P(A^c)$$

# Binomial r.v.

- Let  $X$  be a r.v. representing the number of times a certain event  $A$  occurs in  $n$  trials (number of success in  $n$  trials). Denote  $p$  probability of success.
- Then  $X$  is said to be a binomial r.v. with parameters  $n$  and  $p$

$$X \sim b(n, p)$$

- Pmf

$$P\{X = i\} = \binom{n}{p} p^i (1 - p)^{n-i}, \quad n = 0, 1, \dots, n$$

$$E[X] = np$$

$$Var(X) = np(1 - p)$$

# Binomial r.v.: mean and variance

$$X = I_1 + I_2 + \dots + I_n$$

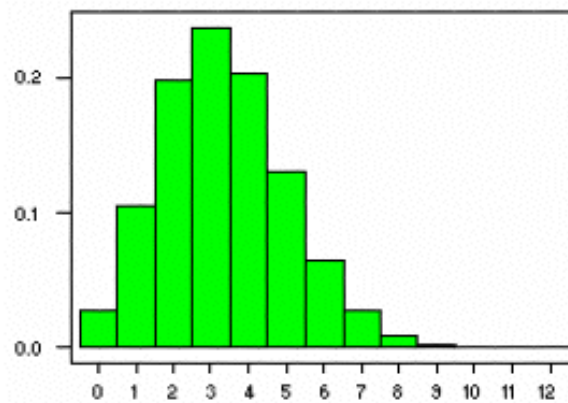
where  $I_i$  is a Bernoulli r.v.

$$\begin{aligned} E[X] &= E[I_1 + I_2 + \dots + I_n] = E[I_1] + E[I_2] + \dots + E[I_n] = \\ &= \underbrace{p + p + \dots + p}_{n \text{ times}} = np \end{aligned}$$

$$\begin{aligned} \text{Var}(X) &= \text{Var}(I_1 + \dots + I_n) = \text{Var}(I_1) + \dots + \text{Var}(I_n) = \\ &= \underbrace{p(1-p) + \dots + p(1-p)}_{n \text{ times}} = np(1-p) \end{aligned}$$



# Binomial r.v.



Parameters  $n=20$ ,  $p=1/6$

# Computing the binomial distribution function

$$P\{X \leq i\} = \sum_{k=0}^i P\{X=k\} = \sum_{k=0}^i \binom{n}{k} p^k (1-p)^{n-k}$$

- To optimize the computational process, we can utilize the following relationship between  $P\{X=k+1\}$  and  $P\{X=k\}$ :

$$P\{X = k + 1\} = \frac{p}{1-p} \frac{n-k}{k+1} P\{X = k\}$$

- Now starting from  $P(X=0)$  we can recursively find all other  $P(X=k)$

# Example: n-component system

- A system consists of  $n$  components. Each component functions independently with probability  $p$ .
- Question: for which values of  $p$  is a 5-component system more likely to operate than a 3-component system?

For 5-component system:

$$\begin{aligned} P_1 &= P\{X \geq 3\} = P\{X=3\} + P\{X=4\} + P\{X=5\} = \\ &= 1 - P\{X < 3\} = 1 - P\{X=0\} - P\{X=1\} - P\{X=2\} = \\ &= \binom{5}{3} p^3 (1-p)^2 + \binom{5}{4} p^4 (1-p) + \binom{5}{5} p^5 (1-p)^0 \end{aligned}$$

# Example (cntd)

For 3-component system:

$$\begin{aligned} P_2 &= P\{Y \geq 2\} = P\{Y=2\} + P\{Y=3\} = \\ &= \binom{3}{2} p^2(1-p) + \binom{3}{3} p^3(1-p)^0 \end{aligned}$$

We would like to choose  $p$  such that  $P_1 \geq P_2$

$$\Rightarrow 10p^3(1-p)^2 + 5p^4(1-p) + p^5 \geq 3p^2(1-p) + p^3$$

$$3(1-p)^2(2p-1) \geq 0$$

$$p \geq \frac{1}{2}$$

# Why is it called binomial?

N / k	0	1	2	3	4
1	p	q			
2	$p^2$	$2pq$	$q^2$		
3	$p^3$	$3p^2q$	$3pq^2$	$q^3$	
4	$p^4$	$4p^3q$	$6p^2q^2$	$4pq^3$	$q^4$

Coefficients are the same as in the binomial expansion  $(p+q)^n$

# The geometric r.v.

- Let a r.v.  $M$  be a number of independent Bernoulli trials until the first occurrence of a success.  $M$  is called a geometric r.v.
- Pmf

$$P\{M = k\} = (1 - p)^{k-1}p$$

$$E[M] = \frac{1}{p} \qquad Var(M) = \frac{1 - p}{p^2}$$

- Example: applications where we are interested in the time that elapses between the occurrence of events in a sequence of independent experiments

# Geometric r.v.: mean and variance

$$P\{M = k\} = (1 - p)^{k-1} p$$

$$E[M] = \sum_{k=1}^{+\infty} k p (1-p)^{k-1}$$

$$\sum_{k=0}^{\infty} x^k = \frac{1}{1-x}$$

$$\sum_{k=0}^{+\infty} k x^{k-1} = \frac{1}{(1-x)^2}$$

$$x = 1-p$$

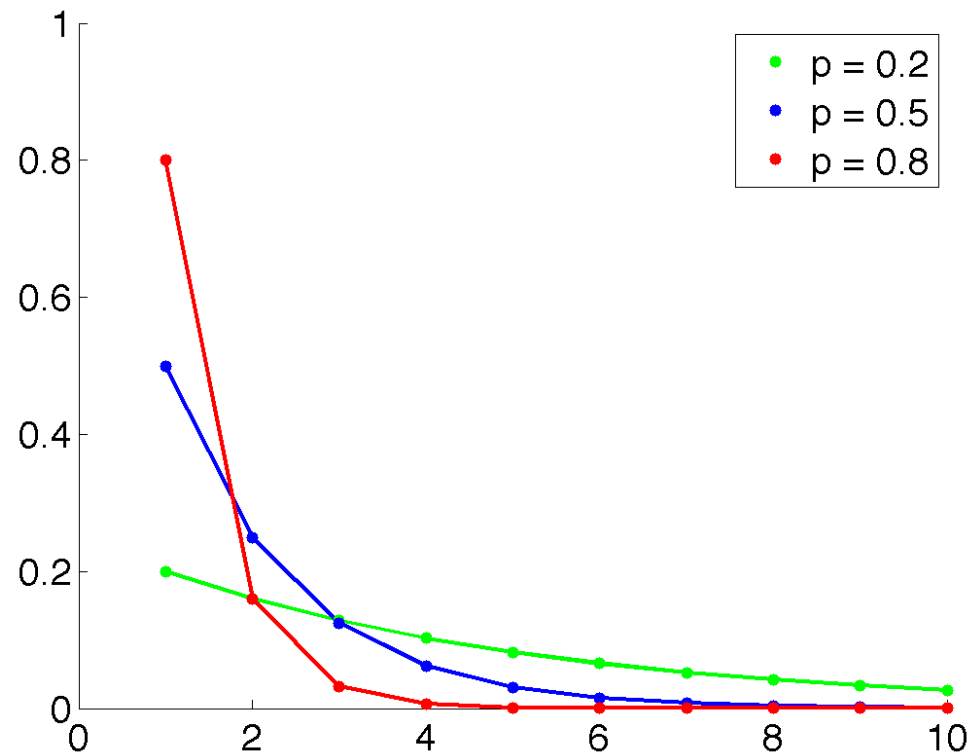
$$E[M] = p \cdot \frac{1}{(1-(1-p))^2} = \frac{p}{p^2} = \frac{1}{p}$$

$$E[M] = \frac{1}{p} \qquad \text{Var}(M) = \frac{1-p}{p^2}$$

M	Probability
1	p
2	(1-p)p
3	(1-p) <sup>2</sup> p
⋮	⋮
k	(1-p) <sup>k-1</sup> p
⋮	⋮

M <sup>2</sup>	Pr
1	p
2 <sup>2</sup>	(1-p)p
3 <sup>2</sup>	(1-p) <sup>2</sup> p
⋮	⋮
k <sup>2</sup>	(1-p) <sup>k-1</sup> p

# The geometric r.v.





# The hypergeometric r.v.

- Suppose we have objects of two types: type 1  $N$  objects (acceptable components) and type 2  $M$  objects (defective components). A sample of  $n$  objects is randomly chosen without replacement.
- A r.v.  $X$  represents a number of type 1 objects in the selection  $\rightarrow$  hypergeometric with parameters  $(N, M, n)$
- Pmf

$$P\{X = i\} = \frac{\binom{N}{i} \cdot \binom{M}{n-i}}{\binom{N+M}{n}}$$

$$E[X] = \frac{nN}{N+M} \quad Var(X) = np(1-p)\left[1 - \frac{n-1}{N+M-1}\right]$$

# Derivation of formulas for mean and variance of hypergeometric r.v.

$$X = I_1 + I_2 + \dots + I_n \quad P\{I_k = 1\} = \frac{N}{N+M} = p$$

$$E[X] = \sum_{k=1}^n E[I_k] = np = \frac{nN}{N+M}$$

$I_k$  are not independent  $\Rightarrow$

$$\text{Var}(X) = \sum_{k=1}^n \text{Var}(I_k) + 2 \sum_{1 \leq k < l \leq n} \text{Cov}(I_k, I_l)$$

$$\text{Var}(I_k) = p(1-p)$$

$$\text{Cov}(I_k, I_l) = E[I_k I_l] - E[I_k]E[I_l]$$

$$\begin{aligned} E[I_k I_l] &= P\{I_k I_l = 1\} = P\{I_k = 1, I_l = 1\} = \\ &= \frac{N}{N+M} \cdot \frac{N-1}{N+M-1} \end{aligned}$$

# Poisson r.v.

- A r.v. taking on one of values 0, 1, 2, ... is said to be a Poisson r.v. with parameter  $\lambda$  ( $\lambda > 0$ ), if pmf is given by

$$P\{X = i\} = e^{-\lambda} \frac{\lambda^i}{i!}$$

- Moment generating function:

$$\phi(t) = E[e^{tX}] = e^{\lambda(e^t - 1)}$$

$$E[X] = \lambda \qquad \text{Var}(X) = \lambda$$

# Examples

- Examples of events that can be modelled by Poisson distribution:
  - Number of typos (misprints) on a page of a book
  - Number of phone calls at a call center per minute
  - Number of times a web server is accessed per minute
  - Number of pine trees per unit area in a mixed forest
  - Number of stars in a given volume of space
  - ...

## When is the Poisson distribution an appropriate model?

- The occurrence of one event does not affect the probability that a second event will occur. That is, events occur independently.
- The rate at which events occur is constant. The rate cannot be higher in some intervals and lower in other intervals.
- The probability of an event in a small interval is proportional to the length of the interval.

## Poisson r.v.

Lets verify that it is truly a pmf:

$$\begin{aligned}\sum_{i=0}^{\infty} P\{X=i\} &= \sum_{i=0}^{\infty} e^{-\lambda} \frac{\lambda^i}{i!} = e^{-\lambda} \sum_{i=0}^{\infty} \frac{\lambda^i}{i!} = \\ &= e^{-\lambda} \cdot e^{\lambda} = 1\end{aligned}$$

Moment generation function:

$$\begin{aligned}\varphi(t) &= E[e^{tX}] = \sum_{i=0}^{\infty} e^{ti} e^{-\lambda} \frac{\lambda^i}{i!} = e^{-\lambda} \sum \frac{(\lambda e^t)^i}{i!} = \\ &= e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)}\end{aligned}$$

$$\varphi'(t) = \lambda e^t e^{\lambda(e^t - 1)} \quad \varphi'(0) = \lambda \cdot 1 \cdot e^{\lambda(1-1)} = \lambda$$

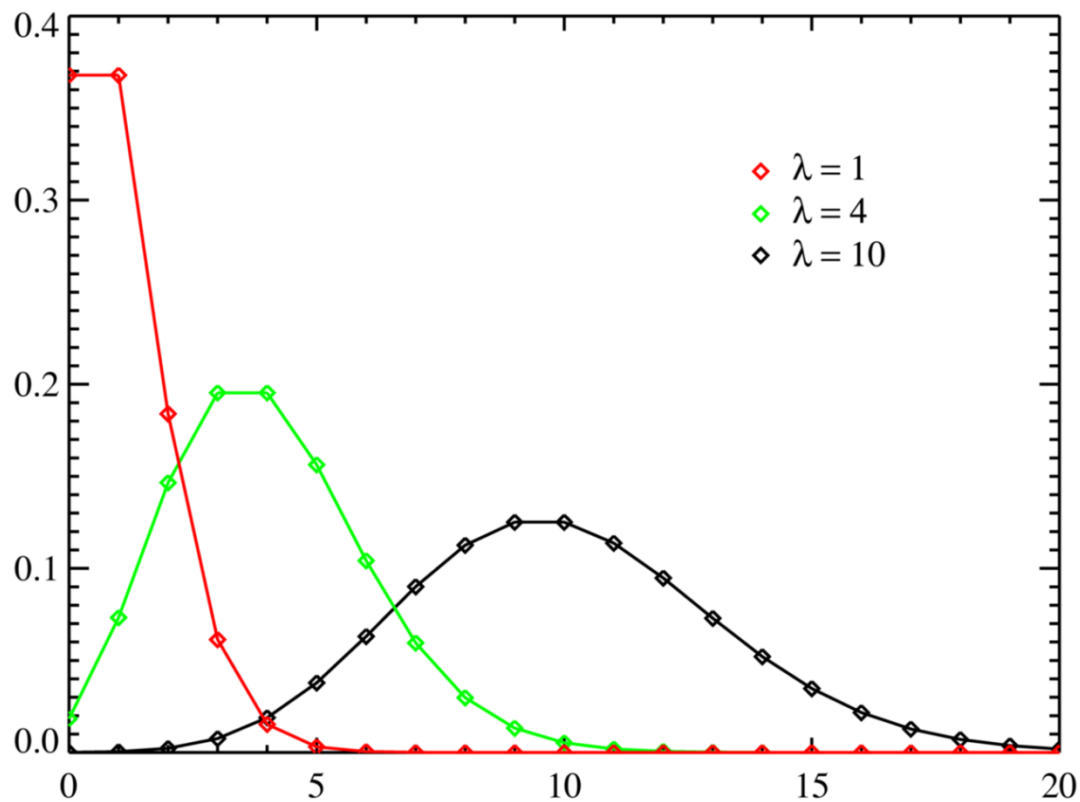
$$\Rightarrow E[X] = \lambda$$

$$\varphi''(t) = (\lambda e^t)^2 e^{\lambda(e^t - 1)} + \lambda e^t e^{\lambda(e^t - 1)}$$

$$\Rightarrow E[X^2] = \varphi''(0) = \lambda^2 + \lambda$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

# Poisson r.v.



# Approximation for a binomial r.v.

- One of the applications of the Poisson probabilities is to approximate the binomial probabilities – when  $n$  is large and  $p$  is small.
- $\lambda = np$



Let  $X \sim b(n, p)$ . We put  $\lambda = np$

$$P\{X=i\} = \frac{n!}{(n-i)!i!} p^i (1-p)^{n-i} = \frac{n!}{(n-i)!i!} \left(\frac{\lambda}{n}\right)^i \left(1 - \frac{\lambda}{n}\right)^{n-i} =$$

$$= \underbrace{\frac{n(n-1)\dots(n-i+1)}{n^i}}_{\xrightarrow{n \rightarrow \infty} 1} \cdot \frac{\lambda^i}{i!} \cdot \frac{\underbrace{\left(1 - \frac{\lambda}{n}\right)^n}_{\xrightarrow{n \rightarrow \infty} e^{-\lambda}}}{\underbrace{\left(1 - \frac{\lambda}{n}\right)^i}_{\xrightarrow{n \rightarrow \infty} 1}} \rightarrow e^{-\lambda} \cdot \frac{\lambda^i}{i!}$$

# Properties of Poisson distribution

- The probability of at least one occurrence of the event in a given time interval is proportional to the length of the interval.
- The probability of two or more occurrences of the event in a very small time interval is negligible.
- The numbers of occurrences of the event in disjoint time intervals are mutually independent.

# Example

- The BER of a communication channel is  $10^{-3}$ . What is a probability that a block of 1000 bits has five or more errors?

$$X \sim b(n, p) \quad n = 1000, p = 10^{-3}$$

Poisson approximation:  $\lambda = np = 1$

$$P\{X \geq 5\} = 1 - P\{X < 5\} = 1 - \sum_{k=0}^4 \frac{\lambda^k}{k!} e^{-\lambda} =$$

$$= 1 - e^{-1} \left( 1 + \frac{1}{1!} + \frac{1}{2!} + \frac{1}{3!} + \frac{1}{4!} \right) \approx 0.0036$$

# Properties of a Poisson r.v.

- Reproductive property
- Proposition.
  - 1) The sum of independent Poisson r.v.s. with parameters  $\lambda_1$  and  $\lambda_2$  is also a Poisson r.v. with parameter  $\lambda_1 + \lambda_2$
  - 2) If a Poisson event having mean  $\lambda$  can be independently classified as being of type 1, 2, ... r with respective probabilities

$$p_1, p_2, \dots, p_r \left( \sum p_i = 1 \right)$$

then type i events are independent Poisson r.v. with means  $\lambda p_1$ ,  $\lambda p_2 \dots \lambda p_r$