total

and

$$E[R] = \frac{1}{\mu - \lambda}$$

 $E[R] = \frac{1}{\mu - \lambda}$  We use the following two properties of an M/M/1 queue:

- 1. From the PASTA property of a Poisson stream, an arriving customer sees the steady-state distribution of the number of customers in system. This arriving customer, with probability  $p_n$ , finds n customers already waiting or in service.
- 2. From the memory less property of the exponential distribution, if the new arrival finds a costumer in service, the remaining service time of that customer is distributed exponentially with mean  $1/\mu$ , i.e., identical to the service requirements of all waiting customers.

The response time of an arriving customer who finds n customers in the system is therefore the sum of (n+1) exponentially distributed random variables, the n already present plus the n arriving customer itself. Such a sum has an (n+1) stage Erlange density function. Then, if  $B_k$  is the service time of customer k, we have

$$Prob\{R > t\} = Prob\left\{\sum_{k=1}^{n+1} B_k > t\right\}.$$

Now, conditioning on the number present when the arrival occurs an using the independence of arrivals and service, we obtain

$$Prob\{R > t\} = \sum_{n=0}^{\infty} \left( Prob\left\{\sum_{k=1}^{n+1} B_k > t\right\} \right) p_n = \sum_{n=0}^{\infty} \left( e^{-\mu t} \sum_{k=0}^{n} \frac{(\mu t)^k}{k!} \right) (1-p) p^n$$

$$= \sum_{k=0}^{n} \sum_{n=0}^{\infty} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) (1-p) p^n = \sum_{k=0}^{n} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) \sum_{n=k}^{\infty} (1-p) p^n$$

$$= \sum_{k=0}^{n} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) p^k = e^{-\mu(1-p)t}, t \ge 0.$$

Hence the probability distribution function for the response time in an M/M/1 queue is  $Prob\{R \le t\} = W_r(t) = 1 - e^{(\mu - \lambda)t},$ 

i.e., the exponential distribution with mean

$$E[R] = \frac{1}{\mu - \lambda} = \frac{1}{\mu(1 - p)}.$$

Notice that is not possible to write this performance measure just in terms of p. It depends on  $\lambda$ and  $\mu$  but not just trough their ratio p. This means that it is possible to assign values to  $\lambda$  and  $\mu$  in such s way that the system can be almost saturated with large queues but still have a very short expected response time.

The probability density function for the response time can be immediately found as the density function of the exponential distribution with parameter  $\mu(1-p)$  or evaluated directly as

$$w_r(t) = \sum_{n=0}^{\infty} P_n g_{n+1}(t) = (1-p)\mu e^{-\mu t} \sum_{n=0}^{\infty} \frac{(p\mu t)^n}{n!} = (\mu - \lambda)e^{-(\mu - \lambda)t},$$

410 Elementary Queueing Theory  $E[R] = \frac{1}{\mu - \lambda}$ We use the following two properties of an M/M/1 queue: From the PASTA property of a Poisson stream, an arriving customer sees the steady-state distribution of the number of customers in the system. This arriving customer, with probability p<sub>s</sub>, finds n customers already waiting or in service.
 From the memoryless property of the exponential distribution, if the new arrival finds a customer in service, the remaining service time of that customer is distributed exponentially with mean 1/μ, i.e., identical to the service requirements of all waiting customers. The response time of an arriving customer who finds n customers in the system is therefore the sum of (n+1) exponentially distributed random variables, the n already present plus the arriving customer itself. Such a sum has an (n+1) stage Erlang density function. Then, if  $B_k$  is the service time of customer k, we have  $\operatorname{Prob}\{R > t\} = \operatorname{Prob}\left\{\sum_{k=1}^{n+1} B_k > t\right\}$ Now, conditioning on the number present when the arrival occurs and using the independence of arrivals and service, we obtain  $Prob\{R \ge t\} = \sum_{n=0}^{\infty} \left(Prob\left\{\sum_{k=1}^{n+1} B_k > t\right\}\right) p_n = \sum_{n=0}^{\infty} \left(e^{-\mu t} \sum_{k=1}^{n} \frac{(\mu t)^k}{k!}\right) (1-\rho) \rho^n$  $= \sum_{k=0}^{n} \sum_{n=k}^{\infty} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) (1-\rho) \rho^n = \sum_{k=0}^{n} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) \sum_{i=1}^{\infty} (1-\rho) \rho^n$  $= \sum_{k=0}^{n} \left( e^{-\mu t} \frac{(\mu t)^k}{k!} \right) \rho^k = e^{-\mu(1-\rho)t}, \quad t \ge 0.$ Hence the probability distribution function for the response time in an M/M/1 queue is  $Prob\{R \le t\} = W_r(t) = 1 - e^{-(\mu - \lambda)t}$ , i.e., the exponential distribution with mean Notice that it is not possible to write this performance measure just in terms of  $\rho$ . It depends on  $\lambda$ and  $\mu$  but not just through their ratio  $\rho$ . This means that it is possible to assign values to  $\lambda$  and  $\mu$ in such a way that the system can be almost saturated with large queues but still have a very short expected response time. The probability density function for the response time can be immediately found as the density function of the exponential distribution with parameter  $\mu(1-\rho)$  or evaluated directly as  $w_r(t) = \sum_{n=0}^{\infty} p_n g_{n+1}(t) = (1-\rho)\mu e^{-\mu t} \sum_{n=0}^{\infty} \frac{(\rho \mu t)^n}{n!} = (\mu - \lambda)e^{-(\mu - \lambda)t},$ 

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