Generating discrete Random variable

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- Consider a r.v.X with pmf $P(X = x_j) = p_j$, $j = 0, 1, \dots, \sum_j p_j = 1$
- To generate a value from this distribution, first we generate a random number U and set

$$X = \begin{cases} x_0, & \text{if } U < p_0 \\ x_1, & \text{if } p_0 \le U < p_0 + p_1 \\ \vdots & & \\ x_j & \text{if } \sum_{i=0}^{j-1} p_i \le U < \sum_{i=0}^{j} p_i \\ \vdots & & \end{cases}$$

- This is called the inverse transform method.
- Proof to be outlined in class.

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$$P(X = x_j) = P(\sum_{i=0}^{j-1} p_i \le U < \sum_{i=0}^{j} p_i)$$

$$= P(U < \sum_{i=0}^{j} p_i) - P(U < \sum_{i=0}^{j-1} p_i)$$

$$= \sum_{i=0}^{j} p_i - \sum_{i=0}^{j-1} p_i$$

$$= p_j$$

Some remarks

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 The method can be written algorithmically as generate a random number U

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if U < p_0, set X = x_0 and STOP
if U < p_0 + p_1, set X = x_1 and STOP
if U < p_0 + p_1 + p_2, set X = x_2 and STOP
\vdots
```

■ If the $X_i's$ are ordered like $X_0 < X_1 < X_2 < \cdots$ so that the cdf $F(X_k) = \sum_{i=0}^k p_i$ and that X equals to X_j if $F(X_{j-1}) \le U < F(X_j)$. Therefore, after generating U, we determine the value of X by looking for the interval. $[F(x_{j-1}), F(X_j)]$ in which it lies (or equivalently finding the inverse of U).

Illustrative example

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Suppose we want to simulate from the discrete distribution with
P(X = 1) = 0.20, P(X = 2) = 0.25, P(X = 3) = 0.40,
P(X = 4) = 0.15
We do the following:
generate a random number U
if U < 0.20, set X = 1 and STOP
if U < 0.45, set X = 2 and STOP
if U < 0.85, set X = 3 and STOP
otherwise set X = 4.
It is suggested the following could be more efficient:
generate a random number U
if U < 0.40, set X = 3 and STOP
if U < 0.65, set X = 2 and STOP
if U < 0.85, set X = 1 and STOP
otherwise set X = 4.
```

- In the discrete uniform distribution, we have equal probabilities: $P(X = j) = \frac{1}{n}$, for j = 1, 2, ..., n.
- To simulate from this distribution, we generate a random number U and then set

$$X = j$$
 if $\frac{j-1}{n} \le U < \frac{j}{n}$.

■ This condition is equivalent to if $j - 1 \le nU < j$, that is X=Int(nU)+1, where Int(X) is the greatest integer part of X.

Generating a geometric random variable

- In a geometric distribution with parameter p, we have $P(X = i) = pq^{i-1}$, for $i \ge 1$.
- Note that the cumulative probability $P(X \le j-1) = \sum_{i=1}^{j-1} P(X=i) = 1 q^{j-1}$.
- It can be shown that with a random number U, then $X=\operatorname{Int}(\frac{log(U)}{log(q)})+1$. is indeed geometric with parameter p.

Generate
$$U$$
 and set $X=j$
if $1-q^{j-1} \le U < 1-q^{j}$
 $U \sim \text{Uniform}(0,1) \Longrightarrow 1-U \sim \text{Uniform}(0,1)$
 $q^{j} < 1-U \le q^{j-1}$

$$\begin{split} X &= & \min\{j| \quad q^j < 1 - U\} \\ &= & \min\{j| \quad j * logq < log(1 - U)\} \\ &= & \min\{j| \quad j > \frac{log(1 - U)}{logq}\} \\ &= & Int(\frac{log(1 - U)}{logq}) + 1 \\ &= & Int(\frac{log(U)}{logq}) + 1 \end{split}$$

- For the case of the Poisson, we exploit the recursion property $p_{i+1} = \frac{\lambda}{i+1} p_i$ for $i \ge 0$.
- The following steps can then be followed to generate from a Poisson with parameter λ :

```
step 1: generate a random number U.

step 2: set i=0, p=e^{-\lambda} and F=p.

step 3: if U < F, set X=i and STOP.

step 4: set p=\frac{\lambda p}{i+1}, F=F+p, and i=i+1.

step 5: return to step 3.
```

- Note that *F* is indeed the cdf $F(i) = P(X \le i)$.
- It can be shown that the average number of searches grows with the square root of λ .(proof to be discussed!)

■ Just as in the Poisson case, we exploit the recursion property for the Binomial distribution:

$$P(X = i + 1) = \frac{n-i}{i+1} \frac{p}{1-p} P(X = i).$$

■ The following steps can then be followed to generate a Binomial random variable with parameters *n* and probability of success *p*:

step 1: generate a random number U.

step 2: set
$$c = \frac{p}{1-p}$$
, $i = 0$, $pr = (1-p)^n$, and $F = pr$.

step 3: if U < F, set X = i and STOP.

step 4: reset
$$pr = \left[\frac{c(n-i)}{i+1}\right]pr$$
, $F = F + pr$, and $i = i+1$.

step 5: return to step 3.

- As an exercise, try to write an *R* routine for generating a Binomial random variable following the above steps.
- Another approach to simulate from a Binomial(n, p) is to use the interpretation that it is equal to the number of success in n independent Bernoulli trials.

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$$Bernoulli(p) = \begin{cases} X = 1 & with probability p \\ X = 0 & with probability 1 - p \end{cases}$$

Let $X_1, X_2, X_3, \dots, X_n$ are all Bernoulli trials, Consider $X_1 + X_2 + X_3 + \dots + X_n = X$, So $X \sim$ Binomial (n, p).

- Consider now simulating from a distribution with mass function $P(X = j) = \alpha p_i^{(1)} + (1 \alpha) p_i^{(2)}, j \ge 0, 0 < \alpha < 1.$
- If X_1 and X_2 are the random variables with respective mass functions $p_j^{(1)}$ and $p_j^{(2)}$, then

$$X = \left\{ \begin{array}{ll} X_1 & \textit{with probability} \quad \alpha \\ X_2 & \textit{with probability} \quad 1-\alpha \end{array} \right.$$

- One approach then to generate from this mixture distribution is:
 - step 1: genrate a random number U_1
 - step 2: generate from X_1 and X_2 distributions.
 - step 3: if $U < \alpha$, set $X = X_1$.
 - step 4: else if $U > \alpha$, set $X = X_2$.

Example 4g

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■ Consider the example of generating *X* from a distribution with mass function

$$p_j = P(X = j) = \begin{cases} 0.05, & for \ j = 1, 2, 3, 4, 5 \\ 0.15, & for \ j = 6, 7, 8, 9, 10 \end{cases}$$

■ Note that this is equivalent to

$$P(X = j) = 0.5p_j^{(1)} + 0.5p_j^{(2)},$$

where $p_j^{(1)} = 0.10$, for $j = 1, 2, \dots, 10$ and

$$p_j^{(2)} = \begin{cases} 0, & for \quad j = 1, 2, 3, 4, 5 \\ 0.2, & for \quad j = 6, 7, 8, 9, 10 \end{cases}$$

Thus, first generate a random number U, and then generate from the discrete uniform over 1, 2, \cdots 10 if U < 0.5 and from the discrete uniform over 6,7,8,9,10 otherwise.

■ Consider the example of generating *X* from a distribution with mass function

$$p_j = P(X = j) = \begin{cases} 0.05, & for \ j = 1, 2, 3, 4, 5 \\ 0.15, & for \ j = 6, 7, 8, 9, 10 \end{cases}$$

Note that this is equivalent to $P(X = j) = 0.5p_j^{(1)} + 0.5p_j^{(2)}$, where $p_j^{(1)} = 0.10$, for $j = 1, 2, \dots, 10$ and

$$p_{j}^{(2)} = \left\{ \begin{array}{ll} 0, & \textit{for} \quad j = 1, 2, 3, 4, 5 \\ 0.2, & \textit{for} \quad j = 6, 7, 8, 9, 10 \end{array} \right.$$

■ Thus, first generate a random number U, and then generate from the discrete uniform over $1, 2, \dots 10$ if U < 0.5 and from the discrete uniform over 6,7,8,9,10 otherwise.

- In the case where the distribution function of X is given by $F(x) = \sum_{i=1}^{n} \alpha_i F_i(x)$, where F_i , $i = 1, \dots, n$ are distribution functions, we have what we call a mixture distribution.
- To simulate from such a mixture distribution, step 1: simulate a random variable I, equal to i with probability α_i , for $i = 1, 2, \dots, 10$.
 - step 2: simulate from the distribution F_i .
- This is also called the composition method. Generate U $U < \alpha_1 \longrightarrow X = X_1$ $U < \alpha_1 + \alpha_2 \longrightarrow X = X_2$ \vdots $U < \sum_{i=1}^j \alpha_i \longrightarrow X = X_j$ \vdots

R codes for simulating from known discrete distributions

- In *R*, there are many functions that generate discrete random variables. Most of them start with *r*.
- Here are a few of them:
 rbinom binomial
 rnbinom negative binomial
 rpois Poisson
 rgeom geometric
 rhyper hypergeometric