

# An Early Days Rough Draft Conversation About MAIHDA

Multilevel Analysis Of Individual Heterogeneity And Discriminatory Accuracy

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## 1 Inspirations

“My conception of the universal is that of a universal enriched by all that is particular, a universal enriched by every particular: the deepening and coexistence of all particulars.” (Cesaire, 1956)

Conceptual inspiration comes from sources such as Cho et al. (2013), Cesaire (1956), UNESCO (1997), Martín-Baró (1998), Montero & Sonn (2009), Antweiler (2016), Stage & Wells (2014), and Scharrer & Ramasubramanian (2021).

Substantively and practically, this handout depends **heavily** on the ideas, example code, and example data set in Evans et al. (2024).

## 2 The Basic Idea of Multilevel Modeling

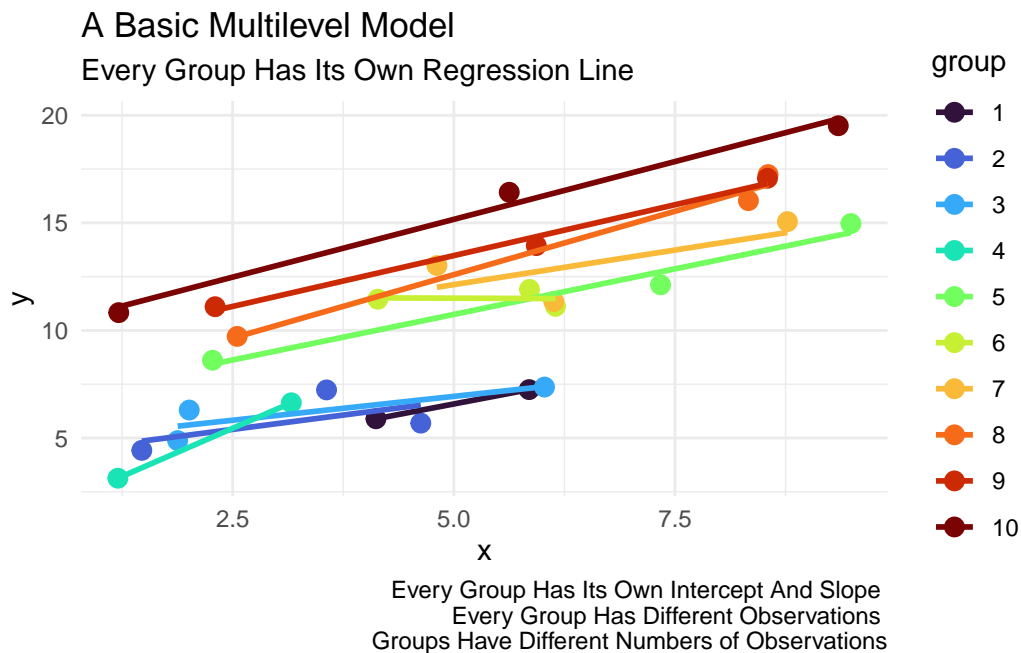


Figure 1: A Basic Multilevel Model

### 3 The Basic Idea Of MAIHDA

#### 💡 Key Ideas

The key ideas are as follows:

1. A *traditional* approach would ask us to generate multiple interaction terms to explore interactions e.g.  $sex \times race \times education \times age$  as well as all of the lower level 2 and 3 -way interactions. Upon first glance this seems unwieldy, though in a future discussion, we could talk about how some software (e.g. Stata) makes this approach less unwieldy than it might initially appear.
2. In contrast, **MAIHDA** uses combinations of variables to generate *strata* which are then treated as the levels in a multilevel model.

### 4 Get The Data

#### 4.1 Stata

```
use "TutorialData.dta" // get the data  
  
describe // describe the data
```

Contains data from TutorialData.dta

Observations: 33,000  
Variables: 7 12 Oct 2023 16:51

Variable name	Storage type	Display format	Value label	Variable label
HbA1c	float	%9.1f		HbA1c (mmol/mol)
diabetic	byte	%9.0g		
sex	byte	%12.0g	sexlabel	
race	byte	%12.0g	racelabel	
education	byte	%21.0g	educationlabel	
income	byte	%12.0g	incomelabel	
age	byte	%12.0g	agelabel	

---

Sorted by:

## 4.2 R

```
library(haven) # read Stata  
TutorialData <- read_dta("TutorialData.dta")
```

## 5 Generating The Strata

### 5.1 Stata

```
generate stratum = 10000*sex + 1000*race + 100*education + 10*income + 1*age // create stratum
```

### 5.2 R

```
library(dplyr) # data wrangling  
TutorialData <- TutorialData %>%  
  mutate(stratum = 10000*sex +  
           1000*race +  
           100*education +  
           10*income +  
           1*age) # create stratum variable
```

## 6 Empty or ANOVA Model

### 6.1 Model

#### 6.1.1 Stata

```
mixed HbA1c || stratum: // ANOVA model
```

Performing EM optimization ...

Performing gradient-based optimization:

Iteration 0: Log likelihood = -121483.73

Iteration 1: Log likelihood = -121483.73

Computing standard errors ...

Mixed-effects ML regression

Group variable: stratum

Number of obs = 33,000

Number of groups = 384

Obs per group:

min = 1

avg = 85.9

max = 683

Wald chi2(0) = .

Prob > chi2 = .

Log likelihood = -121483.73

HbA1c	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
+						
_cons	40.78932	.1760906	231.64	0.000	40.44419	41.13445

Random-effects parameters		Estimate	Std. err.	[95% conf. interval]	
+					
stratum: Identity					
	var(_cons)	9.33452	.8429295	7.820354	11.14186
+					
	var(Residual)	90.26358	.7066668	88.88911	91.65931

LR test vs. linear model: chibar2(01) = 2228.00 Prob >= chibar2 = 0.0000

## 6.1.2 R

```
library(lme4)
```

Loading required package: Matrix

```
fit1 <- lmer(HbA1c ~ (1 | stratum), data = TutorialData)

summary(fit1)
```

```
Linear mixed model fit by REML ['lmerMod']
Formula: HbA1c ~ (1 | stratum)
Data: TutorialData
```

```
REML criterion at convergence: 242969.1
```

```
Scaled residuals:
```

```
      Min       1Q   Median       3Q      Max
-3.7126 -0.6181 -0.0564  0.4822  6.5172
```

```
Random effects:
```

```
Groups   Name             Variance Std.Dev.
stratum  (Intercept)    9.366     3.060
Residual                    90.264    9.501
Number of obs: 33000, groups:  stratum, 384
```

```
Fixed effects:
```

```
              Estimate Std. Error t value
(Intercept)  40.7896     0.1763   231.3
```

## 6.2 IntraClass Correlation Coefficient

### The ICC

The ICC answers the question: “How much clustering is there?”

### 6.2.1 Stata

```
estat icc // intraclass correlation coefficient
```

```
Intraclass correlation
```

```
-----
Level |          ICC   Std. err.   [95% conf. interval]
```



Log likelihood = -121213.02

max = 683  
Wald chi2(12) = 1344.70  
Prob > chi2 = 0.0000

HbA1c	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
sex						
Female	-.516474	.1575437	-3.28	0.001	-.8252541	-.207694
race						
Black	4.44937	.1908202	23.32	0.000	4.07537	4.823371
Hispanic	1.038471	.1939947	5.35	0.000	.6582485	1.418694
education						
High school	-.5410972	.2378408	-2.28	0.023	-1.007257	-.0749378
Some coll..	-.5476019	.2469721	-2.22	0.027	-1.031658	-.0635453
College p..	-1.128424	.2441963	-4.62	0.000	-1.60704	-.649808
income						
Low-mid	.1807317	.2114433	0.85	0.393	-.2336895	.5951529
Mid-high	-.3039793	.2188399	-1.39	0.165	-.7328975	.124939
High	-1.327629	.252043	-5.27	0.000	-1.821624	-.8336339
age						
30-44	1.10363	.2312428	4.77	0.000	.6504026	1.556858
45-59	1.808993	.2356534	7.68	0.000	1.347121	2.270865
60+	5.504486	.2405396	22.88	0.000	5.033037	5.975935
_cons	38.03392	.2980168	127.62	0.000	37.44982	38.61802

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	
stratum: Identity				
var(_cons)	.8002982	.1494763	.5549681	1.154079
var(Residual)	90.2666	.7059543	88.8935	91.6609

LR test vs. linear model: chibar2(01) = 88.75      Prob >= chibar2 = 0.0000



## 7.2 R

```
library(lme4)

TutorialData$sex <- factor(TutorialData$sex)
TutorialData$race <- factor(TutorialData$race)
TutorialData$education <- factor(TutorialData$education)
TutorialData$income <- factor(TutorialData$income)
TutorialData$age <- factor(TutorialData$age)

fit1 <- lmer(HbA1c ~ sex + race + education + income + age + (1 | stratum), data = TutorialData)

summary(fit1)
```

Linear mixed model fit by REML ['lmerMod']

Formula: HbA1c ~ sex + race + education + income + age + (1 | stratum)

Data: TutorialData

REML criterion at convergence: 242446.4

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.7402	-0.6212	-0.0553	0.4837	6.5708

Random effects:

Groups	Name	Variance	Std.Dev.
stratum	(Intercept)	0.8844	0.9404
Residual		90.2660	9.5008

Number of obs: 33000, groups: stratum, 384

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	38.0335	0.3040	125.110
sex2	-0.5154	0.1612	-3.197
race2	4.4504	0.1951	22.816
race3	1.0364	0.1981	5.232
education2	-0.5480	0.2430	-2.256
education3	-0.5448	0.2519	-2.163
education4	-1.1337	0.2492	-4.550
income2	0.1790	0.2163	0.828
income3	-0.3006	0.2237	-1.344
income4	-1.3231	0.2573	-5.143

age2	1.1073	0.2362	4.688
age3	1.8094	0.2405	7.524
age4	5.5112	0.2456	22.442

Correlation matrix not shown by default, as  $p = 13 > 12$ .

Use `print(x, correlation=TRUE)` or  
`vcov(x)` if you need it

## 8 Discussion Questions

1. Clarifying what we mean by intersectionality before moving to quantitative approaches
2. What MAIHDA adds beyond conventional interaction models. (Do interactions fail to capture these ideas? Are multiple interactions so hard to test?)
3. The role of multilevel modeling and partial pooling
4. MAIHDA in descriptive vs. causal contexts
5. Broader model evaluation and potential extensions of the method (CF testing multiple interactions and maximal models (<https://agrogan1.github.io/multilevel/Bayesian-and-frequentist-MLM/Bayesian-and-frequentist-MLM.html#maximal-models>))

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