Multilevel Multilingual

Multilevel Models in Stata, R and Julia

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# 1. Multilevel Multilingual

“This curious world which we inhabit is more wonderful than it is convenient…” (Thoreau, 1975)

“Mathematics is my secret. My secret weakness. I feel like a stubborn, helpless fool in the middle of a problem. Trapped and crazed. Also, thrilled.” (Schanen, 2021)

## 1.1 Introduction

Below, I describe the use of [Stata](https://www.stata.com/) (StataCorp, 2023), [R](https://www.r-project.org/) (Bates et al., 2015; R Core Team, 2023), and [Julia](https://www.julialang.org/) (Bates, 2024; Bezanson et al., 2017) to estimate multilevel models.

All of these software packages can estimate multilevel models and can visualize relationships in the data. However, there are substantial differences between the different packages: Stata is proprietary *for cost* software, which is very well documented and very intuitive. While it costs money to purchase Stata, the price is often very reasonable for academic and educational use. R is *free* open source software which is less intuitive, but there are many excellent resources for learning R. There is often a cost associated with purchasing books and other materials for learning R, which sometimes feels like it offsets the fact that R is free. Julia is newer open source software, also *free*, and ostensibly much faster than either Stata or R, which may be an important advantage when running multilevel models with very large data sets. At this point in time, both Stata and R feel much more *stable* than Julia which is still evolving software.

While any of these software packages can be used for learning and estimating multilevel models, I will offer my own opinion–based upon 15 years of teaching multilevel models at the doctoral level–that Stata offers the quickest pathway for learning the basic and advanced uses of multilevel models. I also believe the intuitive nature of Stata syntax contributes to accurate and replicable work in this area.

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| Table 1.1: Software for Multilevel Modeling   | Software | Cost | Ease of Use | | --- | --- | --- | | Stata | some cost | learning curve, but very intuitive for both multilevel modeling and graphing. | | R | free | learning curve: intuitive for multilevel modeling; but steeper learning curve for graphing (ggplot). | | Julia | free | steep learning curve in general: steep learning curve for multilevel modeling; and very steep learning curve for graphing. Graphics libraries are very much under development and in flux. | |

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| Results Will Vary Somewhat |
| Estimating multilevel models is a complex endeavor. The software details of how this is accomplished are beyond the purview of this book. Suffice it to say that across different software packages there will be differences in estimation routines, resulting in some numerical differences in the results provided by different software packages. Substantively speaking, however, results should agree across software. |

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| Multi-Line Commands |
| Sometimes I have written commands out over multiple lines. I have done this for especially long commands, but have also sometimes done this simply for the sake of clarity. The different software packages have different approaches to multi-line commands.   1. By default, *Stata* ends a command at the end of a line. If you are going to write a multi-line command you should use the /// line continuation characters. 2. *R* is the software that most naturally can be written using multiple lines, as R commands are usually clearly encased in parentheses (()) or continued with + signs. 3. Like *Stata*, *Julia* expects commands to end at the end of a line. If you are going to write a mult-line command, all commands except for the last line should end in a character that clearly indicates continuation, like a + sign. An alternative is to encase the entire Julia command in an outer set of parentheses (()). |

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| Running Statistical Packages in Quarto |
| I used Quarto (Allaire et al., 2024) (<https://quarto.org/>) to create this Appendix. Quarto is a programming and publishing environment that can run multiple programming languages, including Stata, R and Julia, and that can write to multiple output formats including HTML, PDF, and MS Word. To run Stata, I used the Statamarkdown library (Hemken, 2023) in R to connect Stata to Quarto. Quarto has a built in connection to R, and runs R without issue. To run Julia, I used the JuliaCall library (Li, 2019) in R to connect Quarto to Julia.  Of course, each of these programs can be run by itself, if you have them installed on your computer. |

## 1.2 The Data

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| Datasets |
| The examples use the simulated\_multilevel\_data.dta and simulated\_multilevel\_longitudinal\_data.dta files.  Here is a [direct link](https://github.com/agrogan1/multilevel-multilingual/raw/main/simulated_multilevel_data.dta) to download the cross-sectional data.  Here is a [direct link](https://github.com/agrogan1/multilevel-multilingual/raw/main/simulated_multilevel_longitudinal_data.dta) to download the longitudinal data. |

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| Table 1.2: Sample of Simulated Multilevel Data  Table continues below   | country | HDI | family | id | identity | intervention | physical\_punishment | | --- | --- | --- | --- | --- | --- | --- | | 1 | 69 | 1 | 1.1 | 1 | 0 | 3 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 2 | | 1 | 69 | 3 | 1.3 | 0 | 1 | 3 | | 1 | 69 | 4 | 1.4 | 1 | 0 | 0 | | 1 | 69 | 5 | 1.5 | 1 | 0 | 4 | | 1 | 69 | 6 | 1.6 | 0 | 1 | 5 |      | warmth | outcome | | --- | --- | | 3 | 57.47 | | 1 | 50.1 | | 2 | 52.92 | | 5 | 60.17 | | 4 | 55.05 | | 3 | 49.81 | |

## 1.3 An Introduction To Equations and Syntax

To explain statistical syntax for each software, I consider the general case of a multilevel model with dependent variable y, independent variables x and z, clustering variable group, and a random slope for x. *i* is the index for the person, while *j* is the index for the group.

### Stata

In Stata mixed, the syntax for a multilevel model of the form described in [Equation 1.1](#eq-MLMsimple) is:

mixed y x z || group: x

### R

In R lme4, the syntax for a multilevel model of the form described in [Equation 1.1](#eq-MLMsimple) is:

library(lme4)  
  
lmer(y ~ x + z + (1 + x || group), data = ...)

### Julia

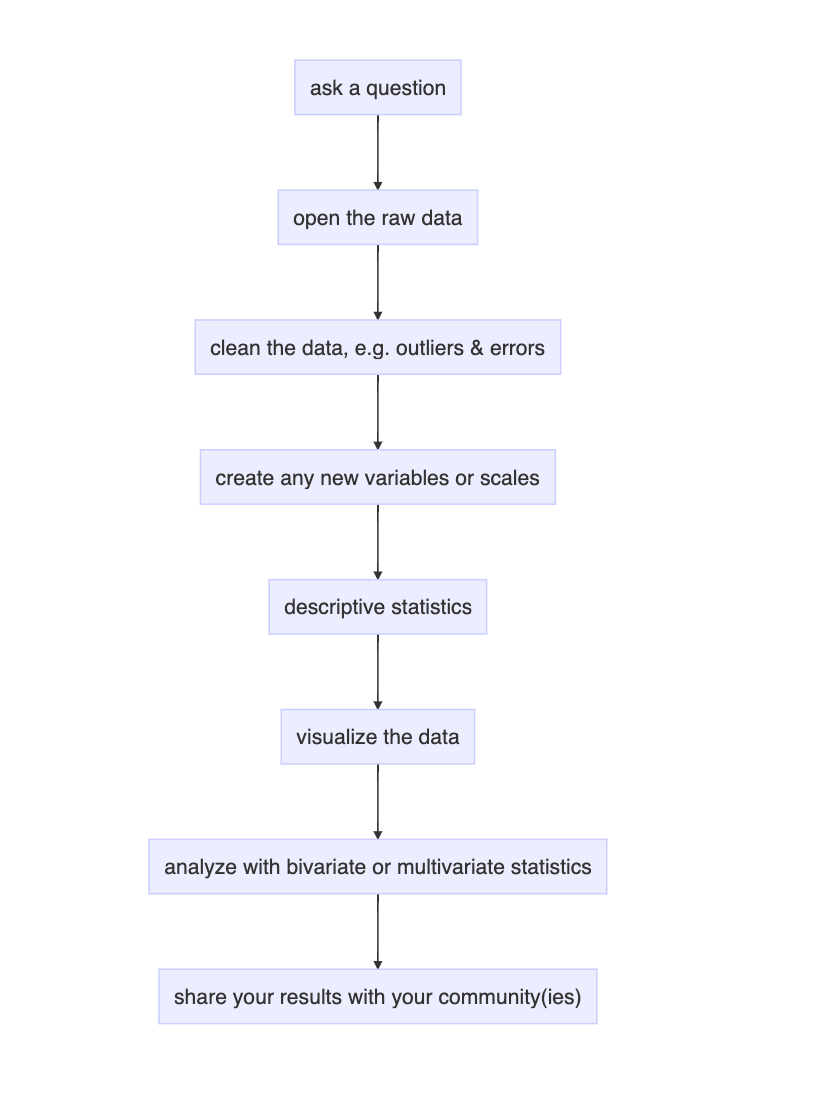
In Julia MixedModels, the syntax for a multilevel model of the form described in [Equation 1.1](#eq-MLMsimple) is:

using MixedModels  
  
fit(MixedModel, @formula(y ~ x + z + (1 + x | group)), data)

# 2. Statistical Workflows

## 2.1 Statistical Software Is Best Run Using a Script

Many statistical workflows–whatever the statistical package being used–follow the same conceptual pattern.



A Common Statistical Workflow

Increasingly, we want to think about workflows that are

* **documentable**, **transparent**, and **auditable**: We have a record of what we did if we want to double check our work, clarify a result, or develop a new project with a similar process. We, or others, can find the inevitable errors in our work, **and correct them**.
* **replicable**: Others can replicate our findings with the same or new data.
* **scalable**: We are developing a process that can be as easily used with *thousands* or *millions* of rows of data as it can with *ten* rows of data. We are developing a process that can be easily repeated if we are *constantly getting new or updated data*, e.g. getting new data every week, or every month.

## 2.2 Scripts

For most statistical workflows, we will often want to write a script or code. Data analysis scripts can be stored in a Quarto document (Allaire et al., 2024) as they are in this Appendix. Alternatively, every statistical package has its own unique format for storing scripts as a text file: in Stata, scripts are stored in .do files; in R, scripts are stored in .R files, and in Julia, scripts are stored in .jl files.

## 2.3 Script Flow

A good practice when writing a script, is to have a script that begins with the raw data, moves through any necessary re-coding or cleaning of the data, generates descriptive statistics, generates the appropriate multivariate results, and then generates any necessary visualizations.

## 2.4 Good Statistical Workflows Allow Multiple Statistical Packages

While this Appendix focuses on the use of each individual statistical package on its own, it is certainly possible to use multiple statistical packages as part of the same workflow. For example, one might employ Stata to carry out data management tasks, and then possibly use R to run a multilevel model with a more complicated multilevel structure, such as a cross-classified model, or Julia to more quickly run a model with a large data.

## 2.5 Good Statistical Workflows Require Safe Workspaces

It is also *very important* to be aware that good complex workflows are *highly iterative* and *highly collaborative*. Good complex workflows require a *safe workspace* in which team members feel free to admit their own errors, and help with others’ mistakes in a non-judgmental fashion. Such a *safe environment* is necessary to build an environment where the *overall error rate* is low.

## 2.6 Good Statistical Workflows Require Patience And Can Be Psychologically Demanding

Developing a good documented and auditable workflow that is implemented in code requires a lot of patience, and often, **many iterations**. Working through these many iterations can be psychologically demanding. It is important to remember that careful attention to getting the details right early in the research process, while sometimes tiring and frustrating, will pay large dividends later on when the research is reviewed, presented, published and read.

## 2.7 Good Statistical Workflows Often Allow Multiple Principled Ways Forward

One of my most recent ideas about statistical workflows is that there are certainly *wrong* decisions that one can make with data.

For example, I would not want to write the paper that says that smoking prevents lung cancer, nor would I want to write a paper saying physical punishment is good for children.

That being said, I think there are often *multiple principled ways forward*.

Often the key is not so much to make the 100% correct decision, but to make one of *several possible principled decisions*.

Then after making a *principled decision*, one is *transparent* and *thorough* about describing the decision that one made.

For example, in implementing a multilevel analysis, I would have many choices: I could estimate only a random intercept; estimate one or more random slopes; or estimate all possible random slopes. The random effects could be correlated or uncorrelated. I could estimate only main effects, or could estimate interactions of several variables. Each of these would be a different, yet principled, approach to analyzing the data.

In science and statistics, we often want an answer that provides one clear direction. Instead, I’m increasingly convinced that the best science (and teaching!) often involves engaging in open discussion about the multiple possible alternatives, and then choosing one principled solution, and being transparent about its implementation.

# 3. Storing Statistical Data

## 3.1 Spreadsheets

Spreadsheets are sometimes used to collect and store data. Spreadsheets may sometimes be used because they are the only program that some individuals or agencies have for storing data. Spreadsheet programs may also be used because spreadsheets can be very intuitive and easy ways of managing small amounts of data.

However, spreadsheets may be problematic as an ultimate data storage solution for a number of reasons detailed below, especially as data sets grow in size. Notably, statistical programs like [Stata](https://www.stata.com/), [R](https://www.r-project.org/), or [Julia](https://julialang.org/) can all store additional information with each variable such as: a *variable label*, describing the contents of the variable, or the survey question that resulted in the variable; and a *value label*, which attaches qualitative information to each possible value of the response.

Spreadsheets do not generally contain this extra information about each variable, or column of data, which may lead to errors in working with quantitative information.

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| If Data Are Stored In A Spreadsheet, Variable Names Should Be Limited To a Single Row of the Spreadsheet |
| If spreadsheets are used to store data, the first row of the data should be used to list the variable names, as is seen in the example below. Rows other than the first row should not contain additional information about the variables, but should only contain data.   |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | | Table 3.1: Example Data As Stored in A Spreadsheet  Table continues below   | country | HDI | family | id | identity | intervention | physical\_punishment | | --- | --- | --- | --- | --- | --- | --- | | 1 | 69 | 1 | 1.1 | 1 | 0 | 3 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 2 | | 1 | 69 | 3 | 1.3 | 0 | 1 | 3 | | 1 | 69 | 4 | 1.4 | 1 | 0 | 0 | | 1 | 69 | 5 | 1.5 | 1 | 0 | 4 | | 1 | 69 | 6 | 1.6 | 0 | 1 | 5 |      | warmth | outcome | | --- | --- | | 3 | 57.47 | | 1 | 50.1 | | 2 | 52.92 | | 5 | 60.17 | | 4 | 55.05 | | 3 | 49.81 | | |

## 3.2 Data in Statistical Format

I load the data from a statistical program.

### 3.2.1 Describe The Data

Notice how a description of the data contains information that helps us to understand the variables.

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| Table 3.2: Variable Labels   | pos | variable | label | | --- | --- | --- | | 1 | country | country id | | 2 | HDI | Human Development Index | | 3 | family | family id | | 4 | id | unique country family id | | 5 | identity | hypothetical identity group variable | | 6 | intervention | recieved intervention | | 7 | physical\_punishment | physical punishment in past week | | 8 | warmth | parental warmth in past week | | 9 | outcome | beneficial outcome | |

### 3.2.2 Descriptive Statistics

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| Variable Labels and Value Labels Help Us Understand Our Data |
| Notice how the descriptive statistics and graph are informative in that they contain information on the *variable label* and *value label*. These help us to get an intuitive sense of the information in the data. We see this information when we list out the data as well. |

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| Table 3.3: Descriptive Statistics  Table continues below   | country | HDI | family | id | | --- | --- | --- | --- | | 1 : 100 | Min. :33.00 | Min. : 1.00 | Length:3000 | | 2 : 100 | 1st Qu.:53.00 | 1st Qu.: 25.75 | Class :character | | 3 : 100 | Median :70.00 | Median : 50.50 | Mode :character | | 4 : 100 | Mean :64.77 | Mean : 50.50 | NA | | 5 : 100 | 3rd Qu.:81.00 | 3rd Qu.: 75.25 | NA | | 6 : 100 | Max. :87.00 | Max. :100.00 | NA | | (Other):2400 | NA | NA | NA |     Table continues below   | identity | intervention | physical\_punishment | warmth | | --- | --- | --- | --- | | Identity B:1507 | no intervention:1547 | Min. :0.000 | Min. :0.000 | | Identity A:1493 | intervention :1453 | 1st Qu.:2.000 | 1st Qu.:2.000 | | NA | NA | Median :2.000 | Median :4.000 | | NA | NA | Mean :2.479 | Mean :3.522 | | NA | NA | 3rd Qu.:3.000 | 3rd Qu.:5.000 | | NA | NA | Max. :5.000 | Max. :7.000 | | NA | NA | NA | NA |      | outcome | | --- | | Min. :29.61 | | 1st Qu.:48.02 | | Median :52.45 | | Mean :52.43 | | 3rd Qu.:56.86 | | Max. :74.84 | | NA | |

### 3.2.3 Graph

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| Figure 3.1: Graph from Data Stored in Statistical Software |

### 3.2.4 List Out A Sample Of The Data

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| Table 3.4: Sample of Data  Table continues below   | country | HDI | family | id | identity | intervention | | --- | --- | --- | --- | --- | --- | | 1 | 69 | 1 | 1.1 | Identity A | no intervention | | 1 | 69 | 2 | 1.2 | Identity A | intervention | | 1 | 69 | 3 | 1.3 | Identity B | intervention | | 1 | 69 | 4 | 1.4 | Identity A | no intervention | | 1 | 69 | 5 | 1.5 | Identity A | no intervention | | 1 | 69 | 6 | 1.6 | Identity B | intervention |      | physical\_punishment | warmth | outcome | | --- | --- | --- | | 3 | 3 | 57.47 | | 2 | 1 | 50.1 | | 3 | 2 | 52.92 | | 0 | 5 | 60.17 | | 4 | 4 | 55.05 | | 5 | 3 | 49.81 | |

## 3.3 Data In Spreadsheet Format

I now import the spreadsheet data file. I use the first row of data as variable names.

We see right away that the data are less informative.

### 3.3.1 Describe The Data

Notice how a description of the data no longer contains much of the information that helped us to understand the variables.

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
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| Table 3.5: Example Data As Stored in A Spreadsheet   | pos | variable | label | | --- | --- | --- | | 1 | country | NA | | 2 | HDI | NA | | 3 | family | NA | | 4 | id | NA | | 5 | identity | NA | | 6 | intervention | NA | | 7 | physical\_punishment | NA | | 8 | warmth | NA | | 9 | outcome | NA | |

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
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| Table 3.6: Example Data As Stored in A Spreadsheet  Table continues below   | country | HDI | family | id | identity | intervention | physical\_punishment | | --- | --- | --- | --- | --- | --- | --- | | 1 | 69 | 1 | 1.1 | 1 | 0 | 3 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 2 | | 1 | 69 | 3 | 1.3 | 0 | 1 | 3 | | 1 | 69 | 4 | 1.4 | 1 | 0 | 0 | | 1 | 69 | 5 | 1.5 | 1 | 0 | 4 | | 1 | 69 | 6 | 1.6 | 0 | 1 | 5 |      | warmth | outcome | | --- | --- | | 3 | 57.47 | | 1 | 50.1 | | 2 | 52.92 | | 5 | 60.17 | | 4 | 55.05 | | 3 | 49.81 | |

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| Warning |
| Adding this valuable information back into the data set may take a great deal of extra effort. |

### 3.3.2 Descriptive Statistics

Notice here how the descriptive statistics and graph are much less informative. For example, it is now not immediately clear what the values of identity or intervention represent. The information on variable labels and value labels will have to be added back into the data when preparing a final product for dissemination.

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
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| Table 3.7: Descriptive Statistics  Table continues below   | country | HDI | family | id | | --- | --- | --- | --- | | Min. : 1.0 | Min. :33.00 | Min. : 1.00 | Length:3000 | | 1st Qu.: 8.0 | 1st Qu.:53.00 | 1st Qu.: 25.75 | Class :character | | Median :15.5 | Median :70.00 | Median : 50.50 | Mode :character | | Mean :15.5 | Mean :64.77 | Mean : 50.50 | NA | | 3rd Qu.:23.0 | 3rd Qu.:81.00 | 3rd Qu.: 75.25 | NA | | Max. :30.0 | Max. :87.00 | Max. :100.00 | NA |     Table continues below   | identity | intervention | physical\_punishment | warmth | | --- | --- | --- | --- | | Min. :0.0000 | Min. :0.0000 | Min. :0.000 | Min. :0.000 | | 1st Qu.:0.0000 | 1st Qu.:0.0000 | 1st Qu.:2.000 | 1st Qu.:2.000 | | Median :0.0000 | Median :0.0000 | Median :2.000 | Median :4.000 | | Mean :0.4977 | Mean :0.4843 | Mean :2.479 | Mean :3.522 | | 3rd Qu.:1.0000 | 3rd Qu.:1.0000 | 3rd Qu.:3.000 | 3rd Qu.:5.000 | | Max. :1.0000 | Max. :1.0000 | Max. :5.000 | Max. :7.000 |      | outcome | | --- | | Min. :29.61 | | 1st Qu.:48.02 | | Median :52.45 | | Mean :52.43 | | 3rd Qu.:56.86 | | Max. :74.84 | |

### 3.3.3 Graph

While the graph has an informative title, as well as informative axis labels, a crucial piece of information is missing: what each status of the intervention represents.

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| Figure 3.2: Graph from Data Stored in Spreadsheet |

## 3.4 A Few Final Issues

Notice, finally, how spreadsheets don’t enforce the idea of whether variables are *numeric*, or *text*, and so would allow storage of different types of information in the same column. Relatedly, *numeric* variables may be improperly stored as *text*, often necessitating recoding before graphical or statistical procedures can be employed.

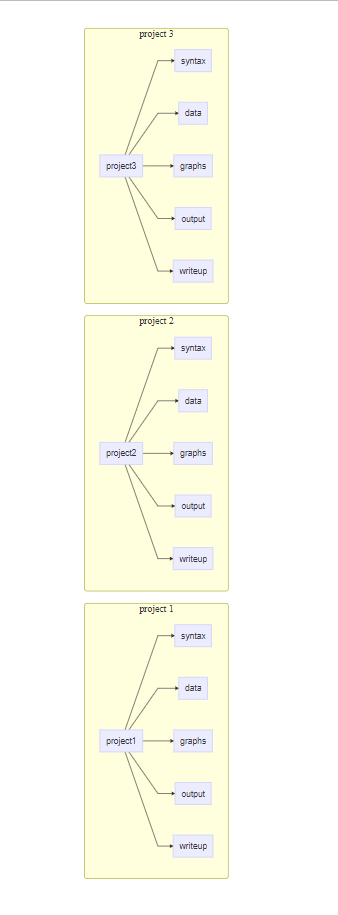
Second, a spreadsheet would allow some of your columns to have the same name, which might make data difficult to work with in other software.

Lastly, spreadsheets do not enforce the idea that the data have a *structure* wherein the *column header* is a *variable name*, while the *other rows* are *data*.

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Table 3.8: A Spreadsheet Table With Problematic Organization   | x | y | verylongvariablename | verylongvariablename | | --- | --- | --- | --- | | 100 | 1 | Smith | 20 | | 200 | 2 | 30 | NA | | not applicable | x | yes | 60 | |  |  | some other random information |  | |

## 3.5 File Organization

Files for all of your work should not be stored all together in downloads. Ideally, you should have a specific set of folders for your work. Each project, should be stored in its own individual folder. Ideally, each project folder would have a separate sub-folder for separate aspects of the project such as data, code or syntax, and various outputs.



A Hypothetical Set of Folders and Subfolders

# 4. Descriptive Statistics

## 4.1 Descriptive Statistics

### Stata

use simulated\_multilevel\_data.dta // use data

We use summarize for *continuous* variables, and tabulate for *categorical* variables.

summarize outcome warmth physical\_punishment HDI  
  
tabulate identity  
  
tabulate intervention

Variable | Obs Mean Std. dev. Min Max  
-------------+---------------------------------------------------------  
 outcome | 3,000 52.43327 6.530996 29.60798 74.83553  
 warmth | 3,000 3.521667 1.888399 0 7  
physical\_p~t | 3,000 2.478667 1.360942 0 5  
 HDI | 3,000 64.76667 17.24562 33 87  
  
  
hypothetica |  
 l identity |  
 group |  
 variable | Freq. Percent Cum.  
------------+-----------------------------------  
 0 | 1,507 50.23 50.23  
 1 | 1,493 49.77 100.00  
------------+-----------------------------------  
 Total | 3,000 100.00  
  
  
 recieved |  
interventio |  
 n | Freq. Percent Cum.  
------------+-----------------------------------  
 0 | 1,547 51.57 51.57  
 1 | 1,453 48.43 100.00  
------------+-----------------------------------  
 Total | 3,000 100.00

### R

library(haven) # read data in Stata format  
  
df <- read\_dta("simulated\_multilevel\_data.dta")

R’s descriptive statistics functions rely heavily on whether a variable is a *numeric* variable, or a *factor* variable. Below, I convert two variables to factors (factor) before using summary[[1]](#footnote-94) to generate descriptive statistics.

df$country <- factor(df$country)  
  
df$identity <- factor(df$identity)  
  
df$intervention <- factor(df$intervention)  
  
summary(df)

country HDI family id identity  
 1 : 100 Min. :33.00 Min. : 1.00 Length:3000 0:1507   
 2 : 100 1st Qu.:53.00 1st Qu.: 25.75 Class :character 1:1493   
 3 : 100 Median :70.00 Median : 50.50 Mode :character   
 4 : 100 Mean :64.77 Mean : 50.50   
 5 : 100 3rd Qu.:81.00 3rd Qu.: 75.25   
 6 : 100 Max. :87.00 Max. :100.00   
 (Other):2400   
 intervention physical\_punishment warmth outcome   
 0:1547 Min. :0.000 Min. :0.000 Min. :29.61   
 1:1453 1st Qu.:2.000 1st Qu.:2.000 1st Qu.:48.02   
 Median :2.000 Median :4.000 Median :52.45   
 Mean :2.479 Mean :3.522 Mean :52.43   
 3rd Qu.:3.000 3rd Qu.:5.000 3rd Qu.:56.86   
 Max. :5.000 Max. :7.000 Max. :74.84

### Julia

using Tables, MixedModels, MixedModelsExtras, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
df = DataFrame(load("simulated\_multilevel\_data.dta"))

Similarly to R, Julia relies on the idea of *variable type*. I use transform to convert the appropriate variables to *categorical* variables.

@transform!(df, :country = categorical(:country))  
  
@transform!(df, :identity = categorical(:identity))  
  
@transform!(df, :intervention = categorical(:intervention))

describe(df) # descriptive statistics

9×7 DataFrame  
 Row │ variable mean min median max nmissing eltype ⋯  
 │ Symbol Union… Any Union… Any Int64 Union ⋯  
─────┼──────────────────────────────────────────────────────────────────────────  
 1 │ country 1.0 30.0 0 Union{ ⋯  
 2 │ HDI 64.7667 33.0 70.0 87.0 0 Union{  
 3 │ family 50.5 1.0 50.5 100.0 0 Union{  
 4 │ id 1.1 9.99 0 Union{  
 5 │ identity 0.0 1.0 0 Union{ ⋯  
 6 │ intervention 0.0 1.0 0 Union{  
 7 │ physical\_punishment 2.47867 0.0 2.0 5.0 0 Union{  
 8 │ warmth 3.52167 0.0 4.0 7.0 0 Union{  
 9 │ outcome 52.4333 29.608 52.449 74.8355 0 Union{ ⋯  
 1 column omitted

## 4.2 Interpretation

Examining descriptive statistics is an important first step in any analysis. It is important to examine your descriptive statistics before skipping ahead to more sophisticated analyses, such as multilevel models.

In examining the descriptive statistics for this data, we get a sense of the data.

* outcome has a mean of approximately 52 and ranges from approximately 30 to 75.
* warmth and physical punishment are both variables that represent the number of times that parents use each of these forms of discipline in a week. The average of the former is about 3.5, while the average of the latter is about 2.5.
* HDI, the Human Development Index has an average of about 65, and a wide range.
* identity is a categorical variable for a hypothetical identity group, and has values of 0 and 1.
* intervention is also a categorical variable, and has values of 0 and 1.

# 5. Unconditional Models

## 5.1 Two Level Model

An *unconditional* multilevel model is a model with no independent variables. One should always run an unconditional model as the first step of a multilevel model in order to get a sense of the way that variation is apportioned in the model across the different levels.

### 5.1.1 The Equation

The Intraclass Correlation Coefficient (ICC) is given by:

In a two level multilevel model, the ICC provides a measure of the proportion of variation attributable to Level 2.

### 5.1.2 Run Models

#### Stata

use simulated\_multilevel\_data.dta // use data  
  
mixed outcome || country: // unconditional model  
  
estat icc // ICC

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -9802.8371   
Iteration 1: Log likelihood = -9802.8371   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 3,000  
Group variable: country Number of groups = 30  
 Obs per group:  
 min = 100  
 avg = 100.0  
 max = 100  
 Wald chi2(0) = .  
Log likelihood = -9802.8371 Prob > chi2 = .  
  
------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
-------------+----------------------------------------------------------------  
 \_cons | 52.43327 .3451217 151.93 0.000 51.75685 53.1097  
------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
country: Identity |  
 var(\_cons) | 3.178658 .9226737 1.799552 5.614658  
-----------------------------+------------------------------------------------  
 var(Residual) | 39.46106 1.024013 37.50421 41.52  
------------------------------------------------------------------------------  
LR test vs. linear model: chibar2(01) = 166.31 Prob >= chibar2 = 0.0000  
  
  
Intraclass correlation  
  
------------------------------------------------------------------------------  
 Level | ICC Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
 country | .0745469 .0201254 .0434963 .1248696  
------------------------------------------------------------------------------

#### R

library(haven)  
  
df <- read\_dta("simulated\_multilevel\_data.dta")

library(lme4) # estimate multilevel models  
  
fit0 <- lmer(outcome ~ (1 | country),  
 data = df) # unconditional model  
  
summary(fit0)

Linear mixed model fit by REML ['lmerMod']  
Formula: outcome ~ (1 | country)  
 Data: df  
  
REML criterion at convergence: 19605.9  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.3844 -0.6655 -0.0086 0.6725 3.6626   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 country (Intercept) 3.302 1.817   
 Residual 39.461 6.282   
Number of obs: 3000, groups: country, 30  
  
Fixed effects:  
 Estimate Std. Error t value  
(Intercept) 52.433 0.351 149.4

library(performance)  
  
performance::icc(fit0) # ICC

# Intraclass Correlation Coefficient  
  
 Adjusted ICC: 0.077  
 Unadjusted ICC: 0.077

#### Julia

using Tables, MixedModels, MixedModelsExtras,   
StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
df = DataFrame(load("simulated\_multilevel\_data.dta"))

@transform!(df, :country = categorical(:country))

m0 = fit(MixedModel,   
 @formula(outcome ~ (1 | country)), df) # unconditional model

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + (1 | country)  
 logLik -2 logLik AIC AICc BIC   
 -9802.8371 19605.6742 19611.6742 19611.6822 19629.6933  
  
Variance components:  
 Column Variance Std.Dev.  
country (Intercept) 3.17863 1.78287  
Residual 39.46106 6.28180  
 Number of obs: 3000; levels of grouping factors: 30  
  
 Fixed-effects parameters:  
──────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
──────────────────────────────────────────────────  
(Intercept) 52.4333 0.345121 151.93 <1e-99  
──────────────────────────────────────────────────

icc(m0) # ICC

0.07454637475695493

### 5.1.3 Interpretation

In each case, the software finds that nearly 8% of the variation in the outcome is explainable by the clustering of the observations in each country.

## 5.2 Three Level Model

### 5.2.1 The Equation

As discussed in the main text, in a three level model, there are two intraclass correlation coefficients (StataCorp, 2023). The formulas for the Intraclass Correlation Coefficient (ICC) are given by (StataCorp, 2023):

Following StataCorp (2023), [Equation 5.4](#eq-ICCunconditional3A) is the correlation of responses for person-timepoints from the same country but different persons.

Again, closely following StataCorp (2023), [Equation 5.5](#eq-ICCunconditional3B) is the correlation of responses for person-timepoints from the same country and same person.

### 5.2.2 Run Models

#### Stata

use simulated\_multilevel\_longitudinal\_data.dta // use data  
  
mixed outcome || country: || id: // unconditional model  
   
estat icc // ICC

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -29058.266   
Iteration 1: Log likelihood = -29058.259   
Iteration 2: Log likelihood = -29058.259   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 9,000  
  
 Grouping information  
 -------------------------------------------------------------  
 | No. of Observations per group  
 Group variable | groups Minimum Average Maximum  
 ----------------+--------------------------------------------  
 country | 30 300 300.0 300  
 id | 3,000 3 3.0 3  
 -------------------------------------------------------------  
  
 Wald chi2(0) = .  
Log likelihood = -29058.259 Prob > chi2 = .  
  
------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
-------------+----------------------------------------------------------------  
 \_cons | 53.37768 .3387943 157.55 0.000 52.71366 54.04171  
------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
country: Identity |  
 var(\_cons) | 3.232092 .8891367 1.885043 5.54174  
-----------------------------+------------------------------------------------  
id: Identity |  
 var(\_cons) | 11.72403 .5747501 10.64996 12.90641  
-----------------------------+------------------------------------------------  
 var(Residual) | 28.23424 .5154843 27.24178 29.26287  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(2) = 1314.88 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.  
  
  
Intraclass correlation  
  
------------------------------------------------------------------------------  
 Level | ICC Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
 country | .0748336 .0190847 .0450028 .1219141  
 id|country | .3462837 .0171461 .3134867 .3806097  
------------------------------------------------------------------------------

#### R

In R, the ICC for a three level model is easiest to estimate “by hand”.

library(haven)  
  
dfL <- read\_dta("simulated\_multilevel\_longitudinal\_data.dta")

library(lme4) # estimate multilevel models  
  
fit0L <- lmer(outcome ~ (1 | country/id),  
 data = dfL) # unconditional model  
  
summary(fit0L)

Linear mixed model fit by REML ['lmerMod']  
Formula: outcome ~ (1 | country/id)  
 Data: dfL  
  
REML criterion at convergence: 58116.8  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.7858 -0.6059 -0.0062 0.6017 3.4348   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 id:country (Intercept) 11.724 3.424   
 country (Intercept) 3.351 1.830   
 Residual 28.234 5.314   
Number of obs: 9000, groups: id:country, 3000; country, 30  
  
Fixed effects:  
 Estimate Std. Error t value  
(Intercept) 53.3777 0.3446 154.9

3.351 / (11.724 + 3.351 + 28.234)

[1] 0.07737422

(3.351 + 11.724) / (11.724 + 3.351 + 28.234)

[1] 0.3480801

#### Julia

In Julia, the ICC for a three level model is also easiest to estimate “by hand”.

using Tables, MixedModels, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
dfL = DataFrame(load("simulated\_multilevel\_longitudinal\_data.dta"))

@transform!(dfL, :country = categorical(:country))

m0L = fit(MixedModel, @formula(outcome ~   
 (1 | country) +   
 (1 | id)), dfL)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + (1 | country) + (1 | id)  
 logLik -2 logLik AIC AICc BIC   
 -29058.2592 58116.5184 58124.5184 58124.5229 58152.9384  
  
Variance components:  
 Column Variance Std.Dev.  
id (Intercept) 11.72401 3.42403  
country (Intercept) 3.23190 1.79775  
Residual 28.23426 5.31359  
 Number of obs: 9000; levels of grouping factors: 3000, 30  
  
 Fixed-effects parameters:  
──────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
──────────────────────────────────────────────────  
(Intercept) 53.3777 0.338785 157.56 <1e-99  
──────────────────────────────────────────────────

3.23190 / (11.72401 + 3.23190 + 28.23426)

0.07482952718176382

(3.23190 + 11.72401) / (11.72401 + 3.23190 + 28.23426)

0.34628041519632824

### 5.2.3 Interpretation

Each software suggests that almost 8% of the variation in the outcome is within time points for different individuals within the same country, while almost 35% of the variation in the outcome is within time points for the same individual within the same country.

# 6. Cross Sectional Multilevel Models

## 6.1 The Equation

Recall the general model of [Equation 1.1](#eq-MLMsimple), and the syntax outlined in [Section 1.3](#sec-syntax). Below in [Equation 6.1](#eq-MLMsubstantive), we consider a more substantive example.

## 6.2 Correlated and Uncorrelated Random Effects

Consider the covariance matrix of random effects (e.g.  and ). In [Equation 6.2](#eq-varcovar) the covariances of the random effects are constrained to be zero.

As discussed in the Chapter on multilevel models with cross-sectional data, however, one can consider a multilevel model in which the random effects are correlated, as is the case in [Equation 6.3](#eq-varcovaruns).

Procedures for estimating models with uncorrelated and correlated random effects are detailed below (Bates et al., 2015; Bates, 2024; StataCorp, 2023).

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Table 6.1: Correlated and Uncorrelated Random Effects   | Software | Uncorrelated Random Effects | Correlated Random Effects | | --- | --- | --- | | Stata | default | add option: , cov(uns) | | R | separate random effects from grouping variable with || | separate random effects from grouping variable with | | | Julia | separate terms for each random effect e.g. (1 | group) + (0 + x | group) | separate random effects from grouping variable with |. | |

All models in the examples below are run with *uncorrelated* random effects, but could just as easily be run with *correlated* random effects.

## 6.3 Run The Models

|  |
| --- |
| Continuous and Categorical Variables |
| Statistically–as noted in the main text–it is important to be clear on whether independent variables in one’s model are continuous or categorical. *Continuous* variables can be entered straightforwardly into statistical syntax. *Categorical* variables, on the other hand usually require specific attention in statistical software. In Stata, categorical variables are indicated in a statistical model by prefixing them with an i.. In R, categorical variables are distinguished by making them into factors e.g. x <- factor(x). In Julia, categorical variables are created by using the @transform syntax detailed below. |

### Stata

#### 6.3.0.1 Get The Data

use simulated\_multilevel\_data.dta

#### 6.3.0.2 Run The Model

mixed outcome warmth physical\_punishment i.identity i.intervention HDI || ///   
country: warmth

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -9626.6279   
Iteration 1: Log likelihood = -9626.607   
Iteration 2: Log likelihood = -9626.607   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 3,000  
Group variable: country Number of groups = 30  
 Obs per group:  
 min = 100  
 avg = 100.0  
 max = 100  
 Wald chi2(5) = 334.14  
Log likelihood = -9626.607 Prob > chi2 = 0.0000  
  
-------------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
--------------------+----------------------------------------------------------------  
 warmth | .8345368 .0637213 13.10 0.000 .7096453 .9594282  
physical\_punishment | -.9916657 .0797906 -12.43 0.000 -1.148052 -.8352791  
 1.identity | -.3004767 .2170295 -1.38 0.166 -.7258466 .1248933  
 1.intervention | .6396427 .2174519 2.94 0.003 .2134448 1.065841  
 HDI | -.003228 .0199257 -0.16 0.871 -.0422817 .0358256  
 \_cons | 51.99991 1.371257 37.92 0.000 49.3123 54.68753  
-------------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
country: Independent |  
 var(warmth) | .0227504 .0257784 .0024689 .2096436  
 var(\_cons) | 2.963975 .9737647 1.556777 5.643163  
-----------------------------+------------------------------------------------  
 var(Residual) | 34.97499 .9097109 33.23668 36.80422  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(2) = 205.74 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

### R

#### 6.3.0.3 Get The Data

library(haven)  
  
df <- read\_dta("simulated\_multilevel\_data.dta")

#### 6.3.0.4 Change Some Variables To Categorical

df$identity <- factor(df$identity)  
  
df$intervention <- factor(df$intervention)

#### 6.3.0.5 Run The Model

|  |
| --- |
| Caution |
| lme4 does not directly provide p values in results, because of some disagreement over exactly how these p values should be calculated. Therefore, in this Appendix, I also call library lmerTest to provide p values for lme4 results. |

|  |
| --- |
| Tip |
| R prefers to use scientific notation when possible. I find that the use of scientific notation can be confusing in reading results. I turn off scientific notation by setting a penalty for its use: options(scipen = 999). |

library(lme4)   
  
library(lmerTest)  
  
options(scipen = 999)   
  
fit1 <- lmer(outcome ~ warmth + physical\_punishment +   
 identity + intervention + HDI +  
 (1 + warmth || country),  
 data = df)  
  
summary(fit1)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula: outcome ~ warmth + physical\_punishment + identity + intervention +   
 HDI + (1 + warmth || country)  
 Data: df  
  
REML criterion at convergence: 19268.8  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.9774 -0.6563 0.0186 0.6645 3.6730   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 country (Intercept) 3.19120 1.786   
 country.1 warmth 0.02464 0.157   
 Residual 35.01779 5.918   
Number of obs: 3000, groups: country, 30  
  
Fixed effects:  
 Estimate Std. Error df t value  
(Intercept) 52.011324 1.414976 30.293141 36.758  
warmth 0.834562 0.064250 41.896457 12.989  
physical\_punishment -0.991893 0.079845 2968.012381 -12.423  
identity1 -0.300354 0.217179 2970.108153 -1.383  
intervention1 0.639060 0.217603 2971.186718 2.937  
HDI -0.003394 0.020598 27.592814 -0.165  
 Pr(>|t|)   
(Intercept) < 0.0000000000000002 \*\*\*  
warmth 0.000000000000000277 \*\*\*  
physical\_punishment < 0.0000000000000002 \*\*\*  
identity1 0.16678   
intervention1 0.00334 \*\*   
HDI 0.87030   
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1  
  
Correlation of Fixed Effects:  
 (Intr) warmth physc\_ idntt1 intrv1  
warmth -0.124   
physcl\_pnsh -0.149 -0.003   
identity1 -0.072 -0.012 -0.003   
interventn1 -0.082 0.034 0.022 -0.018   
HDI -0.943 -0.006 0.009 -0.001 0.000

### Julia

#### 6.3.0.6 Get The Data

using Tables, MixedModels, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
df = DataFrame(load("simulated\_multilevel\_data.dta"))

#### 6.3.0.7 Change Some Variables To Categorical

@transform!(df, :country = categorical(:country))  
  
@transform!(df, :identity = categorical(:identity))  
  
@transform!(df, :intervention = categorical(:intervention))

#### 6.3.0.8 Run The Model

m1 = fit(MixedModel, @formula(outcome ~ warmth + physical\_punishment +   
 identity + intervention + HDI +  
 (1 | country) +  
 (0 + warmth | country)), df)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + warmth + physical\_punishment + identity + intervention + HDI + (1 | country) + (0 + warmth | country)  
 logLik -2 logLik AIC AICc BIC   
 -9626.6070 19253.2140 19271.2140 19271.2742 19325.2713  
  
Variance components:  
 Column Variance Std.Dev. Corr.  
country (Intercept) 2.963849 1.721583  
 warmth 0.022756 0.150852 .   
Residual 34.974984 5.913965  
 Number of obs: 3000; levels of grouping factors: 30  
  
 Fixed-effects parameters:  
─────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
─────────────────────────────────────────────────────────────  
(Intercept) 51.9999 1.37124 37.92 <1e-99  
warmth 0.834537 0.0637228 13.10 <1e-38  
physical\_punishment -0.991665 0.0797906 -12.43 <1e-34  
identity: 1.0 -0.300475 0.217029 -1.38 0.1662  
intervention: 1.0 0.639641 0.217452 2.94 0.0033  
HDI -0.0032286 0.0199255 -0.16 0.8713  
─────────────────────────────────────────────────────────────

## 6.4 Interpretation

Models suggest that parental warmth is associated with increases in the beneficial outcome, while physical punishment is associated with decreases in the beneficial outcome. The intervention is associated with increases in the outcome. There is insufficient evidence that either identity group or the Human Development Index are associated with the outcome.

# 7. Longitudinal Multilevel Models

## 7.1 The Data

The data employed in these examples are a longitudinal extension of the data described in [Section 1.2](#sec-data).

## 7.2 The Equation

## 7.3 Growth Trajectories

Remember, following the discussion in the main text, that in longitudinal multilevel models, the variable for *time* assumes an important role as we are often thinking of a *growth trajectory* over time.

As discussed in the main text, think about a model where *identity* is a (1/0) variable for membership in one of two groups:

Then, each identity group has its own intercept and time trajectory:

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Table 7.1: Slope and Intercept for Each Group   | Group | Intercept | Slope (Time Trajectory) | | --- | --- | --- | | 0 |  |  | | 1 |  |  | |

|  |
| --- |
| Main Effects and Interactions |
| Thus, again following the main text, in longitudinal multilevel models, *main effects* modify the *intercept* of the time trajectory, while *interactions with time*, modify the *slope* of the time trajectory. Below, we run models with *main effects only*, then models with *main effects, and interactions with time*. |

## 7.4 Run The Models

|  |
| --- |
| Warning |
| Remember that we are estimating a model in which time points are nested inside families, who are in turn nested inside countries. For each software package, it is accordingly important to specify the way in which different levels of the data are nested. Pay careful attention to the syntax examples below with regard to id and country |

### Stata

#### 7.4.0.1 Get The Data

use simulated\_multilevel\_longitudinal\_data.dta

#### 7.4.0.2 Run The Models

##### 7.4.0.2.1 Main Effects Only

mixed outcome t warmth physical\_punishment i.identity i.intervention HDI || ///   
country: || id: t

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -28523.888   
Iteration 1: Log likelihood = -28500.378   
Iteration 2: Log likelihood = -28500.105   
Iteration 3: Log likelihood = -28500.086   
Iteration 4: Log likelihood = -28500.086   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 9,000  
  
 Grouping information  
 -------------------------------------------------------------  
 | No. of Observations per group  
 Group variable | groups Minimum Average Maximum  
 ----------------+--------------------------------------------  
 country | 30 300 300.0 300  
 id | 3,000 3 3.0 3  
 -------------------------------------------------------------  
  
 Wald chi2(6) = 1208.49  
Log likelihood = -28500.086 Prob > chi2 = 0.0000  
  
-------------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
--------------------+----------------------------------------------------------------  
 t | .9433792 .0658667 14.32 0.000 .8142828 1.072476  
 warmth | .9140251 .0379154 24.11 0.000 .8397122 .988338  
physical\_punishment | -1.00861 .0497766 -20.26 0.000 -1.106171 -.9110499  
 1.identity | -.1319026 .1516462 -0.87 0.384 -.4291236 .1653184  
 1.intervention | .8592402 .1519616 5.65 0.000 .5614009 1.15708  
 HDI | .0007913 .0200615 0.04 0.969 -.0385285 .040111  
 \_cons | 50.38381 1.367464 36.84 0.000 47.70363 53.06399  
-------------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
country: Identity |  
 var(\_cons) | 3.418565 .9268849 2.009349 5.816108  
-----------------------------+------------------------------------------------  
id: Independent |  
 var(t) | 1.27e-08 2.25e-06 7.3e-160 2.2e+143  
 var(\_cons) | 8.42116 .4720261 7.545013 9.399046  
-----------------------------+------------------------------------------------  
 var(Residual) | 26.02918 .4753157 25.11405 26.97765  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(3) = 1246.06 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

##### 7.4.0.2.2 Interactions With Time

mixed outcome c.t##(c.warmth c.physical\_punishment i.identity i.intervention c.HDI) || country: warmth || id: t

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -28522.21   
Iteration 1: Log likelihood = -28498.677   
Iteration 2: Log likelihood = -28498.468   
Iteration 3: Log likelihood = -28498.31   
Iteration 4: Log likelihood = -28498.309   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 9,000  
  
 Grouping information  
 -------------------------------------------------------------  
 | No. of Observations per group  
 Group variable | groups Minimum Average Maximum  
 ----------------+--------------------------------------------  
 country | 30 300 300.0 300  
 id | 3,000 3 3.0 3  
 -------------------------------------------------------------  
  
 Wald chi2(11) = 1100.25  
Log likelihood = -28498.309 Prob > chi2 = 0.0000  
  
---------------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
----------------------+----------------------------------------------------------------  
 t | .7582075 .326177 2.32 0.020 .1189123 1.397503  
 warmth | .8170757 .082662 9.88 0.000 .6550611 .9790903  
 physical\_punishment | -1.009031 .1112932 -9.07 0.000 -1.227162 -.7909007  
 1.identity | -.2387167 .3039964 -0.79 0.432 -.8345387 .3571053  
 1.intervention | .6607606 .3044503 2.17 0.030 .064049 1.257472  
 HDI | .0013614 .0210842 0.06 0.949 -.0399628 .0426856  
 |  
 c.t#c.warmth | .0483637 .0356074 1.36 0.174 -.0214255 .1181529  
 |  
 c.t#|  
c.physical\_punishment | .0005421 .0494354 0.01 0.991 -.0963496 .0974338  
 |  
 identity#c.t |  
 1 | .0554389 .1317444 0.42 0.674 -.2027754 .3136532  
 |  
 intervention#c.t |  
 1 | .0992811 .131925 0.75 0.452 -.1592872 .3578493  
 |  
 c.t#c.HDI | -.0009551 .0038216 -0.25 0.803 -.0084453 .0065352  
 |  
 \_cons | 50.83632 1.483548 34.27 0.000 47.92862 53.74402  
---------------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
country: Independent |  
 var(warmth) | .0106014 .0127458 .0010046 .1118779  
 var(\_cons) | 3.170088 .9153355 1.80009 5.582753  
-----------------------------+------------------------------------------------  
id: Independent |  
 var(t) | 9.47e-10 2.07e-07 1.5e-195 6.0e+176  
 var(\_cons) | 8.39189 .4724106 7.515234 9.370809  
-----------------------------+------------------------------------------------  
 var(Residual) | 26.01583 .4751602 25.10101 26.964  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(4) = 1247.84 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

### R

#### 7.4.0.3 Get The Data

library(haven)  
  
dfL <- read\_dta("simulated\_multilevel\_longitudinal\_data.dta")

#### 7.4.0.4 Change Some Variables To Categorical

dfL$identity <- factor(dfL$identity)  
  
dfL$intervention <- factor(dfL$intervention)

#### 7.4.0.5 Run The Models

|  |
| --- |
| Caution |
| lme4 does not directly provide p values in results, because of some disagreement over exactly how these p values should be calculated. Therefore, in this Appendix, I also call library lmerTest to provide p values for lme4 results. |

|  |
| --- |
| Tip |
| R prefers to use scientific notation when possible. I find that the use of scientific notation can be confusing in reading results. I turn off scientific notation by setting a penalty for its use: options(scipen = 999). |

##### 7.4.0.5.1 Main Effects Only

library(lme4)   
  
library(lmerTest)  
  
options(scipen = 999)   
  
fit2A <- lmer(outcome ~ t + warmth + physical\_punishment +   
 identity + intervention + HDI +  
 (1 | country/id),  
 data = dfL)  
  
summary(fit2A)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula:   
outcome ~ t + warmth + physical\_punishment + identity + intervention +   
 HDI + (1 | country/id)  
 Data: dfL  
  
REML criterion at convergence: 57022.7  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.6850 -0.6094 -0.0035 0.6133 3.6792   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 id:country (Intercept) 8.438 2.905   
 country (Intercept) 3.675 1.917   
 Residual 26.036 5.103   
Number of obs: 9000, groups: id:country, 3000; country, 30  
  
Fixed effects:  
 Estimate Std. Error df t value  
(Intercept) 50.3842343 1.4139114 29.8246912 35.635  
t 0.9433806 0.0658755 5998.3764548 14.321  
warmth 0.9140307 0.0379336 4745.3497493 24.096  
physical\_punishment -1.0087537 0.0497972 6483.6771808 -20.257  
identity1 -0.1319548 0.1517350 2968.7828107 -0.870  
intervention1 0.8591494 0.1520510 2971.8111995 5.650  
HDI 0.0007909 0.0207656 28.0001855 0.038  
 Pr(>|t|)   
(Intercept) < 0.0000000000000002 \*\*\*  
t < 0.0000000000000002 \*\*\*  
warmth < 0.0000000000000002 \*\*\*  
physical\_punishment < 0.0000000000000002 \*\*\*  
identity1 0.385   
intervention1 0.0000000175 \*\*\*  
HDI 0.970   
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1  
  
Correlation of Fixed Effects:  
 (Intr) t warmth physc\_ idntt1 intrv1  
t -0.092   
warmth -0.091 -0.002   
physcl\_pnsh -0.092 -0.007 -0.012   
identity1 -0.051 0.000 -0.013 -0.003   
interventn1 -0.058 0.000 0.039 0.019 -0.018   
HDI -0.951 0.000 -0.004 0.005 0.000 0.002

##### 7.4.0.5.2 Interactions With Time

fit2B <- lmer(outcome ~ t \*(warmth + physical\_punishment +   
 identity + intervention + HDI) +  
 (1 | country/id),  
 data = dfL)  
  
summary(fit2B)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula:   
outcome ~ t \* (warmth + physical\_punishment + identity + intervention +   
 HDI) + (1 | country/id)  
 Data: dfL  
  
REML criterion at convergence: 57042.8  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.7118 -0.6092 -0.0024 0.6150 3.6779   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 id:country (Intercept) 8.436 2.905   
 country (Intercept) 3.675 1.917   
 Residual 26.046 5.104   
Number of obs: 9000, groups: id:country, 3000; country, 30  
  
Fixed effects:  
 Estimate Std. Error df t value  
(Intercept) 50.7590272 1.5518360 43.2608620 32.709  
t 0.7552909 0.3263028 6176.7440549 2.315  
warmth 0.8170912 0.0805355 8274.9995422 10.146  
physical\_punishment -1.0097729 0.1113557 8084.6084915 -9.068  
identity1 -0.2446453 0.3041604 8695.8966197 -0.804  
intervention1 0.6604671 0.3046286 8697.0843430 2.168  
HDI 0.0026692 0.0221295 36.1037733 0.121  
t:warmth 0.0486211 0.0356217 6404.8723333 1.365  
t:physical\_punishment 0.0004964 0.0494590 6753.0158441 0.010  
t:identity1 0.0563140 0.1318043 5993.4518022 0.427  
t:intervention1 0.0995037 0.1319917 5994.1433001 0.754  
t:HDI -0.0009379 0.0038233 5993.9090880 -0.245  
 Pr(>|t|)   
(Intercept) <0.0000000000000002 \*\*\*  
t 0.0207 \*   
warmth <0.0000000000000002 \*\*\*  
physical\_punishment <0.0000000000000002 \*\*\*  
identity1 0.4212   
intervention1 0.0302 \*   
HDI 0.9047   
t:warmth 0.1723   
t:physical\_punishment 0.9920   
t:identity1 0.6692   
t:intervention1 0.4510   
t:HDI 0.8062   
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1  
  
Correlation of Fixed Effects:  
 (Intr) t warmth physc\_ idntt1 intrv1 HDI t:wrmt t:phy\_  
t -0.421   
warmth -0.178 0.331   
physcl\_pnsh -0.190 0.360 -0.005   
identity1 -0.093 0.166 -0.013 -0.002   
interventn1 -0.107 0.192 0.039 0.019 -0.017   
HDI -0.925 0.264 -0.007 0.012 -0.001 0.003   
t:warmth 0.158 -0.377 -0.882 0.001 0.011 -0.035 0.006   
t:physcl\_pn 0.170 -0.402 0.004 -0.894 -0.001 -0.017 -0.010 -0.003   
t:identity1 0.081 -0.192 0.011 0.000 -0.867 0.014 0.001 -0.013 0.002  
t:intrvntn1 0.093 -0.222 -0.035 -0.017 0.014 -0.867 -0.003 0.041 0.019  
t:HDI 0.322 -0.765 0.015 -0.027 0.002 -0.007 -0.346 -0.016 0.029  
 t:dnt1 t:ntr1  
t   
warmth   
physcl\_pnsh   
identity1   
interventn1   
HDI   
t:warmth   
t:physcl\_pn   
t:identity1   
t:intrvntn1 -0.016   
t:HDI -0.002 0.008

### Julia

#### 7.4.0.6 Get The Data

using Tables, MixedModels, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
dfL = DataFrame(load("simulated\_multilevel\_longitudinal\_data.dta"))

#### 7.4.0.7 Change Some Variables To Categorical

@transform!(dfL, :country = categorical(:country))  
  
@transform!(dfL, :identity = categorical(:identity))  
  
@transform!(dfL, :intervention = categorical(:intervention))

#### 7.4.0.8 Run The Models

##### 7.4.0.8.1 Main Effects Only

m2A = fit(MixedModel, @formula(outcome ~ t + warmth +   
 physical\_punishment +   
 identity + intervention +   
 HDI +  
 (1 | country) +   
 (0 + warmth | country) +  
 (1 | id)), dfL)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + t + warmth + physical\_punishment + identity + intervention + HDI + (1 | country) + (0 + warmth | country) + (1 | id)  
 logLik -2 logLik AIC AICc BIC   
 -28499.6031 56999.2063 57021.2063 57021.2356 57099.3610  
  
Variance components:  
 Column Variance Std.Dev. Corr.  
id (Intercept) 8.387214 2.896069  
country (Intercept) 3.167143 1.779647  
 warmth 0.010762 0.103739 .   
Residual 26.027363 5.101702  
 Number of obs: 9000; levels of grouping factors: 3000, 30  
  
 Fixed-effects parameters:  
───────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
───────────────────────────────────────────────────────────────  
(Intercept) 50.4673 1.33833 37.71 <1e-99  
t 0.943864 0.0658717 14.33 <1e-45  
warmth 0.913496 0.0423744 21.56 <1e-99  
physical\_punishment -1.0079 0.0497622 -20.25 <1e-90  
identity: 1.0 -0.127692 0.151583 -0.84 0.3996  
intervention: 1.0 0.858997 0.151909 5.65 <1e-07  
HDI -0.000566026 0.0196439 -0.03 0.9770  
───────────────────────────────────────────────────────────────

##### 7.4.0.8.2 Interactions With Time

m2B = fit(MixedModel, @formula(outcome ~ t \* (warmth +   
 physical\_punishment +   
 identity + intervention +   
 HDI) +  
 (1 | country) +  
 (0 + warmth | country) +  
 (1 | id)), dfL)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + t + warmth + physical\_punishment + identity + intervention + HDI + t & warmth + t & physical\_punishment + t & identity + t & intervention + t & HDI + (1 | country) + (0 + warmth | country) + (1 | id)  
 logLik -2 logLik AIC AICc BIC   
 -28498.3091 56996.6182 57028.6182 57028.6788 57142.2979  
  
Variance components:  
 Column Variance Std.Dev. Corr.  
id (Intercept) 8.391746 2.896851  
country (Intercept) 3.170032 1.780458  
 warmth 0.010609 0.102999 .   
Residual 26.015906 5.100579  
 Number of obs: 9000; levels of grouping factors: 3000, 30  
  
 Fixed-effects parameters:  
──────────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
──────────────────────────────────────────────────────────────────  
(Intercept) 50.8364 1.48355 34.27 <1e-99  
t 0.758209 0.326177 2.32 0.0201  
warmth 0.817076 0.0826636 9.88 <1e-22  
physical\_punishment -1.00903 0.111293 -9.07 <1e-18  
identity: 1.0 -0.238714 0.303996 -0.79 0.4323  
intervention: 1.0 0.660761 0.30445 2.17 0.0300  
HDI 0.00136065 0.0210842 0.06 0.9485  
t & warmth 0.0483635 0.0356074 1.36 0.1744  
t & physical\_punishment 0.000542203 0.0494355 0.01 0.9912  
t & identity: 1.0 0.0554385 0.131745 0.42 0.6739  
t & intervention: 1.0 0.0992809 0.131925 0.75 0.4517  
t & HDI -0.000955067 0.00382162 -0.25 0.8027  
──────────────────────────────────────────────────────────────────

## 7.5 Interpretation

The *main effects only model* suggests that time is associated with increases in the outcome. In the main effects model, main effects other than time, indicate whether a particular variable is associated with higher or lower intercepts of the time trajectory, at the beginning of the study time. Warmth is again associated with increases in the outcome, while physical punishment is associated with decreases in the outcome. Identity is again not associated with the outcome, while the intervention is associated with higher levels of the outcome. The Human Development Index is again not associated with the outcome.

The second model adds interactions with time to the first model. Results are largely similar to the prior model. However, here we not only examine whether main effects other than time are associated with higher or lower time trajectories, but also whether particular variables are associated with differences in the slope of the time trajectory. In this case, we find insufficient evidence that any independent variable is associated with changes in the slope of the time trajectory.

|  |
| --- |
| Which Interactions To Test? |
| In this example–for the sake of illustration–I test the interaction of *every* independent variable with time. In many cases, it may make sense to test only only one or two interactions of time with particular variables of key interest. Also, after finding, as I did in this model, that none of the interactions of other independent variables with time are significant, I might report the model with interactions, or might report only the results of the model with only main effects. |

It may be illustrative to imagine how we would interpret the results had a particular interaction term been statistically significant. Let us consider one of the interaction terms with the largest coefficient, intervention#time. The interaction of the intervention with time is positive. Had this coefficient been statistically signifcant, it would have indicated that the intervention was associated with more rapid increases in the outcome over time *in addition to* the fact that the intervention is associated with higher initial levels of the outcome.

# 8. Multilevel Logistic Regression

Below, I detail the procedure for multilevel logistic regression models in Stata and R.

## 8.1 The Data

The data employed in these examples are the cross-sectional data described in [Section 1.2](#sec-data).

## 8.2 The Equation

To explain statistical syntax for Stata and R, I consider the general case of a multilevel model with *categorical* dependent variable y, independent variables x and z, clustering variable group, and a random slope for x. *i* is the index for the person, while *j* is the index for the group.

|  |
| --- |
| Correlated and Uncorrelated Random Effects in Logistic Regression |
| The reader is referred to the discussion of correlated and uncorrelated random effects in [Section 6.2](#sec-correlated-uncorrelated) |

### Stata

In Stata mixed, the syntax for a multilevel model of the form described in [Equation 8.1](#eq-MLMsimple-logistic) is:

melogit y x z || group:

### R

In R lme4, the syntax for a multilevel model of the form described in [Equation 8.1](#eq-MLMsimple-logistic) is:

library(lme4)  
  
glmer(y ~ x + z + (1 | group), data = ...)

## 8.3 Run Models

|  |
| --- |
| Less Variation In Logistic Than Linear Models |
| Note that in *logistic* regression models, there is less variation to work with–due to the fact that the outcome is 1/0, than there is in *linear* models. Therefore, in the models below, I do not attempt to estimate a random slope in addition to a random intercept, as I do in [Section 6](#sec-crosssectional). |

### Stata

#### 8.3.0.1 Get The Data

use simulated\_multilevel\_data.dta  
  
generate outcome\_category = outcome > 52 // dichotomous outcome

#### 8.3.0.2 Run The Model

As suggested in [Equation 8.1](#eq-MLMsimple-logistic), odds ratios are obtained by exponentiating the coefficients: . Stata provides the odds ratios automatically with option , or.

melogit outcome\_category warmth physical\_punishment i.identity i.intervention HDI || ///   
country:, or

Fitting fixed-effects model:  
  
Iteration 0: Log likelihood = -1965.6466   
Iteration 1: Log likelihood = -1963.7805   
Iteration 2: Log likelihood = -1963.7791   
Iteration 3: Log likelihood = -1963.7791   
  
Refining starting values:  
  
Grid node 0: Log likelihood = -1908.9697  
  
Fitting full model:  
  
Iteration 0: Log likelihood = -1908.9697 (not concave)  
Iteration 1: Log likelihood = -1903.703   
Iteration 2: Log likelihood = -1902.2851   
Iteration 3: Log likelihood = -1901.3176   
Iteration 4: Log likelihood = -1901.2662   
Iteration 5: Log likelihood = -1901.2661   
  
Mixed-effects logistic regression Number of obs = 3,000  
Group variable: country Number of groups = 30  
  
 Obs per group:  
 min = 100  
 avg = 100.0  
 max = 100  
  
Integration method: mvaghermite Integration pts. = 7  
  
 Wald chi2(5) = 219.75  
Log likelihood = -1901.2661 Prob > chi2 = 0.0000  
-------------------------------------------------------------------------------------  
 outcome\_category | Odds ratio Std. err. z P>|z| [95% conf. interval]  
--------------------+----------------------------------------------------------------  
 warmth | 1.292603 .0278565 11.91 0.000 1.239142 1.34837  
physical\_punishment | .7524276 .0222773 -9.61 0.000 .7100077 .797382  
 1.identity | .9517262 .0748541 -0.63 0.529 .8157636 1.11035  
 1.intervention | 1.191581 .0940459 2.22 0.026 1.020803 1.390929  
 HDI | .9990491 .0061371 -0.15 0.877 .9870928 1.01115  
 \_cons | .9115548 .3901774 -0.22 0.829 .3939478 2.109244  
--------------------+----------------------------------------------------------------  
country |  
 var(\_cons)| .2897697 .0880892 .1596945 .5257944  
-------------------------------------------------------------------------------------  
Note: Estimates are transformed only in the first equation to odds ratios.  
Note: \_cons estimates baseline odds (conditional on zero random effects).  
LR test vs. logistic model: chibar2(01) = 125.03 Prob >= chibar2 = 0.0000

### R

#### 8.3.0.3 Get The Data

library(haven)  
  
df <- read\_dta("simulated\_multilevel\_data.dta")  
  
df$outcome\_category <- 0 # initialize to 0  
  
df$outcome\_category[df$outcome > 52] <- 1 # dichotomous outcome

#### 8.3.0.4 Change Some Variables To Categorical

df$identity <- factor(df$identity)  
  
df$intervention <- factor(df$intervention)

#### 8.3.0.5 Run The Model

|  |
| --- |
| Caution |
| lme4 does not directly provide p values in results, because of some disagreement over exactly how these p values should be calculated. Therefore, in this Appendix, I also call library lmerTest to provide p values for lme4 results. |

|  |
| --- |
| Tip |
| R prefers to use scientific notation when possible. I find that the use of scientific notation can be confusing in reading results. I turn off scientific notation by setting a penalty for its use: options(scipen = 999). |

library(lme4)   
  
library(lmerTest)  
  
options(scipen = 999)   
  
fit3 <- glmer(outcome\_category ~ warmth + physical\_punishment +   
 identity + intervention + HDI +  
 (1 | country),  
 family = binomial(link = "logit"),  
 data = df)  
  
summary(fit3)

Generalized linear mixed model fit by maximum likelihood (Laplace  
 Approximation) [glmerMod]  
 Family: binomial ( logit )  
Formula: outcome\_category ~ warmth + physical\_punishment + identity +   
 intervention + HDI + (1 | country)  
 Data: df  
  
 AIC BIC logLik deviance df.resid   
 3816.6 3858.7 -1901.3 3802.6 2993   
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.0109 -0.8798 0.4369 0.8428 2.8223   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 country (Intercept) 0.2894 0.5379   
Number of obs: 3000, groups: country, 30  
  
Fixed effects:  
 Estimate Std. Error z value Pr(>|z|)   
(Intercept) -0.0926371 0.4277643 -0.217 0.8286   
warmth 0.2566693 0.0215443 11.914 <0.0000000000000002 \*\*\*  
physical\_punishment -0.2844595 0.0295990 -9.610 <0.0000000000000002 \*\*\*  
identity1 -0.0494765 0.0786286 -0.629 0.5292   
intervention1 0.1752879 0.0789030 2.222 0.0263 \*   
HDI -0.0009513 0.0061388 -0.155 0.8769   
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1  
  
Correlation of Fixed Effects:  
 (Intr) warmth physc\_ idntt1 intrv1  
warmth -0.158   
physcl\_pnsh -0.170 -0.082   
identity1 -0.086 -0.014 0.002   
interventn1 -0.102 0.055 0.006 -0.020   
HDI -0.930 -0.007 0.012 -0.001 0.004

#### 8.3.0.6 Calculate Odds Ratios

R requires one to use a bit of extra syntax to extract the odds ratios. As suggested in [Equation 8.1](#eq-MLMsimple-logistic), odds ratios are obtained by exponentiating the coefficients: .

exp(fixef(fit3))

(Intercept) warmth physical\_punishment identity1   
 0.9115242 1.2926176 0.7524208 0.9517275   
 intervention1 HDI   
 1.1915893 0.9990492

# 9. # Models With Three or More Levels and Cross-Classified Models

## 9.1 Introduction

A two level multilevel model imagines that *Level 1* units are nested in *Level 2* units. A three level multilevel model imagines that *Level 1* units are nested in *Level 2* units, which are in turn nested in *Level 3*. As more levels are added to the model (e.g. *Level 4*), we imagine all of these levels to be hierarchically nested.

A *cross classified* model imagines that the nesting is not hierarchical, but rather that there are two sets of clusters or nestings in which individuals may be nested.

Below, I describe the use of Stata, R, and Julia to estimate these models.

## 9.2 Three Or More Levels

### 9.2.1 The Data

I use the *longitudinal* data from *Multilevel Thinking* ([Section 1.2](#sec-data)) to which I have added an extra level of *United Nations Region* (Arel-Bundock et al., 2018). This data thus requires a four level model.

### 9.2.2 The Equation

Here we imagine (region), (country) and (family) are hierarchically nested effects.

### 9.2.3 Run The Models

#### Stata

##### 9.2.3.0.1 Get The Data

use "fourlevel.dta", clear

##### 9.2.3.0.2 Unconditional Model

mixed outcome || UNregion: || country: || family:

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -29061.686   
Iteration 1: Log likelihood = -29061.679   
Iteration 2: Log likelihood = -29061.679   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 9,000  
  
 Grouping information  
 -------------------------------------------------------------  
 | No. of Observations per group  
 Group variable | groups Minimum Average Maximum  
 ----------------+--------------------------------------------  
 UNregion | 5 600 1,800.0 3,600  
 country | 30 300 300.0 300  
 family | 3,000 3 3.0 3  
 -------------------------------------------------------------  
  
 Wald chi2(0) = .  
Log likelihood = -29061.679 Prob > chi2 = .  
  
------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
-------------+----------------------------------------------------------------  
 \_cons | 54.05906 .987367 54.75 0.000 52.12385 55.99426  
------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
UNregion: Identity |  
 var(\_cons) | 4.172687 3.187885 .9334852 18.65194  
-----------------------------+------------------------------------------------  
country: Identity |  
 var(\_cons) | 2.849348 .8710225 1.565093 5.187414  
-----------------------------+------------------------------------------------  
family: Identity |  
 var(\_cons) | 11.72403 .57475 10.64997 12.90641  
-----------------------------+------------------------------------------------  
 var(Residual) | 28.23424 .5154842 27.24177 29.26286  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(3) = 1843.44 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

##### 9.2.3.0.3 Conditional Model

mixed outcome t warmth physical\_punishment i.identity i.intervention HDI ///  
|| UNregion: || country: || id:

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -28503.082   
Iteration 1: Log likelihood = -28503.039   
Iteration 2: Log likelihood = -28503.039   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 9,000  
  
 Grouping information  
 -------------------------------------------------------------  
 | No. of Observations per group  
 Group variable | groups Minimum Average Maximum  
 ----------------+--------------------------------------------  
 UNregion | 5 600 1,800.0 3,600  
 country | 30 300 300.0 300  
 id | 3,000 3 3.0 3  
 -------------------------------------------------------------  
  
 Wald chi2(6) = 1209.42  
Log likelihood = -28503.039 Prob > chi2 = 0.0000  
  
-------------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
--------------------+----------------------------------------------------------------  
 t | .9433791 .0658667 14.32 0.000 .8142827 1.072476  
 warmth | .9140704 .0379156 24.11 0.000 .8397571 .9883837  
physical\_punishment | -1.008615 .0497772 -20.26 0.000 -1.106176 -.9110531  
 1.identity | -.1332133 .1516437 -0.88 0.380 -.4304294 .1640028  
 1.intervention | .8589263 .1519619 5.65 0.000 .5610865 1.156766  
 HDI | .0148561 .0196605 0.76 0.450 -.0236777 .0533899  
 \_cons | 50.16426 1.675219 29.94 0.000 46.88089 53.44763  
-------------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
UNregion: Identity |  
 var(\_cons) | 4.722007 3.585939 1.065898 20.91884  
-----------------------------+------------------------------------------------  
country: Identity |  
 var(\_cons) | 2.863495 .8656459 1.583342 5.178668  
-----------------------------+------------------------------------------------  
id: Identity |  
 var(\_cons) | 8.421131 .4711947 7.546445 9.397199  
-----------------------------+------------------------------------------------  
 var(Residual) | 26.02919 .4752587 25.11417 26.97755  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(3) = 1844.00 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

#### R

##### 9.2.3.0.4 Get The Data

library(haven)  
  
df4 <- read\_dta("fourlevel.dta")

##### 9.2.3.0.5 Change Some Variables To Categorical

df4$identity <- factor(df4$identity)  
  
df4$intervention <- factor(df4$intervention)

##### 9.2.3.0.6 Unconditional Model

|  |
| --- |
| Caution |
| lme4 does not directly provide p values in results, because of some disagreement over exactly how these p values should be calculated. Therefore, in this Appendix, I also call library lmerTest to provide p values for lme4 results. |

|  |
| --- |
| Tip |
| R prefers to use scientific notation when possible. I find that the use of scientific notation can be confusing in reading results. I turn off scientific notation by setting a penalty for its use: options(scipen = 999). |

library(lme4)   
  
library(lmerTest)  
  
options(scipen = 999)   
  
fit4A <- lmer(outcome ~ (1 | UNregion/country/id),  
 data = df4)  
  
summary(fit4A)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula: outcome ~ (1 | UNregion/country/id)  
 Data: df4  
  
REML criterion at convergence: 58121.4  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.7850 -0.6064 -0.0047 0.6020 3.4399   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 id:(country:UNregion) (Intercept) 11.724 3.424   
 country:UNregion (Intercept) 2.842 1.686   
 UNregion (Intercept) 5.478 2.340   
 Residual 28.234 5.314   
Number of obs: 9000, groups:   
id:(country:UNregion), 3000; country:UNregion, 30; UNregion, 5  
  
Fixed effects:  
 Estimate Std. Error df t value Pr(>|t|)   
(Intercept) 54.061 1.112 3.777 48.6 0.00000201 \*\*\*  
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

##### 9.2.3.0.7 Conditional Model

fit4B <- lmer(outcome ~ t + warmth + physical\_punishment +   
 identity + intervention + HDI +   
 (1 | UNregion/country/id),  
 data = df4)  
  
summary(fit4B)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula:   
outcome ~ t + warmth + physical\_punishment + identity + intervention +   
 HDI + (1 | UNregion/country/id)  
 Data: df4  
  
REML criterion at convergence: 57026.4  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.6846 -0.6096 -0.0038 0.6138 3.6850   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 id:(country:UNregion) (Intercept) 8.438 2.905   
 country:UNregion (Intercept) 2.979 1.726   
 UNregion (Intercept) 6.178 2.486   
 Residual 26.036 5.103   
Number of obs: 9000, groups:   
id:(country:UNregion), 3000; country:UNregion, 30; UNregion, 5  
  
Fixed effects:  
 Estimate Std. Error df t value  
(Intercept) 50.11857 1.78086 15.79112 28.143  
t 0.94338 0.06588 5998.37756 14.321  
warmth 0.91406 0.03793 4745.28492 24.096  
physical\_punishment -1.00876 0.04980 6483.46337 -20.257  
identity1 -0.13324 0.15173 2969.00938 -0.878  
intervention1 0.85872 0.15205 2971.85430 5.648  
HDI 0.01560 0.02006 24.39852 0.778  
 Pr(>|t|)   
(Intercept) 0.00000000000000641 \*\*\*  
t < 0.0000000000000002 \*\*\*  
warmth < 0.0000000000000002 \*\*\*  
physical\_punishment < 0.0000000000000002 \*\*\*  
identity1 0.380   
intervention1 0.00000001780521096 \*\*\*  
HDI 0.444   
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1  
  
Correlation of Fixed Effects:  
 (Intr) t warmth physc\_ idntt1 intrv1  
t -0.073   
warmth -0.071 -0.002   
physcl\_pnsh -0.073 -0.007 -0.012   
identity1 -0.040 0.000 -0.013 -0.003   
interventn1 -0.045 0.000 0.039 0.019 -0.018   
HDI -0.738 0.000 -0.005 0.005 -0.001 0.001

#### Julia

##### 9.2.3.0.8 Get The Data

using Tables, MixedModels, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
df4 = DataFrame(load("fourlevel.dta"))

##### 9.2.3.0.9 Change Some Variables To Categorical

@transform!(df4, :country = categorical(:country))  
  
@transform!(df4, :UNregion = categorical(:UNregion))  
  
@transform!(df4, :identity = categorical(:identity))  
  
@transform!(df4, :intervention = categorical(:intervention))

##### 9.2.3.0.10 Unconditional Model

m4A = fit(MixedModel, @formula(outcome ~ t + warmth +   
 physical\_punishment +   
 identity + intervention +   
 HDI +  
 (1 | UNregion) +   
 (1 | country) +   
 (1 | id)), df4)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + t + warmth + physical\_punishment + identity + intervention + HDI + (1 | UNregion) + (1 | country) + (1 | id)  
 logLik -2 logLik AIC AICc BIC   
 -28503.0394 57006.0787 57028.0787 57028.1081 57106.2335  
  
Variance components:  
 Column Variance Std.Dev.  
id (Intercept) 8.42110 2.90191  
country (Intercept) 2.86347 1.69218  
UNregion (Intercept) 4.72082 2.17274  
Residual 26.02921 5.10188  
 Number of obs: 9000; levels of grouping factors: 3000, 30, 5  
  
 Fixed-effects parameters:  
─────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
─────────────────────────────────────────────────────────────  
(Intercept) 50.1643 1.67514 29.95 <1e-99  
t 0.943379 0.0658668 14.32 <1e-45  
warmth 0.91407 0.0379156 24.11 <1e-99  
physical\_punishment -1.00861 0.0497772 -20.26 <1e-90  
identity -0.133213 0.151644 -0.88 0.3797  
intervention 0.858927 0.151962 5.65 <1e-07  
HDI 0.0148553 0.0196604 0.76 0.4499  
─────────────────────────────────────────────────────────────

##### 9.2.3.0.11 Conditional Model

m4B = fit(MixedModel, @formula(outcome ~ t + warmth +   
 physical\_punishment +   
 identity + intervention +   
 HDI +  
 (1 | UNregion) +   
 (1 | country) +   
 (1 | id)), df4)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + t + warmth + physical\_punishment + identity + intervention + HDI + (1 | UNregion) + (1 | country) + (1 | id)  
 logLik -2 logLik AIC AICc BIC   
 -28503.0394 57006.0787 57028.0787 57028.1081 57106.2335  
  
Variance components:  
 Column Variance Std.Dev.  
id (Intercept) 8.42110 2.90191  
country (Intercept) 2.86347 1.69218  
UNregion (Intercept) 4.72082 2.17274  
Residual 26.02921 5.10188  
 Number of obs: 9000; levels of grouping factors: 3000, 30, 5  
  
 Fixed-effects parameters:  
─────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
─────────────────────────────────────────────────────────────  
(Intercept) 50.1643 1.67514 29.95 <1e-99  
t 0.943379 0.0658668 14.32 <1e-45  
warmth 0.91407 0.0379156 24.11 <1e-99  
physical\_punishment -1.00861 0.0497772 -20.26 <1e-90  
identity -0.133213 0.151644 -0.88 0.3797  
intervention 0.858927 0.151962 5.65 <1e-07  
HDI 0.0148553 0.0196604 0.76 0.4499  
─────────────────────────────────────────────────────────────

### 9.2.4 Interpretation

There is group level variation attributable to individual, country, and region.

As in other models, parental warmth, and participation in the intervention are associated with increases in the outcome. Parental use of physical punishment is associated with decreases in the outcome.

## 9.3 Cross-Classified Models

### 9.3.1 The Data

I use the *cross-sectional* data from *Multilevel Thinking* ([Section 1.2](#sec-data)) to which I have added an extra level of a hypothetical language.

### 9.3.2 The Equation

Here (country) and (language) are not nested hierarchically, but are *cross classified*.

### 9.3.3 Run The Models

#### Stata

##### 9.3.3.0.1 Get The Data

use "crossclassified.dta", clear

##### 9.3.3.0.2 Unconditional Model

mixed outcome || \_all: R.country || \_all: R.language

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -9835.8123   
Iteration 1: Log likelihood = -9835.8111   
Iteration 2: Log likelihood = -9835.8111   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 3,000  
Group variable: \_all Number of groups = 1  
 Obs per group:  
 min = 3,000  
 avg = 3,000.0  
 max = 3,000  
 Wald chi2(0) = .  
Log likelihood = -9835.8111 Prob > chi2 = .  
  
---------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
----------------+----------------------------------------------------------------  
 \_cons | 52.43187 .3590214 146.04 0.000 51.7282 53.13554  
---------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
\_all: Identity |  
 var(R.country) | 3.177791 .9244633 1.796798 5.620198  
-----------------------------+------------------------------------------------  
\_all: Identity |  
 var(R.language) | .9566314 .3284087 .4881235 1.87482  
-----------------------------+------------------------------------------------  
 var(Residual) | 39.62877 1.045619 37.63148 41.73206  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(2) = 180.84 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

##### 9.3.3.0.3 Conditional Model

mixed outcome warmth physical\_punishment i.identity i.intervention HDI || \_all: R.country || \_all: R.language

Performing EM optimization ...  
  
Performing gradient-based optimization:   
Iteration 0: Log likelihood = -9663.2195   
Iteration 1: Log likelihood = -9663.2194   
  
Computing standard errors ...  
  
Mixed-effects ML regression Number of obs = 3,000  
Group variable: \_all Number of groups = 1  
 Obs per group:  
 min = 3,000  
 avg = 3,000.0  
 max = 3,000  
 Wald chi2(5) = 367.04  
Log likelihood = -9663.2194 Prob > chi2 = 0.0000  
  
-------------------------------------------------------------------------------------  
 outcome | Coefficient Std. err. z P>|z| [95% conf. interval]  
--------------------+----------------------------------------------------------------  
 warmth | .8331461 .0579811 14.37 0.000 .7195052 .946787  
physical\_punishment | -.9979749 .080268 -12.43 0.000 -1.155297 -.8406525  
 1.identity | -.2922428 .2191421 -1.33 0.182 -.7217534 .1372678  
 1.intervention | .6097458 .2195139 2.78 0.005 .1795064 1.039985  
 HDI | -.0015879 .0204157 -0.08 0.938 -.0416021 .0384262  
 \_cons | 51.92255 1.411069 36.80 0.000 49.15691 54.6882  
-------------------------------------------------------------------------------------  
  
------------------------------------------------------------------------------  
 Random-effects parameters | Estimate Std. err. [95% conf. interval]  
-----------------------------+------------------------------------------------  
\_all: Identity |  
 var(R.country) | 3.361218 .9603072 1.920024 5.884192  
-----------------------------+------------------------------------------------  
\_all: Identity |  
 var(R.language) | 1.121946 .3269535 .6337502 1.986214  
-----------------------------+------------------------------------------------  
 var(Residual) | 35.11959 .9263999 33.35002 36.98306  
------------------------------------------------------------------------------  
LR test vs. linear model: chi2(2) = 227.02 Prob > chi2 = 0.0000  
  
Note: LR test is conservative and provided only for reference.

#### R

##### 9.3.3.0.4 Get The Data

library(haven)  
  
dfCC <- read\_dta("crossclassified.dta")

##### 9.3.3.0.5 Change Some Variables To Categorical

dfCC$identity <- factor(dfCC$identity)  
  
dfCC$intervention <- factor(dfCC$intervention)

##### 9.3.3.0.6 Unconditional Model

library(lme4)   
  
library(lmerTest)  
  
options(scipen = 999)   
  
fitCC\_A <- lmer(outcome ~   
 (1 | country) +  
 (1 | language),  
 data = dfCC)  
  
summary(fitCC\_A)

Linear mixed model fit by REML. t-tests use Satterthwaite's method [  
lmerModLmerTest]  
Formula: outcome ~ (1 | country) + (1 | language)  
 Data: dfCC  
  
REML criterion at convergence: 19671.8  
  
Scaled residuals:   
 Min 1Q Median 3Q Max   
-3.3899 -0.6602 -0.0104 0.6798 3.6924   
  
Random effects:  
 Groups Name Variance Std.Dev.  
 language (Intercept) 0.9604 0.980   
 country (Intercept) 3.2919 1.814   
 Residual 39.6276 6.295   
Number of obs: 3000, groups: language, 100; country, 30  
  
Fixed effects:  
 Estimate Std. Error df t value Pr(>|t|)   
(Intercept) 52.4319 0.3643 33.4284 143.9 <0.0000000000000002 \*\*\*  
---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

##### 9.3.3.0.7 Conditional Model

fitCC\_B <- lmer(outcome ~ t + warmth + physical\_punishment +   
 identity + intervention + HDI +   
 (1 | country) +  
 (1 | language),  
 data = dfCC)

Error in model.frame.default(data = dfCC, drop.unused.levels = TRUE, formula = outcome ~ : invalid type (closure) for variable 't'

summary(fitCC\_B)

Error in h(simpleError(msg, call)): error in evaluating the argument 'object' in selecting a method for function 'summary': object 'fitCC\_B' not found

#### Julia

##### 9.3.3.0.8 Get The Data

using Tables, MixedModels, StatFiles, DataFrames, CategoricalArrays, DataFramesMeta  
  
dfCC = DataFrame(load("crossclassified.dta"))

##### 9.3.3.0.9 Change Some Variables To Categorical

@transform!(dfCC, :country = categorical(:country))  
  
@transform!(dfCC, :language = categorical(:language))  
  
@transform!(dfCC, :identity = categorical(:identity))  
  
@transform!(dfCC, :intervention = categorical(:intervention))

##### 9.3.3.0.10 Unconditional Model

mCCA = fit(MixedModel, @formula(outcome ~   
 (1 | country) +   
 (1 | language)), dfCC)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + (1 | country) + (1 | language)  
 logLik -2 logLik AIC AICc BIC   
 -9835.8111 19671.6222 19679.6222 19679.6356 19703.6477  
  
Variance components:  
 Column Variance Std.Dev.   
language (Intercept) 0.956631 0.978075  
country (Intercept) 3.177768 1.782629  
Residual 39.628773 6.295139  
 Number of obs: 3000; levels of grouping factors: 100, 30  
  
 Fixed-effects parameters:  
──────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
──────────────────────────────────────────────────  
(Intercept) 52.4319 0.35902 146.04 <1e-99  
──────────────────────────────────────────────────

##### 9.3.3.0.11 Conditional Model

mCCA = fit(MixedModel, @formula(outcome ~ warmth +   
 physical\_punishment +   
 identity + intervention +   
 HDI +  
 (1 | country) +   
 (1 | language)), dfCC)

Linear mixed model fit by maximum likelihood  
 outcome ~ 1 + warmth + physical\_punishment + identity + intervention + HDI + (1 | country) + (1 | language)  
 logLik -2 logLik AIC AICc BIC   
 -9663.2194 19326.4388 19344.4388 19344.4990 19398.4962  
  
Variance components:  
 Column Variance Std.Dev.  
language (Intercept) 1.12193 1.05921  
country (Intercept) 3.36119 1.83335  
Residual 35.11960 5.92618  
 Number of obs: 3000; levels of grouping factors: 100, 30  
  
 Fixed-effects parameters:  
──────────────────────────────────────────────────────────────  
 Coef. Std. Error z Pr(>|z|)  
──────────────────────────────────────────────────────────────  
(Intercept) 51.9226 1.41106 36.80 <1e-99  
warmth 0.833146 0.0579811 14.37 <1e-46  
physical\_punishment -0.997975 0.080268 -12.43 <1e-34  
identity -0.292243 0.219142 -1.33 0.1823  
intervention 0.609746 0.219514 2.78 0.0055  
HDI -0.00158794 0.0204156 -0.08 0.9380  
──────────────────────────────────────────────────────────────

### 9.3.4 Interpretation

There is group level variation attributable to both language and country.

As in other models, parental warmth, and participation in the intervention are associated with increases in the outcome. Parental use of physical punishment is associated with decreases in the outcome.

# 10. Reshaping Data

## 10.1 Introduction

Cross-sectional analyses ([Section 6](#sec-crosssectional)) make use of data in *wide* format: every row is a person, or family, and every person, or family, has a single row of data.

Longitudinal analyses ([Section 7](#sec-longitudinal)) make use of *long* data: every row is a person-timepoint, or family-timepoint, and every person, or family, has multiple rows of data.

Data therefore sometimes need to be *reshaped*, most often from *wide* format to *long* format. Almost any software that is capable of estimating multilevel models is capable of reshaping data.

Below, I detail the procedure for reshaping data in Stata and R.

## 10.2 Data in Wide Format

|  |
| --- |
| Note |
| The data below are in *wide* format.  Every individual in the data set has a *single row of data*. Every row in the data set is an *individual* or *family*. |

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Table 10.1: Data in Wide Format  Table continues below   | id | physical\_punishment1 | warmth1 | outcome1 | physical\_punishment2 | | --- | --- | --- | --- | --- | | 1.1 | 3 | 3 | 57.47 | 3 | | 1.10 | 2 | 0 | 62.9 | 3 | | 1.100 | 2 | 5 | 62.71 | 1 | | 1.11 | 4 | 4 | 55.61 | 2 | | 1.12 | 5 | 4 | 41.15 | 5 | | 1.13 | 4 | 5 | 63.66 | 3 |     Table continues below   | warmth2 | outcome2 | physical\_punishment3 | warmth3 | outcome3 | country | HDI | | --- | --- | --- | --- | --- | --- | --- | | 4 | 55.06 | 1 | 2 | 58.77 | 1 | 69 | | 0 | 56.67 | 2 | 0 | 68.22 | 1 | 69 | | 4 | 51.58 | 2 | 5 | 55.51 | 1 | 69 | | 5 | 50.9 | 3 | 3 | 48.02 | 1 | 69 | | 5 | 45.4 | 3 | 4 | 55.86 | 1 | 69 | | 3 | 64.81 | 3 | 3 | 58.3 | 1 | 69 |      | family | identity | intervention | | --- | --- | --- | | 1 | 1 | 0 | | 10 | 1 | 0 | | 100 | 1 | 1 | | 11 | 1 | 1 | | 12 | 0 | 0 | | 13 | 0 | 1 | |

## 10.3 Data Management

Because reshaping your data dramatically changes the structure of your data…

1. It is a good idea to have your raw data saved in a location where it will not be changed, and can be retrieved again if the reshape command does not work correctly, or if you simply want to modify your reshaping data workflow.(CF [Section 2.3](#sec-script-flow))
2. Usually we want to work with only a *subset* of your data, to keep only the data in which we are interested.
   * In Stata, the command to keep only variables of interest would be: keep y x z t.
   * In R, one option would be to use the subset function: mysubset <- subset(mydata, select = c(y, x, z, t))

## 10.4 Reshaping Data From Wide To Long

Usually, we are most interested in reshaping data from *wide* to *long*.

### Stata

In Stata, I only list variables that vary over time, or are *time varying*. Stata assumes that variables that are *not listed* do *not vary over time*, or are *time invariant*.

Notice also that our *time varying* data are in the *stub-time* format, e.g. warmth1, warmth2, physical\_punishment1 physical\_punishment2, etc. Because the variables are named in this way, Stata knows to use the *stub* (e.g. warmth) as the variable name, and the numeric value, (e.g. 1, 2, 3) as the timepoint.

The id variable, whatever it is named, has to uniquely identify the observations. A useful Stata command here is isid, e.g. isid id. If your id variable is not unique, it is often due to missing values. drop if id == . usually solves the problem (assuming that your id variable is indeed named id, and not something else).

use simulated\_multilevel\_longitudinal\_data\_WIDE.dta, clear  
  
describe  
  
reshape long outcome physical\_punishment warmth, i(id) j(wave)

Contains data from simulated\_multilevel\_longitudinal\_data\_WIDE.dta  
 Observations: 3,000   
 Variables: 15 3 Jul 2024 14:29  
-------------------------------------------------------------------------------------------  
Variable Storage Display Value  
 name type format label Variable label  
-------------------------------------------------------------------------------------------  
id str7 %9s unique country family id  
physical\_puni~1 float %9.0g 1 physical\_punishment  
warmth1 float %9.0g 1 warmth  
outcome1 float %9.0g 1 outcome  
physical\_puni~2 float %9.0g 2 physical\_punishment  
warmth2 float %9.0g 2 warmth  
outcome2 float %9.0g 2 outcome  
physical\_puni~3 float %9.0g 3 physical\_punishment  
warmth3 float %9.0g 3 warmth  
outcome3 float %9.0g 3 outcome  
country float %9.0g country id  
HDI float %9.0g Human Development Index  
family float %9.0g family id  
identity float %9.0g hypothetical identity group variable  
intervention float %9.0g recieved intervention  
-------------------------------------------------------------------------------------------  
Sorted by: id  
  
(j = 1 2 3)  
  
Data Wide -> Long  
-----------------------------------------------------------------------------  
Number of observations 3,000 -> 9,000   
Number of variables 15 -> 10   
j variable (3 values) -> wave  
xij variables:  
 outcome1 outcome2 outcome3 -> outcome  
physical\_punishment1 physical\_punishment2 physical\_punishment3->physical\_punishment  
 warmth1 warmth2 warmth3 -> warmth  
-----------------------------------------------------------------------------

### R

In R, I only list variables that vary over time, or are *time varying*.

Notice also that our *time varying* data are in the *stub-time* format, e.g. warmth1, warmth2, physical\_punishment1 physical\_punishment2, etc. In order to facilitate reshaping the data, it is helpful in R to insert an underscore (\_) to separate the *stub* from the *time* variable.

library(dplyr) # data wrangling  
  
library(tidyr) # tidy (reshape data)

# rename variables with "\_" separator  
  
df <- simulated\_multilevel\_longitudinal\_data\_WIDE %>%  
 rename(outcome\_1 = outcome1,  
 outcome\_2 = outcome2,  
 outcome\_3 = outcome3,  
 physical\_punishment\_1 = physical\_punishment1,  
 physical\_punishment\_2 = physical\_punishment2,  
 physical\_punishment\_3 = physical\_punishment3,  
 warmth\_1 = warmth1,  
 warmth\_2 = warmth2,  
 warmth\_3 = warmth3)

# pivot\_longer() to long data  
  
dfL <- df %>%  
 pivot\_longer(cols = c(outcome\_1,  
 outcome\_2,  
 outcome\_3,  
 physical\_punishment\_1,  
 physical\_punishment\_2,  
 physical\_punishment\_3,  
 warmth\_1,  
 warmth\_2,  
 warmth\_3),   
 names\_pattern = "(.+)\_(.+)",  
 names\_to = c(".value", "t"))

## 10.5 Data in Long Format

|  |
| --- |
| Note |
| The data below are in *long* format.  Every individual/family in the data set has a *multiple rows of data*. Every row in the data set is an *individual-timepoint* or *family-timepoint*. |

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Table 10.2: Data in Long Format  Table continues below   | country | HDI | family | id | identity | intervention | t | | --- | --- | --- | --- | --- | --- | --- | | 1 | 69 | 1 | 1.1 | 1 | 0 | 1 | | 1 | 69 | 1 | 1.1 | 1 | 0 | 2 | | 1 | 69 | 1 | 1.1 | 1 | 0 | 3 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 1 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 2 | | 1 | 69 | 2 | 1.2 | 1 | 1 | 3 |      | physical\_punishment | warmth | outcome | | --- | --- | --- | | 3 | 3 | 57.47 | | 3 | 4 | 55.06 | | 1 | 2 | 58.77 | | 2 | 1 | 50.1 | | 3 | 0 | 53.31 | | 3 | 1 | 49.79 | |

# 11. Aggregating Data

In many instances, we may wish to aggregate data. For example, we may wish to create *contextual variables* representing the average level of an indicator across a group. In the examples I am using in this book, the group under consideration is the country. Aggregating data is also an important part of discussions of *within* and *between* variation, and is an important part of the correlated random effects model.

In the examples below, I create a group level variable for warmth, representing the average level of parental warmth in each country. If warmth is denoted by then the country level variable is denoted by .

Below, I detail the procedure for aggregating data in Stata and R.

### Stata

#### 11.0.0.1 Get The Data

use simulated\_multilevel\_data.dta

#### 11.0.0.2 Create A Group Level Variable

bysort country: egen mean\_warmth = mean(warmth)

### R

#### 11.0.0.3 Get The Data

library(haven)  
  
df <- read\_dta("simulated\_multilevel\_data.dta")

#### 11.0.0.4 Create A Group Level Variable

library(dplyr)   
  
df <- df %>%   
 group\_by(country) %>%  
 mutate(mean\_warmth = mean(warmth))

# References

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1. skimr is an excellent new alternative library for generating descriptive statistics in R. [↑](#footnote-ref-94)