

Causal Modeling With GSS Data Using Multiple Approaches

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4 Dec 2020 09:47:56

Research Question

What is the *possibly causal* association of *education* with *job satisfaction*, while accounting for factors that may possibly have an association with *level of education*?

Causality

A variable x can only be considered to have *causal* association with y if the following conditions are met (Holland, 1986):

1. x is correlated with y .
2. x precedes y in time order.
3. The association between x and y can not be accounted for by any third variable z .

Hence, for this particular data, we are exploring:

What happens to the association of *education* and *job satisfaction* when we control for possible confounding variables z using various statistical strategies?

To Be Added To Each Analysis

- Assumptions
- Equation
- Stata Command
- Conclusion

Setup

```
. clear all

. cd "/Users/agrogan/Desktop/newstuff/causal-modeling"
/Users/agrogan/Desktop/newstuff/causal-modeling
```

Get Data

```
. use "GSS_panel2010w123_R6 - stata.dta", clear
```

ID Variable

```
. generate ID = id_1
```

Keep Only Relevant Variables

```
. keep ID satjob_? educ_? race_? incom16_?
```

Describe Data

```
. describe
Contains data from GSS_panel2010w123_R6 - stata.dta
  obs:      2,044
  vars:      13
  size:     32,704
                    5 Jul 2020 13:27
```

variable name	storage type	display format	value label	variable label
educ_1	byte	%8.0g	EDUC_1	educ_1: HIGHEST YEAR OF SCHOOL COMPLETED
educ_2	byte	%8.0g	EDUC_2	educ_2: HIGHEST YEAR OF SCHOOL COMPLETED
educ_3	byte	%8.0g	EDUC_3	educ_3: HIGHEST YEAR OF SCHOOL COMPLETED
incom16_1	byte	%8.0g	INCOM16	incom16_1: RS FAMILY INCOME WHEN 16 YRS OLD
incom16_2	byte	%8.0g	V1318_A	incom16_2: RS FAMILY INCOME WHEN 16 YRS OLD
incom16_3	byte	%8.0g	V1319_A	incom16_3: RS FAMILY INCOME WHEN 16 YRS OLD
race_1	byte	%8.0g	RACE_1	race_1: RACE OF RESPONDENT
race_2	byte	%8.0g	RACE_2	race_2: RACE OF RESPONDENT
race_3	byte	%8.0g	RACE_3	race_3: RACE OF RESPONDENT
satjob_1	byte	%8.0g	SATJOB_1	satjob_1: JOB OR HOUSEWORK
satjob_2	byte	%8.0g	SATJOB_2	satjob_2: JOB OR HOUSEWORK
satjob_3	byte	%8.0g	SATJOB_3	satjob_3: JOB OR HOUSEWORK
ID	float	%9.0g		

```
Sorted by:
Note: Dataset has changed since last saved.
```

Codebook For Selected Variable(s)

```
. codebook satjob_3
```

satjob_3			satjob_3: JOB OR HOUSEWORK		
type:		numeric (byte)			
label:		SATJOB_3			
range:		[1,4]	units:	1	
unique values:		4	missing .:	0/2,044	
unique mv codes:		3	missing .*:	1,086/2,044	
tabulation:	Freq.	Numeric	Label		
	483	1	VERY SATISFIED		
	367	2	MOD. SATISFIED		
	69	3	A LITTLE DISSAT		
	39	4	VERY DISSATISFIED		
	4	.d	DK		
	1,073	.i	IAP		
	9	.n	NA		

Analyses Relying On Wide Data

Correlation

```
. pwcorr satjob_3 educ_3, sig
```

	satjob_3	educ_3
satjob_3	1.0000	
educ_3	-0.0774 0.0166	1.0000

Regression With 1 Independent Variable

```
. regress satjob_3 educ_3
```

Source	SS	df	MS	Number of obs	=	957
Model	3.53828635	1	3.53828635	F(1, 955)	=	5.76
Residual	586.493062	955	.61412886	Prob > F	=	0.0166
				R-squared	=	0.0060
				Adj R-squared	=	0.0050
Total	590.031348	956	.617187602	Root MSE	=	.78366

satjob_3	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ_3	-.0216864	.0090349	-2.40	0.017	-.0394169	-.003956
_cons	1.954439	.1297867	15.06	0.000	1.699739	2.209139

Regression With Multiple Independent Variables

```
. regress satjob_3 educ_3 i.race_3 incom16_3
```

Source	SS	df	MS	Number of obs	=	951
Model	5.81703392	4	1.45425848	F(4, 946)	=	2.36
Residual	582.580442	946	.615835563	Prob > F	=	0.0517
				R-squared	=	0.0099
				Adj R-squared	=	0.0057
Total	588.397476	950	.619365765	Root MSE	=	.78475

satjob_3	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ_3	-.0215151	.0092674	-2.32	0.020	-.0397021	-.0033281
race_3						
black	.1267666	.0708898	1.79	0.074	-.0123528	.2658861
other	.0677238	.0985112	0.69	0.492	-.1256019	.2610495
incom16_3						
_cons	.0115275	.0280601	0.41	0.681	-.0435398	.0665947
	1.89556	.144649	13.10	0.000	1.61169	2.17943

Propensity Score

Data Wrangling Since Propensity Score Requires a Binary Treatment Variable

```
. generate twelve_years_3 = educ_3 >= 12 // 12 or more years of education
. generate twelve_years_2 = educ_2 >= 12 // 12 or more years of education
```

```
. generate twelve_years_1 = educ_1 >= 12 // 12 or more years of education

. label variable twelve_years_3 "12 or more years of education"

. label variable twelve_years_2 "12 or more years of education"

. label variable twelve_years_1 "12 or more years of education"
```

Propensity Score Analysis

```
. teffects psmatch (satjob_3) (twelve_years_3 incom16_3 i.race_3)
Treatment-effects estimation      Number of obs      =      952
Estimator      : propensity-score matching      Matches: requested =      1
Outcome model  : matching                      min =      1
Treatment model: logit                      max =      296
```

	Coef.	AI Robust Std. Err.	z	P> z	[95% Conf. Interval]	
ATE						
twelve_years_3 (1 vs 0)	-.0410168	.1083808	-0.38	0.705	-.2534393	.1714057

Assess Balance of Propensity Score Model ¹

```
. tebalance summarize
note: refitting the model using the generate() option
Covariate balance summary
```

	Raw	Matched
Number of obs =	952	1,904
Treated obs =	854	952
Control obs =	98	952

	Standardized differences		Variance ratio	
	Raw	Matched	Raw	Matched
incom16_3	.5429864	-.0077616	.9418824	.9726307
race_3				
black	-.1354119	-.0199848	.7873145	.9638265
other	-.0248378	.0326166	.9163586	1.114865

```
. tebalance density, scheme(michigan)
note: refitting the model using the generate() option

. graph export mydensity.png, width(500) replace
(file mydensity.png written in PNG format)
```

Cross Lagged Regression

```
. sem (satjob_3 <- twelve_years_2 incom16_2) ///
> (twelve_years_3 <- satjob_2 incom16_2), ///
> cov(e.satjob_3*e.twelve_years_3)
(1164 observations with missing values excluded)

Endogenous variables
Observed: satjob_3 twelve_years_3
```

¹With many thanks to Jorge Cuartas for the ideas for earlier versions of this code.

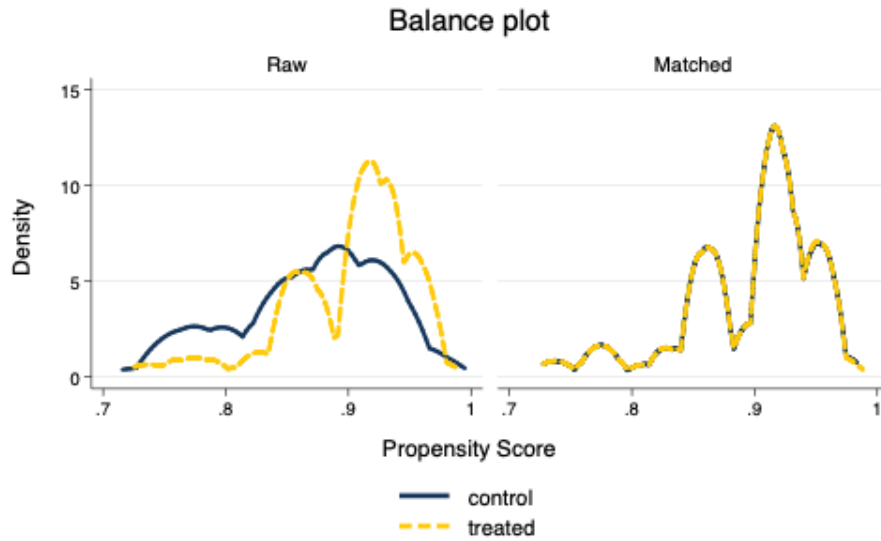


Figure 1: Density Plot of Propensity Score

Exogenous variables

Observed: twelve_years_2 incom16_2 satjob_2

Fitting target model:

Iteration 0: log likelihood = -3673.669

Iteration 1: log likelihood = -3673.4349

Iteration 2: log likelihood = -3673.4342

Iteration 3: log likelihood = -3673.4342

Structural equation model

Number of obs = 880

Estimation method = ml

Log likelihood = -3673.4342

		OIM		z	P> z	[95% Conf. Interval]	
		Coef.	Std. Err.				
Structural satjob_3	twelve_years_2	-.0722619	.1575751	-0.46	0.647	-.3811034	.2365795
	incom16_2	-.0024625	.0304497	-0.08	0.936	-.0621428	.0572177
	_cons	1.734644	.1378498	12.58	0.000	1.464464	2.004825
twelve_years_3	incom16_2	.0607781	.0108369	5.61	0.000	.0395381	.0820181
	satjob_2	-.0054737	.0149476	-0.37	0.714	-.0347704	.0238231
	_cons	.7396908	.0427695	17.29	0.000	.6558642	.8235175
var(e.satjob_3)		.6046871	.0288304			.5507404	.6639181
var(e.twelve_years_3)		.0868092	.0041385			.0790653	.0953116
cov(e.satjob_3,e.twelve_years_3)		.0016217	.0154575	0.10	0.916	-.0286744	.0319178

LR test of model vs. saturated: chi2(2) = 968.48, Prob > chi2 = 0.0000

Analyses Relying On Long Data

Reshape The Data

```
. reshape long satjob_ educ_ twelve_years_ incom16_ race_, i(ID) j(wave)
(note: j = 1 2 3)
```

Data	wide	->	long
Number of obs.	2044	->	6132
Number of variables	16	->	7
j variable (3 values)		->	wave
xij variables:			
satjob_1 satjob_2 satjob_3		->	satjob_
educ_1 educ_2 educ_3		->	educ_
twelve_years_1 twelve_years_2 twelve_years_3		->	twelve_years_
incom16_1 incom16_2 incom16_3		->	incom16_
race_1 race_2 race_3		->	race_

Clean Up Variable Names

```
. rename satjob_ satjob
. rename educ_ educ
. rename incom16_ incom16
. rename race_ race
. rename twelve_years_ twelve_years
```

Multilevel Model

```
. mixed satjob wave educ incom16 i.race || ID:
```

Performing EM optimization:

Performing gradient-based optimization:

Iteration 0: log likelihood = -4161.775

Iteration 1: log likelihood = -4161.7476

Iteration 2: log likelihood = -4161.7476

Computing standard errors:

Mixed-effects ML regression

Group variable: ID

Number of obs = 3,595

Number of groups = 1,661

Obs per group:

min = 1

avg = 2.2

max = 3

Wald chi2(5) = 42.38

Prob > chi2 = 0.0000

Log likelihood = -4161.7476

satjob	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
wave	-.018625	.014015	-1.33	0.184	-.0460938	.0088439
educ	-.018976	.0054133	-3.51	0.000	-.0295859	-.008366
incom16	-.0350535	.0154559	-2.27	0.023	-.0653465	-.0047606
race						
black	.1695589	.0451171	3.76	0.000	.0811311	.2579868
other	.035975	.0543135	0.66	0.508	-.0704776	.1424276
_cons	2.049073	.0843019	24.31	0.000	1.883845	2.214302

Random-effects Parameters		Estimate	Std. Err.	[95% Conf. Interval]	
ID: Identity					
	var(_cons)	.2305185	.0161162	.2009999	.2643722
	var(Residual)	.4174209	.0131143	.3924927	.4439323
LR test vs. linear model: chibar2(01) = 322.95 Prob >= chibar2 = 0.0000					

Fixed effects regression

```
. xtreg satjob wave educ incom16 i.race, i(ID) fe
Fixed-effects (within) regression      Number of obs   =       3,595
Group variable: ID                    Number of groups =       1,661
R-sq:                                Obs per group:
    within = 0.0052                      min =           1
    between = 0.0148                     avg =           2.2
    overall = 0.0122                      max =           3
                                         F(5,1929)       =       2.03
corr(u_i, Xb) = -0.0714                 Prob > F         =       0.0711
```

satjob	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
wave	-.0237842	.0152551	-1.56	0.119	-.0537023	.006134
educ	-.0087664	.0158008	-0.55	0.579	-.0397548	.022222
incom16	-.047186	.0228265	-2.07	0.039	-.0919531	-.0024189
race						
black	.3226033	.2025604	1.59	0.111	-.0746572	.7198637
other	.0383663	.104807	0.37	0.714	-.1671806	.2439132
_cons	1.928458	.227991	8.46	0.000	1.481323	2.375593
sigma_u	.6861769					
sigma_e	.64822634					
rho	.52841711	(fraction of variance due to u_i)				

F test that all u_i=0: F(1660, 1929) = 2.18 Prob > F = 0.0000

“Hybrid” Model

The contention here is that the *between person* coefficient replicates the effect of the fixed effects regression coefficient while the *within person* coefficient is simultaneously estimated.

Generate Within And Between Variables

```
. bysort ID: egen educ_mean = mean(educ)
(6 missing values generated)

. generate educ_deviation = educ - educ_mean
(1,240 missing values generated)
```

Estimate The Model

```
. mixed satjob wave educ_mean educ_deviation incom16 i.race || ID:
Performing EM optimization:
Performing gradient-based optimization:
Iteration 0:   log likelihood = -4161.3224
```

```

Iteration 1:  log likelihood = -4161.2951
Iteration 2:  log likelihood = -4161.2951
Computing standard errors:
Mixed-effects ML regression      Number of obs      =      3,595
Group variable: ID              Number of groups   =      1,661
                                Obs per group:
                                min =          1
                                avg =          2.2
                                max =          3
                                Wald chi2(6)      =      43.30
                                Prob > chi2       =      0.0000
Log likelihood = -4161.2951

```

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
satjob						
wave	-.0197009	.0140588	-1.40	0.161	-.0472556	.0078537
educ_mean	-.0208983	.0057775	-3.62	0.000	-.0322221	-.0095745
educ_deviation	-.0054971	.0151667	-0.36	0.717	-.0352233	.0242292
incom16	-.0343579	.0154712	-2.22	0.026	-.0646809	-.0040349
race						
black	.1684699	.0451261	3.73	0.000	.0800245	.2569154
other	.0342568	.0543368	0.63	0.528	-.0722414	.140755
_cons	2.075849	.088866	23.36	0.000	1.901675	2.250023

Random-effects Parameters		Estimate	Std. Err.	[95% Conf. Interval]	
ID: Identity					
	var(_cons)	.2304651	.0161097	.2009581	.2643046
	var(Residual)	.4173132	.0131099	.3923934	.4438157

LR test vs. linear model: chibar2(01) = 323.08 Prob >= chibar2 = 0.0000

Difference In Difference Model

Combinations of Models

Difference In Difference + Propensity Scores

Cross-Lagged Model With Fixed Effects

References

Holland, P. W. (1986). Statistics and Causal Inference. *Journal of the American Statistical Association*, 81(396), 945–960. <https://doi.org/10.1080/01621459.1986.10478354>