



ACEMS Forecasting Workshop

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1 Forecast Evaluation

Outline

- 1 Introduction
- 2 The statistical forecasting perspective
- 3 Benchmark methods
- 4 Residual diagnostics
- **5** Evaluating point forecast accuracy
- **6** Time series cross-validation
- **7** Prediction intervals
- **8** Evaluating forecast distributions

Resources

robjhyndman.com/acemsforecasting2018

- Slides
- Exercises
- Textbook
- Useful links

Key reference

Hyndman, R. J. & Athanasopoulos, G. (2018) Forecasting: principles and practice, 2nd ed.

OTexts.org/fpp2/

- Free and online
- Data sets in associated R package
- R code for examples

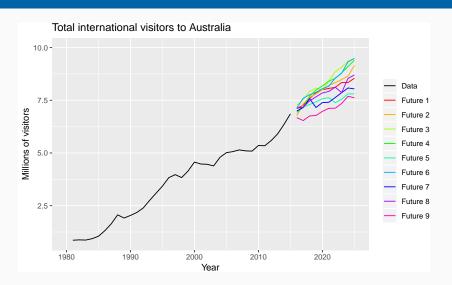
Outline

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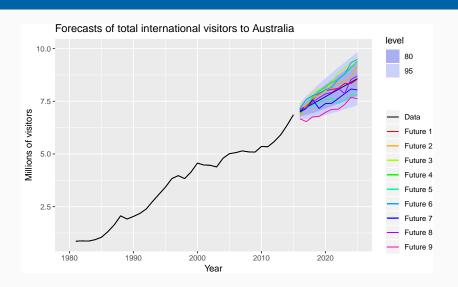
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Sample futures

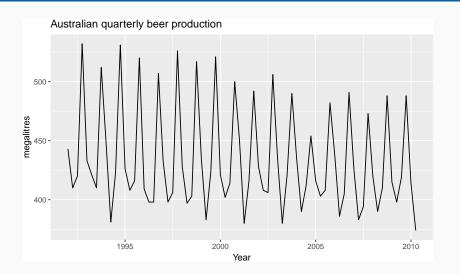


Forecast intervals

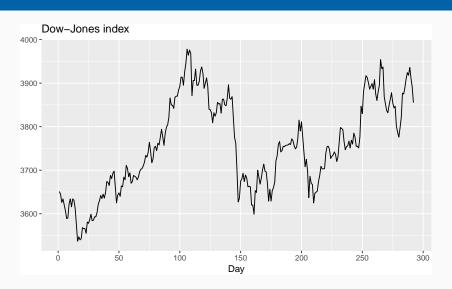


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Average method

- Forecast of all future values is equal to mean of historical data $\{y_1, \dots, y_T\}$.
- Forecasts: $\hat{y}_{T+h|T} = \bar{y} = (y_1 + \cdots + y_T)/T$

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Naïve method

- Forecasts equal to last observed value.
- Forecasts: $\hat{y}_{T+h|T} = y_T$.
- Consequence of efficient market hypothesis.

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Naïve method

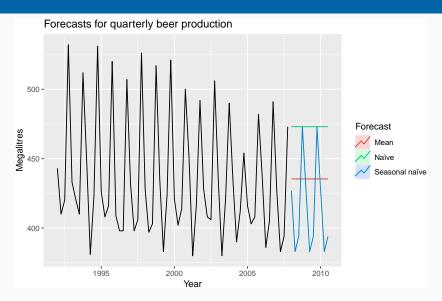
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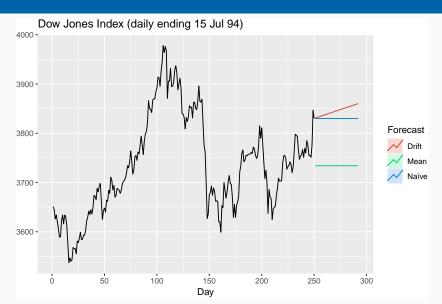
Seasonal naïve method

- Forecasts equal to last value from same season.
- Forecasts: $\hat{y}_{T+h|T} = y_{T+h-km}$ where m = seasonal period and k is integer part of (h-1)/m.

Drift method

- Forecasts equal to last value plus average change.
- Forecasts: $\hat{y}_{T+h|T} = y_T + \frac{h}{T-1} \sum_{t=2}^{T} (y_t y_{t-1})$ = $y_T + \frac{h}{T-1} (y_T - y_1)$.
- Equivalent to extrapolating a line drawn between first and last observations.





- Mean: meanf(y, h=20)
- Naïve: naive(y, h=20)
- Seasonal naïve: snaive(y, h=20)
- Drift: rwf(y, drift=TRUE, h=20)

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Fitted values

- $\hat{y}_{t|t-1}$ is the forecast of y_t based on observations y_1, \dots, y_t .
- We call these "fitted values".
- Sometimes drop the subscript: $\hat{y}_t \equiv \hat{y}_{t|t-1}$.
- Often not true forecasts since parameters are estimated on all data.

For example:

- $\hat{y}_t = \bar{y}$ for average method.
- $\hat{y}_t = y_{t-1} + (y_T y_1)/(T 1)$ for drift method.

Forecasting residuals

Residuals in forecasting: difference between observed value and its fitted value: $e_t = y_t - \hat{y}_{t|t-1}$.

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Assumptions

- $\{e_t\}$ uncorrelated. If they aren't, then information left in residuals that should be used in computing forecasts.
- $\{e_t\}$ have mean zero. If they don't, then forecasts are biased.

Forecasting residuals

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Assumptions

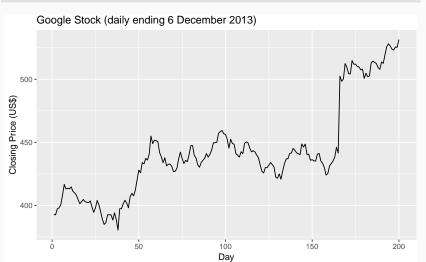
- $\{e_t\}$ uncorrelated. If they aren't, then information left in residuals that should be used in computing forecasts.
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Useful properties (for prediction intervals)

- $\{e_t\}$ have constant variance.
- $\{e_t\}$ are normally distributed.

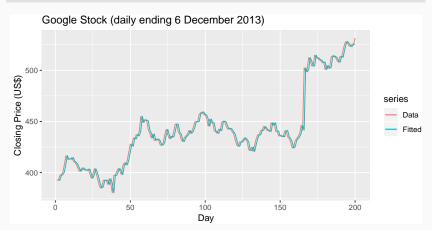
Example: Google stock price

```
autoplot(goog200) +
  xlab("Day") + ylab("Closing Price (US$)") +
  ggtitle("Google Stock (daily ending 6 December 2013)")
```



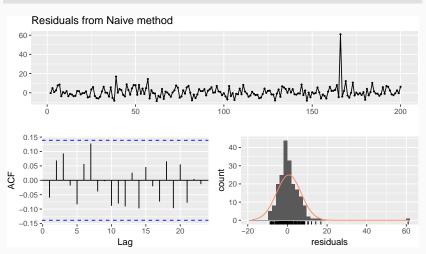
Example: Google stock price

```
fits <- fitted(naive(goog200))
autoplot(goog200, series="Data") +
  autolayer(fits, series="Fitted") +
  xlab("Day") + ylab("Closing Price (US$)") +
  ggtitle("Google Stock (daily ending 6 December 2013)")</pre>
```



checkresiduals function

checkresiduals(naive(goog200), test=FALSE)



ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We *expect* these to look like white noise.

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Training and test sets



- A model which fits the training data well will not necessarily forecast well.
- A perfect fit can always be obtained by using a model with enough parameters.
- Over-fitting a model to data is just as bad as failing to identify a systematic pattern in the data.
- The test set must not be used for *any* aspect of model development or calculation of forecasts.
- Forecast accuracy is based only on the test set.

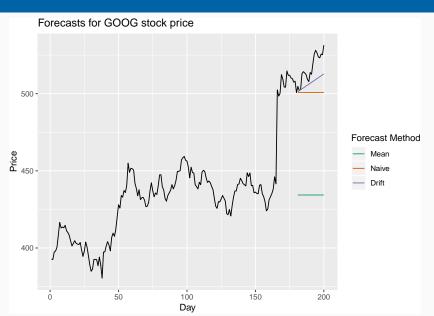
Forecast errors

Forecast "error": the difference between an observed value and its forecast.

$$e_{T+h} = y_{T+h} - \hat{y}_{T+h|T},$$

where the training data is given by $\{y_1, \ldots, y_T\}$

- Unlike residuals, forecast errors on the test set involve multi-step forecasts.
- These are *true* forecast errors as the test data is not used in computing $\hat{y}_{T+h|T}$.



```
y_{T+h} = (T+h)th observation, h = 1, ..., H

\hat{y}_{T+h|T} =  its forecast based on data up to time T.

e_{T+h} = y_{T+h} - \hat{y}_{T+h|T}
```

MAE = mean(
$$|e_{T+h}|$$
)
MSE = mean(e_{T+h}^2) RMSE = $\sqrt{\text{mean}(e_{T+h}^2)}$
MAPE = 100mean($|e_{T+h}|/|y_{T+h}|$)

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```

```
MAE = mean(|e_{T+h}|)

MSE = mean(e_{T+h}^2) RMSE = \sqrt{\text{mean}(e_{T+h}^2)}

MAPE = 100mean(|e_{T+h}|/|y_{T+h}|)
```

- MAE, MSE, RMSE are all scale dependent.
- MAPE is scale independent but is only sensible if $y_t \gg 0$ for all t, and y has a natural zero.

Mean Absolute Scaled Error

MASE =
$$T^{-1} \sum_{t=1}^{T} |y_t - \hat{y}_{t|t-1}|/Q$$

where Q is a stable measure of the scale of the time series $\{y_t\}$.

Proposed by Hyndman and Koehler (IJF, 2006).

For non-seasonal time series,

$$Q = (T - 1)^{-1} \sum_{t=2}^{T} |y_t - y_{t-1}|$$

works well. Then MASE is equivalent to MAE relative to a naïve method.

Mean Absolute Scaled Error

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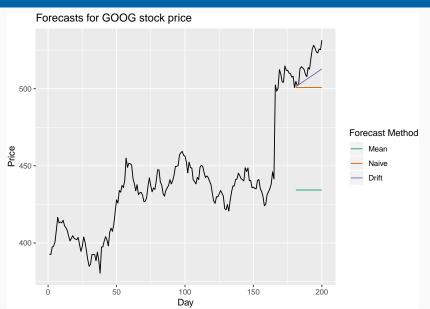
Proposed by Hyndman and Koehler (IJF, 2006).

For seasonal time series,

$$Q = (T - m)^{-1} \sum_{t=1}^{T} |y_t - y_{t-m}|$$

works well. Then MASE is equivalent to MAE relative to a seasonal naïve method.

Measures of forecast accuracy



Measures of forecast accuracy

```
googtrain <- window(goog200,end=180)
googfc1 <- meanf(googtrain,h=20)
googfc2 <- rwf(googtrain,h=20)
googfc3 <- rwf(googtrain,h=20,drift=TRUE)
accuracy(googfc1, goog200)
accuracy(googfc2, goog200)
accuracy(googfc3, goog200)</pre>
```

	RMSE	MAE	MAPE	MASE
Mean method	82.89	82.43	15.93	21.61
Naïve method	18.29	16.04	3.08	4.21
Drift method	11.34	9.71	1.86	2.55

Poll: true or false?

- Good forecast methods should have normally distributed residuals.
- A model with small residuals will give good forecasts.
- The best measure of forecast accuracy is MAPE.
- If your model doesn't forecast well, you should make it more complicated.
- Always choose the model with the best forecast accuracy as measured on the test set.

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Time series cross-validation

Traditional evaluation

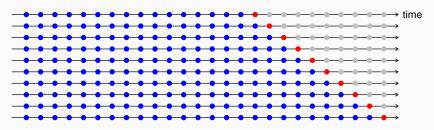


Time series cross-validation

Traditional evaluation



Time series cross-validation

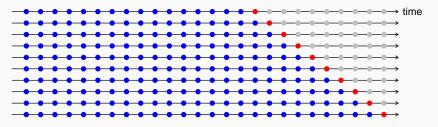


Time series cross-validation

Traditional evaluation



Time series cross-validation



- Forecast accuracy averaged over test sets.
- Also known as "evaluation on a rolling forecasting origin"

tsCV function:

```
goog200 %>%
  tsCV(forecastfunction=meanf, h=1) -> e
e^2 %>% mean(na.rm=TRUE) %>% sqrt
## [1] 37.16
goog200 %>%
  tsCV(forecastfunction=naive, h=1) -> e
e^2 %>% mean(na.rm=TRUE) %>% sqrt
## [1] 6.224
goog200 %>%
  tsCV(forecastfunction=rwf, drift=TRUE, h=1) -> e
e^2 %>% mean(na.rm=TRUE) %>% sgrt
## [1] 6.233
```

A good way to choose the best forecasting model is to find the model with the smallest RMSE computed using time series cross-validation.

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- A forecast $\hat{y}_{T+h|T}$ is (usually) the mean of the conditional distribution $y_{T+h} \mid y_1, \dots, y_T$.
- A prediction interval gives a region within which we expect y_{T+h} to lie with a specified probability.
- Assuming forecast errors are normally distributed, then a 95% PI is

$$\hat{\mathbf{y}}_{T+h|T} \pm 1.96\hat{\sigma}_{h}$$

where $\hat{\sigma}_h$ is the st dev of the *h*-step distribution.

■ When h = 1, $\hat{\sigma}_h$ can be estimated from the residuals.

Drift forecasts with prediction interval:

```
rwf(goog200, level=95, drift=TRUE)
```

```
Point Forecast Lo 95 Hi 95
##
## 201
                532.2 520.0 544.3
## 202
                532.9 515.6 550.1
                533.6 512.4 554.7
## 203
## 204
                534.3 509.8 558.7
## 205
                535.0 507.5 562.4
## 206
                535.7 505.5 565.8
## 207
                536.4 503.7 569.0
                 E27 1 EA2 1 E72 A
```

- Point forecasts are often useless without prediction intervals.
- Prediction intervals require a stochastic model (with random errors, etc).
- Multi-step forecasts for time series require a more sophisticated approach (with PI getting wider as the forecast horizon increases).
- Check residual assumptions before believing them.
- Usually too narrow due to unaccounted uncertainty.

Assume residuals are normal, uncorrelated, sd = $\hat{\sigma}$:

Mean forecasts:
$$\hat{\sigma}_h = \hat{\sigma} \sqrt{1 + 1/T}$$

Naïve forecasts:
$$\hat{\sigma}_h = \hat{\sigma}\sqrt{h}$$

Seasonal naïve forecasts
$$\hat{\sigma}_h = \hat{\sigma}\sqrt{k+1}$$

Drift forecasts:
$$\hat{\sigma}_h = \hat{\sigma} \sqrt{h(1 + h/T)}$$
.

where k is the integer part of (h-1)/m.

Note that when h = 1 and T is large, these all give the same approximate value $\hat{\sigma}$.

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Evaluating prediction intervals

Winkler score

If the 100(1 $-\alpha$)% prediction interval is given by [ℓ , u], and the observed value is y, then the Winkler interval score is

$$(u - \ell) + \frac{2}{\alpha}(\ell - y)1(y < \ell) + \frac{2}{\alpha}(y - u)1(y > u).$$

- penalizes for wide intervals (since $u \ell$ will be large);
- penalizes for non-coverage with observations well outside the interval being penalized more heavily.

Evaluating quantile forecasts

Let q_p be the quantile forecast with probability 1 - p of exceedance.

Pin-ball loss function

$$L(q_p, y) = (1-p)(q_p - y)1(y < q_p) + p(y - q_p)1(y \ge q_p).$$

- average over all target quantiles (e.g., 0.01, 0.02, ..., 0.99) and all forecast horizons.
- Reference: Gneiting and Raftery (JASA, 2007)

Evaluating quantile forecasts